

Text as Data

Justin Grimmer

Professor
Department of Political Science
Stanford University

May 7th, 2019

Three categories of documents

Hand labeled

- Training set (what we'll use to estimate model)
- Validation set (what we'll use to assess model)

Unlabeled

- Test set (what we'll use the model to categorize)

Label more documents than necessary to train model

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta' \mathbf{x}_i \right)^2$$

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2$$

$$\hat{\beta} = \arg \min_{\beta} \left\{ \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2 \right\}$$

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2$$

$$\begin{aligned} \hat{\beta} &= \arg \min_{\beta} \left\{ \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2 \right\} \\ &= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y} \end{aligned}$$

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2$$

$$\begin{aligned} \hat{\beta} &= \arg \min_{\beta} \left\{ \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2 \right\} \\ &= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y} \end{aligned}$$

Problem:

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$\begin{aligned} f(\beta, \mathbf{X}, \mathbf{Y}) &= \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2 \\ \hat{\beta} &= \arg \min_{\beta} \left\{ \sum_{i=1}^N (y_i - \beta' \mathbf{x}_i)^2 \right\} \\ &= (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y} \end{aligned}$$

Problem:

- J will likely be large (perhaps $J > N$)

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$\begin{aligned} f(\beta, \mathbf{X}, \mathbf{Y}) &= \sum_{i=1}^N \left(y_i - \beta' \mathbf{x}_i \right)^2 \\ \hat{\beta} &= \arg \min_{\beta} \left\{ \sum_{i=1}^N \left(y_i - \beta' \mathbf{x}_i \right)^2 \right\} \\ &= \left(\mathbf{X}' \mathbf{X} \right)^{-1} \mathbf{X}' \mathbf{Y} \end{aligned}$$

Problem:

- J will likely be large (perhaps $J > N$)
- There many correlated variables

Regression models

Suppose we have N documents, with each document i having label $y_i \in \{-1, 1\} \rightsquigarrow \{\text{not, credit claiming}\}$

We represent each document i is $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iJ})$.

$$\begin{aligned} f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) &= \sum_{i=1}^N \left(y_i - \boldsymbol{\beta}' \mathbf{x}_i \right)^2 \\ \hat{\boldsymbol{\beta}} &= \arg \min_{\boldsymbol{\beta}} \left\{ \sum_{i=1}^N \left(y_i - \boldsymbol{\beta}' \mathbf{x}_i \right)^2 \right\} \\ &= \left(\mathbf{X}' \mathbf{X} \right)^{-1} \mathbf{X}' \mathbf{Y} \end{aligned}$$

Problem:

- J will likely be large (perhaps $J > N$)
- There many correlated variables

Predictions will be **variable**

Mean Square Error

Suppose θ is some value of the true parameter

Mean Square Error

Suppose θ is some value of the true parameter

Bias:

Mean Square Error

Suppose θ is some value of the true parameter
Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

Mean Square Error

Suppose θ is some value of the true parameter
Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

We may care about average distance from truth

Mean Square Error

Suppose θ is some value of the true parameter
Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

We may care about average distance from truth

$$E[(\hat{\theta} - \theta)^2]$$

Mean Square Error

Suppose θ is some value of the true parameter
Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

We may care about average distance from truth

$$E[(\hat{\theta} - \theta)^2] = E[\hat{\theta}^2] - 2\theta E[\hat{\theta}] + \theta^2$$

Mean Square Error

Suppose θ is some value of the true parameter
Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

We may care about average distance from truth

$$\begin{aligned} E[(\hat{\theta} - \theta)^2] &= E[\hat{\theta}^2] - 2\theta E[\hat{\theta}] + \theta^2 \\ &= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + E[\hat{\theta}]^2 - 2\theta E[\hat{\theta}] + \theta^2 \end{aligned}$$

Mean Square Error

Suppose θ is some value of the true parameter

Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

We may care about average distance from truth

$$\begin{aligned} E[(\hat{\theta} - \theta)^2] &= E[\hat{\theta}^2] - 2\theta E[\hat{\theta}] + \theta^2 \\ &= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + E[\hat{\theta}]^2 - 2\theta E[\hat{\theta}] + \theta^2 \\ &= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + (E[\hat{\theta}] - \theta)^2 \end{aligned}$$

Mean Square Error

Suppose θ is some value of the true parameter

Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

We may care about average distance from truth

$$\begin{aligned} E[(\hat{\theta} - \theta)^2] &= E[\hat{\theta}^2] - 2\theta E[\hat{\theta}] + \theta^2 \\ &= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + E[\hat{\theta}]^2 - 2\theta E[\hat{\theta}] + \theta^2 \\ &= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + (E[\hat{\theta}] - \theta)^2 \\ &= \text{Var}(\hat{\theta}) + \text{Bias}^2 \end{aligned}$$

Mean Square Error

Suppose θ is some value of the true parameter

Bias:

$$\text{Bias} = E[\hat{\theta} - \theta]$$

We may care about average distance from truth

$$\begin{aligned} E[(\hat{\theta} - \theta)^2] &= E[\hat{\theta}^2] - 2\theta E[\hat{\theta}] + \theta^2 \\ &= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + E[\hat{\theta}]^2 - 2\theta E[\hat{\theta}] + \theta^2 \\ &= E[\hat{\theta}^2] - E[\hat{\theta}]^2 + (E[\hat{\theta}] - \theta)^2 \\ &= \text{Var}(\hat{\theta}) + \text{Bias}^2 \end{aligned}$$

To reduce MSE, we are willing to induce bias to decrease variance \rightsquigarrow
methods that **shrink** coefficients toward zero

Ridge Regression

Penalty for model complexity

Ridge Regression

Penalty for model complexity

$$f(\beta, \mathbf{X}, \mathbf{Y})$$

Ridge Regression

Penalty for model complexity

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2$$

Ridge Regression

Penalty for model complexity

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \underbrace{\lambda \sum_{j=1}^J \beta_j^2}_{\text{Penalty}}$$

Ridge Regression

Penalty for model complexity

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \underbrace{\lambda \sum_{j=1}^J \beta_j^2}_{\text{Penalty}}$$

where:

Ridge Regression

Penalty for model complexity

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \underbrace{\lambda \sum_{j=1}^J \beta_j^2}_{\text{Penalty}}$$

where:

- $\beta_0 \rightsquigarrow$ intercept

Ridge Regression

Penalty for model complexity

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \underbrace{\lambda \sum_{j=1}^J \beta_j^2}_{\text{Penalty}}$$

where:

- $\beta_0 \rightsquigarrow$ intercept
- $\lambda \rightsquigarrow$ penalty parameter

Ridge Regression

Penalty for model complexity

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \underbrace{\lambda \sum_{j=1}^J \beta_j^2}_{\text{Penalty}}$$

where:

- $\beta_0 \rightsquigarrow$ intercept
- $\lambda \rightsquigarrow$ penalty parameter
- Standardized \mathbf{X} (coefficients on same scale)

Ridge Regression \rightsquigarrow Optimization

$$\beta^{\text{Ridge}} = \arg \min_{\beta} \{f(\beta, \mathbf{X}, \mathbf{Y})\}$$

Ridge Regression \rightsquigarrow Optimization

$$\begin{aligned}\beta^{\text{Ridge}} &= \arg \min_{\beta} \{f(\beta, \mathbf{X}, \mathbf{Y})\} \\ &= \arg \min_{\beta} \left\{ \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \beta_j^2 \right\}\end{aligned}$$

Ridge Regression \rightsquigarrow Optimization

$$\begin{aligned}\beta^{\text{Ridge}} &= \arg \min_{\beta} \{f(\beta, \mathbf{X}, \mathbf{Y})\} \\ &= \arg \min_{\beta} \left\{ \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \beta_j^2 \right\} \\ &= \arg \min_{\beta} \left\{ (\mathbf{Y} - \mathbf{X}'\beta)'(\mathbf{Y} - \mathbf{X}'\beta) + \lambda \beta' \beta \right\}\end{aligned}$$

Demmean the data and set $\beta_0 = \bar{y} = \sum_{i=1}^N \frac{y_i}{N}$

Ridge Regression \rightsquigarrow Optimization

$$\begin{aligned}\beta^{\text{Ridge}} &= \arg \min_{\beta} \{f(\beta, \mathbf{X}, \mathbf{Y})\} \\ &= \arg \min_{\beta} \left\{ \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \beta_j^2 \right\} \\ &= \arg \min_{\beta} \left\{ (\mathbf{Y} - \mathbf{X}'\beta)'(\mathbf{Y} - \mathbf{X}'\beta) + \lambda \beta' \beta \right\} \\ &= (\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y}\end{aligned}$$

Demmean the data and set $\beta_0 = \bar{y} = \sum_{i=1}^N \frac{y_i}{N}$

Ridge Regression \rightsquigarrow Optimization

$$\begin{aligned}\beta^{\text{Ridge}} &= \arg \min_{\beta} \{f(\beta, \mathbf{X}, \mathbf{Y})\} \\ &= \arg \min_{\beta} \left\{ \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \beta_j^2 \right\} \\ &= \arg \min_{\beta} \left\{ (\mathbf{Y} - \mathbf{X}'\beta)'(\mathbf{Y} - \mathbf{X}'\beta) + \lambda \beta' \beta \right\} \\ &= (\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y}\end{aligned}$$

Demmean the data and set $\beta_0 = \bar{y} = \sum_{i=1}^N \frac{y_i}{N}$

Ridge Regression \rightsquigarrow Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

Ridge Regression \rightsquigarrow Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y}$$

Ridge Regression \rightsquigarrow Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

$$\begin{aligned}\hat{\beta} &= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \\ &= \mathbf{X}'\mathbf{Y}\end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

$$\begin{aligned}\hat{\boldsymbol{\beta}} &= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \\ &= \mathbf{X}'\mathbf{Y} \\ \boldsymbol{\beta}^{\text{ridge}} &= (\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y}\end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

$$\begin{aligned}\hat{\beta} &= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \\ &= \mathbf{X}'\mathbf{Y} \\ \beta^{\text{ridge}} &= (\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y} \\ &= (\mathbf{I}_J + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y}\end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

$$\begin{aligned}\hat{\boldsymbol{\beta}} &= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \\ &= \mathbf{X}'\mathbf{Y} \\ \boldsymbol{\beta}^{\text{ridge}} &= (\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y} \\ &= (\mathbf{I}_J + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y} \\ &= (\mathbf{I}_J + \lambda \mathbf{I}_J)^{-1} \hat{\boldsymbol{\beta}}\end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (1)

Suppose $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$.

$$\begin{aligned}\hat{\beta} &= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \\ &= \mathbf{X}'\mathbf{Y} \\ \beta^{\text{ridge}} &= (\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y} \\ &= (\mathbf{I}_J + \lambda \mathbf{I}_J)^{-1} \mathbf{X}'\mathbf{Y} \\ &= (\mathbf{I}_J + \lambda \mathbf{I}_J)^{-1} \hat{\beta} \\ \beta_j^{\text{Ridge}} &= \frac{\hat{\beta}_j}{1 + \lambda}\end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (2)

$$\beta_j \sim \text{Normal}(0, \tau^2)$$

$$y_i \sim \text{Normal}(\beta_0 + \mathbf{x}_i' \boldsymbol{\beta}, \sigma^2)$$

Ridge Regression \rightsquigarrow Intuition (2)

$$\beta_j \sim \text{Normal}(0, \tau^2)$$

$$y_i \sim \text{Normal}(\beta_0 + \mathbf{x}_i' \boldsymbol{\beta}, \sigma^2)$$

$$p(\boldsymbol{\beta} | \mathbf{X}, \mathbf{Y}) \propto \prod_{j=1}^J p(\beta_j) \prod_{i=1}^N p(y_i | \mathbf{x}_i, \boldsymbol{\beta})$$

