#### warmXtrophic Project: Greenup Analyses

#### February 26, 2021

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DATA INPUT: Data imported as csv files from shared Google drive "SpaCE\_Lab\_warmXtrophic" plant comp folder

DATA OUTPUT: a brief description of the data output from through the script, including what format it's in OVERVIEW: This script ... REQUIRES: any scripts or code sources that are required NOTES: The final\_kbs and final\_umbs dataframes contain data for greenup at each site.

"half\_cover\_date" is the date at which 50% of a species max cover was reached (per plot, per year)

"state" describes each treatment - warmed or ambient

```
# Clear all existing data
rm(list=ls())
#Load packages
library(tidyverse)
library(lme4)
library(olsrr)
library(predictmeans)
library(car)
library(fitdistrplus)
library(ggpubr)
library(rstatix)
library(vegan)
# Set working directory to Google Drive
# **** Update with the path to your Google drive on your computer
setwd("/Volumes/GoogleDrive/Shared drives/SpaCE_Lab_warmXtrophic/data/")
# Read in plant comp data
greenup <- read.csv("L1/greenup/final greenup L1.csv")</pre>
# check variable types
str(greenup)
## 'data.frame':
                    3419 obs. of 13 variables:
```

```
## $ X
                  : int 1 2 3 4 5 6 7 8 9 10 ...
                   : chr "Acru" "Acru" "Acru" "Acru" ...
## $ species
## $ origin
                         "Native" "Native" "Native" ...
                   : chr
## $ duration
                  : chr "Perennial" "Perennial" "Perennial" "Perennial" ...
## $ growth_habit : chr "Tree" "Tree" "Tree" "Tree" ...
## $ plot
                   : chr
                         "A4" "B2" "B2" "B2" ...
                         "warmed" "warmed" "warmed" ...
## $ state
                   : chr
## $ treatment_key : chr "WO" "WO" "WO" "WO" ...
## $ insecticide : chr "no insects" "no insects" "no insects" "no insects" ...
## $ half_cover_date: int 167 135 145 167 175 144 167 141 202 197 ...
## $ site
                   : chr "umbs" "umbs" "umbs" ...
                   : int 2017 2016 2019 2017 2015 2018 2017 2015 2015 2016 ...
## $ year
## $ min_emerg_date : int 167 112 112 112 112 167 141 80 80 ...
```

```
# adding sequential year variable starting at 1
greenup$year1<-greenup$year == 2015] <- 1
greenup$year[greenup$year == 2016] <- 2
greenup$year[greenup$year == 2017] <- 3
greenup$year[greenup$year == 2018] <- 4
greenup$year[greenup$year == 2019] <- 5
greenup$year[greenup$year == 2020] <- 6</pre>
# create dataframes for kbs and umbs
final_kbs <- subset(greenup, site == "kbs")
final_umbs <- subset(greenup, site == "umbs")</pre>
```

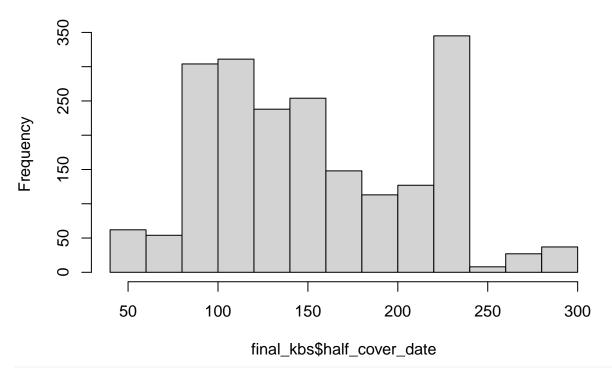
Data Exploration: are there differences between warmed vs. ambient plots?

