practical_4

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Problem 1

a

Importing the data

```
library(readr)
df<- read.csv("C:\\Users\\souma\\Dropbox\\Mstat_CU\\Sem 2\\Regression_analysis_1\\Data Sets\\exp_inc_we
colnames(df)<- c("y","x1","x2")</pre>
colnames(df)
## [1] "y" "x1" "x2"
Fit the model
model1<- lm(y~ x1+x2,data=df)</pre>
summary(model1)
##
## lm(formula = y \sim x1 + x2, data = df)
##
## Residuals:
##
       Min
                1Q Median
                                ЗQ
                                        Max
## -22.637 -5.657
                     2.974
                             8.076
                                      9.237
##
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 18.165539 11.360774 1.599
                                                0.154
## x1
                0.632614
                           1.384489
                                       0.457
                                                0.662
## x2
               -0.009891
                           0.135714 -0.073
                                                0.944
##
## Residual standard error: 11.45 on 7 degrees of freedom
## Multiple R-squared: 0.9104, Adjusted R-squared: 0.8848
## F-statistic: 35.57 on 2 and 7 DF, p-value: 0.0002151
```

All the coefficients are insignificant, but the adj.r-square is very high, so there may be some multicollinearity.

Correlation Matrix

```
cor(df[,-1])
## x1 x2
```

```
## x1 1.0000000 0.9989624
## x2 0.9989624 1.0000000
```

The correlation between x1 and x2 is very high, so there is a high multicollinearity.

VIF

```
library(car)

## Loading required package: carData
vif(model1)

## x1 x2
## 482.1275 482.1275
```

Thumb Rule: The Vif's are quite large (» 10), so tehere is a multicollinearity in between the variables.

$\mathbf{C}\mathbf{N}$

```
# For the Design Matrix
d<-model.matrix(model1)
d_t<- t(d) %*% d # X'X

# Finding the eigen values of matrix
lambda<-eigen(d_t)$values
lambda

## [1] 3.403227e+07 6.795204e+01 1.016483e+00

# CN
cn<- sqrt(max(lambda)/min(lambda))
cn</pre>
```

[1] 5786.226

The CN is too large, so there there is multicolliearity in the data.

Remadial Measures

b: adding two obsns.....

adding 11th and 12th observation

```
a11<- c(160,120,3000) # 11th obns
a12<- c(85,255,920) # 12th obsn
df1<- rbind(df,a11,a12)
dim(df1)
```

[1] 12 3

Fit the model in new data

```
model2<- lm(y~. , data=df1)
summary(model2)</pre>
```

##

```
## Call:
## lm(formula = y ~ ., data = df1)
##
## Residuals:
##
      Min
               1Q Median
                               3Q
                                      Max
## -22.561 -5.413
                    2.400
                                     9.639
                            6.428
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 17.852113 10.173655
                                     1.755
                                             0.1132
               0.100630
                           0.054652
                                     1.841
                                             0.0987 .
               0.042541
                           0.004735
                                     8.984 8.66e-06 ***
## x2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 10.28 on 9 degrees of freedom
## Multiple R-squared: 0.9289, Adjusted R-squared: 0.9131
## F-statistic: 58.8 on 2 and 9 DF, p-value: 6.809e-06
Check for multicollinearity
```

```
vif(model2)
##
                   x2
         x1
## 1.203552 1.203552
# design matrix
d1<- model.matrix(model2)</pre>
dd<- t(d1) %*% d1
# eigen value of matrix
lambda1<-eigen(dd)$values</pre>
cn2<- sqrt(max(lambda1)/min(lambda1))</pre>
cn2
## [1] 6554.492
# correlation matrix
cor(df1[,-1])
##
                         x2
              x1
## x1 1.0000000 0.4112492
## x2 0.4112492 1.0000000
```

Comment: The VIF becomes low, the correlation between the variables becomes low, but the CN is very large.

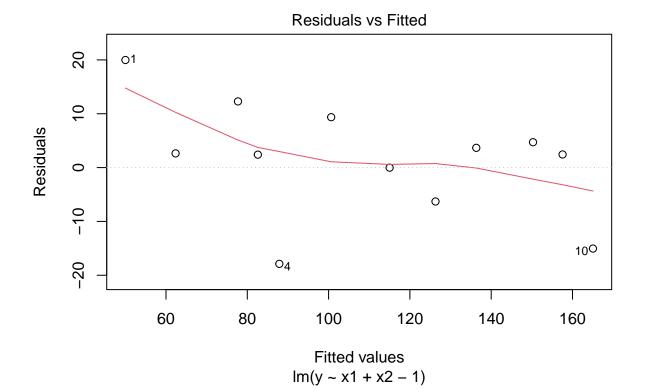
Comment: As the Vifs are low and correlation between the variables is not such high their might not be the problem of multicollinearity.