Ridge Regression \rightsquigarrow Intuition (2)

$$\beta_j \sim \text{Normal}(0, \tau^2)$$

$$y_i \sim \text{Normal}(\beta_0 + \mathbf{x}_i' \boldsymbol{\beta}, \sigma^2)$$

$$\begin{aligned} p(\boldsymbol{\beta} | \mathbf{X}, \mathbf{Y}) &\propto \prod_{j=1}^J p(\beta_j) \prod_{i=1}^N p(y_i | \mathbf{x}_i, \boldsymbol{\beta}) \\ &\propto \prod_{j=1}^J \frac{1}{\sqrt{2\pi}\tau} \exp\left(-\frac{\beta_j^2}{2\tau^2}\right) \prod_{i=1}^N \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y_i - \beta_0 - \mathbf{x}_i' \boldsymbol{\beta})^2}{2\sigma^2}\right) \end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (2)

$$\log p(\boldsymbol{\beta} | \mathbf{X}, \mathbf{Y}) = - \sum_{j=1}^J \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^N \frac{(y_i - \beta_0 - \mathbf{x}'_i \boldsymbol{\beta})^2}{2\sigma^2}$$

Ridge Regression \rightsquigarrow Intuition (2)

$$\begin{aligned}\log p(\boldsymbol{\beta}|\mathbf{X}, \mathbf{Y}) &= -\sum_{j=1}^J \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^N \frac{(y_i - \beta_0 - \mathbf{x}'\boldsymbol{\beta})^2}{2\sigma^2} \\ -2\sigma^2 \log p(\boldsymbol{\beta}|\mathbf{X}, \mathbf{Y}) &= \sum_{i=1}^N (y_i - \beta_0 - \mathbf{x}'\boldsymbol{\beta})^2 + \sum_{j=1}^J \frac{\sigma^2}{\tau^2} \beta_j^2\end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (2)

$$\begin{aligned}\log p(\boldsymbol{\beta}|\mathbf{X}, \mathbf{Y}) &= -\sum_{j=1}^J \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^N \frac{(y_i - \beta_0 - \mathbf{x}'\boldsymbol{\beta})^2}{2\sigma^2} \\ -2\sigma^2 \log p(\boldsymbol{\beta}|\mathbf{X}, \mathbf{Y}) &= \sum_{i=1}^N (y_i - \beta_0 - \mathbf{x}'\boldsymbol{\beta})^2 + \sum_{j=1}^J \frac{\sigma^2}{\tau^2} \beta_j^2\end{aligned}$$

where:

Ridge Regression \rightsquigarrow Intuition (2)

$$\begin{aligned}\log p(\boldsymbol{\beta}|\mathbf{X}, \mathbf{Y}) &= -\sum_{j=1}^J \frac{\beta_j^2}{2\tau^2} - \sum_{i=1}^N \frac{(y_i - \beta_0 - \mathbf{x}'\boldsymbol{\beta})^2}{2\sigma^2} \\ -2\sigma^2 \log p(\boldsymbol{\beta}|\mathbf{X}, \mathbf{Y}) &= \sum_{i=1}^N (y_i - \beta_0 - \mathbf{x}'\boldsymbol{\beta})^2 + \sum_{j=1}^J \frac{\sigma^2}{\tau^2} \beta_j^2\end{aligned}$$

where:

$$- \lambda = \frac{\sigma^2}{\tau^2}$$

Ridge Regression \rightsquigarrow Intuition (3)

Definition

Suppose \mathbf{X} is an $N \times J$ matrix. Then \mathbf{X} can be written as:

$$\mathbf{X} = \underbrace{\mathbf{U}}_{N \times N} \underbrace{\mathbf{S}}_{N \times J} \underbrace{\mathbf{V}'}_{J \times J}$$

Where:

$$\begin{aligned}\mathbf{U}'\mathbf{U} &= \mathbf{I}_N \\ \mathbf{V}'\mathbf{V} &= \mathbf{V}\mathbf{V}' = \mathbf{I}_J\end{aligned}$$

\mathbf{S} contains $\min(N, J)$ singular values, $\sqrt{\lambda_j} \geq 0$ down the diagonal and then 0's for the remaining entries

Ridge Regression \rightsquigarrow Intuition (3)

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

Ridge Regression \rightsquigarrow Intuition (3)

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

Using SVD:

Ridge Regression \rightsquigarrow Intuition (3)

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

Using SVD:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \mathbf{V} \mathbf{S}' \underbrace{(\mathbf{U}' \mathbf{U})}_{\mathbf{I}_J} \mathbf{S} \mathbf{V}'$$

Ridge Regression \rightsquigarrow Intuition (3)

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

Using SVD:

$$\begin{aligned} \frac{1}{N} \mathbf{X}' \mathbf{X} &= \mathbf{V} \mathbf{S}' \underbrace{(\mathbf{U}' \mathbf{U})}_{\mathbf{I}_J} \mathbf{S} \mathbf{V}' \\ &= \frac{1}{N} \mathbf{V} \mathbf{S}' \mathbf{S} \mathbf{V}' \end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (3)

Recall: PCA:

$$\frac{1}{N} \mathbf{X}' \mathbf{X} = \underbrace{\mathbf{W}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{W}'}_{\text{eigenvectors}}$$

Using SVD:

$$\begin{aligned} \frac{1}{N} \mathbf{X}' \mathbf{X} &= \mathbf{V} \mathbf{S}' \underbrace{(\mathbf{U}' \mathbf{U})}_{\mathbf{I}_J} \mathbf{S} \mathbf{V}' \\ &= \frac{1}{N} \mathbf{V} \mathbf{S}' \mathbf{S} \mathbf{V}' \\ &= \underbrace{\mathbf{V}}_{\text{eigenvectors}} \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_J \end{pmatrix} \underbrace{\mathbf{V}'}_{\text{eigenvectors}} \end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (3)

We can write the predicted values for a regular regression as

$$\begin{aligned}\hat{Y} &= \mathbf{X}\hat{\beta} \\ &= \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{U}\mathbf{U}'\mathbf{Y} = \sum_{j=1}^J \mathbf{u}_j\mathbf{u}_j'\mathbf{Y}\end{aligned}$$

Ridge Regression \rightsquigarrow Intuition (3)

We can write the predicted values for a regular regression as

$$\begin{aligned}\hat{Y} &= \mathbf{X}\hat{\beta} \\ &= \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{U}\mathbf{U}'\mathbf{Y} = \sum_{j=1}^J u_j u_j' \mathbf{Y}\end{aligned}$$

We can write β^{ridge} as

Ridge Regression \rightsquigarrow Intuition (3)

We can write the predicted values for a regular regression as

$$\begin{aligned}\hat{Y} &= \mathbf{X}\hat{\beta} \\ &= \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{U}\mathbf{U}'\mathbf{Y} = \sum_{j=1}^J u_j u_j' \mathbf{Y}\end{aligned}$$

We can write β^{ridge} as

$$\hat{Y}^{\text{ridge}} = \mathbf{X}(\mathbf{X}'\mathbf{X} + \lambda I_J)^{-1}\mathbf{X}'\mathbf{Y}$$

Ridge Regression \rightsquigarrow Intuition (3)

We can write the predicted values for a regular regression as

$$\begin{aligned}\hat{Y} &= \mathbf{X}\hat{\beta} \\ &= \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{U}\mathbf{U}'\mathbf{Y} = \sum_{j=1}^J u_j u_j' \mathbf{Y}\end{aligned}$$

We can write β^{ridge} as

$$\begin{aligned}\hat{Y}^{\text{ridge}} &= \mathbf{X}(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{U}\tilde{\mathbf{S}}\mathbf{U}'\mathbf{Y}\end{aligned}$$

Where

$$\tilde{\mathbf{S}} = \left[\mathbf{S}(\mathbf{S}'\mathbf{S} + \lambda \mathbf{I}_J)^{-1}\mathbf{S} \right]$$

Ridge Regression \rightsquigarrow Intuition (3)

We can write the predicted values for a regular regression as

$$\begin{aligned}\hat{Y} &= \mathbf{X}\hat{\beta} \\ &= \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{U}\mathbf{U}'\mathbf{Y} = \sum_{j=1}^J \mathbf{u}_j \mathbf{u}_j' \mathbf{Y}\end{aligned}$$

We can write β^{ridge} as

$$\begin{aligned}\hat{Y}^{\text{ridge}} &= \mathbf{X}(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J)^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{U}\tilde{\mathbf{S}}\mathbf{U}'\mathbf{Y}\end{aligned}$$

Where

$$\tilde{\mathbf{S}} = \left[\mathbf{S}(\mathbf{S}'\mathbf{S} + \lambda \mathbf{I}_J)^{-1}\mathbf{S} \right]$$

Which we can write as:

$$\hat{Y}^{\text{ridge}} = \sum_{j=1}^J \mathbf{u}_j \frac{\lambda_j}{\lambda_j + \lambda} \mathbf{u}_j' \mathbf{Y}$$

Degrees of Freedom for Ridge

We will say that the degrees of freedom for Ridge regression with penalty λ is

$$\text{dof}(\lambda) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \lambda}$$

Lasso Regression Objective Function

Different Penalty for Model Complexity

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \underbrace{|\beta_j|}_{\text{Penalty}}$$

Lasso Regression Objective Function

Different Penalty for Model Complexity

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \underbrace{|\beta_j|}_{\text{Penalty}}$$

Lasso Regression Optimization

Definition

Coordinate Descent Algorithms:

Consider $g : \mathbb{R}^J \rightarrow \mathbb{R}$. Our goal is to find $\mathbf{x}^* \in \mathbb{R}^J$ such that $g(\mathbf{x}^*) \leq g(\mathbf{x})$ for all $\mathbf{x} \in \mathbb{R}$.

