Notes: Double/Debiased Machine Learning

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1 Intution

Consider the partical linear regression model

$$Y_i = D_i \theta + g(\mathbf{X}_i) + U_i, \quad E(U_i \mid \mathbf{X}_i, D_i) = 0$$

$$D_i = m(\mathbf{X}_i) + V_i, \quad E(V_i \mid \mathbf{X}_i) = 0$$
(1)

where $\mathbf{X}_i \in \mathbb{R}^p$, $i = 1, 2, \dots, n$. Suppose

$$m(\mathbf{X}_{i}) = \gamma_{1} x_{i,1} + \gamma_{2} x_{i,2} + \gamma_{3} x_{i,3} + U_{i}$$

$$g(\mathbf{X}_{i}) = \beta_{1} x_{i,1} + \beta_{2} x_{i,2} + \beta_{3} x_{i,3} + V_{i}$$
(2)

and

$$n = 1000, \quad p = 200, \quad \gamma = (3, 2, 1)', \quad \beta = (1, 2, 3)'$$

1.1 Regularization Bias

1. Non Orthogonal: Suppose \hat{g} are estimated using the auxiliary samples indexed by I^c , thus

$$\check{\theta} = \left(\frac{1}{n} \sum_{i \in I} D_i^2\right)^{-1} \frac{1}{n} \sum_{i \in I} D_i \left(Y_i - \hat{g}(X_i)\right)$$
(3)

2. Orthogonal: Suppose \hat{g} and \hat{m} are estimated using the auxiliary samples indexed by I^c , thus

$$\check{\theta} = \left(\frac{1}{n} \sum_{i \in I} \widehat{V}_i D_i\right)^{-1} \frac{1}{n} \sum_{i \in I} \widehat{V}_i \left(Y_i - \hat{g}\left(X_i\right)\right) \tag{4}$$

where $\hat{V} = D - \hat{m}(\mathbf{X})$.

1.2 Overfitting Bias

1.3 Score Functions

Definition 1.1 (Neymaxn Orthogonal Score Function).

Remark. Suppose the score function is the linearity function of in the parameter θ , that,

$$\psi\left(\mathbf{W};\theta,\boldsymbol{\eta}\right) = \psi_a\left(\mathbf{W};\boldsymbol{\eta}\right)\theta + \psi_b\left(\mathbf{W};\boldsymbol{\eta}\right) \tag{5}$$

Hence the estimator can be written as

$$\check{\theta} = -\frac{\mathbb{E}_{N} \left[\psi_{b} \left(\mathbf{W}; \boldsymbol{\eta} \right) \right]}{\mathbb{E}_{N} \left[\psi_{a} \left(\mathbf{W}; \boldsymbol{\eta} \right) \right]} \tag{6}$$

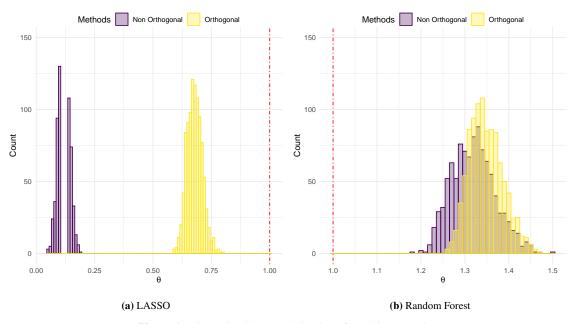


Figure 1: Bias Raised by Regularization of Machine Learning

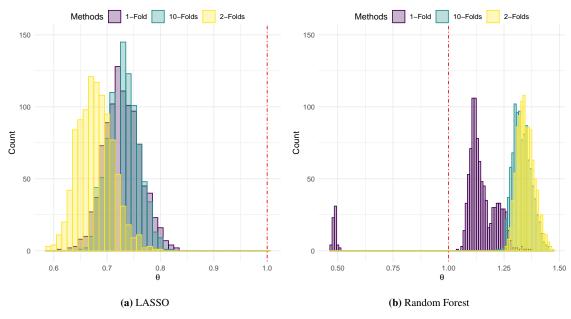


Figure 2: Bias Raised by Overfitting of Machine Learning

1. IV-type: The score function of IV-type is

$$\psi\left(\mathbf{W};\theta,\boldsymbol{\eta}\right) := \left[Y - D\theta - g(\mathbf{X})\right]\left[D - m(\mathbf{X})\right] \tag{7}$$

where $\eta=(g,m)$. Suppose \hat{g} and \hat{m} are estimated using the auxiliary samples indexed by I^c , Thus

$$\check{\theta} = \left(\frac{1}{n} \sum_{i \in \mathcal{I}} \widehat{V}_i D_i\right)^{-1} \frac{1}{n} \sum_{i \in \mathcal{I}} \widehat{V}_i \left(Y_i - \widehat{g}\left(X_i\right)\right) \tag{8}$$

where $\hat{V} = D - \hat{m}(\mathbf{X})$.

2. Partialling-Out: The score function of partialling-out is

$$\psi\left(\mathbf{W}; \theta, \boldsymbol{\eta}\right) := \left[Y - \ell\left(\mathbf{X}\right) - \theta\left(D - m(\mathbf{X})\right)\right] \left[D - m(\mathbf{X})\right] \tag{9}$$

where $\eta = (\ell, m)$. Suppose $\hat{\ell}$ and \hat{m} are estimated using the auxiliary samples indexed by I^c , Thus

$$\check{\theta} = \left(\frac{1}{n}\sum_{i\in\mathcal{I}}\widehat{V}_i^2\right)^{-1}\frac{1}{n}\sum_{i\in\mathcal{I}}\widehat{V}_i\left(Y_i - \widehat{\ell}\left(X_i\right)\right) \tag{10}$$

where $\hat{V} = D - \hat{m}(\mathbf{X})$.

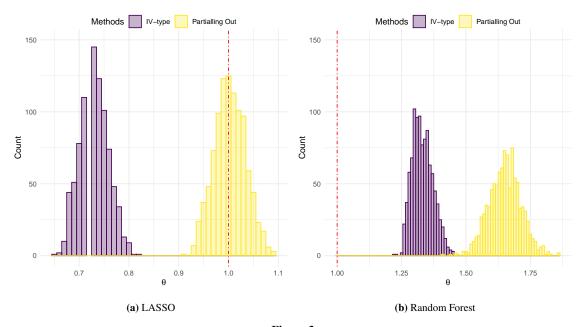


Figure 3

1.4 Complexity of Paramters

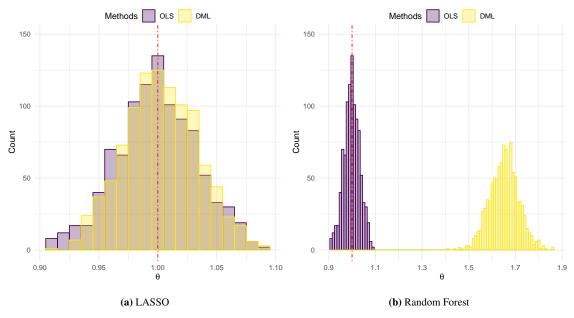


Figure 4