

warmXtrophic Project: Greenup Analyses

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Script Details:

```
# this portion of the script won't knit, so its set to eval=F

script_tbl <- data.frame(Item = c("OVERVIEW", "COLLABORATORS",
  "REQUIRES", "DATA INPUT", "DATA OUTPUT", "NOTES"), Details = c("This script explores and analyses t
  "Moriah Young, Mark Hammond, Pat Bills", "Prior to running this script, make sure plant_comp_clean_1
  "Data imported as csv files from shared Google drive 'SpaCE_Lab_warmXtrophic' plant comp folder",
  "... a brief description of the data output from through the script, including what format it's in"
  "Each row in 'greenup' is the date at which spp_half_cover_date was recorded, per species. The 'gre

kbl(script_tbl) %>% kable_paper(full_width = F) %>% column_spec(1,
  bold = T, border_right = T) %>% column_spec(2, width = "30em",
  background = "lightblue")

metadata_tbl <- data.frame(Variable = c("spp_half_cover_date",
  "plot_half_cover_date", "state"), Definition = c("date at which 50% of a species max cover was reach
  "the date at which 50% of a plot's max cover was reached (per plot, per year)",
  "describes each treatment: warmed or ambient"))

kbl(metadata_tbl) %>% kable_paper(full_width = F) %>% column_spec(1,
  bold = T, border_right = T) %>% column_spec(2, width = "30em",
  background = "lightyellow")

# Clear all existing data
rm(list = ls())

# Load packages
library(tidyverse)
library(ggplot2)
library(lmerTest)
library(olsrr)
library(predictmeans)
library(car)
library(fitdistrplus)
library(ggpubr)
library(rstatix)
library(vegan)
library(interactions)
library(sjPlot)
library(effects)
library(glmmTMB)
```

```

library(bbmle)
library(emmeans)
# install.packages('TMB', type='source')

# Set ggplot2 plots to bw: see here for more options:
# http://www.sthda.com/english/wiki/ggplot2-themes-and-background-colors-the-3-elements
theme_set(theme_bw(base_size = 14))

# Get data
L1_dir <- Sys.getenv("L1DIR")
L2_dir <- Sys.getenv("L2DIR")
greenup <- read.csv(file.path(L2_dir, "greenup/final_greenup_species_L2.csv")) # spp level greenup data
greenup <- greenup %>% select(-X) # get rid of 'X' column that shows up
greenupp <- read.csv(file.path(L2_dir, "greenup/final_greenup_plot_L2.csv")) # plot level greenup data
greenupp <- greenupp %>% select(-X) # get rid of 'X' column that shows up

# check variable types
str(greenup)

## 'data.frame': 2026 obs. of 18 variables:
## $ site : chr "kbs" "kbs" "kbs" "kbs" ...
## $ plot : chr "A1" "A1" "A1" "A1" ...
## $ year : int 2016 2017 2018 2019 2020 2016 2017 2016 2017 2018 ...
## $ species : chr "Acmi" "Acmi" "Acmi" "Acmi" ...
## $ spp_half_cover_date: int 197 101 122 120 127 88 108 97 99 127 ...
## $ min_green_date : int 81 80 122 120 107 81 108 85 80 127 ...
## $ treatment_key : chr "A0" "A0" "A0" "A0" ...
## $ state : chr "ambient" "ambient" "ambient" "ambient" ...
## $ insecticide : chr "no_insects" "no_insects" "no_insects" "no_insects" ...
## $ scientific_name : chr "Achillea millefolium" "Achillea millefolium" "Achillea millefolium" "A
## $ common_name : chr "common yarrow" "common yarrow" "common yarrow" "common yarrow" ...
## $ USDA_species : chr "ACMI2" "ACMI2" "ACMI2" "ACMI2" ...
## $ LTER_species : chr "ACHMI" "ACHMI" "ACHMI" "ACHMI" ...
## $ origin : chr "Native" "Native" "Native" "Native" ...
## $ group : chr "Dicot" "Dicot" "Dicot" "Dicot" ...
## $ family : chr "Fabaceae" "Fabaceae" "Fabaceae" "Fabaceae" ...
## $ duration : chr "Biennial" "Biennial" "Biennial" "Biennial" ...
## $ growth_habit : chr "Forb" "Forb" "Forb" "Forb" ...

# Order warm and ambient so that warm shows up first in
# plotting (and is default is red = warm; blue = ambient).
# First make it a factor.
greenup$state <- as.factor(greenup$state)
levels(greenup$state)

## [1] "ambient" "warmed"

greenup$state <- factor(greenup$state, levels(greenup$state)[c(2,
1)])
levels(greenup$state)

## [1] "warmed" "ambient"

greenupp$state <- as.factor(greenupp$state)
levels(greenupp$state)

```

```
## [1] "ambient" "warmed"
greenupp$state <- factor(greenupp$state, levels(greenupp$state)[c(2,
1)])
levels(greenupp$state)

## [1] "warmed" "ambient"

# adding sequential year variable starting at 1: this is
# because 2016... are large numbers compare with other values
# in the dataset. We can always label axes with these real
# years.
greenup$year_factor[greenup$year == 2016] <- 1
greenup$year_factor[greenup$year == 2017] <- 2
greenup$year_factor[greenup$year == 2018] <- 3
greenup$year_factor[greenup$year == 2019] <- 4
greenup$year_factor[greenup$year == 2020] <- 5

greenupp$year_factor[greenupp$year == 2016] <- 1
greenupp$year_factor[greenupp$year == 2017] <- 2
greenupp$year_factor[greenupp$year == 2018] <- 3
greenupp$year_factor[greenupp$year == 2019] <- 4
greenupp$year_factor[greenupp$year == 2020] <- 5

# create dataframes for kbs and umbs - remember that these
# contain species within plots
green_kbs <- subset(greenup, site == "kbs")
green_umbs <- subset(greenup, site == "umbs")

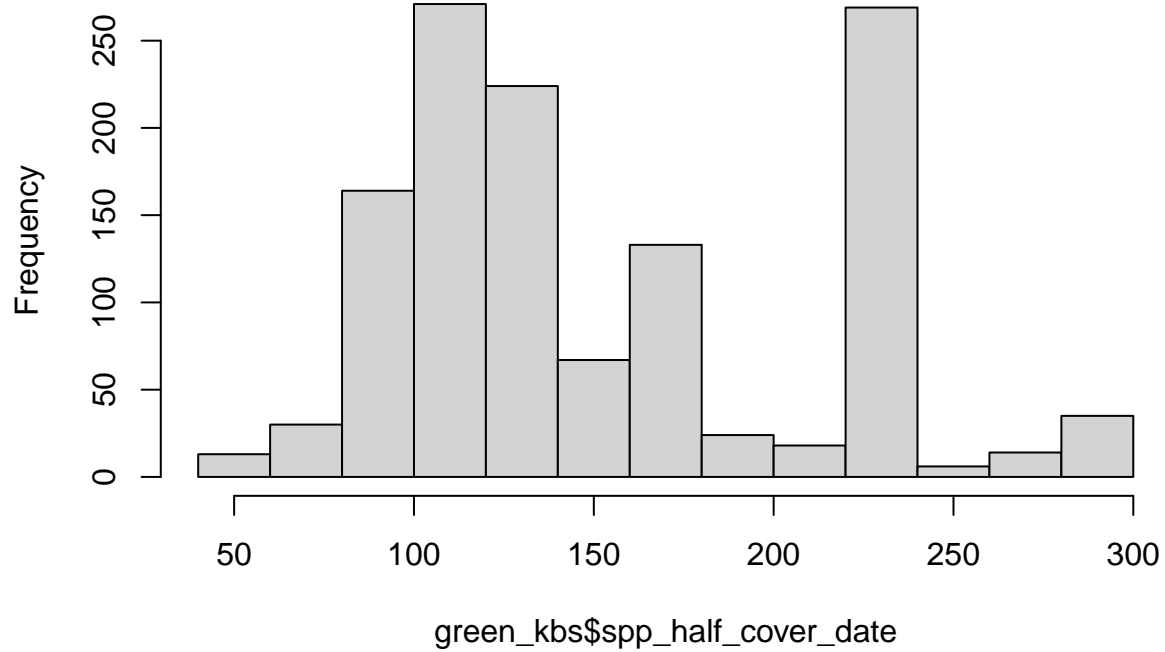
green_kbsp <- subset(greenupp, site == "kbs")
green_umbsp <- subset(greenupp, site == "umbs")
```

Data Exploration:

First, checking for normality in raw data. It's not going to tell you about normality once you fit a model to these data - that's when you really need to investigate the residuals.

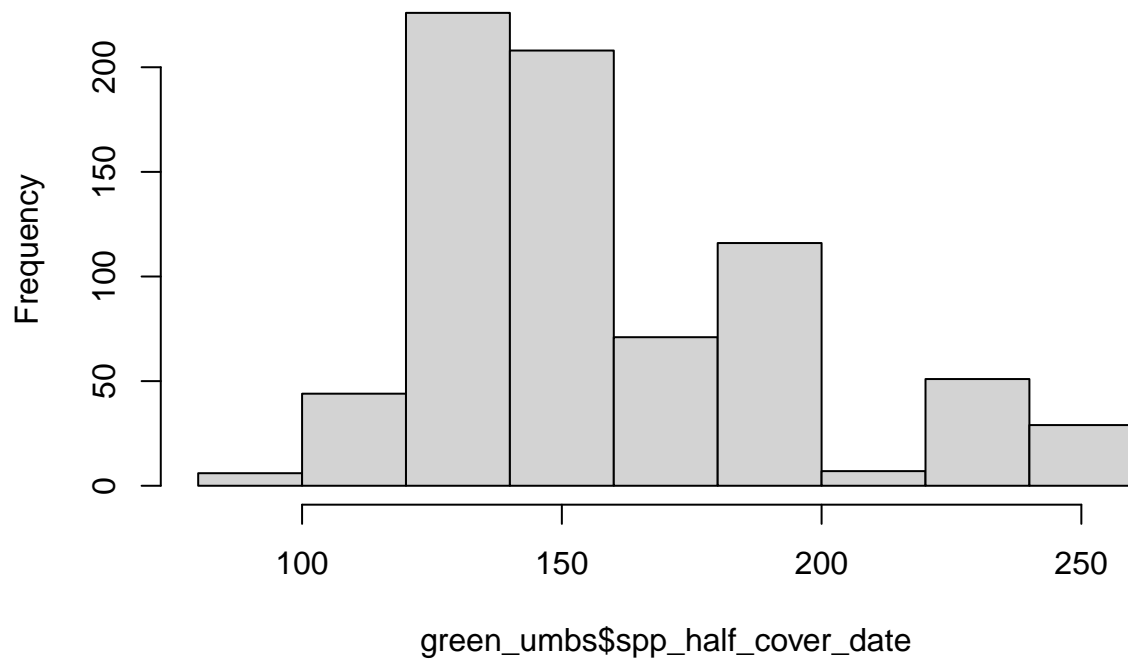
```
# species level
hist(green_kbs$spp_half_cover_date)
```

Histogram of green_kbs\$spp_half_cover_date



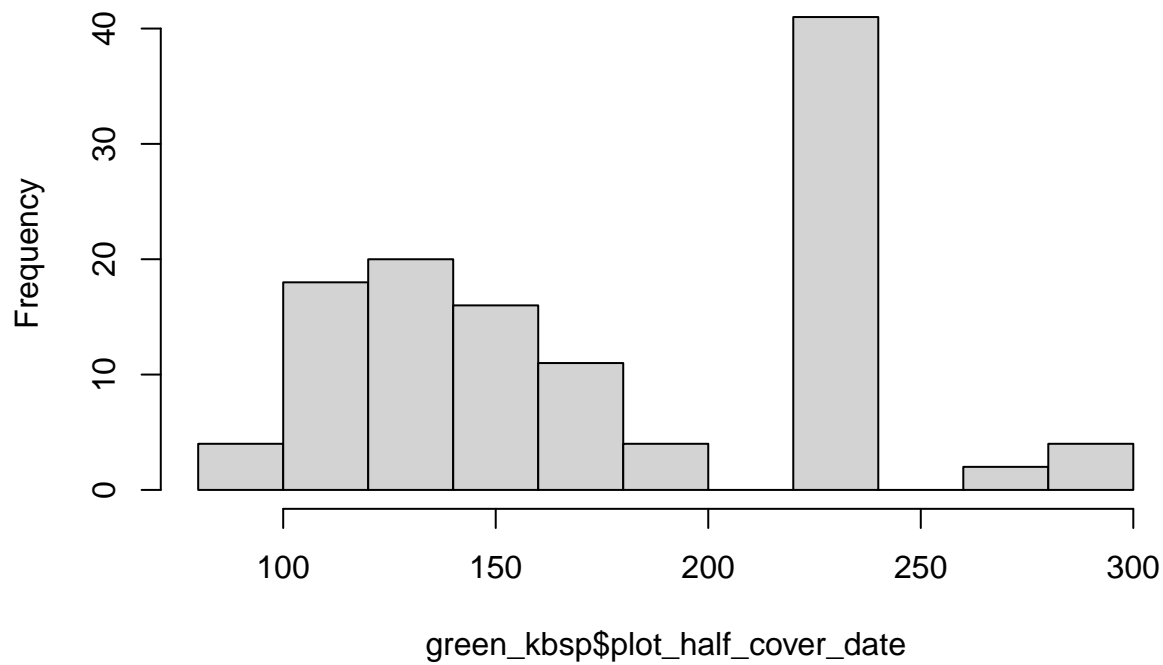
```
hist(green_umbs$spp_half_cover_date)
```

Histogram of green_umbs\$spp_half_cover_date



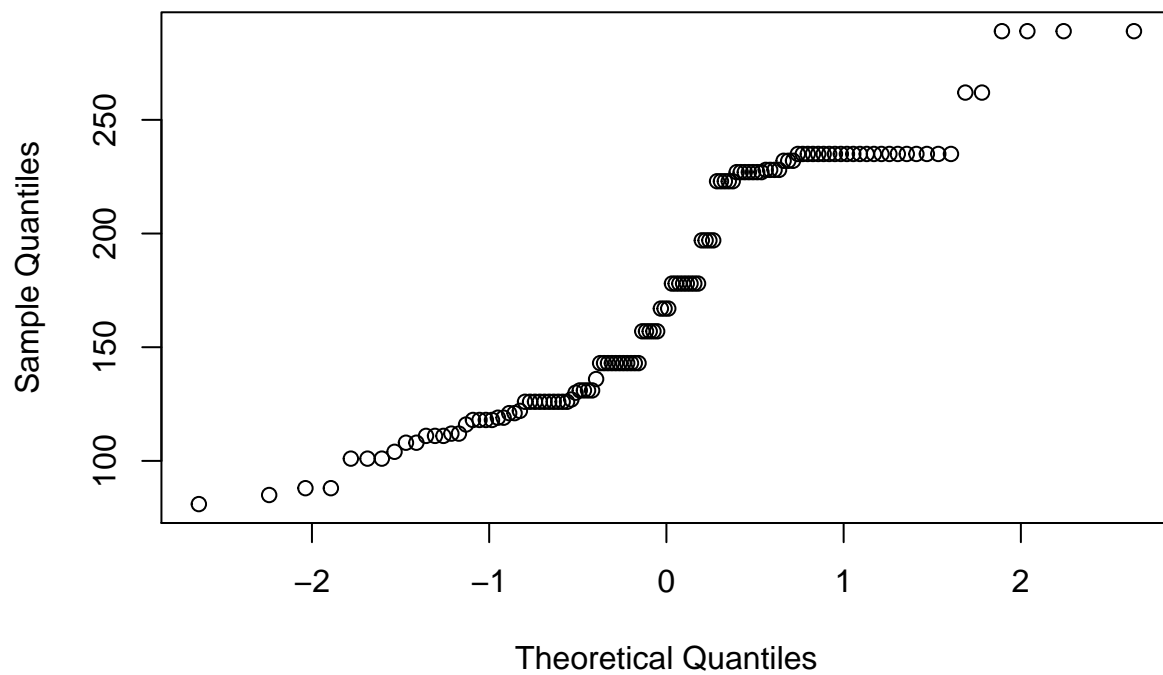
```
# plot level  
hist(green_kbsp$plot_half_cover_date)
```

Histogram of green_kbsp\$plot_half_cover_date



```
qqnorm(green_kbsp$plot_half_cover_date)
```