To find \mathbf{x}^* :

Until convergence: for each iteration t and each coordinate j

$$x_j^{t+1} = \arg \min_{x_j \in \mathbb{R}} g(x_1^{t+1}, x_2^{t+1}, \dots, x_{j-1}^{t+1}, x_j, x_{j+1}^t, \dots, x_J^t)$$

Lasso Regression Optimization: Coordinate Descent

$$\tilde{f}(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J |\beta_j|$$

Lasso Regression Optimization: Coordinate Descent

$$\tilde{f}(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J |\beta_j|$$

- **Case 1:** If $\beta_j = 0 \rightsquigarrow$ not differentiable. But $\beta_j = 0$

Lasso Regression Optimization: Coordinate Descent

$$\tilde{f}(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J |\beta_j|$$

- **Case 1:** If $\beta_j = 0 \rightsquigarrow$ not differentiable. But $\beta_j = 0$
- **Case 2:** If $\beta_j > (<) 0 \rightsquigarrow$ differentiable \rightsquigarrow differentiate and solve for β_j

Lasso Regression Optimization: Coordinate Descent

$$\tilde{f}(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J |\beta_j|$$

- **Case 1:** If $\beta_j = 0 \rightsquigarrow$ not differentiable. But $\beta_j = 0$
- **Case 2:** If $\beta_j > (<) 0 \rightsquigarrow$ differentiable \rightsquigarrow differentiate and solve for β_j

Define $\tilde{y}_i^j = \beta_0 + \sum_{l \neq j} x_{il} \beta_l$

Lasso Regression Optimization: Coordinate Descent

$$\tilde{f}(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J |\beta_j|$$

- **Case 1:** If $\beta_j = 0 \rightsquigarrow$ not differentiable. But $\beta_j = 0$
- **Case 2:** If $\beta_j > (<) 0 \rightsquigarrow$ differentiable \rightsquigarrow differentiate and solve for β_j

Define $\tilde{y}_i^j = \beta_0 + \sum_{l \neq j} x_{il} \beta_l$

$$r^j \equiv \frac{1}{N} \sum_{i=1}^N x_{ij} (y_i - \tilde{y}_i^j)$$

Lasso Regression Optimization: Coordinate Descent

$$\tilde{f}(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J |\beta_j|$$

- **Case 1:** If $\beta_j = 0 \rightsquigarrow$ not differentiable. But $\beta_j = 0$
- **Case 2:** If $\beta_j > (<) 0 \rightsquigarrow$ differentiable \rightsquigarrow differentiate and solve for β_j

Define $\tilde{y}_i^j = \beta_0 + \sum_{l \neq j} x_{il} \beta_l$

$r^j \equiv \frac{1}{N} \sum_{i=1}^N x_{ij} (y_i - \tilde{y}_i^j)$

Update step for β_j is

Lasso Regression Optimization: Coordinate Descent

$$\tilde{f}(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J |\beta_j|$$

- **Case 1:** If $\beta_j = 0 \rightsquigarrow$ not differentiable. But $\beta_j = 0$
- **Case 2:** If $\beta_j > (<) 0 \rightsquigarrow$ differentiable \rightsquigarrow differentiate and solve for β_j

Define $\tilde{y}_i^j = \beta_0 + \sum_{l \neq j} x_{il} \beta_l$

$$r^j \equiv \frac{1}{N} \sum_{i=1}^N x_{ij} (y_i - \tilde{y}_i^j)$$

Update step for β_j is

$$\beta_j \leftarrow \text{sign}(r^j) \max(|r^j| - \lambda, 0)$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) = (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^J |\beta_j|$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$\begin{aligned} f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) &= (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^J |\beta_j| \\ &= -2\mathbf{X}'\mathbf{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^J |\beta_j| \end{aligned}$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$\begin{aligned} f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) &= (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^J |\beta_j| \\ &= -2\mathbf{X}'\mathbf{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^J |\beta_j| \end{aligned}$$

The coefficient is

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$\begin{aligned} f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) &= (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^J |\beta_j| \\ &= -2\mathbf{X}'\mathbf{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^J |\beta_j| \end{aligned}$$

The coefficient is

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$\begin{aligned}f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) &= (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^J |\beta_j| \\&= -2\mathbf{X}'\mathbf{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^J |\beta_j|\end{aligned}$$

The coefficient is

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

- $\text{sign}(\cdot) \rightsquigarrow 1$ or -1

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Suppose again $\mathbf{X}'\mathbf{X} = \mathbf{I}_J$

$$\begin{aligned} f(\boldsymbol{\beta}, \mathbf{X}, \mathbf{Y}) &= (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \sum_{j=1}^J |\beta_j| \\ &= -2\mathbf{X}'\mathbf{Y}\boldsymbol{\beta} + \boldsymbol{\beta}'\boldsymbol{\beta} + \lambda \sum_{j=1}^J |\beta_j| \end{aligned}$$

The coefficient is

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

- $\text{sign}(\cdot) \rightsquigarrow 1 \text{ or } -1$
- $\left(|\hat{\beta}_j| - \lambda \right)_+ = \max(|\hat{\beta}_j| - \lambda, 0)$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Compare soft assignment

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

With hard assignment, selecting M biggest components

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

With hard assignment, selecting M biggest components

$$\beta_j^{\text{subset}} = \hat{\beta}_j \cdot I \left(|\hat{\beta}_j| \geq |\hat{\beta}_{(M)}| \right)$$

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

With hard assignment, selecting M biggest components

$$\beta_j^{\text{subset}} = \hat{\beta}_j \cdot I\left(|\hat{\beta}_j| \geq |\hat{\beta}_{(M)}|\right)$$

Intuition 2: Prior on coefficients \rightsquigarrow Laplace “The Bayesian LASSO”

Lasso Regression \rightsquigarrow Intuition 1, Soft Thresholding

Compare soft assignment

$$\beta_j^{\text{LASSO}} = \text{sign}(\hat{\beta}_j) \left(|\hat{\beta}_j| - \lambda \right)_+$$

With hard assignment, selecting M biggest components

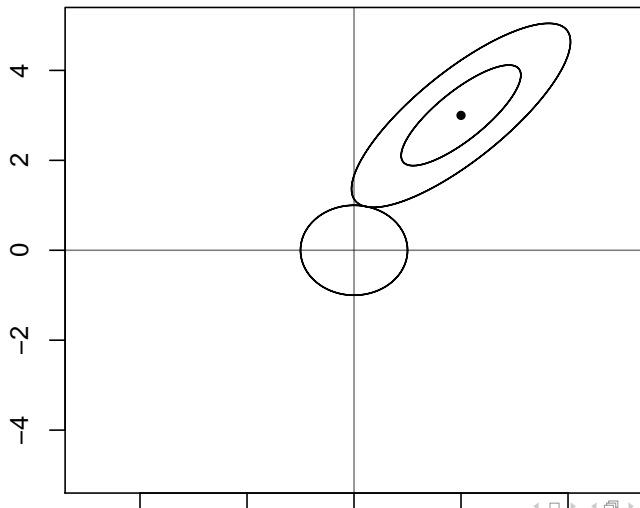
$$\beta_j^{\text{subset}} = \hat{\beta}_j \cdot I\left(|\hat{\beta}_j| \geq |\hat{\beta}_{(M)}|\right)$$

Intuition 2: Prior on coefficients \rightsquigarrow Laplace “The Bayesian LASSO”

Why does LASSO induce sparsity?

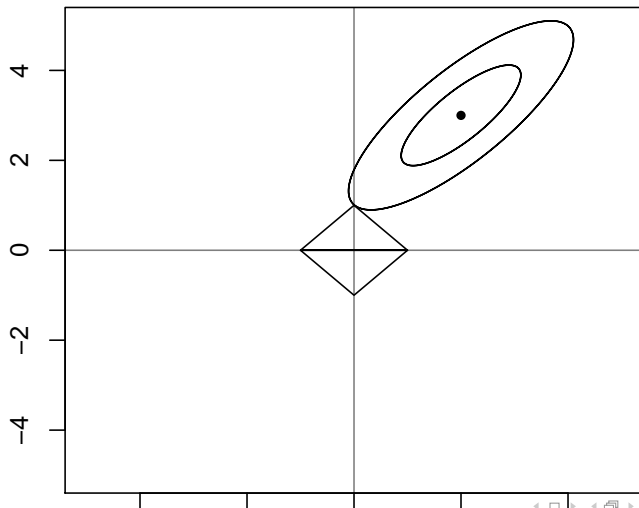
Comparing Ridge and LASSO

Ridge Regression



Comparing Ridge and LASSO

LASSO Regression



Comparing Ridge and LASSO

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$

Comparing Ridge and LASSO

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$

Under ridge:

Comparing Ridge and LASSO

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$

Under ridge:

$$\sum_{j=1}^2 \beta_j^2 = \frac{1}{2} + \frac{1}{2} = 1$$

Comparing Ridge and LASSO

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$

Under ridge:

$$\sum_{j=1}^2 \beta_j^2 = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^2 \tilde{\beta}_j^2 = 1 + 0 = 1$$

Comparing Ridge and LASSO

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$

Under ridge:

$$\sum_{j=1}^2 \beta_j^2 = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^2 \tilde{\beta}_j^2 = 1 + 0 = 1$$

Under LASSO

Comparing Ridge and LASSO

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$

Under ridge:

$$\sum_{j=1}^2 \beta_j^2 = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^2 \tilde{\beta}_j^2 = 1 + 0 = 1$$

Under LASSO

$$\sum_{j=1}^2 |\beta_j| = \frac{1}{\sqrt{2}} + \frac{1}{\sqrt{2}} = \sqrt{2}$$

Comparing Ridge and LASSO

Contrast $\beta = (\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}})$ and $\tilde{\beta} = (1, 0)$

Under ridge:

$$\sum_{j=1}^2 \beta_j^2 = \frac{1}{2} + \frac{1}{2} = 1$$

$$\sum_{j=1}^2 \tilde{\beta}_j^2 = 1 + 0 = 1$$

Under LASSO

$$\sum_{j=1}^2 |\beta_j| = \frac{1}{\sqrt{2}} + \frac{1}{\sqrt{2}} = \sqrt{2}$$

$$\sum_{j=1}^2 |\tilde{\beta}_j| = 1 + 0 = 1$$

Ridge and LASSO: The Elastic-Net

Combining the two criteria \rightsquigarrow Elastic-Net

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \left(\frac{1}{2} (1 - \alpha) \beta_j^2 + \alpha |\beta_j| \right)$$

Ridge and LASSO: The Elastic-Net

Combining the two criteria \rightsquigarrow Elastic-Net

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \left(\frac{1}{2} (1 - \alpha) \beta_j^2 + \alpha |\beta_j| \right)$$

The new update step (for coordinate descent:)

Ridge and LASSO: The Elastic-Net

Combining the two criteria \rightsquigarrow Elastic-Net

$$f(\beta, \mathbf{X}, \mathbf{Y}) = \frac{1}{2N} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^J \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^J \left(\frac{1}{2} (1 - \alpha) \beta_j^2 + \alpha |\beta_j| \right)$$

The new update step (for coordinate descent:)

$$\beta_j \leftarrow \frac{\text{sign}(r^j) \max(|r^j| - \lambda \alpha, 0)}{1 + \lambda(1 - \alpha)}$$

Selecting λ

How do we determine λ ? \rightsquigarrow Cross validation

Selecting λ

How do we determine λ ? \rightsquigarrow Cross validation

Applying models gives score (probability) of document belong to class \rightsquigarrow
threshold to classify

Selecting λ

How do we determine λ ? \rightsquigarrow Cross validation

Applying models gives score (probability) of document belong to class \rightsquigarrow
threshold to classify

Loss Functions and Model Complexity

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Loss Functions and Model Complexity

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

$$\mu_i = f(\boldsymbol{\beta}, \mathbf{x}_i)$$

Use MLE to obtain $\hat{\boldsymbol{\beta}}$.

Loss Functions and Model Complexity

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$Y_i \sim \text{Distribution}(\mu_i, \phi)$$

$$\mu_i = f(\boldsymbol{\beta}, \mathbf{x}_i)$$

Use MLE to obtain $\hat{\boldsymbol{\beta}}$.

Potential **loss** functions:

Loss Functions and Model Complexity

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$\begin{aligned} Y_i &\sim \text{Distribution}(\mu_i, \phi) \\ \mu_i &= f(\boldsymbol{\beta}, \mathbf{x}_i) \end{aligned}$$

Use MLE to obtain $\hat{\boldsymbol{\beta}}$.

