

# Multilevel/hierarchical models: Overview

Silje Synnøve Lyder Hermansen

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Where are we in the course?

## Recap from Monday

**When observations are not i.i.d. (i.e. they share a group identity), we will often consider alternatives to the ordinary linear model**

- ▶ negative take: the assumptions of the linear model are not met.
  - ▶ non-normal residuals,
  - ▶ heteroscedastic residuals
  - ▶ correlation between  $x$  and residuals
- ▶ positive take: we have variation that we want to leverage strategically
  - ▶ within-group variation
  - ▶ between-group variation
  - ▶ more correct estimation of the standard errors

⇒ *see this as an opportunity*

## I pick my models as part of my research design

### **What are the most relevant correlations/variation given my theory?**

- ▶ in experiments: you can create that variation and randomize the rest (cut out confounders)
- ▶ in observational studies: you'll have to “hunt” for the variation you want and control away the rest

# Confounders

- ▶ Control variables that – if absent lead to omitted variable bias – satisfy three criteria:
  - ▶  $z$  correlates with  $y$
  - ▶  $z$  correlates with  $x$
  - ▶  $z$  causes  $x$  and  $y$  (not intermediate/post-treatment)

→ *even when 3 is not satisfied, it might be a sign of a common group identity (e.g. nationality)*
- ▶ Group identities: observations done in the same context share many potential confounders
  - ▶ you might kill several birds with one stone

## The principle

# The principle

**We make the assumption that the residuals are drawn from a normal distribution**

- ▶ **pooled models:** a single distribution

$$y_i = a + bx_i + \epsilon_i$$

$$\epsilon_i \sim N(0, \sigma^2)$$

- ▶ **hierarchical models:** add a hierarchy
  - ▶ assume groups are drawn from different distributions
  - ▶ their mean is drawn from a single distribution that “rules them all”

$$y_i = a + bx_i + \epsilon_{ji}$$

$$\epsilon_j \sim N(\alpha_j, \sigma_j^2)$$

$$\alpha_j \sim N(0, \sigma_\alpha^2)$$

# Untangling the parameters/variation

**This allows me to untangle different sources of variation**

$$y_i = a + bx_i + \epsilon_{ji}$$

$$\epsilon_j \sim N(\alpha_j, \sigma_j^2)$$

$$\alpha_j \sim N(0, \sigma_\alpha^2)$$

- ▶  $\alpha_j$ : grouped mean of residuals: group intercept
- ▶  $\sigma_\alpha^2$ : between-group variation
- ▶  $\sigma_j^2$ : group-level (within) variation



# The promises of a hierarchical structure

## This allows me to leverage different sources of variation

- ▶ leverage **within-group** variation:
  - ▶ by factoring out/control for between-group variation ( $\sigma_j^2$ )
- ▶ leverage **between-group** variation:
  - ▶ by running a second regression on the group means ( $\alpha_\alpha^2$ )
    - ▶ adjusts the standard errors
    - ▶ data augmentation: add variables from other sources that vary by group
    - ▶ predict out of sample even for new groups
- ▶ leverage **both sources** of variation
  - ▶ by borrowing from the more informative variation (“pooling”/“shrinkage”)

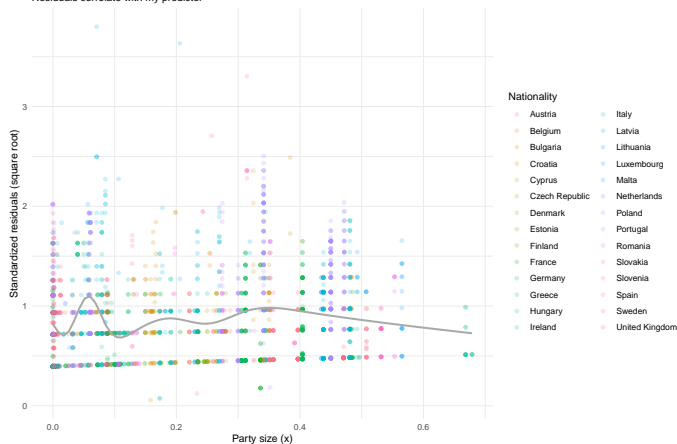
## Labeling the errors: grouped residuals

