

STATG019 – Selected Topics in Statistics 2018

Lecture 5

Validation principles for time series and probabilistic prediction tasks



What is a time series?

Usual mathematical object model:

Time series taking values in \mathcal{X} at times in $\mathcal{T} \subseteq \mathbb{R}$ is a collection of \mathcal{X} -valued random variables X_t i.e., $(X_t; t \in \mathcal{T})$ where X_t t.v.in \mathcal{X}

Usual data storage object model:

 $((x_1,t_1),\ldots,(x_T,t_T))\in \operatorname{seq}(\mathcal{X}\times\mathcal{T})$ "sequences with values in $(\mathcal{X}\times\mathcal{T})$ " t_i are time stamps at which x_i are observed

Mathematical notation will use the *mathematical* object model (i.e., as usual in statistics, all data are random variables)

However, we will also use the data storage model as domain

i.e.,
$$X = (X_t; t \in \mathcal{T})$$
 takes values in $seq(\mathcal{X} \times \mathcal{T})$

 X_t is identified with an entry of the *ordered* tuple X (i.e., also carries knowledge of its own time stamp)

Time series, sequences & Co.

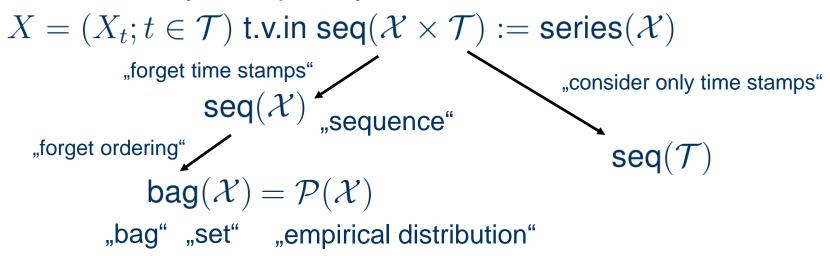
Usual mathematical object model:

complex

simple

Time series taking values in \mathcal{X} at times in $\mathcal{T} \subseteq \mathbb{R}$ is a collection of \mathcal{X} -valued random variables X_t i.e., $(X_t; t \in \mathcal{T})$ where X_t t.v.in \mathcal{X}

Time series carry multiple layers of structural information:



Occam's razor: if ordering is irrelevant, don't use in model, etc perhaps all information is contained in earliest time stamp?



Time series related modelling tasks Crucial distinction: what is the scientific goal?

are there No Yes multiple time series? (univariate or endogeneous) (multivariate or exogeneous) are these i.i.d.? time series analysis No Yes panel temporal Not temporal Prediction? data Yes Yes temporal Prediction? No Yes Yes No (Panel) **Change-point &** Fore-**Nowcasting Panel** Seriesanomaly detection casting **On-line models** forecasting as-features (combinations are possible)

"complexity" and validation depends on the task/setting!



The i.i.d. assumption divide

Two major "classes" of assumptions/settings/tasks

(A) i.i.d. (panel) samples are available

series-as-features, or panel modelling tasks

It is *crucial* to make use of the i.i.d.-ness assumption!

models also trained on other time series will be better

validatory guarantees obtained from sample will be stronger

(B) no i.i.d. assumption can be made

in essence: one object, observed at subsequent time points

Alternative assumptions lead to difficulties, much is open

models using i.i.d. strategies perform badly

validatory guarantees are weak and rely on assumptions

which in addition require checking!



Model validation in the i.i.d. settings



Time series tasks with i.i.d. data: overview

i.i.d. on-line learning: data is temporally revealed but actually i.i.d. Algorithm interface has fit, predict/trafo, and *update* for new data

Validation is as in the i.i.d. case – performance statistics get updated

Below: time-series-within-dataset "panel dataset"

| | energy use (time series) | user type (categorical) |
|---|--|----------------------------|
| 1 | ~~~~~ | residential |
| 2 | ~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~ | factory |

Series-as-features-task: series appear only as features
Sub-cases: supervised prediction, unsupervised learning
Related fields: functional data analysis, kernels, signal processing
Validation almost exactly like the "standard feature type" case

Panel data modelling: prediction/labelling task within series
Sub-cases: panel forecasting, panel anomaly/change-point d.