Starting with KBS

First, checking for normality in raw data

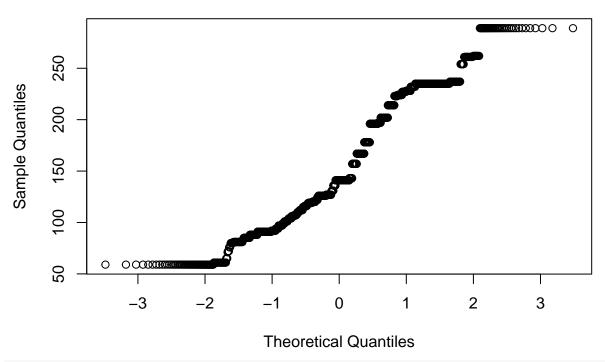
```
hist(final_kbs$half_cover_date)
```

## Histogram of final\_kbs\$half\_cover\_date



qqnorm(final\_kbs\$half\_cover\_date)

#### Normal Q-Q Plot

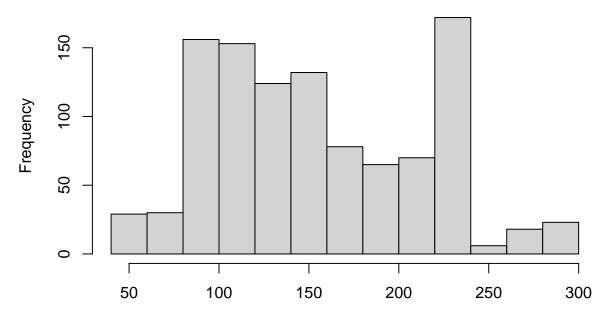


```
shapiro.test(final_kbs$half_cover_date)
```

```
##
## Shapiro-Wilk normality test
##
## data: final_kbs$half_cover_date
## W = 0.94273, p-value < 2.2e-16

# histograms for each treatment separately - look almost identical
hist(final_kbs$half_cover_date[final_kbs$state == "ambient"])</pre>
```

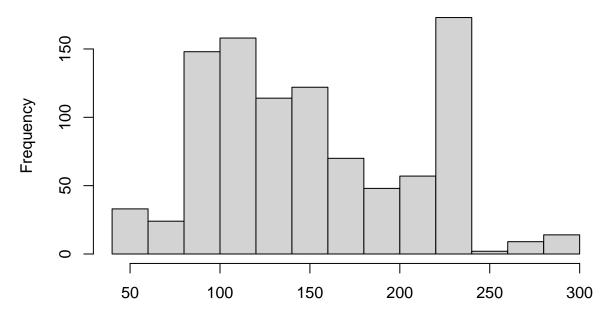
#### Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$state == "ambient"



final\_kbs\$half\_cover\_date[final\_kbs\$state == "ambient"]

hist(final\_kbs\$half\_cover\_date[final\_kbs\$state == "warmed"])

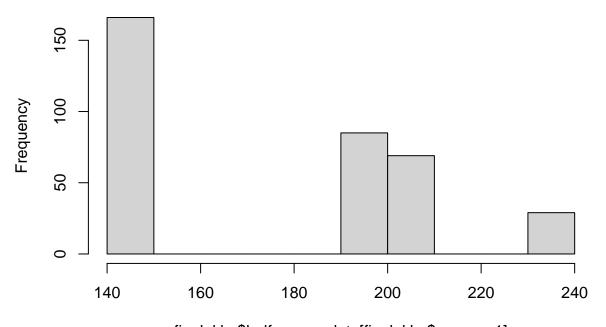
#### Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$state == "warmed



final\_kbs\$half\_cover\_date[final\_kbs\$state == "warmed"]

