

# Multilevel/hierarchical models: Overview

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Where are we in the course?

## Recap from Monday

**When observations are not i.i.d. (i.e. they share a group identity), we will often consider alternatives to the ordinary linear model**

- ▶ negative take: the assumptions of the linear model are not met.
  - ▶ non-normal residuals,
  - ▶ heteroscedastic residuals
  - ▶ correlation between  $x$  and residuals
- ▶ positive take: we have variation that we want to leverage strategically
  - ▶ within-group variation
  - ▶ between-group variation
  - ▶ more correct estimation of the standard errors

⇒ *see this as an opportunity*

## I pick my models as part of my research design

### **What are the most relevant correlations/variation given my theory?**

- ▶ in experiments: you can create that variation and randomize the rest (cut out confounders)
- ▶ in observational studies: you'll have to “hunt” for the variation you want and control away the rest

# Confounders

- ▶ Control variables that – if absent lead to omitted variable bias – satisfy three criteria:
  - ▶  $z$  correlates with  $y$
  - ▶  $z$  correlates with  $x$
  - ▶  $z$  causes  $x$  and  $y$  (not intermediate/post-treatment)

→ *even when 3 is not satisfied, it might be a sign of a common group identity (e.g. nationality)*
- ▶ Group identities: observations done in the same context share many potential confounders
  - ▶ you might kill several birds with one stone

## The principle

# The principle

**We make the assumption that the residuals are drawn from a normal distribution**

- ▶ **pooled models:** a single distribution

$$y_i = a + bx_i + \epsilon_i$$

$$\epsilon_i \sim N(0, \sigma^2)$$

- ▶ **hierarchical models:** add a hierarchy
  - ▶ assume groups are drawn from different distributions
  - ▶ their mean is drawn from a single distribution that “rules them all”

$$y_i = a + bx_i + \epsilon_{ji}$$

$$\epsilon_j \sim N(\alpha_j, \sigma_j^2)$$

$$\alpha_j \sim N(0, \sigma_\alpha^2)$$

# Untangling the parameters/variation

**This allows me to untangle different sources of variation**

$$y_i = a + bx_i + \epsilon_{ji}$$

$$\epsilon_j \sim N(\alpha_j, \sigma_j^2)$$

$$\alpha_j \sim N(0, \sigma_\alpha^2)$$

- ▶  $\alpha_j^2$ : grouped mean of residuals: group intercept
- ▶  $\sigma_\alpha^2$ : between-group variation
- ▶  $\sigma_j^2$ : group-level (within) variation



# The promises of a hierarchical structure

## This allows me to leverage different sources of variation

- ▶ leverage **within-group** variation:
  - ▶ by factoring out/control for between-group variation ( $\sigma_j^2$ )
- ▶ leverage **between-group** variation:
  - ▶ by running a second regression on the group means ( $\alpha_\alpha^2$ )
    - ▶ adjusts the standard errors
    - ▶ data augmentation: add variables from other sources that vary by group
    - ▶ predict out of sample even for new groups
- ▶ leverage **both sources** of variation
  - ▶ by borrowing from the more informative variation (“pooling”/“shrinkage”)

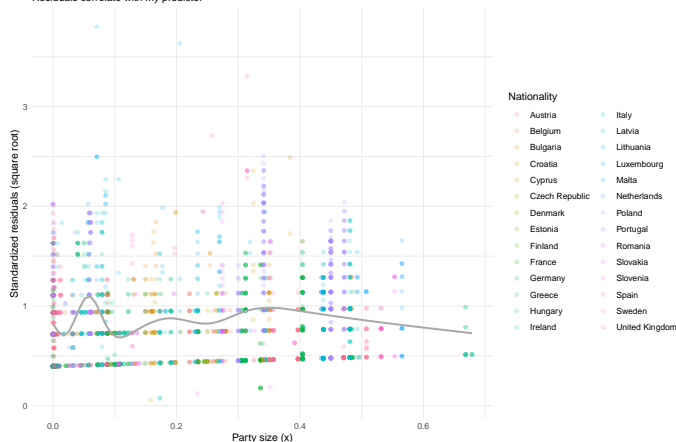
## Labeling the errors: grouped residuals

