Regularization in Regression

Sibylle Hess

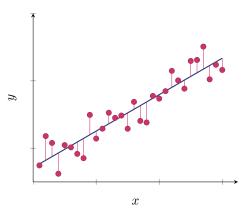




The Problem of Choosing the Right Regression Model



The Regression Optimization Problem



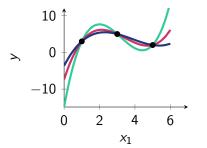
Regression Minimizers

The global minimizers of the regression problem are given by

$$\{\boldsymbol{\beta} \in \mathbb{R}^p \mid \boldsymbol{X}^\top \boldsymbol{X} \boldsymbol{\beta} = \boldsymbol{X}^\top \mathbf{y}\}.$$

If the matrix $X^{T}X$ is invertible, then there is only one minimizer:

$$\boldsymbol{eta} = (X^{ op} X)^{-1} X^{ op} \mathbf{y}$$



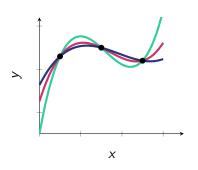
However, there also might be infinitely many local and global minimizers of $RSS(\beta)$. Example: fit the function

$$f(x) = \beta_3 x^3 + \beta_2 x^2 + \beta_1 x^1 + \beta_0$$

to three observations



Toy Example: Regression with p > n



$$f(x) = \beta_3 x^3 + \beta_2 x^2 + \beta_1 x + \beta_0$$

= $\phi(x)^{\top} \beta$,

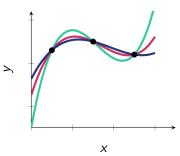
where
$$\phi(x)^{\top} = (1 \ x \ x^2 \ x^3)$$

D	<i>x</i> ₁	У
1	5	2
2	3	5
3	1	3

The design matrix is then given by

$$X = \begin{pmatrix} 1 & 5 & 25 & 125 \\ 1 & 3 & 9 & 27 \\ 1 & 1 & 1 & 1 \end{pmatrix}$$

Toy Example: Regression with p > n



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The global minimizers of the regression problem are given by

$$\{\boldsymbol{\beta} \in \mathbb{R}^p \mid \boldsymbol{X}^\top \boldsymbol{X} \boldsymbol{\beta} = \boldsymbol{X}^\top \mathbf{y}\}.$$

However, the matrix X^TX is in this case not invertible. How do we compute the global minimizers then?



Singular Value Decomposition

Theorem (SVD)

For every matrix $X \in \mathbb{R}^{n \times p}$ there exist orthogonal matrices $U \in \mathbb{R}^{n \times n}$, $V \in \mathbb{R}^{p \times p}$ and $\Sigma \in \mathbb{R}^{n \times p}$ such that

$$X = U\Sigma V^{\top}$$
, where

- $U^{\top}U = UU^{\top} = I_n, V^{\top}V = VV^{\top} = I_p$
- Σ is a rectangular diagonal matrix, $\Sigma_{11} \ge ... \ge \Sigma_{kk}$ where $k = \min\{n, p\}$

The column vectors U_{s} and V_{s} are called left and right singular vectors and the values $\sigma_{i} = \Sigma_{ii}$ are called singular values $(1 \le i \le l)$.

SVD Visualization for p > n

SVD Determines if a Matrix is Invertible

A $(n \times n)$ matrix $A = U\Sigma V^{T}$ is invertible if all singular values are larger than zero. The inverse is given by

$$A^{-1} = V \Sigma^{-1} U^{\mathsf{T}}, \text{ where}$$

$$\Sigma = \begin{pmatrix} \sigma_1 & 0 & \dots & 0 \\ 0 & \sigma_2 & \dots & 0 \\ & & \ddots & \\ 0 & 0 & \dots & \sigma_n \end{pmatrix}, \qquad \Sigma^{-1} = \begin{pmatrix} \frac{1}{\sigma_1} & 0 & \dots & 0 \\ 0 & \frac{1}{\sigma_2} & \dots & 0 \\ & & \ddots & \\ 0 & 0 & \dots & \frac{1}{\sigma_n} \end{pmatrix}$$

