530 陈斯杰 电子信息工程 第19次作业

一、时间序列

1.

实现代码如下

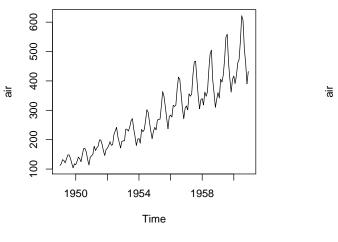
```
data (robot)
    x<-data.frame(robot)
    ro < -ts(x)
3
    > air.fit <- arima(aplog, order=c(0,1,1), seasonal=list(order=c
4
       (0,1,1), period=12))
    > m1=arima(x = ro, order = c(1, 0, 0))
5
  > m2=arima(x = ro, order = c(0, 1, 1))
  > acf(ro, 80)
  > pacf(ro, 30)
 > write.table(robot, "robot.csv", sep=",")
 > plot.ts(ro)
 > plot.ts(diff(aplog)
 |> aplog1 < -diff(aplog, diff=1)
 > aplog222 < -diff(aplog1, lag = 12)
 > plot(rstandard(m1), type="o")
 |> plot(rstandard(m2), type="o")
 | > acf(rstandard(m1))
 > acf(rstandard(m2))
 |> shapiro.test(ro)
  > tsdiag(data.fit)
 > m4 < -arima(ro, order = c(1, 0, 1))
  > airfore4<-forecast.Arima(m4, h=5, level=c(99.5))
```

2.解:导入TSA包并画出原始形式和取对数之后的时间序列图,R语言代码如下:

```
library (TSA)
air <- AirPassengers

plot(air)
air = log(air)
plot(air)</pre>
```

画出的时间序列图如下所示: 由图可见,原始形式和取对数之后的时间序列



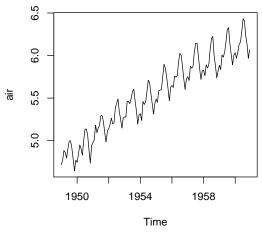


图 1: Initial

图 2: log(Initial)

图是大致相似的因此对数变换在这里是适当的。

求取对数后的时间序列的一次差分时间序列图和样本ACF图的R语言代码为:

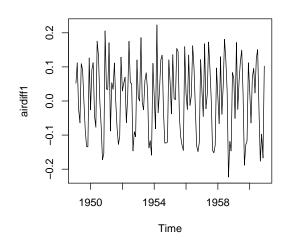
```
airdiff1 <- diff(air, differences=1)

plot(airdiff1)

acf(air, type = "correlation")</pre>
```

画出的一次差分时间序列图和ACF图如下所示: 用ARIMA(0,1,1)*(0,1,1)12拟

Series airdiff1



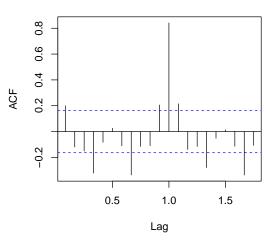


图 3: diff1

图 4: ACF

合对数化序列R语言代码为:

```
fit1<-arima(air, order=c(0,1,1),
seasonal=list(order=c(0,1,1),
period=12))
```

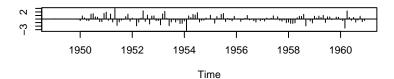
得出结果为: 对模型的自相关性和残差的正态性进行检验的R语言代码为:

```
plot(rstandard(fit1), type="o")
rs = rstandard(fit1)
shapiro.test(rs)
```

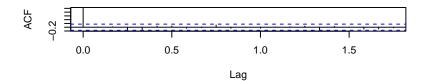
得出结果为: 求出的残差p值为: 0.1674¿0.05,因此残差具有正态性。 预测未来两年的取值R代码为:

```
airHW <-HoltWinters(air, beta=FALSE,
gamma=FALSE)
plot(airHW)
```

Standardized Residuals



ACF of Residuals



p values for Ljung-Box statistic

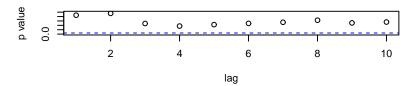
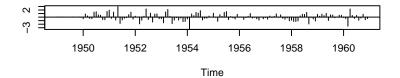
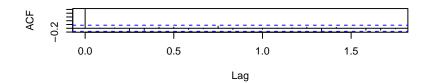