# histograms for each year
hist(final\_kbs\$half\_cover\_date[final\_kbs\$year == 1])

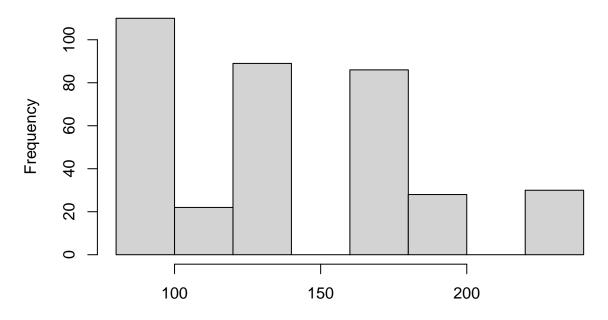
## Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$year == 1]



final\_kbs\$half\_cover\_date[final\_kbs\$year == 1]

hist(final\_kbs\$half\_cover\_date[final\_kbs\$year == 2])

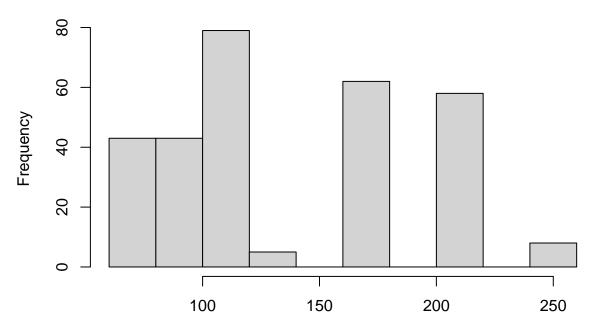
## Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$year == 2]



final\_kbs\$half\_cover\_date[final\_kbs\$year == 2]

hist(final\_kbs\$half\_cover\_date[final\_kbs\$year == 3])

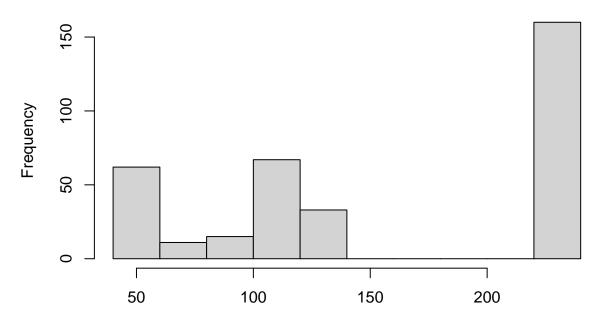
## Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$year == 3]



final\_kbs\$half\_cover\_date[final\_kbs\$year == 3]

hist(final\_kbs\$half\_cover\_date[final\_kbs\$year == 4])

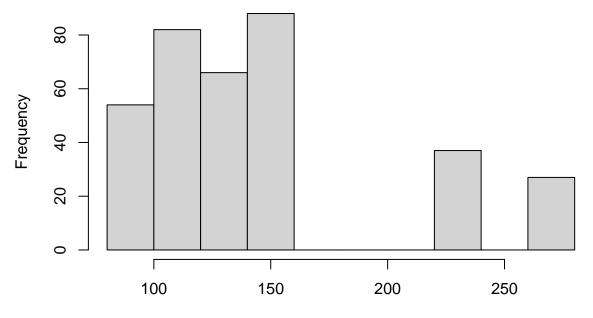
# Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$year == 4]



final\_kbs\$half\_cover\_date[final\_kbs\$year == 4]

hist(final\_kbs\$half\_cover\_date[final\_kbs\$year == 5])

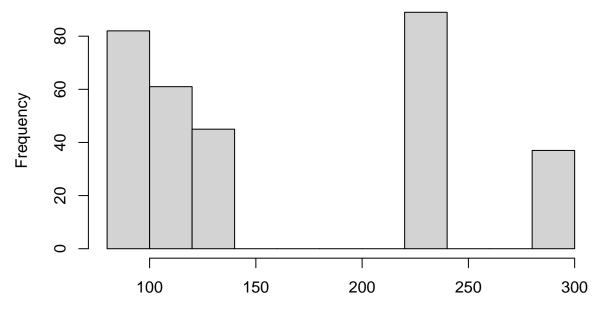
## Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$year == 5]



final\_kbs\$half\_cover\_date[final\_kbs\$year == 5]

hist(final\_kbs\$half\_cover\_date[final\_kbs\$year == 6])

# Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$year == 6]



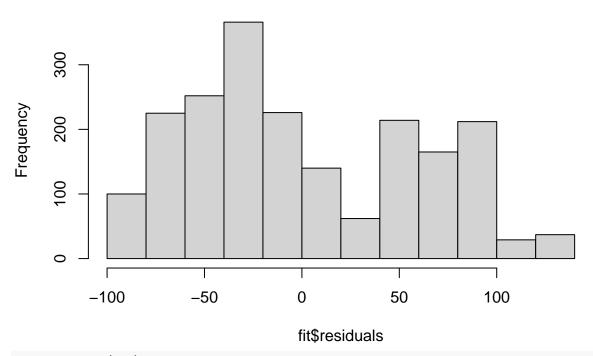
final\_kbs\$half\_cover\_date[final\_kbs\$year == 6]

```
# looks like the 225 spike is from 2018 and 2020
kbs_2018 <- subset(final_kbs, year == 4) # many records on 235
kbs_2020 <- subset(final_kbs, year == 6) # records from 227 & 228</pre>
```

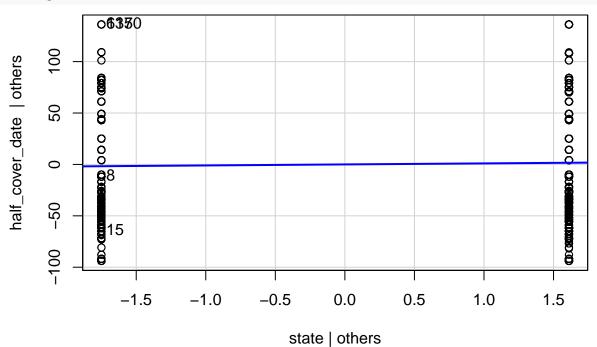
#### Leverage plots?

```
# checking fit for date as a function of state
fit <- lm(half_cover_date~state, data = final_kbs)
hist(fit$residuals)</pre>
```

## Histogram of fit\$residuals

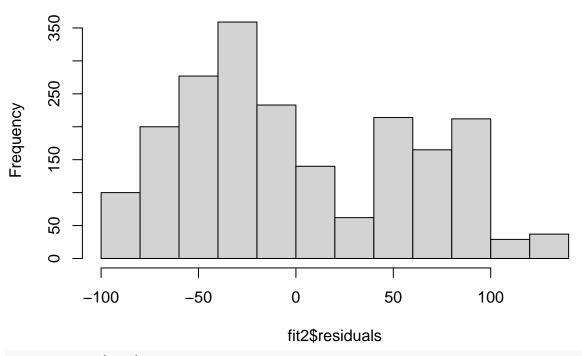


#### leveragePlots(fit)



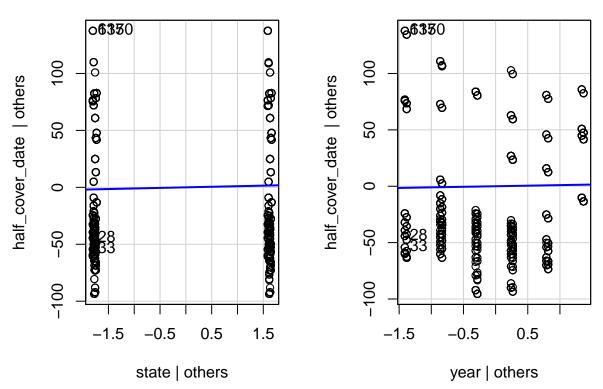
```
# checking fit for date as a function of state and year
fit2 <- lm(half_cover_date~state+year, data = final_kbs)
hist(fit2$residuals)</pre>
```

# Histogram of fit2\$residuals



leveragePlots(fit2)

