

# Exercise 1

TMA4300 Computer Intensive Statistical Models

Mads Adrian Simonsen, William Scott Grundeland Olsen

29 januar, 2021

## Problem A: Stochastic simulation by the probability integral transform and bivariate techniques

### Subproblem 1.

Let  $X \sim \text{Exponential}(\lambda)$ , with the cumulative density function

$$F_X(x) = 1 - e^{-\lambda x}, \quad x \geq 0.$$

Then the random variable  $Y := F_X(X)$  has a  $\text{Uniform}(0, 1)$  distribution. The probability integral transform becomes

$$Y = 1 - e^{-\lambda X} \Leftrightarrow X = -\frac{1}{\lambda} \ln(1 - Y).$$

It is clear that if  $U \sim \text{Uniform}(0, 1)$ , then  $1 - U \sim \text{Uniform}(0, 1)$ , and therefore we may as well say that

$$X = -\frac{1}{\lambda} \ln(Y). \tag{1}$$

Thus, we sample  $Y$  from `runif()` and transform it using Equation (1), to sample from the exponential distribution. Figure 1 shows one million samples drawn from the `generate_from_exp()` function defined in the code chunk below. It also shows the theoretical PDF of the exponential distribution with rate parameter  $\lambda = 2$ .

```
#set.seed(123)

generate_from_exp <- function(n, rate = 1) {
  Y <- runif(n)
  X <- -(1 / rate) * log(Y)
  return(X)
}

# sample
n <- 1000000 # One million samples
lambda <- 2
exp_samp <- generate_from_exp(n, rate = lambda)

# plot
ggplot() +
  geom_histogram(
```

```

data = as.data.frame(exp_samp),
mapping = aes(x = exp_samp, y = ..density..),
binwidth = 0.05,
boundary = 0
) +
stat_function(
  fun = dexp,
  args = list(rate = lambda),
  aes(col = "Theoretical density")
) +
ylim(0, lambda) +
xlim(0, 2) +
ggtitle("Dimulating from an exponential distribution") +
xlab("x") +
ylab("Density") +
theme_minimal() +
theme(plot.title = element_text(hjust = 0.5))

```

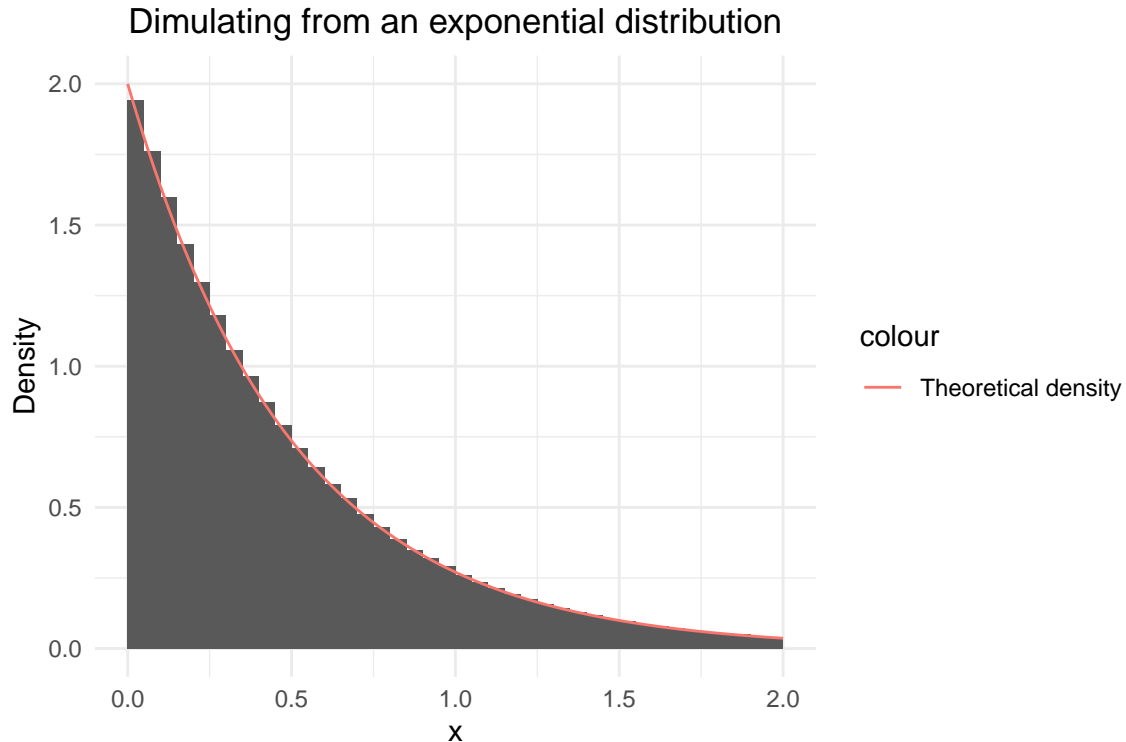


Figure 1: Normalized histogram of one million samples drawn from the exponential distribution, together with the theoretical PDF, with  $\lambda = 2$ .

