

This example assume we know the latent parameter  $Z$ .

\* We observed a sample of 1000 data points. This sample follow a mixture of 2 Poisson distribution. Let  $Z_i$  be the distribution that a data point belong to

$$Z_i \sim \text{Categorical}(\pi_1, 1 - \pi_1),$$

$$X_i | Z_i = 1 \sim \text{Poisson}(\lambda_1) \text{ and } X_i | Z_i = 2 \sim \text{Poisson}(\lambda_2)$$

\* We aim to obtain MLE of parameters  $\Theta = (\pi_1, \lambda_1, \lambda_2)$  using the EM algorithm.

Solution :

$$\text{E-Step: } \Theta^0 = (\pi_1^0, \lambda_1^0, \lambda_2^0)$$

computing posteriors:

$$P(Z_i=1 | X, \Theta) = \frac{P(Z_i=1, X_i | \Theta^0)}{P(X_i | \Theta^0)}$$

$$= \frac{P(X_i | Z_i=1, \Theta^0) P(Z_i=1 | \Theta^0)}{\sum_{K=1}^2 P(X_i | Z_i=K, \Theta^0) \cdot P(Z_i=K | \Theta^0)}$$

$$= \frac{P(X_i | Z_i=1, \Theta^0) P(Z_i=1 | \Theta^0)}{P(X_i | Z_i=1, \Theta^0) P(Z_i=1 | \Theta^0) + P(X_i | Z_i=2, \Theta^0) P(Z_i=2 | \Theta^0)}$$

$$= \frac{(\lambda_1^0)^{x_i} e^{-\lambda_1^0}}{x_i!} \pi_1^0$$


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$$\frac{(\lambda_1^0)^{x_i} e^{-\lambda_1^0}}{x_i!} \pi_1^0 + \frac{(\lambda_2^0)^{x_i} e^{-\lambda_2^0}}{x_i!} (1 - \pi_1^0)$$

\* Now that we have  $P(Z_i = 1 | X, \Theta^0)$  we can get

$$P(Z_i = 0 | X, \Theta^0) = 1 - P(Z_i = 1 | X, \Theta^0)$$

$$= 1 - \frac{(\lambda_1^0)^{x_i} e^{-\lambda_1^0}}{x_i!} \pi_1^0$$


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$$\frac{(\lambda_1^0)^{x_i} e^{-\lambda_1^0}}{x_i!} \pi_1^0 + \frac{(\lambda_2^0)^{x_i} e^{-\lambda_2^0}}{x_i!} (1 - \pi_1^0)$$

## M-Step:

\* We now have to derive new parameter estimate from the previous posteriors:

$$\hat{\Theta} = \operatorname{argmax}_{\Theta} Q(\Theta, \Theta^0)$$

$$\text{where } Q(\Theta, \Theta^0) = E_{Z|X, \Theta^0} [\log(P(X, Z|\Theta))]$$

for  $n = 100$

$$P(X_1, \dots, X_n, Z_1, \dots, Z_n) = \prod_{i=1}^n P(X_i | Z_i, \theta) P(Z_i | \theta)$$

$$= \prod_{i=1}^n \prod_{k=1}^2 [P(X_i | Z_i = k, \theta) P(Z_i = k | \theta)]^{I_{(Z_i=k)}}$$

$$\log(P(X_1, \dots, X_n, Z_1, \dots, Z_n))$$

$$= \sum_{i=1}^n \sum_{k=1}^2 I_{(Z_i=k)} [\log(P(X_i | Z_i = k, \theta)) + \log(P(Z_i = k | \theta))]$$

$$= \sum_{i=1}^n [I_{(Z_i=1)} [X_i \log \lambda_1 - \lambda_1 - \log(X_i!) + \log(\pi_1)] \\ + I_{(Z_i=2)} [X_i \log \lambda_2 - \lambda_2 - \log(X_i!) + \log(1-\pi_1)]]$$

$$Q(\theta, \theta^*) = E_{z|x, \theta^*} [\log(P(x, z | \theta))]$$

$$= \sum_{i=1}^n [P(Z_i=1 | X, \theta) [X_i \log \lambda_1 - \lambda_1 - \log(X_i!) + \log(\pi_1)]$$

$$+ P(Z_i=2 | X, \theta) [X_i \log \lambda_2 - \lambda_2 - \log(X_i!) + \log(1-\pi_1)]]]$$

$$\Theta = (\pi_1, \lambda_1, \lambda_2)$$

\* now we need to get  $\hat{\Theta} = \arg \max_{\Theta} Q(\Theta, \Theta^0)$

$$\frac{dQ(\Theta, \Theta^0)}{d\pi_1} = \sum_{i=1}^n \left[ \frac{P(Z_i=1 | X, \Theta^0)}{\pi_1} - \frac{P(Z_i=2 | X, \Theta^0)}{1-\pi_1} \right]$$

$$= \frac{\sum_{i=1}^n \left[ (1-\pi_1) P(Z_i=1 | X, \Theta^0) - \pi_1 P(Z_i=2 | X, \Theta^0) \right]}{\pi_1 (1-\pi_1)}$$

$$= 0$$

$$\Rightarrow \frac{(1-\pi_1) \sum_{i=1}^n P(Z_i=1 | X, \Theta^0) - \pi_1 \sum_{i=1}^n (1 - P(Z_i=1 | X, \Theta^0))}{\pi_1 (1-\pi_1)}$$

$$= \frac{(1-\pi_1) \sum_{i=1}^n P(Z_i=1 | X, \Theta^0) - \pi_1 n + \pi_1 \sum_{i=1}^n P(Z_i=1 | X, \Theta^0)}{\pi_1 (1-\pi_1)}$$

$$= \frac{\sum_{i=1}^n P(Z_i=1 | X, \Theta^0) - n \pi_1}{\pi_1 (1-\pi_1)} = 0$$

$$= \sum_{i=1}^n P(Z_i=1 | X, \theta^0) - n \pi_1 = 0$$

$$\Rightarrow \hat{\pi}_1 = \frac{\sum_{i=1}^n P(Z_i=1 | X, \theta^0)}{n}$$

\*  $P(Z_i=1 | X, \theta^0)$  was obtained previously, in the E-Step

$$\frac{dQ(\theta, \theta^0)}{d\lambda_1}$$

$$= \frac{\sum_{i=1}^n P(Z_i=1 | X, \theta^0) x_i}{\lambda_1} - \sum_{i=1}^n P(Z_i=1 | X, \theta^0) = 0$$

$$\Rightarrow \hat{\lambda}_1 = \frac{\sum_{i=1}^n P(Z_i=1 | X, \theta^0) x_i}{\sum_{i=1}^n P(Z_i=1 | X, \theta^0)}$$

Similarly,

$$\frac{dQ(\theta, \theta^0)}{d\lambda_2}$$

$$= \frac{\sum_{i=1}^n P(Z_i=2 | X, \theta^0) x_i}{\lambda_2} - \sum_{i=1}^n P(Z_i=2 | X, \theta^0) = 0$$

$$\Rightarrow \hat{\lambda}_2 = \frac{\sum_{i=1}^n P(Z_i=2 | X, \theta^0) x_i}{\sum_{i=1}^n P(Z_i=2 | X, \theta^0)}$$

now we have parameter estimate after 1 iteration