# Leverage Plots

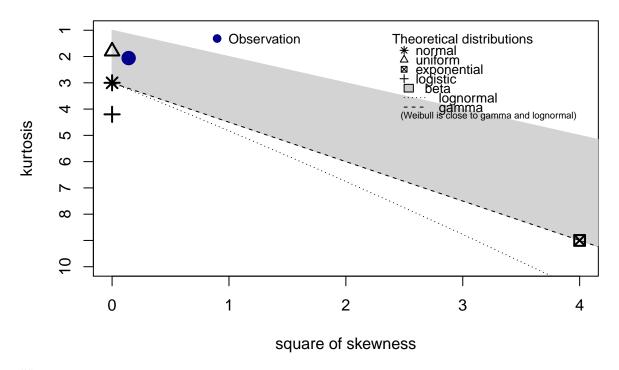


Not normal, and previously attempted transformations don't help (in R script)

#### Seeing what other distribution could fit

```
descdist(final_kbs$half_cover_date, discrete = FALSE)
```

#### **Cullen and Frey graph**

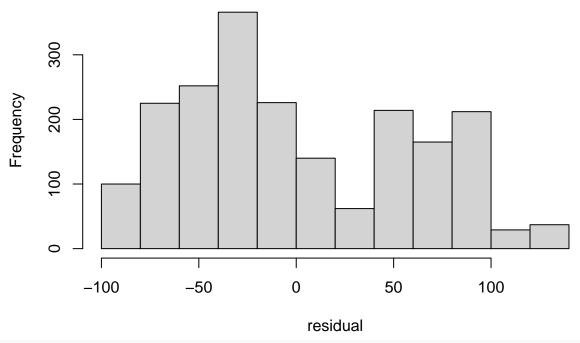


```
## summary statistics
## -----
## min: 59 max: 289
## median: 141
## mean: 152.8644
## estimated sd: 57.73216
## estimated skewness: 0.3763172
## estimated kurtosis: 2.058193
```

#### While uniform looks the closest, I'll try poisson

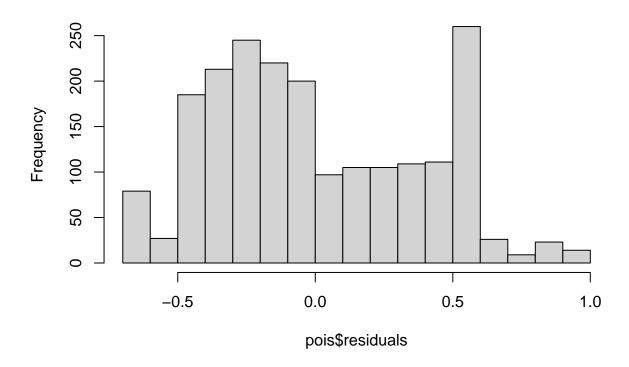
```
fit <- lm(half_cover_date~state, data = final_kbs)
residual <- fit$residuals
hist(residual, main="Raw residuals")</pre>
```

#### Raw residuals



pois <- glm(half\_cover\_date~state, data = final\_kbs, family="poisson")
hist(pois\$residuals, main="Poisson glm residuals")</pre>

