

Last 4 classes

- Probabilistic graphical models
- MC MC / Monte Carlo

ORIE 4742 - Info Theory and Bayesian ML

Bayesian Regression $\begin{cases} \text{today - fixed basis functions} \\ \text{next class - Gaussian processes} \end{cases}$

April 23, 2020

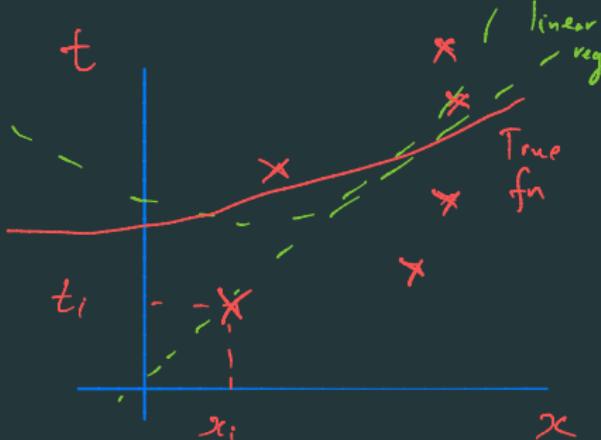
Sid Banerjee, ORIE, Cornell

what is linear regression?

Data - $(x_1, t_1), (x_2, t_2), \dots, (x_N, t_N)$

\uparrow observations \uparrow target

quadratic
reg :



Model

$$y(x) = \sum_{j=0}^{M-1} w_j \phi_j(x)$$

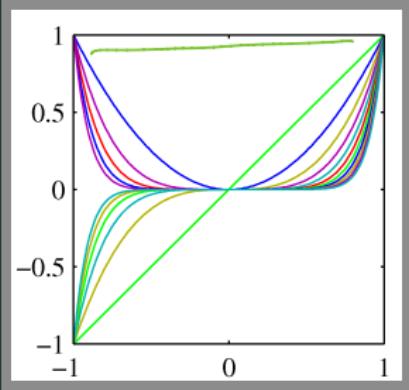
\uparrow regression coefficient \uparrow basis vectors

$$t(x) = y(x) + \varepsilon, \quad \varepsilon \sim N(0, \frac{1}{\beta}) \quad \text{- Noise}$$

$\underbrace{\qquad\qquad\qquad}_{\text{frequentist view of regression}}$ $\underbrace{\qquad\qquad\qquad}_{\text{noise precision}}$

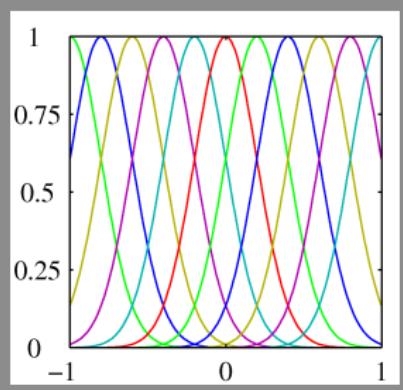
$$\text{- Assume } \phi_0(x) = 1 \quad (w_0 \equiv \text{constant, 'bias'})$$

basis functions



Polynomial basis fns

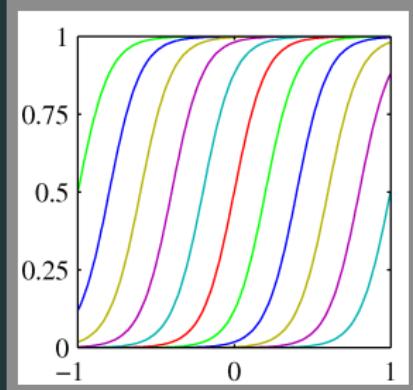
$$\phi_j(x) = x^j$$



Gaussian basis fn

$$\phi_j(x) = e^{-\frac{(x-\mu_j)^2}{s_j^2}}$$

↑
location parameter ↑
scale parameter



Sigmoidal basis fn

$$\phi(x) = \frac{1}{1 + e^{-\frac{(x-\mu)}{s}}}$$

• Fourier basis $\equiv \phi_j(x) = \sin(\omega_j x + \mu_j)$

• Wavelet basis

regression: the frequentist view $y(x) = w_0 + w_1 x$

$$(M) \quad t(x) = \sum_{j=0}^{M-1} w_j \phi_j(x) + \varepsilon, \quad \varepsilon \sim N(0, \frac{1}{\beta})$$