Normal Q-Q Plot

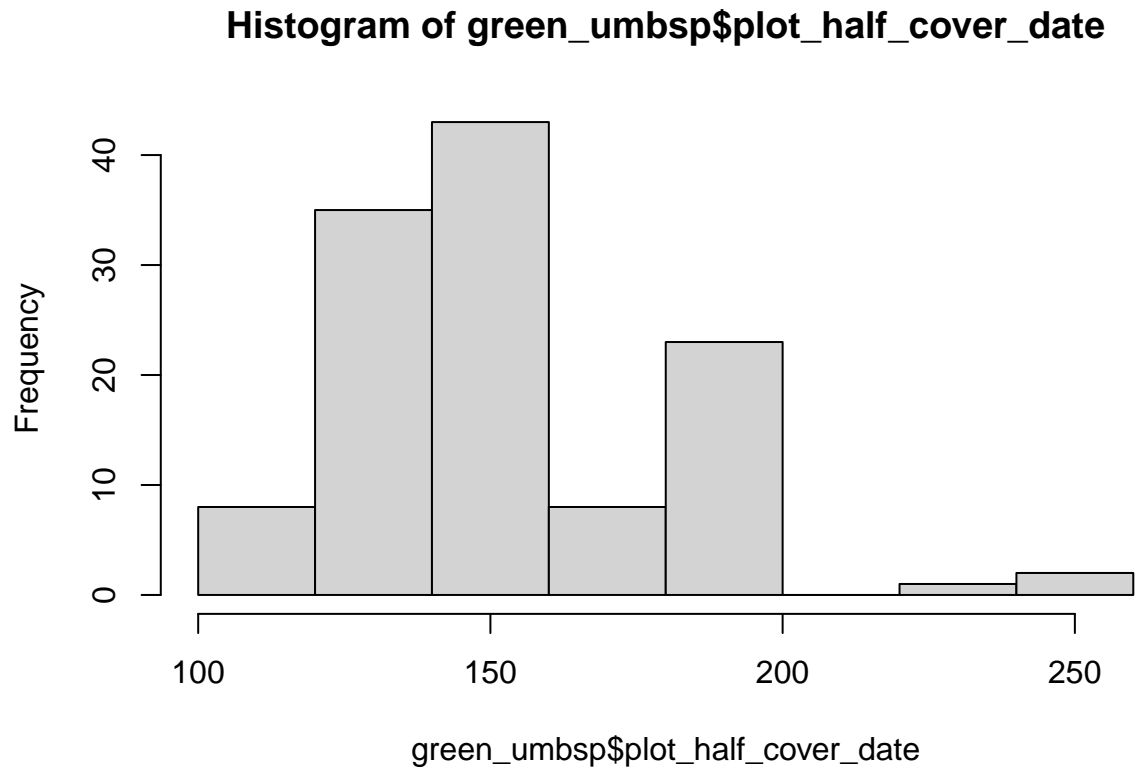


```
shapiro.test(green_kbsp$plot_half_cover_date)
```

```
##
```

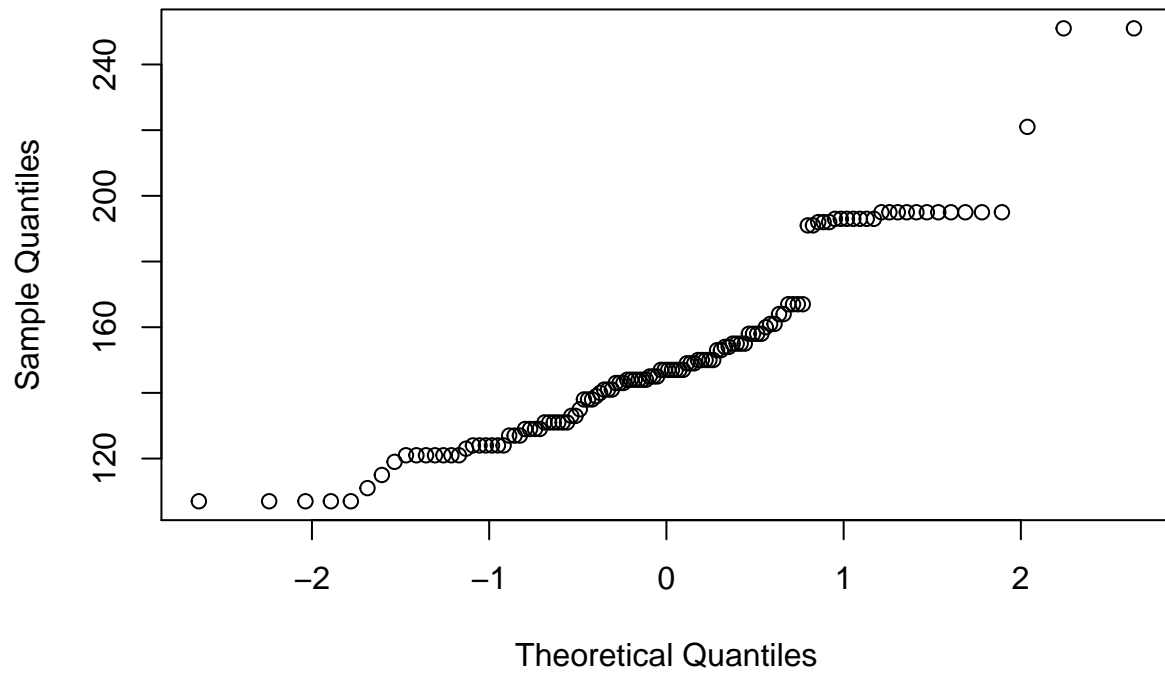
```
## Shapiro-Wilk normality test
##
## data: green_kbsp$plot_half_cover_date
## W = 0.90721, p-value = 4.673e-07
```

```
hist(green_umbsp$plot_half_cover_date)
```



```
qqnorm(green_umbsp$plot_half_cover_date)
```

Normal Q-Q Plot

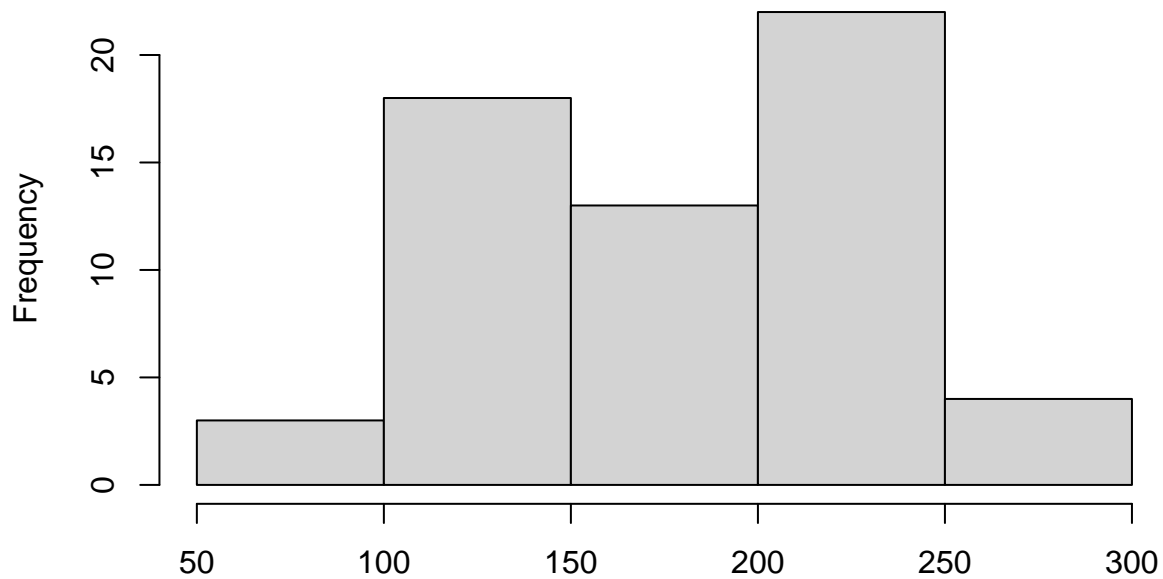


```
shapiro.test(green_umbsp$plot_half_cover_date)
```

```
##  
##  Shapiro-Wilk normality test  
##  
## data:  green_umbsp$plot_half_cover_date  
## W = 0.91432, p-value = 1.143e-06
```

```
# histograms for each treatment separately - plot level  
hist(green_kbsp$plot_half_cover_date[green_kbsp$state == "ambient"])
```

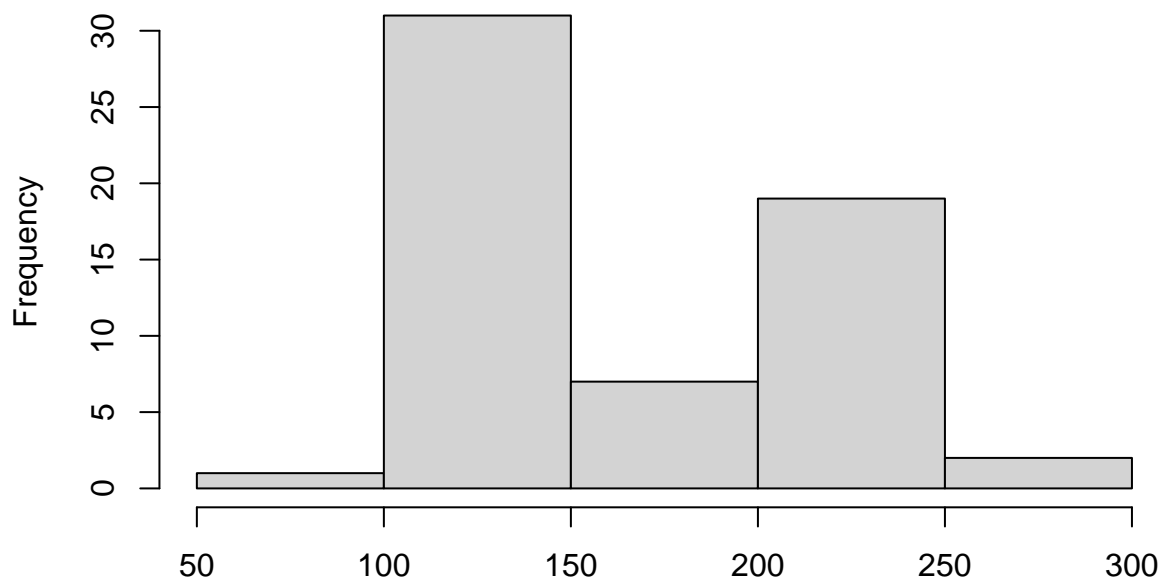
histogram of green_kbsp\$plot_half_cover_date[green_kbsp\$state == "ambient"]



green_kbsp\$plot_half_cover_date[green_kbsp\$state == "ambient"]

```
hist(green_kbsp$plot_half_cover_date[green_kbsp$state == "warmed"])
```

histogram of green_kbsp\$plot_half_cover_date[green_kbsp\$state == "warmed"]

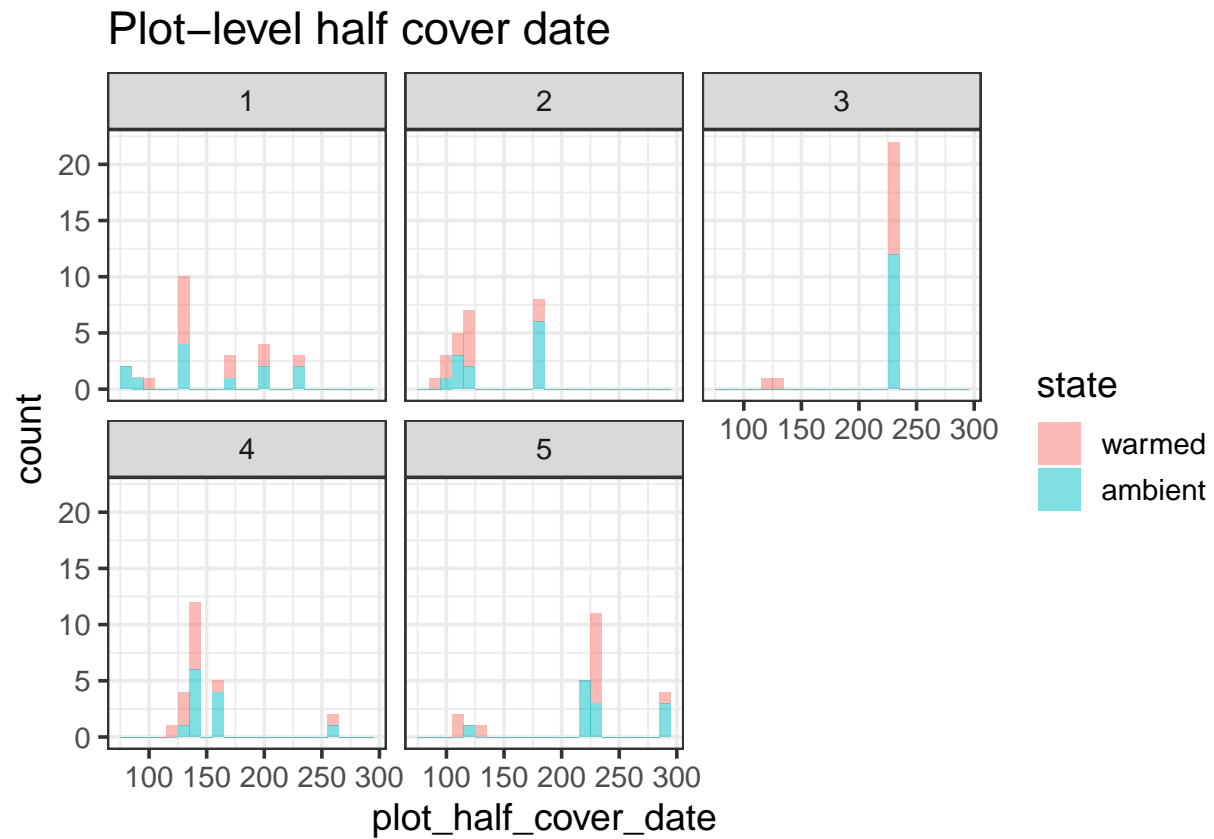


green_kbsp\$plot_half_cover_date[green_kbsp\$state == "warmed"]

```
# histograms for each year - look at them together:
p1 <- ggplot(data = green_kbsp, aes(x = plot_half_cover_date,
  fill = state)) + geom_histogram(alpha = 0.5, binwidth = 10)
```

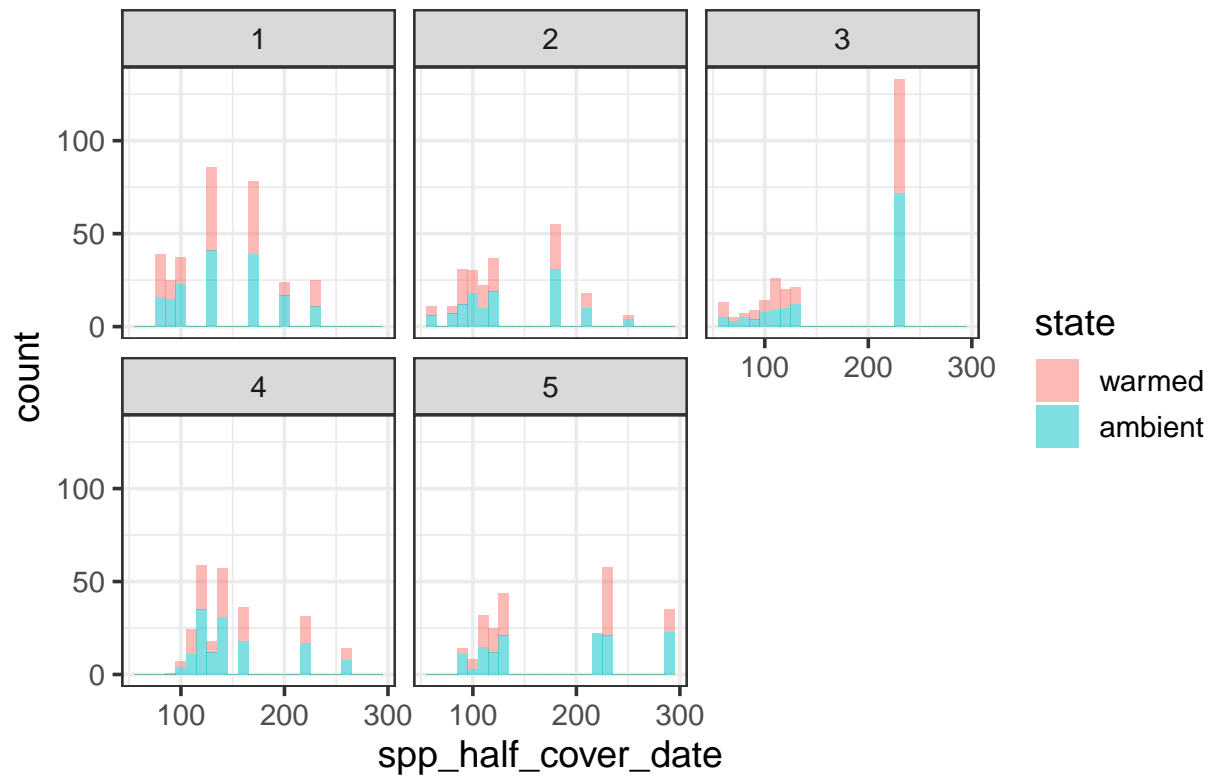


```
p1 + facet_wrap(~year_factor) + labs(title = "Plot-level half cover date")
```



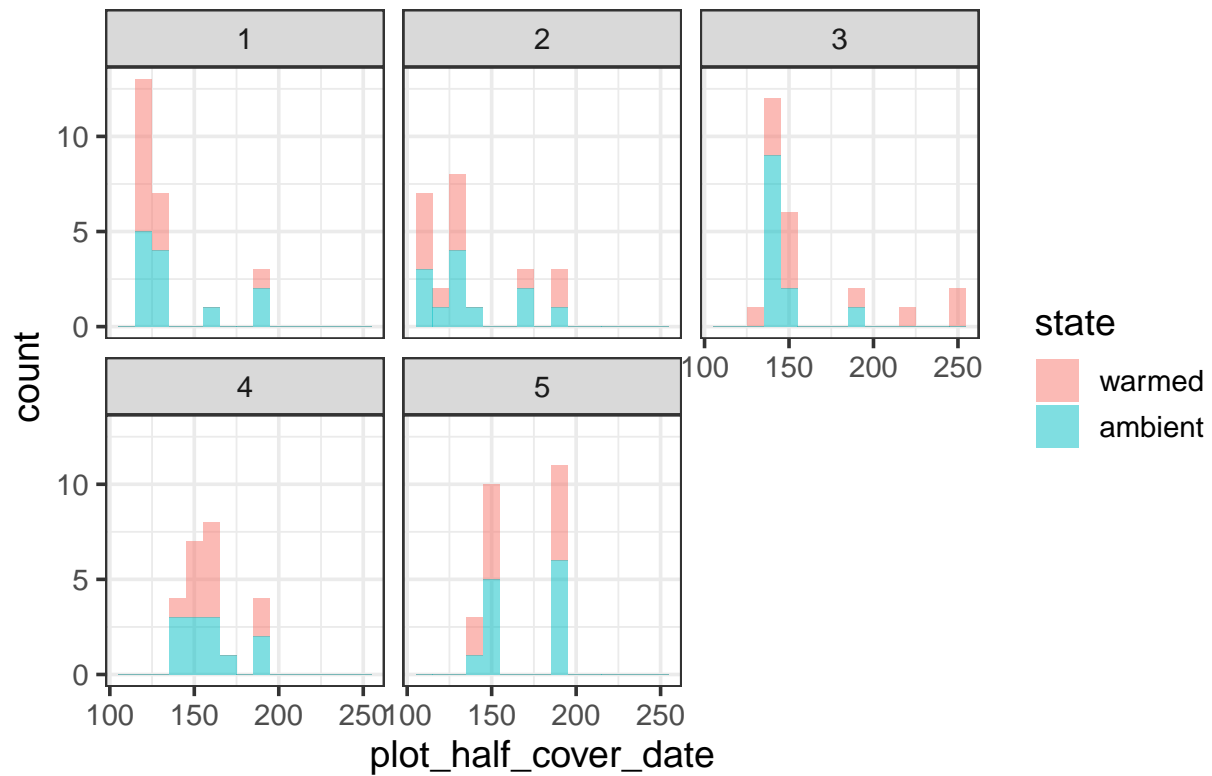
```
p1.2 <- ggplot(data = green_kbs, aes(x = spp_half_cover_date,
  fill = state)) + geom_histogram(alpha = 0.5, binwidth = 10)
p1.2 + facet_wrap(~year_factor) + labs(title = "Species-level half cover date")
```

