Statistical Simulation Assignment 1

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July 19, 2022

Q2. (b) Given: MGF of X is g(t) To find: MGF of Y=kX+d, k,d \in R; $M_Y(t) = \int_{-\infty}^{\infty} e^{yt} f_Y(y) dy =$

$$\int_{-\infty}^{\infty} e^{(ax+b)t} \frac{f_X(x)}{a} a dx$$

$$= e^{bt} \int_{-\infty}^{\infty} e^{axt} f_x(x) dx = e^{bt} M_x(at)$$

(d) Let X have a normal distribution with mean μ_x , variance σ_x^2 , and standard deviation σ_x . Let Y have a normal distribution with mean μ_y , variance σ_y^2 , and standard deviation σ_y . If X and Y are independent, then X-Y will follow a normal distribution with mean $\mu_x-\mu_y$, variance $\sigma_x^2+\sigma_y^2$, and standard deviation $\sqrt{\sigma_x^2+\sigma_y^2}$

(e) Part (i):
$$M(t) = \frac{1}{6}e^{3t} + \frac{1}{2}e^{5t} + \frac{1}{3}e^{7t}$$

$$M_x(t) = E[e^{tx}] = \sum_{x \in S} P(X = x)e^{xt} - eq(1)$$

$$\begin{array}{c|cccc} x & p_x \\ 3 & \frac{1}{6} \\ 5 & \frac{1}{2} \\ 7 & \frac{1}{3} \end{array}$$

Comparing to eq. 1, we see that $S = \{3, 5, 7\}$, and for each for those x values, $P(X = x) = p_x$ from the table above.

Part (ii):
$$M_x(t) = \frac{5}{5-t}$$

Let X be a continuous random variable with an exponential distribution with parameter β for some $\beta \in R > 0$

Then the moment generating function M_X of X is given by: $M_x(t) = \frac{1}{1-\beta t}$

Our given equation can be written as $\frac{1}{1-5t}$, thus $\beta=\frac{1}{5}$

From the definition of the Exponential distribution, X has probability density function: $f_x(x) = \frac{1}{\beta}e^{\frac{-x}{\beta}}$ Thus $pdf f_x(x) = 5e^{-5x}$ or $X \sim Exponential(1/5)$

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 $\mathbf{Q4.}$ By looking at the support S of the DF of a RV that we can determine whether it is continuous or discrete. If the support is an interval then the DF is continuous, if the support is a set of discrete values then the DF is discrete.

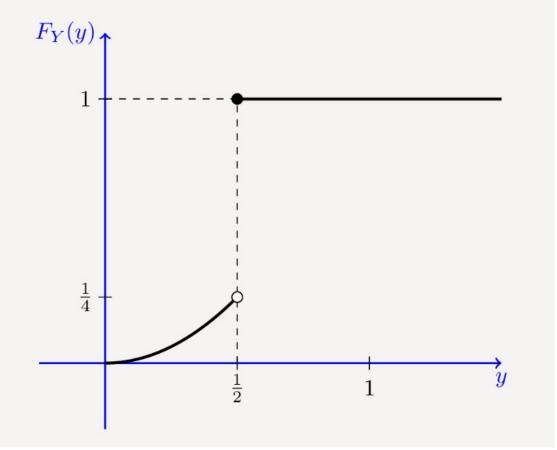
Yes, a RV can be neither discrete or continuous. Such RVs are called mixed random variables. Example:

Let X be a continuous random variable with the following PDF:

$$f_X(x) = egin{cases} 2x & 0 \leq x \leq 1 \ 0 & ext{otherwise} \end{cases}$$

Let also

$$Y=g(X)= egin{cases} X & 0 \leq X \leq rac{1}{2} \ rac{1}{2} & X > rac{1}{2} \end{cases}$$



The CDF of Y has a continuous part and a discrete part.

Q5. DF,mean,variance of following distributions:

(a)
$$f(x) = \lambda e^{-\lambda x}; \forall x \geq 0 \text{ or } X \sim Exponential(\frac{1}{\lambda})$$
 - Valid pdf

DF =
$$F(x) = P(X \le x)$$

= $\int_0^\infty f(w) dw = \int_0 \lambda e^{-\lambda w} dw = 1 - e^{-\lambda x} x > 0$

Mean =
$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx$$

 $=\frac{1}{\lambda}$

$$\begin{split} &=\lambda\left[\left|\frac{-xe^{-\lambda x}}{\lambda}\right|_0^\infty+\frac{1}{\lambda}\int_0^\infty e^{-\lambda x}dx\right]\\ &=\lambda\left[0+\frac{1}{\lambda}\frac{-e^{-\lambda x}}{\lambda}\right]_0^\infty\\ &=\lambda\frac{1}{\lambda^2} \end{split}$$

Hence, the mean of the exponential distribution is $\frac{1}{\lambda}$.

For variance:
$$E[X^2] = \int_0^\infty x^2 \lambda e^{-\lambda x} dx = \frac{2}{\lambda^2}$$

Variance =
$$E[X^2] - (E[X])^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}$$

(b)
$$f(x) = \lambda^x \frac{e^{-\lambda}}{x!}; \forall x \in N_0 \text{ or } X \sim Poisson(\lambda)$$
 - Valid pdf For a poisson distribution, Mean and Variance both are equal to λ .

(d)
$$f(x) = \frac{1}{\pi} \frac{1}{1+x^2}; \forall x \in (-\infty, \infty) \text{ or } X \sim Cauchy(0, 1)$$
 - Valid pdf

$$\int_{-\infty}^{\infty} f_X(x) \, dx = \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{1}{1+y^2} \, dy = \frac{2}{\pi} \int_{0}^{\infty} \frac{1}{1+y^2} \, dy = \frac{2}{\pi} \tan^{-1}(y) \mid_{0}^{\infty} = 1.$$

Mean and variance do not exist for a cauchy distribution.