

RDD and diff-in-diff

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Regression discontinuity design (RDD)

Basic assumption

RDD assumes a running variable (x) with a cut point (c) beyond which treatment is assigned (D).

$$D_i = \begin{cases} 1 & \text{if } x_i \geq c \\ 0 & \text{if } x_i < c \end{cases} \quad (1)$$

Distinction

It has a flavor of logit or propensity scores, but there are some differences:

- ▶ **logit** : x is not latent and we know the cutpoint: Both are observed and included as a *predictors*.
- ▶ **matching** : we have no control/treatment group. However, we assume that units on either side of the treatment are increasingly similar as their x is similar.

Examples

Administrative data are perfect: You have some rule that kicks in at a specific threshold for otherwise almost identical observations.

- ▶ school test scores on school admission, restrictions on class size
- ▶ legal drinking age on alcohol related deaths
- ▶ election of candidates in close races

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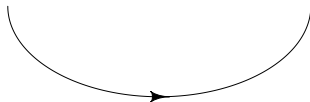
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Two ways of understanding RDD

X is a confounder
...so we only control for X

x \longrightarrow D \longrightarrow Y

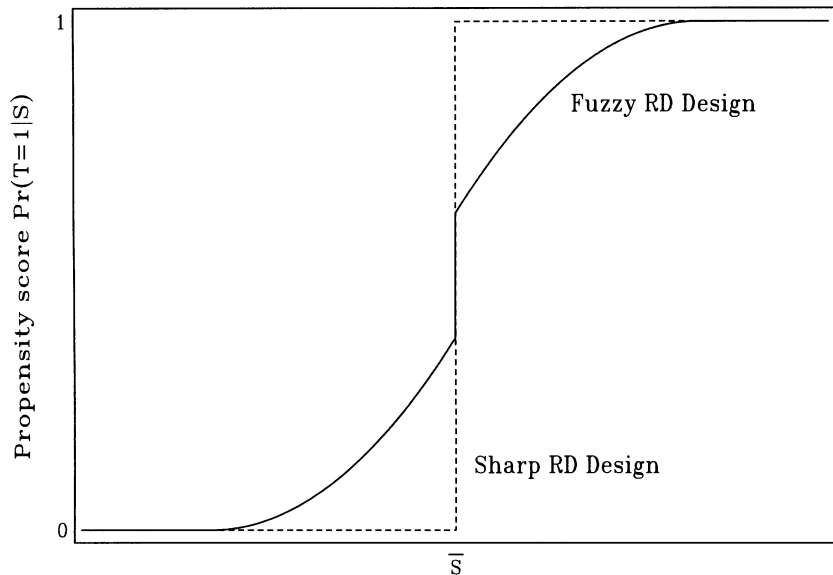


Two designs

We distinguish between two designs depending on how probable the treatment is:

- ▶ **sharp** RD: assignment is deterministic
- ▶ **fuzzy** RD: assignment is probabilistic

A visual representation



The basic model

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Do it in R

Check out alcohol related deaths (y) as a function of legal drinking age (D)

```
load("df.Rda")  
mod1 <- lm(alcohol ~ over21 + age,  
           df)
```

Is that all?

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2. **no omitted variable bias:** x must capture all influence on D .

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... that's not a regression discontinuity.

Ensuring linear effect

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- ▶ recode the $x \rightarrow$ compare with the recoding of y in GLMs.
- ▶ consider a sufficiently small window

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```
df$age2 <- df$age^2
mod2 <- lm(alcchol ~ over21 + age + age2,
           df)
```


Recode the x: symmetric effect

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$$y_i = \alpha + \rho D_i + \gamma_1 x_i + \gamma_2 x_i^2 \quad (4)$$

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df$age2 <- df$age^2  
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```

\Rightarrow Here, x has a symmetrical effect on both sides of the treatment.

Recode the x: assymetric effect

We can assume x has different effects on each side of the treatment

$$y_i = \alpha + \rho D_i + \gamma x_i + \delta x_i D_i \quad (5)$$

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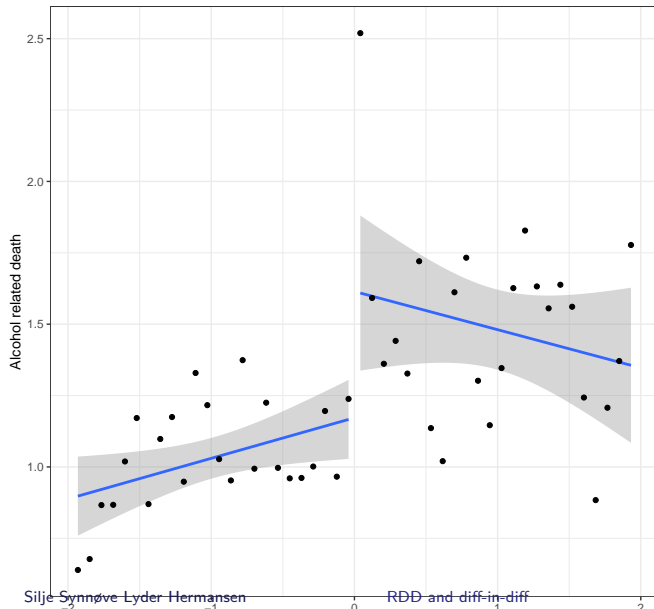
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$$y_i = \alpha + \rho D_i + \gamma x_i + \delta x_i D_i \quad (5)$$

⇒ we center the x on the cutpoint ($x_i - c$) → ρ still reports the change at the cutpoint.

