

# Probability Theory

## Exercise Sheet 10

**Exercise 10.1** (The generalized Borel-Cantelli lemma)

Consider  $(\Omega, \mathcal{F}, P)$  with filtration  $\{\mathcal{F}_n\}_{n \geq 0}$ , and let  $A_n \in \mathcal{F}_n$ ,  $n \geq 1$ , be a sequence of events. Show that, up to a  $P$ -nullset,

$$\limsup_{n \rightarrow \infty} A_n = \left\{ \sum_{n \geq 1} P[A_n | \mathcal{F}_{n-1}] = \infty \right\}.$$

**Hint:** Use Exercise 9.3.

**Exercise 10.2** Consider a Galton-Watson process (see p. 97 of the lecture notes)  $Z_n$ ,  $n \geq 0$ , with offspring distribution  $\nu = \text{Bin}(2, p)$ ,  $p \in [0, 1]$ . We are interested in the probability  $\vartheta(p) = P[Z_n > 0, \forall n \geq 0]$  that the population does not go extinct. Show that

$$\vartheta(p) = \begin{cases} 0 & \text{if } 0 \leq p \leq 1/2; \\ \frac{2p-1}{p^2} & \text{if } 1/2 < p \leq 1. \end{cases}$$

**Hint:** One way to prove this is to use the results for the various cases (subcritical, critical, supercritical) from Section 3.5 A), pp. 97-101 of the lecture notes.

**Exercise 10.3** (Probabilistic solution to the discrete Dirichlet problem)

Let  $A \subseteq \mathbb{Z}^d$  be finite,  $f : \mathbb{Z}^d \setminus A \rightarrow \mathbb{R}$  any function, and  $(S_n)_{n \in \mathbb{N}}$  a simple random walk on  $\mathbb{Z}^d$  with starting point  $S_0 = 0$ . For  $x \in \mathbb{Z}^d$  let  $T_x := \inf\{n \geq 0; |x + S_n \notin A\}$ . Finally, let  $\mathcal{F}_n := \sigma(S_0, \dots, S_n)$  and  $g(x) := E[f(x + S_{T_x})]$ .

(a) Show that  $T_x < \infty$   $P$ -a.s. Thus  $f(x + S_{T_x})$  exists a.s.

**Hint:** Use Exercise 9.3.

(b) Show that  $g$  solves the discrete Dirichlet problem on  $A$  with boundary condition  $f$ , i.e.,

$$g(x) = \begin{cases} f(x) & \text{if } x \in \mathbb{Z}^d \setminus A \\ \frac{1}{2d} \sum_{\substack{\|y-x\|=1 \\ y \in \mathbb{Z}^d}} g(y) & \text{if } x \in A. \end{cases}$$

(c) Show that  $E[f(x + S_{T_x}) | \mathcal{F}_1] = g(x + S_{T_x \wedge 1})$   $P$ -a.s.

**Submission:** until 14:15, Dec 03., during exercise class or in the tray outside of HG G 53.

**Office hours (Präsenz):** Mon. and Thu., 12:00-13:00 in HG G 32.6.

**Class assignment:**

Students	Time & Date	Room	Assistant
Afa-Fül	Tue 13-14	HG F 26.5	Angelo Abächerli
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**Solution 10.1** We define  $X_0 := 0$ ,  $X_n := \sum_{m=1}^n (1_{A_m} - P[A_m|\mathcal{F}_{m-1}])$ ,  $n \geq 1$ . Then  $X_n$  is an  $\mathcal{F}_n$ -martingale, since

$$E[X_{n+1} - X_n | \mathcal{F}_n] = E[1_{A_{n+1}} - P[A_{n+1} | \mathcal{F}_n] | \mathcal{F}_n] = 0.$$

Furthermore  $|X_{n+1} - X_n| \leq 2$ . We apply the result of Exercise 9.3 to obtain  $P[C \cup D] = 1$ . Note that:

- $\sum_{n \geq 1} 1_{A_n} = \infty \iff \sum_{n \geq 1} P[A_n | \mathcal{F}_{n-1}] = \infty$  on  $C$ .
- $\sum_{n \geq 1} 1_{A_n} = \infty$  and  $\sum_{n \geq 1} P[A_n | \mathcal{F}_{n-1}] = \infty$  on  $D$ .

