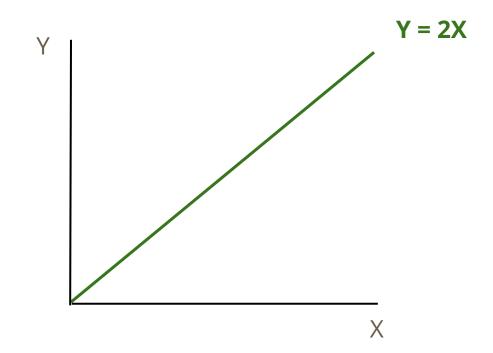
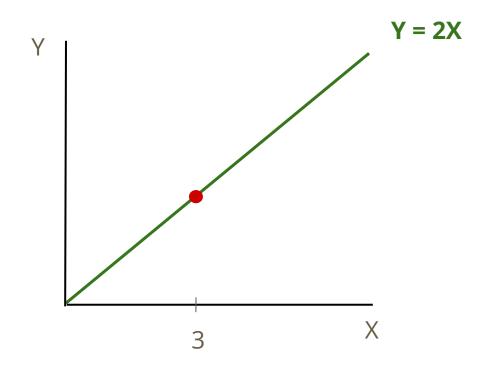
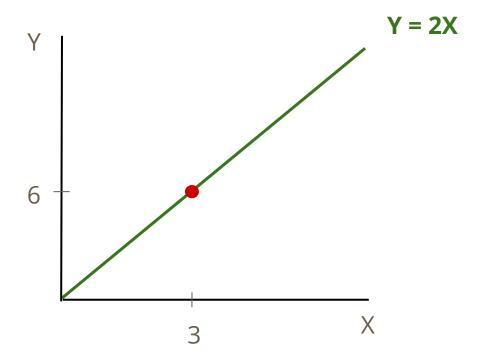
Linear Model Evaluation

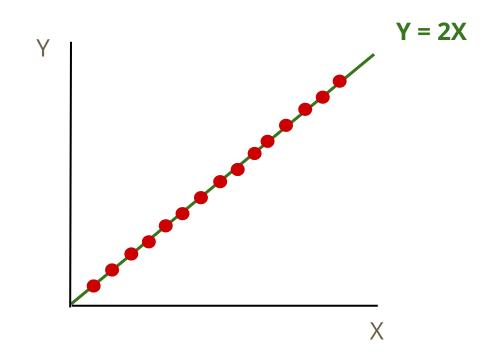
Boston University CS 506 - Lance Galletti



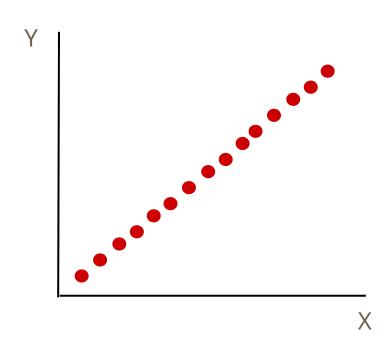


I know for sure what value of Y I'm gonna get



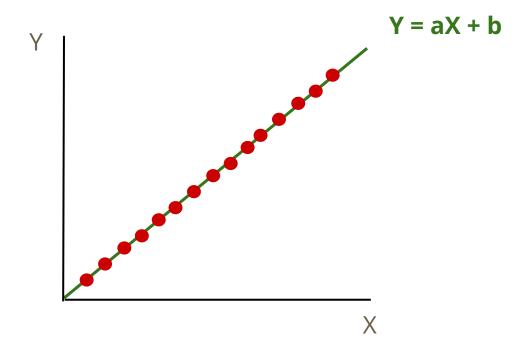


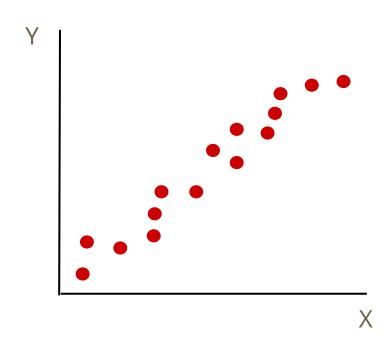
Ideally

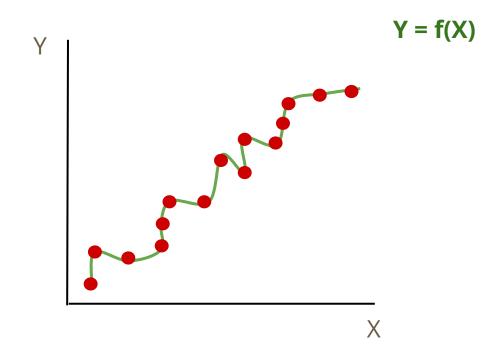


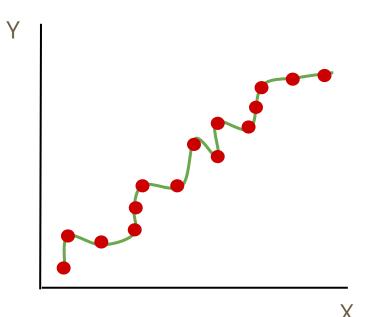
Ideally

Guess the relationship (a and b) from the data



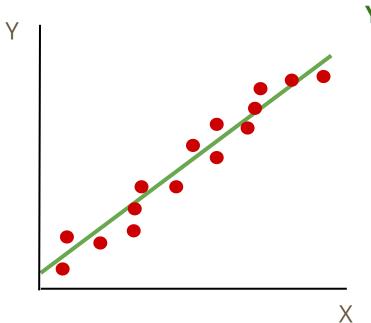






$$Y = f(X)$$

f can be anything: too complex a problem to solve...



