Module 1: Introduction to Bayesian Statistics

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Agenda

- Motivations
- Traditional inference
- Bayesian inference
- Bernoulli, Beta
- Posterior of Bernoulli-Beta
- Conjugacy
- Example with 2012 election data
- Marginal likelihood
- Posterior Prediction
- Additional problems at the end of lecture (derivation + applied)

What should you learn?

- You should learn the main principles of Bayesian inference/prediction and how to apply these to real data analysis.
- You will continue with this in lab/homework to make sure that you understand these key principles.

Traditional inference

You are given data X and there is an **unknown parameter** you wish to estimate θ

How would you estimate θ ?

- \triangleright Find an unbiased estimator of θ .
- ▶ Find the maximum likelihood estimate (MLE) of θ by looking at the likelihood of the data.
- Please review unbiased estimation and finding an MLE.
- Please also review other background material such as likelihoods, sufficient statistics, basic probability concepts, etc. Most of this material can be reviewed in Chapters 1-3 in Hoff.

Bayesian inference

Bayesian methods trace its origin to the 18th century and English Reverend Thomas Bayes, who along with Pierre-Simon Laplace discovered what we now call **Bayes' Theorem**

- \triangleright $p(x \mid \theta)$ likelihood
- $ightharpoonup p(\theta)$ prior
- $ightharpoonup p(\theta \mid x)$ posterior
- \triangleright p(x) marginal distribution

How can we derive $p(\theta \mid x)$?

Derivation of $p(\theta \mid x)$

Bernoulli distribution

The Bernoulli distribution is very common due to binary outcomes.

- Consider flipping a coin (heads or tails).
- We can represent this a binary random variable where the probability of heads is θ and the probability of tails is 1θ .

Consider $X \sim \text{Bernoulli}(\theta)\mathbb{1}(0 < \theta < 1)$

The likelihood is

$$p(x \mid \theta) = \theta^{x} (1 - \theta)^{(1-x)} \mathbb{1}(0 < \theta < 1).$$

- Exercise: what is the mean and the variance of X?
- What is the connection with the Bernoulli and the Binomial distribution?

Bernoulli distribution

▶ Suppose that $X_1, \ldots, X_n \stackrel{iid}{\sim} \mathsf{Bernoulli}(\theta)$. Then for $x_1, \ldots, x_n \in \{0, 1\}$ what is the likelihood?

Notation

- ightharpoonup \propto : means "proportional to"
- \triangleright $x_{1:n}$ denotes x_1, \ldots, x_n

Bernoulli and Binomial Connection

$$X_1, \ldots, X_n \stackrel{iid}{\sim} \text{Bernoulli}(\theta).^1$$

Suppose
$$Y = \sum_{i} X_{i=1}^{n}$$
. Then $Y \sim Binomial(n, \theta).^{2}$

Remark: A binomial random variable with parameter n=1 is equivalent to a Bernoulli random variable, i.e. there is only one trial.

¹This represents n coin flips with success probability θ .

²This represents *n* Bernoulli trials with success probability θ .

Likelihood

$$p(x_{1:n}|\theta) = \mathbb{P}(X_1 = x_1, \dots, X_n = x_n \mid \theta)$$

$$= \prod_{i=1}^n \mathbb{P}(X_i = x_i \mid \theta)$$

$$= \prod_{i=1}^n p(x_i|\theta)$$

$$= \prod_{i=1}^n \theta^{x_i} (1-\theta)^{1-x_i}$$

$$= \theta^{\sum x_i} (1-\theta)^{n-\sum x_i}.$$

Beta distribution

Given a,b>0, we write $\theta \sim \mathrm{Beta}(a,b)$ to mean that θ has pdf

$$p(\theta) = \mathsf{Beta}(\theta|a,b) = \frac{1}{B(a,b)} \theta^{a-1} (1-\theta)^{b-1} \mathbb{1}(0 < \theta < 1),$$

i.e., $p(\theta) \propto \theta^{a-1} (1-\theta)^{b-1}$ on the interval from 0 to 1.

► Here,

$$B(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

- .
- Parameters a, b control the shape of the distribution.
- This distribution models random behavior of percentages/proportions.

Posterior of Bernoulli-Beta

Let's derive the posterior of $\theta \mid x_{1:n}$

Conjugacy

What do you notice about the prior and the posterior from the Bernoulli-Beta example that we just considered?

Conjugacy

A class P of prior distributions for θ is called **conjugate** for the likelihood $p(x \mid \theta)$ if

$$p(\theta) \in P \implies p(\theta \mid x) \in P.$$

Tip: In practice, we check to see if the posterior has an updated form of the prior.

