

Probability Theory and Mathematical Statistics

I. Probability Formulas

- **Addition Formula:** $P(A + B) = P(A \cup B) = P(A) + P(B) - P(AB)$
- **Subtraction Formula:** $P(A - B) = P(A\bar{B}) = P(A) - P(AB)$
- **Conditional Probability:** $P(A|B) = \frac{P(AB)}{P(B)}, P(B) > 0$
- **Compatible Events:** $P(AB) > 0$
- **Mutually Exclusive Events:** $P(AB) = 0$
- **Independent Events:** $P(AB) = P(A)P(B)$
- **Distribution Function:**
 - $F(a) = P\{X \leq a\}$
 - $P\{X < a\} = \lim_{x \rightarrow a^-} F(x)$
 - $f_X(x) = \frac{dF_X(x)}{dx}$
- **Convolution Function** ($Z = X + Y$): $f_Z(z) = \int_{-\infty}^{\infty} f_X(x)f_Y(z - x) dx$
- **Convolution Function** ($Z = X - Y$): $f_Z(z) = \int_{-\infty}^{\infty} f_X(x)f_Y(x - z) dx$

II. Numerical Characteristics

- **Mathematical Expectation:**
 - Continuous Type: If the probability density of X is $f(x)$, then $E(X) = \int_{-\infty}^{+\infty} xf(x) dx$
 - Function of Random Variable: $E[g(X)] = \int_{-\infty}^{+\infty} g(x)f(x) dx$
- **Marginal Expectation in Two-Dimensional Case:** If the joint probability density of (X, Y) is $f(x, y)$, its marginal probability densities are:

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy$$
$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx$$

- **Variance:**
 - Definition: $D(X) = E[(X - E(X))^2]$

- Computational Formula: $D(X) = E(X^2) - [E(X)]^2$
- Property: $D(CX) = C^2 D(X)$
- **Correlation Coefficient:** $\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)} \cdot \sqrt{D(Y)}}$
- **Equivalent Propositions of Uncorrelated:**

$$\text{Cov}(X, Y) = 0 \iff \rho_{XY} = 0 \iff E(XY) = E(X)E(Y) \iff D(X+Y) = D(X) + D(Y)$$
- **Covariance:**
 - Definition: $\text{Cov}(X, Y) = E[(X - E(X))(Y - E(Y))]$
 - Computational Formula: $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$
- **Properties of Covariance**
 1. $\text{Cov}(X, Y) = \text{Cov}(Y, X)$
 2. $\text{Cov}(X, C) = 0$
 3. $\text{Cov}(aX, bY) = ab \text{Cov}(X, Y)$
 4. $\text{Cov}(X, X) = D(X)$
 5. $\text{Cov}(X + Y, Z) = \text{Cov}(X, Z) + \text{Cov}(Y, Z)$
 6. $D(X \pm Y) = D(X) + D(Y) \pm 2 \text{Cov}(X, Y)$
 7. X, Y are independent $\Rightarrow \text{Cov}(X, Y) = 0$

III. Probability Distributions

- **Poisson Distribution:** $X \sim P(\lambda)$
 - Conclusion: If $X \sim P(\lambda_1)$, $Y \sim P(\lambda_2)$, and X and Y are independent, then $X + Y \sim P(\lambda_1 + \lambda_2)$
- **Normal Distribution:** $X \sim N(\mu, \sigma^2)$
 - Standardization and Probability:
$$P\{a < X \leq b\} = P\left\{\frac{a - \mu}{\sigma} < \frac{X - \mu}{\sigma} \leq \frac{b - \mu}{\sigma}\right\} = \Phi\left(\frac{b - \mu}{\sigma}\right) - \Phi\left(\frac{a - \mu}{\sigma}\right)$$
 - Conclusion: If $Z \sim N(0, 1)$, then
$$\forall a > 0, \quad P\{|Z| \leq a\} = 2\Phi(a) - 1$$
- **Exponential Distribution:** $X \sim E(\lambda)$ ($\lambda > 0$)
 - Conclusion:
 - * $P\{X > a\} = e^{-\lambda a}$ ($a > 0$)
 - * $P\{X > s + t \mid X > s\} = P\{X > t\}$, where $s, t > 0$

• **Common Distributions Table:**

Type	Notation	Distribution Law/Density	Expectation	Variance
0-1	$X \sim b(1, p)$	$P\{X = k\} = p^k(1 - p)^{1-k} \quad k = 0, 1$	p	$p(1 - p)$
Binomial	$X \sim B(n, p)$	$P\{X = k\} = C_n^k p^k (1 - p)^{n-k} \quad k = 0, 1, \dots, n$	np	$np(1 - p)$
Poisson	$X \sim \pi(\lambda)$	$P\{X = k\} = \frac{\lambda^k}{k!} e^{-\lambda} \quad k = 0, 1, 2, \dots$	λ	λ
Uniform	$X \sim U(a, b)$	$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Normal	$X \sim N(\mu, \sigma^2)$	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad -\infty < x < +\infty$	μ	σ^2
Exponential	$X \sim E(\lambda)$	$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & x \leq 0 \end{cases}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$

