# Brownian Motion and Stochastic Integration (detailed proofs, split frames)

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#### Outline

- Foundations
- 2 Brownian motion
- 3 Kolmogorov continuity theorem
- Wiener integral for deterministic integrands
- 5 Itô integral for adapted integrands
- 6 Quadratic variation of Brownian motion
- Itô formula (careful sketch)
- Selected exercises (with hints)

## Probability space

- $(\Omega, \mathcal{F}, \mathbb{P})$  is a probability space.
- Filtration  $\{\mathcal{F}_t\}_{t\geq 0}$ : increasing family of sub- $\sigma$ -algebras.

#### Stochastic process

- A stochastic process is a measurable map  $X : [0, \infty) \times \Omega \to \mathbb{R}$ .
- Adapted:  $X_t$  is  $\mathcal{F}_t$ -measurable for each t.

# Definition (standard Brownian motion)

#### **Definition**

A process  $(B_t)_{t\geq 0}$  is a standard Brownian motion if:

- **1**  $B_0 = 0$  a.s.
- ② Independent Gaussian increments: for  $0 \le s < t$ ,  $B_t B_s \sim N(0, t s)$ .
- Continuous paths a.s.

# Immediate consequences (1)

- For each fixed t,  $B_t \sim N(0, t)$ .
- Finite-dimensional distributions are multivariate normal.

# Immediate consequences (2)

Covariance identity:

$$\mathbb{E}[B_sB_t]=\min\{s,t\}.$$

Next frames give the short proof and the joint Gaussian argument.

# Proof: covariance (step 1)

Let  $0 \le s \le t$ . Write

$$B_t = (B_t - B_s) + B_s.$$

The increment  $B_t - B_s$  is independent of  $B_s$  and has mean zero.

# Proof: covariance (step 2)

Therefore

$$\mathbb{E}[B_sB_t] = \mathbb{E}\big[B_s(B_t - B_s)\big] + \mathbb{E}[B_s^2] = 0 + \operatorname{Var}(B_s) = s.$$

So  $\mathbb{E}[B_sB_t] = \min\{s,t\}$  as claimed.

## Proof: joint Gaussianity (idea)

Write the vector  $(B_{t_1}, \ldots, B_{t_n})$  as a linear transform of independent Gaussian increments:

$$B_{t_k} = \sum_{j=1}^k (B_{t_j} - B_{t_{j-1}}), \quad t_0 = 0.$$

A linear transform of a Gaussian vector is Gaussian. Hence finite-dimensional distributions are Gaussian.

# Kolmogorov continuity — statement (part 1)

#### Theorem

Let  $(X_t)_{t\in[0,T]}$  be a real-valued process. If there exist constants C>0,  $p\geq 1,\ \delta>0$  such that

$$\mathbb{E}\big[|X_t - X_s|^p\big] \le C|t - s|^{1 + \delta}$$

for all  $s, t \in [0, T]$ , then...

# Kolmogorov continuity — statement (part 2)

...there exists a continuous modification X whose paths are Hölder continuous of any exponent  $\gamma \in (0, \delta/p)$ .

#### Proof: overview

#### Strategy:

- Use dyadic partitions of [0, T].
- Control probabilities of large dyadic increments.
- Apply Borel-Cantelli to get almost sure control.
- Extend from dyadic points to all times.

#### Proof: dyadic partition and notation

Fix integer  $m \ge 1$  and define dyadic points

$$t_k^{(m)} = k2^{-m}T, \qquad k = 0, \dots, 2^m.$$

Define maximal increment at level *m*:

$$M_m := \max_{0 \le k < 2^m} |X_{t_{k+1}^{(m)}} - X_{t_k^{(m)}}|.$$

#### Proof: tail probability estimate

For a single dyadic interval,

$$\mathbb{P}\{|X_{t_{k+1}}-X_{t_k}|>\varepsilon\}\leq \frac{\mathbb{E}|X_{t_{k+1}}-X_{t_k}|^p}{\varepsilon^p}\leq \frac{C(2^{-m}T)^{1+\delta}}{\varepsilon^p}.$$

(Markov/Chebyshev inequality.)