Potential **loss** functions:

$$L(Y_i, f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i))$$

Loss Functions and Model Complexity

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$\begin{aligned} Y_i &\sim \text{Distribution}(\mu_i, \phi) \\ \mu_i &= f(\boldsymbol{\beta}, \mathbf{x}_i) \end{aligned}$$

Use MLE to obtain $\hat{\boldsymbol{\beta}}$.

Potential **loss** functions:

$$L(Y_i, f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i)) = (Y_i - f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i))^2$$

Loss Functions and Model Complexity

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$\begin{aligned} Y_i &\sim \text{Distribution}(\mu_i, \phi) \\ \mu_i &= f(\boldsymbol{\beta}, \mathbf{x}_i) \end{aligned}$$

Use MLE to obtain $\hat{\boldsymbol{\beta}}$.

Potential **loss** functions:

$$\begin{aligned} L(Y_i, f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i)) &= (Y_i - f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i))^2 \\ &= |Y_i - f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i)| \end{aligned}$$

Loss Functions and Model Complexity

Suppose observations i have dependent variables Y_i and covariates $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iP})$.

Assume:

$$\begin{aligned} Y_i &\sim \text{Distribution}(\mu_i, \phi) \\ \mu_i &= f(\boldsymbol{\beta}, \mathbf{x}_i) \end{aligned}$$

Use MLE to obtain $\hat{\boldsymbol{\beta}}$.

Potential **loss** functions:

$$\begin{aligned} L(Y_i, f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i)) &= (Y_i - f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i))^2 \\ &= |Y_i - f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i)| \\ &= I(Y_i = I(f(\hat{\boldsymbol{\beta}}, \mathbf{x}_i) > \tau)) \end{aligned}$$

Training and Test Sets

The useful “fiction” of training and test sets:

Training and Test Sets

The useful “fiction” of training and test sets:

- Training set: data set used to fit the model

Training and Test Sets

The useful “fiction” of training and test sets:

- Training set: data set used to fit the model
- Test set: data used to evaluate fit of the model

Training and Test Sets

The useful “fiction” of training and test sets:

- Training set: data set used to fit the model
- Test set: data used to evaluate fit of the model

Even if no division, useful to think about **systematic** components of data.

Loss Functions and Model Complexity

Suppose that we have:

=

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$

=

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

=

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

=

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

$$\text{Error}_{\text{in}} =$$

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

$$\text{Error}_{\text{in}} = \sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\beta}, \mathbf{x}_i))$$

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

$$\text{Error}_{\text{in}} = \sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\beta}, \mathbf{x}_i))$$

We'd like to estimate out of sample performance with

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

$$\text{Error}_{\text{in}} = \sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\beta}, \mathbf{x}_i))$$

We'd like to estimate out of sample performance with

$$\text{Error}_{\text{out}} = \mathbb{E}[L(\mathbf{Y}_{i \in \mathcal{O}}, f(\hat{\beta}, \mathbf{x}_{i \in \mathcal{O}})) | \mathcal{T}]$$

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

$$\text{Error}_{\text{in}} = \sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\beta}, \mathbf{x}_i))$$

We'd like to estimate out of sample performance with

$$\text{Error}_{\text{out}} = \mathbb{E}[L(\mathbf{Y}_{i \in \mathcal{O}}, f(\hat{\beta}, \mathbf{x}_{i \in \mathcal{O}})) | \mathcal{T}]$$

where the expectation is taken over **samples** for test sets and supposes we have a training set.

Loss Functions and Model Complexity

Suppose that we have:

- Training sets, \mathcal{T} , with $|\mathcal{T}| = N_{\text{train}}$
- Test sets, \mathcal{O} with $|\mathcal{O}| = N_{\text{test}}$

Training (in-sample) error is:

$$\text{Error}_{\text{in}} = \sum_{i \in \mathcal{T}} \frac{1}{N_{\text{train}}} L(Y_i, f(\hat{\beta}, \mathbf{x}_i))$$

We'd like to estimate out of sample performance with

$$\text{Error}_{\text{out}} = E[L(\mathbf{Y}_{i \in \mathcal{O}}, f(\hat{\beta}, \mathbf{x}_{i \in \mathcal{O}})) | \mathcal{T}]$$

where the expectation is taken over **samples** for test sets and supposes we have a training set.

$$\text{Error} = E \left[E[L(\mathbf{Y}, f(\hat{\beta}, \mathbf{X})) | \mathcal{T}] \right]$$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

$\text{var}(\epsilon_i) = \sigma_\epsilon^2$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

$\text{var}(\epsilon_i) = \sigma_\epsilon^2$

Define $f(\hat{\beta}, \mathbf{x}) = \hat{f}(\mathbf{x})$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

$\text{var}(\epsilon_i) = \sigma_\epsilon^2$

Define $f(\hat{\beta}, \mathbf{x}) = \hat{f}(\mathbf{x})$

With squared error loss:

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

$\text{var}(\epsilon_i) = \sigma_\epsilon^2$

Define $f(\hat{\beta}, \mathbf{x}) = \hat{f}(\mathbf{x})$

With squared error loss:

$$\text{Error}(\mathbf{x}_0) = E[(Y_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]$$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

$\text{var}(\epsilon_i) = \sigma_\epsilon^2$

Define $f(\hat{\beta}, \mathbf{x}) = \hat{f}(\mathbf{x})$

With squared error loss:

$$\begin{aligned}\text{Error}(\mathbf{x}_0) &= E[(Y_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0] \\ &= E[(f(\mathbf{x}_i) + \epsilon_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0]\end{aligned}$$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

$\text{var}(\epsilon_i) = \sigma_\epsilon^2$

Define $f(\hat{\beta}, \mathbf{x}) = \hat{f}(\mathbf{x})$

With squared error loss:

$$\begin{aligned}\text{Error}(\mathbf{x}_0) &= E[(Y_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0] \\ &= E[(f(\mathbf{x}_i) + \epsilon_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0] \\ &= \sigma_\epsilon^2 + \left[f(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)] \right]^2 + E[\left(\hat{f}(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)] \right)^2]\end{aligned}$$

Loss Functions and Model Complexity

Suppose $Y_i = f(\mathbf{x}_i) + \epsilon_i$

Where $E[\epsilon_i] = 0$

$\text{var}(\epsilon_i) = \sigma_\epsilon^2$

Define $f(\hat{\beta}, \mathbf{x}) = \hat{f}(\mathbf{x})$

With squared error loss:

$$\begin{aligned}\text{Error}(\mathbf{x}_0) &= E[(Y_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0] \\ &= E[(f(\mathbf{x}_i) + \epsilon_i - \hat{f}(\mathbf{x}_i))^2 | \mathbf{x}_i = \mathbf{x}_0] \\ &= \sigma_\epsilon^2 + \left[f(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)] \right]^2 + E[\left(\hat{f}(\mathbf{x}_0) - E[\hat{f}(\mathbf{x}_0)] \right)^2] \\ &= \text{Irreducible error} + \text{Bias}^2 + \text{Variance}\end{aligned}$$

Probit Regression (for motivational purposes)

Suppose:

$$\begin{aligned} Y_i &\sim \text{Bernoulli}(\pi_i) \\ \pi_i &= \Phi(\beta' \mathbf{x}_i) \end{aligned}$$

where $\Phi(\cdot)$ is the cumulative normal distribution.

Implies log-likelihood

$$\log L(\beta | \mathbf{X}, \mathbf{Y}) = \sum_{i=1}^N \left[Y_i \log \Phi(\beta' \mathbf{x}_i) + (1 - Y_i) \log(1 - \Phi(\beta' \mathbf{x}_i)) \right]$$

Log-likelihood is a **loss function** \rightsquigarrow overly optimistic: improves with more parameters

How Do We Build A Model?

There are many ways to fit models

And many choices made when performing model fit

How do we choose?

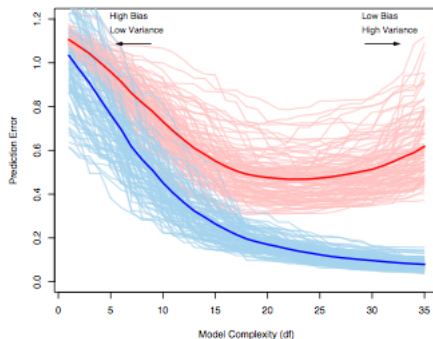


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\hat{\text{err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $\text{E}[\hat{\text{err}}]$.

How Do We Build A Model?

There are many ways to fit models

And many choices made when performing model fit

How do we choose?

Bad way to choose:

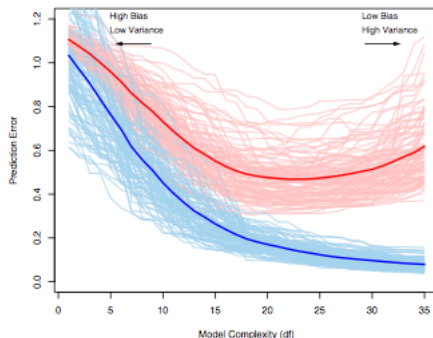


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\hat{\text{err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $\text{E}[\hat{\text{err}}]$.

How Do We Build A Model?

There are many ways to fit models

And many choices made when performing model fit

How do we choose?

Bad way to choose: within sample model fit (HTF Figure 7.1)

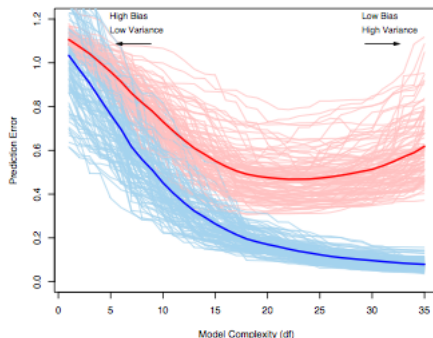


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\hat{\text{err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $\text{E}[\hat{\text{err}}]$.

How Do We Build A Model?

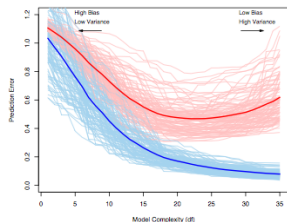


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\overline{\text{Err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $\mathbb{E}[\overline{\text{Err}}]$.

Model **overfit** \rightsquigarrow in sample error is **optimistic**:

How Do We Build A Model?

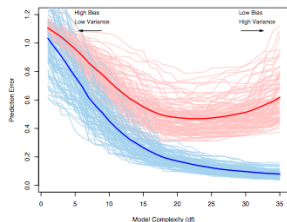


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\overline{\text{Err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $E[\overline{\text{Err}}]$.

Model **overfit** \rightsquigarrow in sample error is **optimistic**:

- Some model complexity captures **systematic** features of the data

How Do We Build A Model?

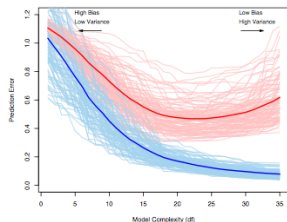


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\overline{\text{Err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $E[\overline{\text{Err}}]$.

Model **overfit** \rightsquigarrow in sample error is **optimistic**:

- Some model complexity captures **systematic** features of the data
- Characteristics found in both training and test set

How Do We Build A Model?

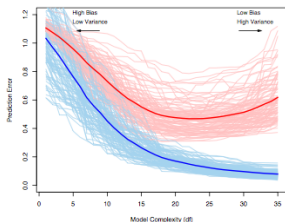


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\overline{\text{Err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $E[\overline{\text{Err}}]$.

