## Exercise L

(a) We want to find the 
$$\hat{\theta}$$
 that minimizes

For this purpose, we will equate the derivative  $\frac{\partial L(\theta)}{\partial \theta}$  to 0.

$$\frac{\partial L(\theta)}{\partial \theta} = 0 \Rightarrow 2\frac{\tilde{z}}{n} (y_n - \hat{\theta} \tilde{x}_n) (-x_n) + 2\lambda \hat{\theta} = 0 \Rightarrow$$

=> 
$$-\frac{2}{x_{n-1}}y_nx_n + \frac{2}{x_n}(\hat{\theta}_{x_n})x_n + \hat{\theta}_{x_n} = 0 = 0$$

$$= \sum_{n=1}^{\infty} (x_n x_n^{\intercal}) \hat{\theta} + \lambda \hat{\theta} = \sum_{n=1}^{\infty} y_n x_n = 0$$

$$= 3\left(\frac{3}{2}\left(x_{n}x_{n}^{*}\right) + \lambda I\right)\hat{\theta} = \frac{3}{2}y_{n}x_{n} \quad (A)$$

## (6) We have shown that (Homework 1):

$$\frac{1}{2} \sum_{n=1}^{\infty} x_n x_n^{\top} = X^{\top} X \qquad \text{where } X = \begin{bmatrix} x_1^{\top} \\ y_2^{\top} \\ \vdots \\ x_n^{\top} \end{bmatrix}$$

Thus, 
$$(A) \Rightarrow (XX + \lambda I) \hat{\theta} = X_y \Rightarrow \hat{\theta} = (XX + \lambda I)^T X_y$$

$$(*) \times \times \times = [\times, \times_2 ... \times_N] \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_N \end{bmatrix} = \underbrace{\sum_{n=1}^N x_n x_n}_{n=1}$$

$$\times^7 y = [\times, \times_2 ... \times_N] \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{bmatrix} = \underbrace{\sum_{n=1}^N y_n x_n}_{n=1} = \underbrace{\sum_{n=1}^N y_n x_n}_{n=$$

## Exercise 2

(a) Since  $\hat{\theta}_{\text{min}}$  is an unbiased estimator of  $\hat{\theta}_{0}$ , we know that  $E[\hat{\theta}_{\text{min}}] = \hat{\theta}_{0}$ Therefore,  $MSE = E[(\hat{\theta} - E(\hat{\theta}))] + (E(\hat{\theta}) - \theta_{0})^{2} = \hat{\theta}_{\text{min}}$ 

=> 
$$MSE = E[(\hat{\theta} - E(\hat{\theta})^2]$$

Bias is equal to zero and MSE is equal to the variance of the estimator Barra.

(b) 
$$\theta_{B} = (1+\alpha)\theta_{MVU} \Rightarrow f[\theta_{B}] = f[(1+\alpha)\theta_{MVU}] \Rightarrow f[\theta_{B}] = f[(1+\alpha)\theta_{MVU}] \Rightarrow f[\theta_{B}] = f[\theta_{MVU}] \Rightarrow f[\theta_{B}] = f[\theta_{B}] =$$

Since a to => f[0] \$ 00, and therefore Do is a biased estimator of 0.

(c) 
$$MSE = \{ [(\hat{G} - F(\hat{G}))] \ (question (a)) \}$$

In order for MSE to be 0, the variance of  $\hat{\theta}$  needs to be 0.

This cannot be achieved for finite N (data sets), because every data set will have an estimator that will be different from the others.

MSE = 
$$E[(\theta_b - E(\theta_b))^2] + (E(\theta_b) - \theta_b)^2 =$$

$$MSE = \left[ \left( (1+a) \theta - \left[ \left( (1+a) \theta \right)^{2} \right] + \left( \left[ \left( (1+a) \theta \right) - \theta_{0} \right]^{2} \right] \right]$$

$$= \mathbb{E}\left[\left((1+\alpha)\theta - (1+\alpha)\mathbb{E}[\theta]\right)^{2}\right] + \left(\mathbb{E}[\theta + \alpha\theta] - \theta_{\delta}\right)^{2} =$$

$$= E[(1+a)^{2}(\theta - E(\theta))^{2}] + (E[\theta_{mvu}] + a E[\theta] - \theta_{0})^{2}$$

$$= (1+a)^{2} E[(\theta - E(\theta))^{2}] + (a E(\theta))^{2}$$

$$= (1+a)^{2} MSE + (a \theta_{0})^{2} = 1$$

$$\theta_{mvu}$$

=> 
$$MSE = (1+a)^2 MSE + (a \theta o)^2$$

(e) We search the values of a for which:

