## 2021 年春季学期/数理统计/第二周/课后作业解答

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4 证明. 样本容量为 n+1 时的样本均值  $\bar{x}_{n+1}$ :

$$\bar{x}_{n+1} = \frac{1}{n+1} \sum_{i=1}^{n+1} x_i = \frac{1}{n+1} \left( \sum_{i=1}^n x_i + x_{n+1} \right) = \frac{1}{n+1} \left( n\bar{x}_n + x_{n+1} \right)$$
$$= \frac{1}{n+1} [(n+1)\bar{x}_n - \bar{x}_n + x_{n+1}] = \bar{x}_n + \frac{1}{n+1} (x_{n+1} - \bar{x}_n)$$

样本容量为 n+1 时的样本方差  $s_{n+1}^2$ :

$$s_{n+1}^{2} = \frac{1}{n} \sum_{i=1}^{n+1} (x_{i} - \bar{x}_{n+1})^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n} \left\{ x_{i} - \left[ \bar{x}_{n} + \frac{1}{n+1} (x_{n+1} - \bar{x}_{n}) \right] \right\}^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n+1} \left[ (x_{i} - \bar{x}_{n}) - \frac{1}{n+1} (x_{n+1} - \bar{x}_{n}) \right]^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n+1} \left\{ (x_{i} - \bar{x}_{n})^{2} - \frac{2}{n+1} (x_{i} - \bar{x}_{n}) (x_{n+1} - \bar{x}_{n}) + \frac{1}{(n+1)^{2}} (x_{n+1} - \bar{x}_{n})^{2} \right\}$$

$$= \frac{1}{n} \left[ \sum_{i=1}^{n} (x_{i} - \bar{x}_{n})^{2} \right] + \frac{1}{n} (x_{n+1} - \bar{x}_{n})^{2} - \frac{2}{n(n+1)} (x_{n+1} - \bar{x}_{n}) \left[ \sum_{i=1}^{n} (x_{i} - \bar{x}_{n}) \right]$$

$$- \frac{2}{n(n+1)} (x_{n+1} - \bar{x}_{n})^{2} + \frac{1}{n(n+1)} (x_{n+1} - \bar{x}_{n})^{2}$$

$$= \frac{n-1}{n} s_{n}^{2} + \frac{1}{n+1} (x_{n+1} - \bar{x}_{n})^{2}$$

(另一种思路)

$$\begin{split} s_{n+1}^2 &= \frac{1}{n} \sum_{i=1}^{n+1} (x_i - \bar{x}_{n+1})^2 = \frac{1}{n} \sum_{i=1}^{n+1} \left[ (x_i - \bar{x}_n) + (\bar{x}_n - \bar{x}_{n+1}) \right]^2 \\ &= \frac{1}{n} \left[ \sum_{i=1}^{n+1} (x_i - \bar{x}_n)^2 \right] + \frac{2}{n} (\bar{x}_n - \bar{x}_{n+1}) \sum_{i=1}^{n+1} (x_i - \bar{x}_n) + \frac{1}{n} \sum_{i=1}^{n+1} (\bar{x}_n - \bar{x}_{n+1})^2 \\ &= \frac{1}{n} \left[ \sum_{i=1}^{n+1} (x_i - \bar{x}_n)^2 \right] - \frac{n+1}{n} (\bar{x}_n - \bar{x}_{n+1})^2 \\ &= \frac{1}{n} \left[ \sum_{i=1}^{n} (x_i - \bar{x}_n)^2 + (x_{n+1} - \bar{x}_n)^2 \right] - \frac{n+1}{n} \frac{1}{(n+1)^2} (x_{n+1} - \bar{x}_n)^2 \\ &= \frac{1}{n} \left[ (n-1) \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x}_n)^2 + \frac{n}{n+1} (x_{n+1} - \bar{x}_n)^2 \right] \\ &= \frac{n-1}{n} s_n^2 + \frac{1}{n+1} (x_{n+1} - \bar{x}_n)^2 \end{split}$$

