

# CL205: AI & DS

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# This handout

Cumulative or probability distribution function: CDF

Probability density function: PDF

- Multiple random variables
- Joint, marginal, conditional distribution and density functions
- Independence

# Parallel Notions of Events

- Multiple random variables  $\equiv$  Multiple events
- Joint distribution/density functions  $\equiv$  Probability of intersection of events
- Conditional distribution/density functions  $\equiv$  Conditional probability
- Independent random variables  $\equiv$  independent events

# Slight Change of Notation

- Random variable denoted by upper-case letter (earlier it was bold-face letter):  $X$  (earlier was  $\mathbf{x}$ )
- Lower case english letter (earlier it was greek letter while writing density or distribution functions:  $x$  (earlier was  $\xi$ ).
- For example:  $F_X(x) = P(\{\omega : X(\omega) \leq x\})$
- $\xi, \rho$ , etc. will be used as variables of integration.

## Extension of Ideas:

- Multiple (Multivariate) Random Variables: Jointly distributed random variables
- Event  $\omega$  occurs in sample space  $\Omega$ . Associate many,  $X_1, X_2, \dots, X_n$ , random variables with  $\omega$ .
- Each random variable is a valid mapping from  $\Omega$  to  $\mathbb{R}$ .

# Bivariate Random Variables

- For simplicity of notation consider two random variables:  $X, Y$ .
- Special case of multiple random variables.
- Examples:
  - ▶ Average number of cigarettes smoked daily and the age at which an individual gets cancer,
  - ▶ Height and weight of an individual,
  - ▶ Height and IQ of an individual.
  - ▶ Flow-rate and pressure drop of a liquid flowing through a pipe.
  - ▶ Number of heads and number of tails in an experiment involving toss of several coins.

# Jointly distributed random variables

Often interested in answering questions on  $X, Y$  taking values in a specified region  $D$  in  $\mathbb{R}^2$  (xy plane).

- The distribution functions  $F_X(x)$  and  $F_Y(y)$  of  $X$  and  $Y$  determine their individual probabilities but not their joint probabilities. The probability of event

$$\begin{aligned} & \{\omega : X(\omega) \leq x\} \cap \{\omega : Y(\omega) \leq y\} \\ &= \{\omega : X(\omega) \leq x, Y(\omega) \leq y\} \end{aligned}$$

cannot be expressed in terms of  $F_X(x)$  and  $F_Y(y)$ .

- Joint probabilities of  $X, Y$  completely determined if probability of above event known for every  $x$  and  $y$ .

# Joint Probability Distribution Function or Joint Cumulative Distribution Function

For random variables (discrete or continuous)  $X, Y$ , the joint (bivariate) probability distribution function is:

$$F_{X,Y}(x,y) = P(\{X \leq x, Y \leq y\})$$

RHS to be interpreted as:  $P(\{\omega : X(\omega) \leq x, Y(\omega) \leq y\})$ ,  
where  $x, y$  are two arbitrary real numbers.

Often, the subscript  $X, Y$  omitted.



# Properties of Joint Probability Distribution Function (Papoulis and Pillai, 2002)

①  $F(-\infty, y) = F(x, -\infty) = 0, \quad F(\infty, \infty) = 1.$

②

$$P(x_1 < X \leq x_2, Y \leq y) = F(x_2, y) - F(x_1, y)$$

$$P(X \leq x, y_1 < Y \leq y_2) = F(x, y_2) - F(x, y_1)$$

③

$$P(x_1 < X \leq x_2, y_1 < Y \leq y_2) = \\ F(x_2, y_2) - F(x_1, y_2) - F(x_2, y_1) + F(x_1, y_1)$$

# Joint Density Function: I

- The joint density of  $X$  and  $Y$  is the function (defn.)

$$f(x, y) = \frac{\partial^2 F(x, y)}{\partial x \partial y}$$

[When both  $X, Y$  are continuous random variables]

- It follows that,

$$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(\xi, \rho) d\xi d\rho$$
$$P((X, Y) \in D) = \int \int_D f(x, y) dx dy$$

## Joint Density Function: II

- In particular, as  $\Delta x \rightarrow 0$  and  $\Delta y \rightarrow 0$ ,

$$P(x < X \leq x + \Delta x, y < Y \leq y + \Delta y) \approx f(x, y)\Delta x\Delta y$$