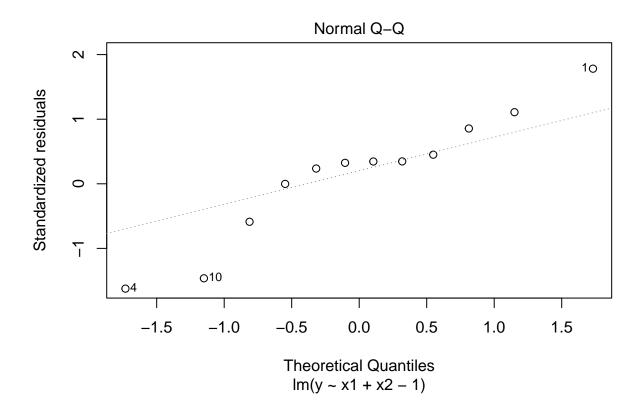
SO we need to re fit the model with out insignificant parameters

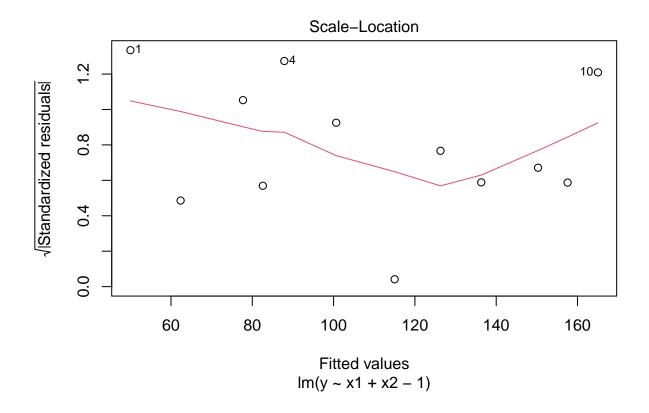
```
model_2.1 \leftarrow lm(y \sim x1 + x2 -1, data=df1)
summary(model_2.1)
```

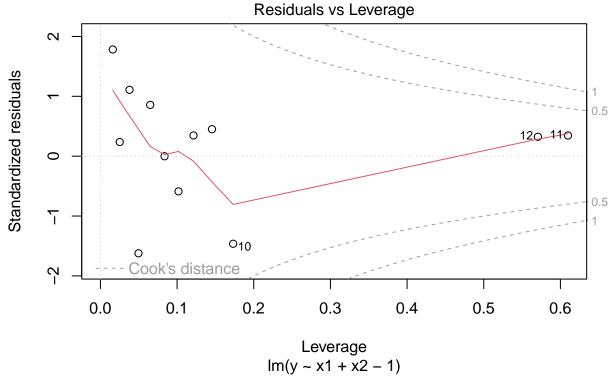
```
##
## Call:
```

```
## lm(formula = y ~ x1 + x2 - 1, data = df1)
##
## Residuals:
##
       Min
                1Q
                                ЗQ
                    Median
                                       Max
##
   -17.882
           -1.589
                     2.537
                             5.871
##
## Coefficients:
      Estimate Std. Error t value Pr(>|t|)
##
## x1 0.157099
                 0.048548
                            3.236 0.00893 **
## x2 0.046237
                 0.004661
                            9.921 1.71e-06 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 11.3 on 10 degrees of freedom
## Multiple R-squared: 0.9921, Adjusted R-squared: 0.9905
## F-statistic: 624.3 on 2 and 10 DF, p-value: 3.167e-11
### Further checking for multicollinearity
vif(model_2.1)
## Warning in vif.default(model_2.1): No intercept: vifs may not be sensible.
##
         x1
## 7.404982 7.404982
The vif is still low, so no multicollineasity is there.
plot(model_2.1)
```









Comment: So, the last model is free from multicollinearity with all the significant parameters with a good Adj R^2 value.

so this model is good enough to work.

c: PCA

PCA

Here we will work with the first data set

```
pca1<- princomp(df[,-1])</pre>
summary(pca1)
## Importance of components:
##
                                 Comp.1
                                               Comp.2
## Standard deviation
                           588.8342894 2.603776e+00
## Proportion of Variance
                             0.9999804 1.955297e-05
## Cumulative Proportion
                             0.9999804 1.000000e+00
prcomp(df[,-1])
## Standard deviations (1, .., p=2):
   [1] 620.685840
                     2.744621
##
## Rotation (n x k) = (2 \times 2):
             PC1
                          PC2
##
## x1 0.09745891 -0.99523955
## x2 0.99523955
                   0.09745891
```

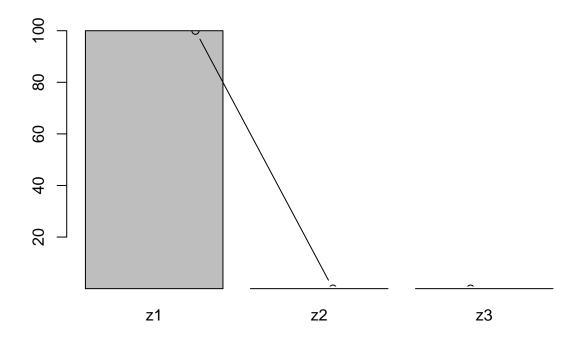
```
so we will use the PC1, component
df$z1<- (0.09745891*df$x1+0.99523955*df$x2)
model_pca<- lm(y~z1 ,data=df )</pre>
summary(model_pca)
##
## Call:
## lm(formula = y \sim z1, data = df)
## Residuals:
##
      Min
                1Q Median
                                3Q
                                       Max
## -22.106 -5.275
                   2.635 7.375 10.310
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) 17.921019 10.769705
                                     1.664
## z1
                0.051810
                          0.005838
                                     8.875 2.05e-05 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 10.87 on 8 degrees of freedom
## Multiple R-squared: 0.9078, Adjusted R-squared: 0.8963
## F-statistic: 78.76 on 1 and 8 DF, p-value: 2.054e-05
model_pca_u<- lm(y~z1-1 ,data=df )</pre>
summary(model_pca_u)
##
## Call:
## lm(formula = y \sim z1 - 1, data = df)
## Residuals:
##
       Min
                1Q Median
                                3Q
                                       Max
## -17.367 -4.347
                    4.083 8.838 20.336
##
## Coefficients:
     Estimate Std. Error t value Pr(>|t|)
## z1 0.061016 0.002038
                           29.93 2.53e-10 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 11.89 on 9 degrees of freedom
## Multiple R-squared: 0.9901, Adjusted R-squared: 0.989
## F-statistic: 896.1 on 1 and 9 DF, p-value: 2.532e-10
The model becomes
                y = 0.061016 \times Z_1 = 0.061016 \times (0.09745891 * x_1 + 0.9952395 * 5x_2)
Sir's algorithm
```

```
#design matrix
d
## (Intercept) x1 x2
```