- design matrix $\bar{\Phi} = \begin{pmatrix} \phi_0(x_1) & \phi_1(x_1) & \dots & \phi_{M-1}(x_1) \\ \phi_0(x_2) & \phi_1(x_2) & \dots & \phi_{M-1}(x_2) \\ \vdots & & & \\ \phi_0(x_N) & \phi_1(x_N) & \dots & \phi_{M-1}(x_N) \end{pmatrix}_{N \times M \text{ matrix}}$
- $D \equiv \bar{\Phi}, \quad t = (t_1, t_2, \dots, t_N)^T \quad \underbrace{\text{Sufficient statistic of the data}}_{\text{N} \times 1 \text{ vector}}$

- likelihood $P(D | M) \propto \exp\left(-\sum_{i=1}^N \frac{\beta(t_i - w^T \phi(x))}{2}\right)$
- maximum likelihood $w_{ML} = \underbrace{\bar{\Phi}^+}_\text{Pseudo inverse} t, \quad \bar{\Phi}^+ = (\bar{\Phi}^T \bar{\Phi})^{-1} \bar{\Phi}^T$

Eg - linear regression (frequentist)

$$t = w_0 + w_1 x + \sum \text{noise } N(0, \frac{1}{\beta})$$

$\uparrow \quad \uparrow$
observed data unknown params

$$(t = y(x) + \epsilon, y(x) = w_0 + w_1 x)$$

$$\Phi = \begin{pmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_N \end{pmatrix}, \quad t = \begin{pmatrix} t_1 \\ t_2 \\ \vdots \\ t_N \end{pmatrix}, \quad w = \begin{pmatrix} w_0 \\ w_1 \end{pmatrix}$$

$$W_{ML} = \underbrace{(\Phi^T \Phi)^{-1} \Phi^T}_A(t_1, t_2, \dots, t_N) \cdot t$$

Alternate
- choose w_0, w_1 to
minimize $\sum_{i=1}^N (t_i - w_0 - w_1 x_i)^2$
i.e., LS estimate

- Output - $y(x) = w_0^{ML} + w_1^{ML} x$

Bayesian linear regression

$$P(t|w) \sim \mathcal{N}(w^T \phi(x), \beta)$$

Model - $t_i = \sum_{j=0}^{M-1} w_j \phi_j(x_i) + \varepsilon_i$

\uparrow \uparrow
 'Unknown' = random variables

- $\varepsilon_i \sim \mathcal{N}(0, \beta)$ \equiv iid for each (x_i, t_i)
 - $w = \begin{pmatrix} w_0 \\ w_1 \\ \vdots \\ w_{M-1} \end{pmatrix} \sim \mathcal{N}(0, T^{-1})$ (prior)
- \uparrow \uparrow
 $\begin{pmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{pmatrix}$ 'Precision' matrix
 Eg - $\alpha^{-1} I$

i.e, $w_j \sim \mathcal{N}(0, \alpha)$, iid $\forall j$

- α, β = model hyper parameters (fixed)

normal-normal model for unknown μ

- data $D = \{X_1, X_2, \dots, X_n\} \in \mathbb{R}^n$
- model \mathcal{M} : X_i i.i.d. from $\mathcal{N}(\mu, \tau)$, with unknown μ , known $\tau = 1/\sigma^2$