Conjugacy

Benefits

- We do minimal or often no math. In fact, https://en.wikipedia.org/wiki/Conjugate_prior provides many conjugate families.
- We have an exact posterior distribution. No approximations are needed.
- Computation is fast and simple!

Downside

Sometimes an unrealistic assumption, however, might provide guidance to us.

Approval ratings of Obama

What is the proportion of people that approve of President Obama in PA?

- ▶ We take a random sample of 10 people in PA and find that 6 approve of President Obama. Likelihood
- ► The national approval rating (Zogby poll) of President Obama in mid-September 2015 was 50%. We'll assume that in PA his approval rating is also 50%. Prior
- ▶ Based on this prior information, we'll use a Beta prior for θ and we'll choose a and b.

Let's consider choosing a prior distribution that is skewed and see what happens.

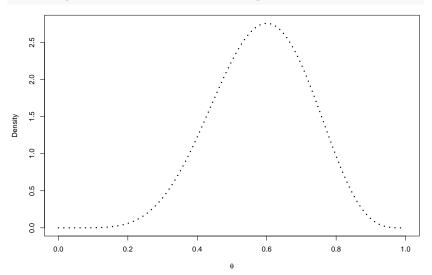
Obama Example

[1] 0.9849906

```
n < -10
# Fixing values of a,b. Chosen skewed Beta.
a = 21/8
b = 0.04
\#a <- 0.25
#b <- 0.25
th <- seq(0, 1, length = 500)
x < -6
# we set the likelihood, prior, and posteriors with
# THETA as the sequence that we plot on the x-axis.
# Beta(c,d) refers to shape parameter
like \leftarrow dbeta(th, x + 1, n - x + 1)
prior <- dbeta(th, a, b)</pre>
print(a / (a + b))
```

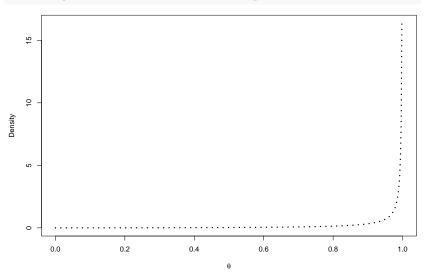
```
post <- dbeta(th, x + a, n - x + b)
```

Likelihood



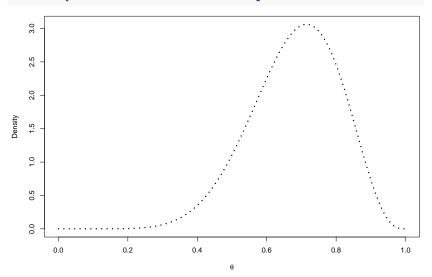
Prior

```
plot(th, prior, type = "l", ylab = "Density",
    lty = 3, lwd = 3, xlab = expression(theta))
```

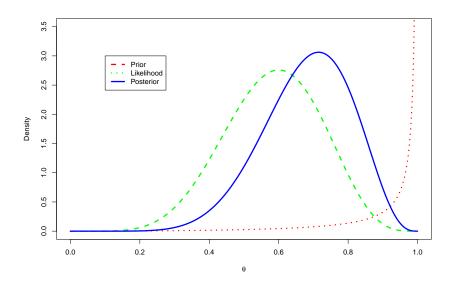


Posterior

```
plot(th, post, type = "l", ylab = "Density",
    lty = 3, lwd = 3, xlab = expression(theta))
```



Likelihood, Prior, and Posterior



Back to the Prior

- ▶ We choose the prior here two different ways. What do you observe?
- ▶ In the supplemental material (end of lecture), find an example where we have more information and can set *a*, *b* from in a more subjective and principled manner.

Cast of characters

- Observed data: x
- This often involves many data points, e.g.,

$$x = x_{1:n} = (x_1, \ldots, x_n).$$

```
likelihood p(x_{1:n}|\theta) prior p(\theta) posterior p(\theta|x_{1:n}) marginal likelihood p(x_{1:n}) posterior predictive p(x_{n+1}|x_{1:n})
```

Marginal likelihood

The marginal likelihood is defined as

$$p(x) = \int p(x|\theta)p(\theta) d\theta$$

Example: Back to the Bernoulli-Beta

$$X_1,\dots,X_n\mid heta\stackrel{\mathit{iid}}{\sim} Bernoulli(heta)$$
 and $heta\sim Beta(a,b).$

What is the marginal likelihood for the Bernoulli-Beta?