图 5: fit1

Standardized Residuals



ACF of Residuals



p values for Ljung-Box statistic

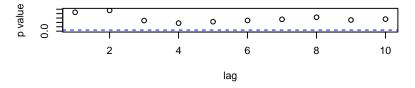


图 6: Residuals

```
library ("forecast")
airforecast = forecast (airHW, h=2)
plot (airforecast)
```

画出预测结果为:

Forecasts from HoltWinters

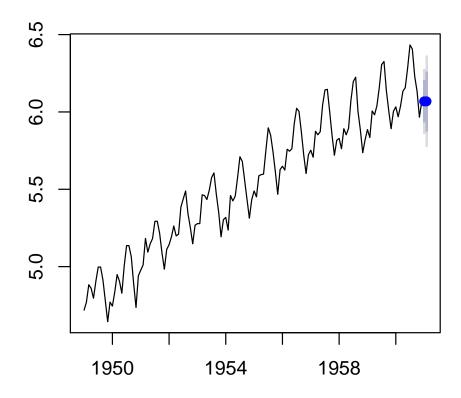


图 7: Airforecast

3.1) 运行以下代码:

```
jj = scan("http://www.stat.pitt.edu/stoffer/tsa2/data/jj.dat");

# read the data

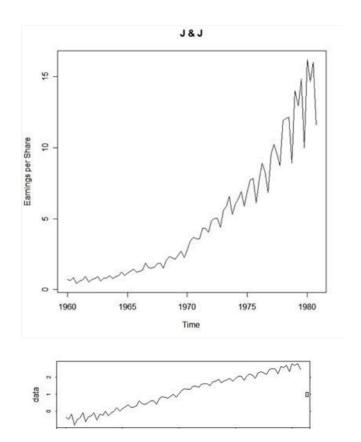
jj = ts(jj, start=1960, frequency=4) #ts()命令

plot(jj, ylab="Earnings_per_Share", main="J_&_J")

dljj = diff(log(jj))

# 做和差的处理log
```

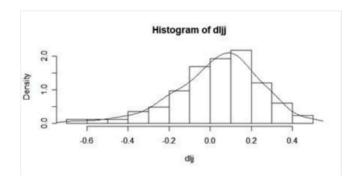
plot (dljj)

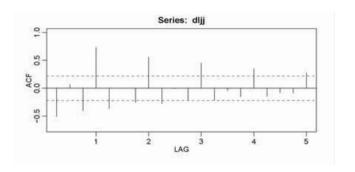


2) 运行以下代码:

```
29 dljj = diff(log(jj))
30 # 做和差的处理log
31 plot(dljj)
32 acf(dljj,20)
```

得以下时间序列图和ACF图表

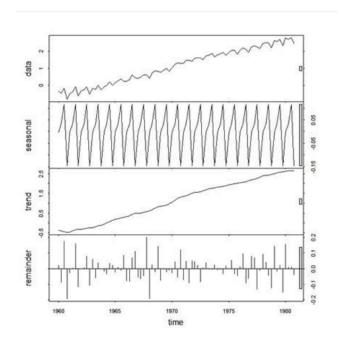




在ACF图中横坐标标记是1,2,3,4,5,但因为数据是季度性的,每年有四个季度,所以1,2,3,4,5的标记代表的是4,8,12,16,20的延迟。3)利用R中结构拆析命令,运行以下语句:

```
33  plot(dog <- stl(log(jj),"per"))
```

得如下图:



这其中第一行代表原来的log(jj)的数据。此数据可以看到总体的上升趋势还存在一定季节循环性的变化。第二行绘图代表拆析后季节循环的作用。第三行绘图代表拆析后将季节循环作用清除剩余的上升趋势,此数据清楚地看到那种循环性变化已经不存在,剩余的只是趋势。可以看出每一年第三季度会出现一个收入高峰,随之而来的第四季度收入就会跌入低谷,然后在一、二季度收入又会逐渐上升。4) 运行以下语句:

```
Q = factor(rep(1:4,21))

trend = time(jj)-1970

reg = lm(log(jj)~0+trend+Q, na.action=NULL)

summary(reg)

#
```

