

$$\frac{1}{\binom{N}{n}} \sum_{x=2}^n \frac{x(x-1) a(a-1)}{x(x-1)(x-2)!(a-x)!} \binom{N-a}{n-x} + \frac{na}{N}$$

$$\frac{a(a-1)}{\binom{N}{n}} \sum_{x=2}^n \frac{(a-2)!}{(x-2)!(a-x)!} \binom{N-a}{n-x} + \frac{na}{N}$$

$y = x - 2$, $x = y + 2$ şeklinde değişken değiştirilir.

$$\frac{a(a-1)}{\binom{N}{n}} \sum_{y=0}^{n-2} \binom{a-2}{y} \binom{N-a}{n-(y+2)} + \frac{na}{N} = \frac{a(a-1)}{\binom{N}{n}} \sum_{y=0}^{n-2} \binom{a-2}{y} \binom{N-a}{n-(y+2)}$$

$$+ \frac{na}{N} = \frac{a(a-1)}{\binom{N}{n}} \binom{N-2}{n-2} + \frac{na}{N}$$

$$= \frac{a(a-1)n(n-1)}{N(N-1)} + \frac{na}{N} = \frac{na[(a-1)(n-1)(N-1)]}{N(N-1)}$$

$$\text{Var}(x) = \frac{na(na-a-n+N)}{N(N-1)} - \frac{n^2 a^2}{N^2} = \frac{N-n}{N-1} \frac{na}{N} \left(1 - \frac{a}{N}\right)$$

POISSON DAĞILIMI

$$f(x) = \begin{cases} \frac{e^{-\lambda} \lambda^x}{x!} & \text{if } x \in \mathbb{R} \\ 0 & \text{else} \end{cases}$$

Beklenen Değer

$$E(x) = \sum_{x=0}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} x \frac{\lambda^x}{x!} = e^{-\lambda} \lambda \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!}$$

$$\frac{\lambda^{x-1}}{x(x-1)!} = e^{-\lambda} \lambda \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!} = e^{-\lambda} \lambda \cdot e^{\lambda} = \lambda$$

Varyans

$$\begin{aligned} E(x^2) &= \sum_{n=0}^{\infty} x^2 \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum_{n=0}^{\infty} x \frac{\lambda^x}{x!} = e^{-\lambda} \lambda \sum_{n=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!} \\ &= e^{-\lambda} \lambda \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} = e^{-\lambda} \lambda e^{\lambda} = \lambda \end{aligned}$$

MOMENT ÜRETEN FONK

$$\begin{aligned} M_x(t) = E[e^{xt}] &= \sum_{x=0}^{\infty} e^{xt} \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!} \\ &= e^{-\lambda} e^{\lambda e^t} = e^{\lambda(e^t - 1)} \end{aligned}$$

Kümülan ÜRETEN FONK

$$K(t) = \ln(M_x(t)) = \frac{d^n \ln(M_x(t))}{dt^n} = \lambda(e^t - 1)$$

ÇARPIMSAL KARAKTERİSTİK FONK $t=1$

$$\begin{aligned} N_x(t) = E(x^x) &= \sum_{x \in \mathbb{R}} x^x \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum_{x \in \mathbb{R}} \frac{(\lambda t)^x}{x!} = e^{\lambda t} \\ &= e^{-\lambda} e^{\lambda t} = e^{\lambda(t-1)} \end{aligned}$$

$$\frac{d^n N_x(t)}{dt^n} = \frac{d^n (e^{\lambda(t-1)})}{dt^n} \Big|_{t=1} = \lambda^n$$

ÇARPIKLIK

$$\frac{\lambda^3(t)}{\lambda^2(t)} = \frac{\lambda}{\lambda^2} = \frac{1}{\lambda}$$

BASIKLIK KURTOSİS

$$K_4(t) \neq E((x - E(x))^4)$$

Hepsi ni de,

$$E[X^4] - 4E[X^3]E[X] + 6E[X^2]E[X]^2 - 3(E[X])^4$$

$$= \frac{\lambda + 3\lambda^2}{\lambda^2} = \frac{1}{\lambda} + 3$$

Bir birim zaman veya alanda X rd λ parametresi ile poisson dağılıyan ise a birim zaman, ad yada alanda X rd a, λ parametresi ile poisson dağılır.