# Labeling the errors: grouped residuals

**Our residuals have group identities that we can “label” as such.**

The ghosts of our regression

Residuals correlate with my predictor

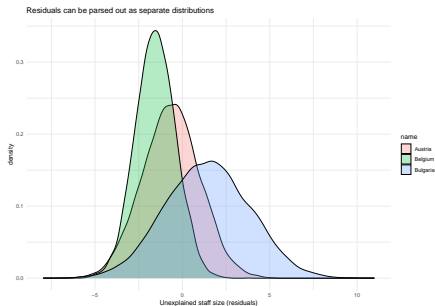


## Group means and group-level variation

**Our residuals have group identities that we can “label” as such.**

- ▶ each group of residuals has a distribution with a mean and a spread

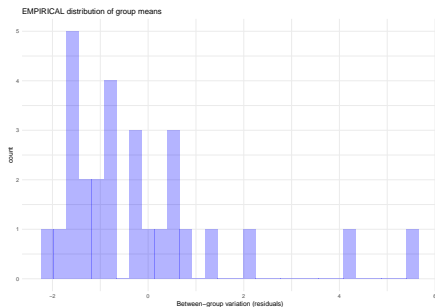
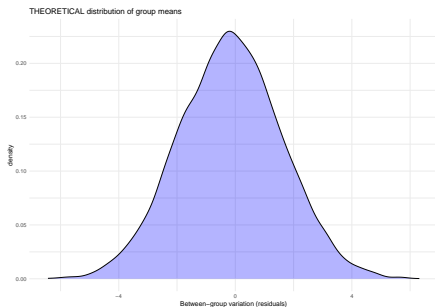
```
## # A tibble: 28 x 3
##   Nationality y_bar_j sigma2_alpha
##   <chr>      <dbl>      <dbl>
## 1 Austria    -0.665      1.65
## 2 Belgium   -1.54       1.15
## 3 Bulgaria    1.41       2.44
## 4 Croatia    0.549      4.10
## 5 Cyprus    -0.253      1.89
## 6 Czech Republic -0.206      1.84
## 7 Denmark    -1.48      1.30
## 8 Estonia    -1.33      0.950
## 9 Finland    -1.47      0.919
## 10 France    -1.11      1.26
## # i 18 more rows
```



*⇒ I can reconstruct their theoretical distribution by calculating the group mean and standard deviation*

## Between-group variation

**The group means are drawn from a common normal distribution with a mean and a spread**



*⇒ I am treating the residuals as if they were a variable, so statistical theory can be applied*

## Varying-intercepts regression: within-group variation

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## **The random/varying-intercept model:**

- ▶ a common slope for all predictors
- ▶ separate intercepts for all group identities
- ▶ a common intercept (grand mean)

## From labelled errors to varying intercepts

**Instead of hiding the groupings in the residuals, we can report them as a series of intercepts (i.e. report their group means)**

$$y_i = a + bx_i + \alpha_{ji}$$

$$\alpha_j \sim N(0, \sigma_\alpha^2)$$

- ▶  $a$ : the **grand mean** (mean of  $\alpha$  means)
- ▶  $\alpha_j$ : **varying intercepts** (deviations from this grand mean)

$\Rightarrow$  *useful for interpretation in R*



## Varying-intercepts

**Now, it is clear that I parse out (control for) between-group variation**

- ▶ **within-group variation** the b coefficients report the effect of observation-level variables
- ▶ **group-level variation** is reported in the varying intercepts, it is the variation that:
  - ▶ has not been accounted for by my main effects
  - ▶ that can be attributed to group identities

## Estimation in R: Varying national intercepts

# Estimation in R: Varying national intercepts

Let's regress MEPs' investment in their district (y) on...

- ▶ x: their party's size in the national parliament (as a proxy for state funding).
- ▶ ... while controlling away between-national variation

## Equation:

$$\text{Staff size} = a + b \times \text{Party size} + \alpha_{\text{Nationality}}$$

$$y_i = a + bx_i + \alpha_{ij}$$

## Estimation:

```
library(lme4)
mod.ran.int <- lmer(y ~ x + (1|Nationality),
  df)
```

## Reading the R output

# Reading the R output

```
summary(mod.ran.int)
```

```
## Linear mixed model fit by REML ['lmerMod']
## Formula: y ~ x + (1 | Nationality)
## Data: df
##
## REML criterion at convergence: 31355.2
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.1127 -0.5387 -0.1435  0.3598 15.2357
##
## Random effects:
## Groups      Name                Variance Std.Dev.
## Nationality (Intercept) 3.125      1.768
## Residual              5.240      2.289
## Number of obs: 6948, groups: Nationality, 28
##
## Fixed effects:
##              Estimate Std. Error t value
## (Intercept)  2.6799    0.3386    7.915
## x            -1.6722    0.1678   -9.965
##
## Correlation of Fixed Effects:
##      (Intr)
## x -0.117
```

R refers to the residuals as “random effects”

$\sigma^2_{\alpha}$ : remaining **between-group variance**: 3.12

- ▶ standard deviation: 1.77
- ▶ the unexplained variation between groups

Residual: remaining **within-group** variance: 5.24

- ▶ standard deviation of within-group distribution: 2.29
- ▶ the unexplained variation within all groups