Marginal Likelihood: Bernoulli-Beta

Then the marginal likelihood is

$$p(x_{1:n})$$

$$= \int p(x_{1:n}|\theta)p(\theta) d\theta$$

$$= \int_{0}^{1} \theta^{\sum x_{i}} (1-\theta)^{n-\sum x_{i}} \frac{1}{B(a,b)} \theta^{a-1} (1-\theta)^{b-1} d\theta$$

$$= \frac{1}{B(a,b)} \int_{0}^{1} \theta^{\sum x_{i}+a-1} (1-\theta)^{n-\sum x_{i}+b-1} d\theta$$

$$= \frac{\frac{B(a+\sum x_{i}, b+n-\sum x_{i})}{B(a,b)} \int_{0}^{1} \frac{\theta^{\sum x_{i}+a-1} (1-\theta)^{n-\sum x_{i}+b-1}}{B(a+\sum x_{i}, b+n-\sum x_{i})} d\theta$$

$$= \frac{B(a+\sum x_{i}, b+n-\sum x_{i})}{B(a,b)},$$

by the integral definition of the Beta function.

Posterior predictive distribution

- At times, we may wish to find the conditional distribution of x_{n+1} given $x_{1:(n+1)}$.
- **Assumption 1**: Assume that $x_{1:(n+1)}$ are independent given θ

$$\begin{split} \rho(x_{n+1}|x_{1:n}) &= \int \rho(x_{n+1},\theta|x_{1:n}) \, d\theta \\ &\int \frac{\rho(x_{n+1},\theta,x_{1:n})}{\rho(x_{1:n})} \, d\theta \quad \text{(Conditional probability)} \\ &\int \frac{\rho(x_{n+1}|\theta,x_{1:n})\rho(\theta|x_{1:n})\rho(x_{1:n})}{\rho(x_{1:n})} \, d\theta \quad \text{(Product rule)} \\ &= \int \rho(x_{n+1}|\theta,x_{1:n})\rho(\theta|x_{1:n}) \, d\theta \\ &= \int \rho(x_{n+1}|\theta)\rho(\theta|x_{1:n}) \, d\theta \quad \text{By Assumption 1.} \end{split}$$

Posterior predictive distribution: Bernoulli-Beta

$$X_1,\ldots,X_n\mid\theta\stackrel{iid}{\sim} Bernoulli(\theta)$$

and

$$\theta \sim Beta(a, b)$$
.

The posterior distribution can be shown to be $p(\theta|x_{1:n}) = \text{Beta}(\theta|a_n, b_n)$, where $a_n = a + \sum x_i$ and $b_n = b + n - \sum x_i$.

Posterior predictive distribution: Bernoulli-Beta

The posterior predictive can be derived to be

$$\mathbb{P}(X_{n+1} = 1 \mid x_{1:n}) = \int \mathbb{P}(X_{n+1} = 1 \mid \theta) p(\theta \mid x_{1:n}) d\theta$$

$$= \int \theta \ \mathsf{Beta}(\theta \mid a_n, b_n) d\theta$$

$$= \frac{a_n}{a_n + b_n} \ \ (\mathsf{Mean of Beta distribution}).$$

Similarly,

$$\mathbb{P}(X_{n+1}=0\mid x_{1:n})=1-\mathbb{P}(X_{n+1}=1\mid x_{1:n})=\frac{b_n}{a_n+b_n}.$$

Posterior predictive distribution (continued)

This implies that

$$p(x_{n+1}|x_{1:n}) = \begin{cases} \frac{a_n}{a_n + b_n} = \frac{(a + \sum_i x_i)}{a + b + n} & \text{if } x_{n+1} = 1\\ \frac{b_n}{a_n + b_n} = \frac{b + \sum_i (1 - x_i)}{a + b + n} & \text{if } x_{n+1} = 0 \end{cases}$$

More formally,

$$p(x_{n+1}|x_{1:n}) = \frac{a_n^{x_{n+1}}b_n^{1-x_{n+1}}}{a_n + b_n} \mathbb{1}(x_{n+1} \in \{0,1\})$$

$$= \frac{(a + \sum_i x_i)^{x_{n+1}}(b + \sum_i (1 - x_i))^{1-x_{n+1}}}{(a + b + n)} \mathbb{1}(x_{n+1} \in \{0,1\})$$

Either solution above is correct. (See page 40 of Hoff for a similar derivation of this result).

Posterior predictive distribution

Observe that the posterior predictive distribution:

- 1. Does not depend on unknown parameters.
- 2. The predictive distribution depends on the observed data.