normal-normal model

- likelihood: $p(D|\mu) \propto \exp(-\tau \sum_{i=1}^n (x_i - \mu)^2 / 2)$
- prior: $\mu \sim \mathcal{N}(m_\mu, 1/\tau_\mu) \propto \exp(-\tau_\mu(\mu - m_\mu)^2 / 2)$ (m_μ, τ_μ - hyperparam)
- posterior: let $\bar{x} = \underbrace{\frac{1}{n} \sum_{i=1}^n x_i}_{\text{Empirical mean}}$, $m_D = \frac{n\tau \bar{x} + \tau_\mu \cdot m_\mu}{n\tau + \tau_\mu}$ and $\tau_D = n\tau + \tau_\mu$
 $p(\mu|D) \sim \mathcal{N}(m_D, 1/\tau_D)$ $\stackrel{\text{Shrinkage estimator}}{\sim} \bar{x} + (1-\bar{x})m_\mu$
(MLE for μ)
- posterior predictive distribution:
 $p(x|D) \sim \mathcal{N}(m_D, 1/\tau + 1/\tau_D)$
 $\stackrel{\text{noise added to } x \text{ by model}}{\sim}$
 $\stackrel{\text{'noise' in parameter } \mu}{\sim}$

Bayesian linear regression

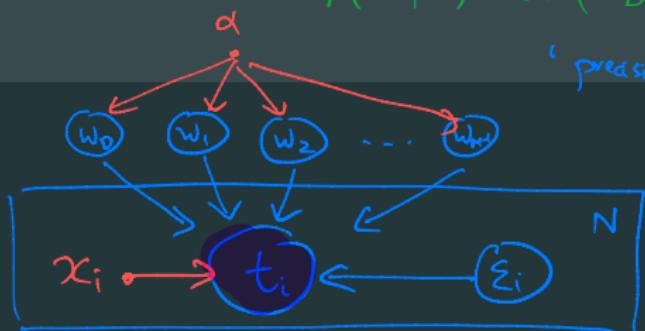
- data $D = \{(t_1, x_1), (t_2, x_2), \dots, (t_N, X_N)\} \in \mathbb{R}^n$
- model \mathcal{M} : $t_i = \sum_{j=0}^{M-1} \underbrace{W_j \phi(x_i)}_{W^\top \phi(x_i)} + \epsilon_i$, where $\epsilon_i \sim \mathcal{N}(0, \beta^{-1})$

Bayesian linear regression model

- likelihood: $p(D|W) \propto \exp\left(-\beta \sum_{i=1}^N (x_i - W^\top \phi(x_i))^2 / 2\right)$
- prior: $W \sim \mathcal{N}(0, \alpha^{-1} I)$ (*i.e.*, $W_j \sim \mathcal{N}(0, \gamma_\alpha)$, *iid*)
 $m_D = T_D^{-1} \beta \sum_{i=1}^N \underbrace{\phi(x_i) t_i}_{M \times 1}$
- posterior: let $m_D = T_D^{-1} \beta \underbrace{\phi^\top t}_{{M \times 1}}$ and $T_D = \beta \Phi^\top \Phi + \alpha I$

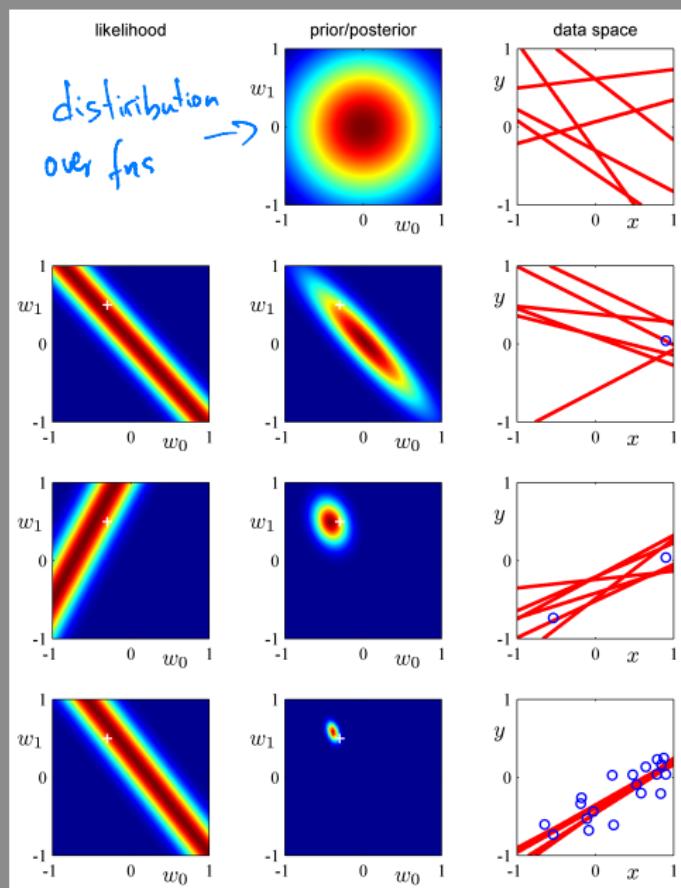
$$p(W|D) \sim \mathcal{N}(m_D, T_D^{-1})$$

