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## Modellierung & Simulation I

### Serie 06

#### Problem 6.1.1

We use the assumption, that the arrivals at the three interfaces are independent and the underlying poisson processes have the same rate  $\lambda$ . Then the superposition of these three poisson processes with rate  $\lambda$  is a poisson process with rate  $3\lambda$ . Therefore the interarrival times follow an exponential distribution and the cumulative distribution function is  $A(t) = 1 - e^{-3\lambda t}$ . ✓

Herleitung fehlt.

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#### Problem 6.1.2

Using the same assumptions as above, the interarrival times at the switch are exponentially distributed. The parameter is  $3\lambda$ . ✓

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#### Problem 6.1.3

For an interval  $\Delta$  of finite duration, the number of arrivals  $\#$  within this timeframe is poisson distributed. With the above assumption the parameter is  $3\lambda$  and the probability formula becomes:

$$P(\# = n) = \frac{(3\lambda)^n}{n!} e^{-3\lambda t}$$

$\Delta$  fehlt in Formel  
 was ist  $t$ ?  
 -2

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#### Problem 6.2.1

If the traffic arrives with exponential inter-arrival times and only every  $n$ -th packet is extracted, the time between 2 extracted packets is Erlang- $n$  distributed. This is, because the sum of  $n$  (independent) samples from the exponential distribution with rate  $\lambda$  determine the time  $t$ , after which the next extraction occurs. Therefor

$$P(D_A = t) = \frac{\lambda^n t^{n-1} e^{-\lambda t}}{(n-1)!}$$

✓

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### Problem 6.2.2

The convolution of 2 functions  $f$  and  $g$  is defined by

$$(f \otimes g) = \int_{-\infty}^{\infty} f(x-y)g(y)dy$$

We use the proposition, that, if  $X$  and  $Y$  are independent random variables with PDFs  $f$  and  $g$ , then  $X + Y$  has the probability density function  $f \otimes g$ . Per definition the sum of 2 exponential distributions with the same parameter  $\lambda$  gives the Erlang 2 distribution (with parameter  $\lambda$ ). Then because these functions are only nonzero for nonnegative arguments the PDF of the Erlang 2 distribution is calculated as:

$$\begin{aligned} g(x, 2, \lambda) &= (f \otimes f)(x, \lambda) \\ &= \int_0^x \lambda e^{-\lambda(x-y)} \lambda e^{-\lambda y} dy \\ &= \lambda^2 \int_0^x e^{-\lambda x} dy \\ &= \lambda^2 x e^{-\lambda x} \end{aligned}$$

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Using the formula from the script for the Erlang- $k$  distribution (3.2.3) we get:

$$\begin{aligned} g(x) &= \frac{(\lambda x)^{k-1}}{(k-1)!} \cdot \lambda \cdot e^{-\lambda x} \quad \text{for } k = 2 \\ &= \frac{(\lambda x)^{2-1}}{(2-1)!} \cdot \lambda \cdot e^{-\lambda x} \\ &= \frac{\lambda x}{1!} \cdot \lambda \cdot e^{-\lambda x} \\ &= \lambda x \cdot \lambda \cdot e^{-\lambda x} \\ &= \lambda^2 x e^{-\lambda x} \end{aligned}$$

Which shows that the two formulas are identical.

✓

### Problem 6.2.3

With  $p = \frac{1}{n}$  and the rate  $p \cdot \lambda$  the pdf of the exponential distribution is:

$$f(x; \lambda) = \frac{\lambda}{n} e^{-\frac{\lambda}{n}x}$$

We get this by using the conditional probability that  $D_B$  takes  $k$  exponential inter-arrival times:

$$\begin{aligned} p(D_B) &= \sum_{i=1}^k p(D_B|e_i) \cdot p(e_i) \\ &= \sum_{i=1}^k \frac{1}{n} e^{-\frac{1}{n}x} \cdot (\lambda_i e^{-\lambda_i x}) \\ &= \sum_{i=1}^k \frac{\lambda_i}{n} e^{-\frac{\lambda_i}{n}x} \end{aligned}$$

nicht definiert

f

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$$= \frac{k \lambda_i}{n} e^{-\frac{\lambda_i}{n}x} \neq \frac{\lambda}{n} e^{-\frac{\lambda}{n}x} \quad \Delta$$

### Problem 6.2.4

Approach A: Erlang distribution with parameters  $n$  and  $\lambda$

$$\begin{aligned} \mathbb{E}[D_A] &= \frac{n}{\lambda} \\ \text{Var}[D_A] &= \frac{n}{\lambda^2} \\ c_{var}[D_A] &= \frac{\sqrt{\text{Var}(D_A)}}{\mathbb{E}(D_A)} \\ &= \frac{\sqrt{\frac{n}{\lambda^2}}}{\frac{n}{\lambda}} \\ &= \frac{1}{\sqrt{n}} \end{aligned}$$

f

Approach B: Exponential distribution with parameters  $p \cdot \lambda$

$$\begin{aligned}\mathbb{E}[D_B] &= \frac{1}{p \cdot \lambda} \\ \mathbb{V}ar[D_B] &= \frac{1}{p^2 \cdot \lambda^2} \\ c_{var}[D_B] &= \frac{\sqrt{\mathbb{V}ar(D_B)}}{\mathbb{E}(D_B)} \\ &= \frac{\sqrt{\frac{1}{p^2 \cdot \lambda^2}}}{\frac{1}{p \cdot \lambda}} \\ &= 1\end{aligned}$$