UNIFORM (DÜZGÜN) DAĞILIM

$$f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{diğer} \end{cases}$$

şeklinde verilmiş ise X rd (a,b) kapalı aralıktaki düzgün veya uniform dağılımıdır denir.

i) $X \in (a,b)$ ise $f(x) = \frac{1}{b-a} > 0$ $b > a$ old göre

ii) $\int_{-\infty}^{\infty} f(x) dx = \int_a^b \frac{1}{b-a} dx = \frac{x}{b-a} \Big|_a^b = \frac{b}{b-a} - \frac{a}{b-a} = 1$ } Ort. İspatı

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx = \int_a^b x \frac{1}{b-a} dx = \frac{x^2}{2(b-a)} \Big|_a^b = \frac{b^2 - a^2}{2(b-a)} = \frac{(b-a)(a+b)}{2(b-a)} = \frac{a+b}{2}$$

$$E(X^2) = \int_{-\infty}^{\infty} x^2 f(x) dx = \int_a^b x^2 \frac{1}{b-a} dx = \frac{x^3}{3(b-a)} \Big|_a^b = \frac{b^3 - a^3}{3(b-a)} = \frac{(b-a)(a^2 + ab + b^2)}{3(b-a)} = \frac{a^2 + ab + b^2}{3}$$

$$\text{Var}(X) = \frac{a^2 + ab + b^2}{3} - \frac{a^2 + b^2}{6} = \frac{b^2 - 2ab + a^2}{12} = \frac{(b-a)^2}{12}$$

MOMENT ÜRETEN FONK

$$M_X(t) = E(e^{tX}) = \int_0^b e^{xt} \cdot \frac{1}{b-a} dx = \int_0^b \frac{e^{xt}}{b-a} dx = \frac{e^{bt} - e^{at}}{t(b-a)}$$

(α, β) GAMA DAĞILIMI

şekil dağılım

Olasılık Yoğunluk Fonk

$$f(x) = \begin{cases} \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} & \text{if } \alpha > 0, x > 0, \beta > 0 \\ 0 & \text{otherwise} \end{cases}$$

BEKLENEN DEĞER

$$E(X) = \int_0^\infty x \cdot \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} dx = \frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^\infty x^\alpha e^{-x/\beta} dx$$

* $\frac{x}{\beta} = u, x = u\beta, dx = \beta du$ değişkenleri değiştirelim

$$\frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^\infty (u\beta)^\alpha e^{-u} \beta du = \frac{1}{\Gamma(\alpha)\beta^\alpha} \beta^{\alpha+1} \int_0^\infty u^\alpha e^{-u} du$$

$$\frac{\beta^{\alpha+1} \Gamma(\alpha+1)}{\Gamma(\alpha)\beta^\alpha} = \frac{\beta^\alpha \cdot \beta \cdot \alpha \Gamma(\alpha)}{\Gamma(\alpha)\beta^\alpha} = \alpha\beta$$

VARYANS

$$E(X^2) = \int_0^\infty x^2 \cdot \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} dx = \frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^\infty x^{\alpha+1} e^{-x/\beta} dx$$

* $u = \frac{x}{\beta}, \beta u = x, dx = \beta du$

$$\frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^\infty (u\beta)^{\alpha+1} e^{-u} \beta du = \frac{1}{\Gamma(\alpha)\beta^\alpha} \beta^{\alpha+2} \int_0^\infty u^{\alpha+1} e^{-u} du$$

$$\frac{\beta^{\alpha+2} \Gamma(\alpha+2)}{\Gamma(\alpha)\beta^\alpha} = \frac{\beta^2 \cdot (\alpha+1)\alpha \Gamma(\alpha)}{\Gamma(\alpha)\beta^\alpha} = (\alpha^2 + \alpha) \beta^2 = \alpha^2 \beta^2 + \alpha \beta^2$$

$$\text{Var}(x) = \alpha^2 \beta^2 + \alpha \beta^2 - \alpha^2 \beta^2 = \alpha \beta^2$$

MOMENT ÜRETEK FONKSİYON

$$M_X(t) = E(e^{xt}) = \int_0^{\infty} e^{xt} \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} dx$$

$$\frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^{\infty} x^{\alpha-1} e^{xt - \frac{x}{\beta}} dx = \frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^{\infty} x^{\alpha-1} e^{-x(1-t)/\beta} dx$$

$$v = \frac{x(1-t)}{\beta}, \quad \frac{\beta v}{(1-t)} = x, \quad dx = \frac{\beta dv}{(1-t)}$$

$$\frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^{\infty} \left(\frac{\beta v}{1-t}\right)^{\alpha-1} e^{-v} \frac{\beta dv}{(1-t)} = \frac{1}{(1-t)} \frac{(1-t)^\alpha}{\beta} \frac{\beta}{(1-t)} \frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^{\infty} v^{\alpha-1} e^{-v} dv$$

$$\frac{1}{\Gamma(\alpha)\beta^\alpha} \frac{\beta^\alpha}{(1-t)^\alpha} \Gamma(\alpha) = \left(\frac{1}{1-t}\right)^\alpha$$