1

1 80 810

```
1 100 1009
## 2
## 3
               1 120 1273
## 4
               1 140 1425
## 5
                1 160 1633
## 6
                1 180 1876
## 7
                1 200 2052
## 8
               1 220 2201
## 9
                1 240 2435
## 10
                1 260 2686
## attr(,"assign")
## [1] 0 1 2
# X'X
d_t
##
               (Intercept)
                                         x2
                                x1
## (Intercept)
                              1700
                                      17400
                        10
## x1
                      1700 322000 3294300
                     17400 3294300 33710326
## x2
# eigen values of X'X
1<- eigen(d_t)$values</pre>
# eigen vectors of X'X
L<- eigen(d_t)$vectors
##
                [,1]
                             [,2]
                                          [,3]
## [1,] -0.000513714 0.005902328 0.999982449
## [2,] -0.097260635 -0.995241903 0.005824382
## [3,] -0.995258813  0.097255936 -0.001085334
# note: L is a otthogonal matrix
## Define Z
z=d %*% L
           [,1]
                      [,2]
                                  [,3]
## 1 -813.941 -0.8361417 0.58681260
## 2 -1013.943 -1.3870485 0.48731881
## 3 -1278.636 4.3836805 0.31727832
## 4 -1431.861 -0.7382552 0.26879522
## 5 -1640.820 -0.4138586 0.15953342
## 6 -1884.613 3.3144958 0.01228494
## 7 -2061.724 0.5267025 -0.06224618
## 8 -2211.963 -4.8870011 -0.10747328
## 9 -2446.798 -2.0339501 -0.24495375
## 10 -2698.553 2.4724517 -0.40088491
## Variability of Z i
pc_var < c(var(z[,1]), var(z[,2]), var(z[,3]))
## plot of variability explained
barplot((pc_var/sum(pc_var))*100, names.arg = c("z1", "z2", "z3"), ylim = c(1,100))
lines((pc_var/sum(pc_var))*100,type = 'b')
```



So the first component explains the most of the variability, so we will take z1 variable for fitting the model.

```
model_p \leftarrow lm(df\$y~z[,1])
summary(model_p)
##
## Call:
## lm(formula = df\$y \sim z[, 1])
##
## Residuals:
##
       Min
                 1Q
                     Median
                                  3Q
                                          Max
                      2.635
                               7.375
##
   -22.106 -5.275
                                      10.310
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept) 17.920992
                            10.769713
                                         1.664
## z[, 1]
                -0.051810
                             0.005838
                                       -8.875 2.05e-05 ***
##
  ---
## Signif. codes:
                    0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 10.87 on 8 degrees of freedom
## Multiple R-squared: 0.9078, Adjusted R-squared: 0.8963
## F-statistic: 78.76 on 1 and 8 DF, p-value: 2.054e-05
The intercept is insignificant, so need to drop that
model_pu \leftarrow lm(df y \sim 0 + z[,1])
summary(model_pu)
```

```
##
## Call:
## lm(formula = df\$y ~ 0 + z[, 1])
##
## Residuals:
##
       Min
                1Q Median
                                 3Q
                                        Max
## -17.367 -4.347
                     4.083
                              8.838 20.336
##
## Coefficients:
##
           Estimate Std. Error t value Pr(>|t|)
## z[, 1] -0.061016
                      0.002038 -29.93 2.53e-10 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 11.89 on 9 degrees of freedom
## Multiple R-squared: 0.9901, Adjusted R-squared: 0.989
## F-statistic: 896.1 on 1 and 9 DF, p-value: 2.532e-10
So, the model is
                                     y = -0.061016 \times Z_1
## Finding estimaton of beta
eta<- -0.061016
beta = L[,1] * eta
round(beta,5)
## [1] 0.00003 0.00593 0.06073
The final model
                           y = 0.00003 + 0.00593 \times X_1 + 0.06073 \times X_2
Ridge Regression
## Removing the z1 variable
df \leftarrow df[,-4]
library(glmnet)
## Loading required package: Matrix
## Loaded glmnet 4.1-4
aa < 10^seq(2, -3, by = -0.1)
ridge_reg = glmnet(df[,-1], df[,1], nlambda = 25, alpha = 0, family = 'gaussian', lambda = aa)
summary(ridge_reg)
##
             Length Class
                               Mode
## a0
              51
                    -none-
                               numeric
             102
                    dgCMatrix S4
## beta
                               numeric
## df
              51
                    -none-
## dim
               2
                    -none-
                               numeric
## lambda
              51
                    -none-
                               numeric
## dev.ratio 51
                    -none-
                               numeric
## nulldev
               1
                    -none-
                               numeric
## npasses
               1
                    -none-
                               numeric
```

```
## call
               7
                     -none-
                               call
## nobs
               1
                     -none-
                               numeric
cv_ridge \leftarrow cv.glmnet(as.matrix(df[,-1]), df[,1], alpha = 0, lambda = aa)
## Warning: Option grouped=FALSE enforced in cv.glmnet, since < 3 observations per
optimal_lambda <- cv_ridge$lambda.min
optimal_lambda
## [1] 1.995262
so, the optimal choice of c is 1.995262
Estimation of Beta
#X'X
d_t
##
                (Intercept)
                                 x1
                                           x2
## (Intercept)
                               1700
                                        17400
                         10
## x1
                       1700 322000 3294300
## x2
                      17400 3294300 33710326
# I matrix
I = diag(rep(1,3))
#beta
beta= (solve(d_t + (optimal_lambda*I))) %*% t(d) %*% df[,1]
beta
##
                       [,1]
## (Intercept) 6.129046444
               0.547649530
## x1
```