# Labeling the errors: grouped residuals

**Our residuals have group identities that we can “label” as such.**

The ghosts of our regression

Residuals correlate with my predictor

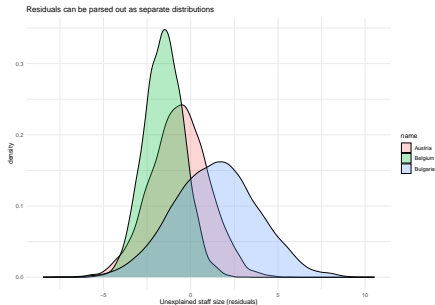


## Group means and group-level variation

**Our residuals have group identities that we can “label” as such.**

- ▶ each group of residuals has a distribution with a mean and a spread

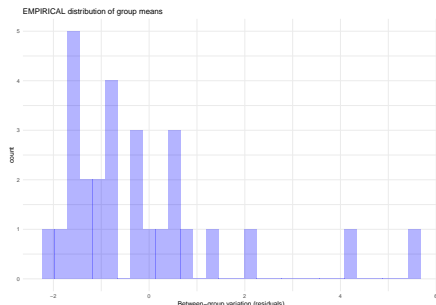
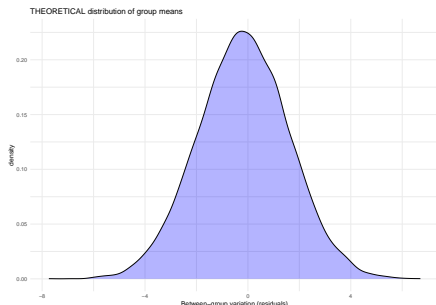
```
## # A tibble: 28 x 3
##   Nationality y_bar_j sigma2_alpha
##   <chr>      <dbl>      <dbl>
## 1 Austria    -0.665      1.65
## 2 Belgium   -1.54       1.15
## 3 Bulgaria   1.41       2.44
## 4 Croatia    0.549      4.10
## 5 Cyprus    -0.253      1.89
## 6 Czech Republic -0.206    1.84
## 7 Denmark   -1.48      1.30
## 8 Estonia   -1.33      0.950
## 9 Finland   -1.47      0.919
## 10 France    -1.11      1.26
## # i 18 more rows
```



⇒ *I can reconstruct their theoretical distribution by calculating the group mean and standard deviation*

# Between-group variation

**The group means are drawn from a common normal distribution with a mean and a spread**



⇒ *I am treating the residuals as if they were a variable, so statistical theory can be applied*

## Varying-intercepts regression: within-group variation

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## **The random/varying-intercept model:**

- ▶ a common slope for all predictors
- ▶ separate intercepts for all group identities
- ▶ a common intercept (grand mean)

## From labelled errors to varying intercepts

**Instead of hiding the groupings in the residuals, we can report them as a series of intercepts (i.e. report their group means)**

$$y_i = a + bx_i + \alpha_j$$

$$\alpha_j \sim N(0, \sigma_\alpha^2)$$

- ▶  $a$ : the **grand mean** (mean of  $\alpha$  means)
- ▶  $\alpha_j$ : **varying intercepts** (deviations from this grand mean)

$\Rightarrow$  *useful for interpretation in R*



## Varying-intercepts

**Now, it is clear that I parse out (control for) between-group variation**

- ▶ **within-group variation** the b coefficients report the effect of observation-level variables
- ▶ **group-level variation** is reported in the varying intercepts, it is the variation that:
  - ▶ has not been accounted for by my main effects
  - ▶ that can be attributed to group identities

## Estimation in R: Varying national intercepts

# Estimation in R: Varying national intercepts

Let's regress MEPs' investment in their district (y) on...

