

Last week (Generative models for discrete)

- The Dirichlet model (for multiclass data)
- The Naive Bayes classifier

Today

- Generative models for continuous data
- Gaussian - Gaussian, Gaussian - Gamma models
- Bayesian regression
- Model selection & the Bayesian Occam's razor
- Gaussian Processes

## generative models for continuous data

### Example - Regression (Bayesian)

Data:  $(x_i, t_i)$ ,  $i = 1, 2, \dots, n$

Model :  $t = \underbrace{w_0 + w_1 x + w_2 x^2}_{\text{polynomial regression model}} + \underbrace{\varepsilon}_{\text{noise, } N(0, \sigma^2)}$

Q: find the 'best' degree 2  
Polynomial that 'approximates' data

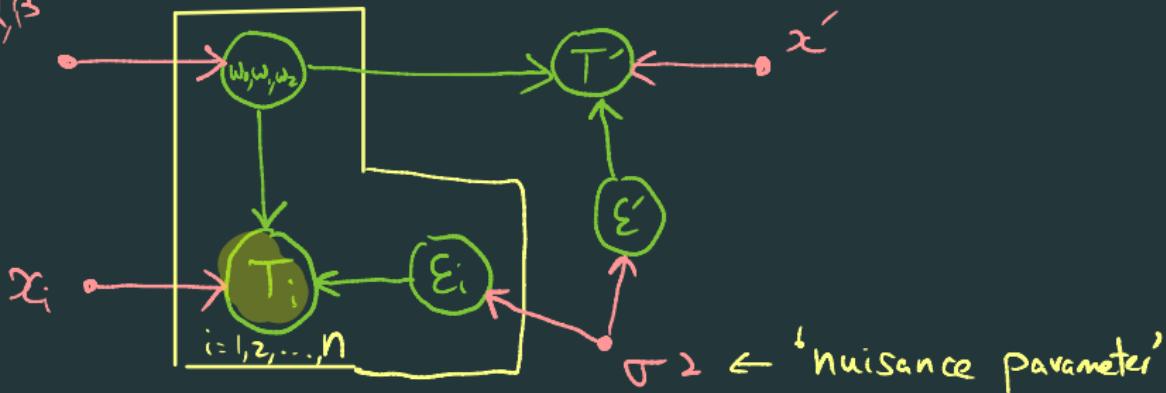
Idea: Assume  $w_0, w_1, w_2$  are random variables, learn from data via Bayesian update



# Bayes Net for regression (for details, see Bishop Ch 8)

hyperparameters prior

$\alpha, \beta$



$$\cdot T' \perp\!\!\!\perp T_i \quad | \quad w$$

- $T'$  not independent of data  $\{T_i\}$  if not conditioned on  $w$

## continuous data and Gaussian priors

- data  $D = \{X_1, X_2, \dots, X_n\} \in \mathbb{R}^n$ ,  $X_i \in \mathbb{R}$  (1-dim data)
- model  $\mathcal{M}$ :  $X_i$  generated i.i.d. from  $\mathcal{N}(\mu, \sigma^2)$  distribution

### Gaussian prior

- $x \in \mathbb{R}$ , parameters:  $\Theta = (\mu, \sigma)$
- pdf:  $\mathbb{P}[X_1 = x_1, X_2 = x_2, \dots, X_n = x_n] \propto \frac{1}{(\sigma^2)^{n/2}} \exp\left(-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}\right)$
- normalizing constant:  $(2\pi)^{-n/2}$

Conditioned on  $\mu, \sigma^2$

$$\frac{1}{(\sigma^2)^{n/2}} \exp\left(-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}\right)$$

Problem: functional form of  $\mu$

and  $\sigma^2$  are different

$$\mu \in \mathbb{R}, \sigma^2 \in \mathbb{R}_+$$

3 options:

1.  $\mu$  unknown,  $\sigma^2$  known most important
2.  $\sigma^2$  unknown,  $\mu$  known
3.  $\mu$  unknown,  $\sigma^2$  unknown

notation: define precision  $\tau = \frac{1}{\sigma^2}$

## case 1: unknown $\mu$

- data  $D = \{X_1, X_2, \dots, X_n\} \in \mathbb{R}^n$  given constant, i.e., hyperparam
- model  $\mathcal{M}$ :  $X_i$  i.i.d. from  $\mathcal{N}(\mu, 1/\tau)$ , with unknown  $\mu$ , known  $\tau = 1/\sigma^2$

normal-normal model

$\uparrow$   
random variable

Gaussian likelihood fn

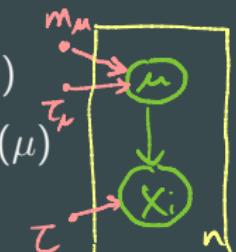
- likelihood:

$$\mathcal{L}(\mu) = p(D|\mu) \propto \tau^{n/2} \exp \left( -\tau \sum_{i=1}^n (x_i - \mu)^2 / 2 \right) = C e^{-c_2(\mu - c_3)^2}$$

- prior parameter:  $\Theta_0 = (m_\mu, 1/\tau_\mu)$  (mean, precision for  $\mu$ )

- Gaussian prior for  $\mu$ :  $\mu \sim \mathcal{N}(m_\mu, \tau_\mu)$ , where  $\tau_\mu = 1/\text{Var}(\mu)$

$$p(\mu|m_\mu, \tau_\mu) \propto \tau_\mu^{1/2} \exp(-\tau_\mu(\mu - m_\mu)^2 / 2)$$



## normal-normal model: posterior

$$P(\mu | D) \propto P(\mu) \cdot L_D(\mu)$$

prior                      likelihood

$\cdot A = \tau_\mu m_\mu + \tau \sum_{i=1}^n \hat{x}_i$ $\cdot \bar{X} = \frac{1}{n} \sum_{i=1}^n \hat{x}_i = \begin{matrix} \text{sample} \\ \text{mean} \end{matrix}$ $(\text{MLE})$ $\cdot m_D = \frac{\tau_\mu m_\mu + n \tau \bar{X}}{\tau_\mu + n \tau}$ $\cdot \tau_D = \tau_\mu + n \tau$	$\propto \exp\left(-\frac{\tau_\mu (\mu - m_\mu)^2}{2}\right) \exp\left(-\frac{\tau \sum_{i=1}^n (x_i - \mu)^2}{2}\right)$ <span style="margin-left: 100px;"><small>want <math>e^{-C(\mu - c_0)^2}</math></small></span> $\propto \exp\left(-\frac{(\tau_\mu + n \tau) \mu^2 + 2A\mu + B}{2}\right)$ $\propto \exp\left(-\frac{\tau_D}{2} (\mu - m_D)^2\right)$
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## normal-normal model: posterior

$$\Rightarrow \text{Posterior} \equiv \mu \sim N(m_D, \tau_D)$$

- posterior mean  $m_D = \frac{\tau_\mu m_\mu + n\tau \bar{x}}{\tau_\mu + n\tau}$   
=  $W_{\text{prior}} m_\mu + (1 - W_{\text{prior}}) \underbrace{\bar{x}}_{\substack{\text{'Shrinkage estimator'} \\ \text{MLE estimate}}}$

- posterior precision  $\tau_D = \tau_\mu + n\tau$   
‘precision on mean adds up under conditioning’

## normal-normal model: posterior predictive distribution

- Prior on  $\mu \sim N(m_\mu, \tau_\mu)$ , posterior given data  $\mu_D \sim N(m_D, \tau_D)$
- $X_{|\mu} \sim N(\mu, \tau) = \underbrace{\mu + \frac{1}{\tau} N(0, 1)}_{\sim N(m_D, \tau_D)}$   
 $= m_D + \frac{1}{\tau_D} N(0, 1) + \frac{1}{\tau} N(0, 1)$   
 $\sim N(m_D, \underbrace{\frac{1}{\tau_D} + \frac{1}{\tau}}_{\sigma_D^2 + \sigma^2})$

'posterior over  $X \sim$  Gaussian, mean = convex comb of  $m_\mu, \bar{x}$   
variance = sum of  $\sigma_D^2, \sigma^2$   
ie - 'uncertainties add up for posterior prediction'