$M_X(t)$ ile 1. ve 2. momentleri ile varyans bulmak

1. MOMENT

$$\frac{d^1 M_X(t)}{dt} = \frac{d^1 (1-t)^{-\alpha}}{dt} = -\alpha(1-t)^{-\alpha-1} \cdot (-1) = \alpha(1-t)^{-\alpha-1} \Big|_{t=0} = \alpha\beta$$

2. MOMENT

$$\frac{d^2 M_X(t)}{dt^2} = \frac{d^2 (1-t)^{-\alpha}}{dt^2} = \alpha\beta(-\alpha-1)(1-t)^{-\alpha-2} \cdot (-1) = \alpha\beta^2(-\alpha-1)(1-t)^{-\alpha-2}$$

$$\alpha^2 \beta^2 - \alpha \beta^2 (1-t)^{-\alpha-2} \Big|_{t=0} = \alpha^2 \beta^2 - \alpha \beta^2$$

VARYANS

$$\text{Var}(x) = \alpha^3 \beta^2 - \alpha \beta^2 - \alpha^2 \beta^2 = \alpha \beta^2$$

$$f(x) = \int x^{\alpha-1} e^{-x/\beta} dx$$

Gamma fonk

Gamma Fonk yap

ÇARPILIK (SKEWNESS)

KARAKTERİSTİK FONK

$$\varphi(t) = E(e^{it}) = \int_0^{\infty} e^{itx} \frac{1}{\Gamma(\alpha)\beta^{\alpha}} x^{\alpha-1} e^{-x/\beta} dx$$

$$\frac{1}{\Gamma(\alpha)\beta^{\alpha}} \int_0^{\infty} x^{\alpha-1} e^{itx} e^{-x/\beta} dx = \frac{1}{\Gamma(\alpha)\beta^{\alpha}} \int_0^{\infty} x^{\alpha-1} e^{-(1-i\beta t)x/\beta} dx$$

$$u = x(1-i\beta t)/\beta \quad \frac{du}{dx} = \frac{1-i\beta t}{\beta} \quad dx = \beta du$$

$$\frac{1}{\Gamma(\alpha)\beta^{\alpha}} \int_0^{\infty} \left(\frac{\beta u}{1-i\beta t}\right)^{\alpha-1} e^{-u} \beta du = \frac{\beta}{(1-i\beta t)^{\alpha}} \frac{1}{\Gamma(\alpha)\beta^{\alpha}} \int_0^{\infty} u^{\alpha-1} e^{-u} du$$

$$\int_0^{\infty} u^{\alpha-1} e^{-u} du = \Gamma(\alpha)$$

KÜMÜLANT ÜRETEN FONK

$$\ln\left(\ln\left(\frac{1}{1-F(x)}\right)\right) = \ln(\ln(1-(1-e^{-x/\beta})^{\alpha}))$$

(E(1)) Kümlant Üreten Fonk 1. Wöreri

$$k'(x) = \frac{d}{dx} \ln(1-F(x)) = \frac{d}{dx} \ln(1-(1-e^{-x/\beta})^{\alpha})$$

$$M(x) = -\alpha(1-i\beta t)^{\alpha-1} \cdot \beta = -\alpha\beta(1-i\beta t)^{\alpha-1}$$

$$k''(x) = \frac{d}{dx} k'(x) = \frac{d}{dx} \left(\frac{-\alpha\beta(1-i\beta t)^{\alpha-1}}{(1-i\beta t)^{\alpha}} \right) = \frac{-\alpha\beta}{(1-i\beta t)} \Big|_{t=0} = -\alpha\beta$$

Var(1) Kümlant Üreten Fonk 2. Wöreri

$$k''(x) = \frac{d}{dx} \ln(1-F(x)) = \frac{d}{dx} \ln(1-(1-e^{-x/\beta})^{\alpha})$$