Ridge Regression Sir's Algorithm

0.004625073

```
# Choice of C
```

 $y = 6.129046444 + 0.547649530 \times x_1 + 0.004625073 \times x_2$

Problem 2

The final model

x2

jerr

offset

1

1

-none-

-none-

numeric

logical

Importing data

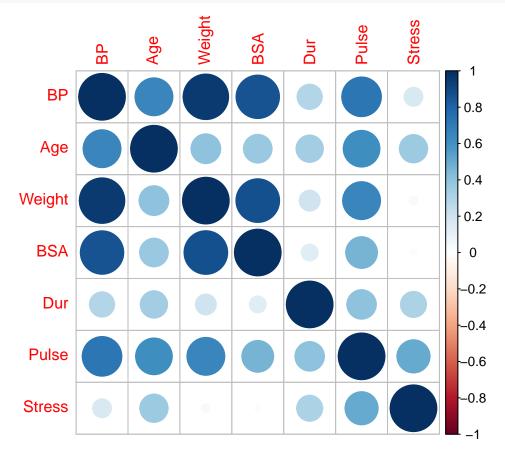
```
df_4<- read.csv("C:\\Users\\souma\\Dropbox\\Mstat_CU\\Sem 2\\Regression_analysis_1\\Data Sets\\problem_
colnames(df_4)[1]<-"pt"

cor(df_4[,-1])

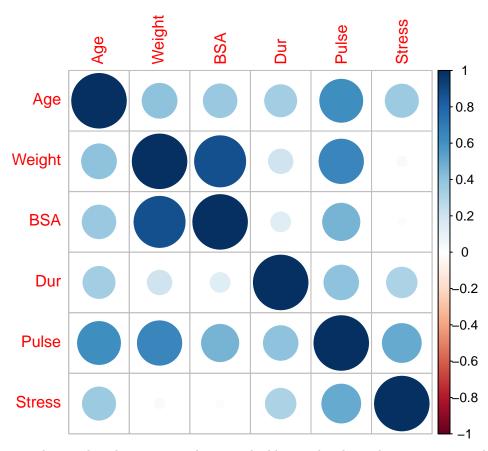
## BP Age Weight BSA Dur Pulse Stress</pre>
```

```
## BP
          1.0000000 0.6590930 0.95006765 0.86587887 0.2928336 0.7214132 0.16390139
          0.6590930\ 1.0000000\ 0.40734926\ 0.37845460\ 0.3437921\ 0.6187643\ 0.36822369
## Age
## Weight 0.9500677 0.4073493 1.00000000 0.87530481 0.2006496 0.6593399 0.03435475
## BSA
          0.8658789\ 0.3784546\ 0.87530481\ 1.00000000\ 0.1305400\ 0.4648188\ 0.01844634
          0.2928336 0.3437921 0.20064959 0.13054001 1.0000000 0.4015144 0.31163982
## Pulse 0.7214132 0.6187643 0.65933987 0.46481881 0.4015144 1.0000000 0.50631008
## Stress 0.1639014 0.3682237 0.03435475 0.01844634 0.3116398 0.5063101 1.00000000
library(corrplot)
```

corrplot 0.92 loaded corrplot(cor(df_4[,-1]))



 $corrplot(cor(df_4[,c(-1,-2)]))$



From the plot we observe that the weight and BSA are highly correlated, so, they may cause multicollinearity.