- ▶ x: their party's size in the national parliament (as a proxy for state funding).
- ▶ ... while controlling away between-national variation

## Equation:

$$\text{Staff size} = a + b \times \text{Party size} + \alpha_{\text{Nationality}}$$

$$y_i = a + bx_i + \alpha_{ij}$$

## Estimation:

```
library(lme4)
mod.ran.int <- lmer(y ~ x + (1|Nationality),
  df)
```

## Reading the R output

# Reading the R output

```
summary(mod.ran.int)
```

```
## Linear mixed model fit by REML ['lmerMod']
## Formula: y ~ x + (1 | Nationality)
## Data: df
##
## REML criterion at convergence: 31355.2
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.1127 -0.5387 -0.1435  0.3598 15.2357
##
## Random effects:
## Groups      Name                Variance Std.Dev.
## Nationality (Intercept)  3.125      1.768
## Residual                5.240      2.289
## Number of obs: 6948, groups: Nationality, 28
##
## Fixed effects:
##              Estimate Std. Error t value
## (Intercept)   2.6799    0.3386   7.915
## x             -1.6722    0.1678  -9.965
##
## Correlation of Fixed Effects:
##      (Intr)
## x -0.117
```

R refers to the residuals as “random effects”

$\sigma^2_{\alpha}$ : remaining **between-group variance**: 3.12

- ▶ standard deviation: 1.77
- ▶ the unexplained variation between groups

Residual: remaining **within-group variance**: 5.24

- ▶ standard deviation of within-group distribution: 2.29
- ▶ the unexplained variation within all groups

R refers to regression coefficients as “fixed effects”

a: intercept/**grand mean**: 2.68

- ▶ a hypothetical intercept for interpretation (mean of means)

b: **slope**: -1.67

- ▶ the marginal effect of party size (x)

## Interpretation

# Interpretation

Interpretation follows normal principles, but there are some complications:

a. we now have two intercepts per scenario:

- ▶ the grand mean ( $\alpha$ ): for focus on general effect of  $x$
- ▶ the group-level mean ( $\alpha_j$ ): for description and prediction
- ▶ sum of the grand mean ( $\alpha$ ) and group-level mean ( $\alpha_j$ ): for prediction

b. all effects are linear

- ▶ so first-difference and marginal effects are the same

## Interpreting marginal effects

**The interpretation of the marginal effect is as with any linear model:**

Table 1: Effect of state funding for parties on MEPs' local staff size

	Dependent variable:
	y
x	-1.672*** (0.168)
Constant	2.680*** (0.339)
Observations	6,948
Log Likelihood	-15,677.610
Akaike Inf. Crit.	31,363.210
Bayesian Inf. Crit.	31,390.600
Note: *p<0.1; **p<0.05; ***p<0.01	

⇒ *A 10% decrease in the national party's seat share would lead every 6th MEP to compensate by hiring an additional local staffer.*



# Prediction

## The varying intercepts are reported as deviations from the grand mean

```
fixef(mod.ran.int); ranef(mod.ran.int)
```

```
## (Intercept)          x
##      2.679887    -1.672226

##              (Intercept)
## Austria             -0.49518857
## Belgium             -1.52249566
## Bulgaria             1.54657524
## Croatia              0.68267309
## Cyprus              -0.05313986
## Czech Republic     -0.10587832
```

Predicted local staff in Austria when national party is not in Parliament:

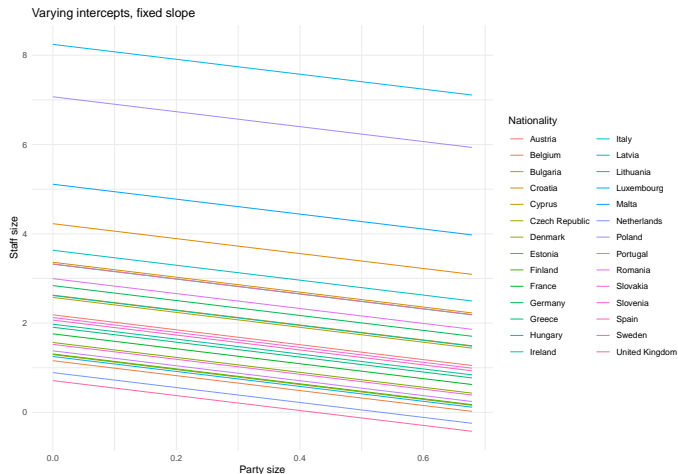
$$\blacktriangleright 2.68 + -0.5 \times 0 = 2.18$$