## Proof: union bound and choice of epsilon

Union bound over  $2^m$  intervals:

$$\mathbb{P}\{M_m > \varepsilon\} \leq 2^m \cdot \frac{C(2^{-m}T)^{1+\delta}}{\varepsilon^p} = CT^{1+\delta}\varepsilon^{-p}2^{-m\delta}.$$

Choose  $\varepsilon_m := 2^{-m\gamma}$  with  $0 < \gamma < \delta/p$ .

# Proof: summability and Borel-Cantelli

With 
$$\varepsilon_m = 2^{-m\gamma}$$
,

$$\mathbb{P}\{M_m > \varepsilon_m\} \leq CT^{1+\delta}2^{-m(\delta-\gamma p)}.$$

Since  $\delta - \gamma p > 0$ ,  $\sum_m \mathbb{P}\{M_m > \varepsilon_m\} < \infty$ . Borel–Cantelli implies  $M_m \leq \varepsilon_m$  eventually a.s.

## Proof: constructing Hölder bound on dyadics

For dyadic points s < t at level m with m large,

$$|X_t - X_s| \leq (\#intervals) \varepsilon_m.$$

As  $(\#intervals) \cdot 2^{-m}T = t - s$ , we get

$$|X_t - X_s| \le \frac{t-s}{2^{-m}T}\varepsilon_m.$$

Selecting m such that  $2^{-(m+1)}T < t-s \le 2^{-m}T$  yields a Hölder estimate.

## Proof: extension to all times and continuity

The Hölder estimate holds for dyadic rationals. Using density of dyadics and completeness of reals, extend the bound to all  $s, t \in [0, T]$ , producing a continuous modification with the claimed exponent.

#### Motivation

Brownian sample paths almost surely have infinite variation, so pathwise Riemann–Stieltjes integrals do not generally make sense. The Wiener integral constructs  $\int f(t) \, dB_t$  for deterministic  $f \in L^2$  as an  $L^2(\Omega)$  limit.

#### Step functions: definition

Let f be a step function on [a, b]:

$$f(t) = \sum_{j=1}^{n} c_j \mathbf{1}_{[t_{j-1},t_j)}(t).$$

Define the stochastic Riemann sum

$$I(f) := \sum_{j=1}^{n} c_j (B_{t_j} - B_{t_{j-1}}).$$

## Step functions: Gaussianity and variance

Each increment  $B_{t_j} - B_{t_{j-1}}$  is Gaussian and independent across j. Thus I(f) is Gaussian, mean zero, and

$$Var(I(f)) = \sum_{j=1}^{n} c_j^2(t_j - t_{j-1}).$$

(Cross-terms vanish by independence.)

#### Isometry for step functions

For step functions f and g (use a common refinement),

$$\mathbb{E}[I(f)I(g)] = \int_a^b f(t)g(t) dt.$$

Hence I preserves the  $L^2$  inner product on step functions.

# Extension to $L^2[a, b]$

Step functions are dense in  $L^2[a,b]$ . If  $f_n \to f$  in  $L^2$ , then  $I(f_n)$  is Cauchy in  $L^2(\Omega)$ :

$$\mathbb{E}|I(f_n)-I(f_m)|^2 = ||f_n-f_m||_{L^2}^2.$$

Define  $\int_a^b f \, dB$  as the  $L^2(\Omega)$  limit of  $I(f_n)$ .

# Properties of the Wiener integral

- $\int_a^b f \ dB$  is Gaussian with mean 0 and variance  $\|f\|_{L^2}^2$ .
- Isometry:  $\mathbb{E}\left[\left(\int_a^b f \, dB\right)^2\right] = \|f\|_{L^2}^2$ .
- Inner product preservation holds for all  $f, g \in L^2$ .

# Compatibility with pathwise integral

If f is continuous and of bounded variation, then for almost every Brownian sample path the pathwise Riemann–Stieltjes integral  $\int f \, dB$  exists and coincides a.s. with the Wiener integral (approximate f by step functions).