- $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1; f(x, y) \geq 0 \forall x, y \in \mathbb{R}.$

# Joint Density Example: Bivariate Gaussian Random Variable

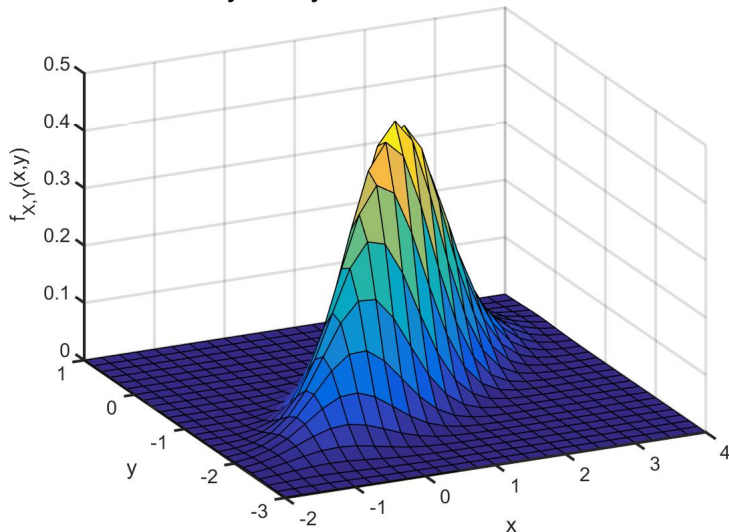
$$f(x, y) = \alpha \exp(-0.5(\xi - \mu)^T P^{-1}(\xi - \mu))$$

with

$$\xi = \begin{bmatrix} x \\ y \end{bmatrix}, \mu = \begin{bmatrix} 1 \\ -1 \end{bmatrix}, P = \begin{bmatrix} 0.9 & 0.4 \\ 0.4 & 0.3 \end{bmatrix},$$
$$\alpha = \frac{1}{2\pi\sqrt{|P|}}$$

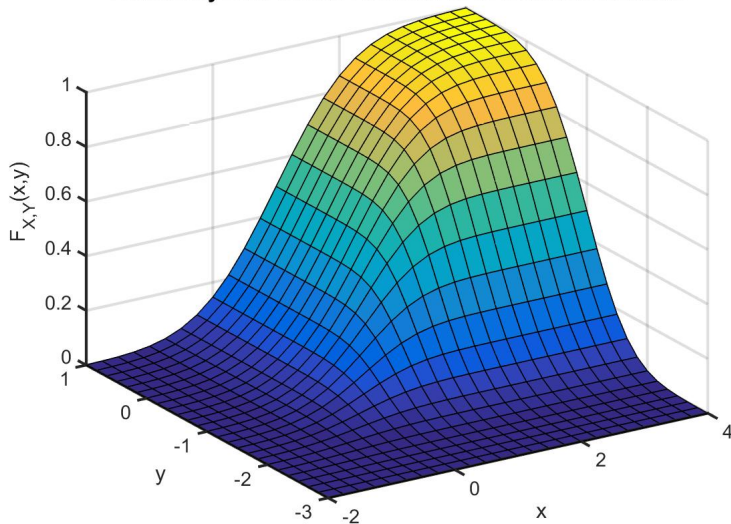
# Joint Density Visualization

Probability Density Function of a Bivariate Gaussian



# Joint Distribution Visualization

**Probability Distribution Function of a Bivariate Gaussian**



# Marginal Distribution or Density Functions of Individual Random Variables: I

- Marginal Probability Distribution Functions:  $F_X(x), F_Y(y)$ :
  - ▶ Extract  $F_X(x)$  from  $F(x, y)$  as:

$$\begin{aligned}F_X(x) &= P(X \leq x) = \\&= P(X \leq x, Y < \infty) = F(x, \infty)\end{aligned}$$

- ▶ Similarly, extract  $F_Y(y)$  as:

$$F_Y(y) = P(Y \leq y) = P(X < \infty, Y \leq y) = F(\infty, y)$$

# Marginal Distribution or Density Functions of Individual Random Variables: II

- Marginal Probability Density Functions:  $f_X(x)$ ,  $f_Y(y)$ :
  - ▶ Extract these from  $f(x, y)$  as:

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$



# Marginal Probability Density

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

Makes sense, since

$$\begin{aligned} P(X \in A) &= P(X \in A, Y \in (-\infty, \infty)) \\ &= \int_A \int_{-\infty}^{\infty} f(x, y) dy dx = \int_A f_X(x) dx \end{aligned}$$

where  $f_X(x)$  is as defined above.