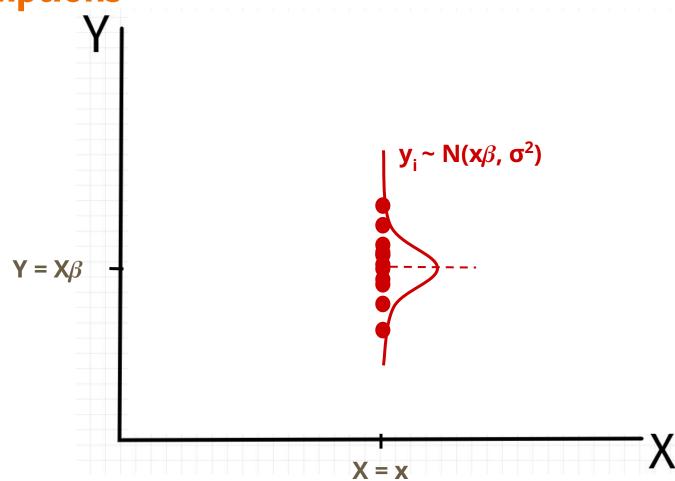
$$Y = f(X) + noise$$

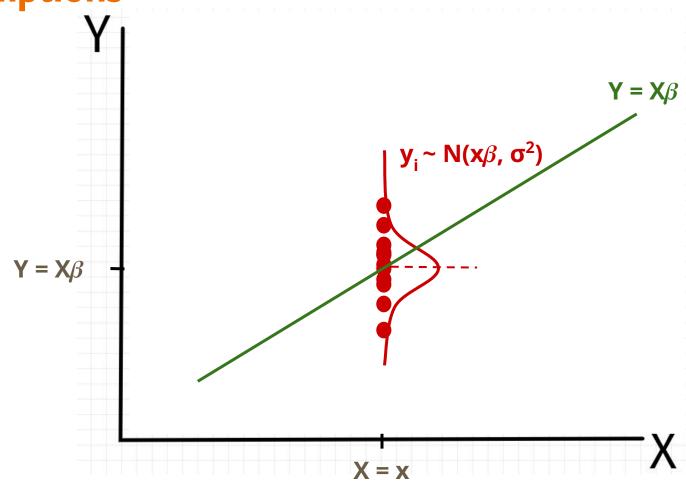
Assume f is linear and the variation we're seeing is noise