'precision' = inverse covariance matrix



Note
 $\{w_i\}_{i=0}^M$ initially indep,
but dependent given t.

Bayesian linear regression: example (from Bishop Ch 3)



$$\text{model} - t_i = w_0 + w_1 x_i + \varepsilon_i$$

$$\begin{pmatrix} w_0 \\ w_1 \end{pmatrix} \sim N(0, \alpha^{-1} I), \quad \varepsilon_i \sim N(0, \beta^{-1})$$

$$\therefore y(x) = -0.3 + 0.1x, \quad t_i = y(x) + \varepsilon_i$$

As N increases

$$\overline{T_D} \rightarrow 0$$

$$M_D \rightarrow \text{true } \begin{pmatrix} w_0 \\ w_1 \end{pmatrix}$$

ground truth: $f(x) = 0.1x - 0.3$

Bayesian linear regression

- data $D = \{(t_1, x_1), (t_2, x_2), \dots, (t_N, X_N)\} \in \mathbb{R}^n$
- model \mathcal{M} : $t_i = \sum_{j=0}^{M-1} W_j \phi(x_i) + \epsilon_i$, where $\epsilon_i \sim \mathcal{N}(0, \beta^{-1})$

Bayesian linear regression model

- likelihood: $p(D|W) \propto \exp\left(-\beta \sum_{i=1}^N (x_i - W^\top \phi(x_i))^2 / 2\right)$
- prior: $W \sim \mathcal{N}(0, \alpha^{-1} I)$
- posterior: let $m_D = T_D^{-1} \beta \Phi^\top t$ and $T_D = \beta \Phi^\top \Phi + \alpha I$

$$p(W|D) \sim \mathcal{N}(m_D, T_D^{-1})$$

- posterior predictive distribution: $(i.e., p(t|x, D))$

$$p(t|D) \sim \mathcal{N}(m_D^\top \phi(x), \beta^{-1} + \phi(x)^\top T_D^{-1} \phi(x))$$