More technical but "easier" than non-panel forecasting etc
Validation "along the i.i.d. axis" easily generalizes tabular i.i.d case



Supervised learning with series-at-features

Observations $(X_1,Y_1),\ldots,(X_N,Y_N) \sim (X,Y)$ "primitive features" t.v.in (series $(\mathcal{X}) \times \mathcal{X}' \times \mathcal{Y}$

Estimate/learn

complex

$$\varepsilon(f):=\mathbb{E}\left[L\left(f(X),Y\right)
ight]$$
 is small where $L:\mathcal{Y}\times\mathcal{Y}\to\mathbb{R}$ is one of the usual losses

Validation: use conditionally i.i.d. test sample $L(f(X_i^*), Y_i^*)$ (just as in the "normal" supervised learning case)

Validation caveat: there is a hierarchy of baselines implied by

 $X = (X_t; t \in \mathcal{T}) \text{ t.v.in seq}(\mathcal{X} \times \mathcal{T}) := \text{series}(\mathcal{X})$ "consider only time stamps" $\text{seq}(\mathcal{X})$ "sequence" $\text{seq}(\mathcal{X}) = \mathcal{P}(\mathcal{X})$ $\text{seq}(\mathcal{T})$



Baselines in the series-as-features setting

Observations $(X_1,Y_1),\ldots,(X_N,Y_N) \sim_{\text{i.i.d.}} (X,Y)$ Estimate/learn t.v.in $(\text{series}(\mathcal{X}) \times \mathcal{X}') \times \mathcal{Y}$ a prediction functional f t.v.in $[(\text{series}(\mathcal{X}) \times \mathcal{X}') \to \mathcal{Y}]$

Validation caveat: there is a hierarchy of baselines/methods in order of "complexity" (later methods should improve on earlier ones)

- **0.** methods not using the series feature at all t.v.in $[\mathcal{X}' \to \mathcal{Y}]$
- 1a. methods using only simple summaries of time stamps
- 1b. methods using only simple summaries of the "bag"/set
- 2. methods using only the sequence of time stamps
- **3.** methods using only the bag/set t.v.in $[bag(\mathcal{X}) \times \mathcal{X}' \to \mathcal{Y}]$
- **4.** methods using only the sequence t.v.in $[seq(\mathcal{X}) \times \mathcal{X}' \to \mathcal{Y}]$
- 5. genuine series methods, not of the above type



The panel data learning setting

Observations $(X_1, Z_1), \ldots, (X_N, Z_N) \sim_{\text{i.i.d.}} (X, Z)$ t.v.in series $(\mathcal{X}) \times \mathcal{Z}$

Estimate/learn a decision rule f t.v.in [series(\mathcal{X}) \times $\mathcal{Z} \to \mathcal{S}$] \mathcal{S} is a set of decisions depending on the task

- e.g., (i) f t.v.in $[\operatorname{series}(\mathcal{X}) \times \mathcal{Z} \to \operatorname{series}(\mathcal{X})]$ "panel forecasting" $(x_{\leq \tau}, z) \mapsto \widehat{x}_{>\tau}$ where $(x_{\leq \tau} \text{ is "past"}, z \text{ "features"}, \widehat{x}_{>\tau} \text{ "future"}$
 - (ii) f t.v.in [series(\mathcal{X}) \times \mathcal{Z} \to series({yes, no})] "anomaly detection on panel data" e.g., heart rate monitor

Such that the expected generalization error

 $\varepsilon(f):=\mathbb{E}\left[L\left(f(X,Z),Y\right)
ight]$ is small—with Y ground truth and $L:\mathcal{S} imes\mathcal{Y} o\mathbb{R}$ an appropriate loss