Species-level half cover date



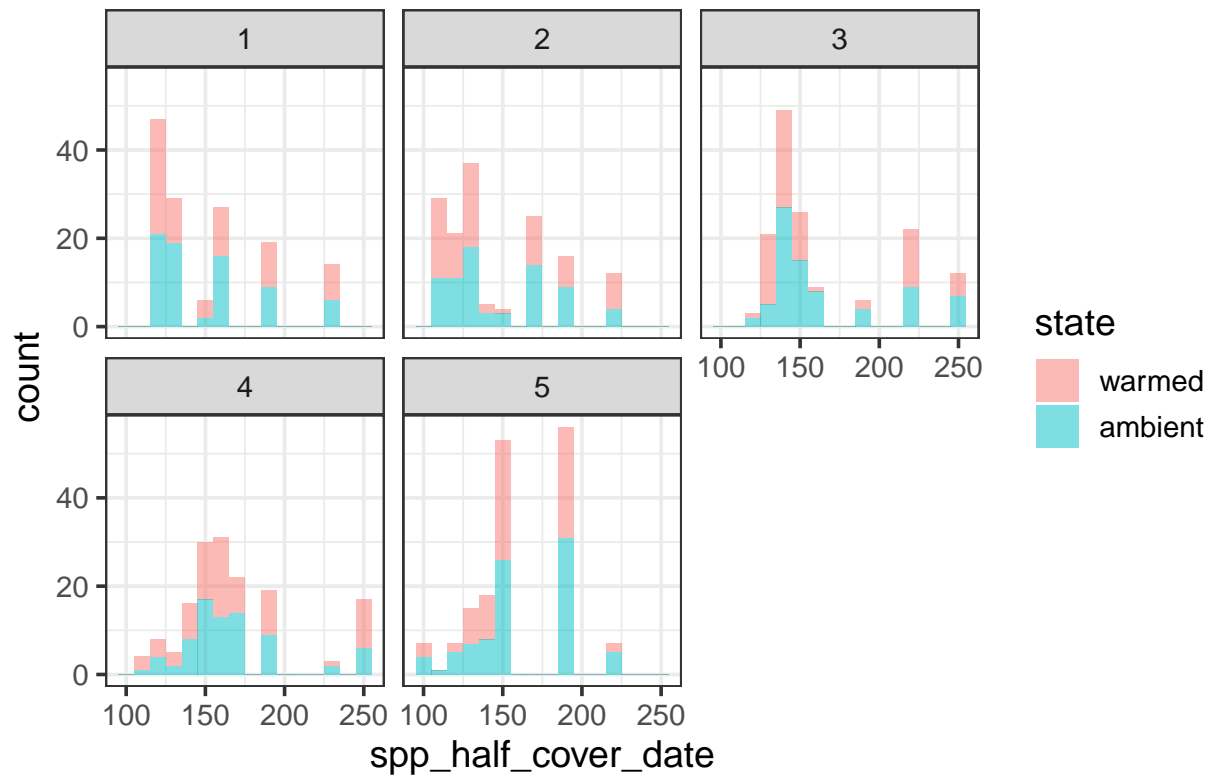
```
p1u <- ggplot(data = green_umbsp, aes(x = plot_half_cover_date,
  fill = state)) + geom_histogram(alpha = 0.5, binwidth = 10)
p1u + facet_wrap(~year_factor) + labs(title = "Plot-level half cover date")
```

Plot-level half cover date

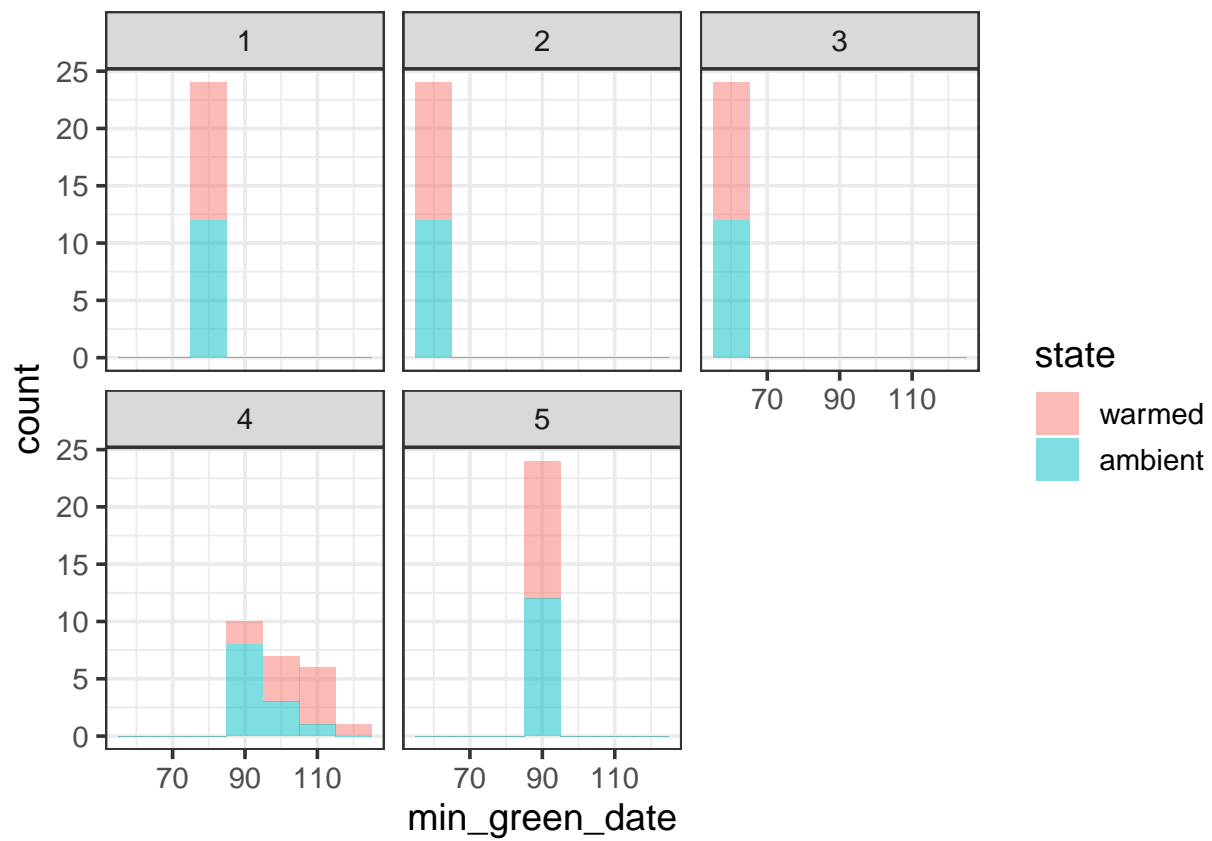


```
p1.2u <- ggplot(data = green_umbs, aes(x = spp_half_cover_date,
  fill = state)) + geom_histogram(alpha = 0.5, binwidth = 10)
p1.2u + facet_wrap(~year_factor) + labs(title = "Species-level half cover date")
```

Species-level half cover date

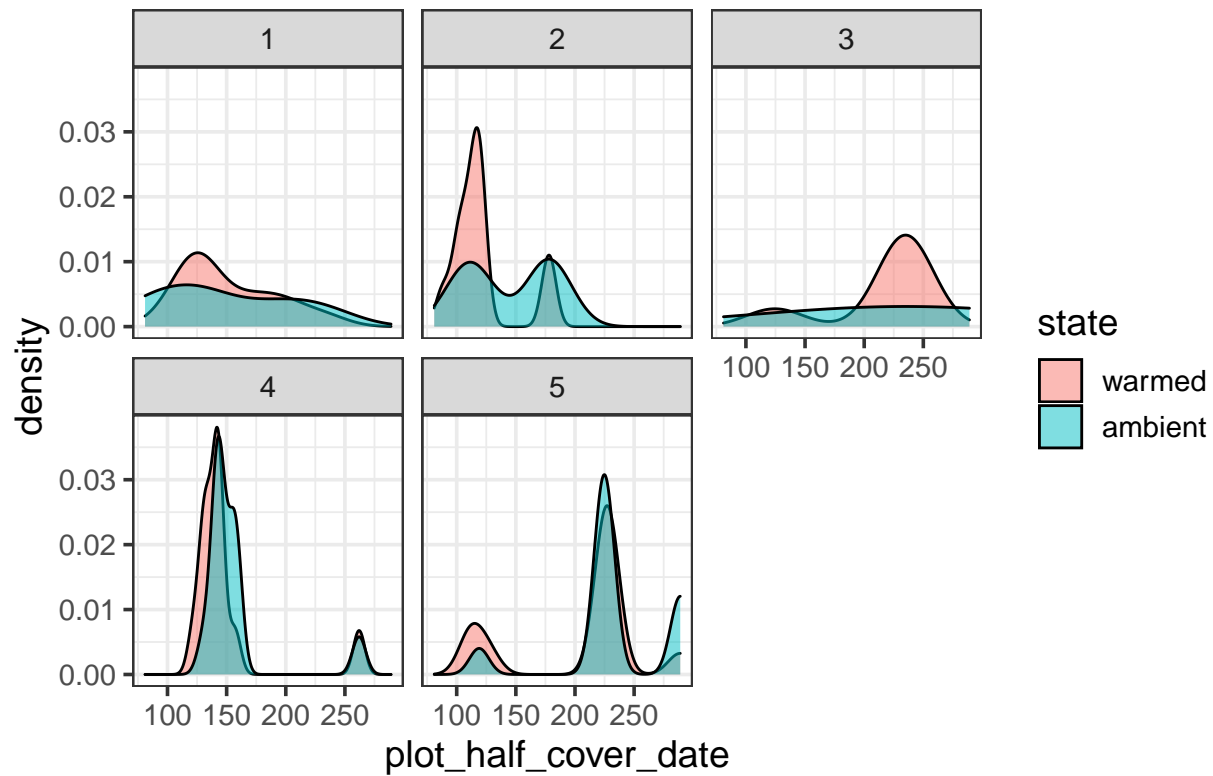


```
# this will just show sampling date artifact
p2 <- ggplot(data = green_kbsp, aes(x = min_green_date, fill = state)) +
  geom_histogram(alpha = 0.5, binwidth = 10)
p2 + facet_wrap(~year_factor)
```



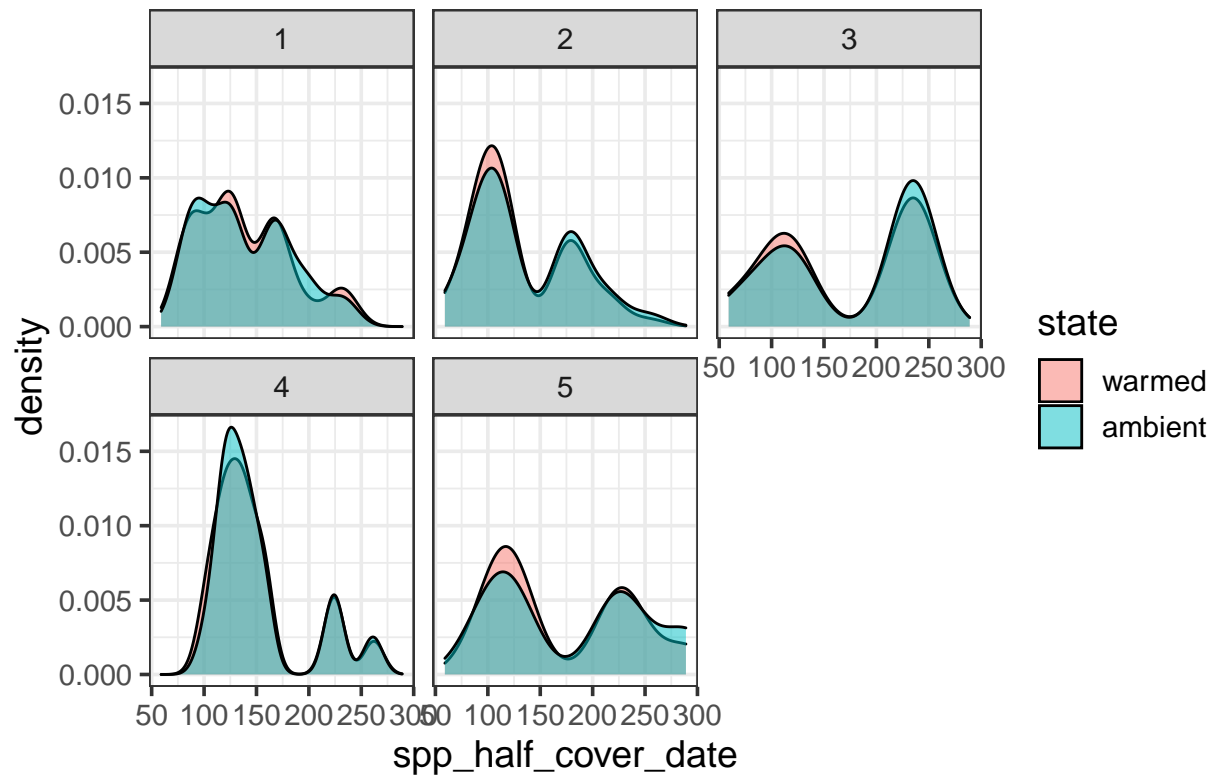
```
# Density plot
p3 <- ggplot(data = green_kbsp, aes(x = plot_half_cover_date,
  fill = state)) + geom_density(alpha = 0.5)
p3 + facet_wrap(~year_factor) + labs(title = "Plot-level half cover date")
```

Plot-level half cover date



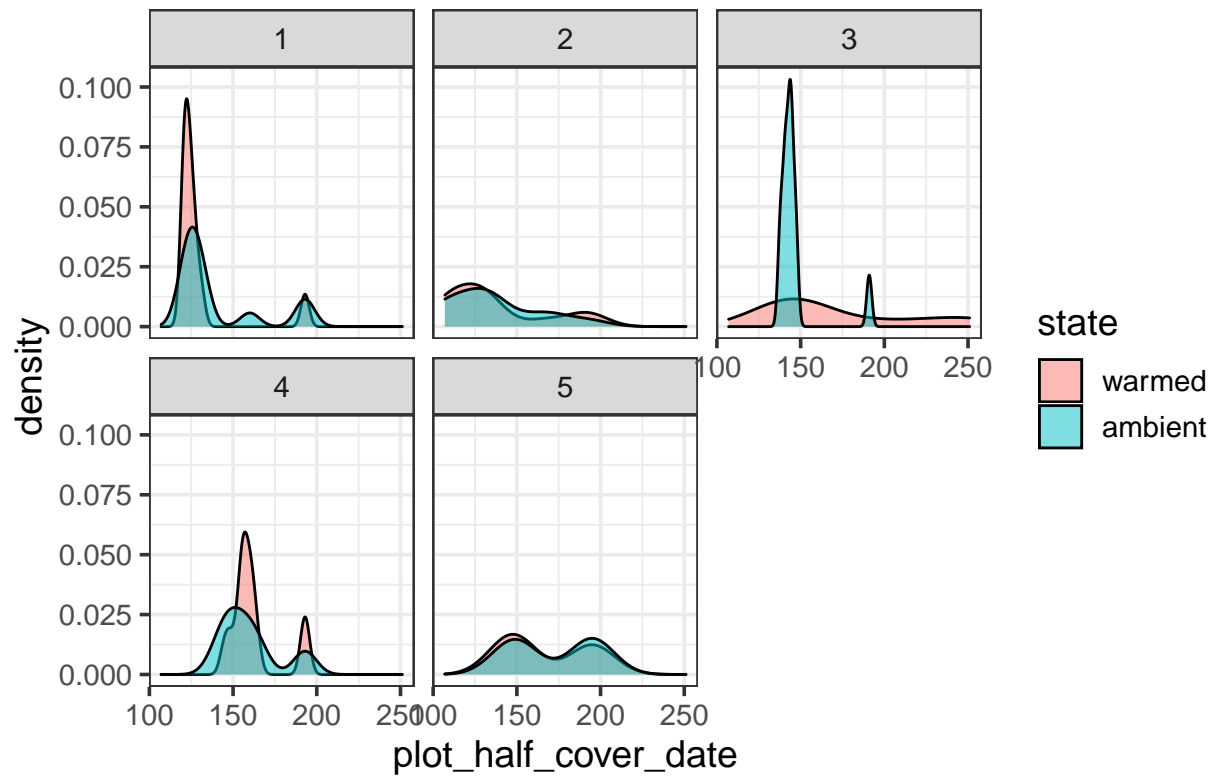
```
p3.2 <- ggplot(data = green_kbs, aes(x = spp_half_cover_date,
  fill = state)) + geom_density(alpha = 0.5)
p3.2 + facet_wrap(~year_factor) + labs(title = "Species-level half cover date")
```

