

Computational Statistics & Machine Learning

Lecture 6

Markov chain Monte Carlo

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Overview

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Lecture Outline

Product Transition
Kernels

Gibbs Sampler

Mixture Transition
Kernels

Data
Augmentation

- ▶ Product Transition Kernels
- ▶ Gibbs Sampler
- ▶ Mixture Transition Kernels
- ▶ Data Augmentation

Lecture Outline

Product Transition Kernels

Gibbs Sampler

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Data Augmentation

Product Transition Kernels

- ▶ Simple illustrative examples are all very well - what if target density is multivariate?
- ▶ Breaking the vector up into sub-blocks is a good strategy to design proposals?
- ▶ Does this ensure convergence to correct target?
- ▶ Let $x = (x_1, x_2)$ where $x \in \mathcal{R}^d$, and each $x_i \in \mathcal{R}^{d_i}$
- ▶ Conditional transition kernel $P_1(x_1, dy_1|x_2)$ with invariant density $\pi_{1|2}(\cdot|x_2)$ for fixed value of x_2
- ▶ Conditional transition kernel $P_2(x_2, dy_2|x_1)$ with invariant density $\pi_{2|1}(\cdot|x_1)$ for fixed value of x_1
- ▶ Standard conditions apply
$$\pi_{1|2}^*(dy_1|x_2) = \int P_1(x_1, dy_1|x_2) \pi_{1|2}(x_1|x_2) dx_1$$
- ▶ Standard conditions apply
$$\pi_{2|1}^*(dy_2|x_1) = \int P_2(x_2, dy_2|x_1) \pi_{2|1}(x_2|x_1) dx_2$$
- ▶ The product kernel $P_1(x_1, dy_1|x_2)P_2(x_2, dy_2|y_1)$ has invariant density $\pi(dy_1, dy_2)$
- ▶ **Wakeup Now** for exercise - prove the above.

Lecture Outline

Product Transition Kernels

Gibbs Sampler

Mixture Transition
KernelsData
Augmentation

$$\begin{aligned}
 & \int \int P_1(x_1, dy_1|x_2) P_2(x_2, dy_2|y_1) \pi(x_1, x_2) dx_1 dx_2 \\
 = & \int P_2(x_2, dy_2|y_1) \left[\int P_1(x_1, dy_1|x_2) \pi_{1|2}(x_1|x_2) dx_1 \right] \pi_2(x_2) dx_2 \\
 = & \int P_2(x_2, dy_2|y_1) \pi_{1|2}^*(dy_1|x_2) \pi_2(x_2) dx_2 \\
 = & \int P_2(x_2, dy_2|y_1) \frac{\pi_{2|1}(x_2|y_1) \pi_1^*(dy_1)}{\pi_2(x_2)} \pi_2(x_2) dx_2 \\
 = & \pi_1^*(dy_1) \int P_2(x_2, dy_2|y_1) \pi_{2|1}(x_2|y_1) dx_2 = \pi_1^*(dy_1) \pi_{2|1}^*(dy_2|y_1) \\
 = & \pi^*(dy_1, dy_2)
 \end{aligned}$$

Lecture Outline

Product Transition
Kernels

Gibbs Sampler

Mixture Transition
KernelsData
Augmentation

- ▶ Set kernels $P_1(x_1, dy_1|x_2) = \pi_{1|2}^*(dy_1|x_2)$ &
 $P_2(x_2, dy_2|y_1) = \pi_{2|1}^*(dy_2|y_1)$
- ▶ Then $\alpha(x, y) = 1$ for the transition of the first coordinate x_1 given x_2 :

$$\begin{aligned}
 \frac{\pi(y)q(y, x)}{\pi(x)q(x, y)} &= \frac{\pi(y_1, y_2)q([y_1, y_2], [x_1, x_2])}{\pi(x_1, x_2)q([x_1, x_2], [y_1, y_2])} \\
 &= \frac{\pi(y_1|y_2)\pi(y_2) \times \pi(x_1|x_2)}{\pi(x_1|x_2)\pi(x_2) \times \pi(y_1|x_2)} \\
 &= \frac{\pi(y_1|x_2)\pi(x_2) \times \pi(x_1|x_2)}{\pi(x_1|x_2)\pi(x_2) \times \pi(y_1|x_2)} \\
 &= 1
 \end{aligned}$$

Lecture Outline

Product Transition
Kernels

Gibbs Sampler

Mixture Transition
KernelsData
Augmentation

- ▶ Target density is $\mathcal{N}(\mathbf{0}, \mathbf{C})$ where $C_{1,1} = C_{2,2} = 1$ and $C_{1,2} = C_{2,1} = \rho$
- ▶ MH Algorithm with proposal $\mathcal{N}(\mathbf{y}|\mathbf{x}, \sigma\mathbf{I})$
- ▶ MH acceptance ratio

$$\alpha(\mathbf{x}, \mathbf{y}) = \min \left[\frac{\exp(-0.5 \times (\mathbf{y}^T \mathbf{C}^{-1} \mathbf{y}))}{\exp(-0.5 \times (\mathbf{x}^T \mathbf{C}^{-1} \mathbf{x}))}, 1 \right]$$

- ▶ What issues will arise with the choice of proposal ?
- ▶ MH with exact conditionals as proposals = Gibbs Sampler
- ▶

$$\begin{aligned} y_1|x_2 &\sim \mathcal{N}(\rho \times x_2, 1 - \rho^2) \\ y_2|y_1 &\sim \mathcal{N}(\rho \times y_1, 1 - \rho^2) \end{aligned}$$