Model fitting

```
m_4<- lm(BP ~Age+Weight+BSA+Dur+Pulse+Stress, data = df_4)</pre>
summary(m_4)
##
## Call:
## lm(formula = BP ~ Age + Weight + BSA + Dur + Pulse + Stress,
##
       data = df_4
##
## Residuals:
##
                  1Q
                       Median
                                             Max
## -0.93213 -0.11314 0.03064 0.21834
##
## Coefficients:
                 Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -12.870476
                            2.556650
                                      -5.034 0.000229 ***
## Age
                 0.703259
                            0.049606
                                      14.177 2.76e-09 ***
                 0.969920
                            0.063108
                                      15.369 1.02e-09 ***
## Weight
## BSA
                 3.776491
                            1.580151
                                       2.390 0.032694 *
                 0.068383
                            0.048441
                                       1.412 0.181534
## Dur
## Pulse
                -0.084485
                            0.051609
                                      -1.637 0.125594
## Stress
                 0.005572
                            0.003412
                                       1.633 0.126491
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

```
##
## Residual standard error: 0.4072 on 13 degrees of freedom
## Multiple R-squared: 0.9962, Adjusted R-squared: 0.9944
## F-statistic: 560.6 on 6 and 13 DF, p-value: 6.395e-15
summary(influence.measures(m_4))
## Potentially influential observations of
    lm(formula = BP ~ Age + Weight + BSA + Dur + Pulse + Stress,
                                                                    data = df_4):
##
##
     dfb.1_ dfb.Age dfb.Wght dfb.BSA dfb.Dur dfb.Puls dfb.Strs dffit cov.r
      0.68 -0.26
## 7
                   -0.61
                            -0.13
                                     1.14_* 0.71
                                                    -0.57
                                                             -1.96 0.71
## 11 -0.10
           0.33
                    0.24
                            -0.24
                                     0.16 -0.35
                                                    -0.02
                                                             -0.51 2.68 *
## 13 -0.07 -0.04
                                    -0.18
                                                             -0.25 2.87_*
                    0.10
                            -0.11
                                            0.02
                                                     0.05
## 15 0.05 -0.01
                   -0.04
                             0.02 -0.03
                                            0.02
                                                    -0.03
                                                             -0.08 2.92_*
## 16 0.05 -0.20
                   0.33
                            -0.23 0.10 -0.27
                                                     0.39
                                                              0.53 2.64_*
## 19 -0.69
             0.66
                   -0.06
                            0.41
                                   -1.03_* -0.10
                                                    -0.80
                                                             -2.14 0.01
## 20 -0.18
             0.26
                   -0.02
                           0.00 -0.08 -0.01
                                                    0.03
                                                              0.33 3.86_*
##
     cook.d hat
## 7
      0.47
             0.54
## 11 0.04
            0.45
## 13 0.01
             0.42
## 15 0.00
            0.40
## 16 0.04
             0.45
## 19 0.33
             0.25
## 20 0.02
             0.57
Vif
```

```
library(car)
vif(m_4)
```

```
Weight
                         BSA
                                   Dur
                                          Pulse
       Age
## 1.762807 8.417035 5.328751 1.237309 4.413575 1.834845
```

CN

```
d_matrix<- model.matrix(m_4)</pre>
x<- t(d_matrix) %*% d_matrix
1<- eigen(x)$values</pre>
cn<- sqrt(max(1)/min(1))</pre>
```

[1] 4016.581

The CN is too large, so there exists multicollieanrity .

Fit a model without BSA

```
m_4.1<- lm(BP ~Age+Weight+Dur+Pulse+Stress, data = df_4)</pre>
summary(m_4.1)
```

##

```
## Call:
## lm(formula = BP ~ Age + Weight + Dur + Pulse + Stress, data = df_4)
## Residuals:
                 1Q
                     Median
                                   3Q
## -1.02600 -0.18526 -0.00077 0.21934 0.72533
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) -15.116781
                          2.748758 -5.499 7.83e-05 ***
                0.731940
                           0.055646 13.154 2.85e-09 ***
                           0.037773 29.093 6.37e-14 ***
## Weight
                1.098958
## Dur
                0.064105
                           0.055965
                                     1.145
                                             0.2712
## Pulse
               -0.137444
                           0.053885 -2.551
                                              0.0231 *
                0.007429
                           0.003841
                                              0.0736 .
## Stress
                                     1.934
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.4708 on 14 degrees of freedom
## Multiple R-squared: 0.9945, Adjusted R-squared: 0.9925
## F-statistic: 502.5 on 5 and 14 DF, p-value: 2.835e-15
# ----- Lets calculated vif for the model -----
vif(m_4.1)
       Age
            Weight
                         Dur
                                Pulse
                                        Stress
## 1.659637 2.256150 1.235620 3.599913 1.739641
Fit a model without Pluse
m_4.2<- lm(BP ~Age+Weight+Dur+Stress, data = df_4)</pre>
summary(m 4.2)
##
## lm(formula = BP ~ Age + Weight + Dur + Stress, data = df_4)
##
## Residuals:
                     Median
       Min
                 1Q
                                   3Q
## -1.11359 -0.29586 0.01515 0.27506 0.88674
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -15.869829 3.195296 -4.967 0.000169 ***
                           0.061195 11.173 1.14e-08 ***
## Age
                0.683741
## Weight
                1.034128
                           0.032672 31.652 3.76e-15 ***
                0.039889
                           0.064486 0.619 0.545485
## Dur
## Stress
                0.002184
                           0.003794 0.576 0.573304
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5505 on 15 degrees of freedom
## Multiple R-squared: 0.9919, Adjusted R-squared: 0.9897
## F-statistic: 458.3 on 4 and 15 DF, p-value: 1.764e-15
```

This is a Failure

Fit a model without stress