ÇARPILIK

$$\text{Var}(x) = \alpha\beta^2, k^2(t) = \frac{\alpha\beta^2}{(1-i\beta t)^2}$$

$$k^3(t) = \frac{d}{dt} \frac{\alpha\beta^2}{(1-i\beta t)^2} = \frac{0(1-i\beta t)^2 - 2(1-i\beta t)\alpha\beta^2(-i\beta)}{(1-i\beta t)^4}$$

$$= \frac{2\alpha\beta^3}{(1-i\beta t)^3} \Big|_{t=0} = 2\alpha\beta^3$$

BASIKLIK

$$k^4(t) = \frac{d}{dt} \frac{2\alpha\beta^3}{(1-i\beta t)^3} = \frac{0(1-i\beta t)^3 - 3(1-i\beta t)^2(-i\beta)2\alpha\beta^3}{(1-i\beta t)^6}$$

$$= \frac{6\alpha\beta^4}{(1-i\beta t)^4} = \frac{6\alpha\beta^4}{(1-i\beta t)^4} \Big|_{t=0} = 6\alpha\beta^4$$

$$\text{Basiklik} = \frac{k^4(t)}{(\text{Var}(x))^2} = \frac{6\alpha\beta^4}{(\alpha\beta^2)^2} = \frac{6\alpha}{\alpha^2} = \frac{6}{\alpha}$$

ÜSTEL DAĞILIM

$$f(x) = \begin{cases} \frac{1}{\theta} e^{-x/\theta} & x > 0, \theta > 0 \\ 0 & \text{diğer} \end{cases}$$

GAMA ~ ÜSTEL İLİŞKİ'Sİ

$$\text{Gamma}(\alpha, \beta) \Rightarrow \alpha=1, \beta=\theta = \text{Üstel}(\theta)$$

$$x \sim \frac{1}{\Gamma(1)\theta} e^{-x/\theta} = \text{Üstel Dağılımı}$$

Gama

$$E(x) = \alpha\beta$$

$$E(\beta) = \alpha(\alpha+1)(\alpha+2)\beta^3$$

$$\text{Skew} = \frac{2}{\alpha}$$

$$k(t) = \ln(1-\beta t)^{-\alpha}$$

$$\alpha=1$$

$$\beta=\theta$$

Üstel

$$E(x) = \theta$$

$$E(x^3) = 6\theta^3$$

$$\text{Skew} = 2$$

$$k(t) = -\ln(1-\theta t)$$

İstatistik vs Mühendislik

$$X \sim \text{Üstel}(\theta) \quad \begin{cases} \frac{1}{\theta} e^{-\frac{x}{\theta}} & x > 0, \theta > 0 \\ 0 & \text{diğer} \end{cases}$$

$$X \sim \text{Üstel}\left(\frac{1}{\theta}\right) \quad \begin{cases} \theta e^{-x\theta} & x > 0, \theta > 0 \\ 0 & \text{diğer} \end{cases}$$

$$X \sim \text{Gamma}(\alpha, \beta) \quad \begin{cases} \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} & x > 0, \beta > 0, \alpha > 0 \\ 0 & \text{diğer} \end{cases}$$

$$X \sim \text{Gamma}\left(\alpha, \frac{1}{\beta}\right) \quad \begin{cases} \frac{1}{\Gamma(\alpha)} x^{\alpha-1} e^{-x\beta} & x > 0, \beta > 0, \alpha > 0 \\ 0 & \text{diğer} \end{cases}$$

$$f(x) = \begin{cases} \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-x\beta} & x > 0, \beta > 0, \alpha > 0 \\ 0 & \text{diğer} \end{cases}$$

ÜSTEL DAĞILIM - DAĞILIM FONK

$$f(x) = \frac{dF(x)}{dx} \quad F(x) = P(X \leq x) = \int_0^x \frac{1}{\theta} e^{-\frac{t}{\theta}} dt$$

$$\frac{t}{\theta} = u \quad dt = \theta du \quad t \rightarrow 0 \rightarrow u = 0 \quad t = x \rightarrow u = \frac{x}{\theta}$$

$$\frac{1}{\theta} \int_0^{x/\theta} e^{-u} \theta du = e^{-u} \Big|_0^{x/\theta} = 1 - e^{-\frac{x}{\theta}}$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-\frac{x}{\theta}} & 0 < x < \infty \\ 1 & x \rightarrow \infty \end{cases}$$

$$X \sim \text{ÜSTEL}(\theta) \cong \text{Gamma}(1, \theta)$$

Ki-Kare Dağı (χ²)

$$x \text{ nd ağıt } f(x) = \begin{cases} \frac{1}{\Gamma(\frac{n}{2})2^{\frac{n}{2}}} x^{\frac{n}{2}-1} e^{-\frac{x}{2}} & x > 0, n > 0 \\ 0 & \text{diğer} \end{cases}$$

şeklinde ise ki-kare Dağılımına sahiptir denir, ve $X \sim \chi^2(n)$ diye edilir.

