Efficient mode jumping MCMC for Bayesian variable selection in GLMM

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Introduction. Issues

- GLMM are addressed for inference and prediction in a wide range of different applications providing a powerful scientific tool for the researchers and analysts from different fields
- More and more sources of data are becoming available introducing a variety of hypothetical explanatory variables for these models to be considered
- Selection of an optimal combination of these variables is crucial.
 Posterior model probabilities is one of the relevant measures to estimate quality of the models
- The number of models to select from is exponential in the number of candidate variables
- The search space in this context has numerous local extrema (potentially sparsely located)
- Hence efficient search algorithms have to be adopted for evaluating the posterior distribution within a reasonable amount of time

Bayesian Generalized Linear Mixed Model

$$Y_t|\mu_t \sim f(y|\mu_t), t \in \{1, ..., T\}$$
 (1)

$$\mu_t = g^{-1} \left(\eta_t \right) \tag{2}$$

$$\eta_t = \gamma_0 \beta_0 + \sum_{i=1}^p \gamma_i \beta_i X_{ti} + \delta_t \tag{3}$$

$$\boldsymbol{\delta} = (\delta_1, ..., \delta_T) \sim N_T (\mathbf{0}, \boldsymbol{\Sigma}_b). \tag{4}$$

- $\beta_i \in \mathbb{R}, i \in \{0,...,p\}$ are regression coefficients
- $\Sigma_b = \Sigma_b(\psi) \in \mathbb{R}^T \times \mathbb{R}^T$ is the covariance of the random effect
- ullet δ_t is the correlation structure between them through random effects
- $g(\cdot)$ is a proper link function
- $\gamma_i \in \{0,1\}, i \in \{0,...,p\}$ are latent indicators defining if covariate X_{ti} is included into the model $(\gamma_i = 1)$ or not $(\gamma_i = 0)$



We use a fully Bayesian approach, hence specify priors

$$\gamma_i \sim Binom(1,q)$$
 (5)

$$q \sim Beta(\alpha_q, \beta_q)$$
 (6)

$$\boldsymbol{\beta}|\boldsymbol{\gamma} \sim N_{\sum_{i=1}^{p} \gamma_{i}}(\boldsymbol{\mu}_{\beta}, \boldsymbol{\Sigma}_{\beta}) \tag{7}$$

$$\psi \sim \varphi(\psi), \tag{8}$$

- q is the prior probability of including a covariate into the model
- α_q, β_q are hyper parameters for the prior on q
- ullet $\mu_eta, oldsymbol{\Sigma}_eta$ are hyper parameters for the prior on $oldsymbol{eta}|\gamma$
- ullet ψ are the hyper parameters of the random effect



Inference on the model

Let:

 $\theta = \{\vec{\beta}, \rho, \sigma_{\epsilon}^2\}$ define parameters of the model and $\gamma : \vec{\gamma}$ define a model itself, i.e. which covariates are addressed.

Then:

- ullet $\theta | \gamma$ define parameters conditioned on fixed models
- $\exists 2^{p+1}$ different models

Goals:

- ullet $p(\gamma, heta|\mathbb{D})$ posterior distribution of parameters and models
- ullet $p(\gamma|\mathbb{D})$ marginal posterior distribution of the models
- Set of estimated models performing well in terms of some model selection criteria (MAP, WAIC, DIC, MLIK)