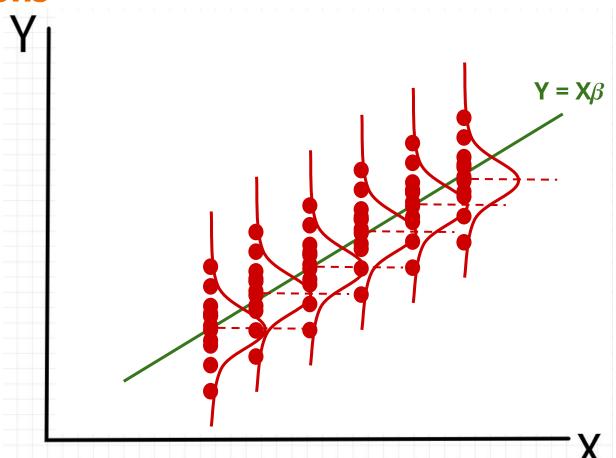
Assumptions X = x

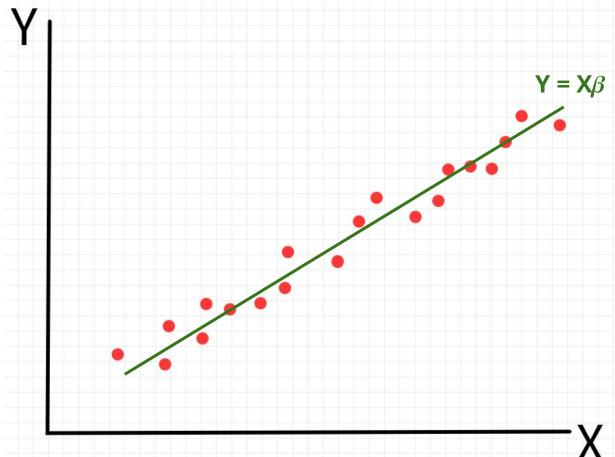
Assumptions X = x

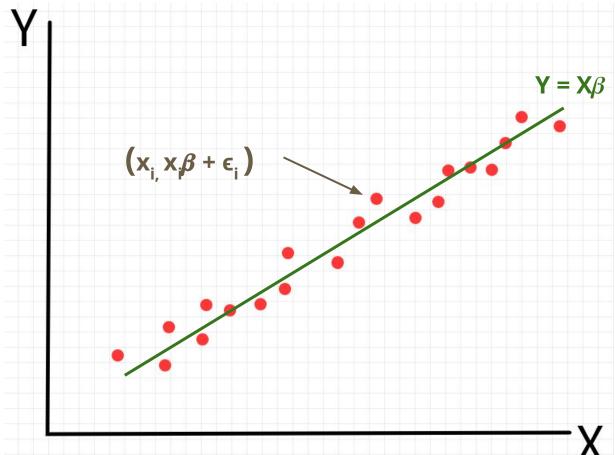
Assumptions X = x

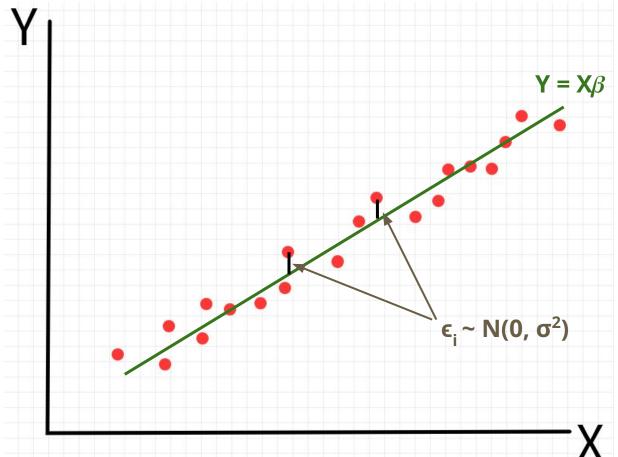




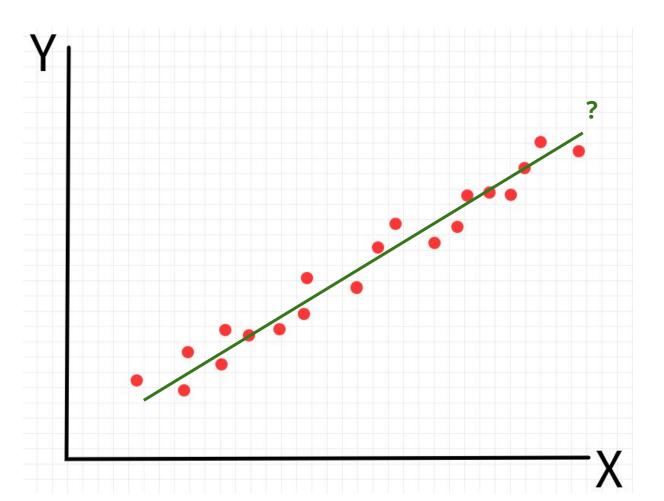


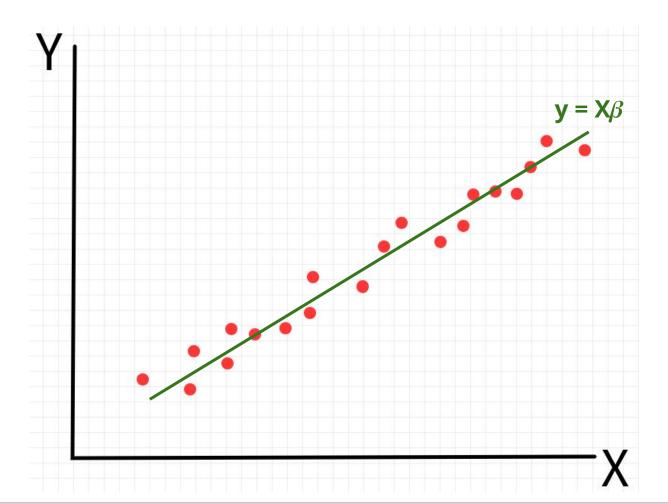


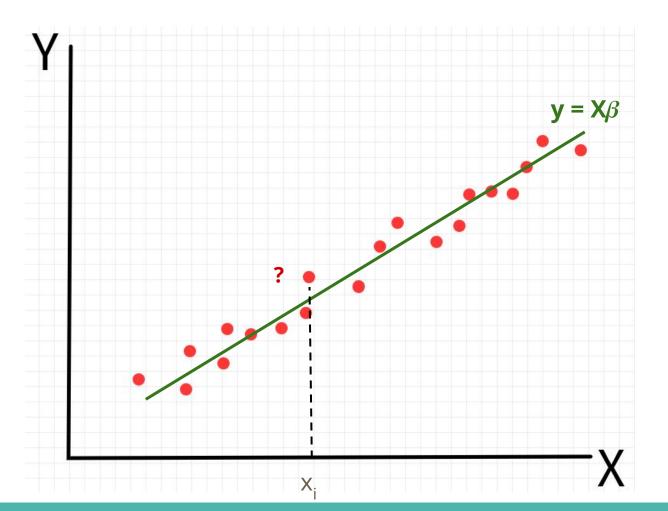


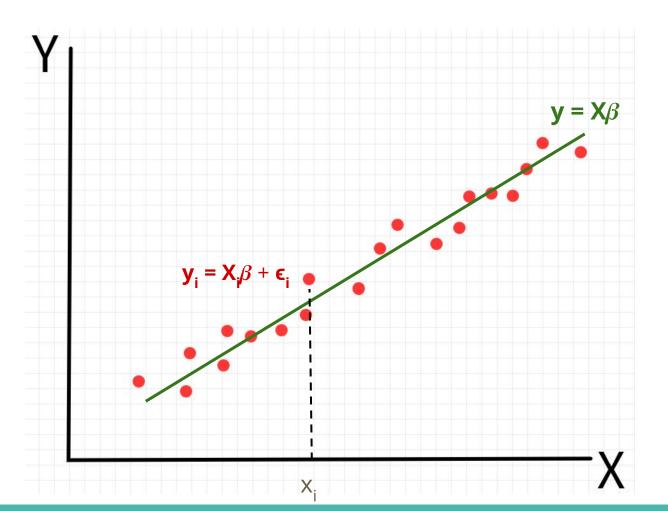


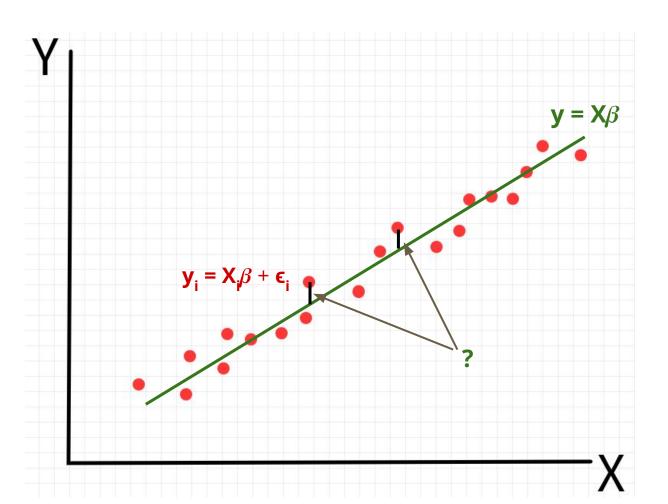
Let's label the diagram

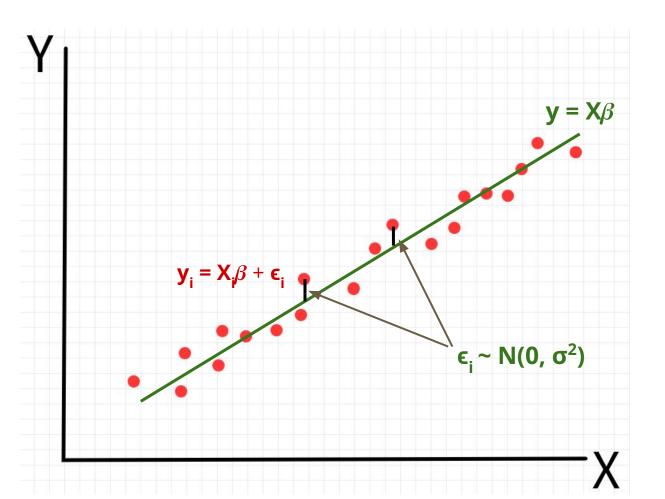


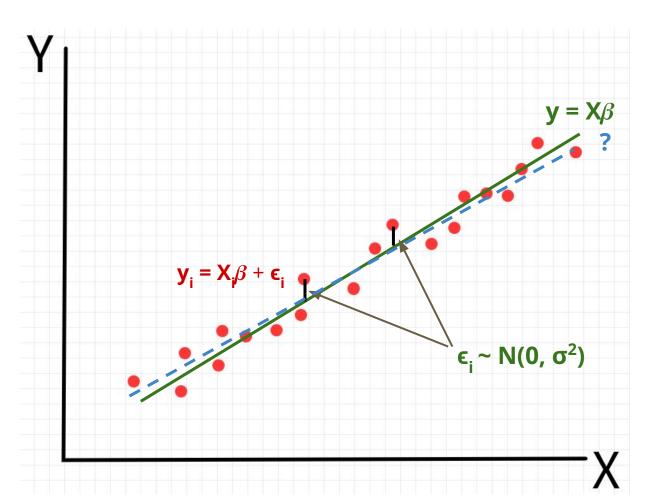


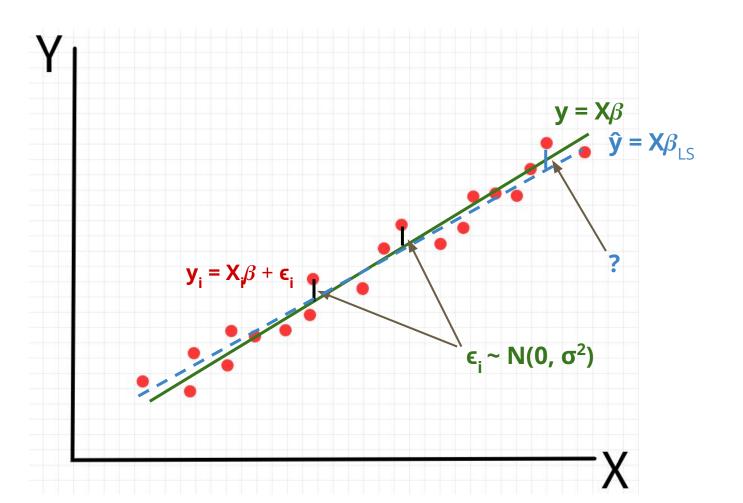


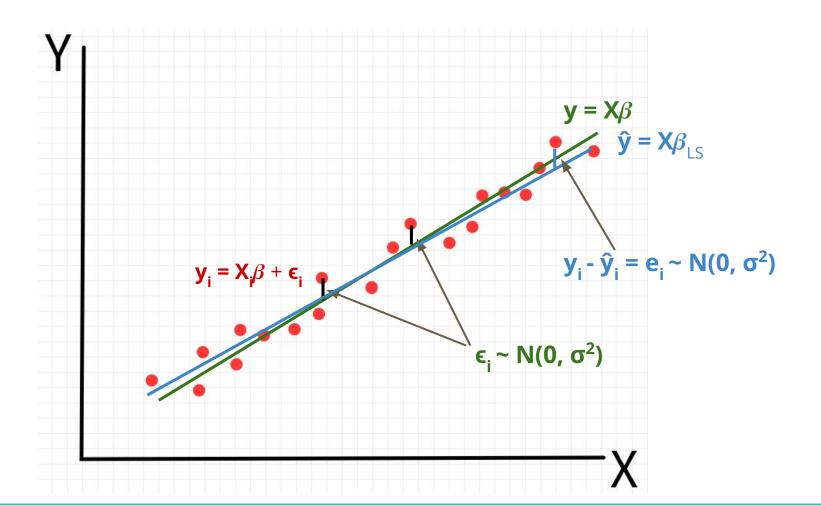


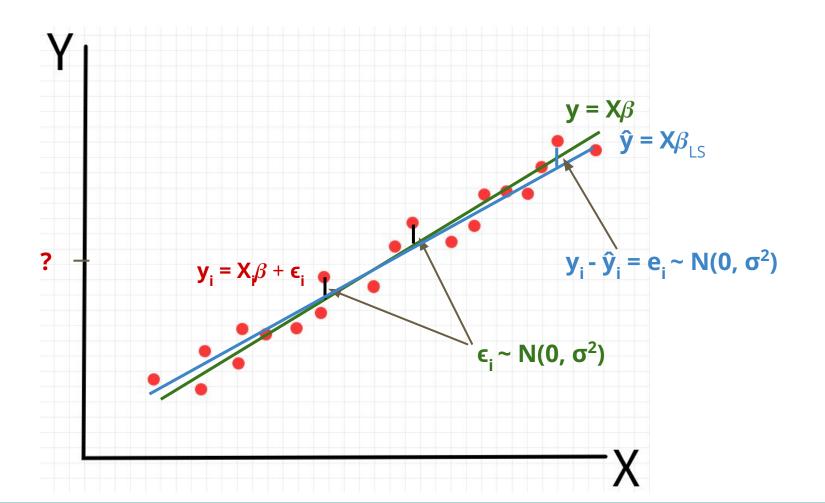


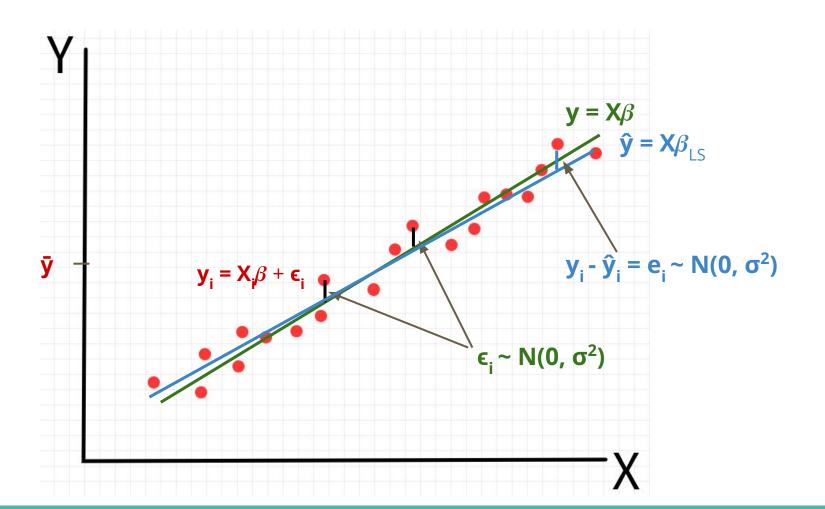












How do we know a linear model is applicable?