# Poisson glm residuals



# Below I try a few different generalized linear models with poisson distribution:

An interaction between state and year, plus insecticide as a fixed effect and species and plot as random effects

```
moda <- glmer(half_cover_date ~ state*year + insecticide + (1|species) + (1|plot),</pre>
              data=final_kbs, family = poisson)
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unide:
## - Rescale variables?
summary(moda)
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
## Family: poisson (log)
## Formula: half_cover_date ~ state * year + insecticide + (1 | species) +
       (1 | plot)
##
##
      Data: final_kbs
##
##
        AIC
                BIC logLik deviance df.resid
##
   42924.6 42963.9 -21455.3 42910.6
##
## Scaled residuals:
##
      Min
               1Q Median
                               3Q
## -9.0161 -2.8352 -0.6425 2.3763 14.4656
##
## Random effects:
## Groups Name
                       Variance Std.Dev.
## species (Intercept) 0.069953 0.26449
            (Intercept) 0.001848 0.04299
## Number of obs: 2028, groups: species, 55; plot, 24
## Fixed effects:
                         Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                         5.020672
                                   0.039631 126.686 < 2e-16 ***
                         0.009389
## statewarmed
                                    0.019347
                                              0.485
                                                       0.6275
                         0.003600
                                    0.001520
                                              2.368
                                                      0.0179 *
## insecticideno_insects 0.028935
                                    0.017936
                                               1.613
                                                       0.1067
## statewarmed:year
                        -0.009698
                                    0.002116 -4.584 4.57e-06 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Correlation of Fixed Effects:
##
              (Intr) sttwrm year
                                   insct
## statewarmed -0.243
## year
              -0.130 0.249
## insctcdn_ns -0.227 -0.002 0.008
## statwrmd:yr 0.089 -0.375 -0.648 0.006
## convergence code: 0
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
```

No interaction between state and year, but with state and insecticide as fixed effects and species and plot as random effects

```
modb <- glmer(half_cover_date ~ state + year + insecticide + (1|species) + (1|plot),</pre>
              data=final_kbs, family = poisson)
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unide:
## - Rescale variables?
summary(modb)
## Generalized linear mixed model fit by maximum likelihood (Laplace
     Approximation) [glmerMod]
  Family: poisson (log)
## Formula: half_cover_date ~ state + year + insecticide + (1 | species) +
##
       (1 | plot)
##
     Data: final kbs
##
##
       AIC
                BIC logLik deviance df.resid
   42943.6 42977.3 -21465.8 42931.6
##
##
## Scaled residuals:
##
               10 Median
                               3Q
## -9.0054 -2.8448 -0.6575 2.3929 14.2130
##
## Random effects:
## Groups Name
                       Variance Std.Dev.
## species (Intercept) 0.069954 0.26449
## plot
            (Intercept) 0.001847 0.04297
## Number of obs: 2028, groups: species, 55; plot, 24
##
## Fixed effects:
##
                          Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                         5.0367370 0.0394718 127.603
                                                        <2e-16 ***
## statewarmed
                        -0.0238778 0.0179289 -1.332
                                                         0.183
                        -0.0009129 0.0011579 -0.788
                                                         0.430
## insecticideno_insects 0.0294349 0.0179300
                                                1.642
                                                         0.101
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Correlation of Fixed Effects:
               (Intr) sttwrm year
## statewarmed -0.227
## year
              -0.096 0.007
## insctcdn_ns -0.228 0.001 0.015
## convergence code: 0
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
```

#### State and insecticide as fixed effects & year, species and plot as random effects

```
modc <- glmer(half_cover_date ~ state + insecticide + (1|year) + (1|species) + (1|plot),</pre>
              data=final kbs, family = poisson)
summary(modc)
## Generalized linear mixed model fit by maximum likelihood (Laplace
     Approximation) [glmerMod]
   Family: poisson (log)
## Formula: half_cover_date ~ state + insecticide + (1 | year) + (1 | species) +
       (1 | plot)
##
##
      Data: final_kbs
##
##
       AIC
                BIC
                      logLik deviance df.resid
   41436.8 41470.5 -20712.4 41424.8
##
##
## Scaled residuals:
##
      Min
               1Q Median
                                3Q
## -9.3418 -2.8134 -0.5215 2.1783 13.1976
## Random effects:
## Groups Name
                       Variance Std.Dev.
## species (Intercept) 0.068408 0.26155
            (Intercept) 0.001820 0.04266
## year
            (Intercept) 0.005867 0.07659
## Number of obs: 2028, groups: species, 55; plot, 24; year, 6
##
## Fixed effects:
                         Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                         5.01945
                                    0.04991 100.577
                                                       <2e-16 ***
## statewarmed
                         -0.02683
                                     0.01781 -1.507
                                                       0.1319
## insecticideno_insects 0.03093
                                    0.01781
                                               1.737
                                                       0.0824 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Correlation of Fixed Effects:
               (Intr) sttwrm
## statewarmed -0.178
## insctcdn ns -0.178 0.001
```