Validation: use conditionally i.i.d. *test* sample $L(f(X_i^*, Z_i^*), Y_i^*)$

Task specific challenges are usually: defining loss, ground truth signal



Model validation in genuinely temporal settings



Time series forecasting setting

All data: $X = (X_t; t \in \mathcal{T})$ with $\mathcal{T} \subseteq \mathbb{R}$ t.v.in series(\mathcal{X})

Observations: $X_{\leq \tau} := (X_t ; t \in \mathcal{T}, t \leq \tau)$

Predict/forecast: $X_{>\tau} := (X_t ; t \in \mathcal{T}, t > \tau)$

by
$$\widehat{X}_{> au}:=\left(\widehat{X}_t\;;\;t\in\mathcal{T},t> au
ight)$$

Such that: the look-forward forecast errors

$$arepsilon(t) := \mathbb{E}\left[L\left(\widehat{X}_t, X_t
ight)\right]$$
 are small, $L: \mathcal{X} imes \mathcal{X} o \mathbb{R}$

Task variant 1: cut-off forecasting τ is fixed "present" $\widehat{X}_{>\tau}$ and $X_{>\tau}$ independent conditional $\widehat{X}_{\leq \tau}$ "don't use the future"

Task variant 2: sliding window forecasting τ moves forward

 \widehat{X}_{τ} and $X_{>\tau}$ independent conditional $\widehat{X}_{\leq \tau}$, for all τ

"don't use the future of a given prediction time point"



Bertrand Russell's turkey

The turkey wants to forecast whether it is still going to be alive tomorrow.

Starting with the first day of November,

$$X_1 = \mathsf{yes}, X_2 = \mathsf{yes}, X_3 = \mathsf{yes}, \ldots,$$

 $X_{
m day\ before\ the\ fourth\ Thursday\ of\ November}={
m yes}$



More generally, what to do if

 $X_1, X_2, \ldots, X_{\tau} \underset{\text{\tiny i.i.d.}}{\sim} X$ but $X_{\tau+1}$ may be completely different?

There's nothing one can do! ... in general.

i.e., generalizability assumptions are needed.



Ways out of the paradox

More generally, what to do if

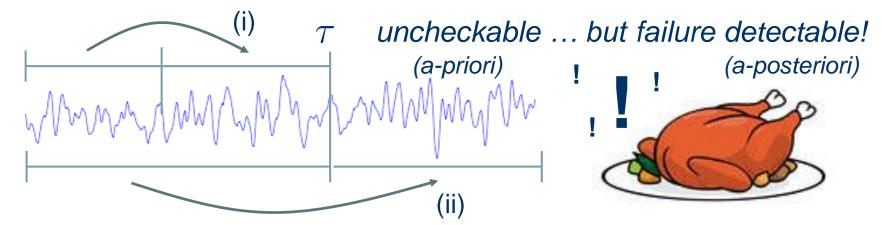
 $X_1, X_2, \ldots, X_{\tau} \sim X$ but $X_{\tau+1}$ may be completely different?

There's nothing one can do! ... in general.

i.e., generalizability assumptions are needed.

Two parts of the assumption: (even in the i.i.d. case!)

- (i) the observed data are similar/related testable
- (ii) future data are similar/related to observed data





Time series – generalizability assumptions

Two parts of the assumption:

- (i) the observed data are similar/related testable
- (ii) future data are similar/related to observed data

Prototypical (but very strong) assumption:

the full series $X = (X_t; t \in \mathcal{T})$ is (strongly) stationary

Definition: $X=(X_t\,;t\in\mathcal{T})$ is called *(strongly) stationary* iff the joint law of $X_{S+\delta}=(X_{t+\delta}\,;t\in S)$ where $\delta\in\mathbb{R},S\subseteq\mathcal{T}$ does not depend on the choice of S,δ

Intuitively: shifting the timestamps by δ does not change statistical properties

Testable by: Unit root tests, spectrum analysis tests (e.g., Dickey-Fuller) (e.g., Priestly-Subba Rao)

Cave: stationarity is strong assumption which is rarely true but validatory framework has all time series characteristic features



Time series – generalizability assumptions

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Including the "feature" that validation in time series is difficult:

