

1. Cauchy-Schwarz inequality:

$$|u \cdot v| \leq \|u\| \|v\| \quad \text{OR} \quad \left(\sum_{i=1}^n u_i v_i \right)^2 \leq \left(\sum_{i=1}^n u_i^2 \right) \left(\sum_{i=1}^n v_i^2 \right)$$

where the equality holds iff u, v are linearly dependent.

proof:

$$\begin{aligned} \sum_{i=1}^n \sum_{j=1}^n (u_i v_j - u_j v_i)^2 &= \sum_{i=1}^n u_i^2 \sum_{j=1}^n v_j^2 + \sum_{i=1}^n v_i^2 \sum_{j=1}^n u_j^2 - 2 \sum_{i=1}^n u_i v_i \sum_{j=1}^n u_j v_j \\ &= 2 \left(\sum_{i=1}^n u_i^2 \right) \left(\sum_{j=1}^n v_j^2 \right) - 2 \left(\sum_{i=1}^n u_i v_i \right)^2 \end{aligned}$$

Since LHS ≥ 0 , we have $\left(\sum_{i=1}^n u_i^2 \right) \left(\sum_{j=1}^n v_j^2 \right) \geq \left(\sum_{i=1}^n u_i v_i \right)^2$.

The equality holds iff $u_i v_j = u_j v_i$, i.e. $\frac{u_i}{u_j} = \frac{v_i}{v_j} \quad \forall i, j = 1, \dots, n$ (linearly dependent).
Q.E.D.

It can be generalized by replacing all finite sums to integrals:

$$\left| \int_{-\infty}^{+\infty} \phi_1(x) \phi_2^*(x) dx \right|^2 \leq \int_{-\infty}^{+\infty} |\phi_1(x)|^2 dx \int_{-\infty}^{+\infty} |\phi_2(x)|^2 dx$$

The proof is similar and thus omitted here.

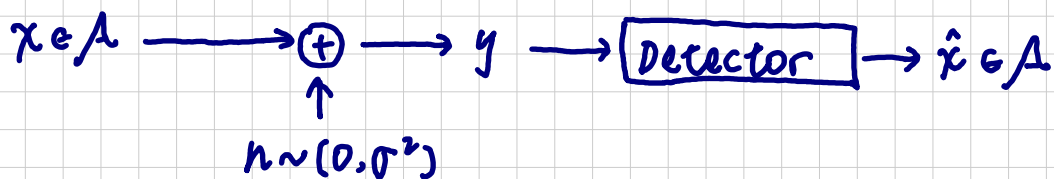
Hölder's inequality generalizes Cauchy-Schwarz inequality:

$$|u \cdot v| \leq \|u\|_p \|v\|_q \quad \text{OR} \quad \sum_{i=1}^n |u_i v_i| \leq \left(\sum_{i=1}^n |u_i|^p \right)^{\frac{1}{p}} \left(\sum_{j=1}^n |v_j|^q \right)^{\frac{1}{q}}$$

where $p, q \in [1, \infty]$ and $\frac{1}{p} + \frac{1}{q} = 1$.

Note that ∞ norm reduces to $\|x\|_\infty = \max_i |x_i|$.

2. Detection in AWGN channel with finite input set \mathcal{A} (alphabet)



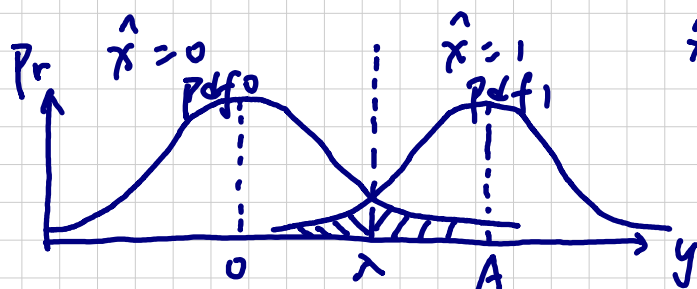
- x is the input letter to transmit
- n is AWGN with zero mean and variance σ^2
- y is the received signal following Gaussian distribution $y \sim (x, \sigma^2)$
- \hat{x} is the detected letter

Success detection: $\hat{x} = x$

For simplicity, let $\mathcal{A} = \{0, A\}$: binary-input $\{0, 1\}$

- If $x=0$, then $y \sim (0, \sigma^2)$
- If $x=A$, then $y \sim (A, \sigma^2)$
- Decision threshold λ :

$$\begin{cases} \hat{x} = 0 & \text{if } y < \lambda \\ \hat{x} = 1 & \text{if } y > \lambda \\ \hat{x} = ? & \text{if } y = \lambda \text{ (uncertain)} \end{cases}$$



Error: $\hat{x} \neq x$

$$\begin{cases} x=0, \hat{x}=1 & \textcircled{1} \\ x=1, \hat{x}=0 & \textcircled{2} \end{cases}$$

①: $x=0, \hat{x}=1$ input probability Gaussian w. pdf 0

$$P_{e0} = P(x=0, \hat{x}=1) = P(x=0) P(\hat{x}=1 | x=0)$$

$$P(\hat{x}=1 | x=0) = \int_{\lambda}^{\infty} \frac{1}{\sigma \sqrt{\pi}} e^{-\frac{n^2}{2\sigma^2}} dn \text{ is a function of } \lambda \cdot \sigma.$$

②: $x=1, \hat{x}=0$

$$P_{e1} = P(x=1, \hat{x}=0) = P(x=1) P(\hat{x}=0 | x=1)$$

$$P(\hat{x}=0 | x=1) = \int_{-\infty}^{\lambda} \frac{1}{\sigma \sqrt{\pi}} e^{-\frac{(n-A)^2}{2\sigma^2}} dn \text{ also depends on } A!$$

$$P_e = P_{e0} + P_{e1} = \underbrace{P_1}_{P(\hat{x}=1)} \int_{-\infty}^{\lambda} \frac{1}{\sigma \sqrt{\pi}} e^{-\frac{(n-A)^2}{2\sigma^2}} dn + (1-P_1) \underbrace{\int_{\lambda}^{\infty} \frac{1}{\sigma \sqrt{\pi}} e^{-\frac{n^2}{2\sigma^2}} dn}_{P(\hat{x}=0)}$$

$$\frac{dP_e}{d\lambda} = 0 \Rightarrow \lambda^* = -\frac{\sigma^2}{A} \ln \frac{P_1}{1-P_1} + \frac{A}{2}, \text{ depends on the input probability.}$$

Consider equiprobable input $P(x=0) = P(x=1) = 0.5$.

$\lambda^* = \frac{A}{2}$ ensures $P_e = P_{e0} = P_{e1}$.

$$z \triangleq \frac{n}{\sigma} \Rightarrow P_{e0} = \frac{1}{\sqrt{\pi}} \int_{\frac{A}{2\sigma}}^{\infty} e^{-\frac{z^2}{2}} dz = Q\left(\frac{A}{2\sigma}\right).$$

$$Q(\pi) = \frac{1}{\sqrt{\pi}} \int_{\pi}^{\infty} e^{-\frac{t^2}{2}} dt$$