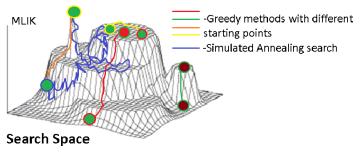
Procedure

- Note that $p(\gamma, \theta|\mathbb{D}) = p(\theta|\gamma, \mathbb{D})p(\gamma|\mathbb{D})$
- ullet $p(heta|\gamma,\mathbb{D})$ and $\log p(\mathbb{D}|\gamma)$ can be efficiently obtained by INLA
- Note that $p(\gamma|\mathbb{D}) = \frac{e^{\log p(\mathbb{D}|\gamma) + \log p(\gamma)}}{\sum_{\gamma' \in \Omega_{\gamma}} e^{\log p(\mathbb{D}|\gamma') + \log p(\gamma')}}$
- $\widehat{p}(\gamma|\mathbb{D}) = \frac{e^{\log p(\mathbb{D}|\gamma) + \log p(\gamma)}}{\sum_{\gamma' \in \mathbb{V}} e^{\log p(\mathbb{D}|\gamma') + \log p(\gamma')}}$
- ullet V is the subspace of Ω_{γ} to be efficiently explored
- Note that for $p(\gamma) = p(\gamma') \forall \gamma, \gamma' \in \Omega_{\gamma}$:
- $p(\gamma|\mathbb{D}) \gg p(\gamma'|\mathbb{D})$ if $\log p(\mathbb{D}|\gamma) > \log p(\mathbb{D}|\gamma')$ often \Longrightarrow
- Near modal values in terms of log MLIK are particularly important for construction of reasonable $\mathbb{V}\subset\Omega_{\gamma}$, missing them can dramatically influence posterior in the original space Ω_{γ}

Possible ways to explore $\mathbb{V} \subset \Omega_{\gamma}$

Main challenges are multimodality in Ω_{γ} and its size.

- ullet Full enumeration of Ω_{γ} infeasible for large dimensions
- Random walk in Ω_{γ} including simple MCMC does not take advantage of the structure of $\Omega_{\gamma} \Longrightarrow$ too slow
- Greedy optimization with numerous initial points end up in local optima
- Random walk with mode jumping proposals seems to be a good idea



MCMC with locally optimized proposals

Tjelmeland and Hegstad [6] suggested continuous mode jumping proposals, Storvik [5] considers a more general setup, we suggest mode jumping proposals in the discrete parameter spaces.

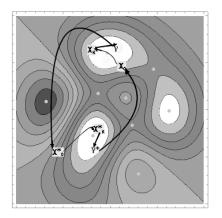


Figure: Locally optimized with randomization proposals

Application of MCMC with mode jumping proposals

We have shown that the detailed balance equation is satisfied for the following acceptance probabilities:

$$r_{m}(\gamma_{j}, \gamma_{k}) = \min \left\{ 1, \frac{p(\mathbb{D}|\gamma_{k})p(\gamma_{k})q_{s}(\gamma_{j}|\gamma_{j_{K-1}})}{p(\mathbb{D}|\gamma_{j})p(\gamma_{j})q_{s}(\gamma_{k}|\gamma_{k_{K-1}})} \right\}.$$
(9)

• $q_s(.|.)$ is the kernel of randomization at the end.

Hence we also obtain alternative MCMC estimators of posterior marginal probabilities

$$\tilde{p}(\gamma|\mathbb{D}) = \frac{\sum_{i=1}^{W} \mathbb{I}(\gamma_i = \gamma)}{W} \xrightarrow{d} p(\gamma|\mathbb{D}). \tag{10}$$

• W is the number of MCMC iterations (after burn-in)

How it looks like in reality

Modes are important: the standard MCMC procedure (right) misses two in this example. Visualization is challenging

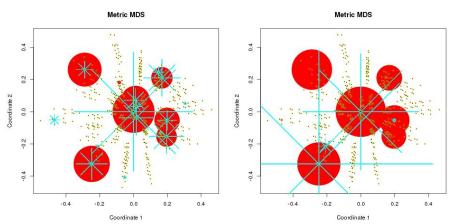


Figure: MDS plots with posterior modes of all found solutions for the approaches

Application to cosmological simulations. Cosmological hydro-simulation data (http://yt-project.org/data/).

Observations (Bernoulli classifiers):

- Galaxy is quenched (or not)
- Star hosts planet (or not)

Variables:

- Dark matter mass
- Gas mass
- Stellar mass
- Star formation rate
- Metallicity
- Gas molecular fraction
- Gas fraction
- Stellar fraction
- Stellar to gas mass ratio
- Other covariates, their interactions, polynomes and etc.

Application to NEO classification. NASA Space Challenge (https://github.com/SpaceApps2016/Resources).

Observations (Bernoulli classifiers):

• Asteroid is a NEO (PHA) object or not (Phocaea)

Variables:

- Rotation period
- Magnitude slope
- Mean anomaly
- Inclination
- Argument of perihelion
- Longitude of the ascending node
- Rms residual
- Semi major axis
- Eccentricity
- Mean motion
- Absolute magnitude
- Other covariates, their interactions, polynomes and etc.