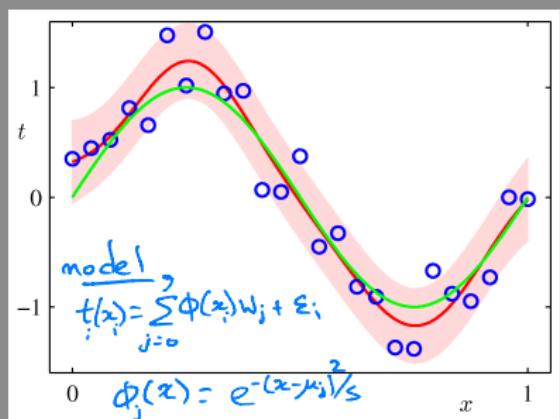
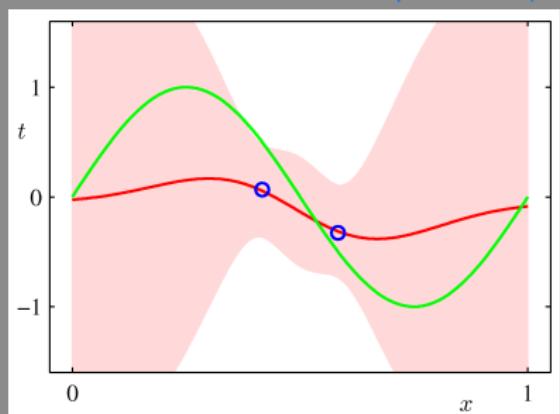
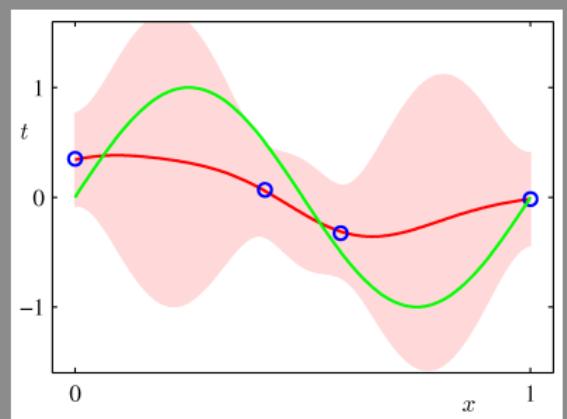
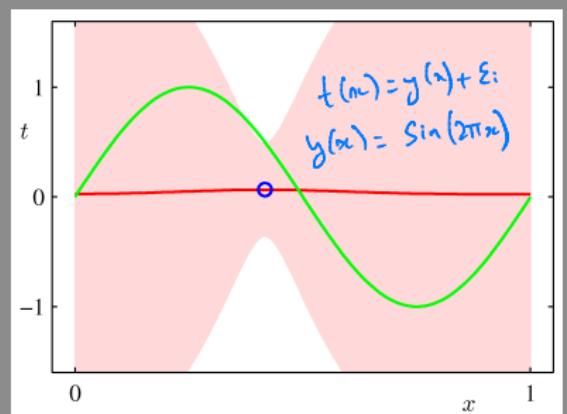
Variance as fn of x

i.e. - $W = m_D + Z$, $Z \sim \mathcal{N}(0, T_D)$, $t = W^\top \phi(x) + \epsilon_i$

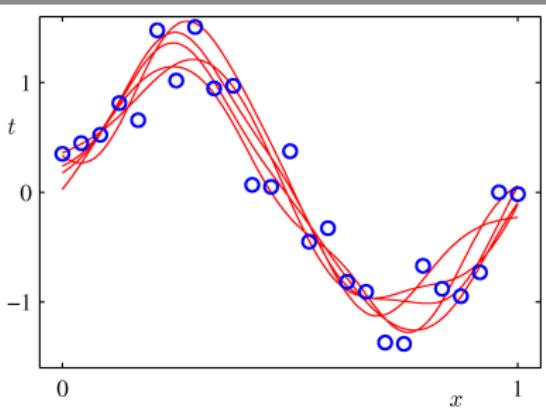