Overall Summary

- ▶ We covered the "cast of characters" needed to work with Bayesian models
- ► These include the likelihood, prior, posterior, marginal likelihood, and posterior predictive distribution
- We derived Bayes' Theorem
- ► Bernoulli-Beta
- Conjugacy
- Marginal distribution
- Posterior predictive

Background Knowledge

- ► Familiar with Discrete and Continuous Distributions
- Can calculate expectations and variances
- Change of variables
- Mean squared error
- Sufficiency
- Confident calculating the likelihood and log-likelihood
- Confident in working with partial derivatives
- Familiar maximizing or minimizing functions (and proving they are global max/min)

Detailed Summary for Exam

- Bayes Theorem
- Likelihood
- Prior
- Posterior derivation
- Marginal likelihood
- Posterior predictive distribution
- Conjugacy
- Proportionality
- Understanding when models are appropriate for data given to you (Ex: Approval ratings for Obama)
- What is an informative prior
- What is a non-informative prior
- Proper posterior
- How do you incorporate a pilot study into your posterior analysis (Ex: See sleep study)

Supplemental Material

Below you will find supplemental material, such as exercises to help you for the exam with solutions provided.

Exercise 1

We write $X \sim \mathsf{Poisson}(\theta)$ if X has the Poisson distribution with rate $\theta > 0$, that is, its p.m.f. is

$$p(x|\theta) = Poisson(x|\theta) = e^{-\theta}\theta^x/x!$$

for $x \in \{0, 1, 2, ...\}$ (and is 0 otherwise). Suppose $X_1, ..., X_n \stackrel{iid}{\sim} \mathsf{Poisson}(\theta)$ given θ , and your prior is

$$p(\theta) = \mathsf{Gamma}(\theta|a,b) = \frac{b^a}{\Gamma(a)} \theta^{a-1} e^{-b\theta} \mathbb{1}(\theta > 0).$$

What is the posterior distribution on θ ?

Since the data is independent given θ , the likelihood factors and we get

$$p(x_{1:n}|\theta) = \prod_{i=1}^{n} p(x_i|\theta)$$
$$= \prod_{i=1}^{n} e^{-\theta} \theta^{x_i} / x_i!$$
$$\propto e^{-n\theta} \theta^{\sum x_i}.$$

Thus, using Bayes' theorem,

$$\begin{split} \rho(\theta|x_{1:n}) &\propto \rho(x_{1:n}|\theta)\rho(\theta) \\ &\propto e^{-n\theta}\theta^{\sum x_i}\theta^{a-1}e^{-b\theta}\mathbb{1}(\theta>0) \\ &\propto e^{-(b+n)\theta}\theta^{a+\sum x_i-1}\mathbb{1}(\theta>0) \\ &\propto \mathsf{Gamma}\;(\theta\mid a+\sum x_i,\;b+n). \end{split}$$

Therefore, since the posterior density must integrate to 1, we have

$$p(\theta|x_{1:n}) = \mathsf{Gamma}(\theta \mid a + \sum x_i, b + n).$$

Exercise 2

Suppose that
$$Y = \sum_{i} X_{i}$$
, where $X_{i} \mid \theta \stackrel{\textit{iid}}{\sim} \mathsf{Bernoulli}(\theta)$ for $i = 1, \dots, n$.

- a. What is the distribution of Y.
- b. What is a conjugate prior? (Provide the distribution and parameters).
- c. What is the posterior update for $\theta \mid Y$ assuming the conjugate prior in part b.
- d. Write the posterior mean $E[\theta \mid Y]$ as a weighted average of the prior mean and the sample mean, where you specify the weights.

- a. $Y \sim \text{Binomial}(n, \theta)$.
- b. A conjugate prior is $\theta \sim \text{Beta}(a, b)$ for a, b > 0 and known.
- c. The posterior update is $\theta \mid Y \sim \text{Beta}(a+y, n+b-y)$.

Recall the prior mean is a/(a+b) and the sample mean is y/n.

d. The posterior mean is

$$E[\theta \mid Y] = \frac{a+y}{a+b+n}$$

$$= \frac{a}{a+b+n} + \frac{y}{a+b+n}$$

$$= \frac{a}{a+b+n} \times \frac{a+b}{a+b} + \frac{y}{a+b+n} \times \frac{n}{n}$$

$$= \frac{a+b}{a+b+n} \times \frac{a}{a+b} + \frac{n}{a+b+n} \times \frac{y}{n}$$

$$= \frac{a+b}{a+b+n} \times \frac{a}{a+b} + \frac{n}{a+b+n} \times \frac{y}{n}$$

$$(4)$$

Above the prior mean and sample mean is in blue and the respective weights are multipled by either prior mean or sample mean.