```
m_4.3<- lm(BP ~Age+Weight+Dur, data = df_4)</pre>
summary(m_4.3)
##
## Call:
## lm(formula = BP ~ Age + Weight + Dur, data = df_4)
## Residuals:
##
       Min
                  1Q
                      Median
                                            Max
## -1.03592 -0.29671 0.05216 0.32551 0.85934
##
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
                                    -5.185 9.04e-05 ***
## (Intercept) -16.09486
                            3.10435
## Age
                 0.69526
                            0.05661 12.280 1.47e-09 ***
                 1.03121
                            0.03159 32.639 4.54e-16 ***
## Weight
## Dur
                 0.04821
                            0.06152
                                      0.784
                                               0.445
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.5388 on 16 degrees of freedom
## Multiple R-squared: 0.9917, Adjusted R-squared: 0.9901
## F-statistic: 637.6 on 3 and 16 DF, p-value: < 2.2e-16
```

Again a Bad model, as the VIF for this model has less VIFs so, there is no multicollinearity, so the parameter is ingnificant not for multicollinearity.

Fitting a model removing dur keeping stress

```
m_4.4<-lm(BP ~Age+Weight+Stress, data = df_4)</pre>
summary(m_4.4)
##
## Call:
## lm(formula = BP ~ Age + Weight + Stress, data = df_4)
##
## Residuals:
##
      Min
                1Q Median
                                3Q
                                       Max
## -1.0252 -0.3277 0.0368 0.2274
                                   0.8901
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept) -16.196316
                            3.090002 -5.242 8.07e-05 ***
                            0.058833 11.748 2.80e-09 ***
                 0.691179
## Weight
                 1.036206
                            0.031865
                                      32.518 4.82e-16 ***
                            0.003625
                                       0.748
## Stress
                 0.002710
                                                0.465
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5397 on 16 degrees of freedom
```

```
## Multiple R-squared: 0.9917, Adjusted R-squared: 0.9901
## F-statistic: 635.4 on 3 and 16 DF, p-value: < 2.2e-16</pre>
```

Fitting model without BSA and Stress

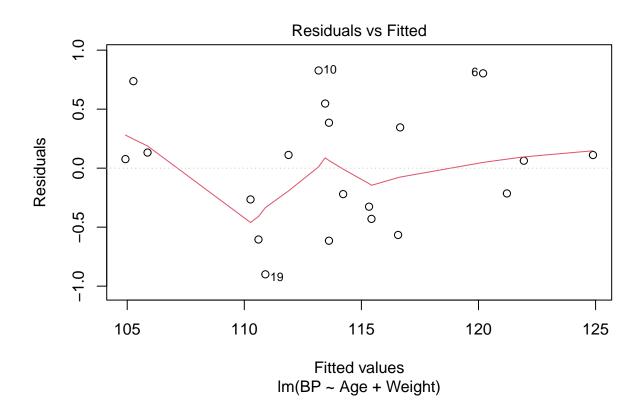
```
m_4.5<- lm(BP~Age+Weight+Pulse+Dur, data = df_4)</pre>
summary(m 4.5)
##
## lm(formula = BP ~ Age + Weight + Pulse + Dur, data = df_4)
##
## Residuals:
##
       Min
                  1Q
                      Median
                                    3Q
                                            Max
## -0.87309 -0.25007 -0.00217 0.30303 0.89738
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -15.96851
                            2.95072 -5.412 7.21e-05 ***
                            0.06033 12.271 3.19e-09 ***
## Age
                0.74032
                1.06556
                            0.03654 29.164 1.26e-14 ***
## Weight
## Pulse
                -0.08165
                            0.04950
                                    -1.650
                                               0.120
                 0.07448
                            0.06058
                                     1.229
                                               0.238
## Dur
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.512 on 15 degrees of freedom
## Multiple R-squared: 0.993, Adjusted R-squared: 0.9911
## F-statistic: 530.3 on 4 and 15 DF, p-value: 5.957e-16
#----VIf-
vif(m_4.5)
        Age
              Weight
                        Pulse
## 1.649586 1.784753 2.568296 1.224274
```

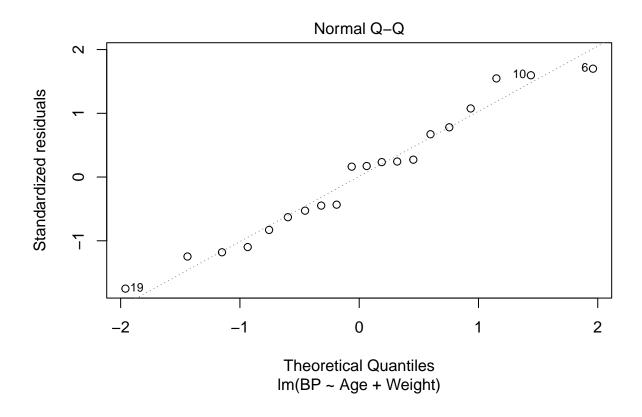
Fitting MOdel without BSA, Stress, DUR

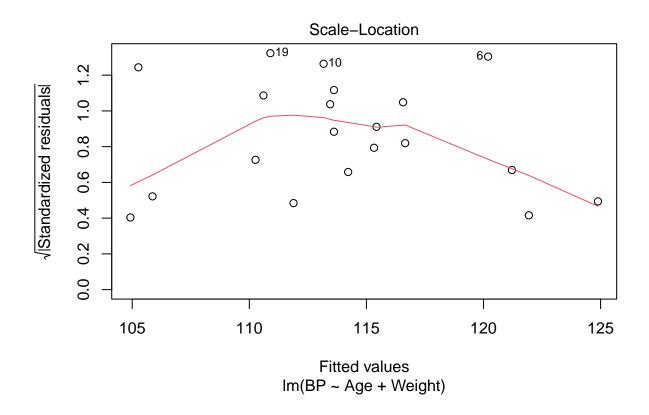
```
m_4.6<- lm(BP ~Age+Weight+Pulse, data = df_4)</pre>
summary(m_4.6)
##
## Call:
## lm(formula = BP ~ Age + Weight + Pulse, data = df_4)
##
## Residuals:
##
       Min
                  1Q
                      Median
                                     3Q
                                             Max
## -0.71174 -0.45422 -0.01909 0.41745 0.88743
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) -16.69000
                            2.93761 -5.681 3.40e-05 ***
## Age
                 0.75018
                            0.06074 12.350 1.36e-09 ***
                            0.03695 28.722 3.40e-15 ***
## Weight
                 1.06135
## Pulse
                -0.06566
                            0.04852 - 1.353
                                                0.195
```

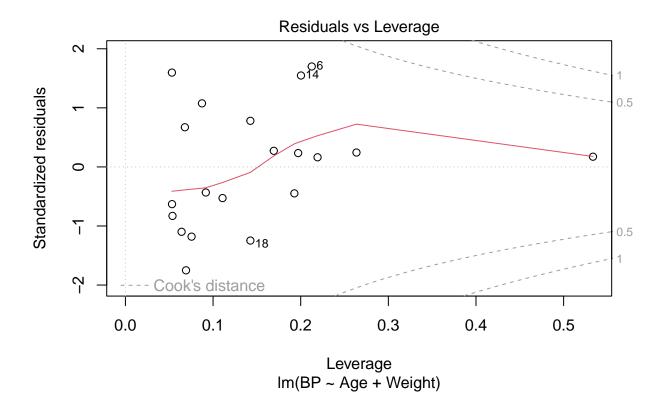
```
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.5201 on 16 degrees of freedom
## Multiple R-squared: 0.9923, Adjusted R-squared: 0.9908
## F-statistic: 684.7 on 3 and 16 DF, p-value: < 2.2e-16
#----VIF----
vif(m_4.6)
##
                       Pulse
       Age
             Weight
## 1.620404 1.769065 2.390933
Again fit the model without Pluse
m_4.6 \leftarrow lm(BP \sim Age + Weight, data = df_4)
summary(m_4.6)
##
## Call:
## lm(formula = BP ~ Age + Weight, data = df_4)
##
## Residuals:
##
       Min
                 1Q
                     Median
                                    3Q
                                            Max
## -0.89968 -0.35242 0.06979 0.35528 0.82781
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
                           3.00746 -5.513 3.80e-05 ***
## (Intercept) -16.57937
                0.70825
                            0.05351 13.235 2.22e-10 ***
## Age
                1.03296
                           0.03116 33.154 < 2e-16 ***
## Weight
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.5327 on 17 degrees of freedom
## Multiple R-squared: 0.9914, Adjusted R-squared: 0.9904
## F-statistic: 978.2 on 2 and 17 DF, p-value: < 2.2e-16
So this is a significant model, lets check is there still multicollinearity or not
#-----vif-----
vif(m_4.6)
        Age
             Weight
## 1.198945 1.198945
The Vifs are also low
So, the data is good enough to work with further.
As we would consider this model, the model will be good enough if it statisfies the assumptions
# ----- Residual Analysis -----
```

 $plot(m_4.6)$









Stepwise Regression Model

FOrward stepwise