Is it true that $\mathbf{y} = \mathbf{x}\boldsymbol{\beta} + \boldsymbol{\epsilon}$ where $\boldsymbol{\epsilon} \sim \mathbf{N}(\mathbf{0}, \boldsymbol{\sigma}^2)$?

How do we know a linear model is applicable?

Is it true that $\mathbf{y} = \mathbf{x}\boldsymbol{\beta} + \boldsymbol{\epsilon}$ where $\boldsymbol{\epsilon} \sim \mathbf{N}(\mathbf{0}, \boldsymbol{\sigma}^2)$?

If this is true then: $\mathbf{y}_i - \hat{\mathbf{y}}_i = \mathbf{e}_i \sim \mathbf{N}(0, \sigma^2)$

Can we check this?

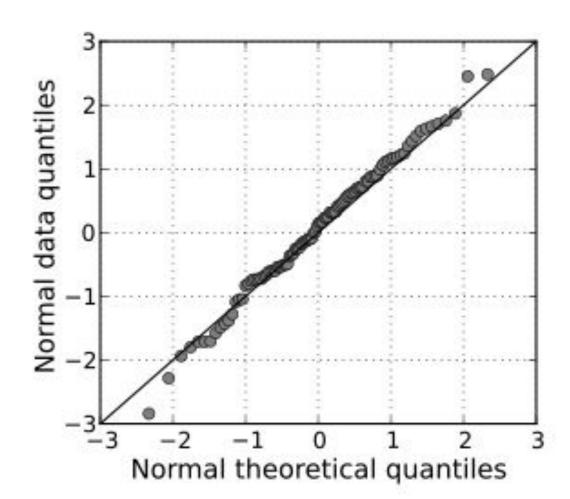
QQ plot

Quantiles are the values for which a particular % of values are contained below it.

For example the 50% quantile of a N(0,1) distribution is 0 since 50% of samples would be contained below 0 were you to sample a large number of times.

Forall quantiles q, if sample.q == known_distribution.q then they have the same distribution.

