

# Models of outcome and choice: The logit model

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## Let's touch base

**We will be using mentimeter (menti.com) to communicate interactively.**

- ▶ answer questions on [www.menti.com](https://www.menti.com) using the access code 8471 19241
- ▶ results show on screen

⇒ *Relax, your answers are anonymous!*

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## Section 2

### A latent variable approach to GLMs

## Many outcomes are not continuous

**OLS assumes a continuous dependent variable. But many phenomena in the social sciences are not like that.**

- ▶ Vote choice, civil conflict onset, legislator performance, court rulings, time to compliance, etc.
- ▶ What phenomena are you interested in?

⇒ *OK. Let's strategize.*

## All regressions are linear(ized)

**The basic formulation in any regression describes a linear relationship between  $x_i$  and  $y_i$ :**

$$y_i = \alpha + \beta x_i + \epsilon_i \quad (1)$$

- ▶ When  $x_i$  increases with one unit,  $y_i$  increases with  $\beta$  units.
- ▶ If that relationship is not linear, we have to make it so:
  - ▶ by recoding the  $x_i$
  - ▶ by recoding the  $y_i \rightarrow$  we *linearize*.

# A latent variable

## **A linear(ized) model requires a continuous dependent variable.**

- ▶ Imagine we are interested in unobservable variable,  $z_i$ , that describes our propensity towards something.
  - ▶ Above a certain threshold ( $\tau$ ) of  $z_i$ , observability kicks in and we can see  $y_i$ .
  - ▶ The regression coefficients ( $\beta$ ) in GLMs describe that relationship.
- ⇒ The latent variable approach is useful when interpreting the results.



## Example: The binomial model

**The logit model is a perfect example:**

$$y_i = \begin{cases} 1 & \Leftrightarrow z_i > \tau \\ 0 & \Leftrightarrow z_i \leq \tau \end{cases} \quad (2)$$

- ▶ The probability ( $z_i$ ) of an outcome  $y_i$  is continuous.
- ▶ Above a certain probability ( $\tau$ ), we observe a positive outcome ( $y_i = 1$ ).

$\Rightarrow$  *but how do we set the value of  $\tau$ ?*

# From latent variable to discrete outcomes

**Statistical theory helps us describe how  $z_i$  leads to  $y_i$ .**

- ▶ What kind of process generated our data? → data generating process (DGP)
- ▶ How can we best describe it? → choice of *probability distribution* (in GLM)

# The three components of GLMs

**When fitting the model, we need to make three choices:**

- ▶ A linear predictor:  $\beta x_i$ .
- ▶ A probability distribution: they're all in the exponential family
- ▶ A recoding strategy

# The three components of GLMs

**In R this translates to two additional arguments compared to your usual OLS.**

- ▶ A linear predictor:  $\rightarrow (y \sim x)$ .
- ▶ A probability distribution:  $\rightarrow (\text{family} = )$
- ▶ A recoding strategy  $\rightarrow (\text{link} = )$ .

# Latent variable approach for interpretation

- ▶ The latent variable approach is useful when interpreting results.
- ▶ That's when we map *from* the latent variable *to* the observed outcome.

⇒ *When estimating the model, we have to go the other way 'round.*

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## Section 3

Recoding: How do we get from a binary to a continuous variable?

# Data structure

**We can only observe the outcome produced by the latent variable.  
There are two data structures for binary data:**

- ▶ classes of observations: e.g.: rats in a cage, coin tosses...
- ▶ case-based: e.g.: legislator votes, Brexit...



# Data structure

**We can only observe the outcome produced by the latent variable.**  
**There are two data structures for binary data:**

- ▶ classes of observations: e.g.: rats in a cage, coin tosses... → *the closest to the latent continuous variable.*
- ▶ case-based: e.g.: legislator votes, Brexit...

⇒ *we know the number of successes and trials in a cage/class/stratum.*  
*That's our starting point.*

## Let's start with the odds

**Despite binary outcomes, we want a continuous variable that is unbounded at both ends. We define a stratum and start comparing:**

- ▶ Odds: Compare number of successes with number of failures within a stratum → *continuous but highly skewed*.
- ▶ Logtransform the odds → *continuous and bell shaped*.