```
library(MASS)
int_only<- lm(BP~ 1, data=df_4)</pre>
s_m_f<- step(int_only, direction = "forward", scope = formula(m_4))</pre>
## Start:
           AIC=68.64
## BP ~ 1
##
##
            Df Sum of Sq
                             RSS
                                     AIC
                   505.47 54.53 24.060
## + Weight
## + BSA
                   419.86 140.14 42.938
## + Pulse
                   291.44 268.56 55.946
                   243.27 316.73 59.247
## + Age
## <none>
                          560.00 68.644
## + Dur
                    48.02 511.98 68.851
             1
                    15.04 544.96 70.099
## + Stress
##
## Step: AIC=24.06
## BP ~ Weight
##
##
                             RSS
                                      AIC
            Df Sum of Sq
## + Age
                   49.704 4.824 -22.443
```

```
9.660 44.868 22.160
## + Stress 1
## + Pulse 1
                8.940 45.588 22.478
## + Dur 1
                6.095 48.433 23.689
## <none>
                       54.528 24.060
## + BSA
                2.814 51.714 25.000
##
## Step: AIC=-22.44
## BP ~ Weight + Age
##
##
                          RSS
           Df Sum of Sq
                                  AIC
## + BSA
           1
               1.76778 3.0561 -29.572
## + Pulse
          1
               0.49557 4.3284 -22.611
## <none>
                       4.8239 -22.443
## + Dur
               0.17835 4.6456 -21.196
         1
## + Stress 1
               0.16286 4.6611 -21.130
##
## Step: AIC=-29.57
## BP ~ Weight + Age + BSA
##
##
           Df Sum of Sq RSS
## + Dur
           1 0.33510 2.7210 -29.894
## <none>
                       3.0561 -29.572
## + Stress 1
               0.21774 2.8384 -29.050
## + Pulse 1 0.04111 3.0150 -27.842
##
## Step: AIC=-29.89
## BP ~ Weight + Age + BSA + Dur
##
           Df Sum of Sq
                         RSS
                                  AIC
## <none>
                       2.7210 -29.894
## + Pulse 1 0.12307 2.5980 -28.820
## + Stress 1
               0.12077 2.6003 -28.802
```

Coefficients