#### Elementary adapted processes

Let  $\{\mathcal{F}_t\}$  be a filtration. An elementary adapted (predictable simple) process is

$$\xi_t(\omega) = \sum_{k=1}^n H_{k-1}(\omega) \mathbf{1}_{(t_{k-1},t_k]}(t),$$

where  $H_{k-1}$  is  $\mathcal{F}_{t_{k-1}}$ -measurable and square-integrable.

# Definition of the Itô integral (elementary)

For such  $\xi$ , define

$$\int_0^T \xi_t \, \mathrm{d} B_t := \sum_{k=1}^n H_{k-1} \big( B_{t_k} - B_{t_{k-1}} \big).$$

This is a random variable in  $L^2(\Omega)$ .

#### Itô isometry (elementary) — statement

For elementary  $\xi$  as above,

$$\mathbb{E}\Big[\Big(\int_0^T \xi_t \, \mathrm{d}B_t\Big)^2\Big] = \mathbb{E}\Big[\int_0^T \xi_t^2 \, \mathrm{d}t\Big].$$

Proof split over the next few frames.

# Itô isometry — proof (part 1)

Expand the square:

$$\left(\sum_{k=1}^{n} H_{k-1} \Delta B_{k}\right)^{2} = \sum_{k=1}^{n} H_{k-1}^{2} (\Delta B_{k})^{2} + 2 \sum_{i < j} H_{i-1} H_{j-1} \Delta B_{i} \Delta B_{j},$$

where 
$$\Delta B_k := B_{t_k} - B_{t_{k-1}}$$
.

# ltô isometry — proof (part 2)

Take expectation. For i < j,  $\Delta B_j$  is independent of  $\mathcal{F}_{t_{j-1}}$ , hence independent of  $H_{i-1}$  and of  $\Delta B_i$ . Moreover  $\mathbb{E}[\Delta B_j] = 0$ . Thus each cross-term has expectation zero.

# ltô isometry — proof (part 3)

For diagonal terms,

$$\mathbb{E}[H_{k-1}^2(\Delta B_k)^2] = \mathbb{E}[H_{k-1}^2] \, \mathbb{E}[(\Delta B_k)^2] = \mathbb{E}[H_{k-1}^2](t_k - t_{k-1}).$$

Summing yields the desired equality.

# Extension to predictable $L^2$ integrands

Let  $\mathcal{H}$  be the closure (under  $\mathbb{E} \int_0^T \xi_t^2 dt$ ) of elementary adapted processes. By the isometry, the map  $\xi \mapsto \int_0^T \xi_t dB_t$  extends uniquely to an isometry on  $\mathcal{H}$ .

# Martingale property (elementary case)

For elementary  $\xi$ , for s < t the increment  $M_t - M_s$  is a sum of terms each independent of  $\mathcal{F}_s$  and mean zero. So  $\mathbb{E}[M_t \mid \mathcal{F}_s] = M_s$ . Thus M is a martingale.

# Martingale property (general case)

For  $\xi \in \mathcal{H}$  pick elementary  $\xi^{(n)} \to \xi$ . The integrals converge in  $L^2$  uniformly in t; conditional expectation is continuous in  $L^2$ , so the martingale property passes to the limit. Hence  $M_t = \int_0^t \xi_s \, dB_s$  is a square-integrable martingale.

#### Quadratic variation — definition

For partition  $\Pi = \{0 = t_0 < t_1 < \cdots < t_n = t\}$  define

$$Q(\Pi) := \sum_{k=0}^{n-1} (B_{t_{k+1}} - B_{t_k})^2.$$

We study the limit as  $\|\Pi\| \to 0$ .

### Q.V. for Brownian motion — statement

For Brownian motion,

$$\lim_{\|\Pi\|\to 0} Q(\Pi) = t \quad \text{in } L^2(\Omega),$$

hence in probability. So  $[B]_t = t$ .

## Proof — expectation

Compute expectation:

$$\mathbb{E}[Q(\Pi)] = \sum_{k} \mathbb{E}[(\Delta B_k)^2] = \sum_{k} (t_{k+1} - t_k) = t.$$

Thus 
$$\mathbb{E}[Q(\Pi) - t] = 0$$
.