QQ plot



demo

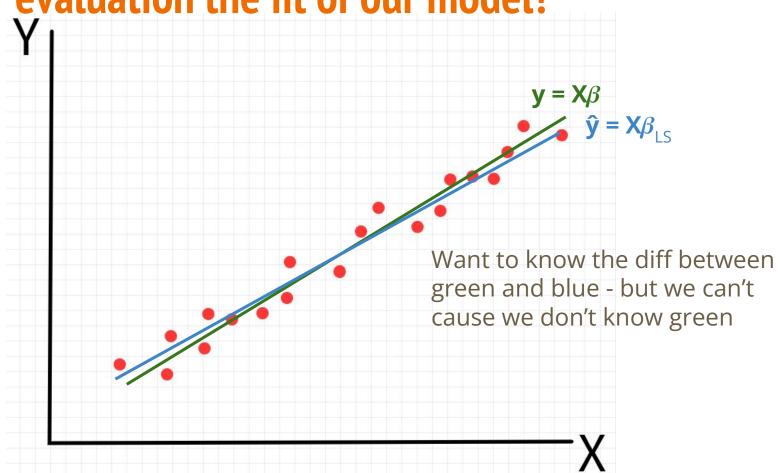
Some Notation:

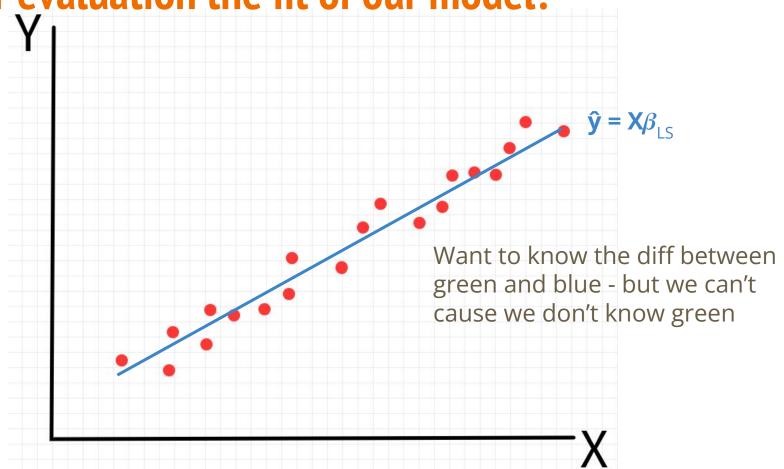
 $\mathbf{y_i}$ is the "true" value from our data set (i.e. $\mathbf{x_i}\boldsymbol{\beta} + \boldsymbol{\epsilon_i}$)

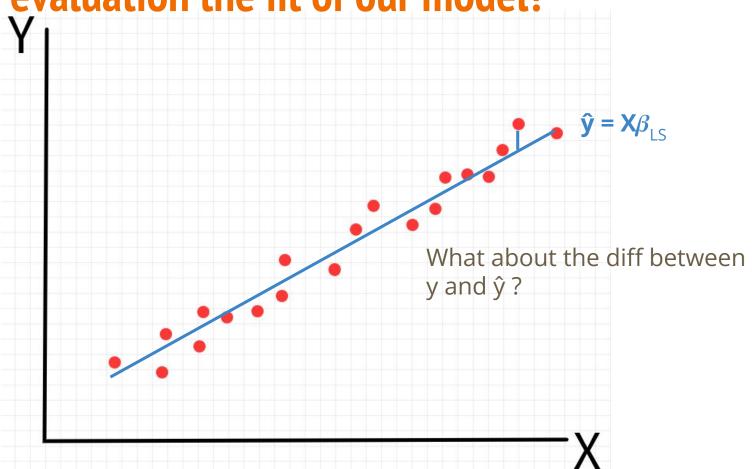
 $\hat{\mathbf{y}}_{i}$ is the estimate of \mathbf{y}_{i} from our model (i.e. $\mathbf{x}_{i}\boldsymbol{\beta}_{LS}$)

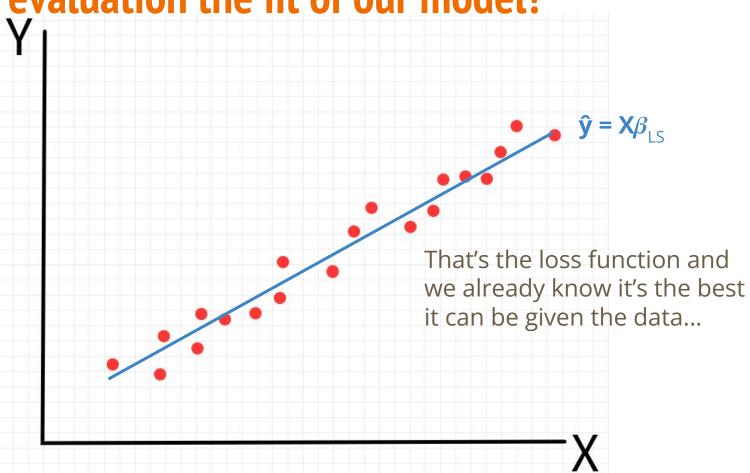
 $ar{\mathbf{y}}$ is the sample mean all $\mathbf{y_i}$

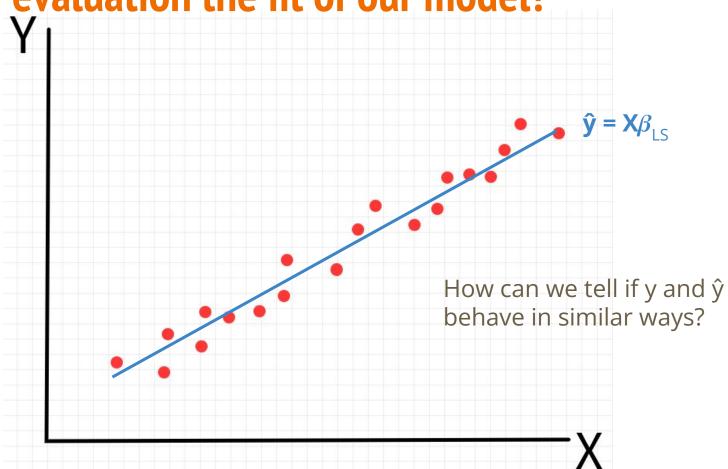
 $\mathbf{y_i}$ - $\mathbf{\hat{y}_i}$ are the estimates of $\mathbf{\epsilon_i}$ and are referred to as residuals