Using SVD to Obtain Solutions to the Regression Problem

The global minimizers $oldsymbol{\beta}$ to the linear regression problem with design matrix X are given by

$$\{\boldsymbol{\beta} \in \mathbb{R}^p \mid \boldsymbol{X}^\top \boldsymbol{X} \boldsymbol{\beta} = \boldsymbol{X}^\top \mathbf{y}\}.$$

Let $X = U\Sigma V^{\top}$ be the SVD of X, then we have

$$X^{\top}X\beta = X^{\top}\mathbf{y} \quad \Leftrightarrow \quad \Sigma^{\top}\Sigma V^{\top}\beta = \Sigma^{\top}U^{\top}\mathbf{y}$$

 $\Sigma^{\top}\Sigma$ does not have an inverse if only r < p singular values are nonzero.

Using SVD to Obtain Solutions to the Regression Problem

The global minimizers β to the linear regression problem with design matrix $X = U\Sigma V^{T}$ are given by

$$\{\boldsymbol{\beta} \in \mathbb{R}^p \mid \boldsymbol{\Sigma}^\top \boldsymbol{\Sigma} \boldsymbol{V}^\top \boldsymbol{\beta} = \boldsymbol{\Sigma}^\top \boldsymbol{U}^\top \mathbf{y}\}.$$

If only r < p singular values are nonzero, we employ the pseudoinverse $(\Sigma^{\top}\Sigma)^+$ defined by

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The Set of all Regression Minimizers

If we have r < p nonzero singular values, then we have infinitely many global optimizers

$$oldsymbol{eta} = V A \Sigma^{ op} U^{ op} \mathbf{y}$$

where

$$A = \left(egin{array}{cccc} rac{1}{\sigma_1^2} & \dots & 0 & & \\ dots & \ddots & dots & \mathbf{0} & \\ 0 & \dots & rac{1}{\sigma_{\ell}^2} & & & \\ A_{r+1,1} & \dots & A_{r+1,p} & & dots & dots & \\ dots & & dots & & dots & \\ A_{
ho, \ 1} & \dots & A_{
ho,p} & \end{array}
ight) \in \mathbb{R}^{
ho imes
ho}$$

The Regression Minimizer by the Pseudo Inverse

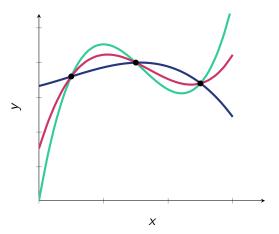
We define the regression solution derived by the pseudo inverse as

$$oldsymbol{eta}_+ = V(oldsymbol{\Sigma}^ op oldsymbol{\Sigma})^+ oldsymbol{\Sigma}^ op U^ op \mathbf{y}$$

where

$$(\mathbf{\Sigma}^{ op}\mathbf{\Sigma})^{+} = \left(egin{array}{cccc} rac{1}{\sigma_{1}^{2}} & \dots & \mathbf{0} & & & & \\ dots & \ddots & dots & \mathbf{0} & & & & \\ 0 & \dots & rac{1}{\sigma_{r}^{2}} & & & & & \\ & \mathbf{0} & & & \mathbf{0} & & & & \end{array}
ight) \in \mathbb{R}^{p imes p}$$

Toy Example: Regression for p > n



So, that's it, sometimes I just have to choose a regression function from infinitely many ones and roll with it?

Well, all regression minimizers are equal, but some minimizers are more equal than others.

Not if d > n!