- ▶ For two transition kernels $P_1(x, dy)$ and $P_2(x, dy)$ that both have $\pi(x)$ as invariant density
- ▶ and a probability $0 \leq \gamma \leq 1$
- ▶ what is the invariant density of the kernel composed of a mixture $\gamma \times P_1(x, dy) + (1 - \gamma) \times P_2(x, dy)$?
- ▶ Convince yourself that it is $\pi(x)$
- ▶ Can you think of situations where this might be useful ?
- ▶ Imagine a target density that may have multiple modes

[Lecture Outline](#)[Product Transition Kernels](#)[Gibbs Sampler](#)[Mixture Transition Kernels](#)[Data Augmentation](#)

- ▶ Monte Carlo estimates require i.i.d samples
- ▶ From our MCMC runs will the samples be i.i.d ?
- ▶ We know that repeated application of the transition kernel will lead to the desired target density
- ▶ How long will it be before we can be sure that the chain has converged ?

Data Augmentation

- ▶ Consider a target density $\pi(\theta)$ where $\theta \in \mathbb{R}^D$
- ▶ Let us now **augment** our model by introducing $\phi \in \mathbb{R}^D$
- ▶ Then the augmented target density is $\pi(\theta, \phi)$
- ▶ The desired density can be recovered as
$$\pi(\theta) = \int \pi(\theta, \phi) d\phi$$
- ▶ Therefore devising a Markov chain whose invariant density is $\pi(\theta, \phi)$
- ▶ and we draw samples $\theta^{(n)}, \phi^{(n)} \sim \pi(\theta, \phi)$
- ▶ then for each $\theta^{(n)}$ it follows that
$$\int \pi(\theta^{(n)}, \phi) d\phi = \pi(\theta^{(n)})$$
- ▶ Nice result..... each $\theta^{(n)}$ is marginally distributed with density $\pi(\cdot)$

[Lecture Outline](#)[Product Transition Kernels](#)[Gibbs Sampler](#)[Mixture Transition Kernels](#)[Data Augmentation](#)

Data Augmentation

- ▶ The main point is that we need to find a representation or completion that will admit exact conditionals for $\pi(\phi|\theta)$ and $\pi(\theta|\phi)$ that can be sampled from directly.
- ▶ Consider a Binary Probit Regression example popular in Machine Learning for classification problems.
- ▶ The probability for binary response variable t is $p(t = 1|x, \beta) = \Phi(\beta'x)$, where Φ is the standard normal CDF
- ▶ Employ DA by introducing the auxiliary variable y such that
- ▶ $y = \beta'x + \epsilon$ where $\epsilon \sim \mathcal{N}(0, 1)$
- ▶ Further define $t_i = 1 : y_i > 0$ and $t_i = 0 : y_i \leq 0$
- ▶ The joint density $p(t_i, y_i|x_i, \beta) = p(t_i|y_i)p(y_i|x_i, \beta)$
- ▶ By definition $p(t_i = 1|y_i) = \delta_{(t_i=1)}\delta_{(y_i>0)}$ and $p(t_i = 0|y_i) = \delta_{(t_i=0)}\delta_{(y_i\leq 0)}$
- ▶ Likewise $p(y_i|x_i, \beta) = \mathcal{N}(y_i|\beta'x_i, 1)$

Lecture Outline

Product Transition
Kernels

Gibbs Sampler

Mixture Transition
KernelsData
Augmentation

- The joint density $p(\mathbf{t}, \mathbf{y} | \beta, \mathbf{X})$ follows as

$$\prod_i [\delta_{(t_i=1)}\delta_{(y_i>0)} + \delta_{(t_i=0)}\delta_{(y_i\leq 0)}] \times \mathcal{N}(y_i | \beta' x_i, 1)$$

- Now derive full conditional $p(\beta | \mathbf{t}, \mathbf{y}, \mathbf{X})$
- $p(\beta | \mathbf{t}, \mathbf{y}, \mathbf{X}) \propto \prod_i \mathcal{N}(y_i | \beta' x_i, 1)$
- $p(\beta | \mathbf{t}, \mathbf{y}, \mathbf{X}) = \mathcal{N}(\beta | (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{y}, (\mathbf{X}' \mathbf{X})^{-1})$
- Now derive full conditional for $p(\mathbf{y} | \mathbf{X}, \mathbf{t}, \beta)$
- This is a product of simple truncated Normals so that

$$\begin{aligned} y_i &\sim TN_{(0, \infty)}(\beta' x_i, 1) : t_i = 1 \\ y_i &\sim TN_{(\infty, 0)}(\beta' x_i, 1) : t_i = 0 \end{aligned}$$

Lecture Outline

Product Transition
Kernels

Gibbs Sampler

Mixture Transition
KernelsData
Augmentation

- ▶ The overall Gibbs sampling scheme follows as

$$\beta | \mathbf{t}, \mathbf{y}, \mathbf{X} \sim \mathcal{N}(\beta | (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}, (\mathbf{X}'\mathbf{X})^{-1})$$

$$y_i \sim TN_{(0,\infty)}(\beta' x_i, 1) : t_i = 1$$

$$y_i \sim TN_{(\infty,0)}(\beta' x_i, 1) : t_i = 0 \text{ for all } i$$

- ▶ A prior can be placed on β i.e. $\pi_0(\beta) = \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$
- ▶ In the examples paper you will derive the Gibbs sampler with a prior on β and this is implemented in the notebooks.