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# **Probability and Applied Statistics**

# Formula Sheet #2

## Chapter 3

#### **Poisson Distribution:**

A random variable Y is said to have a Poisson probability distribution if and only if

$$p(y) = \frac{\lambda^{y}}{y!}e^{-\lambda}, \quad y = 0, 1, 2, ..., \lambda > 0.$$

If Y is a random variable possessing a Poisson distribution with parameter  $\lambda$ , then

$$\mu = E(Y) = \lambda \text{ and } \sigma^2 = V(Y) = \lambda.$$

By definition,

$$E(Y) = \sum_{y} yp(y) = \sum_{y=0}^{\infty} y \frac{\lambda^{y} e^{-\lambda}}{y!}.$$

### Tchebysheff's Theorem:

Tchebysheff's Theorem – Let Y be a random variable with mean  $\mu$  and finite variance  $\sigma^2$ . Then, for any constant k > 0,

$$P(|Y - \mu| < k\sigma) \ge 1 - \frac{1}{k^2}$$
 or  $P(|Y - \mu| \ge k\sigma) \le \frac{1}{k^2}$ .

## Chapter 4

#### The Probability Distribution for a Continuous Random Variable:

Let Y denote any random variable. The distribution function of Y, denoted by F(y), is such that  $F(y) = P(Y \le y) for - \infty < y < \infty$ .

#### **Properties of a Distribution Function:**

If F(y) is a distribution function, then

- 1.  $F(-\infty) \equiv \overline{\lim_{y \to -\infty} F(y) = 0}$ .
- 2.  $F(\infty) \equiv \lim_{y \to \infty} F(y) = 1$ .
- 3. F(y) is a nondecreasing function of y. [If  $y_1$  and  $y_2$  are any values such that  $y_1 < y_2$ , then  $F(y_1) \le F(y_2)$ .]

A random variable Y with distribution function F(y) is said to be continuous if F(y) is continuous for  $-\infty < y < \infty$ .

Let F(y) be the distribution function for a continuous random variable Y. Then f(y) given by

$$f(y) = \frac{dF(y)}{dy} = F'(y)$$

whenever the derivative exists, is called the probability density function for the random variable Y.

It follows that F(y) can be written as

$$F(y) = \int_{-\infty}^{y} f(t)dt.$$

Properties of a Density Function – If f(y) is a density function for a continuous random variable, then

1. 
$$f(y) \ge 0$$
 for all  $y, -\infty < y < \infty$ .

$$2. \int_{-\infty}^{\infty} f(y) dy = 1.$$

If the random variable Y has density function f(y) and a < b, then the probability that Y falls in the interval [a, b] is

$$P(a \le Y \le b) = \int_{a}^{b} f(y)dy.$$

### **Expected Value for Continuous Random Variables:**

The expected value of a continuous random variable Y is

$$E(Y) = \int_{-\infty}^{\infty} y f(y) dy,$$

provided that the integral exists.

Let g(Y) be a function of Y, then the expected value of g(Y) is given by

$$E[g(Y)] = \int_{-\infty}^{\infty} g(y)f(y)dy,$$

provided that the integral exists.

Let c be a constant and let  $g(Y), g_1(Y), g_2(Y), \dots, g_k(Y)$  be functions of a continuous random variable Y. Then the following results hold:

1. 
$$E(c) = c$$
.

2. 
$$E[cg(Y)] = cE[g(Y)].$$

3. 
$$E[g_1(Y) + g_2(Y) + \dots + g_k(Y)] = E[g_1(Y)] + E[g_2(Y)] + \dots + E[g_k(Y)]$$

### The Uniform Probability Distribution:

If  $\theta_1 < \theta_2$ , a random variable Y is said to have a continuous uniform probability distribution on the interval  $(\theta_1, \theta_2)$  if and only if the density function of Y is

$$f(y) = \begin{cases} \frac{1}{\theta_1 - \theta_2}, \theta_1 \le y \le \theta_2. \\ 0, & elsewhere \end{cases}$$

If  $\theta_1 < \theta_2$  and Y is a random variable uniformly distributed on the interval  $(\theta_1, \theta_2)$ , then

$$\mu = E(Y) = \frac{\theta_1 + \theta_2}{2} \text{ and } \sigma^2 = V(Y) = \frac{(\theta_2 - \theta_1)^2}{12}$$