Predicted local staff in Austria when national party holds 10% of the seats

$$\blacktriangleright 2.68 + -0.5 + -1.67 \times 0.1 = 2.02$$

# Visualization

## Effect of x, the slope coefficient

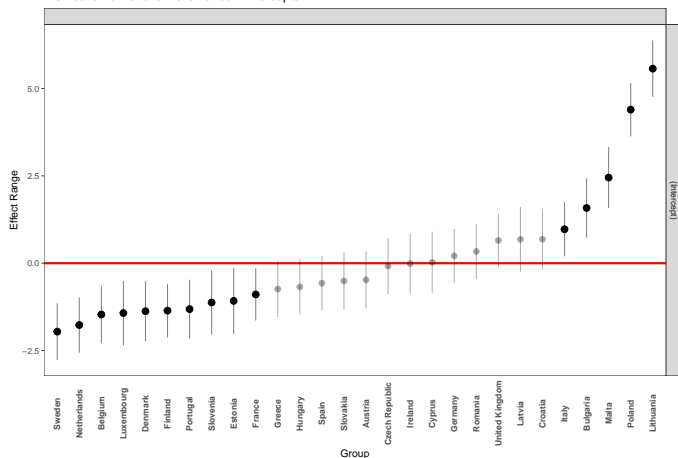


⇒ *the slope is constant, but the intercept changes across nationalities*

# Visualization: as distributions

## The intercepts are distributions in their own right

Distribution of national-level random intercepts



⇒ each varying intercept has a point estimate (regression coefficient) and a distribution. They vary around a normal distribution with mean of 0

## Varying slopes, varying intercepts

## Defintion

We can let the effect of  $z$  vary by group

$$y_i = a + b_1x_i + c_jz_i + \alpha_j$$

- ▶  $c_j$ : **varying slope** (the effect of  $z$  varies by group)
- ▶  $\alpha_j$ : **varying intercepts**
- ▶ we can rewrite to make this explicit

$$y_i = a + bx_i + \epsilon_{ij}$$

$$\epsilon_j \sim N(\alpha_j, \sigma_\alpha)$$

$$\alpha_j = \lambda_j + c_jz_j$$

- ▶  $\lambda_j$ : **varying intercepts**

$\Rightarrow$  *a series of regressions within the regression*

# Estimation in R

**the estimation is done as if it was an interaction effect**

- ▶ fixed-effects model with cross-level interaction

```
mod.ran.slope <- lm(y ~ x + ProxNatElection * Nationality, df)
```

- ▶ random-effects model with varying slope

```
mod.ran.slope <- lmer(y ~ x + (ProxNatElection | Nationality),
```

## Interpretation

# Marginal effects

We can read these coefficients as if they were from separate models

```
ranef(mod.ran.slope)
```

	(Intercept)	ProxNatElection
## Austria	-0.3201686	0.001073577
## Belgium	-1.3944093	-0.018299157
## Bulgaria	1.8027887	0.086786758
## Croatia	0.9352286	0.058233060
## Cyprus	0.1020429	-0.003477477
## Czech Republic	0.1112017	0.024050888

MEPs from Austria hire on average 0.004 ( $= 0.001 * 4$ ) assistants more immediately before an election compared to immediately after, while MEPs from Belgium hire on average 0.073 ( $= 0.018 * 4$ ) fewer assistants.

► These are negligible marginal effects.