Similarly,

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

## Example 4.3c from Ross: I

$$f(x, y) = \begin{cases} 2e^{-x}e^{-2y} & 0 < x < \infty, \quad 0 < y < \infty \\ 0 & \text{otherwise} \end{cases}$$

Compute: (a)  $P(X > 1, Y < 1)$ , (b)  $P(X < Y)$ , (c)  $P(X < a)$

$$\begin{aligned} P(X > 1, Y < 1) &= \int_0^1 \int_1^\infty 2e^{-x}e^{-2y} dx dy \\ &= \int_0^1 2e^{-2y} (-e^{-x} \big|_1^\infty) dy \\ &= e^{-1} \int_0^1 2e^{-2y} dy = e^{-1}(1 - e^{-2}) \end{aligned}$$

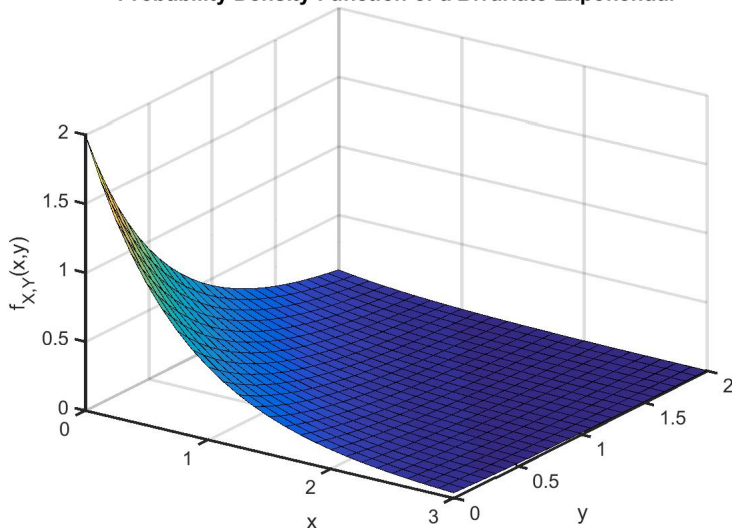
## Example 4.3c from Ross: II

$$P(X < Y) = \int_0^{\infty} \int_0^y 2e^{-x} e^{-2y} dx dy = 1/3$$

$$P(X < a) = \int_0^a \int_0^{\infty} 2e^{-x} e^{-2y} dy dx = 1 - e^{-a}$$

# Joint Density Visualization: Exponential

Probability Density Function of a Bivariate Exponential

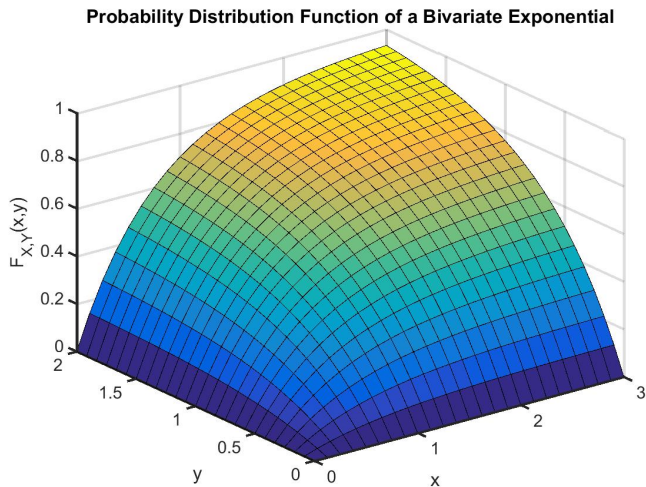


## Example: Joint Distribution Function

For this example:

$$\begin{aligned} F(x, y) &= \int_{-\infty}^x \int_{-\infty}^y f(\xi, \rho) d\xi d\rho \\ &= \begin{cases} (1 - e^{-x})(1 - e^{-2y}), & x > 0, y > 0 \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

# Joint Distribution Visualization: Exponential



# Joint Probability Mass Function (PMF)

- Given two discrete random variables  $X$  and  $Y$  in the same experiment, the joint PMF of  $X$  and  $Y$  is

$$p(x_i, y_j) = P(X = x_i, Y = y_j)$$

for all pairs of  $(x_i, y_j)$  values that  $X$  and  $Y$  can take.