Using that  $P[C \cup D] = 1$ , we get that for an event  $N$  with  $P[N] = 0$ ,

$$\left\{ \sum_{n \geq 1} 1_{A_n} = \infty \right\} \cap N^c = \left\{ \sum_{n \geq 1} P[A_n | \mathcal{F}_{n-1}] = \infty \right\} \cap N^c.$$

Finally, the claim follows since

$$\limsup_{n \rightarrow \infty} A_n = \left\{ \sum_{n \geq 1} 1_{A_n} = \infty \right\}.$$

**Solution 10.2** From (3.5.3), p. 98 of the lecture notes we calculate  $m$ :

$$m = 2p \begin{cases} < 1 & \text{if } p \in [0, \frac{1}{2}), \\ = 1 & \text{if } p = \frac{1}{2}, \\ > 1 & \text{if } p \in (\frac{1}{2}, 1]. \end{cases}$$

Thus, if  $p \in [0, \frac{1}{2})$ , our Galton-Watson process is subcritical, if  $p = \frac{1}{2}$  it is critical, and if  $p \in (\frac{1}{2}, 1]$  it is supercritical.

For a subcritical Galton-Watson process, we have  $P[Z_n = 0 \text{ eventually}] = 1$  by (3.5.7), p. 99, and by (3.5.10), p. 100 of the lecture notes also for a critical process. Hence,

$$\vartheta(p) = 0 \quad \forall p \in [0, 1/2].$$

In the supercritical case, we have, by (3.5.13), p. 101 of the lecture notes,

$$P[Z_n = 0 \text{ eventually}] = \varrho \in [0, 1),$$

where  $\varrho$  is the unique solution to  $\varrho = \varphi(\varrho)$  in  $[0, 1)$ , and let  $X$  be a random variable with distribution  $\nu$ , we have

$$\varphi(z) = E[z^X] = \sum_{k=0}^2 P[X = k] z^k = (1-p)^2 + 2p(1-p)z + p^2 z^2,$$

(see (3.5.11), p. 100 of the lecture notes, and the explanations right below it). Solving the quadratic equation

$$\varphi(z) = ((1-p) + pz)^2 = z$$

for  $z$ , we obtain the solutions  $z = 1$  and  $z = \frac{(1-p)^2}{p^2}$ . Thus, the unique solution to  $\varphi(\varrho) = \varrho$  in  $[0, 1]$  is

$$\varrho = \frac{(1-p)^2}{p^2},$$

from which it follows that

$$\vartheta(p) = 1 - P[Z_n = 0 \text{ eventually}] = 1 - \varrho = 1 - \frac{(1-p)^2}{p^2} = \frac{2p-1}{p^2},$$

for  $p \in (1/2, 1]$ .

**Solution 10.3** Let  $S_n = \sum_{m=1}^n X_m$  ( $S_0 = 0$ ) be the simple random walk on  $\mathbb{Z}^d$  such that  $X_n$ ,  $n \geq 1$  are i.i.d.  $\mathbb{Z}^d$  valued random variable with  $P[X_1 = e] = 1/(2d)$  for any  $e \in \mathbb{R}^d$ ,  $\|e\| = 1$  (i.e.,  $e = e_i$  or  $e = -e_i$  for  $e_i = (0, \dots, 0, 1, 0, \dots, 0)$  with 1 on the  $i$ th entry,  $i = 1, \dots, d$ ). We can easily show that  $S_n$  is an  $\mathcal{F}_n$  martingale taking values in  $\mathbb{Z}^d$  and  $S_n^j$ , the  $j$ th coordinate of  $S_n$ , is also an  $\mathcal{F}_n$  real-valued martingale for each  $j = 1, \dots, d$ .