Example: Gene Expression Analysis

D	Gene 1	Gene 2		Gene 60,000	y: probability of survival
1	0.00	2.75		12.93	0.9
2	0.00	0.00		16.26	0.7
:	:	:	:	:	:
489	0.00	5.38		0.00	0.8

Even if we use a linear function class, we have a design matrix where $p = d = 60,000 \gg 489 = n$.

This introduces the problem of feature selection.

Feature Selection by Sparse Regression Vectors

The regression vector $\boldsymbol{\beta}$ encodes which features are relevant for prediction by nonnegative entries:

$$f(\mathbf{x}) = \mathbf{x}^{\top} \boldsymbol{\beta} = \sum_{i=1}^{p} \beta_i x_i = \sum_{i: \beta_k \neq 0} \beta_i x_i$$

The number of nonnegative entries is given by the L_0 -'norm':

$$\|\beta\|_0 = |\{i \mid \beta_i \neq 0\}|.$$

Be careful: The L_0 -'norm' is not a real norm!

The Sparse Regression Task

Given a data matrix $D \in \mathbb{R}^{n \times d}$, a target vector $\mathbf{y} \in \mathbb{R}^n$, the design matrix $X \in \mathbb{R}^{n \times p}$, where $X_i = \phi(D_i^\top)^\top$ and the integer s.

Find the regression vector β , solving the following objective

$$\min_{\boldsymbol{\beta} \in \mathbb{R}^p} \|\mathbf{y} - \boldsymbol{X}\boldsymbol{\beta}\|^2$$

s.t.
$$\|\beta\|_0 \le s$$
.

Return the predictor function $f: \mathbb{R}^d \to \mathbb{R}$, $f(\mathbf{x}) = \phi(\mathbf{x})^\top \beta$.

That's all nice and stuff, but how are we going to optimize that? The objective is not convex anymore.

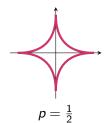
Relax

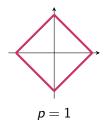


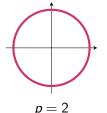
The L_{p} -'norms'

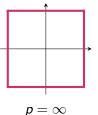
The L_{p} -'norm' is defined for $p \in (0, \infty]$ as follows, and it is a real norm if $p \ge 1$:

$$\|\mathbf{x}\|_{p} = \left(\sum_{k=1}^{d} |x_{k}|^{p}\right)^{1/p}$$









Okay, so we just take an L_p -norm for $p \ge 1$, then the sparse regression problem is convex.

But how do we optimize subject to the constraints?





Solving the dual min f(x) s.t. $c(x) \le 0$

Penalization min $f(x) + \lambda c(x)$

L_p -Norm Penalized Regression

L_p -Constrained Regression

Let $p \in [0, \infty]$, s>0, then the L_p -constrained regression is given as:

$$\min_{\beta} \|\mathbf{y} - X\beta\|^2$$

s.t.
$$\|\boldsymbol{\beta}\|_p \leq s$$

According to the theory of Lagrange multipliers, there exists a parameter $\lambda>0$ such that the objective above is equivalent to

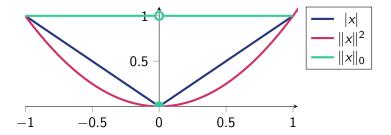
L_p -Penalized Regression

Let $p \in [0, \infty]$, $\lambda > 0$ then the L_p -penalized regression is given as:

$$\min_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2 + \lambda \|\boldsymbol{\beta}\|_{p}$$

Analytical Properties of L_p -norms

norm	continuous	differentiable
$g(\mathbf{x}) = \ \mathbf{x}\ ^2$	✓	✓
$g(\mathbf{x}) = \mathbf{x} $	✓	×
$g(\mathbf{x}) = \ \mathbf{x}\ _0$	×	×



Let us start with a nice and smooth regularization term: the squared L_2 norm.

Given a data matrix $D \in \mathbb{R}^{n \times d}$, a target vector $\mathbf{y} \in \mathbb{R}^n$, the design matrix $X \in \mathbb{R}^{n \times p}$, where $X_i = \phi(D_i^\top)^\top$ and a regularization weight $\lambda > 0$.

