

Exercise 1: σ -Algebra

- a) Let $\Omega = \mathbb{N}$. Then the subset of even numbers $A_0 = \{2, 4, 6, \dots\}$ is not contained in \mathcal{F}_Ω as A_0 is not finite and its complement A_1 is also not finite $A_1 = \{1, 3, 5, \dots\}$. So $\mathcal{F}_\Omega \neq 2^\Omega$. \square
- b) No. the definition says:

$$\forall A \subseteq \Omega : ((|A| < \infty \Rightarrow A \in \mathcal{F}_\Omega) \wedge (A \in \mathcal{F}_\Omega \Rightarrow (\Omega \setminus A) \in \mathcal{F}_\Omega))$$

i.e. (i) means a finite subset is in \mathcal{F}_Ω . (ii) means if the complement of a subset is finite, the subset is in \mathcal{F}_Ω . Thus if neither the set is finite nor its complement, none of them is a member of \mathcal{F}_Ω . To summarize: Even if the largest class of subsets is considered the definition is non-trivial and the example from a) still holds i.e. it may still be that $\mathcal{F}_\Omega \neq 2^\Omega$. \square

- c) (Proof or) disproof that $(\Omega, \mathcal{F}_\Omega)$ is a σ -Algebra.
Consider $A_i = \{2i\}$. $A_i \in \mathcal{F}_\Omega$ as it's finite, but $\cup_i A_i \notin \mathcal{F}_\Omega$ as it's not finite and its complement is also not finite (see a)). So \mathcal{F}_Ω not a σ -Algebra \square

Exercise 2:

- a. Given from slides:

$$\begin{aligned} E[X] &= \sum_{i=0}^n x_i \cdot Pr_X(X = x_i), \\ Var[X] &= E[X^2] - E[X]^2 = E[X(X-1)] + E[X] - E[X^2], \\ \text{Given from Analysis I (Geometric series, first \& second derivative):} \\ \sum_{k=0}^{\infty} q^k &= \frac{1}{1-q}, \\ \sum_{k=0}^{\infty} k \cdot q^{k-1} &= \frac{1}{(1-q)^2}, \\ \sum_{k=0}^{\infty} k(k-1) \cdot q^{k-2} &= \frac{-2}{(1-q)^3}. \end{aligned}$$

With X geometrically distributed and $q = 1 - p$
Expected Value:

$$\begin{aligned} E[X] &= \sum_{k=1}^{\infty} k \cdot (q)^{k-1} \cdot p \\ &\Leftrightarrow p \sum_{k=1}^{\infty} k \cdot q^{k-1} \\ &\Leftrightarrow p \frac{1}{(1-q)^2} = \frac{p}{(1-1+p)(1-1+p)} \\ &\Leftrightarrow \frac{p}{p^2} = \frac{1}{p} \end{aligned}$$

\square

Variance:

$$\begin{aligned}
 \text{Var}[X] &= E[X(X-1)] + E[X] - E[X^2] = \sum_{k=1}^{\infty} k(k-1) \cdot q^{k-1} \cdot p + \frac{1}{p} - \frac{1}{p^2} \\
 &\Leftrightarrow p \sum_{k=1}^{\infty} k(k-1) \cdot q^{k-1} + \frac{1}{p} - \frac{1}{p^2} \\
 &\Leftrightarrow p \sum_{k=1}^{\infty} k(k-1) \cdot q^{k-1} q^{-1} q^1 + \frac{1}{p} - \frac{1}{p^2} \\
 &\Leftrightarrow pq \sum_{k=1}^{\infty} k(k-1) \cdot q^{k-2} + \frac{1}{p} - \frac{1}{p^2} \\
 &\Leftrightarrow pq \frac{-2}{(1-q)^3} + \frac{1}{p} - \frac{1}{p^2} \\
 &\Leftrightarrow \frac{-2p(1-p)}{(1-1-p)^3} + \frac{1}{p} - \frac{1}{p^2} \\
 &\Leftrightarrow \frac{-2(1-p)}{-p^2} + \frac{1}{p} - \frac{1}{p^2} \\
 &\Leftrightarrow \frac{2(1-p)}{p^2} - \frac{1}{p^2} + \frac{p}{p^2} \\
 &\Leftrightarrow \frac{2(1-p) + p - 1}{p^2} = \frac{2 - 2p + p - 1}{p^2} \\
 &\Leftrightarrow \frac{1-p}{p^2}
 \end{aligned}$$

□

b. Prove that $\Pr(X = k + m | X > m) = \Pr(X = k)$ for any $m, k \in \mathbb{N}$

$$\begin{aligned}
 \Pr(X = k + m | X > m) &= \frac{\Pr(X = k + m \cap X > m)}{\Pr(X > m)} \\
 \frac{\Pr(X = k + m | X > m) \Pr(X > m)}{\Pr(X > m)} &= \frac{\Pr(X > m | X = k + m) \Pr(X = k + m)}{\Pr(X > m)} \\
 \frac{1 \cdot (1-p)^{k+m-1} p}{(1-p)^m} &= (1-p)^{k-1} p
 \end{aligned}$$

Regarding $\Pr(X > m | X = k + m) = 1$ is trivial as $k, m \in \mathbb{N} \wedge k + m > m$

□.

Exercise 3:

- Given:
 $\Pr(D) = 0.01$, $\Pr(T|D) = \Pr(Tc|Dc) = 0.95$ accuracy of T
- Wanted: $\Pr(D|T)$
- Ansatz: Conditional Probability, Bayes rule and true positives + false positives for the denominator.

$$\begin{aligned}\Pr(D|T) &= \frac{\Pr(T|D)\Pr(D)}{\Pr(T)} = \frac{\Pr(T|D)\Pr(D)}{\Pr(T|D)\Pr(D) + \Pr(T|\overline{D})\Pr(\overline{D})} \\ &= \frac{0.95 \cdot 0.01}{0.95 \cdot 0.01 + 0.05 \cdot 0.99} = \frac{0.0095}{0.0095 + 0.0495} = \frac{0.0095}{0.059} \\ &= 0.161016949153 = 16.1016949153\%\end{aligned}$$