Model **overfit** \rightsquigarrow in sample error is **optimistic**:

- Some model complexity captures **systematic** features of the data
- Characteristics found in both training and test set
- Reduces error in both training and test set

How Do We Build A Model?

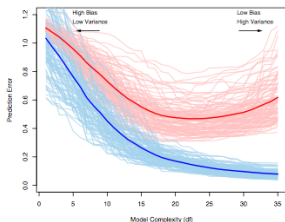


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\overline{\text{Err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $E[\overline{\text{Err}}]$.

Model **overfit** \rightsquigarrow in sample error is **optimistic**:

- Some model complexity captures **systematic** features of the data
- Characteristics found in both training and test set
- Reduces error in both training and test set
- Additional model complexity: **idiosyncratic** features of the training set

How Do We Build A Model?

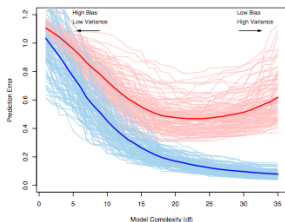


FIGURE 7.1. Behavior of test sample and training sample error as the model complexity is varied. The light blue curves show the training error $\overline{\text{Err}}$, while the light red curves show the conditional test error Err_T for 100 training sets of size 50 each, as the model complexity is increased. The solid curves show the expected test error Err and the expected training error $\mathbb{E}[\overline{\text{Err}}]$.

Model **overfit** \rightsquigarrow in sample error is **optimistic**:

- Some model complexity captures **systematic** features of the data
- Characteristics found in both training and test set
- Reduces error in both training and test set
- Additional model complexity: **idiosyncratic** features of the training set
- Reduces error in training set, increases error in test set

How Do We Choose Covariates?

Best model **depends on task**

- Causal inference observational study: make treatment assignment ignorable
- Prediction: improve predictive performance

Stepwise Regression

Suppose we have P covariates.
 2^P potential models

Stepwise Regression

Suppose we have P covariates.

2^P potential models

Stepwise procedures

Stepwise Regression

Suppose we have P covariates.

2^P potential models

Stepwise procedures

1) Forward selection

- a) No variables in model.
- b) Check all variables p-value if include, include lowest p-value
- c) Repeat until included p-value is above some threshold

Stepwise Regression

Suppose we have P covariates.

2^P potential models

Stepwise procedures

1) Forward selection

- a) No variables in model.
- b) Check all variables p-value if include, include lowest p-value
- c) Repeat until included p-value is above some threshold

2) Backward elimination

- a) Fit model with all variables (if possible)
- b) Remove variable with largest p-value
- c) Repeat until potentially excluded p-value is below some threshold

Stepwise Regression

Suppose we have P covariates.

2^P potential models

Stepwise procedures

1) Forward selection

- a) No variables in model.
- b) Check all variables p-value if include, include lowest p-value
- c) Repeat until included p-value is above some threshold

2) Backward elimination

- a) Fit model with all variables (if possible)
- b) Remove variable with largest p-value
- c) Repeat until potentially excluded p-value is below some threshold

Problematic:

- 1) Not optimal model selection (path dependent)
- 2) P-value \neq objective of model

Analytic Solutions

Approximate optimism and compensate in loss function.

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

$$-2\mathbb{E}[\log P_{\hat{\beta}}(Y)] = -2 \left[\mathbb{E}[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y})] - d \right]$$

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

$$\begin{aligned} -2\mathbb{E}[\log P_{\hat{\beta}}(Y)] &= -2 \left[\mathbb{E}[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y})] - d \right] \\ \text{AIC} &= -2 \left[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y}) - d \right] \end{aligned}$$

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

$$\begin{aligned} -2\mathbb{E}[\log P_{\hat{\beta}}(Y)] &= -2 \left[\mathbb{E}[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y})] - d \right] \\ \text{AIC} &= -2 \left[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y}) - d \right] \end{aligned}$$

where d is the number of parameters in the model

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

$$\begin{aligned} -2\mathbb{E}[\log P_{\hat{\beta}}(Y)] &= -2 \left[\mathbb{E}[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y})] - d \right] \\ \text{AIC} &= -2 \left[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y}) - d \right] \end{aligned}$$

where d is the number of parameters in the model

- Intuition: balances model fit with penalty for complexity

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

$$\begin{aligned} -2\mathbb{E}[\log P_{\hat{\beta}}(Y)] &= -2 \left[\mathbb{E}[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y})] - d \right] \\ \text{AIC} &= -2 \left[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y}) - d \right] \end{aligned}$$

where d is the number of parameters in the model

- Intuition: balances model fit with penalty for complexity
- Derived from method to estimate **optimism** in likelihood based models

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

$$\begin{aligned} -2\mathbb{E}[\log P_{\hat{\beta}}(Y)] &= -2 \left[\mathbb{E}[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y})] - d \right] \\ \text{AIC} &= -2 \left[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y}) - d \right] \end{aligned}$$

where d is the number of parameters in the model

- Intuition: balances model fit with penalty for complexity
- Derived from method to estimate **optimism** in likelihood based models
- Derived from a method to compute similarity between estimated model and true model (under assumptions of course)

Analytic Solutions

Approximate optimism and compensate in loss function.

Akaike Information Criterion (AIC) \rightsquigarrow Minimize

As $N \rightarrow \infty$

$$\begin{aligned} -2\mathbb{E}[\log P_{\hat{\beta}}(Y)] &= -2 \left[\mathbb{E}[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y})] - d \right] \\ \text{AIC} &= -2 \left[\log L(\hat{\beta}|\mathbf{X}, \mathbf{Y}) - d \right] \end{aligned}$$

where d is the number of parameters in the model

- Intuition: balances model fit with penalty for complexity
- Derived from method to estimate **optimism** in likelihood based models
- Derived from a method to compute similarity between estimated model and true model (under assumptions of course)
- Can be extended to general models, though requires estimate of irresolvable error

Analytic Solutions

Bayesian Information Criterion (BIC) [Schwarz Criterion]

Analytic Solutions

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$\text{BIC} = -2 \log L(\hat{\beta} | \mathbf{X}, \mathbf{Y}) + (\log N)d$$

Analytic Solutions

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$\text{BIC} = -2 \log L(\hat{\beta} | \mathbf{X}, \mathbf{Y}) + (\log N)d$$

where d is again the effective number of parameters

Analytic Solutions

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$\text{BIC} = -2 \log L(\hat{\beta} | \mathbf{X}, \mathbf{Y}) + (\log N)d$$

where d is again the effective number of parameters

- Intuition: balances model fit with penalty for complexity

Analytic Solutions

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$\text{BIC} = -2 \log L(\hat{\beta} | \mathbf{X}, \mathbf{Y}) + (\log N)d$$

where d is again the effective number of parameters

- Intuition: balances model fit with penalty for complexity
- Derived from **Bayesian** approach to model selection

Analytic Solutions

Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$\text{BIC} = -2 \log L(\hat{\beta} | \mathbf{X}, \mathbf{Y}) + (\log N)d$$

where d is again the effective number of parameters

- Intuition: balances model fit with penalty for complexity
- Derived from **Bayesian** approach to model selection
- Approximation to Bayes' factor

Analytic Solutions

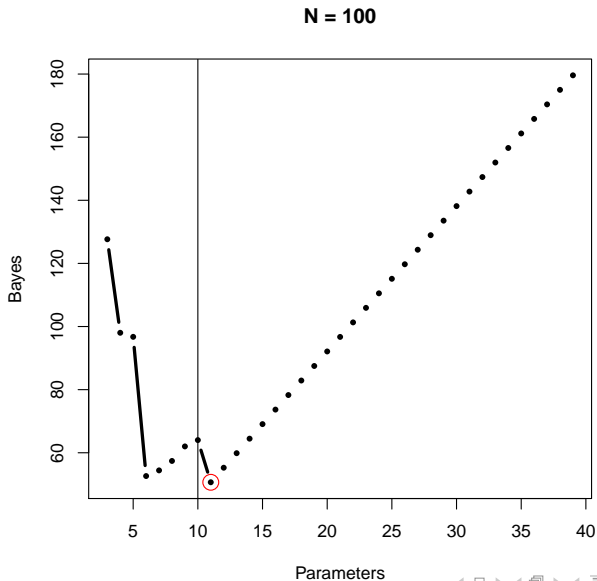
Bayesian Information Criterion (BIC) [Schwarz Criterion]

$$\text{BIC} = -2 \log L(\hat{\beta} | \mathbf{X}, \mathbf{Y}) + (\log N)d$$

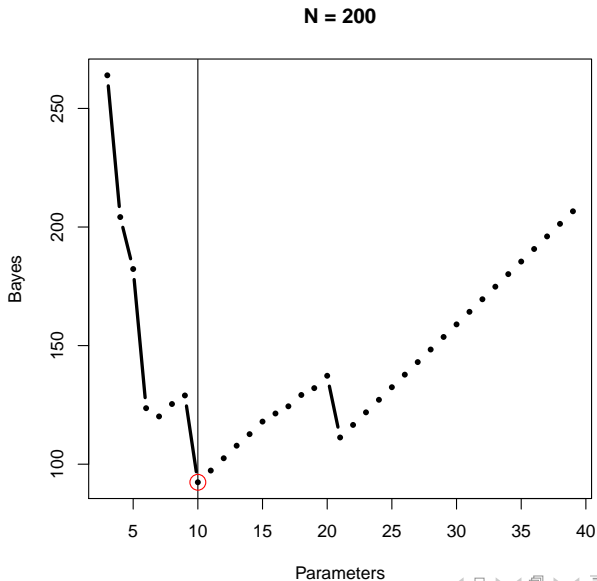
where d is again the effective number of parameters

- Intuition: balances model fit with penalty for complexity
- Derived from **Bayesian** approach to model selection
- Approximation to Bayes' factor
- **Penalizes more heavily than AIC**

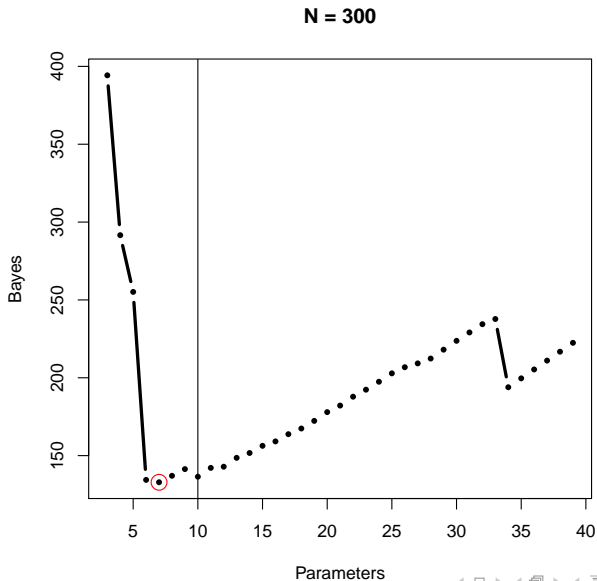
BIC or AIC?