#### The Gamma Probability Distribution:

A random variable Y is said to have a gamma distribution with parameters  $\alpha>0$  and  $\beta>0$  if and only if the density function of Y is

$$f(y) = \begin{cases} \frac{y^{\alpha - 1}e^{-\frac{y}{\beta}}}{\beta^{\alpha}\Gamma(\alpha)}, 0 \le y \le \infty \\ 0, & elsewhere \end{cases}$$

where

$$\Gamma(\alpha) = \int_0^\infty y^{\alpha - 1} e^{-y} dy.$$

If Y has a gamma distribution with parameters  $\alpha$  and  $\beta$ , then

$$\mu = E(Y) = \alpha \beta$$
 and  $\sigma^2 = V(Y) = \alpha \beta^2$ 

By definition,

$$E(Y) = \int_{-\infty}^{\infty} y f(y) dy = \int_{0}^{\infty} y \left( \frac{y^{\alpha - 1} e^{-y/\beta}}{\beta^{\alpha} \Gamma(\alpha)} \right) dy \text{ and } V(Y) = E[Y^{2}] - [E(Y)]^{2}$$

A random variable Y is said to have an exponential distribution with parameter  $\beta > 0$  if and only if the density function of Y is

$$f(y) = \begin{cases} \frac{1}{\beta} e^{-y/\beta}, & 0 \le y \le \infty \\ 0 & elsewhere \end{cases}$$

If Y is an exponential random variable with parameter  $\beta$ , then

$$\mu = E(Y) = \beta$$
 and  $\sigma^2 = V(Y) = \beta^2$ 

with  $\alpha = 1$ 

## Chapter 5

# **Bivariate and Multivariate Probability Distributions:**

Let  $Y_1$  and  $Y_2$  be discrete random variables. The joint probability function for  $Y_1$  and  $Y_2$  is given by

$$p(y_1, y_2) = P(Y_1 = y_1, Y_2 = y_2), \quad -\infty < y_1 < \infty, -\infty < y_2 < \infty$$

If  $Y_1$  and  $Y_2$  are discrete random variables with joint probability function  $p(y_1, y_2)$ , then

- 1.  $p(y_1, y_2) \ge 0$  for all  $y_1, y_2$
- 2.  $\sum_{y_1,y_2} p(y_1,y_2) = 1$ , where the sum is over the values  $(y_1,y_2)$  that are assigned nonzero probabilities.

For any random variables  $Y_1$  and  $Y_2$ , the joint distribution function  $F(y_1, y_2)$  is

$$F(y_1,y_2) = P(Y_1 \leq y_1,Y_2 \leq y_2), \qquad -\infty < y_1 < \infty, -\infty < y_2 < \infty$$

Let  $Y_1$  and  $Y_2$  be continuous random variables with joint distribution function  $F(y_1, y_2)$ . If there exists a nonnegative function  $f(y_1, y_2)$ , such that

$$F(y_1, y_2) = \int_{-\infty}^{y_1} \int_{-\infty}^{y_2} f(t_1, t_2) dt_2 dt_1$$

for all  $-\infty < y_1 < \infty$ ,  $-\infty < y_2 < \infty$ , then  $Y_1$  and  $Y_2$  are said to be jointly continuous random variables. The function  $f(y_1, y_2)$  is called the joint probability density function

If  $Y_1$  and  $Y_2$  are random variables with joint distribution function  $F(y_1, y_2)$ , then

- 1.  $F(-\infty, -\infty) = F(-\infty, y_2) = F(y_1, -\infty) = 0$ .
- 2.  $F(\infty, \infty) = 1$ .
- 3. If  $y_1^* \ge y_1$  and  $y_2^* \ge y_2$ , then  $F(y_1^*, y_2^*) F(y_1^*, y_2) F(y_1, y_2^*) + F(y_1, y_2) \ge 0$

If  $Y_1$  and  $Y_2$  are jointly continuous random variables with a joint density function given by  $f(y_1, y_2)$ , then

- 1.  $f(y_1, y_2) \ge 0$  for all  $y_1, y_2$
- 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(y_1, y_2) dy_1 dy_2 = 1$ .

