A Measure of Random Utility Model Violations and Econometric Test

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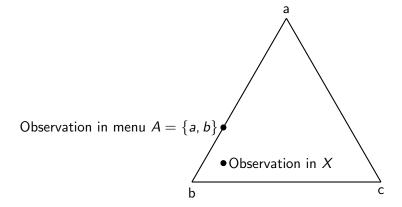
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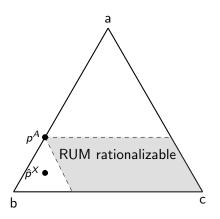
2024 Economic Science Association (ESA) North American Meeting at Columbus

- Random Utility Model (RUM) is popular model of individual/aggregate choice
- Few methods to evaluate violation of model

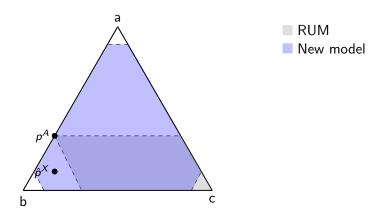
• Consider a set of alternatives $X = \{a, b, c\}$



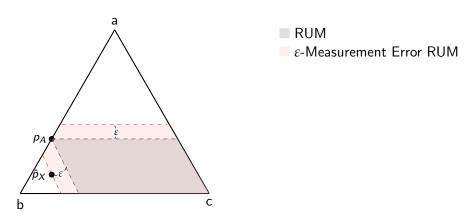
Choice probabilities must not increase as the menu enlarges
 ⇒ Data violates RUM (Regularity Axiom)



- Could rationalize data with a larger model
- E.g. Rational Inattention, Consideration Set models, etc.



- What if the data has measurement error?
- Extend RUM to accommodate measurement error ε-Measurement Error RUM (ε-MERUM)



Contribution

• Generate method to quantify measurement error in RUM (MERUM)

• Minimum distance (ε^*) gives ε^* -MERUM rationalization

 \bullet ε^* -MERUM gives statistical test via multinomial distributions

Stochastic choice primitives

- X: a finite nonempty set of alternatives
 - e.g. $X = \{a, b, c\}$
- $A \subseteq 2^X \setminus \{\emptyset\}$: a collection of menus
 - e.g. $A = \{\{a, b\}, X\}$
- $\mathcal{O} = \{(a, A) \in X \times \mathcal{A} | A \in \mathcal{A}, a \in A\}$: the collection of all pairs of a menu and an alternative in the menu
 - e.g. $\mathcal{O} = \{(a, \{a, b\}), (b, \{a, b\}), (a, X), (b, X), (c, X)\}$
- $m{\circ}$ $p:\mathcal{O}
 ightarrow [0,1]$ is a stochastic choice function
 - $\forall A \in \mathcal{A}, \ \sum_{a \in A} p(a, A) = 1$

Rationalization by RUM

• \mathcal{R} : the set of all strict preference ordering on X

• e.g.
$$\mathcal{R} = \{\succ_{abc}, \succ_{acb}, \succ_{bac}, \succ_{bca}, \succ_{cab}, \succ_{cba}\}$$

- ullet Random Utility Model (RUM): probability distribution $\mu \in \Delta(\mathcal{R})$
 - Assigns frequency of each preference ordering occurrence
- e.g. p: **RUM rationalizable** if we find μ s.t.

$$\begin{split} p(a,\{a,b\}) &= \mu(\succ_{abc}) + \mu(\succ_{acb}) + \mu(\succ_{cab}) \\ p(b,\{a,b\}) &= \mu(\succ_{bac}) + \mu(\succ_{bca}) + \mu(\succ_{cba}) \\ p(a,X) &= \mu(\succ_{abc}) + \mu(\succ_{acb}) \\ p(b,X) &= \mu(\succ_{bac}) + \mu(\succ_{bca}) \\ p(c,X) &= \mu(\succ_{cab}) + \mu(\succ_{cba}) \end{split}$$