General tests for the assumption only known in univariate case Change-point analysis ("stationarity breaks") open research topic



Guarantees and confidence intervals

Recall reason for main guarantees for the i.i.d. case:

Central Limit Theorem: Let
$$X_1, \ldots, X_N \underset{\text{i.i.d.}}{\sim} X$$
 (and assume all moments exist) Let $\widehat{\mu} = \frac{1}{N} \sum_{i=1}^N X_i$. Then $\sqrt{N} \left(\widehat{\mu} - \mathbb{E}[X] \right) \overset{d}{\to} \mathcal{N} \left(0, \text{Var}(X) \right)$ as $N \to \infty$

Central Limit Theorem for stationary time series:

Let
$$X=(X_t\;;t\in\mathbb{Z})$$
 be (strongly) stationary (and assume joint mgf is total) Let $\widehat{\mu}=\frac{1}{N}\sum_{i=1}^N X_i$. Then $\sqrt{N}\left(\widehat{\mu}-\mathbb{E}[X_{42}]\right)\overset{d}{\to}\mathcal{N}\left(0,\rho\cdot \mathrm{Var}(X_7)\right)$ as $N\to\infty$ where $\rho=\sum_{i=-\infty}^\infty \mathrm{Corr}(X_0,X_i)=1+2\sum_{i=1}^\infty \mathrm{Corr}(X_0,X_i)$ and where $\mu=\mathbb{E}[X_{42}]$

Proof sketch: Consider joint mgf $M_X(t_1,\ldots,t_N)$ of $(X_1,\ldots,X_N)-\mu$ and mgf $M_{\widehat{\mu}}(t)$ of $\widehat{\mu}-\mu$ and $M_{\widehat{\mu}}(t)=M_X\left(\frac{t}{N},\ldots,\frac{t}{N}\right)=\frac{t^2}{N^2}\cdot\mathbb{1}^{\top}\left(\begin{array}{cccc} \rho_0 & \rho_1 & \rho_2 & \ldots \\ \rho_{-1} & \rho_0 & \rho_1 & \ldots \\ \rho_2 & \rho_{-1} & \rho_0 & \ldots \\ \vdots & \ddots & \end{array}\right)$ "avg of entries" $\mathbb{1}^{\top} +O(N^{-3})$ where $\mathbb{1}^{\top} +O(N^{-3})$ where $\mathbb{1}^{\top} +O(N^{-3})$ and $\mathbb{1}^{\top} +O(N^{-3})$ where $\mathbb{1}^{\top} +O(N^{-3})$ and $\mathbb{1}^{\top} +O(N^{-3})$ where $\mathbb{1}^{\top} +O(N^{-3})$ and $\mathbb{1}^{\top} +O(N^{-3$



Guarantees and confidence intervals

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Consequence: $\widehat{\mu}$ estimates μ consistently, with normal asymptotic

 $\rho \cdot \text{Var}[X_7]$ may be estimated as mean of $\sum_s X_s X_t - \widehat{\mu}^2$

Confidence intervals similar to the i.i.d. setting (additional assumption: decay of autocorrelation)

 N/ρ is an "effective sample size": more correlation, wider CI

CI and tests for location comparisons as in i.i.d. setting but with effective sample size correction for variance



The Diebold-Mariano test for comparing forecasts

Forecast test set: $X = (X_t ; t \in \mathcal{T})$ where $\mathcal{T} = \{1, \ldots, T\}$

Two predictions to compare: $\widehat{X}^{(1)}$ and $\widehat{X}^{(2)}$, times also \mathcal{T}