# **Marginal and Conditional Probability Distributions:**

Let  $Y_1$  and  $Y_2$  be jointly discrete random variables with probability function  $p(y_1, y_2)$ . Then the marginal probability functions of  $Y_1$  and  $Y_2$ , respectively, are given by

$$p_1(y_1) = \sum_{\text{all } y_2} p(y_1, y_2) \text{ and } p_2(y_2) \neq = \sum_{\text{all } y_1} p(y_1, y_2)$$

Let  $Y_1$  and  $Y_2$  be jointly continuous random variables with joint density function  $f(y_1, y_2)$ . Then the marginal density functions of  $Y_1$  and  $Y_2$ , respectively, are given by

$$f_1(y_1) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_2$$
 and  $f_2(y_2) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_1$ 

If  $Y_1$  and  $Y_2$  are jointly discrete random variables with joint probability function  $p(y_1, y_2)$  and marginal probability functions  $p_1(y_1)$  and  $p_2(y_2)$ , respectively, then the conditional discrete probability function of  $Y_1$  and  $Y_2$  is

$$p(y_1|y_2) = P(Y_1 = y_1|Y_2 = y_2) = \frac{P(Y_1 = y_1, Y_2 = y_2)}{P(Y_2 = y_2)} = \frac{p(y_1, y_2)}{p_2(y_2)}$$

provided that  $p_2(y_2) > 0$ .

If  $Y_1$  and  $Y_2$  are jointly continuous random variables with joint density function  $f(y_1, y_2)$ , then the conditional distribution of  $Y_1$  given  $Y_2 = y_2$  is

$$F(y_1|y_2) = P(Y_1 \le y_1|Y_2 = y_2).$$

Let  $Y_1$  and  $Y_2$  be jointly continuous random variables with join density  $f(y_1, y_2)$  and marginal densities  $f_1(y_1)$  and  $f_2(y_2)$  respectively. For any  $y_2$  such that  $f_2(y_2) > 0$ , the conditional density of  $Y_1$  given  $Y_2 = y_2$  is given by

$$f(y_1|y_2) = \frac{f(y_1, y_2)}{f_2(y_2)}$$

And, for  $y_1$  such that  $f_1(y_1) > 0$ , the conditional density of  $Y_2$  given  $Y_1 = y_1$  is given by

$$f(y_2|y_1) = \frac{f(y_1, y_2)}{f_1(y_1)}$$

### **Independent Random Variables:**

Let  $Y_1$  have distribution function  $F_1(y_1)$ ,  $Y_2$  have distribution function  $F_2(y_2)$ , and  $Y_1$  and  $Y_2$  have joint distribution function  $F(y_1, y_2)$ . Then  $Y_1$  and  $Y_2$  are said to be independent if and only if

$$F(y_1, y_2) = F_1(y_1)F_2(y_2)$$

For every pair of real numbers  $(y_1, y_2)$ .

If  $Y_1$  and  $Y_2$  are not independent, they are said to be dependent.

If  $Y_1$  and  $Y_2$  are discrete random variables with joint probability function  $p(y_1, y_2)$  and marginal probability functions  $p_1(y_1)$  and  $p_2(y_2)$ , respectively, then  $Y_1$  and  $Y_2$  are independent if and only if

$$p(y_1, y_2) = p_1(y_1)p_2(y_2)$$

For all pairs of real numbers  $(y_1, y_2)$ .

If  $Y_1$  and  $Y_2$  are continuous random variables with joint density function  $f(y_1, y_2)$  and marginal density functions  $f_1(y_1)$  and  $f_2(y_2)$ , respectively, then  $Y_1$  and  $Y_2$  are independent if and only if

$$f(y_1, y_2) = f_1(y_1)f_2(y_2)$$

For all pairs of real numbers  $(y_1, y_2)$ .

Let  $Y_1$  and  $Y_2$  have a joint density  $f(y_1,y_2)$  that is positive if and only if  $a \le y_1 \le b$  and  $c \le y_2 \le d$ , for constants a, b, c, and d; and  $f(y_1,y_2) = 0$  otherwise. Then  $Y_1$  and  $Y_2$  are independent random variables if and only if

$$f(y_1, y_2) = g(y_1)h(y_2)$$

where  $g(y_1)$  is a nonnegative function of  $y_1$  alone and  $h(y_2)$  is a nonnegative function of  $y_2$  alone.