Species-level half cover date



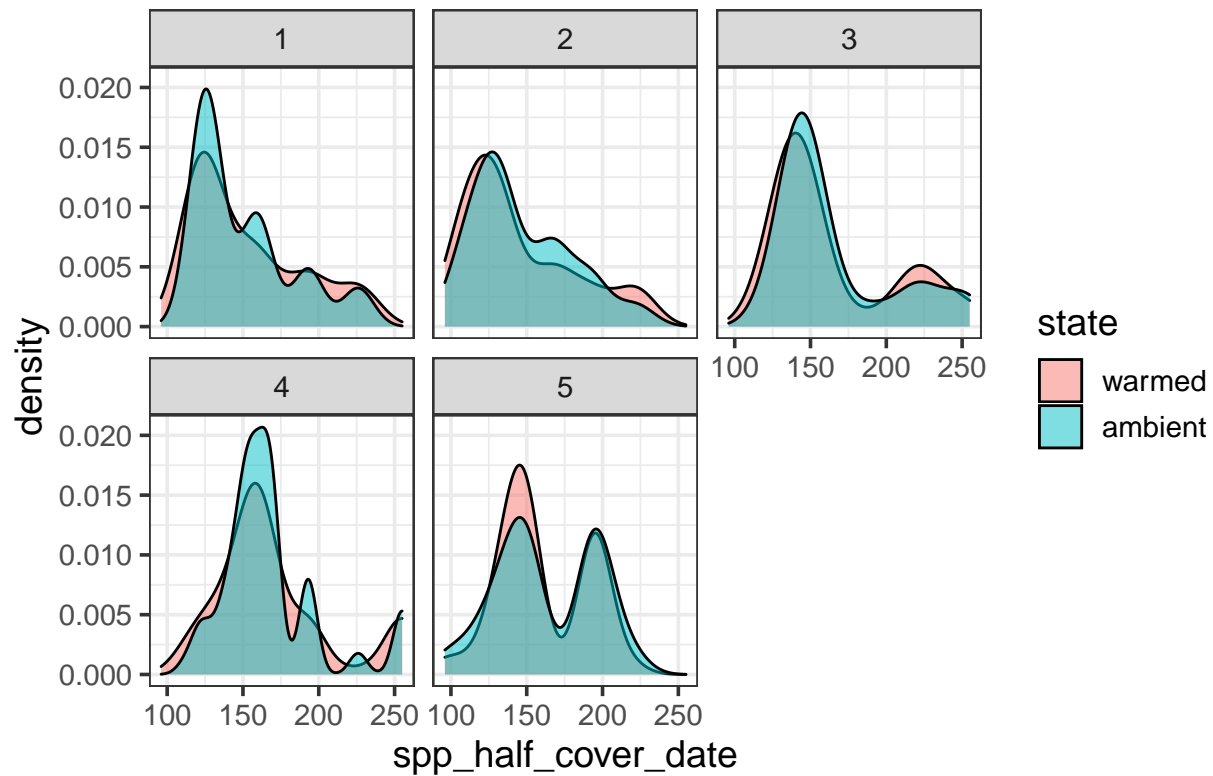
```
p3u <- ggplot(data = green_umbsp, aes(x = plot_half_cover_date,
  fill = state)) + geom_density(alpha = 0.5)
p3u + facet_wrap(~year_factor) + labs(title = "Plot-level half cover date")
```

Plot-level half cover date

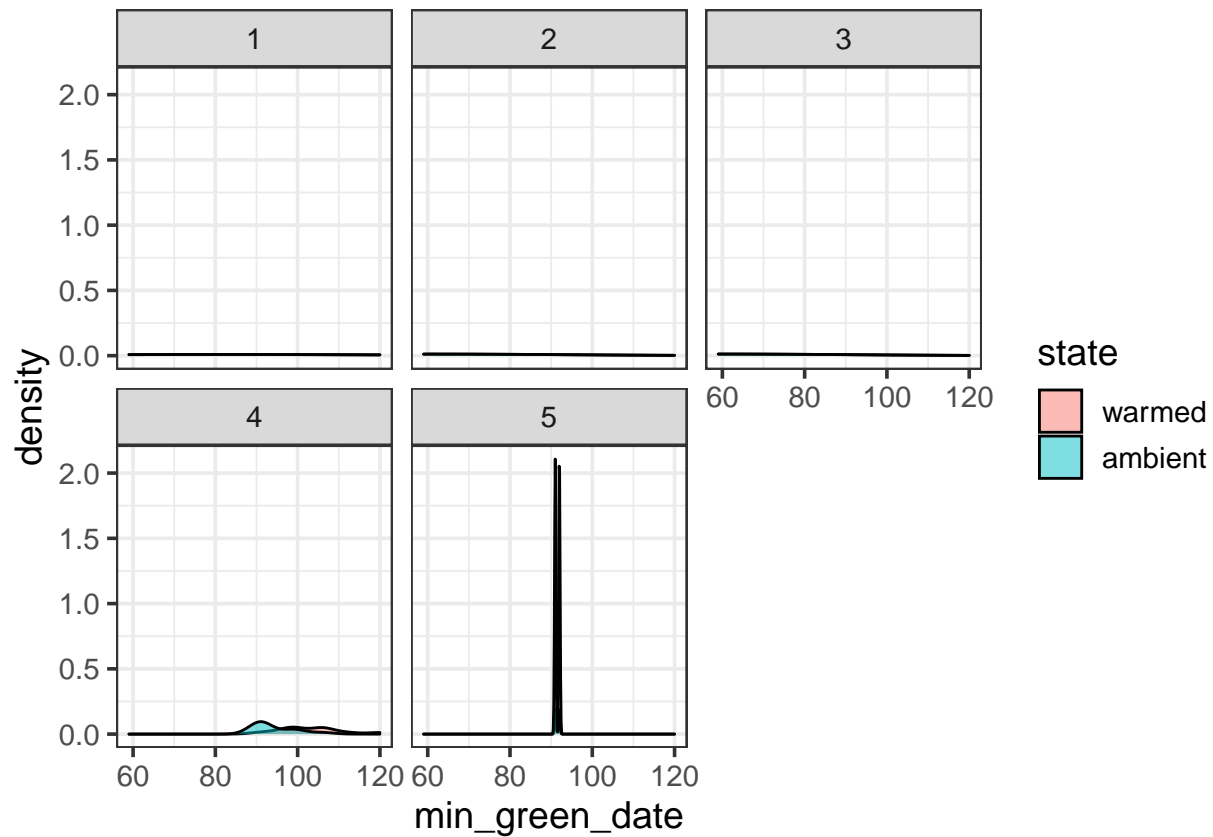


```
p3.2u <- ggplot(data = green_umbs, aes(x = spp_half_cover_date,
  fill = state)) + geom_density(alpha = 0.5)
p3.2u + facet_wrap(~year_factor) + labs(title = "Species-level half cover date")
```


Species-level half cover date



```
# this will just show sampling date artifact
p4 <- ggplot(data = green_kbsp, aes(x = min_green_date, fill = state)) +
  geom_density(alpha = 0.5)
p4 + facet_wrap(~year_factor)
```



```
# code below won't run: Or try with tidyverse format
# green_kbsp.t<-as_tibble(green_kbsp) green_kbsp.t %>%
# gather(state, plot_half_cover_date, year_factor) %>%
# ggplot(aes(plot_half_cover_date, fill = state)) +
# geom_histogram() + facet_wrap(~year_factor)

# looks like the 225 spike is from 2018 and 2020 - what's
# going on here is that you are treating all species-plot
# records as independent observations, so the influence of
# species differences is likely coming through here.
kbs_2018 <- subset(green_kbs, year == 4) # many records on 235
kbs_2020 <- subset(green_kbs, year == 6) # records from 227 & 228
```

Leverage plots and detecting Outliers. <https://www.statmethods.net/stats/riagnostics.html>

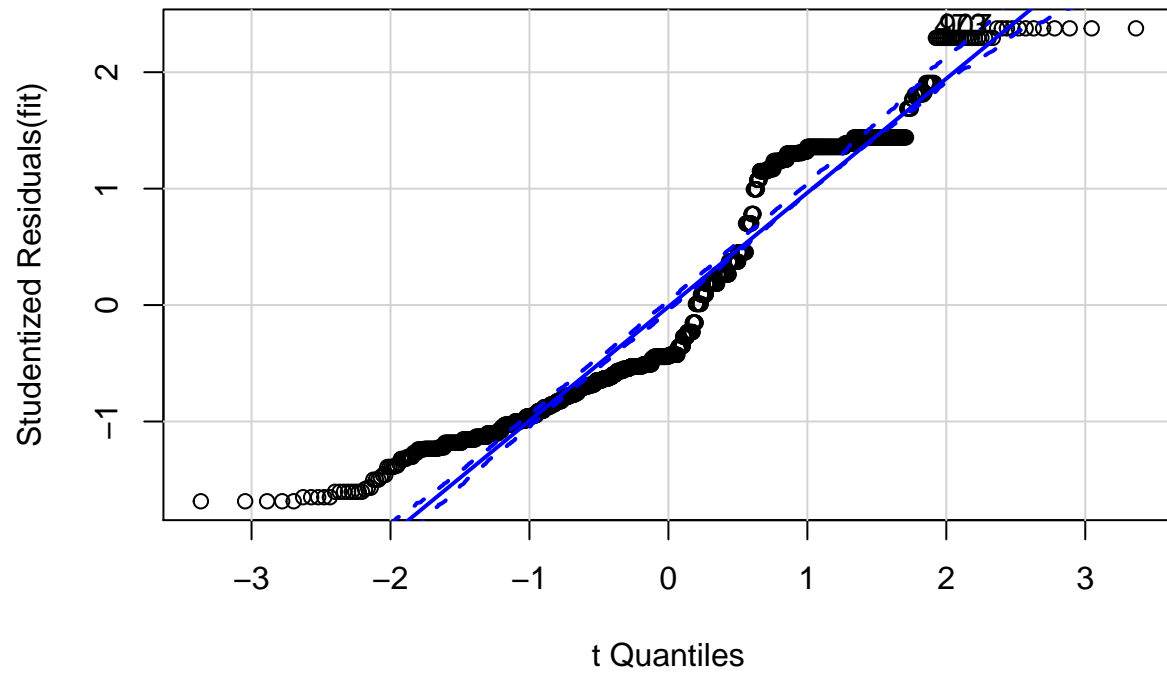
These illustrate whether certain data points have more leverage (more influence), and thus could be outliers. It's a way of detecting outliers. Leverage plots can help identify whether a point has high or low influence, based on its leverage and residual and determining model fit with and without the point in question. Ultimately you decide whether the points are outliers or not, based on the knowledge of the system and how much it changes the model when included vs. excluded from the data used to fit the model. Here is a good overview of the combination of leverage and residual: scroll down to sections beginning at "13.3 Unusual Observations": <https://davidalpiaz.github.io/appliedstats/model-diagnostics.html>

```
# checking fit for date as a function of state and species -  
# bringing in species here makes it obvious that that is  
# explaining some of the variation compared with the  
# state-only model you had previously.
```

```
# species-level data KBS State-only model  
fit <- lm(spp_half_cover_date ~ state, data = green_kbs)  
outlierTest(fit) # no outliers
```

```
## No Studentized residuals with Bonferroni p < 0.05  
## Largest |rstudent|:  
##      rstudent unadjusted p-value Bonferroni p  
## 473 2.376821      0.017611      NA  
qqPlot(fit, main = "QQ Plot")
```

QQ Plot

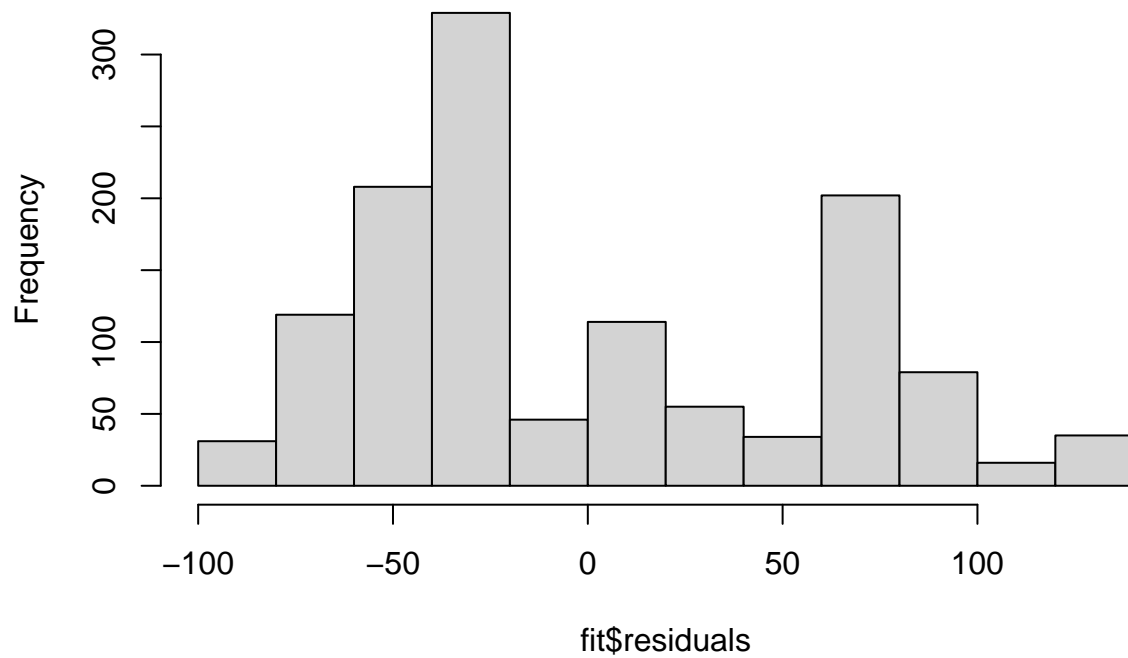


```
## 473 907
```

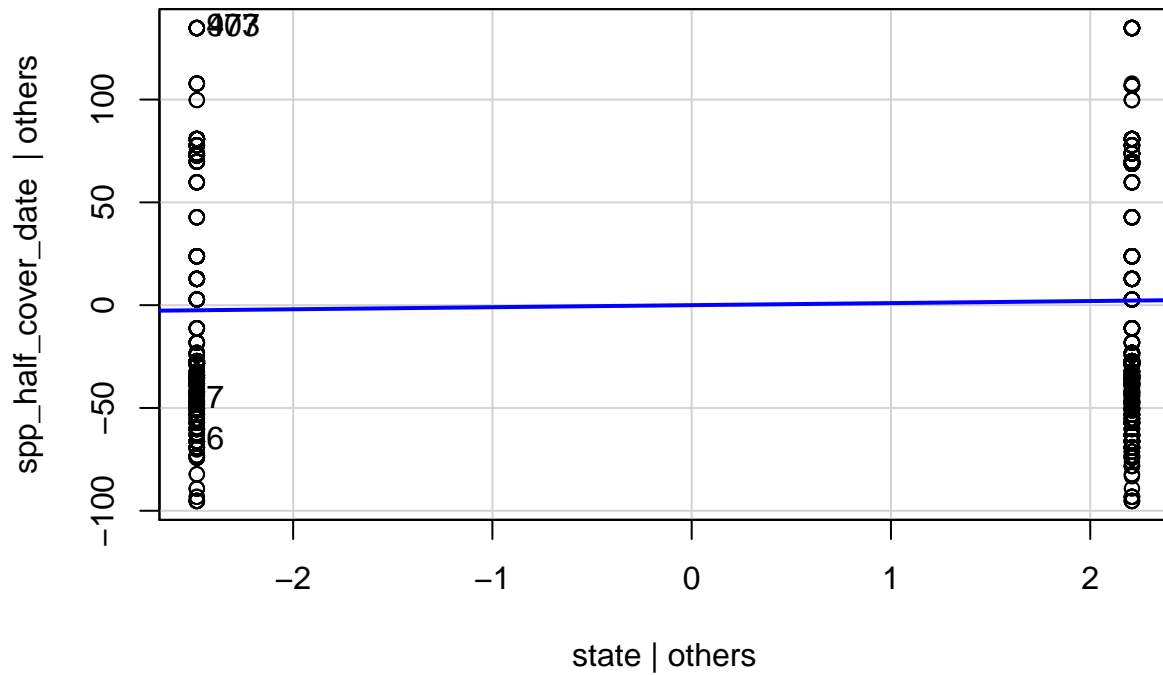
```
## 283 552
```

```
hist(fit$residuals)
```