(4)

The weights are proportional to a + b for the prior mean and n for the sample mean.

This leads to an interpretation of a and b as "prior data":

- ▶ $a \approx$ "prior number of 1's."
- ▶ $b \approx$ "prior number of 0's."
- $ightharpoonup a + b \approx$ "prior sample size"

Remark: If n >> a+b then we would inform θ according to the data. However, if n << a+b, we would inform θ according to our prior sample or historical data. (This is explained more in depth on page 39 of Hoff).

Module 1 Derivations

Class notes from Module 1 can be found below:

https://github.com/resteorts/modern-bayes/blob/master/lecturesModernBayes20/lecture-1/notes-module1.pdf

https://github.com/resteorts/modern-bayes/blob/master/lecturesModernBayes20/lecture-1/notes-module1.pdf

Additional Applied Example

Below, there is an additional applied example that you may find useful regarding this material.

How Much Do You Sleep Example

We are interested in a population of American college students and the proportion of the population that sleep at least eight hours a night, which we denote by θ .

How Much Do You Sleep Example

- ► The Gamecock, at the USC printed an internet article "College Students Don't Get Enough Sleep" (2004).
 - Most students spend six hours sleeping each night.
- 2003: University of Notre Dame's paper, Fresh Writing.
 - ▶ The article reported took random sample of 100 students:
 - "approximately 70% reported to receiving only five to six hours of sleep on the weekdays,
 - 28% receiving seven to eight,
 - ▶ and only 2% receiving the healthy nine hours for teenagers."

- ▶ Have a random sample of 27 students is taken from UF.
- ▶ 11 students record that they sleep at least eight hours each night.
- **\triangleright** Based on this information, we are interested in estimating θ .

- From USC and UND, believe it's probably true that most college students get less than eight hours of sleep.
- Want our prior to assign most of the probability to values of $\theta < 0.5$.
- From the information given, we decide that our best guess for θ is 0.3, although we think it is very possible that θ could be any value in [0,0.5].

Our Model

Our model can be summarized by the Binomial-Beta distribution

$$X|\theta \sim \mathsf{Binomial}(n,\theta)$$
 (5)

$$\theta \sim \text{Beta}(a, b)$$
 (6)

You can show that the posterior of

$$\theta \mid X \sim \text{Beta}(x+a, n-x+b)$$

Choice of a,b for Beta Prior

- ▶ Given this information, we believe that the median of θ is 0.3 and the 90th percentile is 0.5.
- ► Knowing this allows us to estimate the unknown values of *a* and *b*.
- ► How do we actually calculate *a* and *b*?

Choice of a,b for Beta Prior

We would need to solve the following equations:

$$\int_0^{0.3} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} d\theta = 0.5$$
$$\int_0^{0.5} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} d\theta = 0.9$$

In non-calculus language, this means the 0.5 quantile (50th percentile)=0.3. The 0.9 quantile (90th percentile) = 0.5.

The equations are written as percentiles above!

- ► We can easily solve this numerically in R using a numerical solver BBsolve using the BB package. .
- ▶ The documentation for this package is not great, so beware.

```
## load the BB package
library(BB)
## using percentiles
myfn <- function(shape) {</pre>
    test \leftarrow pbeta(q = c(0.3, 0.5), shape1 = shape[1],
     shape2 = shape[2]) - c(0.5, 0.9)
   return(test) }
BBsolve(c(1, 1), myfn)
##
     Successful convergence.
## $par
## [1] 3.263743 7.185121
##
## $residual
## [1] 5.905161e-08
##
## $fn.reduction
```

Using our calculations from the Beta-Binomial our model is

$$X \mid \theta \sim \text{Binomial}(27, \theta)$$

 $\theta \sim \text{Beta}(3.3, 7.2)$
 $\theta \mid x \sim \text{Beta}(x + 3.3, 27 - x + 7.2)$
 $\theta \mid 11 \sim \text{Beta}(14.3, 23.2)$

```
th \leftarrow seq(0,1,length=500)
a <- estimated$par[1]</pre>
b <- estimated$par[2]</pre>
n < -27
x < -11
prior <- dbeta(th, a, b)</pre>
like \leftarrow dbeta(th, x + 1, n - x + 1)
post \leftarrow dbeta(th, x + a, n - x + b)
plot(th, post, type = "l", ylab = "Density", lty = 2, lwd =
xlab = expression(theta))
lines(th, like, lty = 1, lwd = 3)
lines(th, prior, lty = 3, lwd = 3)
legend(0.7, 4, c("Prior", "Likelihood", "Posterior"),
lty = c(3,1,2), lwd = c(3,3,3)
```

