

Suppose X_1, \dots, X_n are an iid random sample of size n w/ sample mean \bar{X} and sample variance $S^2 = 5$
 $\bar{X} = 12$

a) Let $n=5$ and suppose samples are drawn from a normal distribution w/ unknown mean μ and known variance $\sigma^2 = 9$

Let the null hypothesis be $H_0: \mu = 10$; $H_a: \mu \neq 10$

Calculate the relevant test statistic value and p-value.

Want to use z-test. The relevant test statistic

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} = \frac{12 - 10}{3/\sqrt{5}} = \frac{2\sqrt{5}}{3}$$

P-value will be smallest significance level at which the null hypothesis would be rejected.

$$p = \Pr \{ \text{event more contradictory than observed} \mid H_0 \text{ is true} \}$$

$$= \Pr \{ Z > z \cup Z < -z \mid H_0 \text{ is true} \}$$

Due to symmetry under null hypothesis

$$= 2 \Pr \{ Z > |z| \mid H_0 \text{ is true} \}$$

$$= 2 (1 - \Pr \{ Z \leq |z| \mid H_0 \text{ is true} \})$$

$$= 2 (1 - \Phi(|z|))$$

$$= 2 (1 - \Phi(\frac{2\sqrt{5}}{3})) = 2 (1 - \Phi(1.5)) = .13 = p\text{-value}$$

b) Using the acceptance region of this test, construct a 95% confidence interval.

let $\alpha = .05$ for 95% confidence

$$1 - \alpha = \Pr \left(-z \leq \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \leq z \right)$$

from Normal table

$$1 - .05 = \Pr \left(-1.96 \leq \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \leq 1.96 \right)$$

$$.95 = \Pr \left(-1.96(\sigma/\sqrt{n}) - \bar{X} \leq -\mu_0 \leq 1.96(\sigma/\sqrt{n}) - \bar{X} \right)$$

$$.95 = \Pr \left(-1.96(\sigma/\sqrt{n}) + \bar{X} \geq \mu_0 \geq -1.96(\sigma/\sqrt{n}) + \bar{X} \right)$$

$$.95 = \Pr \left(\bar{X} - 1.96(\sigma/\sqrt{n}) \leq \mu_0 \leq \bar{X} + 1.96(\sigma/\sqrt{n}) \right)$$

substitute \bar{X} , σ , and n from known values.

$$.95 = \Pr \left(12 - 1.96 \left(\frac{3}{\sqrt{5}} \right) \leq \mu_0 \leq 12 + 1.96 \left(\frac{3}{\sqrt{5}} \right) \right)$$

\Rightarrow 95% confidence interval for μ_0

$$\left[12 - 1.96 \left(\frac{3}{\sqrt{5}} \right), 12 + 1.96 \left(\frac{3}{\sqrt{5}} \right) \right]$$

$$[9.37, 14.63]$$

(a) Determine decision rule for $\alpha = .05$

$$.05 = \Pr \{ \text{type I error} \}$$

$$= \Pr \{ H_0 \text{ is rejected} \mid H_0 \text{ is true} \}$$

$$= \Pr \{ Z < -z_{\alpha} \cup Z > z_{\alpha} \mid H_0 \text{ is true} \}$$

$$= 2 \Pr \{ Z < -z_{\alpha} \mid H_0 \text{ is true} \}$$

$$.05 = 2 \Phi(-z_{\alpha})$$

$$\frac{.05}{2} = \Phi(-z_{\alpha})$$

$$\Phi^{-1}(.05/2) = -z_{\alpha}$$

$$-\Phi^{-1}(\frac{.05}{2}) = z_{\alpha} = 1.96$$

Accept H_0 if $-z_{\alpha} < Z < z_{\alpha}$ $-1.96 < Z < 1.96$

reject otherwise.

1c) Now let $n=5$ again but suppose the samples are drawn from an unspecified distribution w/ unknown mean μ and unknown variance σ^2 . Let the null hypothesis be $H_0: \mu=10$ and the alternate hypothesis be $H_a: \mu>10$. Calculate the relevant test statistic value and p-value. Determine the decision rule for $\alpha=.05$.

The relevant test statistic will be t , because we do not know the distribution and n is small.

$$T = \frac{\bar{x} - \mu_0}{s/\sqrt{n}} = \frac{12 - 10}{\sqrt{5}/\sqrt{5}} = \frac{2\sqrt{5}}{\sqrt{5}} = 2$$

$$p\text{-value} = \text{Prob}\{T > T_0 \mid H_0 \text{ is true}\} \dots \text{one-sided}$$

$$= 2(1 - \text{Prob}\{T \leq T_0 \mid H_0 \text{ is true}\})$$

$$= 2(1 - t_{n-1}(T)) = 1 - t_4(T) = 1 - t_4(2)$$

CDF of t -dist with 4 deg of freedom

$$= .06$$

$$= .12$$

$$\alpha = P\{\text{Type I error} \mid H_0 \text{ is true}\}$$

should be $t_{\alpha, n-1}$

1.2.1.

$$.05 = P\{T \geq T_{\alpha} \mid H_0 \text{ is true}\}$$

but I am getting $t_{\alpha, n-1}$

$$.05 = P\{T \geq T_{\alpha} \mid H_0 \text{ is true}\}$$

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$$\text{critical value} = T_{\alpha} = t_{\alpha}^{-1}(1 - .05) = t_{\alpha}^{-1}(.95) = 2.13$$

Suppose that Y_1, \dots, Y_n are an iid random sample of size n drawn from a Poisson distribution w/ unknown parameter λ .