Application to cosmological simulations

Logistic regression addressed by

$$y_t = y | p_t \sim Binom(1, p_t) \tag{11}$$

$$p_{t} = \frac{e^{\gamma_{0}\beta_{0} + \sum_{i=1}^{p} \gamma_{i}\beta_{i}X_{t,i}}}{1 + e^{\gamma_{0}\beta_{0} + \sum_{i=1}^{p} \gamma_{i}\beta_{i}X_{t,i}}}$$
(12)

$$\beta|\gamma \sim N_{\sum_{i=1}^{p} \gamma_{i}}(\mu_{\beta}, \Sigma_{\beta})$$
 (13)

$$\gamma_i \sim Binom(1,q)$$
 (14)

NEO objects classification. NASA space challenge

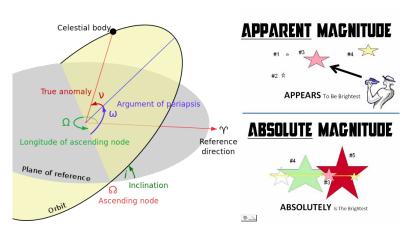


Figure: Orbital elements(left) by Lasunncty (talk), CC BY-SA 3.0 and absolute vs apparent magnitude (right) by Mrscreath(http://mrscreath.blogspot.com)

NEO objects classification. NASA space challenge

20 covariates addressed in the experiment (both reasonable and heuristic): Mean anomaly ∈ [0°; 360°); Argument of perihelion $\in [0^{\circ}; 360^{\circ});$ Longitude of the ascending node $\in [0^{\circ}; 360^{\circ});$ Inclination $\in [0^{\circ}; 180^{\circ}];$ Semi major axis $\in \mathbb{R}^{+};$ Eccentricity $\in \mathbb{R}^{+};$ Mean motion $\in \mathbb{R}^{+};$ Absolute magnitude $\in \mathbf{R}$ (brightness); Rms residual $\in \mathbf{R}^+$ (brightness error); Eccentricity² $\in \mathbb{R}^+$; Absolute magnitude² $\in \mathbb{R}^+$; Semi major axis² $\in \mathbb{R}^+$; Semi major axis³ $\in \mathbb{R}^+$; Mean anomaly×Semi major axis; Mean anomaly×Semi major axis² $\in \mathbb{R}^+$; Mean anomaly×Semi major axis³ $\in \mathbb{R}^+$; Argument of perihelion×Semi major axis $\in \mathbb{R}^+$; Argument of perihelion \times Semi major axis² $\in \mathbb{R}^+$; Argument of perihelion \times Semi major axis³ $\in \mathbb{R}^+$: Longitude of the ascending node×Semi major axis $\in \mathbb{R}^+$.

Training set includes 32 NEO and 32 non-NEO objects, **test set** includes 20720 objects (14099 NEO, 6621 non-NEO), **validation sets** were used as some random subsets of a 100 elements from these 20720 **objects**

 2^{20} models in total, algorithm was run until ca 2500 models are visited.

NEO objects classification. Inference

Posterior inclusion probabilities and posterior model probabilities

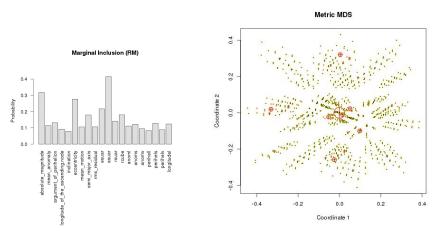


Figure: Comparison of marginal inclusion probabilities of the covariates (left) and models on the whole (right)

NEO objects classification. Bayesian classification

Choice of \mathbb{V}^* is crucial, $\mathbb{V}^* = \Omega_{\gamma}$ - often in-feasible, $\mathbb{V}^* = \mathbb{V}$ - very precise can be too slow, $\mathbb{V}^* = \mathbb{V} \cap p(\gamma|\mathbb{D}) \geq \alpha$ - often precise, but is a way faster!!!