$$\text{Gamma}(\alpha, \beta) \cong \alpha = \frac{n}{2} \quad \beta = 2 \quad \text{Gamma}\left(\frac{n}{2}, 2\right) \cong \chi^2(n)$$

$$\text{Gamma}(1, 2) \cong \chi^2(2) \quad \text{Gamma}\left(\frac{2}{2}, 2\right) = \chi^2(2)$$

-NOT: Parametre çıkmasını sebebi: deneysel dağılımı, olasılık, chi², normal, t, F dağı

$$\text{Gamma}(2, 4) \rightarrow \beta \text{ hep 2 almalı}$$

NOT: analitik olarak $F(x)$ 'i yoktur

Gamma - Kikare

Gamma(α, β)

$$E(x) = \alpha \cdot \beta$$

$$E(x^2) = \alpha(\alpha+1)\beta^2$$

$$\text{Var}(x) = \alpha\beta^2$$

$$M_x(t) = (1 - \beta t)^{-\alpha}$$

$$k_x(t) = -\alpha(1 - \beta t)^{-\alpha-1}$$

$$\text{Skew}(x) = \frac{2}{\alpha}$$

$$\text{Kurt}(x) = \frac{6}{\alpha}$$

$$\text{Kurt}(x) = \frac{6}{\alpha}$$

NOT: Ki-kare'nin $F(x)$ 'i yoktur

Ki-kare $\chi^2(n) \cong \text{Gamma}(\frac{n}{2}, 2)$

$$E(x) = \frac{n}{2} \cdot 2 = n$$

$$E(x^2) = \frac{n}{2}(\frac{n}{2}+1)2^2 = n(n+2)$$

$$E(x^3) = \frac{n}{2}(\frac{n}{2}+1)(\frac{n}{2}+2)2^3$$

$$\text{Var}(x) = 2n$$

$$M_x(t) = (1 - 2t)^{-\frac{n}{2}}$$

$$M_x(t) = (1 - 2t)^{-\frac{n}{2}}$$

$$\text{Skew}(x) = \frac{3}{\sqrt{n}}$$

Beta Dağılımı

Beta Fonk.

$$\text{Beta}(\alpha, \beta) = \int_0^1 x^{\alpha-1} \cdot (1-x)^{\beta-1} dx \quad \alpha > 0, \beta > 0$$

$$= \frac{\Gamma(\alpha) \Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$$\text{Beta}(\alpha, \beta) = \frac{\Gamma(\alpha) \Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$$\int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx$$

$$(\alpha+\beta) = \sum_{k=0}^{\infty} \binom{\beta-1}{k} x^{\alpha+k} \cdot (-1)^k$$

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx$$

$$\text{Beta}(\alpha+1, \beta) = \frac{\Gamma(\alpha+1) \Gamma(\beta)}{\Gamma(\alpha+\beta+1)}$$

$$= \frac{\alpha \Gamma(\alpha) \Gamma(\beta)}{(\alpha+\beta) \Gamma(\alpha+\beta)} \text{Beta}(\alpha, \beta)$$

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx$$

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx$$

Bu şekilde Beta fonk. en ufak e ile edebilir.

Beta Dağılımı

$$f(x) = \begin{cases} \frac{1}{\text{Beta}(\alpha, \beta)} x^{\alpha-1} \cdot (1-x)^{\beta-1} & 0 < x < 1, \alpha > 0, \beta > 0 \\ 0 & \text{diğer} \end{cases}$$

X'nin beta dağılımı sahiptir denir ve kısaca $X \sim \text{Beta}(\alpha, \beta)$ şeklinde gösterilir.

Burada α parametresi şekil parametresidir.

Beta dağılımı en küçük beklenti 0,1 gibi

sonlu aralıkta tanımlıdır.

BETA DAĞI E(x), Var(x),

$$E(x) = \int_0^1 x^{\alpha} \cdot \frac{1}{\text{Beta}(\alpha, \beta)} \cdot x^{\alpha-1} \cdot (1-x)^{\beta-1} dx$$

$$\frac{1}{\text{Beta}(\alpha, \beta)} \int_0^1 x^{\alpha+1-1} \cdot (1-x)^{\beta-1} dx = \frac{\text{Beta}(\alpha+1, \beta)}{\text{Beta}(\alpha, \beta)}$$

$$\text{Beta}(\alpha+1, \beta) = \frac{\Gamma(\alpha+1) \Gamma(\beta)}{\Gamma(\alpha+\beta+1)}$$

$$\text{Beta}(\alpha, \beta) = \frac{\Gamma(\alpha) \Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$E(x)$ için

$$\frac{\Gamma(\alpha+1) \Gamma(\beta)}{\Gamma(\alpha+\beta+1)} \cdot \frac{1}{\Gamma(\alpha) \Gamma(\beta)} = \frac{\alpha \Gamma(\alpha) \Gamma(\beta)}{(\alpha+\beta) \Gamma(\alpha) \Gamma(\beta)} = \frac{\alpha}{\alpha+\beta}$$