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✓

### Problem 6.2.5

The mean time between two extracted packets of both approaches is the same:

$$\frac{1}{p \cdot \lambda} = \frac{1}{\frac{1}{n} \lambda} = \frac{n}{\lambda}$$

But because the coefficient of variation for approach A is usually lower than for approach B, the latter needs a faster packet processing engine.

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✓

### Problem 6.2.6

We proof the claims by squaring the terms.

For approach A the discrete random variable N is a deterministic one, because only every n-th packet is extracted. Therefore we obtain:

$$\begin{aligned}c_{var}^n[A]^2 &= \frac{\mathbb{V}ar^n[A]}{\mathbb{E}^n[A]^2} \\ &= \frac{\mathbb{E}[N]\mathbb{V}ar[X] + \mathbb{V}ar[N]\mathbb{E}[X]^2}{\mathbb{E}[N]^2\mathbb{E}[X]^2} \\ &= \frac{n\mathbb{V}ar[X] + 0\mathbb{E}[X]^2}{n^2\mathbb{E}[X]^2} \\ &= \frac{\mathbb{V}ar[X]}{n\mathbb{E}[X]^2} \\ &= \frac{c_{var}[X]^2}{n}\end{aligned}$$

✓

For approach B the discrete random variable N is geometric distributed with  $p = \frac{1}{n}$ .

$$\begin{aligned}
 c_{var}^n[B]^2 &= \frac{\mathbb{V}ar^n[B]}{\mathbb{E}^n[B]^2} \\
 &= \frac{\mathbb{E}[N]\mathbb{V}ar[X] + \mathbb{V}ar[N]\mathbb{E}[X]^2}{\mathbb{E}[N]^2\mathbb{E}[X]^2} \\
 &= \frac{\frac{1}{p}\mathbb{V}ar[X] + \frac{1-p}{p^2}\mathbb{E}[X]^2}{(\frac{1}{p})^2\mathbb{E}[X]^2} \\
 &= \frac{n\mathbb{V}ar[X] + (1 - \frac{1}{n})n^2\mathbb{E}[X]^2}{n^2\mathbb{E}[X]^2} \\
 &= \frac{\frac{\mathbb{V}ar[X]}{\mathbb{E}[X]^2} + (1 - \frac{1}{n})}{n} \quad - 2 \\
 &= \frac{c_{var}[X]^2 + (1 - \frac{1}{n})}{n} \quad (\checkmark)
 \end{aligned}$$

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# Problem 6.2.7

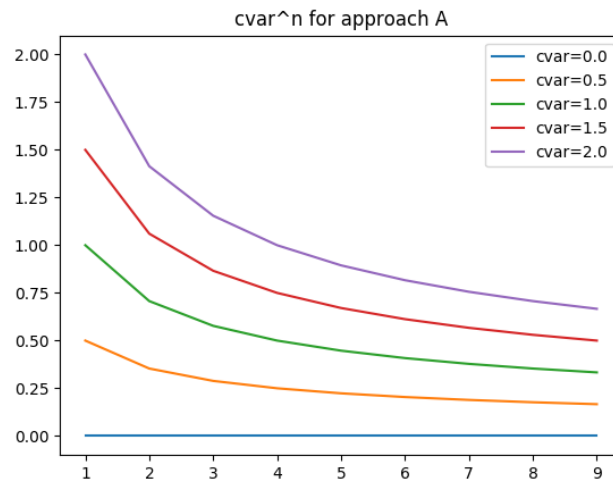


Figure 1: Plot for Approach A

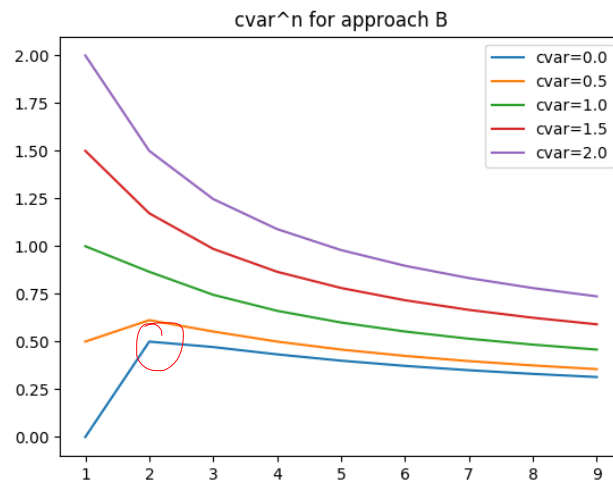


Figure 2: Plot for Approach B

Diff game idij

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**Problem 6.2.8**

This exercise was moved to part II of the lecture.