- **1.** compute loss residuals $L^{(i)} := \left(L(\widehat{X}_t^{(i)}, X_t) : t \in \mathcal{T}\right)$
- 2. compute mean estimates $\widehat{\mu}^{(i)} := \sum_{t=1}^T L_t^{(i)}$ "performances" difference = "effect size "
- **3.** compute average autocorrelation $\widehat{v}^{(i)} := \sum_{s=1}^{\infty} \sum_{t=1}^{\infty} L_t^{(i)} L_s^{(i)} \left(\widehat{\mu}^{(i)}\right)^2$
- **4.** α -CI: $\left[\widehat{\mu}^{(i)} + \Phi^{-1}(\alpha/2) \cdot \sqrt{\widehat{v}^{(i)}/T} , \widehat{\mu}^{(i)} \Phi^{-1}(\alpha/2) \cdot \sqrt{\widehat{v}^{(i)}/T} \right]$
- 5. run two-sided t-test with statistic $\frac{\widehat{\mu}^{(1)}-\widehat{\mu}^{(2)}}{\sqrt{\widehat{v}^{(1)}+\widehat{v}^{(2)}}}$ "DM-statistic"

Cave: compares (sliding) predictions and not strategies! (same restrictions as in supervised setting, lecture 2 apply)



Forecasting workflow & validation principles

Checking/testing & keeping track of assumptions is important

For comparison, need to hold only on losses or loss differences i.i.d.-ness can hold for losses: (WW-)run tests, spectrum analysis tests Frequent assumptions: (weak/strong) stationarity

ergodicity & mixingness to take care of "coefficient decay"
plus local, periodic and difference versions of the above

Re-sampling schemes need to be temporal! "don't predict from future"
Sliding window/shift aggregation compatible with CLTs
necessary for validation, and usually beneficial in tuning

Generally, a broadly open (and difficult) research field...

Full understanding of a model-agnostic validatory workflow is missing the "right assumptions" for broad usefulness are unknown/open should be weaker than stationarity, but stronger than turkey (= none)



Exhaustive list of ML toolboxes for time series:

(i.e., with a unified sklearn-like interface)

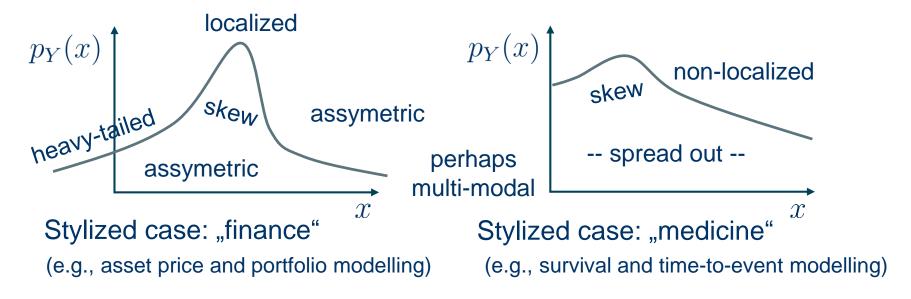


Theory of loss functionals for probabilistic prediction



The need for probabilistic prediction

arising from specifics of the data generative process in applications:



Both cases: modelling of shape, skewness, and spread crucial

Finance: "need to model higher-order moments" or "tail behaviour"

Medicine: "need to model hazard" or "survival function"

Mathematical approach: model the *conditional law* of Y|X=x i.d., conditional distribution functional $p_{Y|X}(y|x)$



Point prediction losses & elicitation

Can we use the standard setting for modelling & validation?

Setting: want to predict Y t.v.in \mathcal{Y} prediction $y \in \mathcal{Y}$ (temporarily: $\mathcal{Y} \subseteq \mathbb{R}$) goodness assessed through convex loss $L: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$ (recall: convex in 1st argument = prediction, by definition)

Lemma: $L: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$ is convex iff $L(\mathbb{E}[Z], y) \leq \mathbb{E}[L(Z, y)]$ for all r.v. Z t.v.in \mathcal{Y} and $y \in \mathcal{Z}$

Proof: Jensen's inequality/lemma applied to $[z \mapsto L(z,y)]$

Examples: $L:(\widehat{y},y)\mapsto (\widehat{y}-y)^2$ $L:(\widehat{y},y)\mapsto |\widehat{y}-y|$ (of convex losses) "squared loss" "absolute loss"

Definition: Let $L: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$ be a convex loss functional.