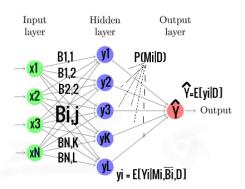


Figure: Bayesian Artificial Neuron Network for Classification

$$\hat{Y} = \mathbb{I}\{\hat{\mathsf{E}}\big[\,Y|\mathbf{D}\big] \geq 0.5\}, \hat{\mathsf{E}}\big[\,Y|\mathbf{D}\big] = \sum_{\gamma \in \mathbb{V}^*} \hat{\mathsf{E}}\big[\,y_\gamma|\gamma,\mathbf{D}\big]\hat{\rho}\big(\gamma|\mathbf{D}\big)$$

NEO objects classification. Results

Quite impressive actually... Surprisingly or not?.. Comments?.. Remember: $||\mathbf{training\ set}|| = 64$, $||\mathbf{test\ set}|| = 20720$

99.95656% 99.80212%	0.1208277% 0.1208277%	0.00709270% 0.2340592%	619.89 min	1.795 sec
99.80212%	0.1208277%	0.22405020/		
		0.234059276	172.66 min	0.499 sec
99.46429%	0.0906208%	0.7447337%	29.166 min	0.084 sec
90.00483%	0.1057242%	14.639340%	4.7789 min	0.014sec
82.83301%	0.1510346%	25.15781%	4.5222 min	0.013 sec
93.86271%	1.0000000%	17.000000%	-	-
	90.00483% 82.83301%	90.00483% 0.1057242% 82.83301% 0.1510346%	90.00483% 0.1057242% 14.639340% 82.83301% 0.1510346% 25.15781%	90.00483% 0.1057242% 14.639340% 4.7789 min 82.83301% 0.1510346% 25.15781% 4.5222 min

Table: Comparison of performance (Precision, FDR, FNR, Time) of different models

N/B: the best model includes eccentricity², eccentricity, absolute magnitude², absolute magnitude

Multicore and shared memory issues

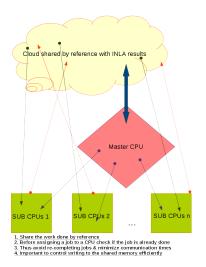


Figure: Multiprocessing architecture

The protein activity data. 288 models. Multiple modes

Comparison to other algorithms: BAS, RS (simper MCMC) on 2^{20} unique models visited for MJMCMC and BAS and 88×2^{20} iterations of RS.

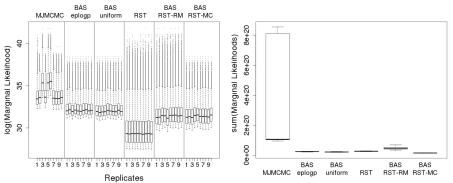


Figure: 100000 best mliks found (left) and posterior masses captured (right). Bayesian linear regression with a g-prior is addressed, since no other packages (to our awareness) manage model selection in GLMM

The protein activity data. 288 models. Multiple modes

Checking convergence. Marginal inclusion probabilities

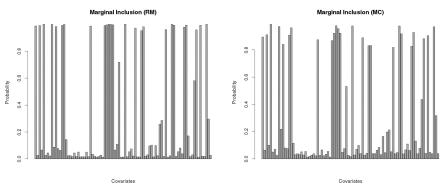


Figure: Comparison of marginal inclusion probabilities obtained by the Bayes formula and MCMC approximations from the best run of MJMCMC with 8.56e + 20 posterior mass captured

Concluding remarks

- We introduced the MJMCMC approach for estimating posterior model probabilities and Bayesian model averaging and selection.
- It incorporates the ideas of MCMC with possibility of large jumps combined with local optimizers to generate proposals in the discrete space of models
- EMJMCMC R-package is developed and available from the GitHub repository: http://aliaksah.github.io/EMJMCMC2016/
- The developed package gives a user high flexibility in the choice of methods to obtain marginal likelihoods and model selection criteria within GLMM
- Extensive parallel computing for both MCMC moves and local optimizers is available within the developed package
- Based on the obtained in the experimental part results, we can claim MJMCMC to be a rather competitive novel algorithm that both performs well in terms of the search quality and addressed a more general class of statistical models than the competing approaches

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The End.



Thank you.