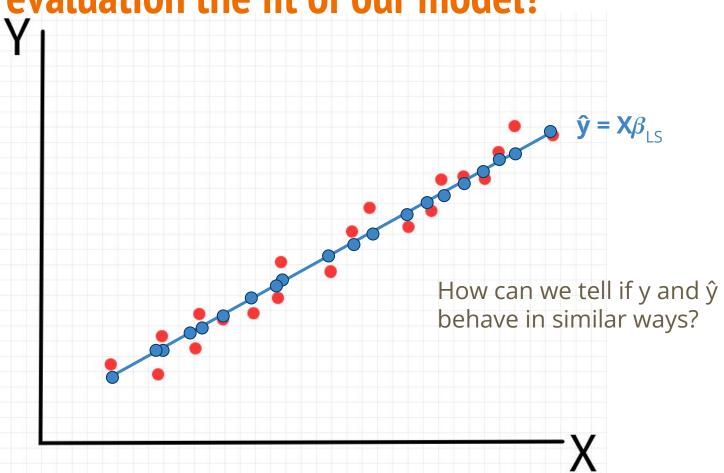


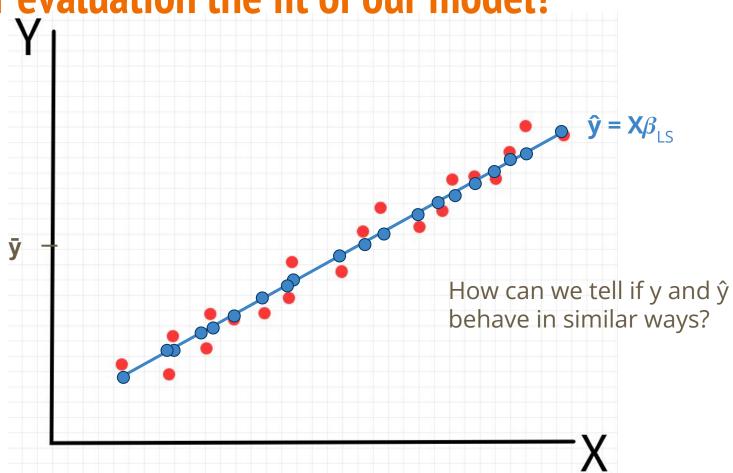


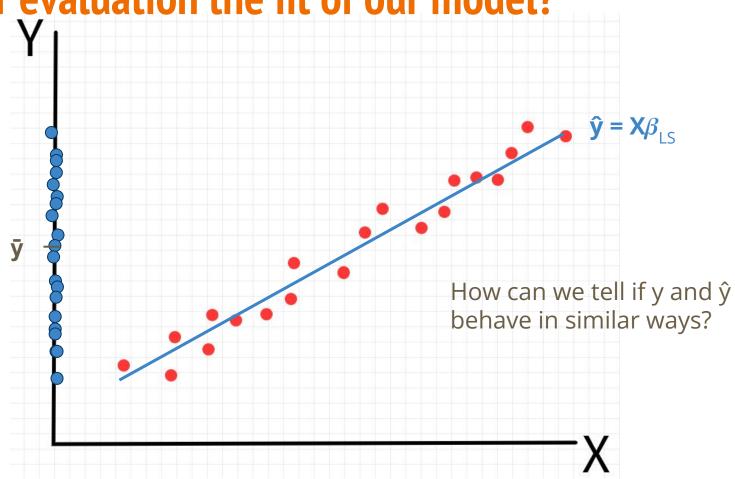


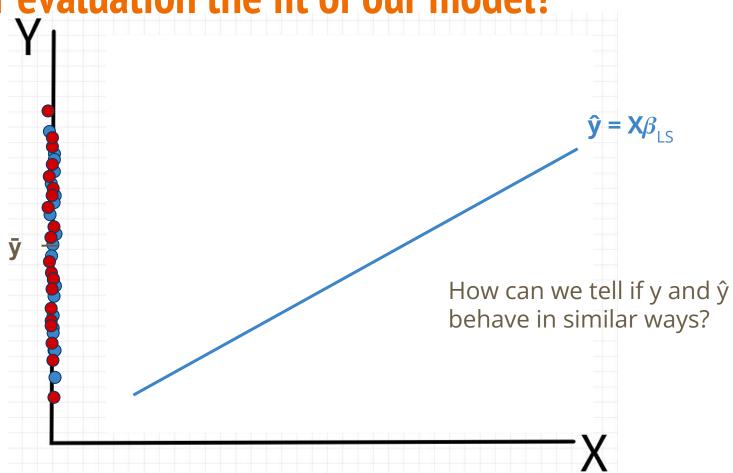












Metric for evaluation the fit of our model? $\hat{\mathbf{y}} = \mathbf{X}\boldsymbol{\beta}_{\mathsf{LS}}$ How can we tell if y and ŷ behave in similar ways?

Is the value of the loss function sufficient? i.e.

$$||y - X\beta||_2^2 = \sum_i (y_i - \hat{y_i})^2$$

$$TSS = \sum_i^n (y_i - \bar{y})^2$$
 This is a measure of the spread of $\mathbf{y}_{_{\! i}}$ around the mean of \mathbf{y}

$$TSS = \sum_i^n (y_i - \bar{y})^2$$
 This is a measure of the spread of \mathbf{y}_i around the mean of \mathbf{y}

$$ESS = \sum_{i}^{n} (\hat{y}_i - \bar{y})^2$$

This is a measure of the spread of our model's estimates of y_i around the mean of y

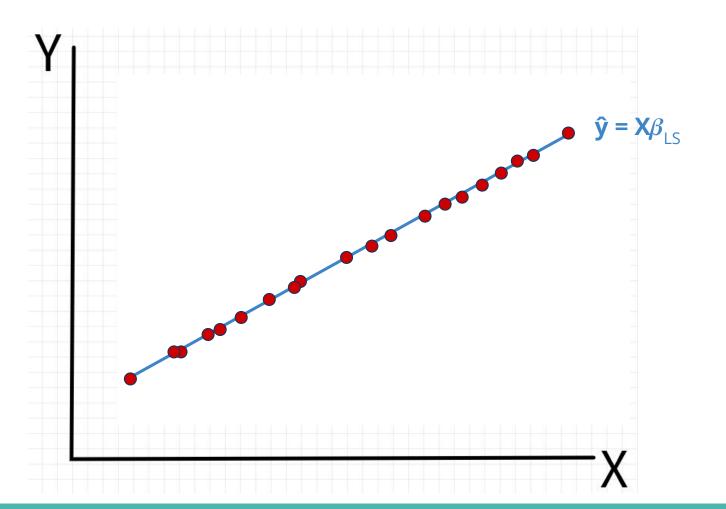
$$TSS = \sum_{i}^{n} (y_i - \bar{y})^2$$

$$R^2 = \frac{ESS}{TSS}$$

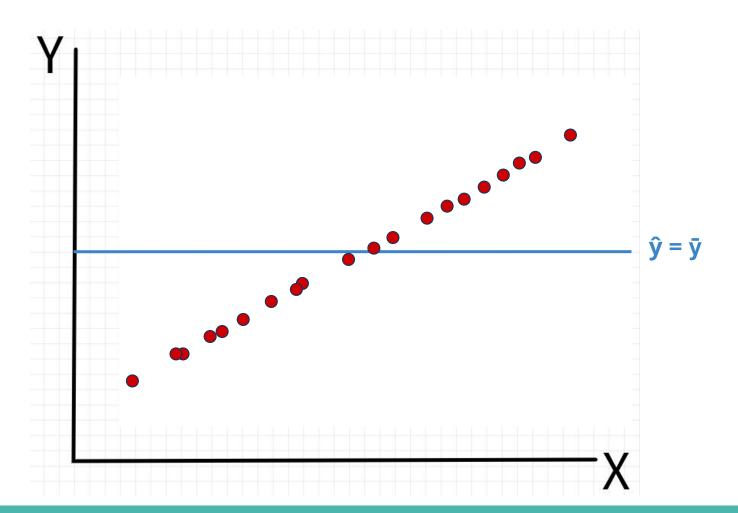
$$ESS = \sum_{\dot{}} (\hat{y_i} - \bar{y})^2$$