## Let's exemplify with rats

**We kept a 1000 rats in a cage and a number of them died (failure) while others are still alive (success). How can we model this?**

## We calculate the odds

### **We calculate the odds of surviving in a cage in a 1000 cages**

- ▶ Let's consider a series of 1000 trials where we let the successes go from complete failure (success = 0) to complete success (success = 1000)

```
success = 0:1000
tries = 1000
#remember: failure = tries - success
odds <- success/(tries - success)

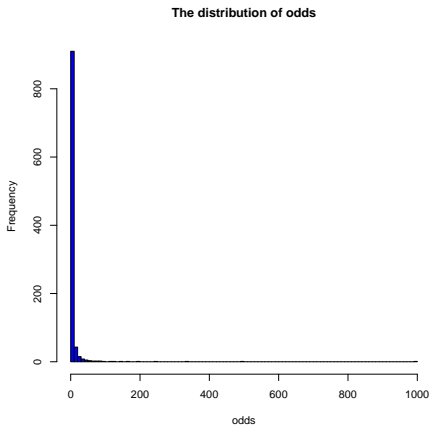
hist(odds, breaks = 100, col = "blue")

hist(log(odds), breaks = 101, col = "blue")

plot(log(odds), success, type = "l")
```

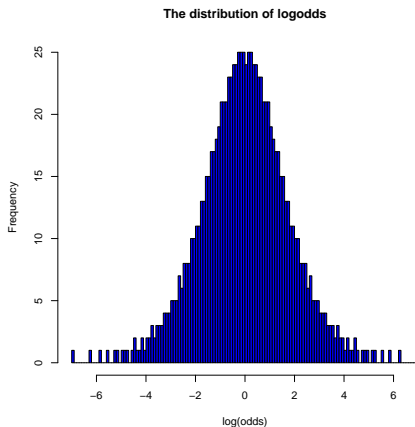
## Let's start with the odds

**We get a continuous but skewed variable.**



Now, let's logtransform the odds

**We get a nice, bellshaped curve.**



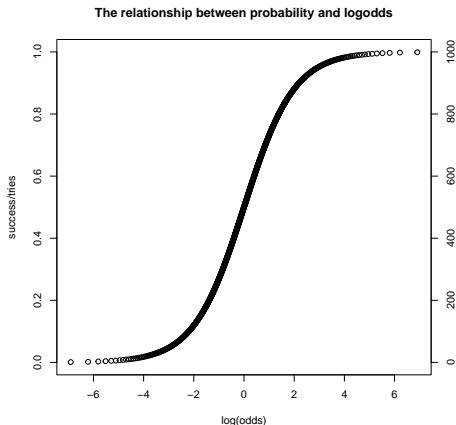
Now, let's logtransform the odds

**This, we can run regressions on!**



# The famous S shape

We can plot the logodds of success against the number of successes or their probability (it's the same).



# Probability distributions for binary variables

**There are two, closely related probability distributions for binary outcomes:**

- ▶ The binomial distribution:  $B(n, p)$ 
  - ▶  $p$  is the probability of success tells where on the x-axis (trials) the distribution is placed.
  - ▶  $n$  is the number of trials and defines the precision (width) of the distribution.
- ▶ The Bernoulli distribution:  $Ber(p)$ : when we only have only one trial.

## Subsection 2

Why all the fuzz? Why not OLS?

# Distributions in OLS and maximum likelihood

- ▶ In OLS: The residuals must be normally distributed (but not the  $y_i$ )
- ▶ In ML: The  $z_i$  must follow a known probability distribution.

⇒ *This what allows us to translate the latent variable to outcomes.*

# What happens if I run a linear model on binary outcomes?

- ▶ The model predicts out of the possible boundaries
  - ▶ Predictions are wrong.
  - ▶ Regression coefficients are wrong.
  - ▶ Standard errors are wrong.
- ▶ The relationship between  $x_i$  and  $y_i$  is constant across all values.

⇒ *This last element has a bearing for the interpretation.*

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## Section 4

Interpretation: So... what did I find?

### Subsection 1

Back and forth: Logistic and logit transformation

# The logit transformation

**When we go from outcomes to latent variable we use the logit transformation.**

$$\text{logit}(p) = \log\left(\frac{p}{1-p}\right) \quad (3)$$

*⇒ This what R does when estimating our model*



# The logistic transformation

**When we go from the latent variable to outcomes we use the logistic transformation.**

$$\text{logit}^{-1}(\text{logodds}) = \frac{\exp(\text{logodds})}{1 + \exp(\text{logodds})} = \frac{1}{1 + \exp(-\text{logodds})} \quad (4)$$