Histogram of fit\$residuals



```
leveragePlots(fit)
```

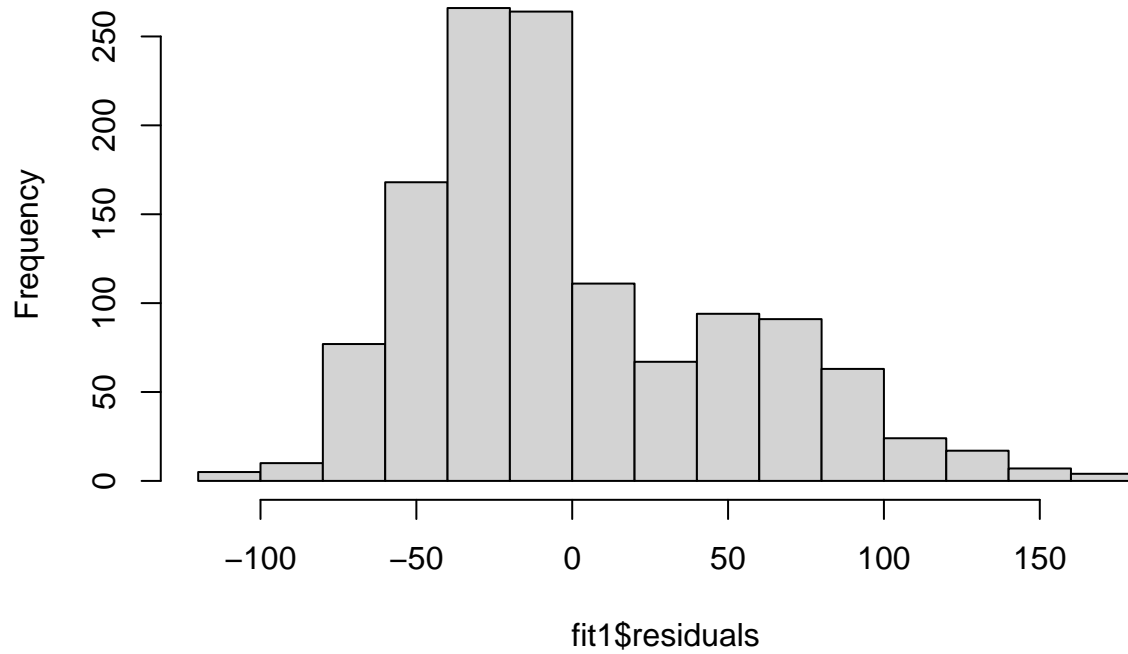


```
# KBS State and species model
fit1 <- lm(spp_half_cover_date ~ state + species, data = green_kbs)
outlierTest(fit1) # no outliers
```

```
## No Studentized residuals with Bonferroni p < 0.05
## Largest |rstudent|:
##      rstudent unadjusted p-value Bonferroni p
## 1910 3.455976      0.00056677      0.71866
```

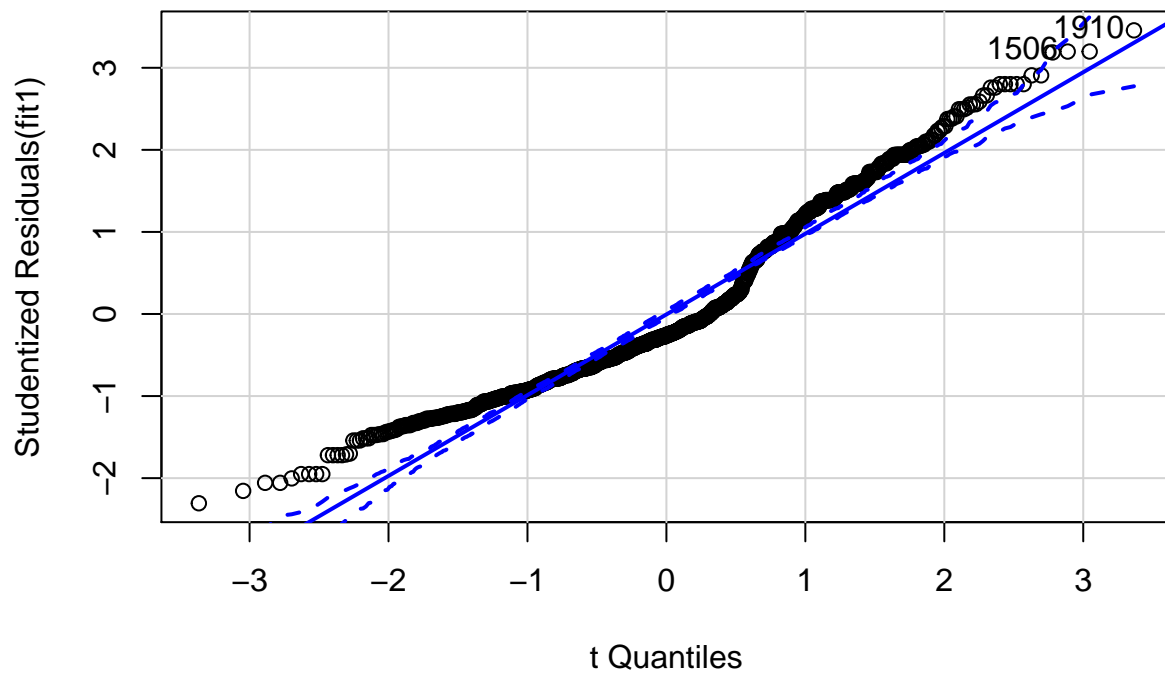
```
hist(fit1$residuals)
```

Histogram of fit1\$residuals



```
qqPlot(fit1, main = "QQ Plot")
```

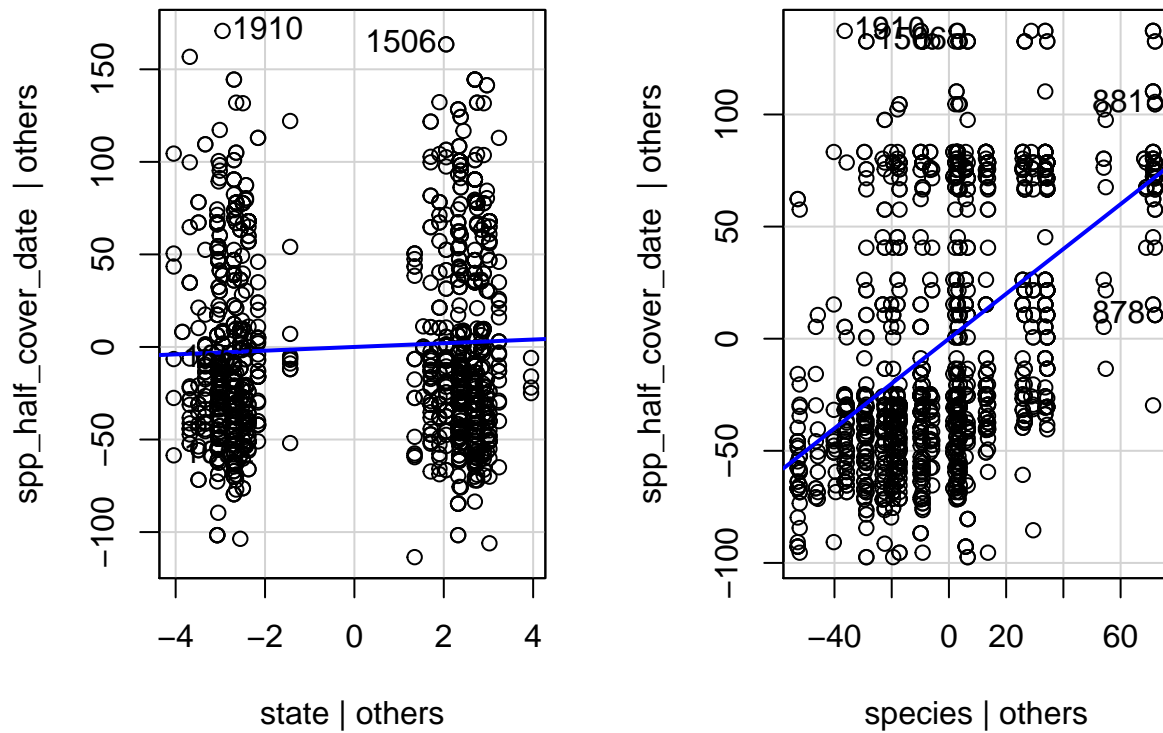
QQ Plot



```
## 1506 1910  
## 943 1152
```

```
leveragePlots(fit1)
```

Leverage Plots

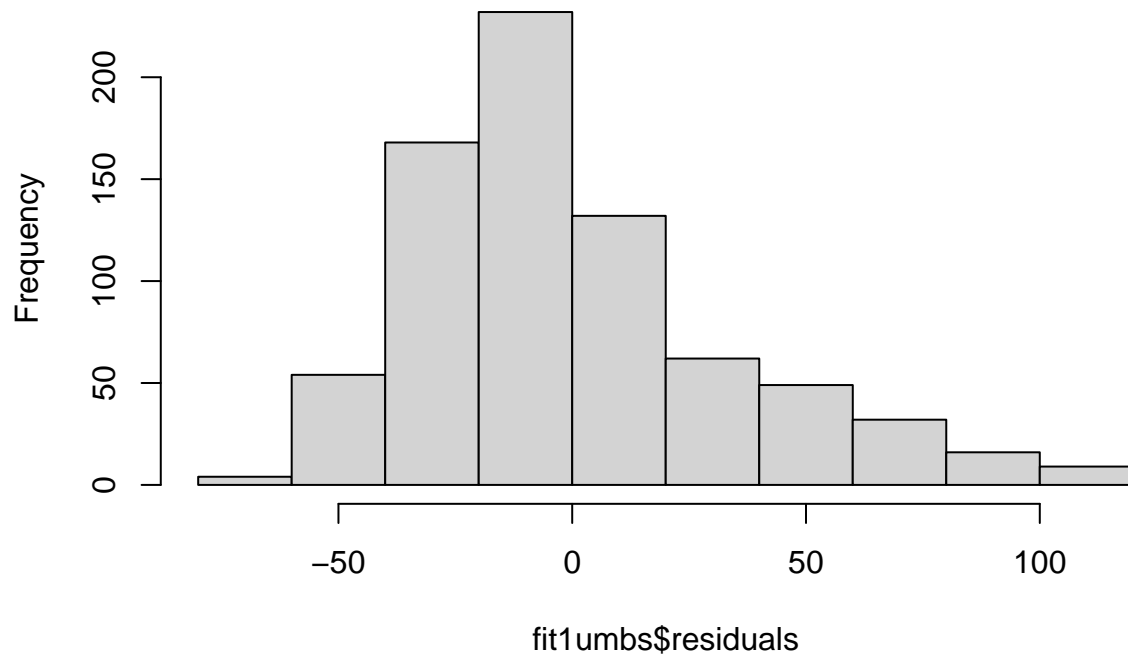


```
# UMBS State and species model
fitlumbs <- lm(spp_half_cover_date ~ state + species, data = green_umbs)
outlierTest(fitlumbs) # no outliers

## No Studentized residuals with Bonferroni p < 0.05
## Largest |rstudent|:
##      rstudent unadjusted p-value Bonferroni p
## 647 3.194361      0.0014607      NA

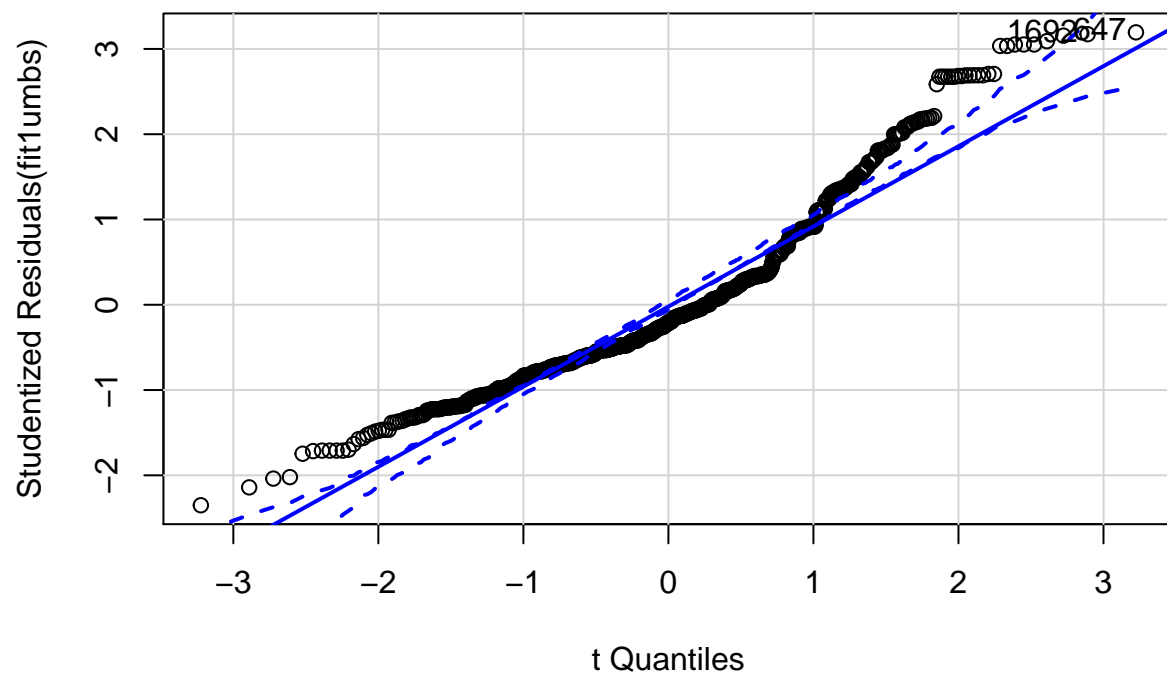
hist(fitlumbs$residuals)
```

Histogram of fit1umbs\$residuals



```
qqPlot(fit1umbs, main = "QQ Plot")
```

QQ Plot

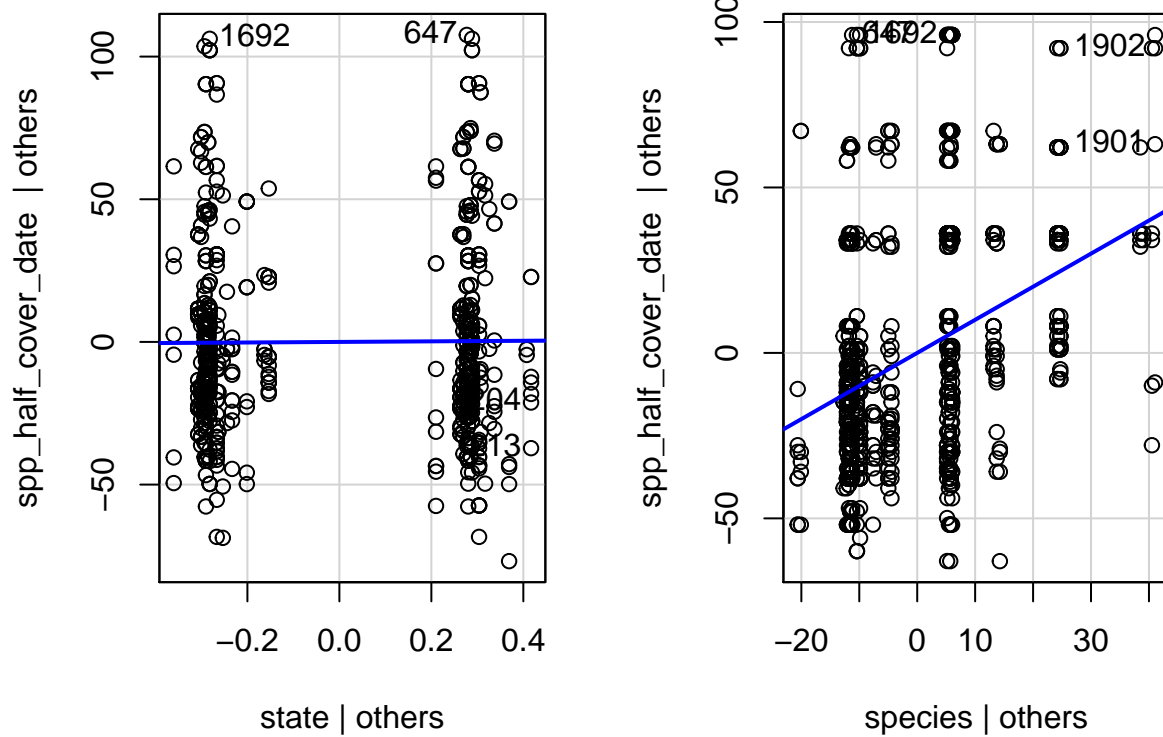


```
## 647 1692
## 267 666
```



```
leveragePlots(fit1umbs)
```

Leverage Plots



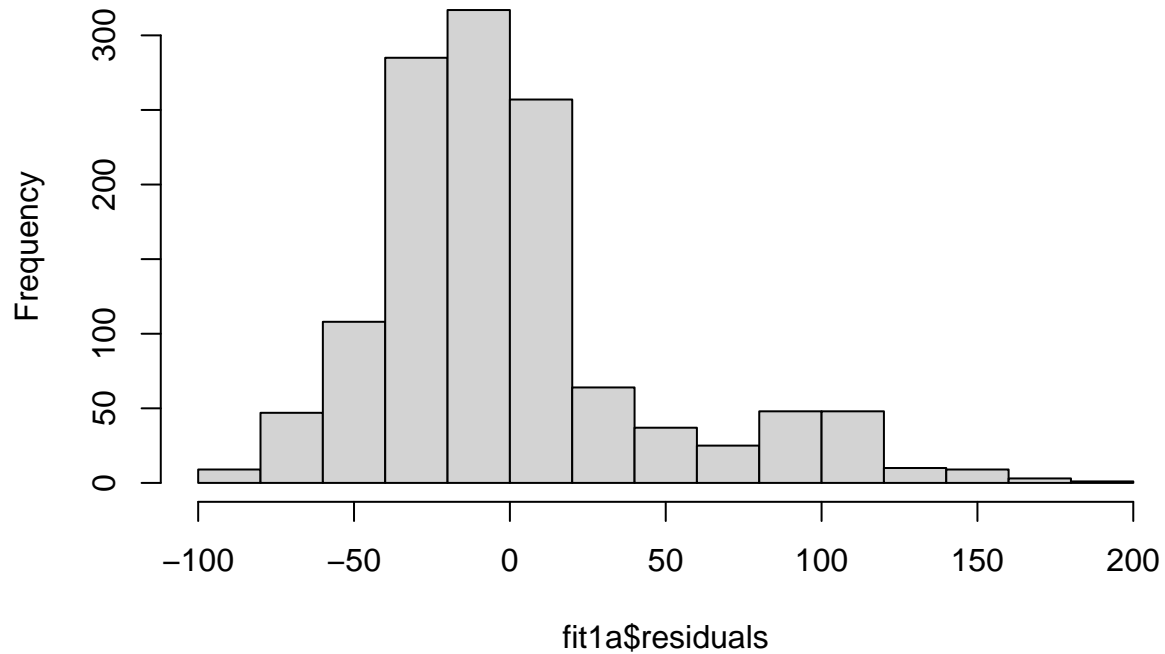
```
fit1a <- lm(min_green_date ~ state + species, data = green_kbs)
```

```
outlierTest(fit1a) # record 473 is an outlier, could probably ignore since we aren't using min date as
```

```
##      rstudent unadjusted p-value Bonferroni p
## 473 4.211351      2.7212e-05      0.034505
```

```
hist(fit1a$residuals)
```