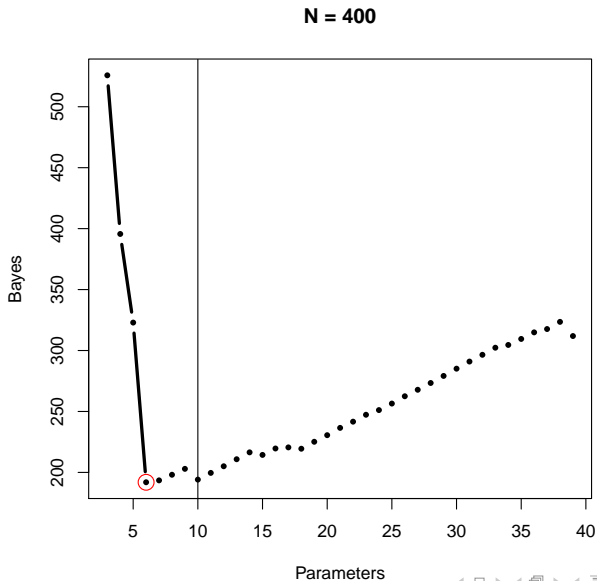
BIC or AIC?



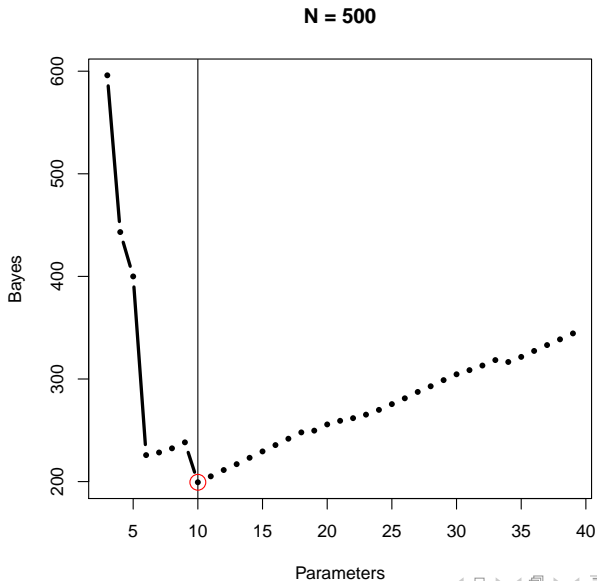
BIC or AIC?



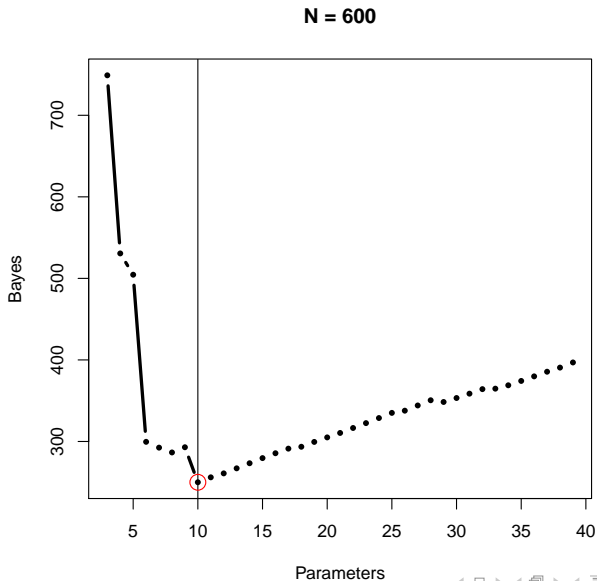
BIC or AIC?



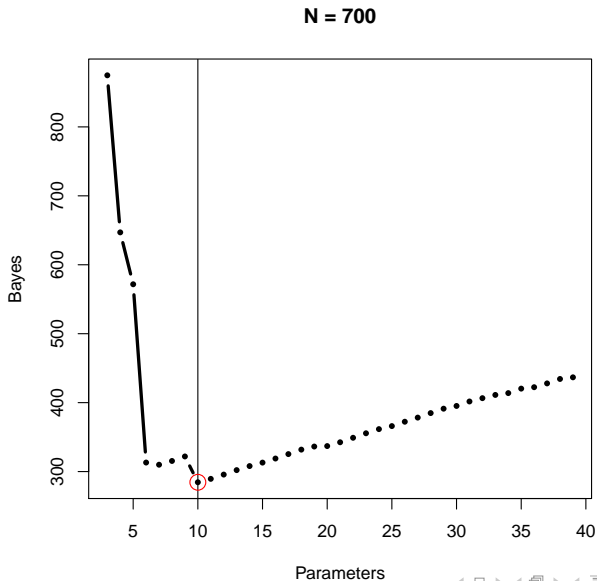
BIC or AIC?



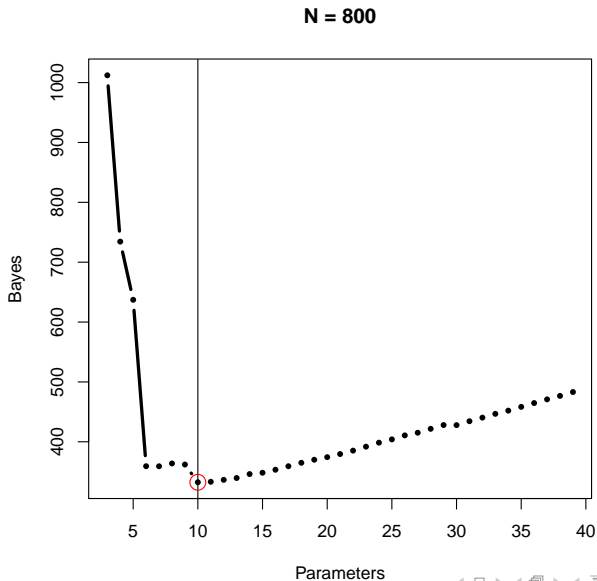
BIC or AIC?



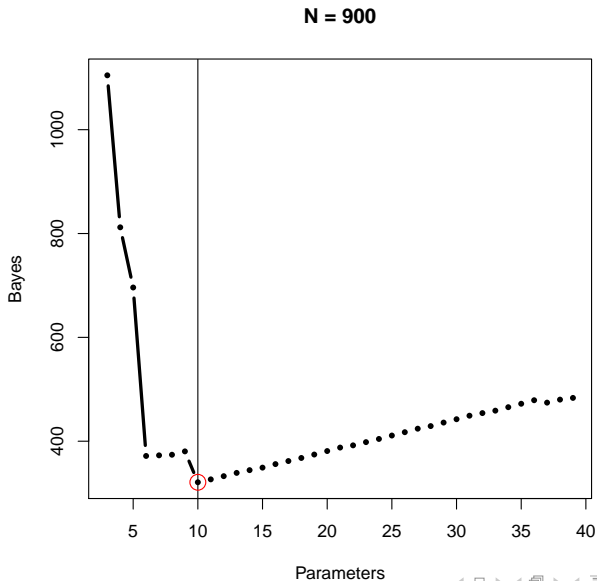
BIC or AIC?



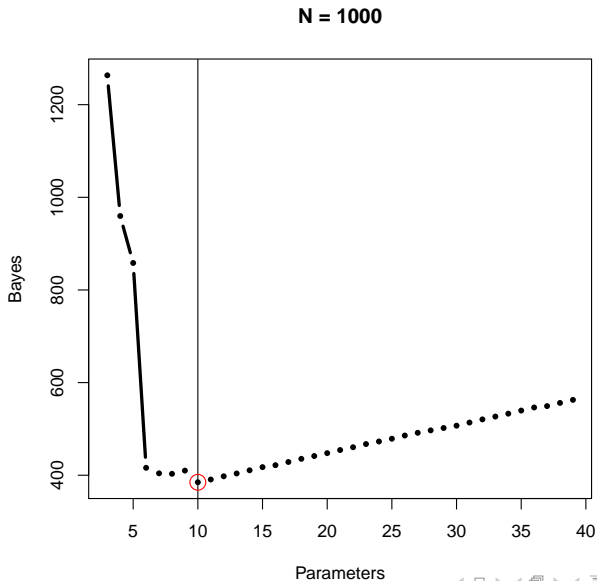
BIC or AIC?



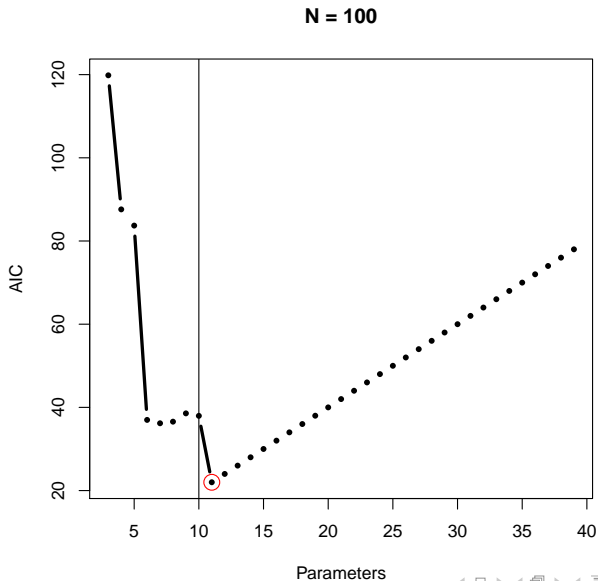
BIC or AIC?



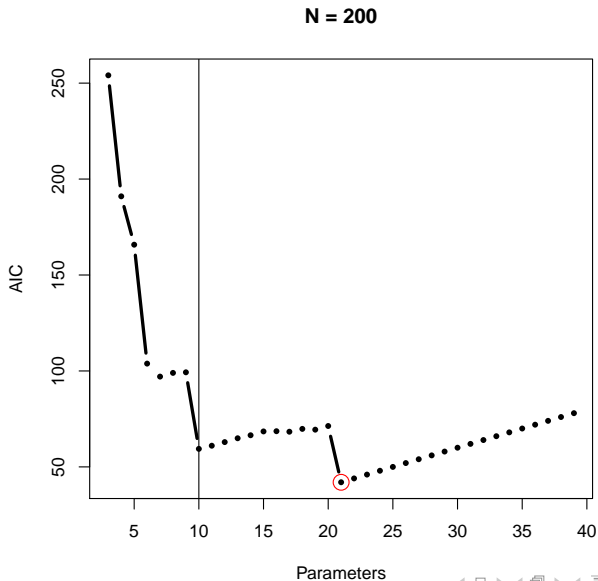
BIC or AIC?



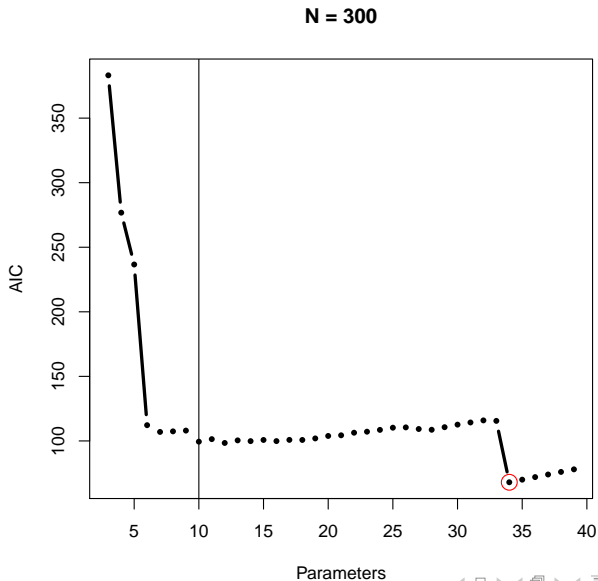
BIC or AIC?