 R^2 measures the fraction of variance that is explained by \hat{y} (our model)

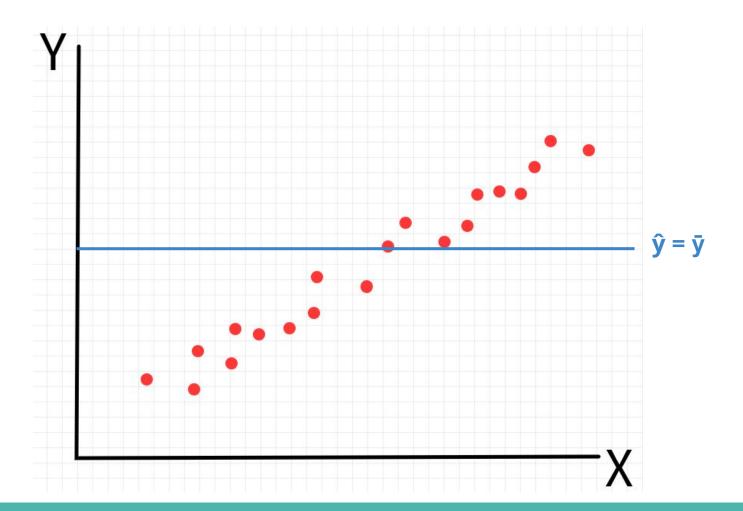
 $R^2 = 1$



 $\mathbf{R}^2 = \mathbf{0}$



 $\mathbf{R}^2 = \mathbf{0}$



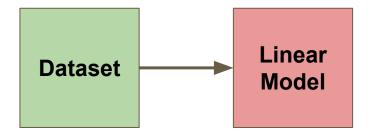
```
OLS Regression Results
Dep. Variable:
                                       R-squared:
                                                                       0.840
Model:
                                       Adj. R-squared:
                                 OLS
                                                                       0.836
Method:
                       Least Squares F-statistic:
                                                                       254.1
Date:
                    Sun, 20 Mar 2022
                                       Prob (F-statistic):
                                                                    2.72e-39
Time:
                                       Log-Likelihood:
                                                                     -482.37
                            11:36:16
No. Observations:
                                       AIC:
                                                                       970.7
                                 100
Df Residuals:
                                  97
                                       BTC:
                                                                       978.5
Df Model:
Covariance Type:
                           nonrobust
                coef
                        std err
                                                P>|t|
                                                           [0.025
                                                                      0.975]
const
              2.1912
                          3.162
                                     0.693
                                                0.490
                                                         -4.085
                                                                       8.467
             29.3912
x1
                          3.274
                                    8.977
                                                0.000
                                                          22.893
                                                                      35.889
                                                0.000
             78.1391
                          3.594
                                    21.741
                                                           71.006
                                                                      85.272
Omnibus:
                                       Durbin-Watson:
                               1.279
                                                                       1.824
Prob(Omnibus):
                               0.527
                                       Jarque-Bera (JB):
                                                                       1.065
Skew:
                               0.253
                                       Prob(JB):
                                                                       0.587
                                       Cond. No.
Kurtosis:
                                                                        1.38
```

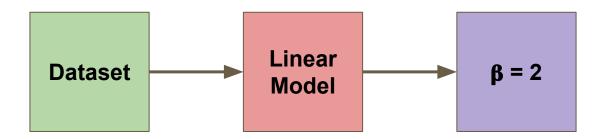
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                                                  0.490
                                                             -4.085
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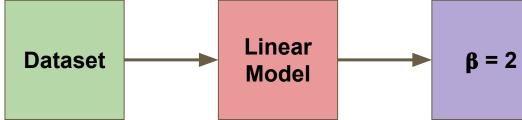
Dataset





Could the real beta be 5?





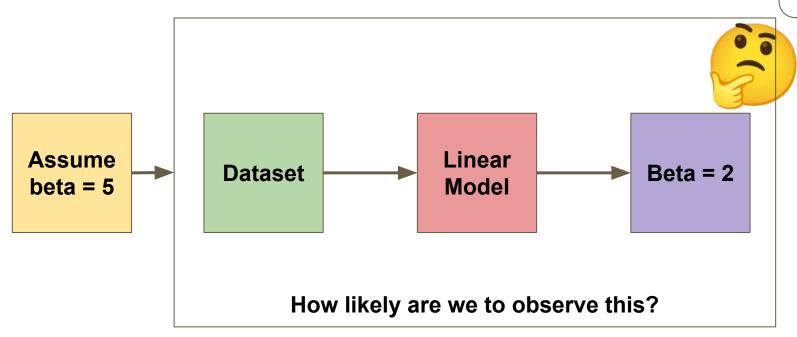
ннннннн

The coin is probably not fair

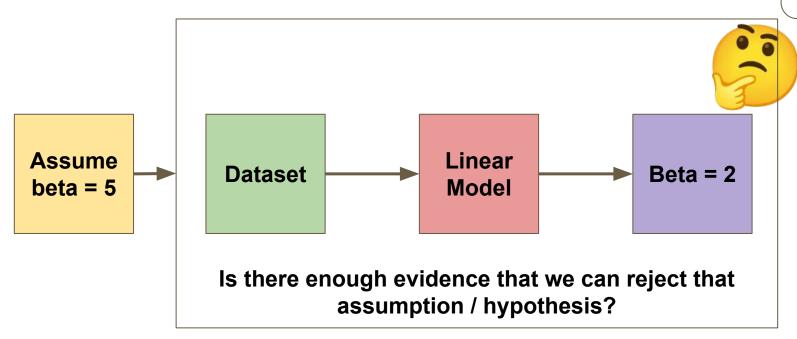
HTHHTHTT



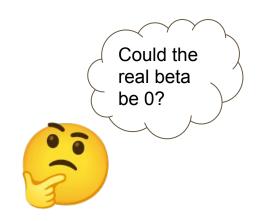
Could the real beta be 5?



Could the real beta be 5?



worksheet



Each parameter of an independent variable \mathbf{x} has an associated confidence interval and t-value + p-value.

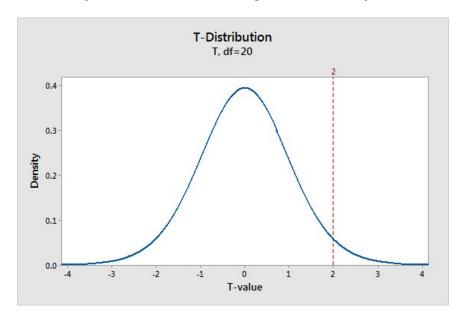
If the parameter / coefficient is not significantly distinguishable from 0 then we cannot assume that there is a significant linear relationship between that independent variable and the observations \mathbf{y} (i.e. if the interval includes 0 or if the p-value is too large)

We want to know if there is evidence to reject the hypothesis H0 : β = 0 (i.e. that there is no linear relation between X and Y) using the information from β hat.