$p(x_i, y_j)$  also denoted as  $p_{X,Y}(x_i, y_j)$ .

- The marginal probability mass functions for  $X$  and  $Y$  are

$$p_X(x) = P(X = x) = \sum_y p_{X,Y}(x, y)$$

$$p_Y(y) = P(Y = y) = \sum_x p_{X,Y}(x, y)$$

# Computation of Marginal PMF from Joint PMF: I

Formally:

$$\{X = x_i\} = \bigcup_j \{X = x_i, Y = y_j\}$$

All events on RHS are mutually exclusive. Thus,

$$\begin{aligned} p_X(x_i) &= P(X = x_i) = \sum_j P(X = x_i, Y = y_j) \\ &= \sum_j p(x_i, y_j) \end{aligned}$$



## Computation of Marginal PMF from Joint PMF: II

Similarly,

$$p_Y(y_j) = P(Y = y_j) = \sum_i p(x_i, y_j).$$

**Note:**  $P(X = x_i, Y = y_j)$  cannot be constructed from knowledge of  $P(X = x_i)$  and  $P(Y = y_j)$ .

## Example: 4.3a, Ross

3 batteries are randomly chosen from a group of 3 new, 4 used but still working, and 5 defective batteries. Let  $X$ ,  $Y$  denote the number of new, and used but working batteries that are chosen, respectively. Find

$$p(x_i, y_j) = P(X = x_i, Y = y_j).$$

Solution: Let  $T = {}^{12}C_3$

$$p(0, 0) = ({}^5C_3) / T$$

$$p(0, 1) = ({}^4C_1)({}^5C_2) / T$$

$$p(0, 2) = ({}^4C_2)({}^5C_1) / T$$

$$p(0, 3) = ({}^4C_3) / T$$

$$p(1, 0) = ({}^3C_1)({}^5C_2) / T$$

$$p(1, 1) = ({}^3C_1)({}^4C_1)({}^5C_1) / T$$

$$p(1, 2) = \dots, p(2, 0) = \dots, p(2, 1) = \dots, p(3, 0) = \dots$$

# Tabular Form

	0	1	2	3	Row Sum ( $P(X = i)$ )
0	10/220	40/220	30/220	4/220	84/220
1	30/220	60/220	18/220	0	108/220
2	15/220	12/220	0	0	27/220
3	1/220	0	0	0	1/220
Col sum ( $P(Y = j)$ )	56/220	112/220	48/220	4/220	

$i$  represents row and  $j$  represents column:

Both row and column sums add upto 1.

Marginal probabilities in the margins of the table.

## $n$ Random Variables: I

- Joint cumulative probability distribution function  $F(x_1, x_2, \dots, x_n)$  of  $n$  random variables  $X_1, X_2, \dots, X_n$  is defined as:

$$F(x_1, x_2, \dots, x_n) = P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n)$$

- If random vars. discrete: joint probability mass function

$$p(x_1, x_2, \dots, x_n) = P(X_1 = x_1, X_2 = x_2, \dots, X_n = x_n)$$

## $n$ Random Variables: II

- If random vars. continuous: joint probability density function  $f(x_1, x_2, \dots, x_n)$  such that for any set  $C$  in  $n$ -dimensional space

$$P((X_1, X_2, \dots, X_n) \in C) = \int \int \dots \int_{(x_1, \dots, x_n) \in C} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n$$

where,

$$f(x_1, x_2, \dots, x_n) = \frac{\partial^n F(x_1, x_2, \dots, x_n)}{\partial x_1 \partial x_2 \dots \partial x_n}$$

# Obtaining Marginals

$$F_{X_1}(x_1) = F(x_1, \infty, \infty, \dots, \infty)$$

$$f_{X_1}(x_1) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_n) dx_2 dx_3 \dots dx_n$$

$$p_{X_1}(x_1) = \sum_{x_2} \sum_{x_3} \dots \sum_{x_n} p(x_1, x_2, \dots, x_n)$$

# Independence of Random Variables: I

- Random variables  $X$  and  $Y$  are independent if for any two sets of real numbers  $A$  and  $B$ :

$$P(X \in A, Y \in B) = P(X \in A)P(Y \in B)$$

i.e. events  $E_A = \{X \in A\}$  and  $E_B = \{Y \in B\}$  are independent.