## normal-normal model: posterior predictive distribution

- data  $D = \{X_1, X_2, \dots, X_n\} \in \mathbb{R}^n$
- model  $\mathcal{M}$ :  $X_i$  i.i.d. from  $\mathcal{N}(\mu, \tau)$ , with unknown  $\mu$ , known  $\tau = 1/\sigma^2$
- thus we have  $Z_1, Z_2 \sim \mathcal{N}(0, 1)$  i.i.d.

$$X_i = \mu + \sigma Z_1$$

$$\mu = m_\mu + \sigma_\mu Z_2 = \mu + \sigma Z_1 + \sigma_\mu Z_2$$

$$\Rightarrow \mathbb{E}[X_i] = \mathbb{E}[\mu] = m_D, \text{Var}(X_i) = \sigma^2 + \sigma_D^2$$

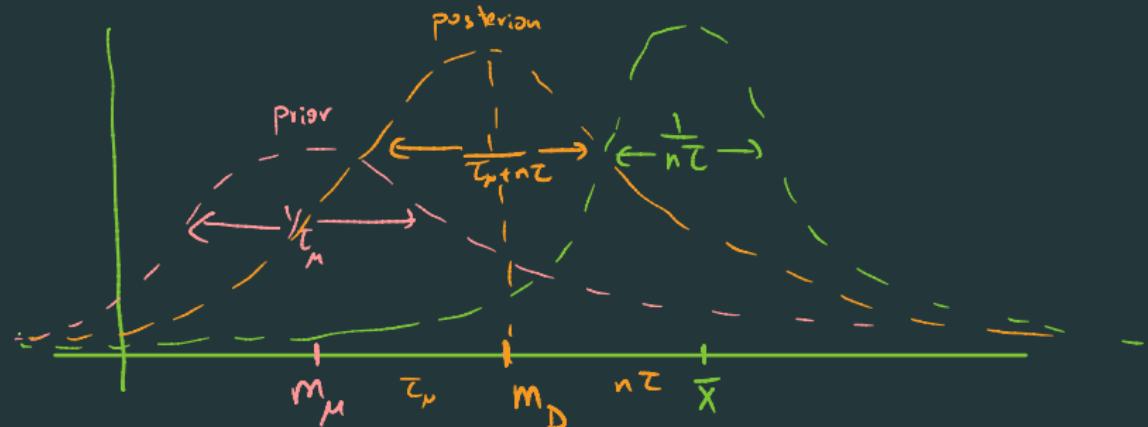
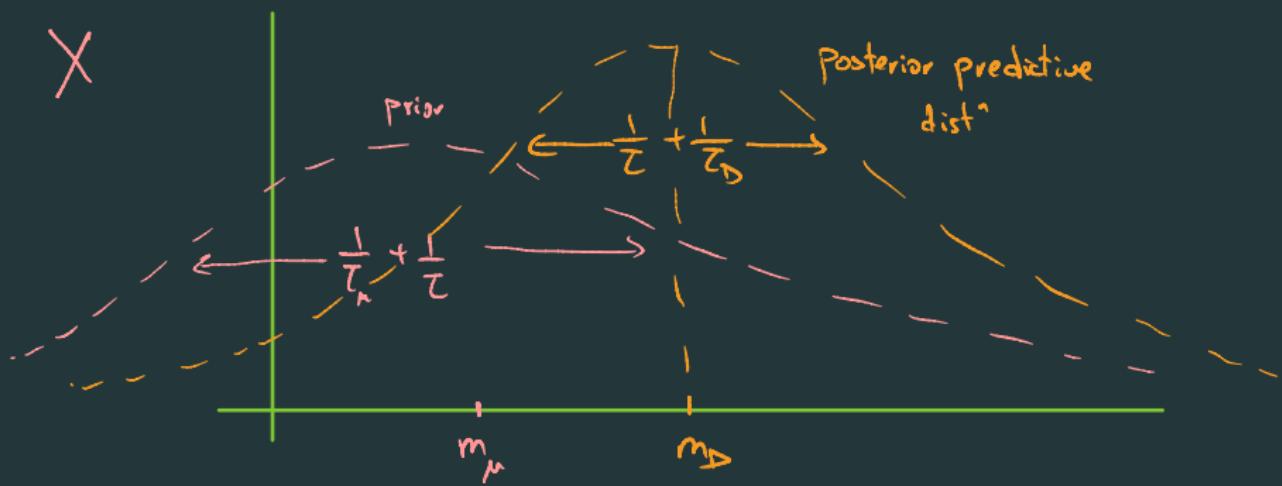
## normal-normal model for unknown $\mu$ (see Bishop Ch 2, sec 3?)