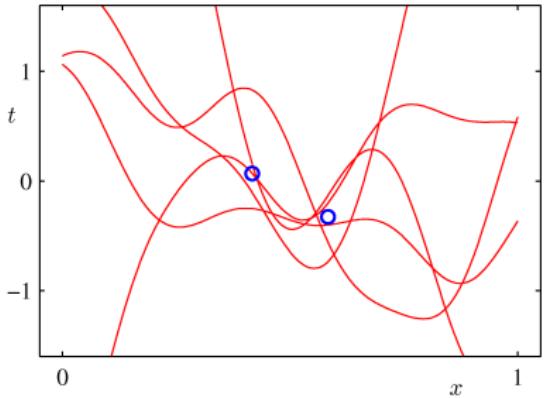
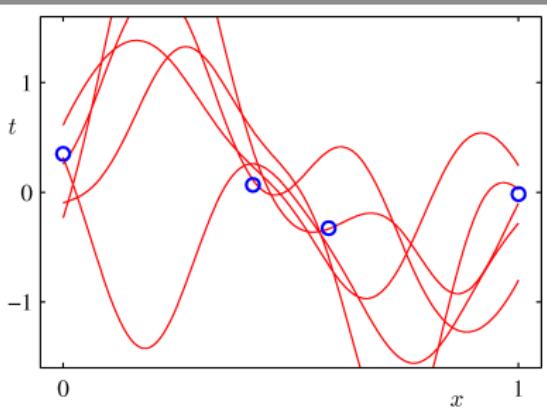
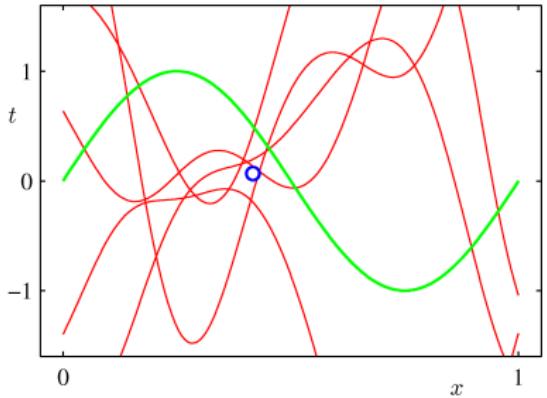
$$\Rightarrow t = m_D^\top \phi(x) + \underbrace{Z^\top \phi(x) + \epsilon_i}_{\sim \mathcal{N}(0, \frac{1}{\beta} + \phi(x)^\top T_D^{-1} \phi(x))}$$