(a) Consider the martingale

$$M_n = \sum_{i=1}^d S_n^i,$$

with bounded increments

$$M_n - M_{n-1} = \sum_{i=1}^d (S_n^i - S_{n-1}^i) \in \{-1, 1\}.$$

By Exercise 9.3, it thus holds that

$$P[\{\lim M_n \in \mathbb{Z} \text{ exists}\} \cup \{\liminf M_n = -\infty \text{ and } \limsup M_n = \infty\}] = 1.$$

However, since  $|M_n - M_{n-1}| = 1$ ,  $M_n$  cannot converge, from which it follows that  $\limsup M_n = \infty$   $P$ -a.s. and  $\liminf M_n = -\infty$   $P$ -a.s. This, in turn, implies that  $S_n$  must exit the finite set  $A$ , since  $\max_{(a_1, \dots, a_d) \in A} \sum_{i=1}^d |a_i| < \infty$ .

(b) If  $x \notin A$ , then  $T_x = 0 \Rightarrow f(x + S_{T_x}) = f(x)$ .

If  $x \in A$ , then

$$\begin{aligned} E[f(x + S_{T_x})] &= \sum_{\|e\|=1} E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] \\ &= \frac{1}{2d} \sum_{\|e\|=1} E\left[f\left(x + e + S_{T_{x+e}}\right)\right]. \end{aligned}$$

To see the last equality, we note that if  $x + e \notin A$ , then in view of the definition of  $T_{x+e}$  (resp.  $T_x$ ) we have  $T_{x+e} = 0$  (resp.  $T_x = 1$ ) and in this case it holds that  $E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] = f(x + e)P[S_1 = e] = \frac{f(x+e)}{2d} = E\left[f\left(x + e + S_{T_{x+e}}\right)\right]$ . On

the other hand, if  $x + e \in A$ , then we must have  $T_x \geq 2$  and  $T_{x+e} \geq 1$ . In this case we have

$$\begin{aligned} E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] &= \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{\{S_1=e\} \cap \{T_x=n\}}] \\ &= \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{\{S_1=e\} \cap B_n}], \end{aligned}$$

where  $B_n = \{\forall 2 \leq m < n, x + e + (S_m - S_1) \in A; x + e + S_n - S_1 \notin A\}$ . But since  $S_m - S_1 = \sum_{j=2}^m X_j$  is independent of  $S_1 = X_1$  for all  $m \geq 2$ , the event  $B_n$  is independent of  $\{S_1 = e\}$ . Moreover, it is easy to see that  $x + e + S_m - S_1 = x + e + \sum_{j=2}^m X_j$ ,  $m \geq 1$  is the simple random walk started from  $x + e$ , which implies that  $B_n = \{T_{x+e} = n - 1\}$  for all  $n \geq 2$  and  $x + e + S_n - S_1 = x + e + S_{T_{x+e}}$  on  $B_n$ . Hence, we can deduce that

$$\begin{aligned} \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{\{S_1=e\} \cap B_n}] &= \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{B_n}] P[S_1 = e] \\ &= \frac{1}{2d} E[f(x + e + S_{T_{x+e}}) \mathbf{1}_{\{T_{x+e} \geq 1\}}]. \end{aligned}$$

Now we complete our proof by writing  $y = x + e$  for some  $e$  with  $\|e\| = 1$ .

(c) In case  $x \notin A$  the statement is trivial.

If  $x \in A$ , then  $S_{T_x \wedge 1} = S_1$ , whence it follows that

$$E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] = \frac{1}{2d} E[f(x + e + S_{T_{x+e}})] = \frac{1}{2d} g(x + e)$$

and

$$E[g(x + S_1) \mathbf{1}_{S_1=e}] = \frac{1}{2d} E[g(x + e)] = \frac{1}{2d} g(x + e).$$

This concludes the proof.