$E(x^2)$ için

$$\frac{\Gamma(\alpha+2) \Gamma(\beta)}{\Gamma(\alpha+\beta+2)} \cdot \frac{1}{\Gamma(\alpha) \Gamma(\beta)} = \frac{\alpha(\alpha+1) \Gamma(\alpha) \Gamma(\beta)}{(\alpha+\beta)(\alpha+\beta+1) \Gamma(\alpha) \Gamma(\beta)} = \frac{\alpha(\alpha+1)}{(\alpha+\beta)(\alpha+\beta+1)}$$

$$\text{Var}(x) = \frac{\alpha(\alpha+1)}{(\alpha+\beta)(\alpha+\beta+1)} - \left(\frac{\alpha}{\alpha+\beta} \right)^2 = \frac{\alpha \beta}{(\alpha+\beta)^2 (\alpha+\beta+1)}$$

$$\frac{\alpha(\alpha+1) \beta}{(\alpha+\beta)^2 (\alpha+\beta+1)} = \frac{\alpha \beta}{(\alpha+\beta)^2 (\alpha+\beta+1)}$$

$$x^2 \sim \text{Beta}(\alpha+2, \beta)$$

$$X \sim \text{Beta}(\alpha, \beta) \approx U(0,1)$$

$$\alpha=1, \beta=1$$

CAUCHY DAG

$$f(x) = \begin{cases} \frac{1}{\pi} \cdot \frac{1}{1+(x-\theta)^2} & -\infty < x < \infty, x \in \mathbb{R} \\ 0 & \text{and} \end{cases}$$

skilade de X rel cauchy dag skript 1.

$$\int_{-\infty}^{\infty} \frac{1}{1+x^2} \cdot \arctan(x) \cdot \frac{1}{1+(x-\theta)^2} dx = \arctan(x-\theta)$$

$$i) f(x) \geq 0$$

$$f(x) = \frac{1}{\pi} \cdot \frac{1}{1+(x-\theta)^2} > 0 \quad x > 0 \quad \begin{cases} f(x) \geq 0 \\ \text{do } x \in \mathbb{R} \end{cases}$$

$$ii) \int_{-\infty}^{\infty} f(x) = 1$$

$$\int_{-\infty}^{\infty} \frac{1}{\pi} \cdot \frac{1}{1+(x-\theta)^2} = \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{1}{1+(x-\theta)^2} = \frac{1}{\pi} \arctan(x-\theta)$$

$$\lim_{x \rightarrow \infty} \frac{1}{\pi} [\arctan(x-\theta)] \quad \lim_{x \rightarrow -\infty} \frac{1}{\pi} [\arctan(x-\theta)]$$

$$= \frac{1}{\pi} \left[\frac{\pi}{2} - \left(-\frac{\pi}{2}\right) \right] = \frac{\pi}{\pi} = 1$$

$$\begin{aligned} \tan x &= y \\ \arctan y &= x \\ x &= \arctan(y) \end{aligned}$$

$$\begin{aligned} \cos x (90^\circ) &= \frac{\pi}{2} > 0 \\ \arctan(90^\circ) &= \frac{\pi}{2} \end{aligned}$$

Cauchy F(x)

$$\begin{aligned} F(x) &= P(X \leq x) = \int_{-\infty}^x f(t) dt = \int_{-\infty}^x \frac{1}{\pi} \cdot \frac{1}{1+(t-\theta)^2} dt \\ &= \frac{1}{\pi} (\arctan(x-\theta)) \Big|_{-\infty}^x = \frac{1}{\pi} \left[\arctan(x-\theta) - \left(-\frac{\pi}{2}\right) \right] \\ &= \frac{1}{\pi} \arctan(x-\theta) + \frac{1}{2} \end{aligned}$$

$$F(x) = \begin{cases} 0 & x > -\infty \\ \frac{1}{\pi} \arctan(x-\theta) + \frac{1}{2} & x \in \mathbb{R} \\ 1 & x \rightarrow \infty \end{cases}$$