# Prediction

## The prediction is done per group, but follows normal rules

- ▶ two intercepts:  
grand mean + group-level intercept
- ▶ one slope per group

```
fixef(mod.ran.slope); ranef(mod.ran.slope)
```

```
## (Intercept)      x
##    2.513035  -1.691321
```

```
##              (Intercept) ProxNatElection
## Austria          -0.3201686      0.001073577
## Belgium          -1.3944093     -0.018299157
## Bulgaria          1.8027887      0.086786758
## Croatia           0.9352286      0.058233060
## Cyprus            0.1020429     -0.003477477
## Czech Republic    0.1112017      0.024050888
```

Austria after election:

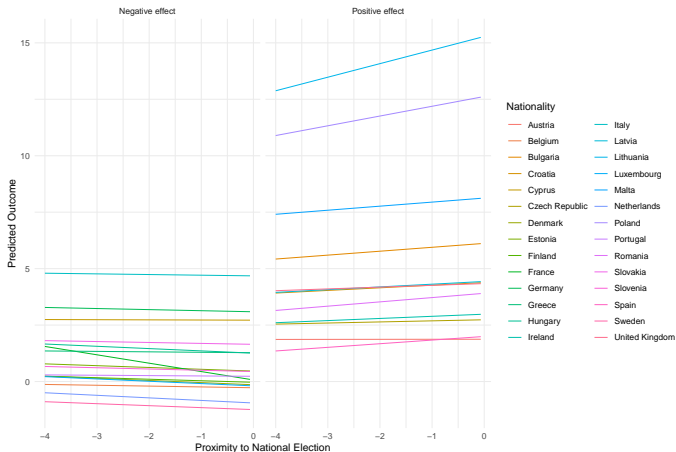
$$\text{▶ } 2.51 + -0.32 + 0.001 \times -4 = 2.189$$

Austria before election:

$$\text{▶ } 2.51 + -0.32 + 0.001 \times 0 = 2.193$$

# Visualization

Varying slopes and intercepts: Effect of electoral calendar by nationality



## Level-2 regression: between-group variation

## Definition

## Definition

We can think of the residuals/group intercepts as a variable in their own right

$$y_i = bx_i + \epsilon_{ji}$$

- ▶ they are generated by draws from J number of distributions:

$$\epsilon_{ji} \sim N(\alpha_j, \sigma_\alpha^2)$$

- ▶ ... and therefore we can model them

$$\alpha_j = a + dz_j$$

- ▶ a: a **single intercept**
- ▶ d: a **single slope coefficient**

⇒ *we run a second regression on the residuals*

# Implications

## We explicitly model between-group variation

- ▶  $z$ , the level-2 predictor only varies at the group level
  - ▶ standard errors for  $z$  reflect the number of groups
  - ▶ the more groups, the more the approach makes sense
- ▶ data augmentation
  - ▶ we can add information from other to the model
  - ▶ contextual elements
  - ▶ improves prediction

## Estimation in R: Electoral system

# Estimation in R: Electoral system

## Let's add electoral system (z) as a predictor

- ▶ it never changes in a country (in this study)

## R handles this automatically

- ▶ same data frame
  - ▶ all variables that don't vary within groups are regressed as a level 2
- ▶ coefficients reported the same way
- ▶ estimation of coefficients and standard errors is different

```
mod.two.levels <- lmer(y ~ x + z + (1|Nationality), df)
```



## Reading the R output

# Reading the R output

## The R output looks exactly the same as for the varying-intercept model.

```
## Linear mixed model fit by REML ['lmerMod']
## Formula: y ~ x + z + (1 | Nationality)
## Data: df
##
## REML criterion at convergence: 31353.9
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.1145 -0.5388 -0.1434  0.3599 15.2339
##
## Random effects:
##   Groups      Name      Variance Std.Dev.
## Nationality (Intercept) 3.235     1.799
## Residual                5.240     2.289
## Number of obs: 6948, groups: Nationality, 28
##
## Fixed effects:
##              Estimate Std. Error t value
## (Intercept)   2.5268    0.6030   4.191
## x             -1.6719    0.1678  -9.962
## z              0.2263    0.7311   0.310
##
## Correlation of Fixed Effects:
##      (Intr) x
## x  -0.077
## z  -0.821  0.013
```

The level-2 regression coefficient appears as “fixed effects”

a: **grand mean:** 2.53

▶ the “mean of means”

d: **slope:** 0.23

▶ the marginal effect of electoral system (z)

Check the **change in between-group variance:**

▶ the between-group variance ( $\sigma_{\alpha}^2$ , 3.23) should normally decrease

▶ it is not the case here ( $3.12 \leq 3.23$ )

→ *increase in variance indicates “complexities” between levels (interactions)*

Correlation of Fixed Effects:

▶ negative correlation between predictor (z) and intercept (-0.82): high level of z correlates with low base-line value of y.