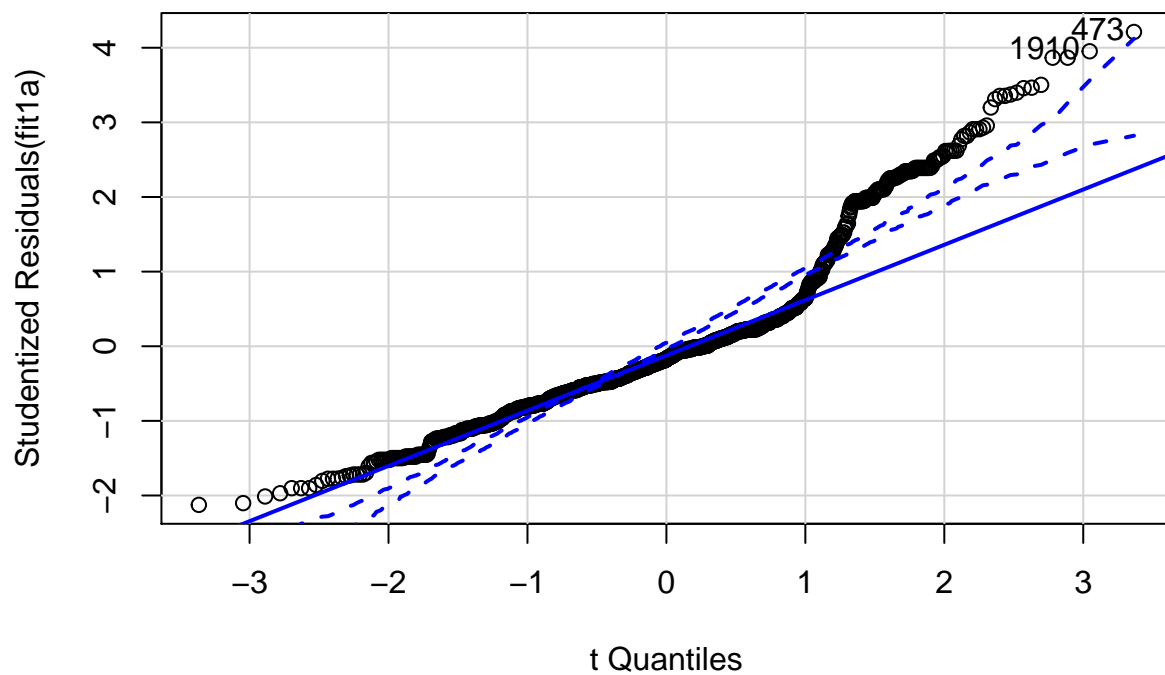
Histogram of fit1a\$residuals



```
qqPlot(fit1a, main = "QQ Plot")
```

```
## Warning in rlm.default(x, y, weights, method = method, wt.method = wt.method, :  
## 'rlm' failed to converge in 20 steps
```

QQ Plot

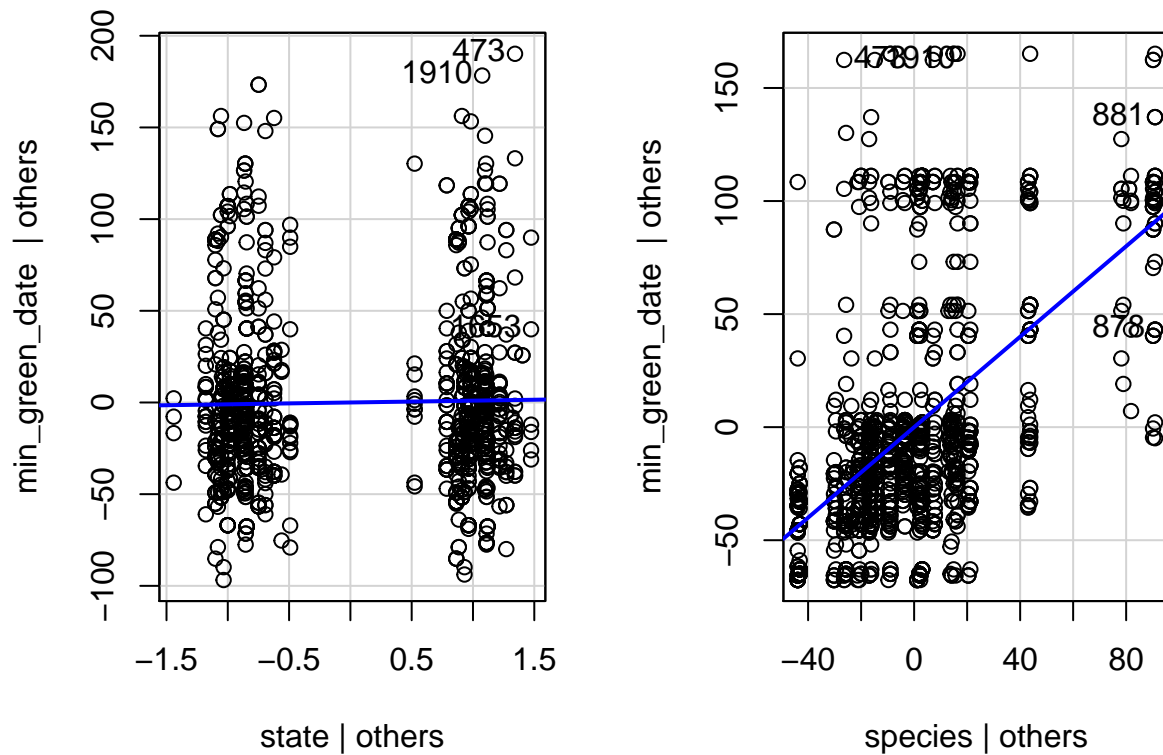


```
## 473 1910
```

```
## 283 1152
```

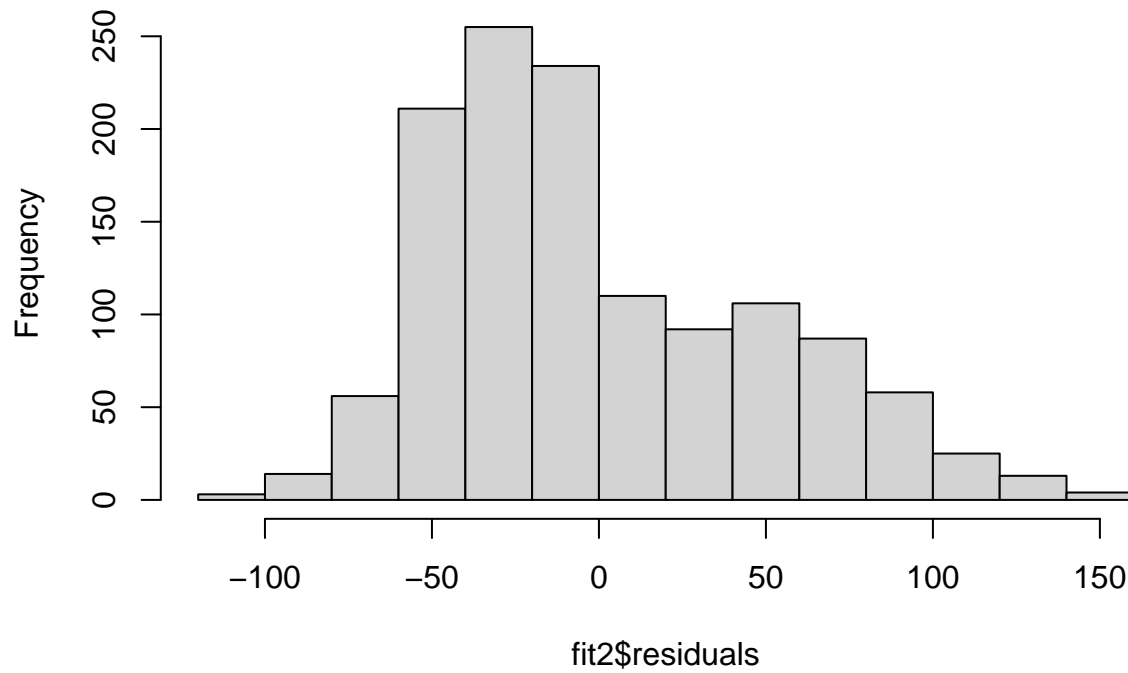
```
leveragePlots(fit1a)
```

Leverage Plots



```
# checking fit for date as a function of state and year  
fit2 <- lm(spp_half_cover_date ~ state + species + year, data = green_kbs)  
hist(fit2$residuals)
```

Histogram of fit2\$residuals

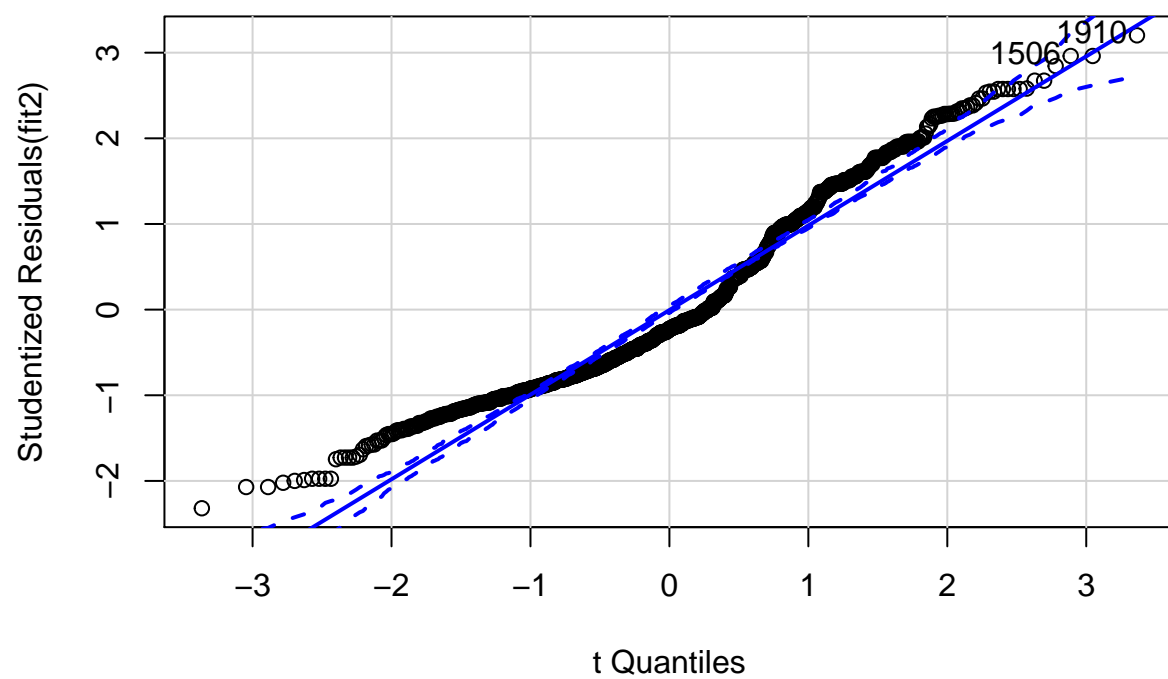


```
outlierTest(fit2) # no outliers
```

```
## No Studentized residuals with Bonferroni p < 0.05  
## Largest |rstudent|:  
##      rstudent unadjusted p-value Bonferroni p  
## 1910 3.202366      0.0013976      NA
```

```
qqPlot(fit2, main = "QQ Plot")
```

QQ Plot

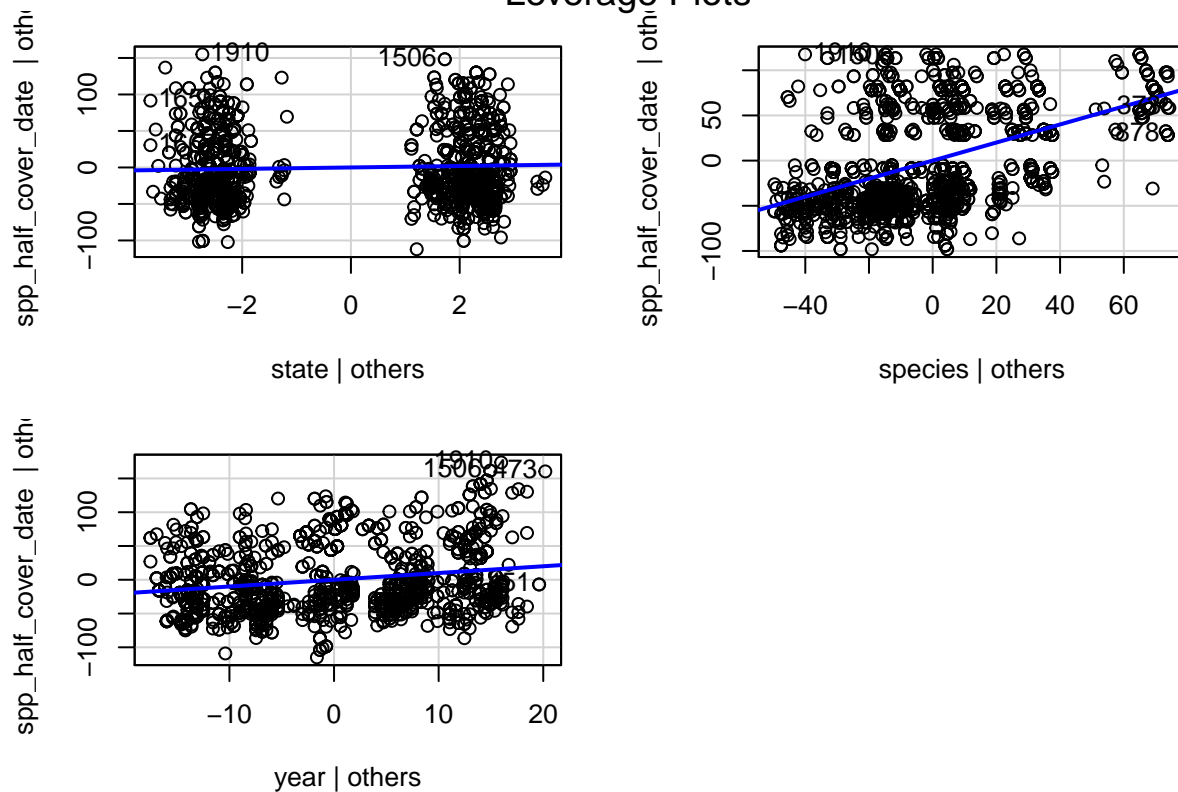


```
## 1506 1910
```

```
## 943 1152
```

```
leveragePlots(fit2)
```

Leverage Plots

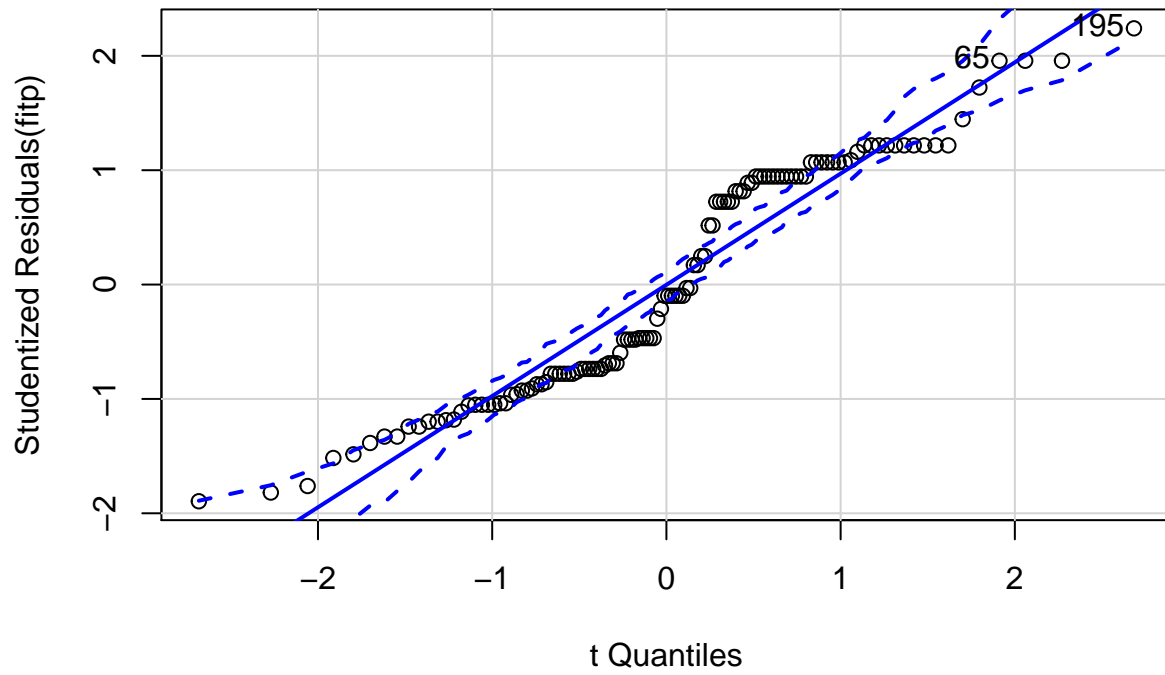


```
# plot level data KBS State-only model
fitp <- lm(plot_half_cover_date ~ state, data = green_kbsp)
outlierTest(fitp) # no outliers

## No Studentized residuals with Bonferroni p < 0.05
## Largest |rstudent|:
##      rstudent unadjusted p-value Bonferroni p
## 195 2.240402      0.026953      NA

qqPlot(fitp, main = "QQ Plot")
```