parametriserad
generell
funktion

$F(x)$, $f(x)$
generell
funktion

$$X \sim \mathcal{C}(\frac{\pi}{2}, 2)$$

$$E(X) = \frac{\pi}{2} \cdot 2 = \pi$$

$$\text{Var}(X) = \frac{\pi}{2} \cdot 4 = 2\pi$$

$$M_X(t) = (1-2t)^{\frac{\pi}{2}}$$

$$K(t) = \frac{\pi}{2} \ln(1-2t)$$

$$\text{Skew} = \frac{2\sqrt{2}}{\sqrt{\pi}}$$

$$\text{kurtosis} = \frac{12}{\pi}$$

$$X \sim \text{Cauchy}(\alpha, \beta)$$

$$E(X) = \alpha$$

$$\text{Var}(X) = \alpha^2 \beta^2$$

$$M_X(t) = (1-t^2 \beta^2)^{-\frac{\pi}{2}}$$

$$K(t) = -\alpha \ln(1-t^2 \beta^2)$$

$$\text{Skew} = \frac{2}{\sqrt{\pi}}$$

$$\text{kurtosis} = \frac{6}{\pi}$$

$$X \sim \text{Cauchy}(\theta, 1)$$

$$E(X) = 1 \cdot \theta = \theta$$

$$\text{Var}(X) = 1 \cdot \theta^2 = \theta^2$$

$$M_X(t) = (1-t^2)^{-\frac{\pi}{2}}$$

$$K(t) = -\ln(1-t^2)$$

$$\text{Skew}(X) = \frac{2}{\sqrt{\pi}}$$

$$\text{kurtosis} = \frac{6}{\pi}$$

NORMAL DAĞILIM

O.Y.F.
 $f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$

STANDARD NORMAL DAĞILIM

$$Z \sim N(0,1)$$

$\mu=0$, varyansı 1 den, $z = \frac{x-\mu}{\sigma}$ seklinde standartlaştırılır.
 $\sigma^2=1$
 $\mu=0$

$$f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, z \in \mathbb{R}$$

BEKLENEN DEĞER

$$E[X] = \int_{-\infty}^{\infty} x \cdot f(x) dx = \int_{-\infty}^{\infty} x \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

$$\frac{1}{\sqrt{2\pi}} \left[\int_{-\infty}^0 x \cdot e^{-\frac{x^2}{2}} dx + \int_0^{\infty} x \cdot e^{-\frac{x^2}{2}} dx \right] \quad \begin{matrix} e^{-x} \cdot x \rightarrow 0 \\ e^{-x} \cdot x = 0 \end{matrix}$$

$$\frac{1}{\sqrt{2\pi}} \left[\int_{-\infty}^0 x \cdot e^{-\frac{x^2}{2}} dx + \int_0^{\infty} x \cdot e^{-\frac{x^2}{2}} dx \right] = \frac{1}{\sqrt{2\pi}} [0-1] + \frac{1}{\sqrt{2\pi}} [1+0] = 0$$

$$E[X] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \left[-e^{-\frac{x^2}{2}} \right]_{-\infty}^{\infty} = 0$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \left[-e^{-\frac{x^2}{2}} \right]_{-\infty}^{\infty} = 0$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx = 0$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx = 0$$

Moment Üreten Fonksiyon

$$M_X(t) = E(e^{xt}) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} e^{xt} dx \quad \begin{matrix} \sqrt{2\pi}\sigma = 1, \mu = x \\ t = \frac{x-\mu}{\sqrt{2\pi}\sigma} \end{matrix}$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{xt} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{xt} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx =$$

$$\frac{e^{\mu t}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-(u-\frac{\sqrt{2\pi}\sigma t}{2})^2} \frac{2\sigma^2 t}{2} du$$

Kümülatif Üreten Fonksiyon

X rd kümülatif üreten fonksiyonu $K_X(t)$

$K_X(t) = \ln(M_X(t))$ olarak tanımlanır.

