

## Homework - October 1, 2024

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### Problem 1: Baye's Theorem

Prove Baye's Theorem:

$$P(Y|X) = \frac{P(X|Y)P(Y)}{\sum_j P(X_j)P(Y|X_j)}$$

To prove Baye's Theorem, we recall the definition of conditional probability:

$$P(X|Y) = \frac{P(X \cap Y)}{P(Y)}$$
$$P(Y|X) = \frac{P(X \cap Y)}{P(X)}$$

Solving for  $P(X \cap Y)$  in both equations, we can then substitute, and set the two systems equal to each other

$$P(X|Y)P(Y) = P(Y|X)P(Y) = P(X \cap Y)$$

Rearranging gives the default form of the equation

$$P(Y|X) = \frac{P(X|Y)P(Y)}{P(X)}$$

We can then use the multiplication rule for conditional probabilities to write

$$P(Y|X) = \frac{P(X|Y)P(Y)}{\sum_j P(X_j)P(Y|X_j)} \quad (1)$$

as required. □

### Problem 2: Monty Hall Problem

Consider the Monty Hall game in a TV show. There are three closed doors, behind which are a car and two goats placed randomly.

1. You are asked to open a door by the host. Say, you would like to open Door 1. What is the probability of getting a car?
2. The host knows where the car is but he/she does not tell you. Instead, the host will open another door with a goat.
  - (a) If the car is behind Door 2, the host can only open Door 3.
  - (b) If the car is behind Door 1, the host can open either Door 2 or Door 3. He/she will do it with equal probability.

Say, the host has opened Door 3. What is the probability of having the car behind

Door 1 now? What is the probability of having the car behind Door 2 now?

3. If my goal is to get the car, should I change my first choice (i.e., open Door 1 or Door 2)?

We solve each section in turn.

## 1 Part 1

The probability of the car being behind the door is  $\frac{1}{3}$

## 2 Part 2

The assignment provides the following hints:

1. We can analyze this using Baye's Theorem:

$$P(A = 1|B = 1) = \frac{P(A = 1)P(B = 1|A = 1)}{P(A = 1)P(B = 1|A = 1) + P(A = 2)P(B = 1|A = 2) + P(A = 3)P(B = 1|A = 3)}$$

2. Numerator:  $\frac{1}{6}$

3. Denominator:  $\frac{3}{6}$

We can plug the numbers in and we get that  $P(A = 1|B = 1) = \frac{2}{3}$

## 3 Part 3

We should switch doors. We can compute  $P(A = 2|B = 1) = \frac{1}{3}$ , and so it is more likely that the car is behind the other door.

### Problem 3: Linearity of Expectation

Prove that the Expectation function is linear:

$$\mathbb{E}[\alpha f(X) + \beta g(X)] = \alpha \mathbb{E}[f(X)] + \beta \mathbb{E}[g(X)]$$

Recall the independence of observations for the expectation function, and the rule of constants:

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

$$\mathbb{E}[\alpha] = \alpha$$

From these two rules, we can first separate the two functions from each other:

$$\mathbb{E}[\alpha f(X) + \beta g(X)] = \mathbb{E}[\alpha f(X)] + \mathbb{E}[\beta g(X)]$$

Since we know that the expectation of a constant is that constant, we can extract the constants

$$\mathbb{E}[\alpha f(X) + \beta g(X)] = \alpha \mathbb{E}[f(X)] + \beta \mathbb{E}[g(X)]$$

as required. □

#### Problem 4: Parameter Estimation

Let  $X \sim U[a, b]$  be a continuous random variable uniformly distributed in the interval  $[a, b]$ , where  $a$  and  $b$  are unknown parameters.

We have a dataset  $\{x^{(m)}\}_{m=1}^M$ , where each data sample is iid drawn from the above distribution, and we would like to estimate the parameters  $a$  and  $b$ .

Give the likelihood of parameters.

Give the maximum likelihood estimation of parameters.

If  $X$  is a uniform distribution, then the probability density function of  $X$  is

$$f(x \in X|a, b) = \frac{1}{b-a} \quad a \leq x \leq b$$

We can compute the likelihood of the data given the parameters  $a$  and  $b$  with the likelihood function  $\mathcal{L}(a, b)$  which is the joint probability of observing all data in our set  $x^{(m)}$ . We can thus write

$$\begin{aligned} \mathcal{L}(a, b) &= \prod_{m=1}^M f(x|a, b) \\ &= \prod_{m=1}^M \frac{1}{b-a} \\ &= \left( \frac{1}{b-a} \right)^M \end{aligned}$$

and we need to add a term to ensure that all data is within  $a$  and  $b$  since observing data outside of the bounds has probability 0

$$\mathcal{L}(a, b) = \left( \frac{1}{b-a} \right)^M \cdot \mathbb{I}_{a \leq \min \vec{x}} \cdot \mathbb{I}_{b \geq \max \vec{x}}$$

and thus, minimizing  $b - a$  maximizes the likelihood function, so our most likely estimate for the parameters  $a$  and  $b$  is

$$\begin{aligned} a &= \min \vec{x} \\ b &= \max \vec{x} \end{aligned}$$

□

#### Problem 5: MSE

Suppose  $M$  samples, each of which  $x^{(m)} \sim \mathcal{N}(\mu, 1)$  is iid generated. Show that the estimate

$$\hat{\mu} = \frac{1}{M} \sum_{m=1}^M x^{(m)}$$

is an unbiased estimate of  $\mu$ .

This is a duplicate of the proof we covered in class for the MSE being an unbiased estimate of the value  $\mu$ .

First, the sample mean for the set is

$$\bar{\mu} = \frac{1}{M} \sum_{m=1}^M x^{(m)}$$

Now, we need to prove that the mean is a fair estimator, so let us compute the expectation of the gaussian normal distribution as given:

$$\begin{aligned} \mathbb{E}[\mathcal{N}(\mu, \sigma^2)] &= \operatorname{argmax}_{\theta} \Pi P(x|\mu, \sigma) \\ &= \operatorname{argmax}_{\theta} \log [\operatorname{argmax}_{\theta} \Pi P(x|\mu, \sigma)] \\ &= \operatorname{argmax}_{\mu} \sum_{m=1}^M \log \left[ \frac{1}{\sqrt{2\pi}\sigma} \exp \left\{ -\frac{1}{2} \left( \frac{t^{(m)} - \mu^T x^{(m)}}{\sigma} \right)^2 \right\} \right] \\ \mathbb{E}[\mathcal{N}(\mu, \sigma^2)] &= \frac{1}{M} \sum_{m=1}^M x^{(m)} \end{aligned}$$

Which we can find by a slight change to the derivation done in class. □