 $\begin{array}{ll} \text{(Gneiting,} & L \textit{ elicits} \text{ the statistic } T: \mathsf{Distr}(\mathcal{Y}) \to \mathbb{R} \\ & \text{if} & T(F_Y) = \arg\min_{y \in \mathcal{Y}} \mathbb{E}[L(y,Y)] & \text{for all r.v. } Y \text{ t.v.in } \mathcal{Y} \\ & & \text{(where } F_Y \text{ denotes cdf of } Y) \end{array}$



Point prediction losses & elicitation

Setting: want to predict Y t.v.in \mathcal{Y} prediction $y \in \mathcal{Y}$ (temporarily: $\mathcal{Y} \subseteq \mathbb{R}$) goodness assessed through convex loss $L: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$

Examples:
$$L_{sq}:(\widehat{y},y)\mapsto(\widehat{y}-y)^2$$
 $L_{abs}:(\widehat{y},y)\mapsto|\widehat{y}-y|$ (of convex losses) "squared loss" "absolute loss"

Definition: Let $L: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$ be a convex loss functional.

$$\begin{array}{ccc} \text{(Gneiting,} & L \textit{ elicits} \text{ the statistic } T: \mathsf{Distr}(\mathcal{Y}) \to \mathbb{R} \\ & \text{if} & T(F_Y) = \mathop{\arg\min}_{y \in \mathcal{Y}} \mathbb{E}[L(y,Y)] \text{ for all r.v. } Y \text{ t.v.in } \mathcal{Y} \end{array}$$

Examples: squared loss elicits mean: $\mathbb{E}(Y) = \arg\min_{y \in \mathcal{Y}} \mathbb{E}[L_{sq}(y,Y)]$ (of elicitation)

absolute loss elicits median: $median(Y) = \underset{y \in \mathcal{Y}}{\arg\min} \ \mathbb{E}[L_{abs}(y, Y)]$

Intuitively: squared and absolute loss

measure how well *location* is predicted!

useless to validate prediction of distributional features!

tails, skew, etc



Quantile losses for tail predictions

Setting: want to predict Y t.v.in \mathcal{Y} prediction $y \in \mathcal{Y}$ (temporarily: $\mathcal{Y} \subseteq \mathbb{R}$) goodness assessed through convex loss $L: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$

Definition: Let $L: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$ be a convex loss functional. L elicits the statistic $T: \mathrm{Distr}(\mathcal{Y}) \to \mathbb{R}$ if $T(F_Y) = \arg\min_{y \in \mathcal{Y}} \mathbb{E}[L(y,Y)]$ for all r.v. Y t.v.in \mathcal{Y}

Definition: $L_{\alpha}: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R} \; ; \; (\widehat{y}, y) \mapsto \alpha \cdot m(y, \widehat{y}) + (1 - \alpha) \cdot m(\widehat{y}, y)$ is called $(\alpha$ -)quantile loss where $m(x, z) = \min(x - z, 0)$

Proposition: L_{α} elicits the α -quantile that is, $F_Y^{-1}(\alpha) = \operatorname*{arg\,min}_{y \in \mathcal{Y}} \mathbb{E}[L_{\alpha}(y,Y)]$ useful to validate "value at risk" predictions!

Problems: for "full picture", multiple predictions/losses needed quantile measures may have high loss variance (thesis?)



Fully probabilistic predictions

Setting: want to predict Y t.v.in \mathcal{Y} e.g., absolutely continuous (pdf exist) prediction $p \in \mathcal{P} \subseteq \mathsf{Distr}(\mathcal{Y})$ i.e., p is a distribution in \mathcal{P} goodness assessed through loss $L: \mathcal{P} \times \mathcal{Y} \to \mathbb{R}$

Definition: Let $L: \mathcal{P} \times \mathcal{Y} \to \mathbb{R}$ be a convex loss functional.

