

- Beta - Bernoulli Model (for binary data)
- Dirichlet Model (for multi-class data)

• Naive Bayes classifier

Problem - email classification

$\left. \begin{matrix} \text{regular} \\ \text{important} \\ \text{spam} \end{matrix} \right\}$ classes

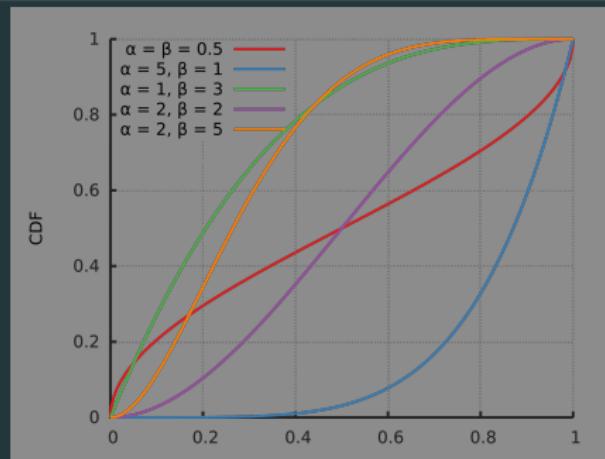
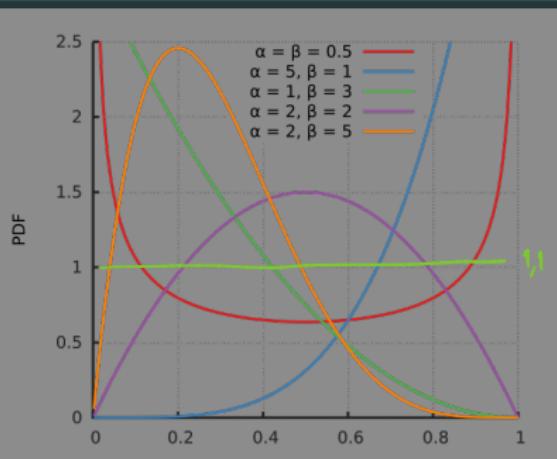
- Input to Naive Bayes - 1) prior over classes

<p>'Hi lets meet after class'</p> <p>\Rightarrow</p> <table border="1"> <tr> <td>Hi</td> <td>1</td> </tr> <tr> <td>meet</td> <td>1</td> </tr> <tr> <td>lets</td> <td>1</td> </tr> <tr> <td>.</td> <td></td> </tr> </table>	Hi	1	meet	1	lets	1	.		<p>2) dataset of labelled examples $(D_1, c_1), (D_2, c_2), \dots, (D_n, c_n)$</p> <p style="text-align: center;">↑ email \equiv 'bag of words'</p> <p>3) new email D</p> <p>\Rightarrow output = distn over $C(D)$</p>
Hi	1								
meet	1								
lets	1								
.									

the Beta distribution

Beta distribution

- $x \in [0, 1]$, parameters: $\Theta = (\alpha, \beta) \in \mathbb{R}^+$ ('# ones'+1, '# zeros'+1)
- pdf: $p(x) \propto x^{\alpha-1}(1-x)^{\beta-1}$ \leftarrow same form as the Bernoulli
- normalizing constant: $\frac{1}{B(\alpha, \beta)} = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} = \frac{1}{\int_0^1 x^{\alpha-1}(1-x)^{\beta-1} dx}$ likelihood!



Beta-Bernoulli prior and updates

- data $D = \{X_1, X_2, \dots, X_n\} \in \{0, 1\}^n$, contains N_1 ones and N_0 zeros
- model \mathcal{M} : X_i are generated i.i.d. from a $Ber(\theta)$ distribution