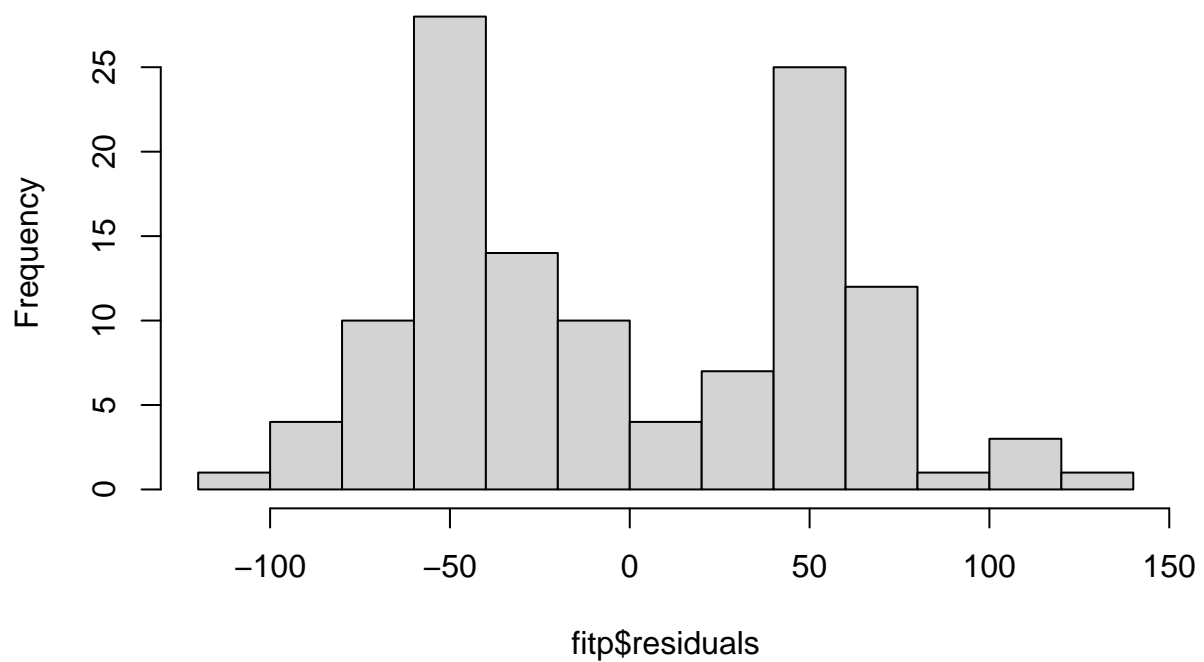
QQ Plot



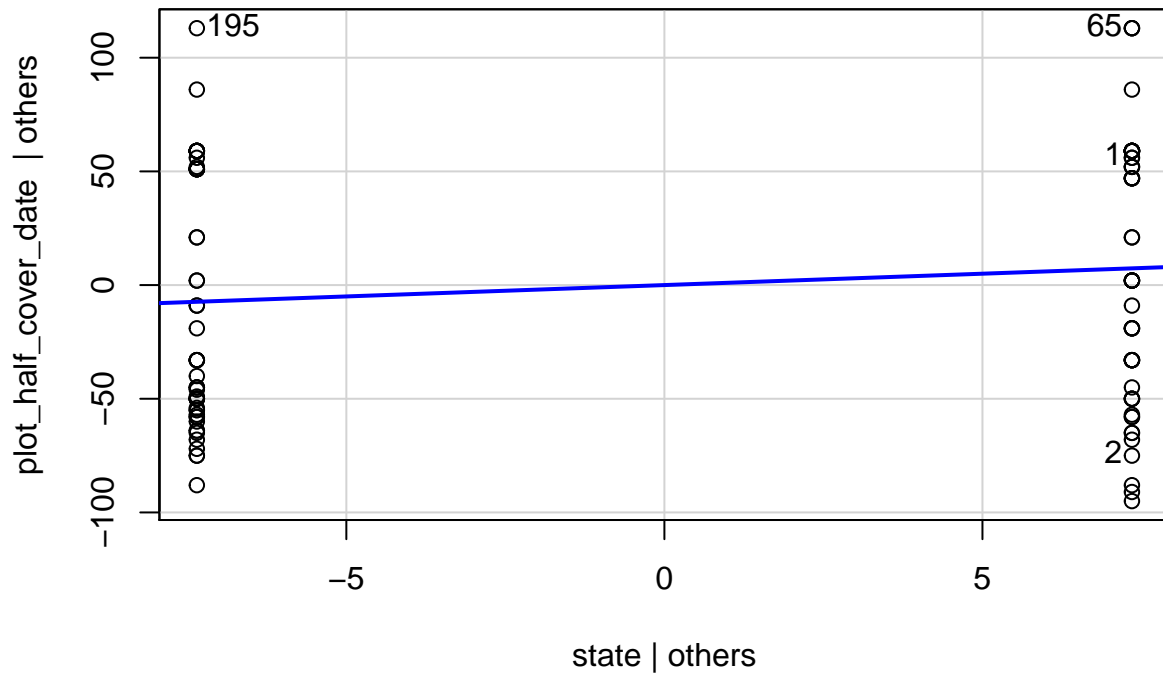
```
## 65 195  
## 35 100
```

```
hist(fitp$residuals)
```

Histogram of fitp\$residuals



```
leveragePlots(fitp)
```

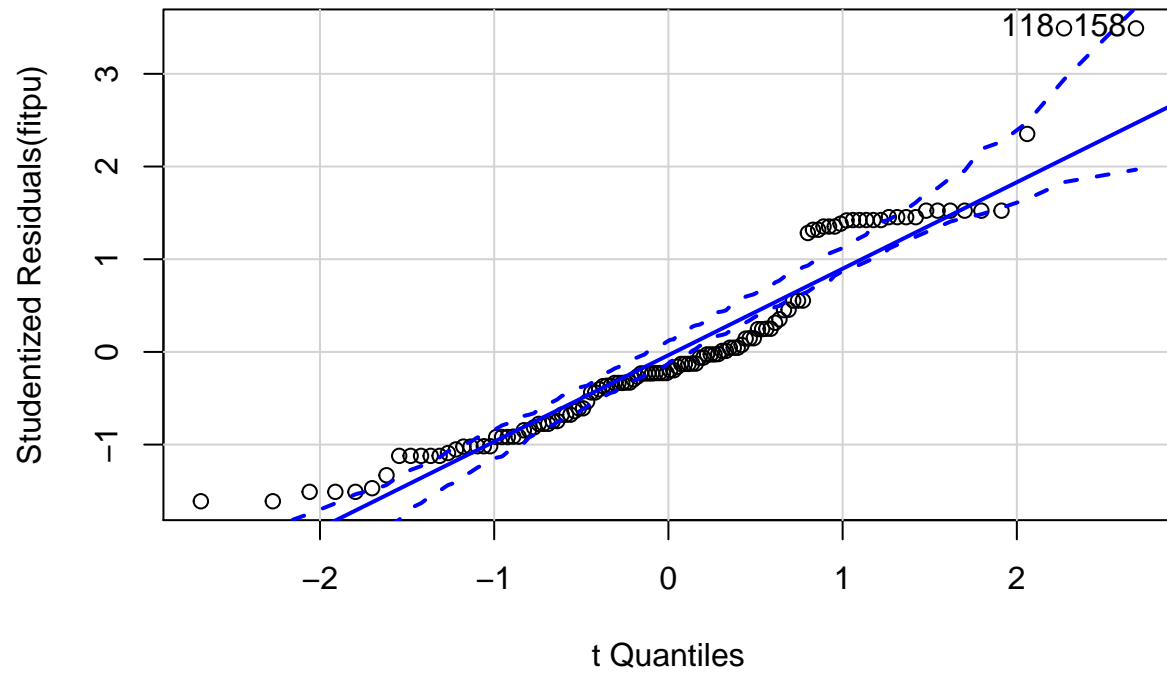


```
# UMBS State-only model
fitpu <- lm(plot_half_cover_date ~ state, data = green_umbsp)
outlierTest(fitpu) # no outliers
```

```
## No Studentized residuals with Bonferroni p < 0.05
## Largest |rstudent|:
##      rstudent unadjusted p-value Bonferroni p
## 118 3.490806      0.00068005      0.081605
```

```
qqPlot(fitpu, main = "QQ Plot")
```

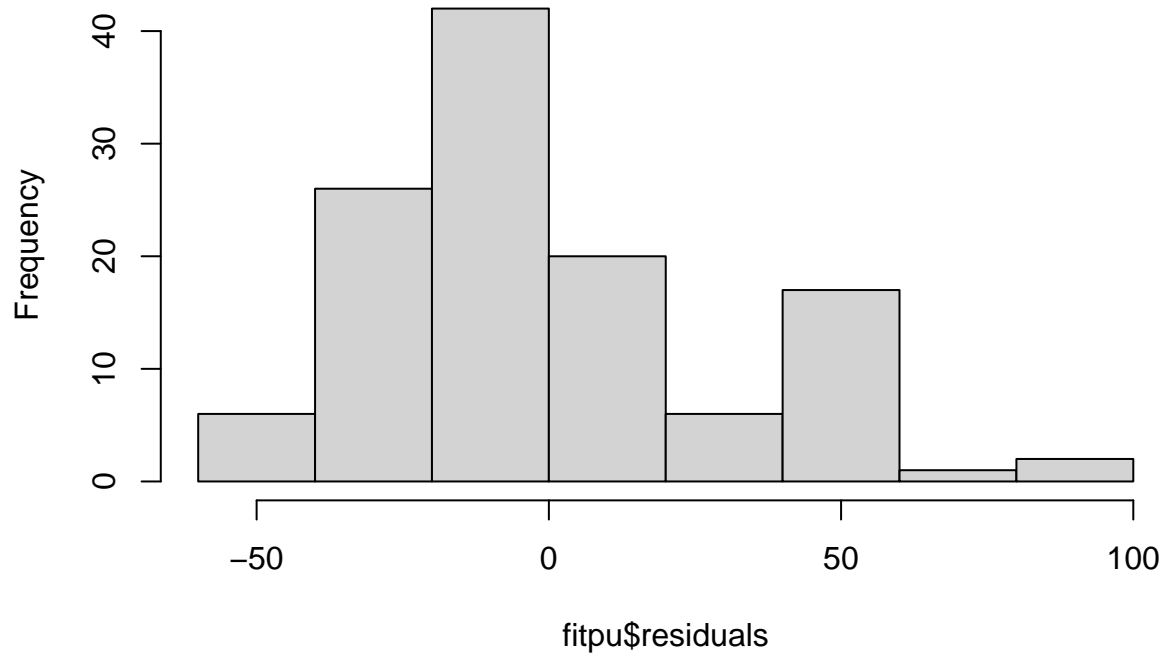

QQ Plot



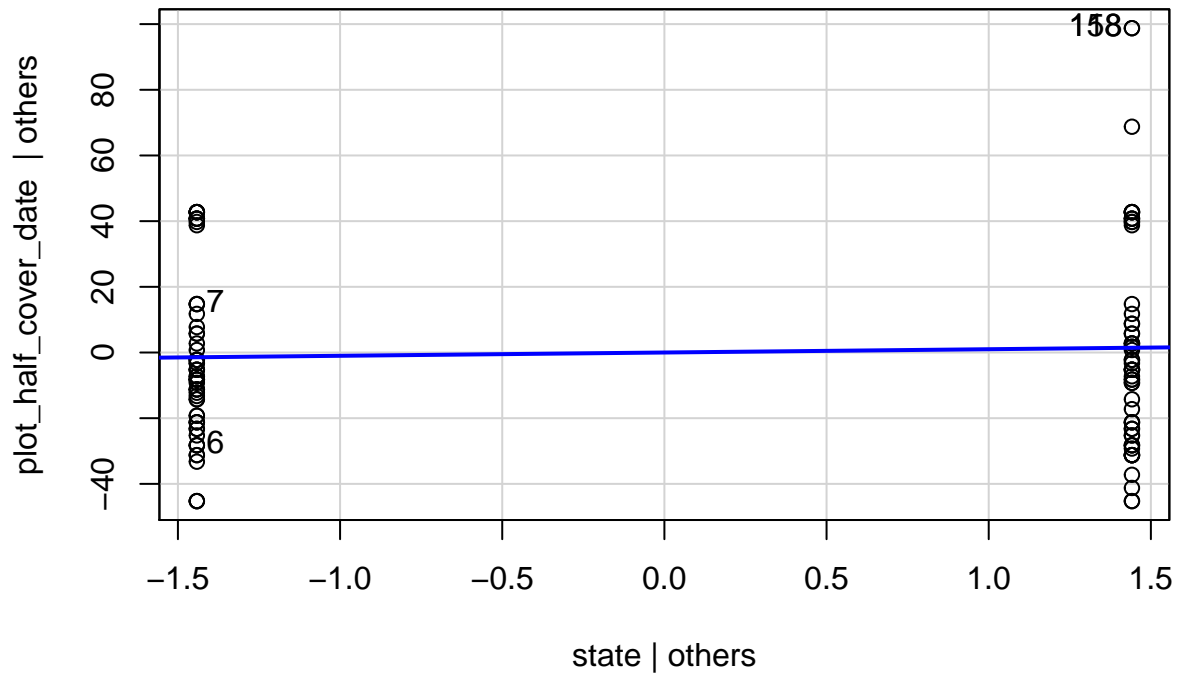
```
## 118 158  
## 58 78
```

```
hist(fitpu$residuals)
```

Histogram of fitpu\$residuals



```
leveragePlots(fitpu)
```

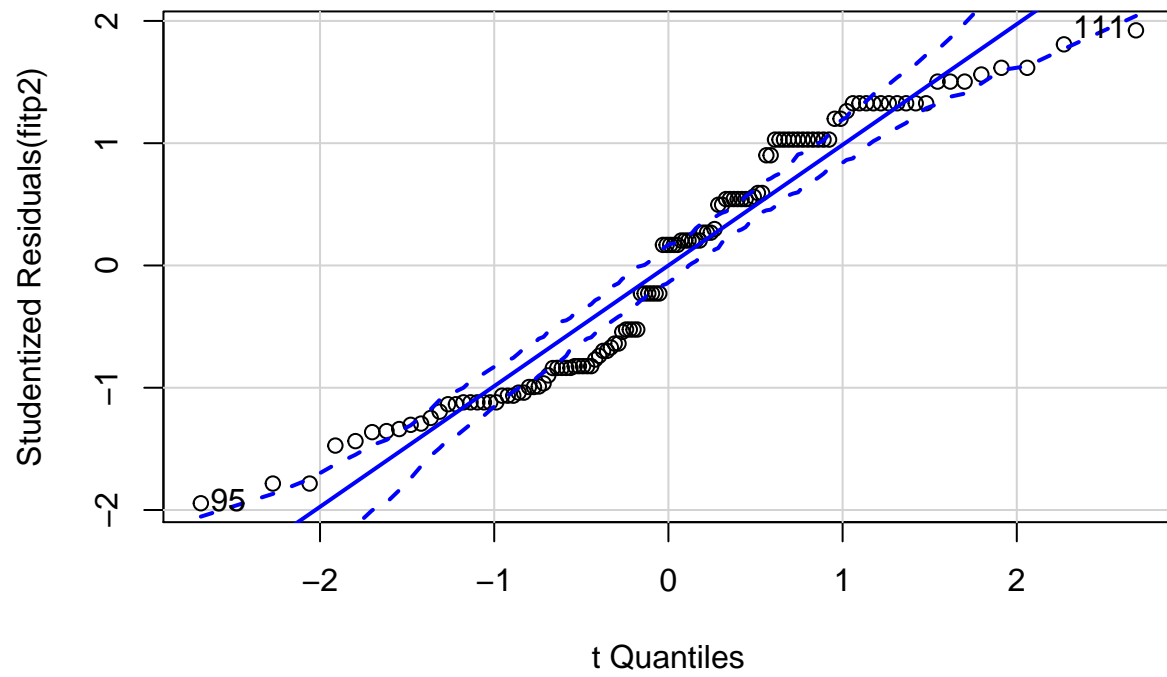


```
# KBS State and year model
fitp2 <- lm(plot_half_cover_date ~ state + year, data = green_kbsp)
outlierTest(fitp2) # no outliers
```

```
## No Studentized residuals with Bonferroni p < 0.05
## Largest |rstudent|:
##      rstudent unadjusted p-value Bonferroni p
## 95 -1.944735      0.054227      NA
```

```
qqPlot(fitp2, main = "QQ Plot")
```

