一、填空题

1.
$$\frac{361}{400}$$
; 2. $\frac{1}{4}$; 3. $e^{-2} - e^{-3}$; 4. 2; 5. 1; 6. $\frac{1}{3}$

二、解答题

- 1. 设 $A = \{ 考试及格 \}, B = \{ 按时交作业 \}$
 - (1) 由全概率公式得

$$P(B) = P(A)P(B|A) + P(\overline{A})P(B|\overline{A}) = 0.9 \times 0.85 + 0.1 \times 0.35 = 0.8$$

(2) 由贝叶斯公式得

$$P(A \mid B) = \frac{P(A)P(B \mid A)}{P(B)} = \frac{0.9 \times 0.85}{0.8} = \frac{153}{160} = 0.95625$$

2. (1)
$$P(X=k) = \frac{C_n^k p^k (1-p)^{n-k}}{1-(1-p)^n}, k=1, 2, ..., n.$$

(2)
$$E(X) = \sum_{k=1}^{n} k \frac{C_n^k p^k (1-p)^{n-k}}{1 - (1-p)^n} = \frac{1}{1 - (1-p)^n} \sum_{k=1}^{n} k C_n^k p^k (1-p)^{n-k}$$
$$= \frac{1}{1 - (1-p)^n} \sum_{k=1}^{n} k C_n^k p^k (1-p)^{n-k} = \frac{np}{1 - (1-p)^n}$$

3. 设 X_i 为第 i 位经手人加入的游戏币数目, i = 1, 2, ..., 200. X 为最终箱子中游戏币增加的数目.则:

$$P(X_i = k) = \frac{1}{5}, k = -2, -1, 0, 1, 2;$$
 且 $X_1, X_2, ..., X_{200}$ 相互独立,
$$E(X_i) = 0, D(X_i) = E(X_i^2) = \frac{1}{5} \times [(-2)^2 + (-1)^2 + 0 + 1^2 + 2^2] = 2$$

$$X = \sum_{i=1}^{200} X_i, E(X) = 0, D(X) = 400$$

由中心极限定理, X 近似服从 N(0, 400), 于是

$$P(|X| \le 10) = 2\Phi\left(\frac{10}{\sqrt{400}}\right) - 1 = 2\Phi(0.5) - 1 = 2 \times 0.6915 - 1 = 0.383$$

4. (1)
$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^x 15xy^2 dy, & 0 < x < 1 \\ 0, & 其他 \end{cases} = \begin{cases} 5x^4, & 0 < x < 1 \\ 0, & 其他 \end{cases}$$

(2) 当 0 < x < 1 时

$$f_{Y|X}(y|x) = \frac{f(x,y)}{f_X(x)} = \begin{cases} \frac{15xy^2}{5x^4}, & 0 < y < x \\ 0, & \text{其他} \end{cases} = \begin{cases} \frac{3y^2}{x^3}, & 0 < y < x \\ 0, & \text{其他} \end{cases}$$

$$f_{Y|X}\left(y \left| \frac{1}{2} \right) = \begin{cases} 24y^2, & 0 < y < \frac{1}{2} \\ 0, & \text{#ide} \end{cases}, \quad P\left(Y < \frac{1}{4} \middle| X = \frac{1}{2} \right) = \int_0^{\frac{1}{4}} 24y^2 dy = \frac{1}{8}$$

(3) Z 的取值范围为(0, 1). 当 z < 0 时, $F_Z(z) = 0$; 当 $z \ge 1$ 时, $F_Z(z) = 1$. 当 $0 \le z < 1$ 时,

$$F_Z(z) = P(Z \le z) = P\left(\frac{Y}{X} \le z\right) = \int_0^1 dx \int_0^{zx} 15xy^2 dy = \int_0^1 5z^3 x^4 dx = z^3$$

于是

$$f_Z(z) = F_Z'(z) = \begin{cases} 3z^2, & 0 < z < 1 \\ 0, & 其他 \end{cases}$$

5. X的密度函数为 $f(x) = \begin{cases} e^{-x}, & x > 0 \\ 0, & x \le 0 \end{cases}$, $Y = \min\{X, 1\}$.

$$E(Y) = \int_0^1 x e^{-x} dx + \int_1^{+\infty} e^{-x} dx = 1 - e^{-1};$$

$$E(Y^2) = \int_0^1 x^2 e^{-x} dx + \int_1^{+\infty} e^{-x} dx = 2 - 4e^{-1};$$

$$D(Y) = E(Y^2) - [E(Y)]^2 = 1 - 2e^{-1} - e^{-2}$$

$$\stackrel{\underline{}}{=} z \ge 0 \text{ iff}, F_{Z}(z) = P(Z_1 \le z) = P(|X_1 - \mu| \le z) = 2\Phi\left(\frac{z}{\sigma}\right) - 1,$$

$$f_Z(z) = F_Z'(z) = \frac{2}{\sigma} \varphi\left(\frac{z}{\sigma}\right) = \frac{2}{\sqrt{2\pi\sigma}} e^{-\frac{z^2}{2\sigma^2}}$$

$$\text{ BD } \quad f_Z(z) = \begin{cases} \frac{2}{\sqrt{2\pi}\sigma} e^{-\frac{z^2}{2\sigma^2}}, & z \ge 0\\ 0, & z < 0 \end{cases}$$

(2)
$$L(\sigma) = \prod_{i=1}^{n} \frac{2}{\sqrt{2\pi}\sigma} e^{-\frac{z_i^2}{2\sigma^2}} = \frac{2^n}{(\sqrt{2\pi}\sigma)^n} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^{n} z_i^2}$$

$$\ln L(\sigma) = n \ln 2 - \frac{n}{2} \ln(2\pi) - n \ln \sigma - \frac{1}{2\sigma^2} \sum_{i=1}^{n} z_i^2$$

$$\frac{d \ln L(\sigma)}{d \sigma} = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} z_i^2 = 0, \quad \text{MFR} \hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^{n} z_i^2}$$

故
$$\sigma$$
的极大似然估计量为 $\hat{\sigma} = \sqrt{\frac{1}{n}\sum_{i=1}^{n}Z_{i}^{2}}$.

7. (1)
$$\frac{\overline{X} - \mu}{S/\sqrt{n}} = \frac{4(\overline{X} - \mu)}{S} \sim t(15)$$
, $P\left(-t_{0.975}(15) < \frac{4(\overline{X} - \mu)}{S} < t_{0.975}(15)\right) = 0.95$

$$\mathbb{EP}\left(\overline{X} - \frac{S}{4}t_{0.975}(15) < \mu < \overline{X} + \frac{S}{4}t_{0.975}(15)\right) = 0.95$$

故μ的置信度为95%的置信区间为

$$\left(\overline{x} - \frac{s}{4} \times t_{0.975}(15), \overline{x} - \frac{s}{4} \times t_{0.975}(15)\right) = \left(5.2 - \frac{1.6}{4} \times 2.1315, 5.2 + \frac{1.6}{4} \times 2.1315\right)$$

$$= (4.3474, 6.0526)$$

(2) 检验统计量为
$$\chi^2 = \frac{15S^2}{\sigma_0^2} = \frac{15S^2}{2.5} = 6S^2$$
,

检验拒绝域为 $W = \{\chi^2 = 6S^2 > \chi^2_{0.95}(15)\} = \{\chi^2 = 6S^2 > 24.996\}$ 带入 $S^2 = 2.56$,计算得 $6S^2 = 6 \times 2.56 = 15.36 < 24.996$ 故不能拒绝原假设,不能认为 σ 显著大于 2.5.