Rationalization by RUM

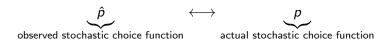
$$\underbrace{\begin{bmatrix} p(a,\{a,b\}) \\ p(b,\{a,b\}) \\ p(a,\{a,b,c\}) \\ p(b,\{a,b,c\}) \\ p(c,\{a,b,c\}) \end{bmatrix}}_{=:p_{(|\mathcal{O}|\times 1)}} = \underbrace{\begin{bmatrix} 1 & 1 & 0 & 0 & 1 & 0 \\ 0 & 0 & 1 & 1 & 0 & 1 \\ 1 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 1 \end{bmatrix}}_{=:M_{(|\mathcal{O}|\times|\mathcal{R}|)}} \underbrace{\begin{bmatrix} \mu(\succ_{abc}) \\ \mu(\succ_{acb}) \\ \mu(\succ_{bca}) \\ \mu(\succ_{cab}) \\ \mu(\succ_{cba}) \\ \mu(\succ_{cba}) \end{bmatrix}}_{=:\mu_{(|\mathcal{R}|\times 1)}}$$

Definition

A stochastic choice function p is **RUM rationalizable** if there exists a $\mu \in \Delta(\mathcal{R})$ s.t. $p = M\mu$.

Extension of RUM with errors

To accommodate the measurement error,



• \hat{p} can be described as p containing error: for each $(a, A) \in \mathcal{O}$,

$$p(a, A) - \varepsilon \le \hat{p}(a, A) \le p(a, A) + \varepsilon$$

where $\varepsilon \in [0,1]$ is an error uniform across $(a,A) \in \mathcal{O}$

Extension of RUM with errors

• Assume that true data is RUM rationalizable: $p = M\mu$

$$\begin{split} & \rho(a,A) - \varepsilon \leq \hat{p}(a,A) \leq \rho(a,A) + \varepsilon \\ \text{plug-in:} & & (M\mu)_{(a,A)} - \varepsilon \leq \hat{p}(a,A) \leq (M\mu)_{(a,A)} + \varepsilon \\ \text{rearrange:} & & \hat{p}(a,A) - \varepsilon \leq (M\mu)_{(a,A)} \leq \hat{p}(a,A) + \varepsilon \\ \text{stack:} & & \underbrace{\begin{bmatrix} \hat{p}(a_1,A_1) \\ \vdots \\ \hat{p}(a_{|\mathcal{O}|},A_{|\mathcal{O}|}) \end{bmatrix}}_{=:\hat{p}_{(|\mathcal{O}|\times 1)}} - \varepsilon \cdot \underbrace{\begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}}_{=:1_{|\mathcal{O}|,1}} \leq M\mu \leq \underbrace{\begin{bmatrix} \hat{p}(a_1,A_1) \\ \vdots \\ \hat{p}(a_{|\mathcal{O}|},A_{|\mathcal{O}|}) \end{bmatrix}}_{=:\hat{p}_{(|\mathcal{O}|\times 1)}} + \varepsilon \cdot \underbrace{\begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}}_{=:1_{|\mathcal{O}|,1}} \\ \Leftrightarrow & & \hat{p} - \varepsilon \cdot \mathbf{1}_{|\mathcal{O}|,1} \leq M\mu \leq \hat{p} + \varepsilon \cdot \mathbf{1}_{|\mathcal{O}|,1} \end{split}$$

• $\exists (\varepsilon, \mu)$ that satisfy (1): ε -MERUM rationalization of $\hat{\rho}$!