6 证明. 样本 B 的均值  $\bar{y}_B$ :

$$\bar{y}_B = \frac{1}{n} \sum_{i=1}^n y_i = \frac{1}{n} \sum_{i=1}^n (ax_i + b) = \frac{1}{n} \left( a \sum_{i=1}^n x_i + nb \right)$$
$$= a \cdot \frac{1}{n} \sum_{i=1}^n x_i + b = a\bar{x}_A + b$$

样本 B 的标准差  $s_B$ :

$$s_B = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y}_B)^2} = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (ax_i + b - a\bar{x}_A - b)^2}$$
$$= |a| \cdot \sqrt{\frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x}_A)^2} = |a| s_A$$

样本 B 的极差  $R_B$ :

$$R_B = y_{(n)} - y_{(1)} = ax_{(n)} + b - ax_{(1)} - b = a \left[ x_{(n)} - x_{(1)} \right] = aR_A$$

样本 B 的中位数  $m_{0.5B}$ :

(a). 当 n 为偶数时,

$$m_{0.5B} = y_{\left(\frac{n+1}{2}\right)} = ax_{\left(\frac{n+1}{2}\right)} + b = am_{0.5A} + b$$

(b). 当 n 为奇数时,

$$\begin{split} m_{0.5B} = & \frac{1}{2} \left[ y_{\left(\frac{n}{2}\right)} + y_{\left(\frac{n}{2}+1\right)} \right] = \frac{1}{2} \left[ a x_{\left(\frac{n}{2}\right)} + b + a x_{\left(\frac{n}{2}+1\right)} + b \right] \\ = & \frac{a}{2} \left[ x_{\left(\frac{n}{2}\right)} + x_{\left(\frac{n}{2}+1\right)} \right] + b = a m_{0.5A} + b \end{split}$$

因此, 样本 B 的中位数  $m_{0.5B}$  为

$$m_{0.5B} = am_{0.5A} + b.$$

8 证明. 由定理 5.3.2 有,

$$E(\bar{X}) = E(X_i) = \mu = 0, \quad Var(\bar{X}) = \frac{\sigma^2}{n} = \frac{1}{3n}.$$

9 **证明.** 因为样本  $X_i$ , i = 1, 2, ..., n 是相互独立的,所以

$$Cov(X_i, X_j) = 0, \quad (i \neq j).$$

因此,

$$\operatorname{Cov}(X_i, \bar{X}) = \operatorname{Cov}\left(X_i, \frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \operatorname{Cov}(X_i, X_i) = \frac{\sigma^2}{n},$$
$$\operatorname{Cov}(\bar{X}, \bar{X}) = \operatorname{Var}(\bar{X}) = \frac{1}{n} \sum_{i=1}^n \operatorname{Cov}(X_i, \bar{X}) = \frac{\sigma^2}{n}.$$

$$\operatorname{Cov}\left(X_{i} - \bar{X}, X_{j} - \bar{X}\right) = \operatorname{Cov}\left(X_{i}, X_{j}\right) - \operatorname{Cov}\left(X_{i}, \bar{X}\right) - \operatorname{Cov}\left(X_{j}, \bar{X}\right) + \operatorname{Cov}(\bar{X}, \bar{X})$$

$$= 0 - \frac{1}{n}\sigma^{2} - \frac{1}{n}\sigma^{2} + \frac{1}{n}\sigma^{2} = -\frac{1}{n}\sigma^{2}$$

$$\operatorname{Var}\left(X_{i} - \bar{X}\right) = \operatorname{Var}\left(X_{i}\right) + \operatorname{Var}(\bar{X}) - 2\operatorname{Cov}\left(X_{i}, \bar{X}\right)$$

$$= \sigma^{2} + \frac{1}{n}\sigma^{2} - \frac{2}{n}\sigma^{2} = \frac{n-1}{n}\sigma^{2}$$