NSE < MSE =>

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=> (1+a)2 MSE + (a (00))2 < MSE =>

-> MSE + 2a MSE + a<sup>2</sup> MSE - MSE + a<sup>2</sup> fo<sup>2</sup> LO =s

=> a2 (MSE + 402) + 2a MSE (0

The above is equal to 0 for a = 0 or  $a = \frac{-2MSE_{\theta mvn}}{MSE} + \theta_0^2$ Therefore it is less than 0 for  $a \in (-2\frac{MSE_{\theta mvn}}{MSE} + \theta_0^2)$ , 0)

 $(f) \frac{-2MSE_{mvu}}{MSE} < a < 0 \Rightarrow 1 - \frac{2MSE_{mvu}}{MSE} < a+1 < 1 \Rightarrow 0$   $\theta_{mvu} = \theta_0^2$   $\theta_{mvu} = \theta_0^2$ 

=> |a+1| < 1 and Since  $\theta_b = (1+\alpha) \theta_{mvn}$  we have that  $|\theta_b| < |\theta_{mvn}|$ 

$$\frac{\partial MSE}{\partial a} = 0 \Rightarrow 2(1+a)MSE + 2abo^2 = 0 \Rightarrow$$

$$MSE + a MSE + a \theta_0^2 = 0 = 3$$
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$$MSE + a (MSE + \theta_0^2) = 0 = 3$$

$$= \frac{1}{2} = \frac{-ME_{mvu}}{ME + \theta_0^2}$$

(h) Because we do not know the value of to. If we knew the real value of to we would not need to try finding estimators.

Exercise 3

(a) 
$$(\frac{3}{2}x_nx_n^T)\theta = \frac{3}{2}y_nx_n \Rightarrow N\hat{\theta}_{is} = \frac{3}{2}y_n \Rightarrow \hat{\theta}_{is} = \frac{1}{N}\sum_{n=1}^{\infty}y_n$$

(c) 
$$\bar{y} = \frac{1}{N} \sum_{n=1}^{\infty} y_n = 1$$
  $f(\bar{y}) = \frac{1}{N} \cdot N \cdot R = 1$ 

(e) 
$$(\widetilde{Z}(x_nx_n^{-1}) + \lambda I)\widehat{\theta} = \widetilde{Z}_{n=1} y_n x_n$$
  
But  $x_n=1$ , thus:  
 $(N+\lambda)\widehat{\theta}_{ridge} = \widetilde{Z}_{n=1} y_n = \widehat{\theta}_{ridge} = \widetilde{X}_{n+\lambda} y_n$ 

(f) We know that 
$$\hat{\theta}_{LS} = \frac{1}{N} \sum_{n=1}^{\infty} y_n = N \hat{\theta}_{LS}$$
 or  $\sum_{n=1}^{\infty} y_n = N \hat{\theta}_{LS}$ 

(g) 
$$F[\hat{\theta}] = f\left[\frac{N}{N+2} \hat{\theta} + \frac{\hat{\theta}}{N+2} +$$

=> 
$$f[\hat{\theta}] = \frac{N}{N+2}\theta_0$$
 For  $\hat{\theta}$  to be unbiased estimator it needs to hold that  $f[\hat{\theta}] = \theta_0 \Rightarrow \frac{N}{N+2} = 1 \Rightarrow \lambda = 0$   
But  $\lambda \neq 0$ , thuy  $\hat{\theta}$  is biased estimator of  $\theta_0$ 

(h) 
$$|\hat{\theta}| = \left|\frac{N}{N+2}\right| \left|\hat{\theta}\right|$$
 For  $|\hat{\theta}|$  to be smaller than  $|\hat{\theta}|$  it needs to hold that  $\left|\frac{N}{N+2}\right| < 1$ . Indeed for  $2 > 0$   $\left|\frac{N}{N+2}\right| < 1$ . Thus,  $|\hat{\theta}| < \left|\hat{\theta}\right|$ 

(i) 
$$\hat{\partial}_{b} = (1+\alpha)\hat{\partial}_{mvn} \Rightarrow 1+\alpha = \frac{N}{N+\lambda} \Rightarrow \alpha = \frac{-\lambda}{N+\lambda}$$

$$a > \frac{-2MSE(\theta_{MVN})}{MSE(\theta_{MVN})} = > \frac{\lambda}{2} > \frac{\lambda}{2} > \frac{-2MSE(\theta_{MVN})}{MSE(\theta_{MVN})} = > \frac{\lambda}{2} > \frac{\lambda}{2$$

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=> 
$$\lambda \angle \frac{2MSE(\theta_{MVu})}{\theta_0^2 - MSE(\theta_{MVu})}$$
, also  $\lambda > 0$