- data  $D = \{X_1, X_2, \dots, X_n\} \in \mathbb{R}^n \rightarrow$  suff stat  $X_{\text{MLE}} = \bar{X} = \sum x_i/n$
- model  $\mathcal{M}$ :  $X_i$  i.i.d. from  $\mathcal{N}(\mu, \tau)$ , with unknown  $\mu$ , known  $\tau = 1/\sigma^2$

### normal-normal model

- likelihood:  $p(D|\mu) \propto \exp(-\tau \sum_{i=1}^n (x_i - \mu)^2 / 2)$
- prior:  $\mu \sim \mathcal{N}(m_\mu, 1/\tau_\mu) \propto \exp(-\tau_\mu(\mu - m_\mu)^2 / 2)$
- posterior: let  $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$ ,  $m_D = \frac{n\tau \cdot \bar{x} + \tau_\mu \cdot m_\mu}{n\tau + \tau_\mu}$  and  $\tau_D = \frac{1}{n\tau + \tau_\mu}$ ,  
 $p(\mu|D) \sim \mathcal{N}(m_D, 1/\tau_D)$   
precisions add for inference
- posterior predictive distribution:

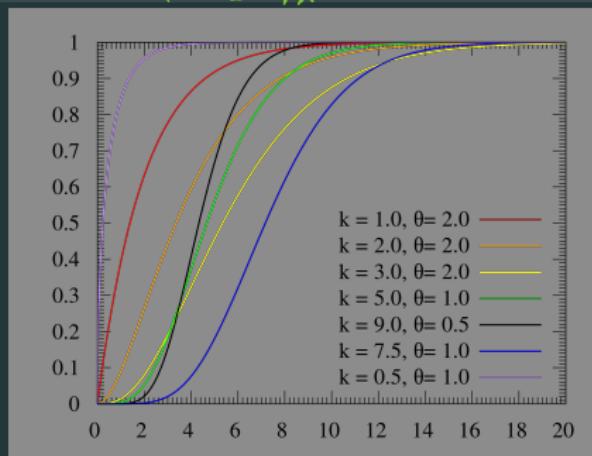
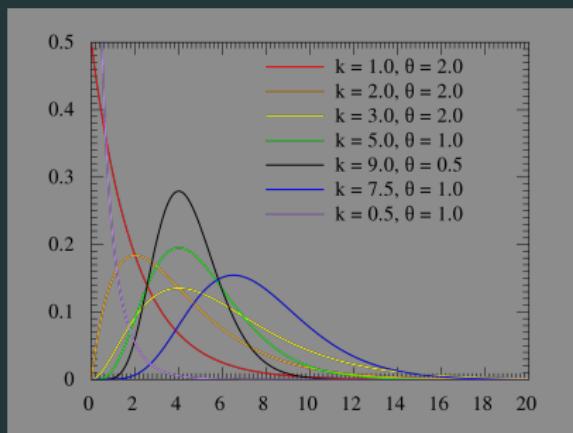
$$p(x|D) \sim \mathcal{N}(m_D, \underbrace{1/\tau + 1/\tau_D}_{\text{variances add for prediction}})$$

$\mu$  $X$ 

# the gamma distribution

## gamma distribution

- $x \in (0, \infty)$ , parameters:  $\Theta = (\alpha, \beta) \in \mathbb{R}^+$  ('shape,rate')
  - pdf of  $\text{Gamma}(\alpha, \beta)$ :  $p(x) \propto x^{\alpha-1} e^{-\beta x}$
  - normalizing constant:  $\frac{1}{Z(\alpha, \beta)} = \frac{\beta^\alpha}{\Gamma(\alpha)}$
- Eg - If  $X_i \sim \text{Exp}(\lambda)$  iid, then  $X_1 + X_2 \sim \text{Gamma}$  with mean  $2/\lambda$