Beta-Bernoulli model

- prior parameters: $\Theta_0 = (\alpha, \beta) \in \mathbb{R}^+$ (hyperparameters)
- Beta-Bernoulli prior: $Beta(\alpha, \beta) \sim p(\theta) \propto \theta^{\alpha-1}(1-\theta)^{\beta-1}$
- likelihood: $p(D|\theta) = \theta^{N_1}(1-\theta)^{N_0}$

then via Bayesian update we get $p(\theta|D) = \mathcal{L}(\theta|D) p(\theta)$

- posterior:

$$p(\theta|D) \propto \theta^{\alpha-1}(1-\theta)^{\beta-1} \theta^{N_1}(1-\theta)^{N_0} \sim Beta(\alpha + N_1, \beta + N_0)$$

the Beta distribution: getting familiar

Beta(α, β) distribution

$$p(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

properties of $\Gamma(\alpha)$

$$\frac{1}{\Gamma(\alpha)} , \quad B(\alpha, \beta) = \int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx$$

$$\Gamma(\alpha) = \int_0^\infty e^{-y} y^{\alpha-1} dy \quad \left| \begin{array}{l} \text{more importantly} \\ \boxed{\Gamma(\alpha+1) = \alpha \Gamma(\alpha)} \quad \boxed{\Gamma(1)=1} \end{array} \right.$$

- If α is an integer $\Rightarrow \Gamma(\alpha) = (\alpha-1)!$

the Beta distribution: mean and mode

Beta(α, β) distribution

$$p(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

$$\begin{aligned}\cdot \mathbb{E}[x] &= \int_0^1 x \cdot \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} dx \\ &= \frac{1}{B(\alpha, \beta)} \int_0^1 x^\alpha (1-x)^{\beta-1} dx = \frac{B(\alpha+1, \beta)}{B(\alpha, \beta)} = \frac{\Gamma(\alpha+1)}{\Gamma(\alpha)\Gamma(\beta)} \cdot \cancel{\frac{\Gamma(\alpha+1)\Gamma(\beta)}{\Gamma(\alpha+\beta+1)}} \\ &= \frac{\alpha}{\alpha+\beta}\end{aligned}$$

$$\Rightarrow \boxed{\text{Mean} (\text{Beta}(\alpha, \beta)) = \frac{\alpha}{\alpha+\beta}}$$

$$\text{mode} \equiv \underset{x \in [0,1]}{\arg \max} \frac{x^{\alpha-1} (1-x)^{\beta-1}}{B(\alpha, \beta)}$$

$$\begin{aligned} \frac{d}{dx} (x^{\alpha-1} (1-x)^{\beta-1}) &= (\alpha-1) x^{\alpha-2} (1-x)^{\beta-1} - (\beta-1) x^{\alpha-1} (1-x)^{\beta-2} \\ &= x^{\alpha-2} (1-x)^{\beta-2} ((\alpha-1)x - (\beta-1)(1-x)) = 0 \end{aligned}$$

$$\Rightarrow x^* = 0, 1 \text{ or } \frac{\alpha-1}{\alpha+\beta-2}$$

Mode of Beta(α, β) dist' is $\frac{\alpha-1}{\alpha+\beta-2}$

Beta-Bernoulli model: what should we report?

- data $D = \{X_1, X_2, \dots, X_n\} \in \{0, 1\}^n$, contains N_1 ones and N_0 zeros
- model \mathcal{M} : X_i are generated i.i.d. from a $Ber(\theta)$ distribution
- prior: $p(\theta) \sim Beta(\underline{\alpha}, \underline{\beta})$ posterior: $p(\theta|D) \sim Beta(\underline{\alpha + N_1}, \underline{\beta + N_0})$
 $\underline{\alpha}$ prior hyperparams $\underline{\alpha + N_1}$ posterior hyperparams

Correct answer \equiv report $\begin{pmatrix} \text{Prior} + \\ \text{Posterior} \end{pmatrix} \equiv Beta(\underline{\alpha + N_1}, \underline{\beta + N_0})$

- In practice - 'people' want a single number
For this - [ask what the 'loss function' is
decision theory - [return the Θ which minimizes expected loss over the posterior]

decision theory

Idea - Choose 'action' \rightarrow minimize expected loss over $p(\theta|D)$

Eg - Let θ be the 'true answer', $\hat{\theta}$ be the report

1) $L(\theta, \hat{\theta}) = \mathbb{I}\{\theta \neq \hat{\theta}\}$ (L_0 loss) : $\hat{\theta} = \text{mode of } p(\theta|D)$