(Dawid, 1990s) L is called *strictly proper* if for all r.v. Y t.v.in ${\mathcal Y}$

$$p_Y = \operatorname*{arg\,min}_{p \in \mathcal{P}} \mathbb{E}[L(p,Y)]$$
 where $Y \sim p_Y$ (note: definition is w.r.t. \mathcal{P})

Intuition: best prediction (w.r.t. expected loss) is "true" distribution

(of strictly proper losses) $L_{log}:(p,y)\mapsto -\log p(y)$ "logarithmic loss" "cross-entropy loss" $L_{sg}:(\widehat{y},y)\mapsto -2p(y)+\|p\|_2^2$

Note: defined only for discrete or absolutely continuous distributions

"squared integrated loss" "Brier loss (if discrete)"

where p(y) is shorthand for pdf/pdf at y and $\|p\|_2^2 := \int_{\mathcal{V}} p(y)^2 \; dy$



Fully probabilistic predictions

Setting: want to predict Y t.v.in \mathcal{Y} e.g., absolutely continuous (pdf exist) prediction $p \in \mathcal{P} \subseteq \mathsf{Distr}(\mathcal{Y})$ i.e., p is a distribution in \mathcal{P} goodness assessed through loss $L: \mathcal{P} \times \mathcal{Y} \to \mathbb{R}$

$$\begin{array}{ll} \textbf{Examples:} & L_{log}:(p,y)\mapsto -\log p(y)\\ \text{(of strictly proper losses)} & \text{"logarithmic loss" "cross-entropy loss"}\\ & L_{isq}:(p,y)\mapsto -2p(y)+\|p\|_2^2\\ & \text{"squared integrated loss"} & \text{"Brier loss (if discrete)"} \end{array}$$

Note: defined only for discrete or absolutely continuous distributions

Short proofs of strict properness:

$$\mathbb{E}[L_{log}(p,Y)] - \mathbb{E}[L_{log}(p_Y,Y)] = \int_{\mathcal{Y}} p_Y \cdot \log \frac{p_Y(y)}{p(y)} \ dy = \mathsf{D}_{\mathit{KL}}(p||p_Y)$$
 "cross-entropy" "Kullback-Leibler-divergence"

$$\mathbb{E}[L_{isq}(p,Y)] - \mathbb{E}[L_{isq}(p_Y,Y)] = \int_{\mathcal{Y}} \left(p(y) - p_Y(y)\right)^2 \ dy = \mathsf{D}_{sq}(p,p_Y)$$
 then use school math "integrated squared distance/divergence"



Example: validation of unsupervised methods

Unsupervised setting (full probabilistic model):

Given data
$$X_1, \ldots, X_N \sim X \sim p \in \mathcal{P} \subseteq \mathsf{Distr}(\mathcal{X})$$

Estimate p via $\widehat{p} \in \mathcal{P}$ (using the data X_i) (assume discrete, or absolutely continuous)

Goodness assessed through probabilistic loss $L: \mathcal{P} \times \mathcal{Y} \to \mathbb{R}$ i.e., $\mathbb{E}[L(\widehat{p}, X)]$ is small.

Validation: use conditionally i.i.d. test sample $L(\widehat{p}, X_i^*)$

Note:
$$\frac{1}{M} \sum_{i=1}^{M} L(\widehat{p}, X_i^*)$$
 for the log/cross-entropy loss

is the negative out-of-sample likelihood of \widehat{p} , up to a constant factor

Cave: this validation strategy applies only to unsupervised methods with a full probabilistic model



WORQ - widely open research questions!

(aka BSc/MSc/PhD topics for theory/practice cross-over inclined)

Completing the model-agnostic validation workflows

What is the right formulation for a fully model-agnostic setting?
Which assumptions are general enough but not too general to be useful?
Right combination of change-point detection and assumption testing?

Basic toolbox design for the time series related tasks

What are the right data containers and data pipelines?
How to encapsulate and interface the different tasks, models, strategies?
Combination of time series, on-line, probabilistic, and heterogeneity?

Meta-learning for time series related tasks

What are good re-sampling schemes, and related learning guarantees? Tuning, ensembling and composition for time series & probabilistic losses Composition of prediction and anomaly/change-point detection methods?

Also: systematic study of which methods *really* work

Deep learning for everything? Conditional on complete validation workflow ...