# Proof — second moment expansion

Expand

$$Q(\Pi)^2 = \sum_k (\Delta B_k)^4 + 2 \sum_{i < j} (\Delta B_i)^2 (\Delta B_j)^2.$$

Take expectation and use independence for cross terms.

### Proof — use Gaussian fourth moment

For 
$$Y \sim N(0, \sigma^2)$$
 we have  $\mathbb{E}[Y^4] = 3\sigma^4$ . Thus  $\mathbb{E}[(\Delta B_k)^4] = 3(t_{k+1} - t_k)^2$ . Combine to get 
$$\mathbb{E}[Q(\Pi)^2] = t^2 + 2\sum_k (t_{k+1} - t_k)^2.$$

#### Proof — variance tends to zero

Therefore

$$\operatorname{Var}(Q(\Pi)) = \mathbb{E}[Q(\Pi)^2] - t^2 = 2\sum_k (t_{k+1} - t_k)^2 \le 2t \|\Pi\| \to 0.$$

Hence  $Q(\Pi) \to t$  in  $L^2$  as desired.

# ltô formula — setup

Suppose

$$X_t = X_0 + \int_0^t \mu_s \, ds + \int_0^t \sigma_s \, dB_s,$$

with adapted  $\mu, \sigma$  satisfying integrability assumptions. Let  $f \in C^2(\mathbb{R})$ .

## Itô formula — discrete Taylor

On a partition  $0 = t_0 < \cdots < t_n = t$  apply Taylor:

$$f(X_{t_{k+1}}) - f(X_{t_k}) = f'(X_{t_k}) \Delta X_k + \frac{1}{2} f''(X_{t_k}) (\Delta X_k)^2 + r_k,$$

where  $\Delta X_k = X_{t_{k+1}} - X_{t_k}$  and  $r_k$  is the remainder.

### Itô formula — analyze terms: first term

Sum of first terms approximates

$$\sum_{k} f'(X_{t_k}) \Delta X_k \to \int_0^t f'(X_s) \, dX_s$$

which splits into Lebesgue and Itô parts corresponding to  $\mu$  and  $\sigma$  dB.

# Itô formula — analyze terms: quadratic term

Note  $(\Delta X_k)^2$  contains  $(\Delta M_k)^2$  where M is the martingale part. Using quadratic variation,  $\sum (\Delta M_k)^2 \to [M]_t = \int_0^t \sigma_s^2 ds$ . Hence quadratic term yields  $\frac{1}{2} \int_0^t f''(X_s) \sigma_s^2 ds$ .

### Itô formula — remainder control

Remainder  $r_k$  is of higher order: typically  $O((\Delta t)^{3/2})$  because  $\Delta M_k = O(\sqrt{\Delta t})$ . Sum of  $r_k$  tends to zero in probability (or in  $L^1$ ) under standard integrability assumptions, by direct  $L^2$  estimates.

### Itô formula — final form

Putting pieces together:

$$f(X_t) = f(X_0) + \int_0^t f'(X_s) \mu_s \, ds + \int_0^t f'(X_s) \sigma_s \, dB_s + \frac{1}{2} \int_0^t f''(X_s) \sigma_s^2 \, ds.$$

This is the Itô formula in one dimension.

## Exercise A

Show 
$$\mathbb{E}|B_t - B_s|^4 = 3|t - s|^2$$
.

### Exercise A — hint

Use: if  $Z \sim N(0, \sigma^2)$  then  $\mathbb{E}Z^4 = 3\sigma^4$ . Apply with  $\sigma^2 = |t - s|$ .

### Exercise B

Compute the distribution of  $Y = \int_0^1 B_t dt$ .

### Exercise B — hint

Integrate by parts (or use kernel representation):

$$\int_0^1 B_t \, dt = \int_0^1 (1-t) \, dB_t,$$

so variance is  $\int_{0}^{1} (1-t)^{2} dt = 1/3$ .