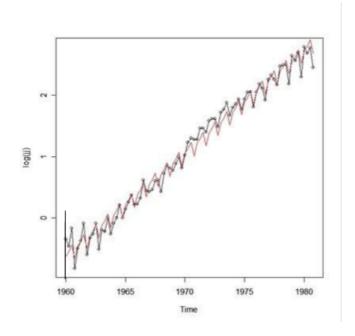
```
Residuals:
                         Median
      Min
                1Q
                                   3Q
                                           Max
40
   -0.29318~-0.09062~-0.01180~0.08460~0.27644
41
42
   Coefficients:
43
        Estimate Std. Error t value Pr(\$>|t|\$)
44
                      0.002259
                                  74.00
                                          <2e-16~***
   trend 0.167172
         1.052793
                      0.027359
                                  38.48
                                          <2e-16^****
   Q1
46
   Q2
          1.080916
                      0.027365
                                  39.50
                                          <2e-16^****
47
   Q3
          1.151024
                      0.027383
                                  42.03
                                          <2e-16^****
   Q4
          0.882266
                      0.027412
                                  32.19
                                          <2e-16^****
```

得下面等式:

Log(jj) = 0.167172*time + 1.052793*Q1 + 1.080916*Q2 + 1.151024*Q3 + 0.882266*Q4 绘图比较预测数据和实际数据:

```
plot(log(jj),type="o")
lines(fitted(reg),col=2)
```

得以下图:

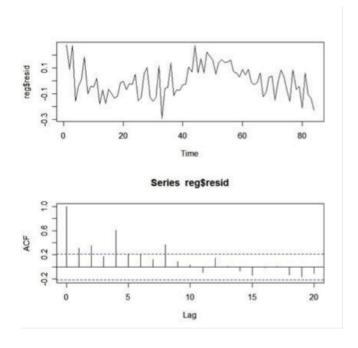


5) 运行如下语句:

```
_{52} \left| \begin{array}{c} \operatorname{par}\left( \operatorname{mfrow=c}\left( 2 , 1 \right) \right) \end{array} \right|
```

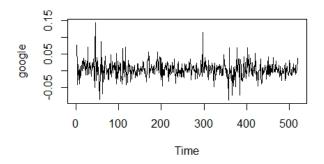
```
53 | plot(resid(reg))
54 | acf(resid(reg),20)
```

得以下图:



第一个图可以看到误差的变化范围在允许范围之内,并且没大变化。第二张图只有在lag0上有数值为1的关联性,其他lag上的关联性都非常小,即说明误差和误差之间是没关系的。

4.1)时间序列图如图所示。



2)要对序列的异方差性进行检验,可先将序列平方后进行白噪声检验,得到 $\chi^2 = 21.362, p-value = 3.802 \times 10^{-6}$

因此可认为存在异方差。

3) 由题意,建立GARCH(1,1)模型,得出

```
Error Analysis:

Estimate Std. Error t value Pr(>|t|)

mu 2.736e-03 9.190e-04 2.977 0.00291 **

ar1 6.249e-02 4.946e-02 1.263 0.20650

ar2 3.139e-03 4.808e-02 0.065 0.94795

ar3 -2.062e-02 4.651e-02 -0.443 0.65755

omega 5.236e-05 2.090e-05 2.505 0.01224 *

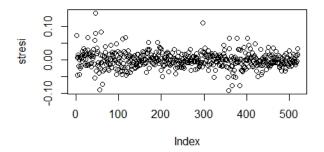
alphal 1.426e-01 4.589e-02 3.107 0.00189 **

betal 7.684e-01 6.697e-02 11.474 < 2e-16 ***

---

signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

对其进行残差分析, 可得出下图



对残差进行LB检验得出 $\chi^2=2.8863, p-value=0.08933$