Find the regression vector β , solving the following objective

$$\min_{\boldsymbol{\beta} \in \mathbb{R}^p} \|\mathbf{y} - X\boldsymbol{\beta}\|^2 + \lambda \|\boldsymbol{\beta}\|^2.$$

Return the predictor function $f: \mathbb{R}^d \to \mathbb{R}$, $f(\mathbf{x}) = \phi(\mathbf{x})^\top \beta$.

Minimizers of Ridge Regression

Ridge Regression Objective

$$\min_{oldsymbol{eta} \in \mathbb{R}^{
ho}} extit{RSS}_{L_2}(oldsymbol{eta}) = \|\mathbf{y} - Xoldsymbol{eta}\|^2 + \lambda \|oldsymbol{eta}\|^2$$

The solution to ridge regression is given by the stationary points (RSS_{L_2}) is convex as weighted sum of convex functions):

$$\nabla_{\boldsymbol{\beta}} RSS_{L_2}(\boldsymbol{\beta}) = -2X^{\top}(\mathbf{y} - X\boldsymbol{\beta}) + 2\lambda\boldsymbol{\beta} = 0$$

$$\Leftrightarrow (X^{\top}X + \lambda I)\boldsymbol{\beta} = X^{\top}\mathbf{y}$$

Is this now better?

Yes

Minimizers of Ridge Regression are Unique

The matrix $X^{T}X + \lambda I$ is invertible for all $\lambda > 0$! Let $X = U\Sigma V^{T}$ be the singular value decomposition of X, then

$$X^{\top}X + \lambda I = V(\Sigma^{\top}\Sigma + \lambda I)V^{\top}$$

Hence, the uniquely defined global minimizer of Ridge Regression is given by

$$\boldsymbol{\beta}_{L_2} = (\boldsymbol{X}^{\top} \boldsymbol{X} + \lambda \boldsymbol{I})^{-1} \boldsymbol{X}^{\top} \mathbf{y}$$

Ridge Regression and Regression Minimizers

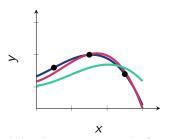
Given the SVD of the design matrix $X = U\Sigma V^{T}$, the ridge regression solution $\beta_{L_{2}}$ with small regularization weight $\lambda > 0$ is similar to one of the global minimizers of regression:

$$\beta_{L_2} = (X^\top X + \lambda I)^{-1} X^\top \mathbf{y} = V(\Sigma^\top \Sigma + \lambda I)^{-1} \Sigma^\top U^\top \mathbf{y}$$

$$\approx V A \Sigma^\top U^\top \mathbf{y}$$

$$\text{if } \lambda > 0 \text{ is small and } A = \begin{pmatrix} \frac{1}{\sigma_1^2} & \dots & 0 \\ \vdots & \ddots & \vdots & \mathbf{0} \\ 0 & \dots & \frac{1}{\sigma_\ell^2} & & \\ & \mathbf{0} & & \ddots & \\ & & & \frac{1}{\lambda} & \end{pmatrix} \in \mathbb{R}^{p \times p}$$

Toy Example: Ridge Regression for p > n



$$f(x_1) = \beta_3 x^3 + \beta_2 x^2 + \beta_1 x^1 + \beta_0$$

$$X = \begin{pmatrix} 1 & 5 & 25 & 125 \\ 1 & 3 & 9 & 27 \\ 1 & 1 & 1 & 1 \end{pmatrix}, \ y = \begin{pmatrix} 2 \\ 5 \\ 3 \end{pmatrix}$$

We obtain as a result for the regression parameters

$$\boldsymbol{\beta}_{+} = \begin{pmatrix} 1.6 \\ 1.2 \\ 0.3 \\ -0.1 \end{pmatrix}, \ \boldsymbol{\beta}_{L_{2}(\lambda=1)} = \begin{pmatrix} 0.8 \\ 0.8 \\ 0.7 \\ -0.2 \end{pmatrix}, \ \boldsymbol{\beta}_{L_{2}(\lambda=20)} = \begin{pmatrix} 0.2 \\ 0.3 \\ 0.4 \\ -0.1 \end{pmatrix}$$

Regularization with the L_1 norm.