*⇒ This what we do when interpreting our model*

# My three stages of interpretation

## I go through three stages of interpretation

- ▶ Inspect the marginal effects from regression table
  - ▶ Logodds: check direction and significance.
  - ▶ Odds ratio (for large coefficients) and percentage change (for smaller coefficients).
- ▶ Formulate scenarios using point estimates (in text)
- ▶ Formulate more scenarios with uncertainty using graphics.

```
load("MEP2016.rda")
df <- MEP2016

mod <- glm(PoolsLocal ~
            OpenList +
            SeatsNatPal.prop +
            LaborCost,
            family = binomial(link = "logit"),
            df)

stargazer::stargazer(mod,
                      # label = "tab:regression",
                      title = "MEPs' propensity to share local
                      out = "results_table.tex",
                      type = "latex")
```

# The regression table: marginal effects

## I interpret the regression coefficient itself

- ▶ Change in logodds: check direction and significance.
- ▶ Odds ratio (for large coefficients) and percentage change (for smaller coefficients).

⇒ *A first stab at hypothesis testing.*

## The regression table: marginal effects

**Now, you try!** What statements would you make using the change in logodds, the odds ratio and percentage change?

**Table:** MEPs' propensity to share local assistants (a binomial logit)

	<i>Dependent variable:</i>
	PoolsLocal
OpenList	-1.124*** (0.181)
SeatsNatPal.prop	-1.930*** (0.527)
LaborCost	0.056*** (0.009)
Constant	-1.094*** (0.286)
Observations	686
Log Likelihood	-392.832
Akaike Inf. Crit.	793.665
<i>Note:</i> * p<0.1; ** p<0.05; *** p<0.01	

```
#Change in logodds for MEPs in candidate-centered systems
mod$coefficients[2]

##      OpenList
## -1.124427

# Odds ratio: <1 is negative; > 1 is positive
exp(mod$coefficients[2])

##      OpenList
## 0.3248387

# Percentage change
(exp(mod$coefficients[2]) - 1)*100

##      OpenList
## -67.51613
```

# The regression table: marginal effects

## Typical statements about marginal effects

- ▶ Change in logodds: "MEPs from candidate-centered systems are less likely to share local assistants. Both effects are statistically significant."
- ▶ Percentage change (for smaller coefficients; -1.93)." The likelihood that an MEP shares a local assistant with a party colleague is 68% lower when they compete in a candidate-centered system compared to those that compete in party-centered systems."

⇒ *A first stab at hypothesis testing.*

## Predicted values

**If you believe the model describes reality appropriately, you can learn more about it by interpreting more thoroughly**

- ▶ Odds ratios are notoriously hard to understand.
- ▶ The effect depends on the value of  $y_i$  and all the other  $x$ s.

⇒ *Interpret the predicted values*



## Predicted point estimates (text)

### Formulate scenarios using point estimates (in text)

- ▶ Take an all-else-equal approach: Let one  $x$  change and keep all others constant (on a typical value).
- ▶ Find the typical representative of two  $x$  values and set the other  $x$ s accordingly.

⇒ *Which one you use depends on your objective: A theoretical point, assess effect of intervention on groups...*

## Predicted point estimates/first difference (text)

**Now you try!** What is the predicted effect of changing electoral system on MEPs' propensity to share local assistants ...

- ▶ In Bulgaria (Labor cost == 4.4); when the party is small (Seat share == 0.1).
- ▶ In Denmark (Labor cost == 42); when the party is small (Seat share == 0.1).
- ▶ Is this a realistic set of scenarios?