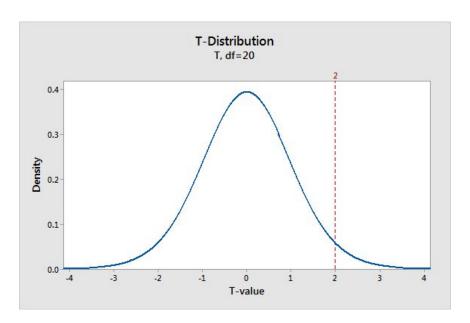
We want to know the largest probability of obtaining the data observed, under the assumption that the null hypothesis is correct.

How do we obtain that probability?

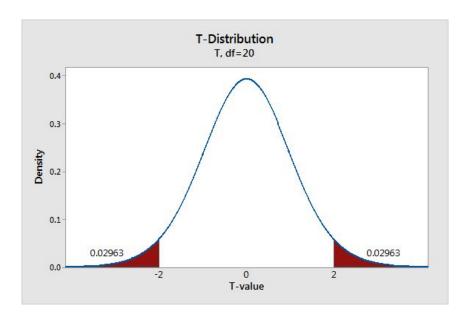
Under the null hypothesis what should be the distribution of the normalized estimates? T-distribution (parametrized by the sample size)



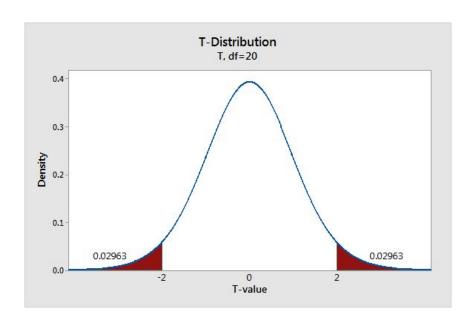
We can then compute the t-value that corresponds to the sample we observed.



And then compute the probability of observing estimates of β at least as extreme as the one observed. (i.e. trying to find evidence against H0)



This probability is called a p-value.



A p-value **smaller than a given threshold** would mean the data was unlikely to be observed under H0 so we can reject the hypothesis H0. If not, then we lack the evidence to reject H0.

| | ======== coef | ======= std err | t | P> t | [0.025 | 0.975] |
|-------------------|------------------------------|-------------------------|--------------------------|-------------------------|----------------------------|---------------------------|
| const x1 x2 | 2.1912 29.3912 78.1391 | 3.162 3.274 3.594 | 0.693 8.977 21.741 | 0.490 0.000 0.000 | -4.085 22.893 71.006 | 8.467 35.889 85.272 |
| | | | | | | |

Which parameters should we not include in our linear model?

| ======= | coef | std err | t | P> t | [0.025 | 0.975] |
|-------------------|------------------------------|-------------------------|--------------------------|-------------------------|----------------------------|---------------------------|
| const x1 x2 | 2.1912 29.3912 78.1391 | 3.162 3.274 3.594 | 0.693 8.977 21.741 | 0.490 0.000 0.000 | -4.085 22.893 71.006 | 8.467 35.889 85.272 |
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const
               2.1912
                           3.162
                                      0.693
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Skew:
                                        Prob(JB):
                                0.253
                                                                          0.587
                                         Cond. No.
Kurtosis:
                                                                           1.38
```

Confidence Intervals

Goal: for a given confidence level (let's say 90%), construct an interval around an estimate such that, if the estimation process were repeated indefinitely, the interval would contain the true value (that the estimate is estimating) 90% of the time.

| | coef | std err | t | P> t | [0.025 | 0.975] |
|-----------|---------|---------|--------|-------|--------|--------|
| const | 2.1912 | 3.162 | 0.693 | 0.490 | -4.085 | 8.467 |
| ×1 | 29.3912 | 3.274 | 8.977 | 0.000 | 22.893 | 35.889 |
| x2 | 78.1391 | 3.594 | 21.741 | 0.000 | 71.006 | 85.272 |

Z-values

These are the number of standard deviations from the mean of a N(0,1) distribution required in order to contain a specific % of values were you to sample a large number of times.

To find the .95 z-value (the value z such that 95% of the observations lie within z standard deviations of the mean ($\mu \pm z * \sigma$)) you need to solve:

$$\int_{-z}^{z} \frac{1}{2\pi} e^{-\frac{1}{2}x^2} dx = .95$$

Z-values

The .95 z-value is 1.96.

This means 95% of observations from a N(μ , σ) lie within 1.96 standard deviations of the mean (μ ± 1.960 * σ)

If we get a sample from a $N(\mu, \sigma)$ of size n, how would we create a confidence interval around the estimated mean?

Confidence Intervals

How do we build a confidence interval?

Assume $Y_i \sim N(5, 25)$, for $1 \le i \le 100$ and $y_i = \mu + \epsilon$ where $\epsilon \sim N(0, 25)$. Then the Least Squares estimator of μ (μ_{LS}) is

the sample mean \bar{y} $SE(\mu_{LS}) = \sigma_{\varepsilon} / \sqrt{n}$ What is the 95% confidence interval for μ_{LS} ? $= 5 / \sqrt{100}$ = .5 $CI_{.95} = [\bar{y} - 1.96 \times SE(\mu_{LS}), \bar{y} + 1.96 \times SE(\mu_{LS})]$ $= [\bar{y} - 1.96 \times .5, \bar{y} + 1.96 \times .5]$ Z-value for 95% Confidence Interval

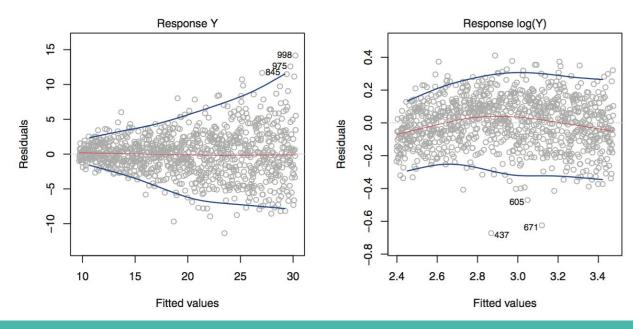
Checking our Assumptions

- 1. Normal Distribution?
- 2. Constant Variance?

Constant Variance

One of our assumptions was that our noise had constant variance. How can we verify this?

We can plot residuals (noise estimates) for each fitted value $\hat{\mathbf{y}}_{\mathbf{i}}$



Extending our Linear Model

Changing the assumptions we made can drastically change the problem we are solving. A few ways to extend the linear model:

- 1. Non-constant variance used in WLS (weighted least squares)
- Distribution of error is not Normal used in GLM (generalized linear models)