- Height and IQ
- In particular:  $P(X \leq a, Y \leq b) = P(X \leq a)P(Y \leq b)$ , or
- In terms of joint cumulative distribution function  $F$  of  $X$  and  $Y$ :

$$F(a, b) = F_X(a)F_Y(b); \quad \forall a, b \in \mathbb{R}$$

# Independence of Random Variables: II

Random variables that are not independent are called dependent



# Independence: Probability Mass and Density Functions

Random variables  $X, Y$  independent if:

- Discrete random variables: Probability mass function

$$p(x_i, y_j) = p_X(x_i)p_Y(y_j) \text{ for all } x_i, y_j$$

- Continuous random variables: Probability density function

$$f(x, y) = f_X(x)f_Y(y) \text{ for all } x, y$$

# Independence: Equivalent Statements

- 1)  $P(X \in A, Y \in B) = P(X \in A)P(Y \in B);$   
 $\forall A, B$  sets in  $\mathbb{R}$
- 2)  $F(x, y) = F_X(x)F_Y(y); \quad \forall x, y$
- 3)  $f(x, y) = f_X(x)f_Y(y); \quad \forall x, y; \text{ continuous RVs}$
- 3)  $p(x_i, y_j) = p_X(x_i)p_Y(y_j); \quad \forall x_i, y_j; \text{ discrete RVs}$

## Example 5.2 (Ogunnaike, 2009): I

The reliability of the temperature control system for a commercial, highly exothermic polymer reactor is known to depend on the lifetimes (in years) of the control hardware electronics,  $X_1$ , and of the control valve on the cooling water line,  $X_2$ . If one component fails, the entire control system fails. The random phenomenon in question is characterized by the two-dimensional random variable  $(X_1, X_2)$  whose joint probability distribution is given as:

$$f(x_1, x_2) = \begin{cases} \frac{1}{50} e^{-(0.2x_1 + 0.1x_2)}, & 0 < x_1 < \infty, 0 < x_2 < \infty \\ 0, & \text{otherwise} \end{cases}$$

## Example 5.2 (Ogunnaike, 2009): II

- ① Establish that above is a legitimate joint probability density function,  
To show:  $\int_0^\infty \int_0^\infty f(x_1, x_2) dx_1 dx_2 = 1$ .

$$\begin{aligned} & \int_0^\infty \int_0^\infty \frac{1}{50} e^{-(0.2x_1 + 0.1x_2)} dx_1 dx_2 \\ &= \frac{1}{50} (-5e^{-0.2x_1} \Big|_0^\infty) (-10e^{-0.1x_2} \Big|_0^\infty) = 1 \end{aligned}$$

## Example 5.2 (Ogunnaike, 2009): III

- 1 What's the probability of the system lasting more than 2 years.

To find:

$$P(X_1 > 2, X_2 > 2) = \int_2^\infty \int_2^\infty \frac{1}{50} e^{-(0.2x_1 + 0.1x_2)} dx_1 dx_2 = 0.549.$$

- 2 Find marginal density function of  $X_1$ .

$$f_{X_1}(x_1) = \int_0^\infty \frac{1}{50} e^{-(0.2x_1 + 0.1x_2)} dx_2 = \frac{1}{5} e^{-(0.2x_1)}$$

- 3 Find marginal density function of  $X_2$ ?

$$f_{X_2}(x_2) = \int_0^\infty \frac{1}{50} e^{-(0.2x_1 + 0.1x_2)} dx_1 = \frac{1}{10} e^{-(0.1x_2)}$$

## Example 5.2 (Ogunnaike, 2009): IV

4 Are  $X_1, X_2$  independent? Yes, since  $f(x_1, x_2) = f_{X_1}(x_1)f_{X_2}(x_2)$ .

# Independence of $n$ Random Variables: I

Random variables  $X_1, X_2, \dots, X_n$  are said to be independent if

- For all sets of real numbers  $A_1, A_2, \dots, A_n$ :

$$P(X_1 \in A_1, X_2 \in A_2, \dots, X_n \in A_n) = \prod_{i=1}^n P(X_i \in A_i)$$

