

2022 年春季学期/数理统计/第三周/课后作业解答

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4 **证明.** 样本容量为 $n+1$ 时的样本均值 \bar{x}_{n+1} :

$$\begin{aligned}\bar{x}_{n+1} &= \frac{1}{n+1} \sum_{i=1}^{n+1} x_i = \frac{1}{n+1} \left(\sum_{i=1}^n x_i + x_{n+1} \right) = \frac{1}{n+1} (n\bar{x}_n + x_{n+1}) \\ &= \frac{1}{n+1} [(n+1)\bar{x}_n - \bar{x}_n + x_{n+1}] = \bar{x}_n + \frac{1}{n+1} (x_{n+1} - \bar{x}_n)\end{aligned}$$

样本容量为 $n+1$ 时的样本方差 s_{n+1}^2 :

$$\begin{aligned}s_{n+1}^2 &= \frac{1}{n} \sum_{i=1}^{n+1} (x_i - \bar{x}_{n+1})^2 \\ &= \frac{1}{n} \sum_{i=1}^n \left\{ x_i - \left[\bar{x}_n + \frac{1}{n+1} (x_{n+1} - \bar{x}_n) \right] \right\}^2 \\ &= \frac{1}{n} \sum_{i=1}^{n+1} \left[(x_i - \bar{x}_n) - \frac{1}{n+1} (x_{n+1} - \bar{x}_n) \right]^2 \\ &= \frac{1}{n} \sum_{i=1}^{n+1} \left\{ (x_i - \bar{x}_n)^2 - \frac{2}{n+1} (x_i - \bar{x}_n) (x_{n+1} - \bar{x}_n) + \frac{1}{(n+1)^2} (x_{n+1} - \bar{x}_n)^2 \right\} \\ &= \frac{1}{n} \left[\sum_{i=1}^n (x_i - \bar{x}_n)^2 \right] + \frac{1}{n} (x_{n+1} - \bar{x}_n)^2 - \frac{2}{n(n+1)} (x_{n+1} - \bar{x}_n) \left[\sum_{i=1}^n (x_i - \bar{x}_n) \right] \\ &\quad - \frac{2}{n(n+1)} (x_{n+1} - \bar{x}_n)^2 + \frac{1}{n(n+1)} (x_{n+1} - \bar{x}_n)^2 \\ &= \frac{n-1}{n} s_n^2 + \frac{1}{n+1} (x_{n+1} - \bar{x}_n)^2\end{aligned}$$

(另一种思路)

$$\begin{aligned}
s_{n+1}^2 &= \frac{1}{n} \sum_{i=1}^{n+1} (x_i - \bar{x}_{n+1})^2 = \frac{1}{n} \sum_{i=1}^{n+1} [(x_i - \bar{x}_n) + (\bar{x}_n - \bar{x}_{n+1})]^2 \\
&= \frac{1}{n} \left[\sum_{i=1}^{n+1} (x_i - \bar{x}_n)^2 \right] + \frac{2}{n} (\bar{x}_n - \bar{x}_{n+1}) \sum_{i=1}^{n+1} (x_i - \bar{x}_n) + \frac{1}{n} \sum_{i=1}^{n+1} (\bar{x}_n - \bar{x}_{n+1})^2 \\
&= \frac{1}{n} \left[\sum_{i=1}^{n+1} (x_i - \bar{x}_n)^2 \right] - \frac{n+1}{n} (\bar{x}_n - \bar{x}_{n+1})^2 \\
&= \frac{1}{n} \left[\sum_{i=1}^n (x_i - \bar{x}_n)^2 + (x_{n+1} - \bar{x}_n)^2 \right] - \frac{n+1}{n} \frac{1}{(n+1)^2} (x_{n+1} - \bar{x}_n)^2 \\
&= \frac{1}{n} \left[(n-1) \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x}_n)^2 + \frac{n}{n+1} (x_{n+1} - \bar{x}_n)^2 \right] \\
&= \frac{n-1}{n} s_n^2 + \frac{1}{n+1} (x_{n+1} - \bar{x}_n)^2
\end{aligned}$$

6 **证明.** 样本 B 的均值 \bar{y}_B :

$$\begin{aligned}
\bar{y}_B &= \frac{1}{n} \sum_{i=1}^n y_i = \frac{1}{n} \sum_{i=1}^n (ax_i + b) = \frac{1}{n} \left(a \sum_{i=1}^n x_i + nb \right) \\
&= a \cdot \frac{1}{n} \sum_{i=1}^n x_i + b = a\bar{x}_A + b
\end{aligned}$$

样本 B 的标准差 s_B :

$$\begin{aligned}
s_B &= \sqrt{\frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y}_B)^2} = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (ax_i + b - a\bar{x}_A - b)^2} \\
&= |a| \cdot \sqrt{\frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x}_A)^2} = |a|s_A
\end{aligned}$$

样本 B 的极差 R_B :

$$R_B = y_{(n)} - y_{(1)} = ax_{(n)} + b - ax_{(1)} - b = a[x_{(n)} - x_{(1)}] = aR_A$$

样本 B 的中位数 $m_{0.5B}$:

(a). 当 n 为偶数时,

$$m_{0.5B} = y_{(\frac{n+1}{2})} = ax_{(\frac{n+1}{2})} + b = am_{0.5A} + b$$

(b). 当 n 为奇数时,

$$\begin{aligned}
m_{0.5B} &= \frac{1}{2} \left[y_{(\frac{n}{2})} + y_{(\frac{n}{2}+1)} \right] = \frac{1}{2} \left[ax_{(\frac{n}{2})} + b + ax_{(\frac{n}{2}+1)} + b \right] \\
&= \frac{a}{2} \left[x_{(\frac{n}{2})} + x_{(\frac{n}{2}+1)} \right] + b = am_{0.5A} + b
\end{aligned}$$

因此, 样本 B 的中位数 $m_{0.5B}$ 为

$$m_{0.5B} = am_{0.5A} + b.$$

23 证明. (可参考例 3.3.4 与例 3.3.5)

$$P(X \leq k) = \sum_{i=1}^k pq^{i-1} = \frac{p(1-q^k)}{1-q} = 1-q^k, \quad k=1,2,\dots,$$

对于 $X_{(n)}$, 有

$$\begin{aligned} P(X_{(n)} = k) &= P(X_{(n)} \leq k) - P(X_{(n)} \leq k-1) \\ &= \prod_{i=1}^n P(X_i \leq k) - \prod_{i=1}^n P(X_i \leq k-1) \\ &= (1-q^k)^n - (1-q^{k-1})^n \end{aligned}$$

对于 $X_{(1)}$, 有

$$\begin{aligned} P(X_{(1)} = k) &= P(X_{(1)} \leq k) - P(X_{(1)} \leq k-1) \\ &= 1 - P(X_{(1)} > k) - [1 - P(X_{(1)} > k-1)] \\ &= \prod_{i=1}^n P(X_i > k-1) - \prod_{i=1}^n P(X_i > k) \\ &= q^{n(k-1)} - q^{nk} = q^{n(k-1)}(1-q^n) \end{aligned}$$

32 证明. 总体 X 的密度函数和分布函数分别为

$$p(x) = \begin{cases} 3x^2, & 0 < x < 1; \\ 0, & \text{其他.} \end{cases} \quad F(x) = \begin{cases} 0, & x < 0; \\ x^3, & 0 \leq x < 1; \\ 1, & x \geq 1. \end{cases}$$

所以, 根据 P242 (5.3.16), $(X_{(2)}, X_{(4)})$ 的联合密度函数为

$$\begin{aligned} p(x, y) &= \frac{5!}{1! \cdot 1! \cdot 1!} [F(x)]^{2-1} [F(y) - F(x)]^{4-2-1} [1 - F(y)]^{5-4} p(x) p(y) \\ &= 120x^3 (y^3 - x^3) (1 - y^3) \cdot 3x^2 \cdot 3y^2 \\ &= 1080x^5 y^2 (y^3 - x^3) (1 - y^3), \quad 0 < x < y < 1 \end{aligned}$$

为求 $\left(\frac{X_{(2)}}{X_{(4)}}, X_{(4)}\right)$ 的联合密度, 令

$$\begin{cases} \mu = \frac{x}{y} \\ \nu = y \end{cases} \implies \begin{cases} x = \mu\nu \\ y = \nu \end{cases}$$

则, 其雅可比行列式为

$$J = \frac{\partial(x, y)}{\partial(\mu, \nu)} = \begin{vmatrix} \nu & \mu \\ 0 & 1 \end{vmatrix} = \nu$$

由 $0 < x < y < 1$ 可得 $0 < \mu < 1, 0 < \nu < 1$ 。于是, $\left(\frac{X_{(2)}}{X_{(4)}}, X_{(4)}\right)$ 的联合密度函数为

$$\begin{aligned} p(\mu, \nu) &= p(\mu\nu, \nu) \cdot |J| = 1080 (\mu\nu)^5 \nu^2 [\nu^3 - (\mu\nu)^3] (1 - \nu^3) \\ &= 1080 \mu^5 (1 - \mu^3) \cdot \nu^{11} (1 - \nu^3) \end{aligned}$$

可求得其边际密度函数为,

$$\begin{aligned}U = \frac{X_{(2)}}{X_{(4)}} \sim p(\mu) &= \int_0^1 p(\mu, \nu) d\nu = 1080\mu^5 (1 - \mu^3) \int_0^1 \nu^{11} (1 - \nu^3) d\nu \\&= 18\mu^5 (1 - \mu^3), \quad 0 < \mu < 1\end{aligned}$$

类似可得

$$V = X_{(4)} \sim p(\nu) = 60\nu^{11} (1 - \nu^3), \quad 0 < \nu < 1$$

因此, 可以证明

$$p(\mu, \nu) = p(\mu) \cdot p(\nu)$$

即, $\frac{X_{(2)}}{X_{(4)}}$ 与 $X_{(4)}$ 独立。

评论 也可利用变量独立的一个推论, 判断 $p(\mu, \nu)$ 可分离变量, 省去求解边际密度函数的步骤。

推论 令 $(X, Y) \sim p(x, y)$, 则 X 与 Y 相互独立的充要条件是, $p(x, y)$ 可分离变量, 即存在非负函数 $g(x)$, $h(y)$ 使得 $p(x, y) = g(x) \cdot h(y)$ 。