Consider a classification problem with N Tyler Olivier HWG #1 liffort classes. of class n eN be Th let prior probability observations in class n are drawn from tenale fo(x)= P-(X=X/4=n) N(pn, En) where £ = E Use bayes theorem to And Pr (Y=n |X=X) a) P. PK- (XJ-Y) (=) PG (XJ4) Pr(X | y=n) p(y=n) = Pr(y | X=X) Pr(X=X) Pr(y|x=x) = P(x|y=n)p(y=n) Pr(x=x)> P(y=n |x=x) = P(x=x/y=n)p(y=n) PCX=X) = fp(x) (Th) p(x=x) $= f_n(x) \pi_n$ P (x=x, y=i) FA(x) Tr. = { p(x=x/y=i)p(y=i) $P(y=n|x=x) = \frac{f_n(x) \pi_n}{5^E P(x=x|y=i) \Pi_i}$

$$P\left(y=n\mid x=x\right) = \frac{f_n(x) \prod_{i=1}^n f_i(x_i)}{\sum_{i=1}^n f_i(x_i) \prod_{i=1}^n f_i(x_i)}$$

b) Derive the linear descriminant function, In(x) and write the classification rule for the predicted class, if for an LDA in

+(LK)) = TEMP(Z) exp (-1/2 (X-Pn) {-1 (X-Pn))

Pr(y=n|X=x) = (xp(-1/2(x-Pn) \(\frac{1}{(x-Pn)}\)) ITn

109 Pr(y=n|x=x) = log Tin + log [TCITIPIZI exp (-1/2 (x-1/2)]

- log p(x=x)

+ log (exp (-1/2 x = x + 1 x = p) = log Tn + log (TCZTIPIEI) - 42 PM TZ-1 Wn)

-109 p(x=x)

5 n(x) = loggr(y=n)x=x) + log (1/21) (-105 P(x=x) will drop out

the poclassification rule should the the LUA in the case of yell on show taking the most with y to removes constants in foch). y = argmax (lostin to xt & pn - 12 Pn = 2 Pn) where we estimate To stor with the spin the C) Derive the decision boundary for the LDA in the case of two classes, on and b. The decision boundary for two classes a,b is the set of points where fa(x) = Sb(x) 169 Ta + xt & Pa - 1/2 10 at & Pa = logITb + xt & Pb - 1/2 1/6 TE Pb

This is an equation of a line to the (Substitute estimates Go TTa, Tib, Pa, Pb)

d) Consider two classes, a=1 and b=2. You are Given $\hat{\pi}_a = .6$, $\hat{\pi}_b = .4$ $\hat{\gamma}_i = [-2]$ and £ = [1.6 .1 Find the decision boundary and classification Corresponding LOA. How would the observation x=[i] be classified?

The decision boundary are the points x 5.t (0g(.4) + xT[1.5.1][0] - 1/2 (1 0) [1.5.1][0] $= \log(.4) + x^{T} \begin{bmatrix} 1.5 & 1 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} -2 \\ 1 \end{bmatrix} - \frac{1}{2} \begin{bmatrix} -2 \\ 1 \end{bmatrix} \begin{bmatrix} 1.5 & 1 \\ -1 \end{bmatrix} \begin{bmatrix} -2 \\ 1 \end{bmatrix}$ $-.51 + x^{T} \left(.71\right) -.316 = -.97 + x^{T} \left(-7.14\right) - 2.75$ -.87 + x + (.71) = -3.67 + x + (1.71)clasify x as class a if $-187 + x^{T} \begin{bmatrix} .71 \\ -.07 \end{bmatrix} > -3.67 + x^{T} \begin{bmatrix} -2.41 \\ 1.21 \end{bmatrix}$ and b otherwise -.87 + [1][.71] = -.23 -3.27 + [1][-7.14] = -41.6 since = 237 -4.6 = classify X as a.

This relates to logistic regression under the binary case in the following:

with two classes the decision boundary for gaussian bayes is when $\log p(y=1|X) = \log p(y=0|X)$ so thoose doss I when $\log p(y=1|X) - \log p(y=0|X) > 0$ $\Rightarrow \log \frac{p(y=1|X)}{p(y=0|X)} > 0$ the quantity $\log \frac{p(y=1|X)}{p(y=0|X)}$ is the $\log p(y=0|X)$

Which is the basis of logistic regression

One naive Bayes classifier is similar to LDA, except it assumes each predictor is conditionally independent of every other predictor given class n. Derive the classification rule for 9 under this classifier. How does this relate to logistic regression in the binary case (i.e for two classes?)

p(X1y) = ft p(X:1y) under naive bayes conditionally independent assumption.

 $P(y=p(x=x)) = \frac{f_n(x) T_n}{P(x=x)} = \frac{\prod_{i=1}^n p(x_i|y=n) \prod_{i=1}^n p(x_i|y=n) \prod_{i=1}^n$

 $\log p(y=n|X=x) = \log \left(\prod_{i=1}^{p} p(x_i|y=n) \right) + \log \prod_{i=1}^{p} -\log p(x=x)$ $= \sum_{i=1}^{p} (\log p(x_i|y=n)) + \log \prod_{i=1}^{p} + \log p(x=x)$

= $\frac{2}{2} \log \frac{1}{100} \exp \left(-\frac{1}{2\sigma_n^2} (x_0 - \mu_n)^2\right) + \log \pi_n + \log \rho(x=x)$

 $\hat{y} = \underset{y \in \{1, 2, \dots, n\}}{\operatorname{log}} p(y = n| x = x)$ is classification rule

Substitute above expression except for log plank) as it does not effect maximization wity. Substitue estimates for necessary parameters.