- In particular:  $\forall a_1, a_2, \dots, a_n \in \mathbb{R}$

$$P(X_1 \leq a_1, X_2 \leq a_2, \dots, X_n \leq a_n) = \prod_{i=1}^n P(X_i \leq a_i),$$

$$\text{or, } F(a_1, a_2, \dots, a_n) = \prod_{i=1}^n F_{X_i}(a_i)$$

# Independence of $n$ Random Variables: II

- For discrete random variables: probability mass function factorizes:

$$p(x_1, x_2, \dots, x_n) = p_{X_1}(x_1)p_{X_2}(x_2)\dots p_{X_n}(x_n)$$

- For continuous random variables: probability density function factorizes:

$$f(x_1, x_2, \dots, x_n) = f_{X_1}(x_1)f_{X_2}(x_2)\dots f_{X_n}(x_n)$$



# Independent, Repeated Trials

- In statistics, one usually does not consider just a single experiment, but that the same experiment is performed several times.
- Associate a separate random variable with each of those experimental outcomes.
- If the experiments are independent of each other, then we get a set of independent random variables.
- Example: Tossing a coin  $n$  times. Random variable  $X_i$  is the outcome (0 or 1) in the  $i^{th}$  toss.

# Independent and Identically Distributed (IID) Variables: I

A collection of random variables is said to be IID if

- The variables are independent
- The variables have the same probability distribution
- Example 1: Tossing a coin  $n$  times. The probability of obtaining a head in a single toss does not vary and all the tosses are independent.
  - ▶ Each toss leads to a random variable with the same probability distribution function. The random variables are also independent. Thus, IID.

# Independent and Identically Distributed (IID) Variables: II

- Example 2: Measuring temperature of a beaker at  $n$  time instances in the day. The true water temperature changes throughout the day. The sensor is noisy.
  - ▶ Each sensor reading leads to a random variable.
  - ▶ Variables are independent but not identically distributed.

# Conditional Distributions

**Remember** for two events  $A$  and  $B$ : conditional probability of  $A$  given  $B$  is:

$$P(A \mid B) = \frac{P(A, B)}{P(B)}$$

for  $P(B) > 0$ .

# Conditional Probability Mass Function

- For  $X, Y$  discrete random variables, define the conditional probability mass function of  $X$  given  $Y = y$  by

$$\begin{aligned} p_{X|Y}(x|y) &= P(X = x \mid Y = y) \\ &= \frac{P(X = x, Y = y)}{P(Y = y)} = \frac{p(x, y)}{p_Y(y)} \end{aligned}$$

for  $p_Y(y) > 0$ .

## Examples 4.3b,f from Ross: I

**Question:** In a community, 15% families have no children, 20% have 1, 35% have 2 and 30% have 3 children. Each child is equally likely to be a boy or girl. We choose a family at random. Given that the chosen family has one girl, compute the probability mass function of the number of boys in the family?

G: number of girls, B: number of boys, C: number of children

To find:  $P(B = i | G = 1)$ ,  $i = 0, 1, 2, 3$ .

$$P(B = i | G = 1) = \frac{P(B = i, G = 1)}{P(G = 1)}, \quad i = 0, 1, 2, 3$$

## Examples 4.3b,f from Ross: II

First find  $P(G = 1)$

$$\{G = 1\} = \{G = 1\} \cap (\{C = 0\} \cup \{C = 1\} \cup \{C = 2\} \cup \{C = 3\})$$

$$\begin{aligned} P(G = 1) &= P(G = 1, C = 0) + P(G = 1, C = 1) \\ &\quad + P(G = 1, C = 2) + P(G = 1, C = 3) \end{aligned}$$

since  $C = 0, C = 1, C = 2, C = 3$  are mutually exclusive events with union as  $\Omega$ .

Then,

$$\begin{aligned} P(G = 1) &= P(G = 1 \mid C = 0)P(C = 0) \\ &\quad + P(G = 1 \mid C = 1)P(C = 1) + \dots \\ &= 0 + (1/2) \times 0.2 + \dots = 0.3875 \end{aligned}$$

## Examples 4.3b,f from Ross: III

Then,

$$P(B = 0 \mid G = 1) = \frac{P(B = 0, G = 1)}{P(G = 1)}$$

Numerator

$$= P(G = 1 \text{ and } C = 1) = P(G = 1 \mid C = 1)P(C = 1) = (1/2)0.2 = 0.1.$$

Then,

$$P(B = 0 \mid G = 1) = 0.1/0.3875 = 8/31$$

Similarly:  $P(B = 1 \mid G = 1) = 14/31$ ,  $P(B = 2 \mid G = 1) = 9/31$ ,  $P(B = 3 \mid G = 1) = 0$ . Check: Sum of conditional probabilities is 1.