## case 2: unknown $\sigma$

- data  $D = \{X_1, X_2, \dots, X_n\} \in \mathbb{R}^n$
- model  $\mathcal{M}$ :  $X_i$  i.i.d. from  $\mathcal{N}(\mu, 1/\tau)$ , with **unknown**  $\tau = 1/\sigma^2$ , **known**  $\mu$

### normal-gamma model

- likelihood:

$$p(D|\theta) \propto \tau^{n/2} \exp\left(-\tau \sum_{i=1}^n (x_i - \mu)^2 / 2\right)$$

- prior parameters:  $\Theta_0 = (\alpha, \beta)$
- gamma prior for  $\tau$ :  $\tau \sim \text{Gamma}(\alpha, \beta)$

$$p(\tau|\alpha, \beta) \propto \tau^{\alpha-1} e^{-\beta\tau}$$

## normal-gamma model: posterior

$$P(\tau | D) \propto \underbrace{\tau^{\alpha-1} e^{-\beta\tau}}_{\text{prior}} \underbrace{\tau^{n/2 - \frac{1}{2}} e^{-\frac{\sum(x_i - \mu)^2}{2}}}_{\text{Likelihood}}$$

$$\propto \tau^{\alpha + \frac{n}{2} - 1} e^{-\tau \left( \beta + \frac{\sum(x_i - \mu)^2}{2} \right)}$$

$$= \text{Gamma} \left( \alpha + \frac{n}{2}, \beta + \frac{\sum(x_i - \mu)^2}{2} \right)$$

## normal-gamma model: posterior predictive distribution

$$P(x | \alpha, \beta, D) = \underbrace{\int_0^\infty}_{\text{gamma}} p(\tau | \alpha, \beta, D) \cdot \underbrace{p(x | \tau)}_{\text{Gaussian}} d\tau$$

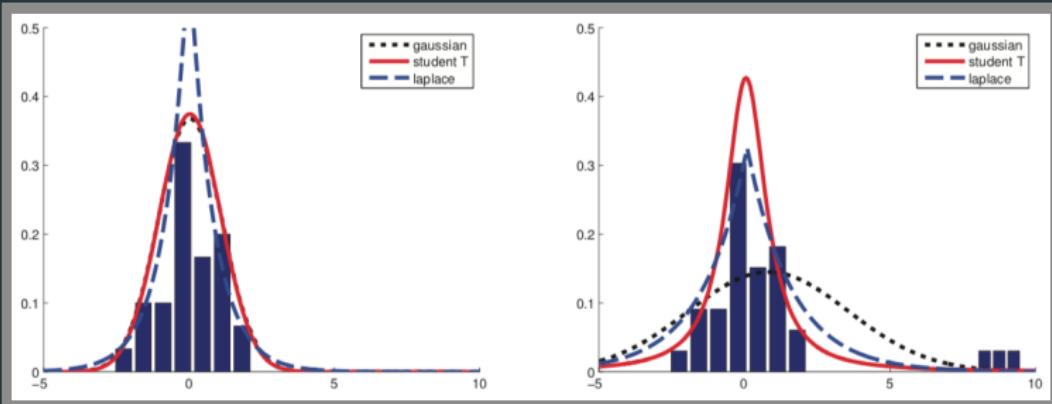
$$= \frac{1}{Z} \underbrace{\frac{1}{\left(1 + \frac{(x-\mu)^2}{2\beta}\right)^{\alpha_D + \frac{1}{2}}}}_{\text{Student's t-dist}}$$

# the Student-t distribution

('naturally robust' distribution)