Because no distributions seems to match well, I'll try a Friedman's test

```
#friedman_kbs <- final_kbs %>%
# friedman_test(half_cover_date ~ state)
```

Error: Must extract column with a single valid subscript. x Subscript var can't be NA

Can't figure out what this means

If I include the blocks portion of the formula (from the documentation) I get this error

```
#friedman_kbs <- final_kbs %>%
# friedman_test(half_cover_date ~ state / plot)
```

Error in friedman.test.default(c(141L, 202L, 122L, 101L, 127L, 120L, 197L, : not an unreplicated complete block design

#### Permanova?

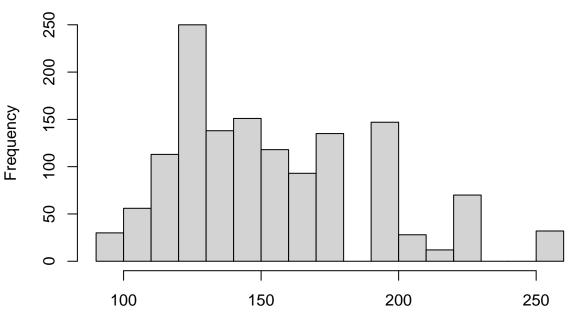
```
per1 <- adonis2(final_kbs$half_cover_date ~ state*year + insecticide, data = final_kbs)</pre>
per1
## Permutation test for adonis under reduced model
## Terms added sequentially (first to last)
## Permutation: free
## Number of permutations: 999
##
## adonis2(formula = final_kbs$half_cover_date ~ state * year + insecticide, data = final_kbs)
                Df SumOfSqs
##
                               R2
                                         F Pr(>F)
## state
                1
                     0.051 0.00074 1.5046 0.212
                     0.487 0.00698 14.2700 0.001 ***
## year
                1
               1 0.168 0.00241 4.9326 0.021 *
## insecticide
               1
                     0.025 0.00036 0.7290 0.405
## state:year
              2023
                     69.058 0.98951
## Residual
## Total
              2027
                     69.790 1.00000
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

#### UMBS

## Checking for normality

hist(final\_umbs\$half\_cover\_date)

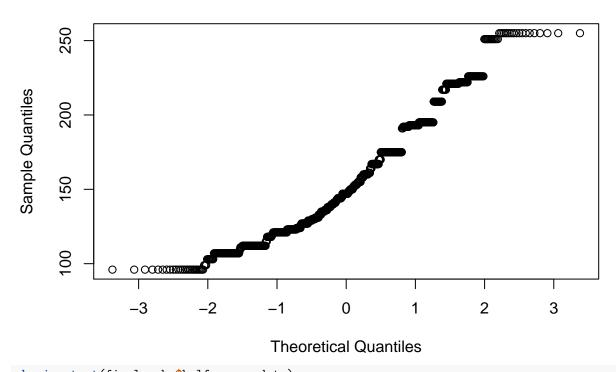
# Histogram of final\_umbs\$half\_cover\_date



final\_umbs\$half\_cover\_date

qqnorm(final\_umbs\$half\_cover\_date)

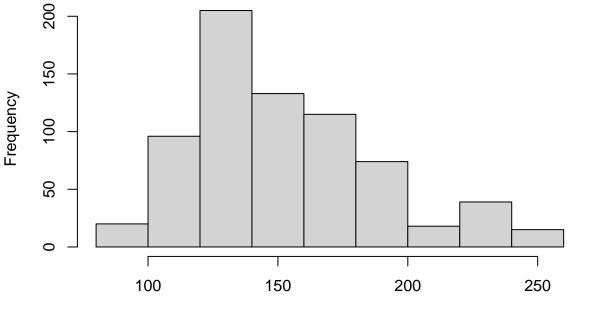
#### Normal Q-Q Plot