R refers to regression coefficients as “fixed effects”

a: intercept/**grand mean**: 2.68

- ▶ a hypothetical intercept for interpretation (mean of means)

b: **slope**: -1.67

- ▶ the marginal effect of party size (x)

## Interpretation

# Interpretation

Interpretation follows normal principles, but there are some complications:

- a. we now have two intercepts per scenario:
  - ▶ the grand mean ( $\alpha$ ): for focus on general effect of  $x$
  - ▶ the group-level mean ( $\alpha_j$ ): for description and prediction
  - ▶ the grand mean ( $\alpha$ ): for focus on general effect of  $x$
  - ▶ sum of the grand mean ( $\alpha$ ) group-level mean ( $\alpha_j$ ): for prediction
- b. all effects are linear
  - ▶ so first-difference and marginal effects are the same

## Interpreting marginal effects

**The interpretation of the marginal effect is as with any linear model:**

Table 1: Effect of state funding for parties on MEPs' local staff size

<i>Dependent variable:</i>	
	<i>y</i>
x	-1.672*** (0.168)
Constant	2.680*** (0.339)
Observations	6,948
Log Likelihood	-15,677.610
Akaike Inf. Crit.	31,363.210
Bayesian Inf. Crit.	31,390.600
Note: * $p < 0.1$ ; ** $p < 0.05$ ; *** $p < 0.01$	

⇒ *A 10% decrease in the national party's seat share would lead every 6th MEP to compensate by hiring an additional local staffer.*



# Prediction

## The varying intercepts are reported as deviations from the grand mean

```
fixef(mod.ran.int); ranef(mod.ran.int)
```

```
## (Intercept)      x
##      2.679887  -1.672226

##      (Intercept)
## Austria      -0.49518857
## Belgium     -1.52249566
## Bulgaria      1.54657524
## Croatia       0.68267309
## Cyprus       -0.05313986
## Czech Republic -0.10587832
```

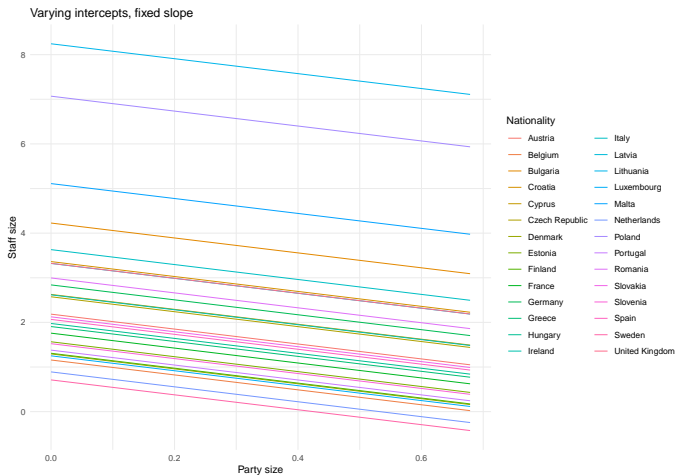
Predicted local staff in Austria when national party is not in Parliament:

$$\blacktriangleright 2.68 + -0.5 \times 0 = 2.18$$

Predicted local staff in Austria when national party holds 10% of the seats

$$\blacktriangleright 2.68 + -0.5 + -1.67 \times 0.1 = 2.02$$

# Visualization



⇒ *the slope is constant, but the intercept changes across nationalities*

## Varying slopes, varying intercepts

# Defintion

**We can let the effect of x vary by group through an interaction effect**

$$y_i = a + bx_i + c_jz_i + \alpha_{ji}$$

- ▶  $c_j$ : **varying slope** (the effect of z varies by group)
- ▶  $\alpha_j$ : **varying intercepts**

$\Rightarrow$  *a series of regressions within the regression*

# Estimation in R

- ▶ the estimation is done as if it was an interaction effect
- ▶ the R syntax is somewhat different

```
mod.ran.slope <- lmer(y ~ x + (ProxNatElection | Nationality),
```

## Interpretation

# Marginal effects

**We can read these coefficients as if they were from separate models**

```
ranef(mod.ran.slope)
```

##	(Intercept)	ProxNatElection
## Austria	-0.3201686	0.001073577
## Belgium	-1.3944093	-0.018299157
## Bulgaria	1.8027887	0.086786759
## Croatia	0.9352286	0.058233060
## Cyprus	0.1020429	-0.003477477
## Czech Republic	0.1112017	0.024050889

MEPs from Austria hire on average 0.004 assistants more immediately before an election compared to immediately after, while MEPs from Belgium hire on average 0.073 fewer assistants.

► These are negligible marginal effects.