Theoretically, the mean and variance of  $X \sim \text{Exponential}(\lambda)$  is  $E(X) = \lambda^{-1}$  and  $\text{Var}(X) = \lambda^{-2}$ . So for  $\lambda = 2$  we would expect  $E(X) = 1/2$  and  $\text{Var}(X) = 1/4$ . For the simulation we get the mean and variance as calculated in the code block below, showing what we would expect.

```

mean(exp_samp)

## [1] 0.5001832

```

```
var(exp_samp)
```

```
## [1] 0.2503854
```

## Subproblem 2.

### Subsubproblem (a)

We are considering the probability density function

$$g(x) = \begin{cases} cx^{\alpha-1} & \text{if } 0 < x < 1, \\ ce^{-x} & \text{if } x \geq 1, \\ 0 & \text{otherwise,} \end{cases} \quad (2)$$

where  $c$  is a normalizing constant and  $\alpha \in (0, 1)$ . If  $x \leq 0$  the cumulative distribution function is zero. In the interval  $0 < x < 1$  it becomes

$$G(x) = \int_{-\infty}^x g(\xi) d\xi = \int_0^x c\xi^{\alpha-1} d\xi = \frac{c}{\alpha} [\xi^\alpha]_0^x = \frac{c}{\alpha} x^\alpha,$$

and finally for  $x \geq 1$  we have

$$G(x) = \int_{-\infty}^x g(\xi) d\xi = \int_0^1 c\xi^{\alpha-1} d\xi + \int_1^x ce^{-\xi} d\xi = \left[ \frac{c}{\alpha} \xi^\alpha \right]_0^1 - [ce^{-\xi}]_1^x = c \left( \frac{1}{\alpha} - e^{-x} + \frac{1}{e} \right),$$

for  $\alpha \in (0, 1)$ . That is, the cumulative density function is

$$G(x) = \begin{cases} 0 & \text{if } x \leq 0, \\ \frac{c}{\alpha} x^\alpha & \text{if } 0 < x < 1, \\ c \left( \frac{1}{\alpha} - e^{-x} + \frac{1}{e} \right) & \text{if } x \geq 1. \end{cases}$$

In this case it is trivial to find  $c$ . We solve

$$1 = \int_{\mathbb{R}} g(x) dx = \int_0^1 cx^{\alpha-1} dx + \int_1^\infty ce^{-x} dx = \frac{c}{\alpha} + \frac{c}{e},$$

which gives that

$$c = \frac{\alpha e}{\alpha + e}.$$

Writing the cumulative density function using this as  $c$  we obtain

$$G(x) = \begin{cases} 0 & \text{if } x \leq 0, \\ \frac{e}{\alpha+e} x^\alpha & \text{if } 0 < x < 1, \\ 1 - \frac{\alpha}{\alpha+e} e^{1-x} & \text{if } x \geq 1, \end{cases}$$

for  $\alpha \in (0, 1)$ .

We may then find the inverse cumulative function. For  $x \leq 0$  this is just zero, and for  $0 < x < 1$ , that is  $0 < G(x) < \frac{e}{\alpha+e}$ , we solve  $x = \frac{e}{\alpha+e} y^\alpha$  for  $y$  giving  $G^{-1}(x) = \left( \frac{\alpha+e}{e} x \right)^{1/\alpha}$ . Similarly for  $x \geq 1$ , that is  $G(x) \geq 1 - \frac{\alpha}{\alpha+e} = \frac{e}{\alpha+e}$ , we solve  $x = 1 - \frac{\alpha}{\alpha+e} e^{1-y}$  for  $y$ , such that

$$G^{-1}(x) = \begin{cases} \left( \frac{\alpha+e}{e} x \right)^{1/\alpha} & \text{if } 0 \leq x < \frac{e}{\alpha+e}, \\ \ln \left[ \frac{\alpha e}{(1-x)(\alpha+e)} \right] & \text{if } x \geq \frac{e}{\alpha+e}, \end{cases}$$

for  $\alpha \in (0, 1)$ .