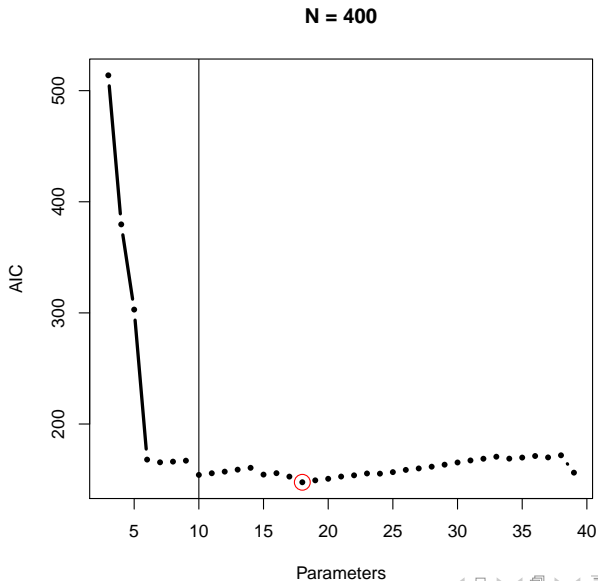
BIC or AIC?



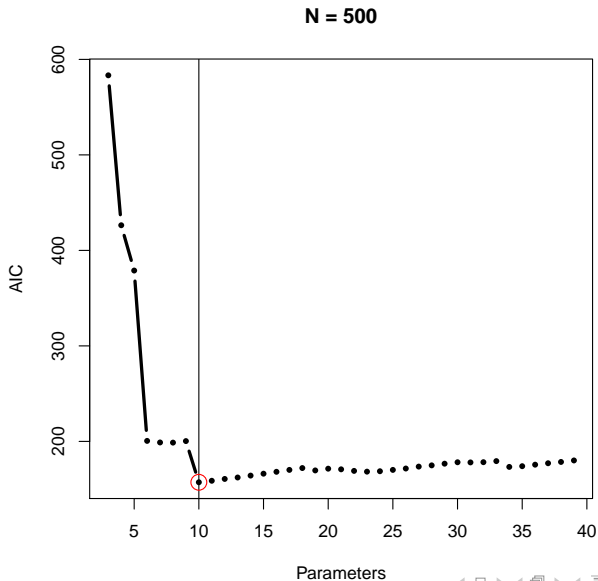
BIC or AIC?



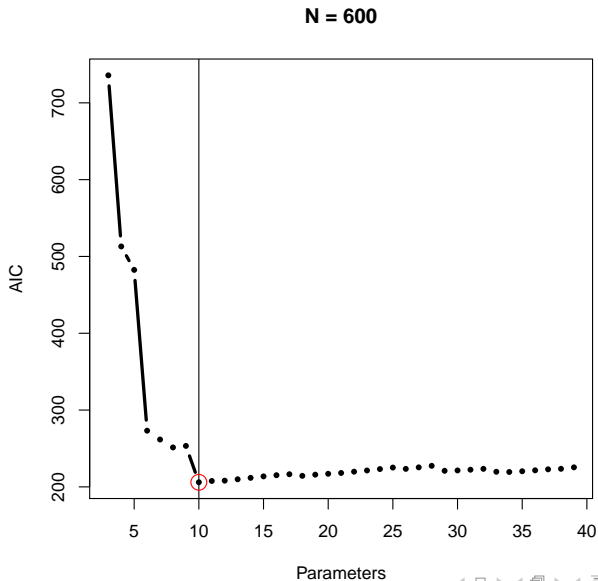
BIC or AIC?



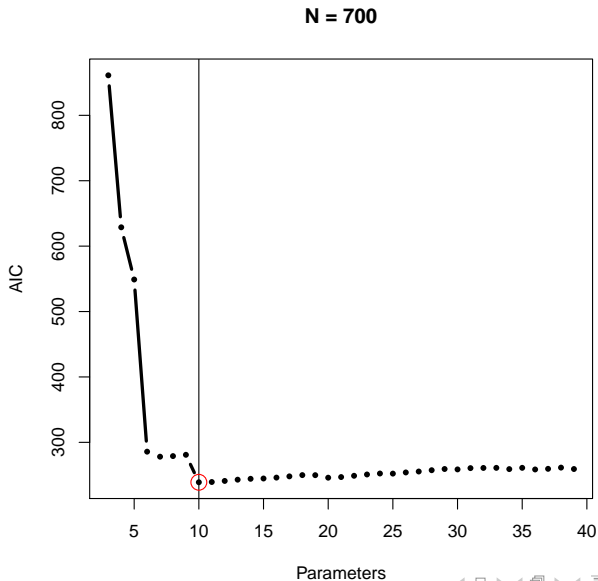
BIC or AIC?



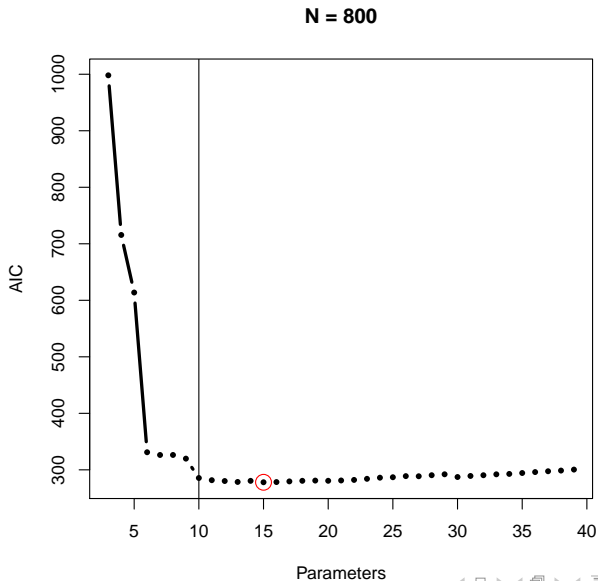
BIC or AIC?



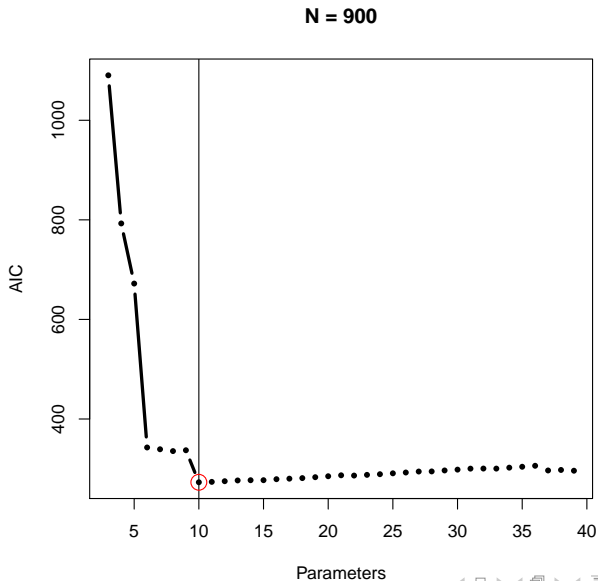
BIC or AIC?



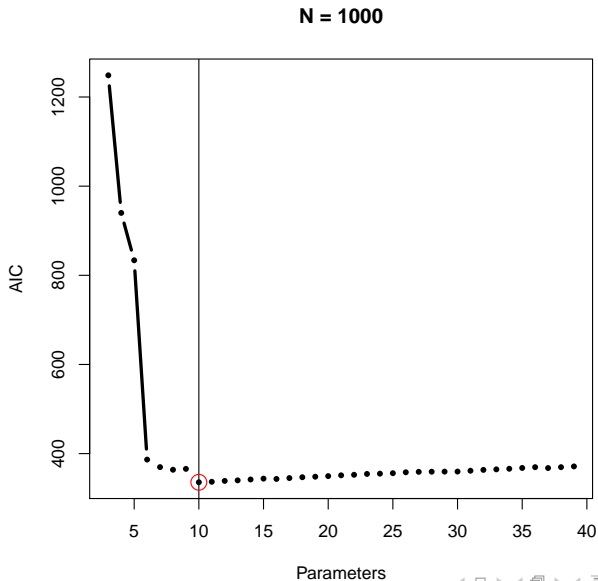
BIC or AIC?



BIC or AIC?



BIC or AIC?



BIC or AIC?

- BIC
 - Asymptotically consistent **if true model is in choice set**
 - As $N \rightarrow \infty$ will choose correct model with probability 1 (if available)
 - Small samples \rightsquigarrow overpenalize
- AIC
 - No asymptotic guarantees \rightsquigarrow derivation doesn't require truth in set. (KL-criteria)
 - In large samples \rightsquigarrow favors complexity
 - Small samples \rightsquigarrow avoids over penalization

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

- Rely on specific loss function

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

- Rely on specific loss function
- Rely on asymptotic argument

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

- Rely on specific loss function
- Rely on asymptotic argument
- Rely on estimate of number of parameters

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

- Rely on specific loss function
- Rely on asymptotic argument
- Rely on estimate of number of parameters
- **Extremely model dependent**

How Do We Select A Model?

Analytic statistics for selection, include penalty for complexity

- AIC : Akaka Information Criterion
- BIC: Bayesian Information Criterion
- DIC: Deviance Information Criterion

Can work well, but...

- Rely on specific loss function
- Rely on asymptotic argument
- Rely on estimate of number of parameters
- **Extremely model dependent**

Need: general tool for evaluating models, **replicates** decision problem

Cross-Validation: Some Intuition

Optimal division of data for prediction:

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model
- Validation: assess model

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model
- Validation: assess model
- Test: predict remaining data

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model
- Validation: assess model
- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model
- Validation: assess model
- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model
- Validation: assess model
- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets
- Each step: use held out data to evaluate performance

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model
- Validation: assess model
- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets
- Each step: use held out data to evaluate performance
- **Avoid overfitting** and have context specific penalty

Cross-Validation: Some Intuition

Optimal division of data for prediction:

- Train: build model
- Validation: assess model
- Test: predict remaining data

K-fold Cross-validation idea: create many training and test sets.

- Idea: use observations both in training and test sets
- Each step: use held out data to evaluate performance
- **Avoid overfitting** and have context specific penalty

Estimates:

$$\text{Error} = E \left[E[L(\mathbf{Y}, f(\hat{\beta}, \mathbf{X})) | \mathcal{T}] \right]$$

Cross-Validation: A How To Guide

Process:

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, . . . , Group K)

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, \dots , Group K)
- Rotate through groups as follows

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, . . . , Group K)
- Rotate through groups as follows

Step Training

Validation (“Test”)

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
⋮	⋮	⋮

Cross-Validation: A How To Guide

Process:

- Randomly partition data into K groups.
(Group 1, Group 2, Group3, ..., Group K)
- Rotate through groups as follows

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
⋮	⋮	⋮
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups
- Train data on $K - 1$ groups. Estimate $\hat{f}^{-K}(\beta, \mathbf{X})$

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups
- Train data on $K - 1$ groups. Estimate $\hat{f}^{-K}(\beta, \mathbf{X})$
- Predict values for K^{th}

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups
- Train data on $K - 1$ groups. Estimate $\hat{f}^{-K}(\beta, \mathbf{X})$
- Predict values for K^{th}
- Summarize performance with loss function: $L(\mathbf{Y}_i, \hat{f}^{-k}(\beta, \mathbf{X}))$

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups
- Train data on $K - 1$ groups. Estimate $\hat{f}^{-K}(\beta, \mathbf{X})$
- Predict values for K^{th}
- Summarize performance with loss function: $L(\mathbf{Y}_i, \hat{f}^{-k}(\beta, \mathbf{X}))$
 - Mean square error, Absolute error, Prediction error, ...