```
s_m_f$coefficients
```

```
## (Intercept) Weight Age BSA Dur
## -12.85206440 0.89700637 0.68335254 4.86037186 0.06652958
```

Backward Stepwise

```
s_m_b<- step(m_4, direction = "backward", scope = formula(int_only))</pre>
## Start: AIC=-30.55
## BP ~ Age + Weight + BSA + Dur + Pulse + Stress
##
##
           Df Sum of Sq
                          RSS
## <none>
                        2.156 -30.551
## - Dur
            1
                 0.330 2.486 -29.698
## - Stress 1
                 0.442 2.598 -28.820
## - Pulse
            1
                 0.444 2.600 -28.802
## - BSA
                 0.947 3.103 -25.267
            1
## - Age
          1 33.331 35.486 23.468
```

```
## - Weight 1 39.172 41.328 26.516
```

Both side

```
s_m<- step(int_only, direction = "both", scope = formula(m_4))</pre>
## Start: AIC=68.64
## BP ~ 1
##
           Df Sum of Sq RSS
## + Weight 1 505.47 54.53 24.060
## + BSA
           1
                419.86 140.14 42.938
## + Pulse 1
                291.44 268.56 55.946
## + Age
            1
                243.27 316.73 59.247
## <none>
                       560.00 68.644
## + Dur 1
                48.02 511.98 68.851
## + Stress 1
                15.04 544.96 70.099
##
## Step: AIC=24.06
## BP ~ Weight
##
##
           Df Sum of Sq
                         RSS
                               AIC
## + Age
          1
               49.70
                        4.82 -22.443
## + Stress 1
                9.66 44.87 22.160
## + Pulse 1
                 8.94 45.59 22.478
## + Dur
                 6.09 48.43 23.689
           1
## <none>
                        54.53 24.060
## + BSA 1
                 2.81 51.71 25.000
## - Weight 1
                505.47 560.00 68.644
##
## Step: AIC=-22.44
## BP ~ Weight + Age
##
           Df Sum of Sq
##
                         RSS
              1.768
## + BSA
                         3.06 -29.572
           1
## + Pulse
          1
                0.496
                         4.33 -22.611
                         4.82 -22.443
## <none>
## + Dur
                 0.178
                         4.65 -21.196
           1
## + Stress 1
                 0.163
                        4.66 -21.130
## - Age 1
                49.704 54.53 24.060
## - Weight 1
               311.910 316.73 59.247
##
## Step: AIC=-29.57
## BP ~ Weight + Age + BSA
##
##
           Df Sum of Sq
                          RSS
## + Dur
          1 0.335 2.721 -29.894
## <none>
                        3.056 -29.572
                 0.218 2.838 -29.050
## + Stress 1
## + Pulse 1
                0.041 3.015 -27.842
## - BSA
            1
                1.768 4.824 -22.443
## - Age
                48.658 51.714 25.000
           1
## - Weight 1
                65.303 68.359 30.581
##
```

```
## Step: AIC=-29.89
## BP ~ Weight + Age + BSA + Dur
##
##
           Df Sum of Sq
                           RSS
                                    AIC
## <none>
                          2.721 -29.894
## - Dur
                  0.335 3.056 -29.572
            1
## + Pulse
                  0.123 2.598 -28.820
           1
## + Stress 1
                  0.121 2.600 -28.802
## - BSA
                  1.925 4.646 -21.196
            1
## - Age
            1
                 42.021 44.742 24.104
## - Weight 1
                  62.878 65.599 31.756
s_m$coefficients
## (Intercept)
                     Weight
                                      Age
                                                   BSA
                                                                Dur
## -12.85206440
                  0.89700637
                              0.68335254
                                            4.86037186
                                                         0.06652958
summary(s_m)
##
## Call:
## lm(formula = BP ~ Weight + Age + BSA + Dur, data = df_4)
##
## Residuals:
##
       Min
                 1Q
                     Median
                                    3Q
                                            Max
## -0.86420 -0.26320 0.08341 0.25020 0.58272
##
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -12.85206
                          2.64804 -4.853 0.000211 ***
## Weight
                 0.89701
                           0.04818 18.618 8.88e-12 ***
                 0.68335
                           0.04490 15.220 1.58e-10 ***
## Age
## BSA
                 4.86037
                           1.49220
                                    3.257 0.005305 **
## Dur
                 0.06653
                            0.04895
                                    1.359 0.194184
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.4259 on 15 degrees of freedom
## Multiple R-squared: 0.9951, Adjusted R-squared: 0.9938
                 768 on 4 and 15 DF, p-value: < 2.2e-16
## F-statistic:
vif(s_m)
    Weight
                 Age
                         BSA
## 4.484932 1.320201 4.344272 1.154968
## CN
a<- model.matrix(s m)</pre>
a<- t(a)%*%a
a <- eigen(a) $values
cn<- sqrt(max(a)/min(a))</pre>
cn
```

[1] 3015.973

Multicollinearity is still there.