Tyler Olivieri HWG # 2 K-means clustering a) Show that setting the objective to the the sum of the square Euclidean distances of points from the center of their cluster dj = \(\frac{1}{2} \) \(\(\chi_{i} - \chi_{i} \)^{2} results in an update rule where the optimal centroid is the mean of the points in the cluster. min obj wit Cic dobi = 2 2 2 (Cki - Xi) = 0 2 2 2 (CK: - X.) =0 2 & C(CK1 - Xi) = 0 E E CK: - E E X; =0 XECEIZI = ZZZXI XECEIZI # points assigned to duster K. XEGE.

YECK IT E TOTAL TOTAL TO CONSTRUCT K.

So set Cri to be the mean of all data

points assigned to durter k.

CK: = \(\times \times \) ignored to the color of the col

for each feeture, set cluster center to be mean of data points assigned to the cluster center.

Show that setting the objective function to the Sum of the manhattan distances of points from the center of their dustos, results in an appearte rule where the applicat centraid is the median of the cluster. $\frac{dobj}{dCk} = \underbrace{\sum_{X \in Ck} \sum_{i=1}^{k} \frac{Cki - Xi}{|Cki - Xi|}}_{= 0}$ ther (CKi + Xi = | CKi - Xi =) [CKi + Xi] = | CKi - Xi when E Cx; >0 when $C_{|C|} - X_i \subset D$ then $C_{|C|} - X_i = C_{|C|} - C_{|C|}$ $\frac{1}{|C_{E_i}-X_i|} = \frac{sign(C_{E_i}-X_i)}{|C_{E_i}-X_i|} \quad \text{where} \quad \frac{sign(x)}{|X_i|} = \frac{1}{|X_i|} \text{ when } \frac{1}{|X_i|}$ S (Z sign(Ck: -Xi)) = 0 This can only equal zero when the number of the positive elements equals the number of regative elements. So for Cx; So Cki = median (X & Chai) than

Tyler Olivice, HWG P3 Consider the dataset (1) Normalize the data and derive the two principal components in Sorted order. 0+17+2+2+3+3+4 = 2.14 Ry = 1+1+1+3+2+3+5 = 22.29 $\frac{1}{2} = \frac{1}{1} \frac{2}{5} (x_1 - y_2)^2 = \frac{1}{6} \frac{2}{1} (x_1 - 2.14)^2 = 1.55$ Ox = Jo2x = 1.55 = 1.28 ση= 1 } (9,-μη) = 1 } (9,-2.27) = 197 og= Jog = 1.92 = 7.38 Normalized data X: = X: - Px ory stolev=1 X-norm 60 ~ (0,1) Xnow Buow 1-norm 11 -1.72 -.92 .69 -.21

let
$$A = \begin{bmatrix} -1.77 & -.93 \\ -.92 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.11 & -.93 \\ -.12 & -.52 \\ -.12 & -.54 \\ -.2 & -.54 \\ -.36 & -.06 \\ -.2 & -.54 \\ -.36 & -.06 \\ -.2 & -.54 \\ -.36 & -.06 \\ -.2 & -.54 \\ -.36 & -.32 \end{bmatrix}$$

Now we have $T = U \mathcal{E} = \begin{bmatrix} -.57 & -.54 \\ -.56 & -.66 \\ -.2 & -.54 \\ -.36 & -.32 \end{bmatrix}$

Now we have $T = U \mathcal{E} = \begin{bmatrix} -.57 & -.54 \\ -.56 & -.66 \\ -.2 & -.54 \\ -.36 & -.32 \end{bmatrix}$

$$T = U\Sigma = \begin{bmatrix} -1.88 & -1.56 \\ -1.31 & .06 \\ -.74 & .58 \\ -.74 & .58 \\ -.44 & .34 \\ -.86 & .12 \\ -2.44 & -.33 \end{bmatrix}$$

The new transformed lataset using the first principal component is

Which is the first column of T.

b) repeat the previous analysis but do not normalize the data. when you multiply UE, they are irrelevant. Is pea scale-invariant (removing A-r scolumns of U, and of A-unnormalized: 0's 1700 and columns of (E) because python libs give SVD this way except numpy, which I did not use A - unnormalized = $\begin{bmatrix} 9.57 & 0 \\ 0 & 1.54 \end{bmatrix} \begin{bmatrix} .68 & .73 \\ -.73 & .68 \end{bmatrix}$ T-unnormalized = UE

=> PCA is not scale-invariant X + X-unnoquelized

$$= 02 = \begin{bmatrix} -1.88 & -1.512 \\ -1.31 & .06 \\ -.74 & .58 \\ -.74 & .58 \\ -.44 & .34 \\ .64 & .34 \\ .86 & .12 \\ -2.44 & -.33 \end{bmatrix}$$

The new transformed lataset using the first principal component is

Which is the first column of T.