IV. Mathematical Statistics

• **Common Statistics:**

- Sample Mean: $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$
- Sample Variance: $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

• **Three Major Sampling Distributions:**

- **χ^2 Distribution:**
 - * Definition: $X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} N(0, 1)$, then $\sum_{i=1}^n X_i^2 \sim \chi^2(n)$
 - * Property: $E(\chi^2) = n$, $D(\chi^2) = 2n$
- **t Distribution:**
 - * Definition: $X \sim N(0, 1)$, $Y \sim \chi^2(n)$ independent, then $T = \frac{X}{\sqrt{Y/n}} \sim t(n)$
 - * Property: Probability density $f(t)$ is an even function; if $T \sim t(n)$, then $T^2 \sim F(1, n)$
- **F Distribution:**
 - * Definition: $X \sim \chi^2(n_1)$, $Y \sim \chi^2(n_2)$ independent, then $F = \frac{X/n_1}{Y/n_2} \sim F(n_1, n_2)$
 - * Property: If $F \sim F(n_1, n_2)$, then $\frac{1}{F} \sim F(n_2, n_1)$

• **Properties of Statistics:** Let X_1, X_2, \dots, X_n be a sample from population X , with $E(X) = \mu$, $D(X) = \sigma^2$, then:

- $E(X_i) = \mu$, $D(X_i) = \sigma^2$, $E(\bar{X}) = \mu$, $D(\bar{X}) = \frac{\sigma^2}{n}$

- **Testing Methods for Mean and Variance of Normal Population:**

Null Hypothesis H_0	Test Statistic	Rejection Region
$\mu \leq \mu_0$	$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$	$z \geq z_\alpha$
$\mu = \mu_0$		$ z \geq z_{\alpha/2}$
$\mu \geq \mu_0$		$z \leq -z_\alpha$
$\mu \leq \mu_0$	$t = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$	$t \geq t_\alpha(n-1)$
$\mu = \mu_0$		$ t \geq t_{\alpha/2}(n-1)$
$\mu \geq \mu_0$		$t \leq -t_\alpha(n-1)$
$\sigma^2 \leq \sigma_0^2$	$\chi^2 = \frac{(n-1)S^2}{\sigma_0^2}$	$\chi^2 \geq \chi_\alpha^2(n-1)$
$\sigma^2 = \sigma_0^2$		$\chi^2 \leq \chi_{1-\alpha/2}^2(n-1)$ or $\chi^2 \geq \chi_{\alpha/2}^2(n-1)$
$\sigma^2 \geq \sigma_0^2$		$\chi^2 \leq \chi_{1-\alpha}^2(n-1)$

V. Parameter Estimation and Hypothesis Testing

- **Unbiasedness:** If $E(\hat{\theta}) = \theta$, then $\hat{\theta}$ is called an unbiased estimator of θ .
- **Efficiency:** If $\hat{\theta}_1$ and $\hat{\theta}_2$ are both unbiased estimators of θ , and $D(\hat{\theta}_1) < D(\hat{\theta}_2)$, then $\hat{\theta}_1$ is more efficient than $\hat{\theta}_2$.
- **Method of Moments:** $E(X) = \bar{X}$

$$- \int_{-\infty}^{+\infty} xf(x) dx = \frac{1}{n} \sum_{i=1}^n X_i$$

- **Steps of Maximum Likelihood Estimation:**

$$- L(\theta) = \prod_{i=1}^n p(x_i; \theta) \Rightarrow \ln L(\theta) = \sum_{i=1}^n \ln p(x_i; \theta) \Rightarrow \frac{d[\ln L(\theta)]}{d\theta} = 0$$

- **Confidence Interval for μ of Single Normal Population** (Confidence Level $1 - \alpha$):

$$1. \sigma^2 \text{ known: } \left(\bar{X} - Z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}, \bar{X} + Z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} \right)$$

$$2. \sigma^2 \text{ unknown: } \left(\bar{X} - t_{\alpha/2}(n-1) \cdot \frac{S}{\sqrt{n}}, \bar{X} + t_{\alpha/2}(n-1) \cdot \frac{S}{\sqrt{n}} \right)$$

- **Two Types of Errors in Hypothesis Testing:**

- **Type I Error (Rejecting Truth):** When H_0 is true, H_0 is rejected.
- **Type II Error (Accepting False):** When H_0 is false, H_0 is accepted.