```
shapiro.test(final_umbs$half_cover_date)

##
## Shapiro-Wilk normality test
##
## data: final_umbs$half_cover_date
## W = 0.94754, p-value < 2.2e-16
hist(final_umbs$half_cover_date[final_kbs$state == "ambient"])</pre>
```

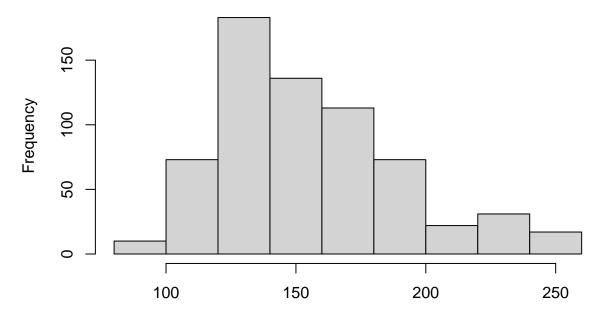
#### Histogram of final\_umbs\$half\_cover\_date[final\_kbs\$state == "ambier



final\_umbs\$half\_cover\_date[final\_kbs\$state == "ambient"]

hist(final\_umbs\$half\_cover\_date[final\_kbs\$state == "warmed"])

#### Histogram of final\_umbs\$half\_cover\_date[final\_kbs\$state == "warme



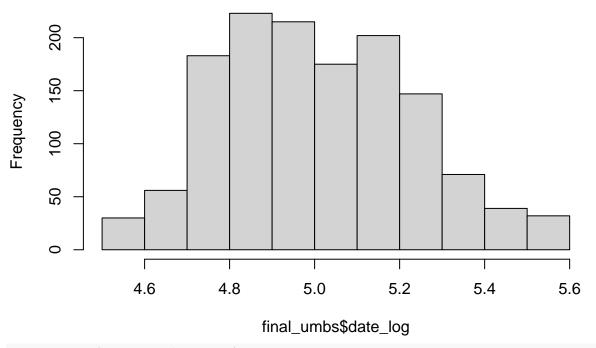
final\_umbs\$half\_cover\_date[final\_kbs\$state == "warmed"]

These look pretty good

#### Trying log transformation

```
final_umbs$date_log <- log(final_umbs$half_cover_date)
hist(final_umbs$date_log)</pre>
```

## Histogram of final\_umbs\$date\_log



shapiro.test(final\_umbs\$date\_log)

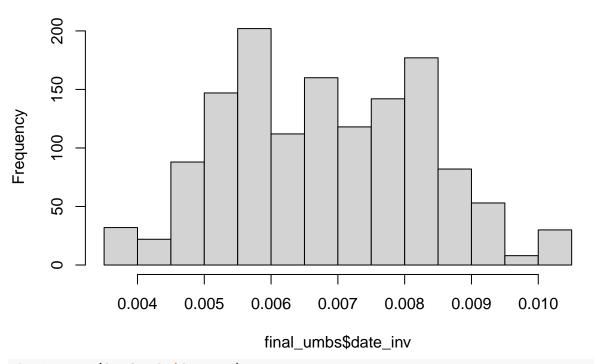
```
##
## Shapiro-Wilk normality test
##
## data: final_umbs$date_log
## W = 0.97728, p-value = 6.765e-14
```

I think this looks good but shapiro-wilk is lower than 0.05

#### Trying inverse tranformation

```
final_umbs$date_inv <- 1/(final_umbs$half_cover_date)
hist(final_umbs$date_inv)</pre>
```

# Histogram of final\_umbs\$date\_inv



shapiro.test(final\_umbs\$date\_inv)

```
##
## Shapiro-Wilk normality test
##
## data: final_umbs$date_inv
## W = 0.97993, p-value = 6.679e-13
```

This also looks good but is also still low for shapiro-wilk