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups
- Train data on $K - 1$ groups. Estimate $\hat{f}^{-K}(\beta, \mathbf{X})$
- Predict values for K^{th}
- Summarize performance with loss function: $L(\mathbf{Y}_i, \hat{f}^{-k}(\beta, \mathbf{X}))$
 - Mean square error, Absolute error, Prediction error, ...

$$\text{CV}(\text{ind. classification}) = \frac{1}{N} \sum_{i=1}^N L(\mathbf{Y}_i, f^{-k}(\beta, \mathbf{X}_i))$$

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups
- Train data on $K - 1$ groups. Estimate $\hat{f}^{-K}(\beta, \mathbf{X})$
- Predict values for K^{th}
- Summarize performance with loss function: $L(\mathbf{Y}_i, \hat{f}^{-k}(\beta, \mathbf{X}))$
 - Mean square error, Absolute error, Prediction error, ...

$$\text{CV(ind. classification)} = \frac{1}{N} \sum_{i=1}^N L(\mathbf{Y}_i, \hat{f}^{-k}(\beta, \mathbf{X}_i))$$

$$\text{CV(proportions)} =$$

$$\frac{1}{K} \sum_{j=1}^K \text{Mean Square Error Proportions from Group } j$$

Cross-Validation: A How To Guide

Step	Training	Validation ("Test")
1	Group2, Group3, Group 4, ..., Group K	Group 1
2	Group 1, Group3, Group 4, ..., Group K	Group 2
3	Group 1, Group 2, Group 4, ..., Group K	Group 3
\vdots	\vdots	\vdots
K	Group 1, Group 2, Group 3, ..., Group K - 1	Group K

Strategy:

- Divide data into K groups
- Train data on $K - 1$ groups. Estimate $\hat{f}^{-K}(\beta, \mathbf{X})$
- Predict values for K^{th}
- Summarize performance with loss function: $L(\mathbf{Y}_i, \hat{f}^{-k}(\beta, \mathbf{X}))$
 - Mean square error, Absolute error, Prediction error, ...

$$\text{CV(ind. classification)} = \frac{1}{N} \sum_{i=1}^N L(\mathbf{Y}_i, \hat{f}^{-k}(\beta, \mathbf{X}_i))$$

$$\text{CV(proportions)} =$$

$$\frac{1}{K} \sum_{j=1}^K \text{Mean Square Error Proportions from Group } j$$

- Final choice: model with highest CV score

How Do We Select K ? (HTF, Section 7.10)

Common values of K

- $K = 5$: Five fold cross validation
- $K = 10$: Ten fold cross validation
- $K = N$: Leave one out cross validation

Considerations:

- How sensitive are inferences to number of coded documents? (HTF, pg 243-244)
- 200 labeled documents
 - $K = N \rightarrow 199$ documents to train,
 - $K = 10 \rightarrow 180$ documents to train
 - $K = 5 \rightarrow 160$ documents to train
- 50 labeled documents
 - $K = N \rightarrow 49$ documents to train,
 - $K = 10 \rightarrow 45$ documents to train
 - $K = 5 \rightarrow 40$ documents to train
- How long will it take to run models?
 - K -fold cross validation requires $K \times$ One model run
- What is the correct loss function?

If you cross validate, you really need to cross validate (Section 7.10.2, ESL)

- Use CV to estimate prediction error
- **All** supervised steps performed in cross-validation
- **Underestimate** prediction error
- **Could lead to selecting lower performing model**

Example from Facebook Data

What do people say to legislators? (Franco, Grimmer, and Lee 2017)

1) Example: estimating classification error

a) Accuracy in legislator posts: 75%

b) Accuracy in public posts: 66.25%

Credit Claiming (Back to Ridge/Lasso, Grimmer, Westwood, and Messing 2014)

```
library(glmnet)
set.seed(8675309) ##setting seed
folds<- sample(1:10, nrow(dtm), replace=T) ##assigning to fold
out_of_samp<- c() ##collecting the predictions
```

Credit Claiming (Back to Ridge/Lasso, Grimmer, Westwood, and Messing 2014)

```
for(z in 1:10){
train<- which(folds!=z) ##the observations we will use to train the model

test<- which(folds==z) ##the observations we will use to test the model
part1<- cv.glmnet(x = dtm[train,], y = credit[train], alpha = 1, family =
binomial) ##fitting the LASSO model on the data.
## alpha = 1 -> LASSO
## alpha = 0 -> RIDGE
## 0<alpha<1 -> Elastic-Net
out_of_samp[test]<- predict(part1, newx= dtm[test,], s = part1$lambda.min,
type =c"class") ##predicting the labels
print(z) ##printing the labels
}

conf_table<- table(out_of_samp, credit) ##calculating the confusion table
> round(sum(diag(conf_table))/len(credit), 3)
[1] 0.844
```

Generalized Cross Validation and Ridge Regression

In some special cases there are analytic solutions:

Generalized Cross Validation and Ridge Regression

In some special cases there are analytic solutions:

$$\beta^{\text{Ridge}} = \left(\mathbf{X}'\mathbf{X} + \lambda I_J \right)^{-1} \mathbf{X}'\mathbf{Y}$$

Generalized Cross Validation and Ridge Regression

In some special cases there are analytic solutions:

$$\begin{aligned}\beta^{\text{Ridge}} &= \left(\mathbf{X}'\mathbf{X} + \lambda I_J\right)^{-1} \mathbf{X}'\mathbf{Y} \\ \hat{\mathbf{Y}} &= \mathbf{X}(\beta)^{\text{Ridge}}\end{aligned}$$

Generalized Cross Validation and Ridge Regression

In some special cases there are analytic solutions:

$$\begin{aligned}\beta^{\text{Ridge}} &= \left(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J \right)^{-1} \mathbf{X}'\mathbf{Y} \\ \hat{\mathbf{Y}} &= \mathbf{X}(\beta)^{\text{Ridge}} \\ &= \underbrace{\mathbf{X} \left(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J \right)^{-1} \mathbf{X}'}_{\text{Hat Matrix}} \mathbf{Y}\end{aligned}$$

Generalized Cross Validation and Ridge Regression

In some special cases there are analytic solutions:

$$\begin{aligned}\beta^{\text{Ridge}} &= \left(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J \right)^{-1} \mathbf{X}'\mathbf{Y} \\ \hat{\mathbf{Y}} &= \mathbf{X}(\beta)^{\text{Ridge}} \\ &= \underbrace{\mathbf{X} \left(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J \right)^{-1} \mathbf{X}'}_{\text{Hat Matrix}} \mathbf{Y} \\ \hat{\mathbf{Y}} &= \underbrace{\mathbf{H}}_{\text{Smoother Matrix}} \mathbf{Y}\end{aligned}$$

Generalized Cross Validation and Ridge Regression

In some special cases there are analytic solutions:

$$\begin{aligned}\beta^{\text{Ridge}} &= \left(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J \right)^{-1} \mathbf{X}'\mathbf{Y} \\ \hat{\mathbf{Y}} &= \mathbf{X}(\beta)^{\text{Ridge}} \\ &= \underbrace{\mathbf{X} \left(\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_J \right)^{-1} \mathbf{X}'}_{\text{Hat Matrix}} \mathbf{Y} \\ \hat{\mathbf{Y}} &= \underbrace{\mathbf{H}}_{\text{Smoother Matrix}} \mathbf{Y}\end{aligned}$$

Generalized Cross Validation and Ridge Regression

Why do we care?

Generalized Cross Validation and Ridge Regression

Why do we care?

Leave one out cross validation

Generalized Cross Validation and Ridge Regression

Why do we care?

Leave one out cross validation

$$\text{Cross Validation}(1) = \frac{1}{N} \sum_{i=1}^N (Y_i - f(\mathbf{X}_{-i}, \mathbf{Y}_{-i}, \lambda, \hat{\beta}))^2$$

Generalized Cross Validation and Ridge Regression

Why do we care?

Leave one out cross validation

$$\begin{aligned}\text{Cross Validation(1)} &= \frac{1}{N} \sum_{i=1}^N (Y_i - f(\mathbf{X}_{-i}, \mathbf{Y}_{-i}, \lambda, \hat{\beta}))^2 \\ &= \frac{1}{N} \sum_{i=1}^N \left(\frac{Y_i - f(\mathbf{X}, \mathbf{Y}, \lambda, \hat{\beta})}{1 - H_{ii}} \right)^2\end{aligned}$$

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H}) = \text{Effective number of parameters (class regression = number of independent variables + 1)}$

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H}) = \text{Effective number of parameters (class regression = number of independent variables + 1)}$
- For Ridge regression:

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H}) = \text{Effective number of parameters (class regression = number of independent variables + 1)}$
- For Ridge regression:

$$\text{Tr}(\mathbf{H}) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\text{Penalty}}}$$

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H}) = \text{Effective number of parameters (class regression = number of independent variables + 1)}$
- For Ridge regression:

$$\text{Tr}(\mathbf{H}) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\text{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}'\mathbf{X}$

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H}) = \text{Effective number of parameters (class regression = number of independent variables + 1)}$
- For Ridge regression:

$$\text{Tr}(\mathbf{H}) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\text{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}'\mathbf{X}$ (!!!!!)

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H}) = \text{Effective number of parameters (class regression = number of independent variables + 1)}$
- For Ridge regression:

$$\text{Tr}(\mathbf{H}) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\text{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}'\mathbf{X}$ (!!!!!)

Define generalized cross validation:

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H}) = \text{Effective number of parameters (class regression = number of independent variables + 1)}$
- For Ridge regression:

$$\text{Tr}(\mathbf{H}) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\text{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}'\mathbf{X}$ (!!!!!)

Define generalized cross validation:

$$\text{GCV} = \frac{1}{N} \sum_{i=1}^N \left(\frac{Y_i - \hat{Y}_i}{1 - \frac{\text{Tr}(\mathbf{H})}{N}} \right)^2$$

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H})$ = Effective number of parameters (class regression = number of independent variables + 1)
- For Ridge regression:

$$\text{Tr}(\mathbf{H}) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\text{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}'\mathbf{X}$ (!!!!!)

Define generalized cross validation:

$$\text{GCV} = \frac{1}{N} \sum_{i=1}^N \left(\frac{Y_i - \hat{Y}_i}{1 - \frac{\text{Tr}(\mathbf{H})}{N}} \right)^2$$

Applicable in any setting where we can write **Smoother** matrix

Generalized Cross Validation and Ridge Regression

Calculating \mathbf{H} can be computationally expensive

- $\text{Trace}(\mathbf{H}) \equiv \text{Tr}(\mathbf{H}) = \sum_{i=1}^N H_{ii}$
- $\text{Tr}(\mathbf{H})$ = Effective number of parameters (class regression = number of independent variables + 1)
- For Ridge regression:

$$\text{Tr}(\mathbf{H}) = \sum_{j=1}^J \frac{\lambda_j}{\lambda_j + \underbrace{\lambda}_{\text{Penalty}}}$$

where λ_j is the j^{th} Eigenvalue from $\mathbf{\Sigma} = \mathbf{X}'\mathbf{X}$ (!!!!!)

Define generalized cross validation:

$$\text{GCV} = \frac{1}{N} \sum_{i=1}^N \left(\frac{Y_i - \hat{Y}_i}{1 - \frac{\text{Tr}(\mathbf{H})}{N}} \right)^2$$

Applicable in any setting where we can write **Smoother** matrix