# Conditional Probability Density Function

For Random Variables  $X, Y$ , conditional probability density of  $X$  given that  $Y = y$  is defined as:

$$f_{X|Y}(x | y) = \frac{f(x, y)}{f_Y(y)}$$

for  $f_Y(y) > 0$ .

Hence, can make statements on probabilities of  $X$  taking values in some set  $A$  given the value obtained by  $Y$  as:

$$P(X \in A | Y = y) = \int_A f_{X|Y}(x | y) dx$$

# Independence and Conditional Probabilities

If  $X, Y$  are independent, then

$$p_{X|Y}(x|y) = p_X(x)$$

$$f_{X|Y}(x|y) = f_X(x)$$

## Temperature Control Example (Continued), Example 5.2 (Ogunnaike, 2009) Earlier

- ① Find Conditional density function:  $f_{X_1|X_2}(x_1|x_2)$ .

$$f(x_1, x_2)/f_{X_2}(x_2) = \frac{1}{5}e^{-0.2x_1}$$

which is same as  $f_{X_1}(x_1)$  in this example.

- ② Similarly,  $f_{X_2|X_1}(x_2|x_1) = f_{X_2}(x_2)$  in this example.

Generic Question: If  $f_{X_1|X_2}(x_1|x_2) = f_{X_1}(x_1)$ , then is  $f_{X_2|X_1}(x_2|x_1) = f_{X_2}(x_2)$ ?

Answer: Yes

## Example 5.5 (Ogunnaike, 2009): I

$$f_{X_1, X_2} = \begin{cases} x_1 - x_2, & 1 < x_1 < 2, 0 < x_2 < 1 \\ 0, & \text{otherwise} \end{cases}$$

Find: Conditional probability densities.

Answer: Compute marginals

$$f_{X_1}(x_1) = \begin{cases} (x_1 - 0.5), & 1 < x_1 < 2 \\ 0, & \text{otherwise} \end{cases}$$
$$f_{X_2}(x_2) = \begin{cases} (1.5 - x_2), & 0 < x_2 < 1 \\ 0, & \text{otherwise} \end{cases}$$

## Example 5.5 (Ogunnaike, 2009): II

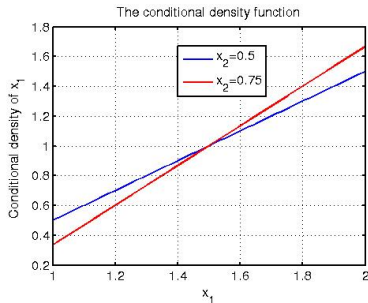
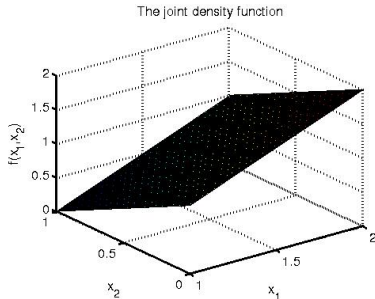
Then compute conditionals

$$f_{X_1|X_2}(x_1|x_2) = \frac{(x_1 - x_2)}{(1.5 - x_2)}, \quad 1 < x_1 < 2$$

$$f_{X_2|X_1}(x_2|x_1) = \frac{(x_1 - x_2)}{(x_1 - 0.5)}, \quad 0 < x_2 < 1$$

The random variables  $X_1, X_2$  are not independent.

# Plots



# Independence of Transformations

If random variables  $X, Y$  are independent, then the random variables

$$Z = g(X), U = h(Y)$$

are also independent.

Proof: Let  $A_z$  denote the set of points on the x-axis such that  $g(x) \leq z$  and  $B_u$  denote the set of points on the y-axis such that  $h(y) \leq u$ . Then,

$$\{Z \leq z\} = \{X \in A_z\}; \quad \{U \leq u\} = \{Y \in B_u\}$$

Thus, the events  $\{Z \leq z\}$  and  $\{U \leq u\}$  are independent because events  $\{X \in A_z\}$  and  $\{Y \in B_u\}$  are independent.

THANK YOU