# Pooling

# Pooling

## What is the difference between a fixed-effects and a random-effects model, then?

- ▶ the fixed-effects model only compares within groups

```
mod.fix <- lm(y ~ a + Nationality, df)
```

- ▶ the random-effects (hierarchical) model borrows information between and within groups → *pools*

```
mod.fix <- lmer(y ~ a + (1|Nationality), df)
```

⇒ *both are varying-intercepts models*

# What is pooling?

**The hierarchical model calculates a weighted average of between- and within-group variation for each coefficient**

$$\frac{\frac{n_j}{\sigma_y^2} \bar{y}_j + \frac{1}{\sigma_\alpha^2} \bar{y}_{all}}{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}}$$

- ▶ the denominator is there to normalize  $(\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}) \rightarrow$  *ignore it*
- ▶  $\bar{y}_{all}$ : the pooled mean
  - ▶ its weight  $(\frac{1}{\sigma_\alpha^2})$ 
    - ▶  $\sigma_\alpha^2$ : between-group variation
- ▶  $\bar{y}_j$ : the group mean
  - ▶ its weight  $(\frac{n_j}{\sigma_y^2})$ 
    - ▶  $n_j$ : size of the group (number of observations)
    - ▶  $\sigma_y^2$ : residual variation not explained by the between-group variation

## The weights in pooling

The hierarchical model calculates a **weighted average of between- and within-group variation for each coefficient**

$$\frac{\frac{n_j}{\sigma_y^2} \bar{y}_j + \frac{1}{\sigma_\alpha^2} \bar{y}_{all}}{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}}$$

- ▶  $\sigma_\alpha^2$ : as the **between-group variation** increases, the weight of the pooled mean decreases
- ▶  $n_j$ : as **the size of the group** (number of observations) increases, the weight of the non-pooled (within-) group mean increases

# What to do?

## Sooo... what do I choose?

Condition	Fixed	Random	Advantage	Limitation
plenty of within-group variation	x		stringent comparison	no weighing of groups
		x	weighing by group size	groups should be distinct (between-group variation is high)
variables only vary by group		x	standard errors are corrected	fixed effects will be non-identified
mix of between- and within-group variation		x	pooling/borrows information	no idea where the info comes from
data augmentation/prediction		x	infers from group-level predictors	fixed effects don't perform out of sample

# How many groups and how many observations?

## Random/hierarchical model

- ▶ if you want level 2 variables:
  - ▶ **many groups** → *you run a second regression*
- ▶ if you want within-group variation:
  - ▶ **distinct groups** (large between-group variation, size matters less) → *similar to fixed-effects*
  - ▶ not distinct groups (little between-group variation) → *similar to pooled model*
- ▶ if you think the smaller groups are less representative
  - ▶ larger groups count more for within-group variation → *unbalanced panels*

## Fixed-effects model

- ▶ only the observations with variation within the groups count towards the estimate → *your  $N$  may be deceptive*



## Recap

# Recap

## Hierarchical models leverage variation according to the structure in the data (groupings)

- ▶ varying-intercepts models (fixed and random effects)
  - ▶ *one slope*, but control for group identities
- ▶ varying-intercept, varying slope (fixed and random effects)
  - ▶ one intercept and one slope *per group*,
- ▶ level-2 regression (random effects)
  - ▶ one slope per group predictor, but adjusts *standard errors*,
- ▶ pooling (all random effects models)
  - ▶ regression coefficients are a weighted average of between- and within-group variation

⇒ *Pick the variation you want, then pick the model you need.*