2) $L(\theta, \hat{\theta}) = (\theta - \hat{\theta})^2$ (L_2 loss) : $\hat{\theta} = \text{mean of } p(\theta|D)$

3) $L(\theta, \hat{\theta}) = |\theta - \hat{\theta}|$ (L_1 loss) : $\hat{\theta} = \text{median of } p(\theta|D)$

General rule - $\hat{\theta} = \underset{\theta}{\operatorname{arg\min}} \mathbb{E}_{P(\theta|D)} [L(\theta, \hat{\theta})]$

decision theory in a nutshell

Bayesian decision theory in learning

given prior F on θ , choose 'action' $\hat{\theta}$ to minimize loss function $\mathbb{E}_F[L(\theta, \hat{\theta})]$

examples

- L_0 loss: $L(\theta, \hat{\theta}) = \mathbb{1}_{\{\theta \neq \hat{\theta}\}} \Rightarrow \hat{\theta}_{L_0} = \text{mode of } F$
- L_1 loss: $L(\theta, \hat{\theta}) = \|\theta - \hat{\theta}\|_1 \Rightarrow \hat{\theta}_{L_1} = \text{median of } \theta \text{ under } F$
- L_2 loss: $L(\theta, \hat{\theta}) = \|\theta - \hat{\theta}\|_2 \Rightarrow \hat{\theta}_{L_2} = \mathbb{E}_F[\theta]$

in general 'decision-making'

given prior F on X , choose 'action' $a \in \mathcal{A}$ to minimize loss, i.e.

$$a^* = \arg \min_{a \in \mathcal{A}} \mathbb{E}_{X \sim F}[L(a, X)]$$

Beta-Bernoulli model: posterior mean

- data $D = \{X_1, X_2, \dots, X_n\} \in \{0, 1\}^n$, contains N_1 ones and N_0 zeros
- model \mathcal{M} : X_i are generated i.i.d. from a $Ber(\theta)$ distribution
- prior: $p(\theta) \sim Beta(\alpha, \beta)$ posterior: $p(\theta|D) \sim Beta(\alpha + N_1, \beta + N_0)$

posterior mean: $\mathbb{E}[\theta|\alpha, \beta, N_0, N_1] = \mathbb{E}[Beta(\alpha+N_1, \beta+N_0)] = \frac{\alpha + N_1}{\alpha + \beta + N_0 + N_1}$

Defn - $n = N_1 + N_0$
 $m = \alpha + \beta$
 $m = \text{'# of prior samples'}$

$\frac{\alpha}{m} = \text{prior mean}$
 $\frac{N_1}{n} = \text{Data mean (also, MLE)}$
 $w = \frac{m}{m+n} \text{'strength of prior'}$

$= \frac{\alpha + N_1}{m+n} = \left(\frac{\alpha}{m}\right) \cdot \left(\frac{m}{m+n}\right) + \left(\frac{N_1}{n}\right) \left(\frac{n}{m+n}\right)$

$= \underbrace{\left(\text{Prior mean}\right) \cdot w}_{\text{regularizer}} + \underbrace{\left(\text{MLE}\right) \left(1-w\right)}_{\text{shrinkage}}$

Note: For fixed m , as $n \rightarrow \infty$, then
Posterior mean \rightarrow MLE

Beta-Bernoulli model: posterior mode (**MAP estimation**)

max a posteriori

- data $D = \{X_1, X_2, \dots, X_n\} \in \{0, 1\}^n$, contains N_1 ones and N_0 zeros
- model \mathcal{M} : X_i are generated i.i.d. from a $Ber(\theta)$ distribution
- prior: $p(\theta) \sim Beta(\alpha, \beta)$ posterior: $p(\theta|D) \sim Beta(\alpha + N_1, \beta + N_0)$

posterior mode: $\max_{\theta \in [0,1]} p(\theta|\alpha, \beta, N_0, N_1) = \frac{\alpha + N_1 - 1}{\alpha + \beta + N_1 + N_0 - 2} = \frac{\alpha + N_1}{m+n-2}$