⇒ *Compare the two predicted probabilities for each pairs of scenarios.*

- ▶ Go to Padlet to provide your answer:  
(<https://padlet.com/siljesynnove/logit>)

```
dfp <- df %>% select(PoolsLocal,  
                     SeatsNatPal.prop,  
                     OpenList,  
                     VoteShare_LastElection,  
                     LaborCost)
```

```
## Error in df %>% select(PoolsLocal, SeatsNatPal.prop,  
OpenList, VoteShare_LastElection, : could not find  
function "%>%"
```

```
stargazer::stargazer(dfp,  
                     title = "Descriptive statistics",  
                     out = "desc_table.tex")
```

```
## Error in .stargazer.wrap(..., type = type, title =  
title, style = style, : object 'dfp' not found
```

```
##Bulgaria; party-centered  
logodds1 <- mod$coefficients[1] * 1 + mod$coefficients[2] * 0 +
```

```

  mod$coefficients[3] * 0.1 + mod$coefficients[4] * 4.4
prob1 <- exp(logodds1)/(1+exp(logodds1))
prob1

## (Intercept)
##    0.2610522

##Bulgaria; candidate-centered
logodds2 <- mod$coefficients[1] * 1 + mod$coefficients[2] * 1 +
  mod$coefficients[3] * 0.1 + mod$coefficients[4] * 4.4
prob2 <- exp(logodds2)/(1+exp(logodds2))
prob2

## (Intercept)
##    0.102944

#First difference
prob2 - prob1

## (Intercept)
##   -0.1581082

#Denmark; party-centered
logodds3 <- mod$coefficients[1] * 1 + mod$coefficients[2] * 0 +
  mod$coefficients[3] * 0.1 + mod$coefficients[4] * 42
prob3 <- exp(logodds3)/(1+exp(logodds3))
prob3

## (Intercept)
##    0.7430194

#Denmark; candidate-centered
logodds4 <- mod$coefficients[1] * 1 + mod$coefficients[2] * 1 +
  mod$coefficients[3] * 0.1 + mod$coefficients[4] * 42
prob4 <- exp(logodds4)/(1+exp(logodds4))
prob4

```

```
## (Intercept)
## 0.4843289
```

```
#First difference
prob4 - prob3
```

```
## (Intercept)
## -0.2586905
```

```
#Alternative
```

```
newdata <- data.frame(OpenList = c(0, 1, 0, 1),
                        SeatsNatPal.prop = 0.1,
                        LaborCost = c(4.4, 4.4, 42, 42))
```

```
preds <- predict(mod, newdata, type = "response")
preds
```

```
##           1           2           3           4
## 0.2610522 0.1029440 0.7430194 0.4843289
```

```
diff(preds[1:2]); diff(preds[3:4])
```

```
##          2
## -0.1581082
##          4
## -0.2586905
```

## Predicted point estimates (text)

**Notice how the absolute effect of the electoral system changes!**

- ▶ **Marginal effect:** MEPs' likelihood of sharing assistants decreases by 68% % when we change electoral system. → *holds for all values of  $x$ .*
- ▶ **First difference (scenario 1a and b):** We see that changing electoral system when labor cost is *low* corresponds to a predicted 16 percentage points shift in likelihood of sharing assistants (from 26 percentage points to 10 percentage points).
- ▶ **First difference (scenario 2a and b):** We see that changing electoral system when labor cost is *high* corresponds to a predicted 26 percentage points shift in likelihood of sharing assistants (from 74 percentage points to 48 percentage points).

⇒ *This is an implicit interaction effect.*

## Predicted values (graphic)

### **Formulate scenarios using point estimates and put them on speed**

- ▶ Predict  $y$  values for the entire range of  $x$  and plot it.
- ▶ Simulate confidence and plot that too.
- ▶ You can do this for two scenarios.

⇒ *You get a sense of the actual differences in the data.*



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## Section 5

Model assessment: How well is reality described?

# Model assessment

**Model assessments aim to gauge how well we describe the data (i.e. the  $y$ ).**

- ▶ comparison between predicted and observed values (as in OLS).
- ▶ mapping outcomes to the recoded, "latent" variable (GLM).

⇒ *You have a few additional "tricks" to the standard OLS assessment.*

## Brier score

**Describes the "average size" of the residuals.**

$$B_b \equiv \frac{1}{n} \sum_{i=1}^n (\hat{\theta}_i - y_i)^2 \quad (5)$$

$\Rightarrow$  *Lower scores imply better predictions.*

## How well do I discriminate?

**The real question for logits is how well do I distinguish 0s from 1s.**  
 $\Rightarrow$  *Several strategies.*

## Table comparison

**The real question for logits is how well do I distinguish 0s from 1s.**

- ▶ Table (e.g.  $2 \times 2$ ) with proportion of predicted against observed values for 0s and 1s.
- ▶ It is  $\chi^2$  distributed (ref. the Hosmer-Lemeshow test)

$\Rightarrow$  *But how do I set the cut values (the  $\tau$ )?*

# The ROC curve

**The ROC lets the cut values vary and displays how wrong we are on each side (true positive vs. false positive).**

- ▶ A model with good predictions has a curve tending towards the upper left corner.
- ▶ The actual cut value depends on our priorities

⇒ *The graphic is useful in and of itself*

# The separation plot

**The separation plot show how the density of observed "successes" increases as our predicted values increase.**

*⇒ Another graphic that is useful in and of itself*