$X \sim N(\mu, \sigma^2)$ için $M_X(t) = e^{t\mu + \frac{t^2\sigma^2}{2}}$

$$K_X(t) = \ln(M_X(t)) = \ln(e^{t\mu + \frac{t^2\sigma^2}{2}}) = t\mu + \frac{t^2\sigma^2}{2}$$

Benzerlik, Momentler

$$\frac{d}{dt} (M_X(t)) \Big|_{t=0} = E(X) \text{ olarak elde edilir.}$$

$X \sim N(\mu, \sigma^2)$ için $M_X(t) = e^{t\mu + \frac{t^2\sigma^2}{2}}$ old

$$\frac{d}{dt} (M_X(t)) = (\mu + t\sigma^2) e^{t\mu + \frac{t^2\sigma^2}{2}}$$

$$E(X) = \frac{d}{dt} (M_X(t)) \Big|_{t=0} = (\mu + t\sigma^2) e^{t\mu + \frac{t^2\sigma^2}{2}} \Big|_{t=0} = \mu$$

2.10.0 X rassal değişkenin ortan $kx(1)^{1/2}$ ver ise

$$\frac{d}{dx} (kx(1)^{1/2}) = kx \quad E(X) = k_1 \quad Var(X) = k_2$$

$$E(X - E(X))^3 = k_3 \quad E(X - E(X))^4 = k_4 + 3(k_2)^2$$

X ~ N(μ, σ²) için
 (Çarpıklık) (Skewness) $(X) = \frac{E(X - E(X))^3}{(Var(X))^{3/2}} = \frac{0}{(\sigma^2)^{3/2}} = 0$

$$Basıklık = \frac{E(X - E(X))^4}{(Var(X))^2} = \frac{3\sigma^2}{(\sigma^2)^2} = 3$$

Büyüklik Logit Fonk.

$$f(x) = p(x|x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt$$

$$F(x) = p(x \leq x) = \int_{-\infty}^x f(t) dt$$

ORTAK TANIMLAR

ORTAK OLASILIK DAĞILIM FONKSİYONLARI.

$$F(x,y) = \begin{cases} \frac{\pi^2 f(x,y)}{\pi x y} & f \text{ in tanımlanmış olduğu noktalarda} \\ 0 & \text{diğer} \end{cases}$$

$$\int_{-\infty}^{\infty} f(x) dx + f(y) = 1$$

ORTAK OLASILIK

FONKSİYONU/ YOĞUNLUK FONKSİYONU

ORTAK OLASILIK FONKSİYONU (kesikli)

i) $\forall x \in R, y \in R$ için $f(x,y) \geq 0$

$$ii) \sum_{x \in R} \sum_{y \in R} f(x,y) = 1$$

ORTAK OLASILIK YOĞUNLUK FONKSİYONU (sürekli)

i) $\forall x \in R, y \in R$ için $f(x,y) \geq 0$

$$ii) \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dx dy = 1 \quad \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dy dx = 1$$

MARJİNAL OLASILIK FONKSİYONLARI

Kesikli marjinal olasılık fonksiyonları

$$f(x) = \begin{cases} \sum_{y \in R_y} f(x,y) & \forall x \in R_x \text{ için} \\ 0 & \text{diğer} \end{cases} \quad f(y) = \begin{cases} \sum_{x \in R_x} f(x,y) & \forall y \in R_y \text{ için} \\ 0 & \text{diğer} \end{cases}$$

Sürekli marjinal olasılık yoğunluk fonksiyonları

$$f(x) = \begin{cases} \int_{-\infty}^{\infty} f(x,y) dy & y \in R_y \\ 0 & \text{diğer} \end{cases} \quad f(y) = \begin{cases} \int_{-\infty}^{\infty} f(x,y) dx & x \in R_x \\ 0 & \text{diğer} \end{cases}$$

KOŞULLU OLASILIK / OYF

$$f_{X|Y=y} = f(X|Y) = \begin{cases} \frac{f(X,Y)}{f(Y)} & f(Y) > 0 \\ 0 & \text{diğer} \end{cases}$$

$$f_{Y|X=x} = f(Y|X) = \begin{cases} \frac{f(X,Y)}{f(X)} & f(X) > 0 \text{ ve } y \in D_{Y|X} \\ 0 & \text{diğer} \end{cases}$$

KOVARYANS (iki değişkenin birbirine göre değişimi)

$$\text{COV}(X,Y) = E[(X - E(X))(Y - E(Y))]$$

$$E(XY) - E(X)E(Y) \quad \text{bağımsız ise } 0$$

KORELASYON KATSAYISI doğrusal ilişki

$$P_{X,Y} = \frac{\text{COV}(X,Y)}{\sqrt{\text{Var}(X) \cdot \text{Var}(Y)}} \quad |P_{X,Y}| \leq 1 \quad \text{ise ilişki vardır.}$$

$-1 < P_{X,Y} < 1$ 0 ise bağımsız eğer 0 ise veya yaklaşıyorsa ilişki yok.

Bağımsızlık: $f(X,Y) = f(X) \cdot f(Y)$

Varyans: bir değişkenin varyansı

Verilen ortalamadan sapmaların karesi ortalaması