Main Proposition

Let \hat{p} : an observed s.c.f. and ε^* , where

$$\begin{split} (\varepsilon^*, \mu^*) &= \arg \min_{\mu \in \Delta(\mathcal{R}), \varepsilon \in [0,1]} \varepsilon \\ \text{s.t. } \hat{\rho} &- \varepsilon \cdot \mathbf{1}_{|\mathcal{O}|,1} \leq M \mu \leq \hat{\rho} + \varepsilon \cdot \mathbf{1}_{|\mathcal{O}|,1} \end{split}$$

Then,

- **①** ε^* always exists for any observed s.c.f. \hat{p} with $\varepsilon^* \in [0,1]$, and
- 2 $\varepsilon^* = 0 \Leftrightarrow \hat{p}$: RUM rationalizable

Building the statistical test

- Suppose \hat{p} is ε^* -MERUM rationalized with optimizer (ε^*, μ^*)
- Then, consider the following hypothesis test:

$$H_0: p = M\mu^*$$
 $H_1: p \neq M\mu^*$

- Observe that $p = M\mu^*$ is a collection of **multinomial** distribution
 - ▶ Focus on *certain menu* $A \in A$ where |A| = K
 - ▶ The choice probability of k-th alternative in A will be $p_k := p(a_k, A)$

Building the statistical test

$$H_0: p = M\mu^*$$
 $H_1: p \neq M\mu^*$

• Construct a confidence interval (CI) under the null (H_0) :

$$CI(\hat{p}_k^-(\alpha), \hat{p}_k^+(\alpha)) = P(\hat{p}_k^-(\alpha) \le p_k \le \hat{p}_k^+(\alpha), \forall k = 1, \dots, K)$$

• We have the representation (Sison and Glaz [1995]) of

$$\hat{\rho}_k^- = \hat{\rho}_k - \frac{c(\alpha)}{n}$$
 $\hat{\rho}_k^+ = \hat{\rho}_k + \frac{c(\alpha)}{n}$

we can **directly compare** ε^* with $\frac{c(\alpha)}{n}$, since

$$\hat{\rho}_k - \frac{c(\alpha)}{n} \le p_k \le \hat{\rho}_k + \frac{c(\alpha)}{n}$$
 $\hat{\rho}_k - \varepsilon^* \le p_k \le \hat{\rho}_k + \varepsilon^*$

Rationalization problem for simulation

• We consider the following DGP with marginal regularity:

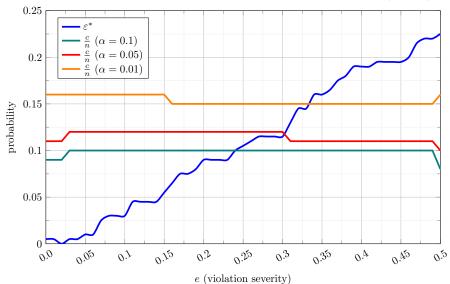
	а	Ь	С
$\{a,b\}$	0.5 − <i>e</i>	0.5 + e	•
$\{a,b,c\}$	0.5	0.5	0

Table: True data generating process (DGP), p

- Estimate \hat{p} through counting numbers: $\hat{p}(a, A) = \frac{n_a}{n}$
 - ▶ n: the number of choice drawn in total
 - \triangleright n_a : the number of alternative a is chosen

Simulation result of the comparison

Comparison of Consistency Measure and Confidence Bounds (n = 100)



Simulation result of the comparison

e n	20	30	50	100	200	500
0.0	0	0	0	0	0	0
0.1	0	0	0	0	0.006	0.14
0.2	0.002	0	0.002	0.056	0.422	0.994
0.3	0	0.008	0.052	0.474	0.966	1
0.4	0.008	0.056	0.328	0.97	1	1
0.5	0.028	0.304	0.742	1	1	1

Table: Frequency of ε^* rejection with $\alpha=0.01$ and 500 simulations

- Measure falls into the rejection region with the following properties
 - ▶ The more *severe* the violation (e), the more *frequent* the rejection
 - ► The *larger* the sample size (n), the more *frequent* the rejection (for $e \neq 0$)

Conclusion

Our results can be summarized as follows:

- The consistency measure achieved
 - rationalization with error as linear programming problem
- 2 The measure captures severity of RUM violation of data
 - stochastic version of CCEI
- The measure enters to the rejection region with nice properties
 - measure falls into rejection region as e and/or n increases

Thank you!