- If $\alpha = \beta = 1$ (uniform prior)

$$\Rightarrow \theta_{MAP} = \frac{N_1}{n} = \theta_{MLE}$$

i.e.: MLE is the 'right' answer if Lo loss
+ uniform prior

Beta-Bernoulli model: posterior prediction (**marginalization**)

- data $D = \{X_1, X_2, \dots, X_n\} \in \{0, 1\}^n$, contains N_1 ones and N_0 zeros
- model \mathcal{M} : X_i are generated i.i.d. from a $Ber(\theta)$ distribution
- prior: $p(\theta) \sim Beta(\alpha, \beta)$ posterior: $p(\theta|D) \sim Beta(\alpha + N_1, \beta + N_0)$

posterior prediction: $\mathbb{P}[X = 1|D] = \int_0^1 p(\theta|D) \cdot \theta \cdot d\theta$

$$= \frac{\alpha + N_1}{m+n}$$

- If $\alpha = \beta = 1$

$$\boxed{\mathbb{P}[X=1|D] = \frac{N_1+1}{N_0+N_1+2}}$$

Laplace Estimator

$$\text{If } N_1 = 0, \quad \mathbb{P}[X=1|D] = \frac{1}{n+2}$$

the black swan

- In some datasets, a rare event is never seen
 - If n samples, $N_1=0, N_0=n$
then Laplace estimator $P[X=1|D] = \frac{1}{n+2}$
 $P[X=0|D] = \frac{n+1}{n+2}$
prob fo. the
'unseen'
- One important application = crypto
(Good-Turing estimator)

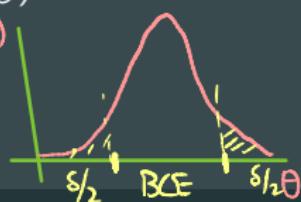
summarizing the posterior

model \mathcal{M} + prior $p(\Theta)$ + data $D \Rightarrow$ posterior $p(\Theta|D)$

summarizing $p(\Theta|D)$

- posterior mean $\widehat{\theta}_{mean} = \mathbb{E}[\Theta|D]$
- posterior mode (or MAP estimate) $\widehat{\theta}_{MAP} = \arg \max_{\Theta} p(\Theta|D)$
- posterior median $\widehat{\theta}_{median} = \min\{\Theta : p(\Theta|D) \geq 0.5\}$
- Bayesian credible intervals: given $\delta > 0$, want $(\ell_{\Theta}, u_{\Theta})$ s.t.

$$\underbrace{\mathbb{P}[\ell_{\Theta} \leq \Theta \leq u_{\Theta}|D]}_{> 1 - \delta}$$



marginal likelihood (evidence)

- data $D = \{X_1, X_2, \dots, X_n\} \in \{0, 1\}^n$, contains N_1 ones and N_0 zeros
- model \mathcal{M} : X_i are generated i.i.d. from a $Ber(\theta)$ distribution

marginal likelihood

$$p(D) = \frac{p(\theta)p(D|\theta)}{p(\theta|D)} = \frac{\text{prior} \times \text{likelihood}}{\text{posterior}}$$

$$P(D) = \frac{\Gamma(m)}{\Gamma(n+m)} \cdot \frac{\Gamma(N_1 + \alpha)}{\Gamma(\alpha)} \cdot \frac{\Gamma(N_0 + \beta)}{\Gamma(\beta)}$$

- Q: Given data, is it from $Ber(0.5)$ vs $Ber(\rho)$ for $\rho \notin [0.5-\delta, 0.5+\delta]$
- marginal likelihood lets you compare models'

multiclass data

- data $D = \{X_1, X_2, \dots, X_n\} \in \{1, 2, \dots, K\}^n$
- for $i \in [K]$, data D contains N_i copies of type i
- model \mathcal{M} : X_i generated i.i.d. from $Multinomial(\theta_1, \theta_2, \dots, \theta_K)$ distn