## Student-t distribution

- $x \in \mathbb{R}$ , parameter:  $\mu \in \mathbb{R}, \nu > 0$  (mean, 'degrees of freedom')
- pdf of student-t( $\mu, \nu$ ):  $p(x) \propto \left(1 + \frac{(x-\mu)^2}{\nu}\right)^{\frac{\nu+1}{2}}$
- normalizing constant:  $\frac{1}{Z(\mu, \nu)} = \frac{\Gamma(\nu+1)/2}{\sqrt{\nu}\pi\Gamma(\nu/2)}$



robustness of student-t to outliers

## normal-gamma model for unknown $\tau$

- data  $D = \{X_1, X_2, \dots, X_n\} \in \mathbb{R}^n$
- model  $\mathcal{M}$ :  $X_i$  i.i.d. from  $\mathcal{N}(\mu, \sigma^2)$ , with unknown  $\tau = 1/\sigma^2$ , known  $\mu$

### normal-gamma model

- likelihood:  $p(D|\theta) \propto \exp(-\tau \sum_{i=1}^n (x_i - \mu)^2 / 2)$
- prior for  $\tau$ :  $\tau \sim \text{gamma}(\alpha, \beta)$
- posterior: let  $\alpha_D = \alpha + \frac{n}{2}$  and  $\beta_D = \beta + \frac{1}{2} \sum_{i=1}^n (x_i - \mu)^2$

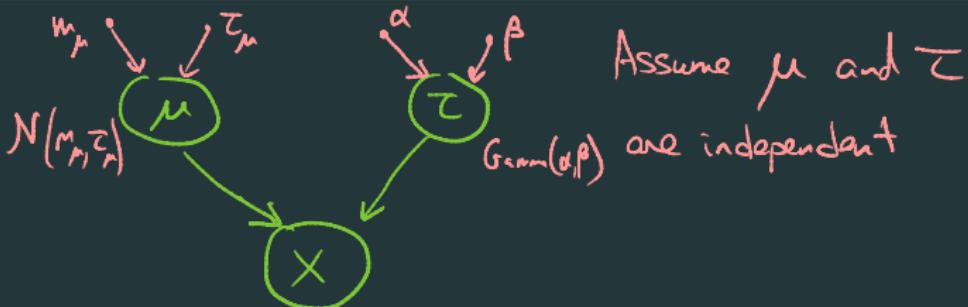
$$p(\tau|D) \sim \text{gamma}(\alpha_D, \beta_D)$$

- posterior predictive distribution:

$$p(x|D) \sim \text{student-t}$$

### case 3: unknown $\mu$ and $\sigma^2$

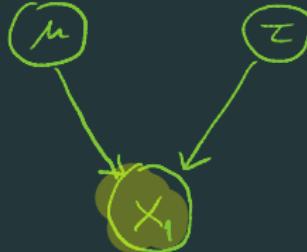
Idea 1



Now given data  $X_1$ , what happens to  $\mu, \tau$ ?

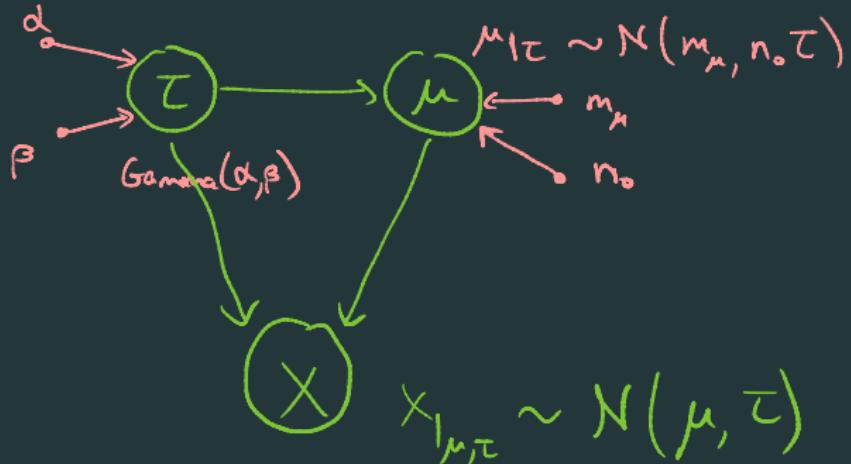
want posterior to be

this is not true!



Conditioned on  $X_1$ ,  $\mu$  and  $\tau$  are no longer indep  
(explaining away)

case 3: unknown  $\mu$  and  $\sigma^2$



Bayesian update for this is known in closed form ('correct' conjugate prior)