# Prediction

## The prediction is done per group, but follows normal rules

- ▶ two intercepts:  
grand mean + group-level intercept
- ▶ one slope per group

```
fixef(mod.ran.slope); ranef(mod.ran.slope)
```

```
## (Intercept)      x
##    2.513035  -1.691321
```

```
##          (Intercept) ProxNatElection
## Austria      -0.3201686      0.001073577
## Belgium     -1.3944093     -0.018299157
## Bulgaria      1.8027887      0.086786759
## Croatia       0.9352286      0.058233060
## Cyprus        0.1020429     -0.003477477
## Czech Republic 0.1112017      0.024050889
```

Austria after election:

$$\text{▶ } 2.51 + -0.32 + 0.001 \times -4 = 2.189$$

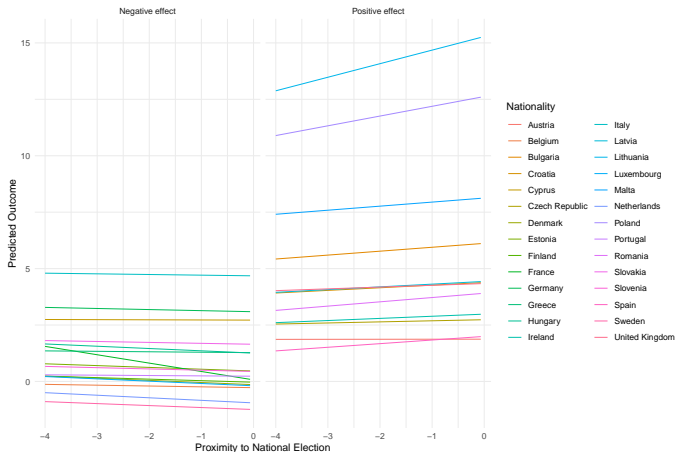
Austria before election:

$$\text{▶ } 2.51 + -0.32 + 0.001 \times 0 = 2.193$$



# Visualization

Varying slopes and intercepts: Effect of electoral calendar by nationality



## Level-2 regression: between-group variation

## Definition

# Definition

**We can think of the residuals/group intercepts as a variable in their own right**

$$y_i = a + bx_i + \epsilon_{ji}$$

- ▶ they are generated by draws from J number of distributions:  $\epsilon_{ji} \sim N(\alpha_j, \sigma_\alpha^2)$
- ▶ ... and therefore we can model them

$$\alpha_j = c_1 + c_2 Z_j$$

*⇒ we run a second regression on the residuals*

# Implications

## We explicitly model between-group variation

- ▶  $z$ , the level-2 predictor only varies at the group level
  - ▶ standard errors for  $z$  reflect the number of groups
  - ▶ the more groups, the more the approach makes sense
- ▶ data augmentation
  - ▶ we can add information from other to the model
  - ▶ contextual elements
  - ▶ improves prediction

## Estimation in R: Electoral system

# Estimation in R: Electoral system

## Let's add electoral system (z) as a predictor

- ▶ it never changes in a country (in this study)

## R handles this automatically

- ▶ same data frame
  - ▶ all variables that don't vary within groups are regressed as a level 2
- ▶ coefficients reported the same way
- ▶ estimation of coefficients and standard errors is different

```
mod.two.levels <- lmer(y ~ x + z + (1|Nationality), df)
```

## Reading the R output



# Reading the R output

```
## Linear mixed model fit by REML ['lmerMod']
## Formula: y ~ x + z + (1 | Nationality)
##      Data: df
##
## REML criterion at convergence: 31353.9
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.1145 -0.5388 -0.1434  0.3599 15.2339
##
## Random effects:
##      Groups      Name      Variance Std.Dev.
## Nationality (Intercept) 3.235    1.799
## Residual              5.240    2.289
## Number of obs: 6948, groups: Nationality, 28
##
## Fixed effects:
##              Estimate Std. Error t value
## (Intercept)  2.5268    0.6030    4.191
## x            -1.6719    0.1678   -9.962
## z             0.2263    0.7311    0.310
##
## Correlation of Fixed Effects:
##      (Intr) x
## x -0.077
## z -0.821  0.013
```

The level-2 regression coefficient appears as “fixed effects”

c: **slope:** 0.23

- ▶ the marginal effect of electoral system (z)

Check the **change in between-group variance**:

- ▶ the between-group variance ( $\sigma_\alpha^2$ , 3.23) should decrease
- ▶ it is not the case here ( $3.12 \leq 3.23$ )

→ *increase in variance indicates “complexities” between levels (interactions)*

Correlation of Fixed Effects:

- ▶ negative correlation between predictor (z) and intercept (-0.82): high level of z correlates with low base-line value of y.