```
mod3 <- lm(alcohol ~ over21 * age,  
           df)
```

Do it in R

Legal drinking age 21 ($x = 0$) and alcohol related deaths

Extrapolation

We do this to estimate the effect at the cutpoint (ρ)

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- ▶ but we can also extrapolate y beyond the cutpoint with x : $\rho + \delta(x - c)$

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- ▶ recoding the x is a **the parametric approach**.
- ▶ subsetting the data to tweak the window around the cutpoint is a **non-parametric approach**.

Bandwidth : the idea

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- ▶ There's a tradeoff between linearity and statistical power (we need sufficient N).

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```
mod5 <- lm(alcohol ~ over21 * age,  
           df[df$agecell > 20 & df$agecell < 22,])
```

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```
mod5 <- lm(alcohol ~ over21 * age,  
           df[df$agecell > 20 & df$agecell < 22,])
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⇒ *When you narrow down, do you get a weaker or stronger effect?*

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- ▶ We can do it by hand
- ▶ ...or we can make an algorithm do it:
 - ▶ run a local weighted regression line
 - ▶ bandwidth is estimated accordingly

⇒ *the point is to show robustness, not p-hack!*

Omitted variable bias

Fuzzy RD

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Often the D increases the probability of a treatment, but we don't know!

⇒ This is a Instrumental Variable approach (more on Thursday)

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$\Rightarrow \gamma$ is the causal effect of D in a fuzzy design.

Differences-in-differences

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⇒ *Treatment effect is that difference*

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⇒ *based on panel data (units are observed several times).*

Example: States' monetary policy and number of banks

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⇒ *What was the treatment effect?*

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⇒ *Basically a 2-by-2 table*

How to do it?

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- ▶ T represents the treatment group: differences *between* units
- ▶ β_3 is the causal effect

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- ▶ Requires panel data → which means correcting the standard errors.
- ▶ Common panel types: state-year/administrative unit-time period; people over time ...

⇒ *we want to know the trend before and after the break*

Another example: drinking age and death

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Does the legal drinking age has an effect on death rates among the young?

- ▶ y is number of deaths per 100 000
- ▶ P is post-treatment dummy
- ▶ T is dummies for states
- ▶ $trend$ is year dummies

Another example: step 1 → calculate differences

The authors have two tricks:

- ▶ Hardcode the interaction effect (dummy before/after treatment)
- ▶ They remove the intercept to retain all dummies

```
load("df2.Rda")  
##with intercept  
mod <- lm(mrate ~ legal +  
          state +  
          year_fct,  
          df)  
##without intercept; with all dummies  
mod <- lm(mrate ~ 0 +  
          legal +  
          state +  
          year_fct,  
          df)
```

Another example: step 2 → calculate errors

Calculate robust standard errors:

```
library(clusterSandwich)

## Registered S3 method overwritten by 'clusterSandwich':
##   method      from
##   bread.mlm sandwich

vcov <- vcovCR(mod, cluster = df[["state"]],
               type = "CR2")
robust <- coef_test(mod, vcov = vcov)$SE
```

Another example: step 3 → interpretation

Display the results and interpret:

```
library(stargazer)
stargazer(mod, se = robust,
           omit = "state|year",
           type = "html",
           out = "regtable.html")
```


Another example: step 3 → interpretation

Table 1: Death rates among young as a function of legal drinking age

	<i>Dependent variable:</i>
	mrate
Legal drinking age (causal effect)	10.804** (4.479)
Observations	714
R ²	0.986
Adjusted R ²	0.985
Residual Std. Error	17.339 (df = 649)
F Statistic	726.005*** (df = 65; 649)
<i>Note:</i>	
*p<0.1; **p<0.05; ***p<0.01	

⇒ *What did we find?*

The parallel trends assumption

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⇒ *We do that with an interaction effect!*

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⇒ *There's a tradeoff: treatment is at the unit level, statistical power at the subunit level.*