### Subsubproblem (b)

```

generate_from_pieewise <- function(n, alpha) {
  U <- runif(n) # Generate n Uniform(0, 1) variables
  bound <- exp(1) / (alpha + exp(1)) # Boundary where  $G^{(-1)}$  changes
  left <- U < bound # The left of the boundary
  U[left] <- (U[left] / bound)^(1 / alpha) # Left CDF
  U[!left] <- 1 + log(alpha) - log(1 - U[!left]) - log(alpha + exp(1)) # Right CDF
  return(U)
}

# Sample
n <- 1000000 # One million samples
alpha <- 0.75
x <- generate_from_pieewise(n, alpha)

# The theoretically correct PDF
theo_PDF <- function(x, alpha) {
  const <- alpha * exp(1) / (alpha + exp(1)) # Normalizing constant
  func <- rep(0, length(x))
  left <- x > 0 & x < 1 # The PDF has one value for  $0 < x < 1$ 
  right <- x >= 1 # ... and one value for  $x >= 1$ 
  func[left] <- const * x[left]^(alpha - 1) # The value to the left
  func[right] <- const * exp(-x[right]) # The value to the right
  return(func)
}

# Plot
ggplot() +
  geom_histogram(
    data = as.data.frame(x),
    mapping = aes(x = x, y = ..density..),
    binwidth = 0.05,
    boundary = 0
  ) +
  stat_function(
    fun = theo_PDF,
    args = list(alpha = alpha),
    aes(col = "Theoretical density")
  ) +
  xlim(0, 5) +
  ggtitle("Simulating from  $g(x)$  given in Equation (2)") +
  ylab("Density") +
  theme_minimal() +
  theme(plot.title = element_text(hjust = 0.5))

```

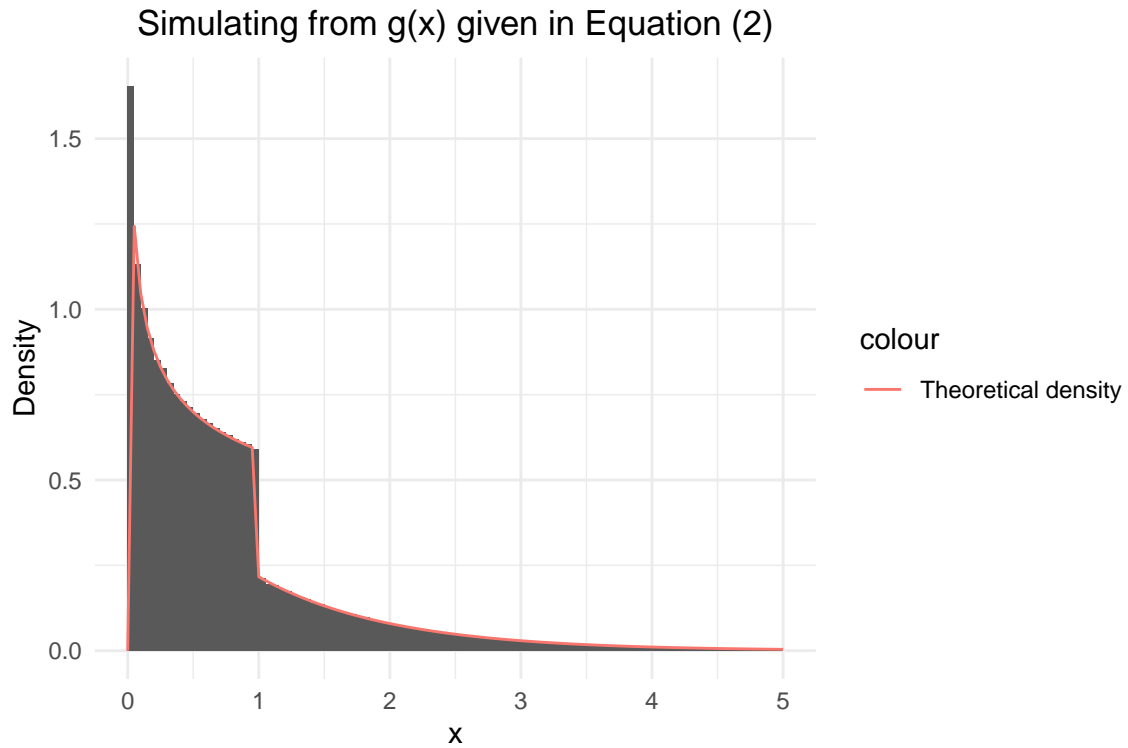


Figure 2: Normalized histogram of one million samples drawn from  $g(x)$  given in Equation (2), together with the theoretical PDF, with  $\alpha = 0.75$ .

3.

(a)

(b)

(c)

4.

5

## Problem B: The gamma distribution

1.

(a)

(b)

2.

(a)

(b)

3.

(a)

(b)

4.

5.

(a)

(b)