Using $\sum_{i=1}^n Y_i$ as the test statistic, find the critical value and rejection region at level α for the test.

let $T = \sum_{i=1}^n Y_i$

$$\begin{cases} H_0: \lambda = \lambda_0 \\ H_a: \lambda = \lambda^* \end{cases} \text{ where } \lambda^* > \lambda_0$$

let c be the critical value

$$\alpha = \Pr \{ \text{Type I error occurs} \}$$

$$\alpha = \Pr \{ \text{reject } H_0 \mid H_0 \text{ is true} \}$$

reject H_0 when $T > c$

$$\alpha = \Pr \{ T > c \mid H_0 \text{ is true} \}$$

$$\alpha = 1 - \Pr \{ T \leq c \mid H_0 \text{ is true} \}$$

Since $T \sim \text{poisson}(\lambda)$, given H_0 is true $\lambda = \lambda_0$

$\Pr \{ T \leq c \mid H_0 \text{ is true} \}$ is cdf of poisson with λ_0 where $IE[Y_i] = \lambda$ and $Var(Y_i) = \lambda$

However, as $n \rightarrow \infty$, $\frac{T - n\lambda_0}{\sqrt{n\lambda_0}} \xrightarrow{d} N(0,1)$

Piazza note allows assumption of n large
then we can say $\frac{T - n\lambda_0}{\sqrt{n\lambda_0}}$ is approximately distributed as gaussian conditioned under the null hypothesis being true.

anything larger

critical value

$$\text{let } z = \frac{T - n\lambda_0}{\sqrt{\frac{\lambda_0}{n}}}$$

reject when $z > \tau$ where τ is a threshold

$$\frac{T - n\lambda_0}{\sqrt{\frac{\lambda_0}{n}}} > \tau$$

$$T > \tau \sqrt{\frac{\lambda_0}{n}} + n\lambda_0 \quad (*)$$

$$\alpha = \Pr \{ z > \tau \mid H_0 \text{ is true} \}$$

$$1 - \alpha = \Pr \{ z \leq \tau \mid H_0 \text{ is true} \} \quad z \sim N(0,1) \text{ given } H_0 \text{ is true}$$

$$\alpha = 1 - \Phi(\tau)$$

$$1 - \alpha = \Phi(\tau)$$

$$\tau = \Phi^{-1}(1 - \alpha) = \Phi^{-1}(\alpha)$$

we can translate this threshold to use the statistic T instead of z using $(*)$

The rejection region is any $z > \tau$ so (τ, ∞)

The rejection region for T is $T > \tau \sqrt{\frac{\lambda_0}{n}} + n\lambda_0$

$$(\tau \sqrt{\frac{\lambda_0}{n}} + n\lambda_0, \infty)$$

Consider a random sample of size $n=100$ w/ sample proportion $\hat{p} = .2$ from a population w/ a true unknown proportion P .

- a) For the test $H_0: P = .25$ versus $H_a: P < .25$, calculate the relevant test statistic value and p-value. Determine the decision rule for $\alpha = .05$ and $\alpha = .01$

Test statistic: We can use the z-test here because we have large $\frac{n \cdot P}{s^2}$. From CLT \hat{P} calculation will be distributed approx normal.

$$\text{test } z = \frac{\hat{p} - P}{\sqrt{\frac{P(1-P)}{n}}}$$

$$\text{Variance of prop} = \frac{P(1-P)}{n}$$

$$= \frac{.2 - .25}{\sqrt{\frac{.25(1-.25)}{100}}} = \frac{-.05}{.043} = -1.463$$

$$\alpha = \Pr \{ \text{Type I error} \} = \Pr \{ H_0 \text{ is rejected} \mid H_0 \text{ is true} \}$$

$$\alpha = \Pr \{ Z < z \mid H_0 \text{ is true} \} \quad z \text{ is threshold.}$$

This is just one-sided gaussian.

$Z \sim N(0, 1)$ when H_0 is true.

$$\alpha = \Phi(z)$$

$$z = \Phi^{-1}(\alpha)$$

when $\alpha = .05$

$$z = \Phi^{-1}(.05) = -1.65$$

when $\alpha = .01$

$$z = \Phi^{-1}(.01) = -2.33$$

$$p\text{-value} = \Pr\{z < z \mid H_0 \text{ is true}\}$$

$$= \Phi(-1.163)$$

$$= .12$$

cho accept H_0 when $z > z$
reject H_0 when $z < z$

b)

For the test $H_0: p = .25$ versus $H_a: p \neq .25$
calculate the relevant test statistic value and p-value.

Determine the decision rule for $\alpha = .05$ and $\alpha = .01$

The test statistic does not change

$$z = \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}} = -1.163$$

$$p\text{-value} = 2 \Pr\{z < z \mid H_0 \text{ is true}\}$$

$$= 2\Phi(-1.163) = .1213$$

$$\alpha = \Pr\{\text{Type I error}\} = \Pr\{H_0 \text{ is rejected} \mid H_0 \text{ is true}\}$$

$$\alpha = 2(1 - \Phi(z))$$

$$\Rightarrow z = \pm \Phi^{-1}(1 - \alpha/2)$$

for $\alpha = .05$

$$z = \Phi^{-1}(1 - .05/2), -\Phi^{-1}(1 - .05/2)$$

$$\alpha/2 = 1 - \Phi(z)$$

$$1 - \alpha/2 = \Phi(z)$$

$$\text{for } z = 1.96, -1.96$$

accept if $-z < z < z$

reject
D.W

for $\alpha = .015$

$$z = \Phi^{-1}(1 - .01/2), -\Phi^{-1}(1 - .01/2) = 2.518, -2.518$$

accept

H_0

when $-2 < z < 2$

$-2.58 < z < 2.58$

reject

H_0

otherwise.