同理,我们有

$$\operatorname{Var}\left(X_{j} - \bar{X}\right) = \frac{n-1}{n}\sigma^{2};$$

所以,

$$\operatorname{Corr}\left(X_{i} - \bar{X}, X_{j} - \bar{X}\right) = \frac{\operatorname{Cov}\left(X_{i} - \bar{X}, X_{j} - \bar{X}\right)}{\sqrt{\operatorname{Var}\left(X_{i} - \bar{X}\right)} \cdot \sqrt{\operatorname{Var}\left(X_{j} - \bar{X}\right)}}$$
$$= \frac{-\frac{1}{n}\sigma^{2}}{\sqrt{\frac{n-1}{n}\sigma^{2}} \cdot \sqrt{\frac{n-1}{n}\sigma^{2}}} = -\frac{1}{n-1}$$

18 证明. 由定理 5.3.2 有,

$$\operatorname{Var}(\bar{X}) = \frac{\sigma^2}{n} = \frac{9}{8}, \quad \sigma(\bar{X}) = \sqrt{\operatorname{Var}(\bar{X})} = \frac{3\sqrt{2}}{4}.$$

23 证明. (可参考例 3.3.4 与例 3.3.5)

$$P(X \le k) = \sum_{i=1}^{k} pq^{i-1} = \frac{p(1-q^k)}{1-q} = 1 - q^k, \quad k = 1, 2, \dots,$$

对于 $X_{(n)}$ ,有

$$P(X_{(n)} = k) = P(X_{(n)} \le k) - P(X_{(n)} \le k - 1)$$

$$= \prod_{i=1}^{n} P(X_{i} \le k) - \prod_{i=1}^{n} P(X_{i} \le k - 1)$$

$$= (1 - q^{k})^{n} - (1 - q^{k-1})^{n}$$

对于  $X_{(1)}$ ,有

$$\begin{split} P\left(X_{(1)} = k\right) &= P\left(X_{(1)} \le k\right) - P\left(X_{(1)} \le k - 1\right) \\ &= 1 - P\left(X_{(1)} > k\right) - \left[1 - P\left(X_{(1)} > k - 1\right)\right] \\ &= \prod_{i=1}^{n} P\left(X_{i} > k - 1\right) - \prod_{i=1}^{n} P\left(X_{i} > k\right) \\ &= q^{n(k-1)} - q^{nk} = q^{n(k-1)} \left(1 - q^{n}\right) \end{split}$$

25 证明. 韦布尔分布的总体分布函数 F(x) 为

$$F(x) = \int_0^x p(t) dt = \int_0^x \frac{mt^{m-1}}{\eta^m} e^{-\left(\frac{t}{\eta}\right)^m} dt = \int_0^x e^{-\left(\frac{t}{\eta}\right)^m} d\left(\frac{t}{\eta}\right)^m$$
$$= -e^{-\left(\frac{t}{\eta}\right)^m} \Big|_0^x = 1 - e^{-\left(\frac{x}{\eta}\right)^m}$$

因此,

$$p_{(1)}(x) = n[1 - F(x)]^{n-1}p(x) = ne^{-(n-1)\left(\frac{x}{\eta}\right)^m} \cdot \frac{mx^{m-1}}{\eta^m} e^{-\left(\frac{x}{\eta}\right)^m}$$
$$= \frac{mnx^{m-1}}{\eta^m} e^{-n\left(\frac{x}{\eta}\right)^m} = \frac{mx^{m-1}}{(\eta/\sqrt[m]{\eta})^m} e^{-\left(\frac{x}{\eta/\sqrt[m]{\eta}}\right)^m}$$

所以, $X_{(1)}$  服从参数为  $\left(m, \frac{\eta}{\sqrt[\eta]{n}}\right)$  的韦布尔分布。