Lasso Regression

Given a data matrix $D \in \mathbb{R}^{n \times d}$, a target vector $\mathbf{y} \in \mathbb{R}^n$, the design matrix $X \in \mathbb{R}^{n \times p}$, where $X_i = \phi(D_i^\top)^\top$ and a regularization weight $\lambda > 0$.

Find the regression vector β , solving the following objective

$$\min_{\boldsymbol{\beta} \in \mathbb{R}^p} \lVert \mathbf{y} - \boldsymbol{X}\boldsymbol{\beta} \rVert^2 + \lambda |\boldsymbol{\beta}|.$$

Return the predictor function $f: \mathbb{R}^d \to \mathbb{R}$, $f(\mathbf{x}) = \phi(\mathbf{x})^\top \beta$.

Minimizers of Lasso

Lasso Objective

$$\min_{oldsymbol{eta} \in \mathbb{R}^p} RSS_{L_1}(oldsymbol{eta}) = \|\mathbf{y} - Xoldsymbol{eta}\|^2 + \lambda |oldsymbol{eta}|$$

The L_1 -norm has a subgradient:

$$\frac{\partial |\boldsymbol{\beta}|}{\partial \beta_k} \in \begin{cases} \{1\}, & \text{if } \beta_k > 0\\ \{-1\}, & \text{if } \beta_k < 0\\ [-1, 1], & \text{if } \beta_k = 0 \end{cases}$$

Minimizers of objective functions which have a subgradient satisfy $\mathbf{0} \in \nabla f(\mathbf{x})$ (FONC for subgradients).

Gradients are a special case of subgradients.



Regression Minimizers

Lasso

We could do subgradient descent but then we have to deal with step-sizes and additional difficulties of applying just the subgradient. Luckily, the function is simple enough to derive the minimizers subject to one coordinate, enabling cooordinate descent

Coordinate-Wise Minimizers of Lasso

The minimizer of Lasso subject to the coordinate β_k

$$\beta_k^* = \underset{\beta_k \in \mathbb{R}}{\arg\min} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2 + \lambda |\boldsymbol{\beta}|$$

is given for $c_k = X_{\cdot k}^{\top} \mathbf{y} - \sum_{i \neq k} X_{\cdot k}^{\top} X_{\cdot i} \beta_i$ by

$$eta_k^* = egin{cases} rac{1}{\|X_{\cdot k}\|^2}(c_k - \lambda) & ext{if } c_k > \lambda \ rac{1}{\|X_{\cdot k}\|^2}(c_k + \lambda) & ext{if } c_k < -\lambda \ 0 & ext{if } -\lambda \leq c_k \leq \lambda. \end{cases}$$

FONC for subgradients $\mathbf{0} \in \frac{\partial}{\partial \beta_k} RSS_{L_1}$ yields the solutions to the coordinate-wise minimization problems.



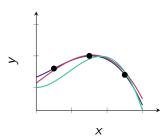
Coordinate Descent for Lasso

return β

end function

8:

1: function Lasso(X, λ, β)
2: while not converged do
3: for $k \in \{1, \dots, p\}$ do
4: $c_k \leftarrow X_{\cdot k}^{\top} \mathbf{y} - \sum_{i \neq k} X_{\cdot k}^{\top} X_{\cdot i} \beta_i$ 5: $\beta_k \leftarrow \begin{cases} \frac{1}{\|X_{\cdot k}\|^2} (c_k - \lambda) & \text{if } c_k > \lambda \\ \frac{1}{\|X_{\cdot k}\|^2} (c_k + \lambda) & \text{if } c_k < -\lambda \\ 0 & \text{if } -\lambda \leq c_k \leq \lambda \end{cases}$ 6: end for
7: end while



$$f(x) = \beta_3 x^3 + \beta_2 x^2 + \beta_1 x^1 + \beta_0$$

$$X = \begin{pmatrix} 1 & 5 & 25 & 125 \\ 1 & 3 & 9 & 27 \\ 1 & 1 & 1 & 1 \end{pmatrix}, \ y = \begin{pmatrix} 2 \\ 5 \\ 3 \end{pmatrix}$$

We obtain as a result for the regression parameters

$$\boldsymbol{\beta}_{+} = \begin{pmatrix} 1.6 \\ 1.2 \\ 0.3 \\ -0.1 \end{pmatrix}, \ \boldsymbol{\beta}_{L_{1}(\lambda=0.1)} = \begin{pmatrix} 0.7 \\ 2.1 \\ 0. \\ -0.07 \end{pmatrix}, \ \boldsymbol{\beta}_{L_{1}(\lambda=1)} = \begin{pmatrix} 0. \\ 0. \\ 1.1 \\ -0.2 \end{pmatrix}$$

Lasso

L_1 vs. L_2 Regularization

The penalized Lasso and Ridge Regression objectives are equivalent to constrained optimization problems.

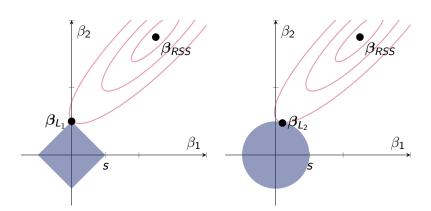
That is, for every $\lambda>0$ there exists a radius s>0 and vice versa, such that the following optimization problems are equivalent:

$$\min \|\mathbf{y} - X\boldsymbol{\beta}\|^2 + \lambda \|\boldsymbol{\beta}\|^2 \qquad \text{s.t. } \boldsymbol{\beta} \in \mathbb{R}^p$$
$$\min \|\mathbf{y} - X\boldsymbol{\beta}\|^2 \qquad \text{s.t. } \|\boldsymbol{\beta}\|^2 \le s^2, \boldsymbol{\beta} \in \mathbb{R}^p$$

Similarly, for every $\lambda>0$ there exists a radius s>0 and vice versa, such that the following optimization problems are equivalent:

$$\begin{split} \min & \|\mathbf{y} - X\boldsymbol{\beta}\|^2 + \lambda |\boldsymbol{\beta}| & \text{s.t. } \boldsymbol{\beta} \in \mathbb{R}^p \\ \min & \|\mathbf{y} - X\boldsymbol{\beta}\|^2 & \text{s.t. } |\boldsymbol{\beta}| \leq s, \boldsymbol{\beta} \in \mathbb{R}^p \end{split}$$

L1-Regularization Tends to Sparser Solutions than L2



Summary L_1 vs. L_2 Regularization

Ridge Regression

$$\min_{oldsymbol{eta} \in \mathbb{R}^p} \mathit{RSS}_{L_2}(oldsymbol{eta}) = \|\mathbf{y} - Xoldsymbol{eta}\|^2 + \lambda \|oldsymbol{eta}\|^2$$

Lasso

$$\min_{oldsymbol{eta} \in \mathbb{R}^p} extit{RSS}_{L_1}(oldsymbol{eta}) = \|\mathbf{y} - Xoldsymbol{eta}\| + \lambda |oldsymbol{eta}|$$

- The solution of Ridge Regression is computable very fast, analyically. The Ridge Regression minimizer is uniquely defined, but usually not sparse.
- **2** Lasso is optimized with coordinate descent, which is a theoretically well-founded optimization procedure. Lasso regression is more likely to return sparse regression vectors β .