Bayesian linear regression: posterior prediction

Bishop Ch 3
 - ground truth - Sin(2πx)
 - basis fn - Gaussian, M=10



Bayesian linear regression: posterior sampling



the 'equivalent' kernel (distance fn defined by data)

- given $D = \{(t_i, x_i)\}$, posterior mean $\hat{t}(x) = \sum_{i=1}^N t_i K(x, x_i)$
- data $D = \{(t_1, x_1), (t_2, x_2), \dots, (t_N, x_N)\} \in \mathbb{R}^n$
- model $M: t_i = \sum_{j=0}^{M-1} W_j \phi(x_i) + \epsilon_i$, where $\epsilon_i \sim \mathcal{N}(0, \beta^{-1})$
- prior: $W \sim \mathcal{N}(0, \alpha^{-1} I)$
- posterior: let $m_D = T_D^{-1} \beta \Phi^\top t$ and $T_D = \beta \Phi^\top \Phi + \alpha I$, then

$$t(x|D) = m_D^\top \phi(x) + \epsilon_D$$

noise in model

$$\text{where } \epsilon_D \sim \mathcal{N}(0, \beta^{-1} + \Phi^\top T_D^{-1} \Phi)$$

noise in params W

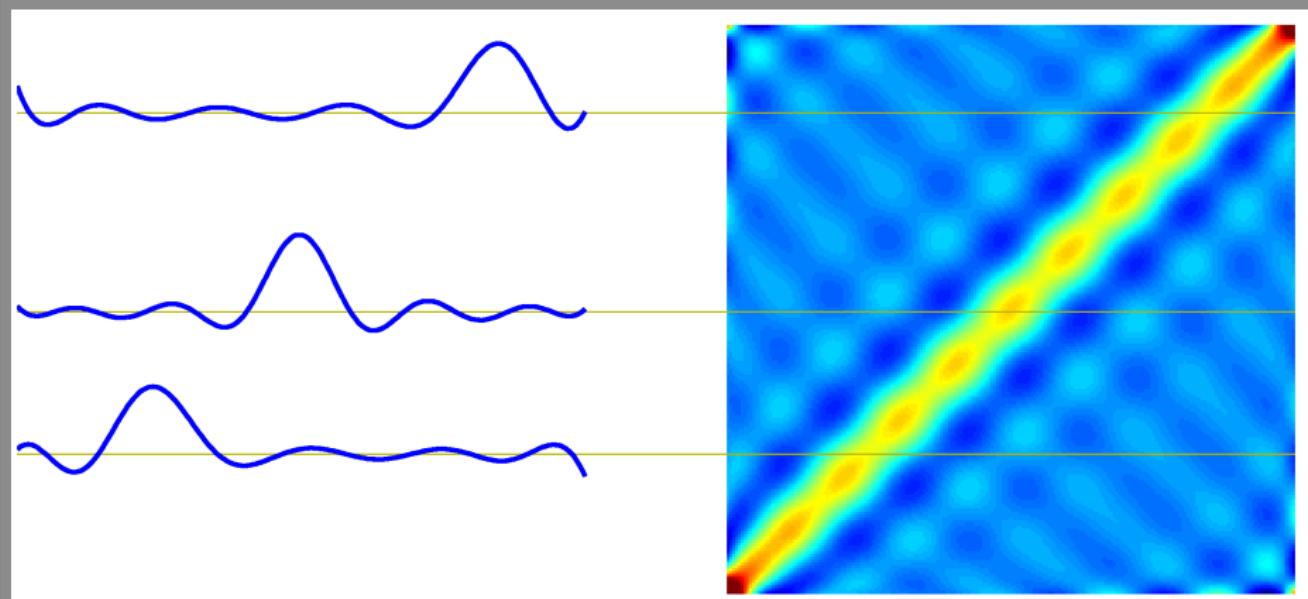
alternately, $y(x|D) = \sum_{n=1}^N k(x, x_n) t_n$, where $k(x, y) = \beta \phi(x)^\top S_D \phi(y)$

$$\begin{aligned}
 y(x|D) &= m_D^\top \phi(x) = \beta (T_D^{-1} \Phi^\top t)^\top \begin{pmatrix} \phi_0(x) \\ \phi_1(x) \\ \vdots \\ \phi_{M-1}(x) \end{pmatrix} \\
 &= t^\top \left(\beta \Phi_{\substack{\in \\ (\phi(x))_{i,j}}}^\top T_D^{-1} \right) \phi(x)
 \end{aligned}$$

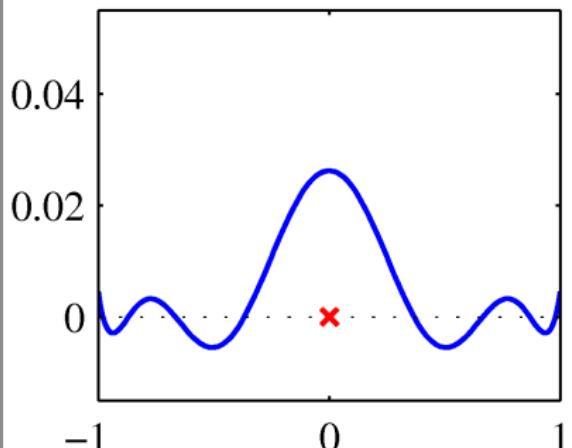
\$AB^\top = B^\top A^\top\$

the equivalent kernel: example

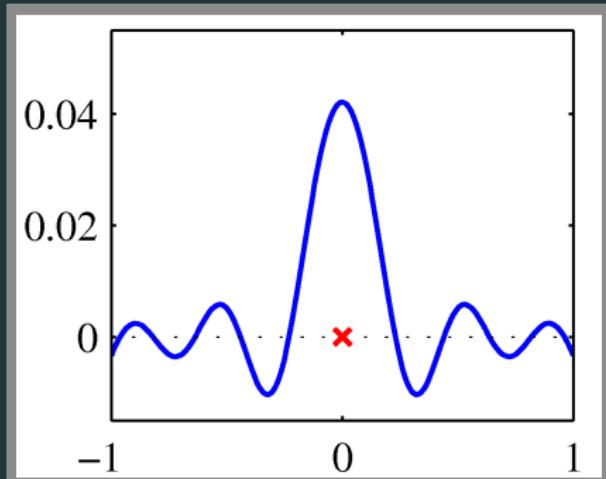
(sinusoid ground truth, Gaussian basis fn)



equivalent kernels



Polynomial kernel



Sigmoidal kernel