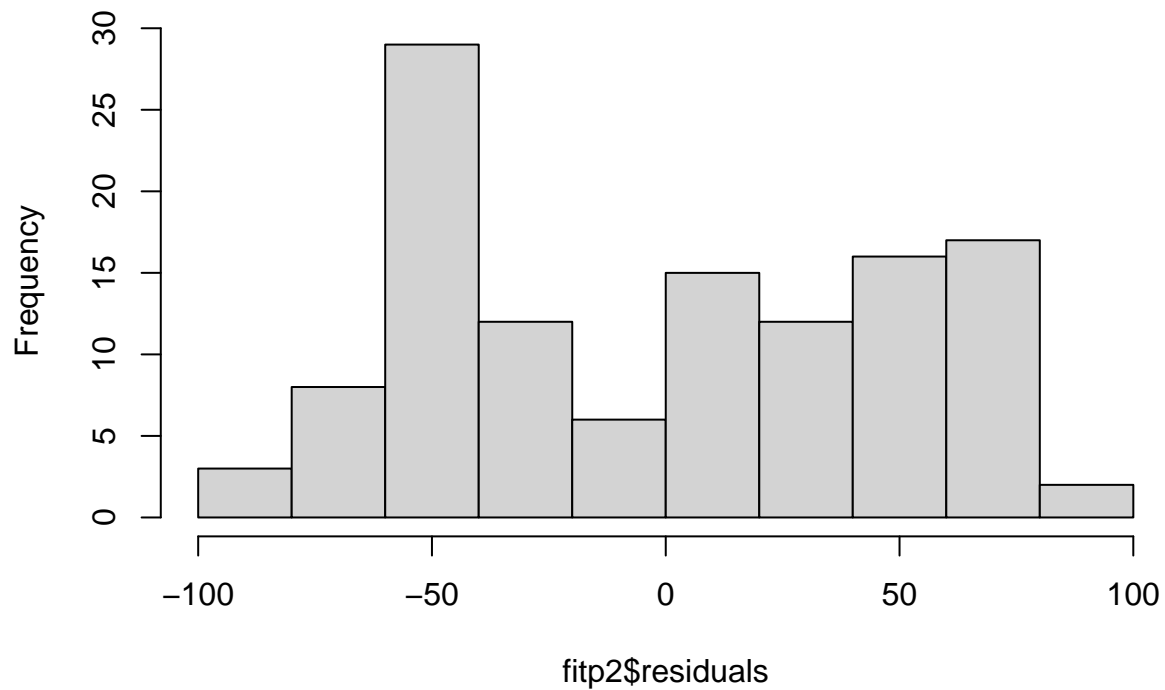
QQ Plot



```
## 95 111  
## 50 56
```

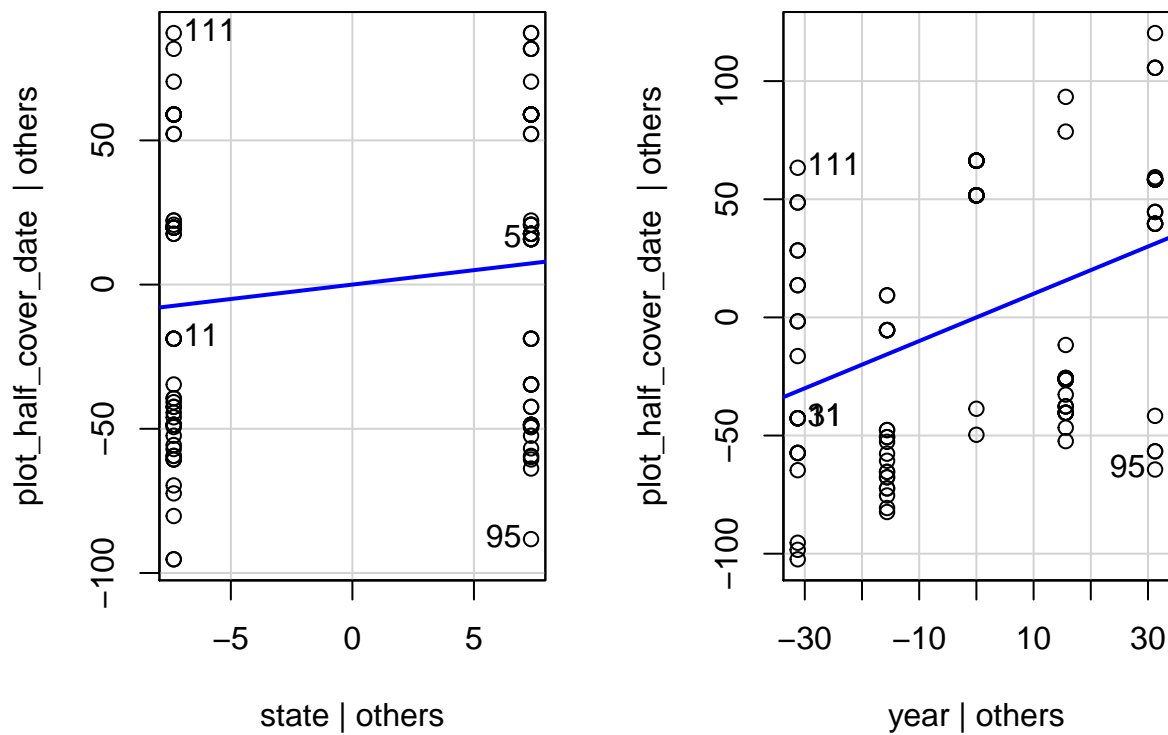
```
hist(fitp2$residuals)
```

Histogram of fitp2\$residuals



```
leveragePlots(fitp2)
```

Leverage Plots



```
# UMBS State and year model
```

```
fitp2u <- lm(plot_half_cover_date ~ state + year, data = green_umbsp)
```

```
outlierTest(fitp2u) # 118 and 158 are outliers - remove? look like the only plots with a half cover da
```

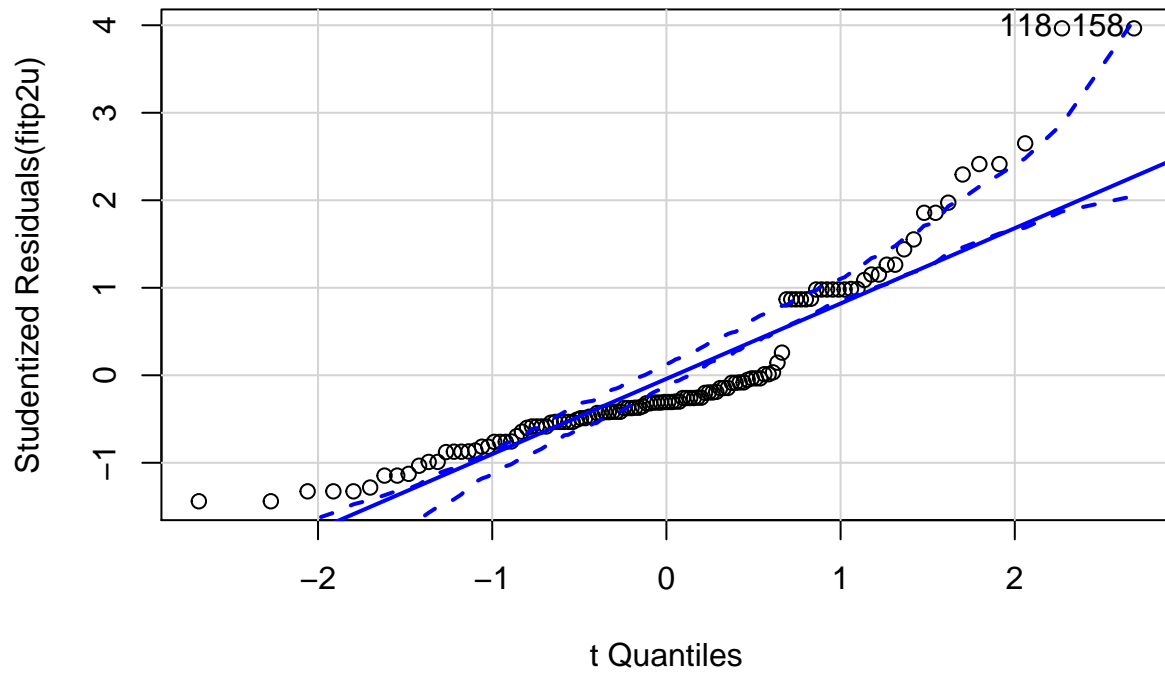
```
##      rstudent unadjusted p-value Bonferroni p
```

```
## 118 3.964625      0.00012753      0.015303
```

```
## 158 3.964625      0.00012753      0.015303
```

```
qqPlot(fitp2u, main = "QQ Plot")
```

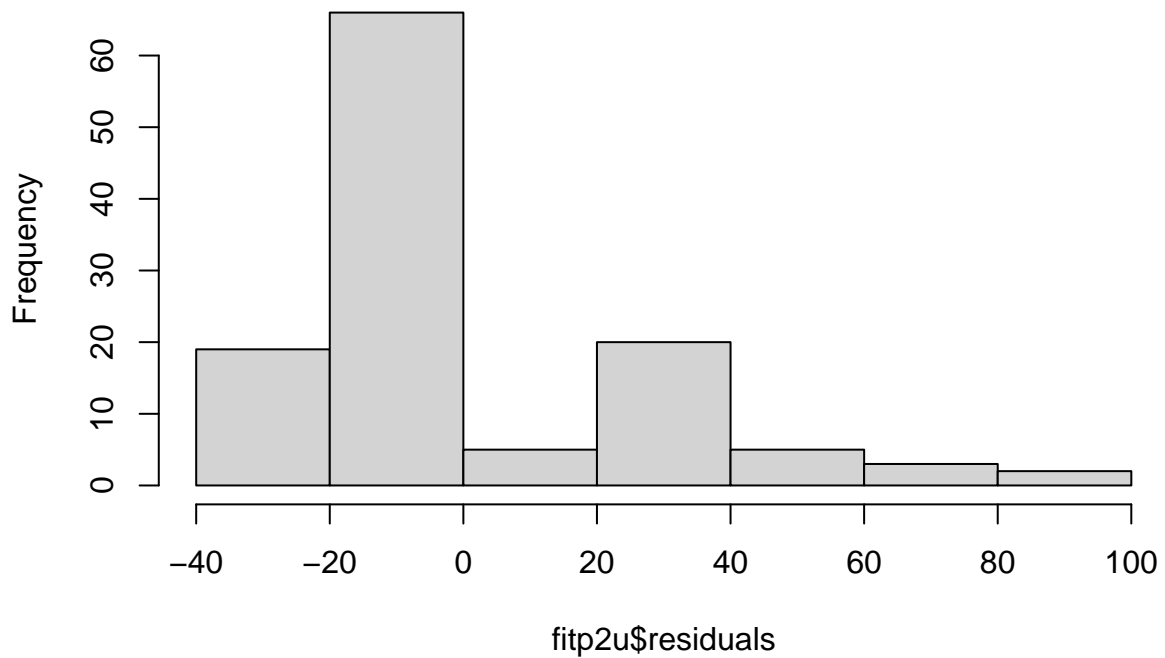
QQ Plot



```
## 118 158  
## 58 78
```

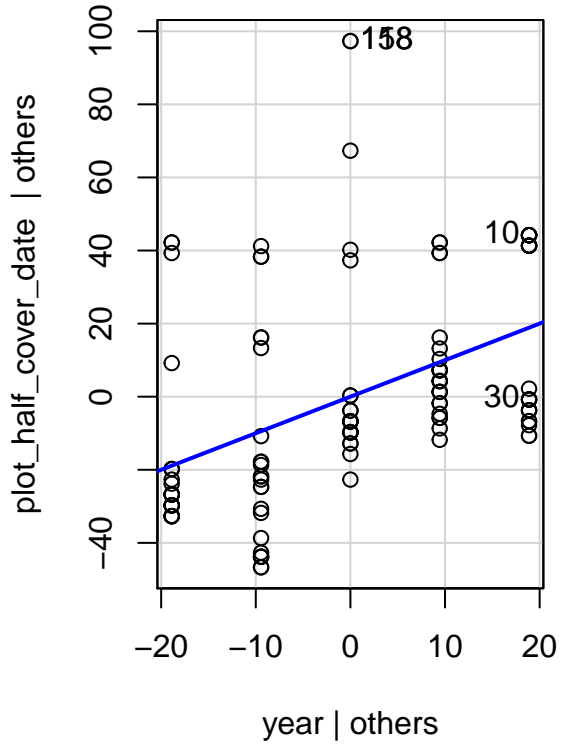
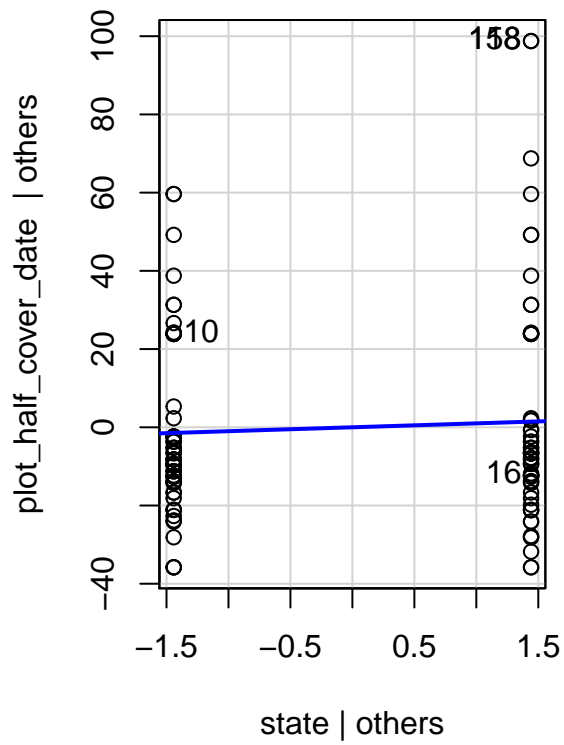
```
hist(fitp2u$residuals)
```

Histogram of fitp2u\$residuals



```
leveragePlots(fitp2u)
```

Leverage Plots



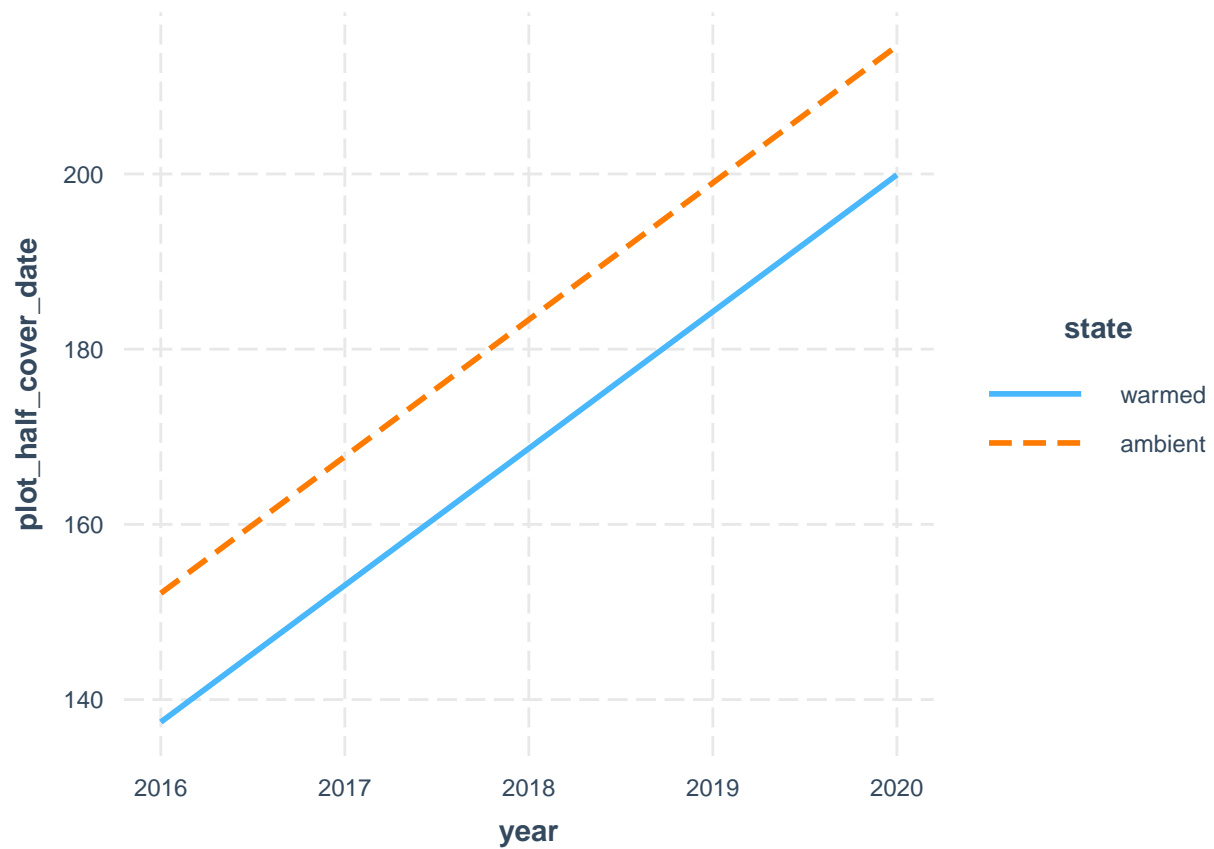
Normal distribution after accounting for species - we will be using species as a random effect to account for their variation. Set up some linear mixed effects models to evaluate. From Ben Bolker: “The traditional view of random effects is as a way to do correct statistical tests when some observations are correlated. ... Random effects are especially useful when we have (1) lots of levels (e.g., many species or blocks), (2) relatively little data on each level (although we need multiple samples from most of the levels), and (3) uneven sampling across levels. People sometimes say that random effects are “factors that you aren’t interested in.” This is not always true. While it is often the case in ecological experiments (where variation among sites is usually just a nuisance), it is sometimes of great interest.” In our case, variation among plots is a nuisance, and not something we’re interested in. For some questions, variation among species is also a nuisance for us. It’s possible that variation among years is a nuisance if we only care about warm vs. ambient, but I think time is an interesting variable to consider with this study.

We should also think about how we’re treating year. Some of the models have a state * year interaction as a fixed effect, which means that the warming or ambient treatment could affect the half_cover_date differently over time (there would be a different slope for each state in the relationship between half_cover_date (y) and year (x)). If we just had state + year, the states would have the same slope, indicating that they have no interaction in their effect on half_cover_date (but they could still have different intercepts).

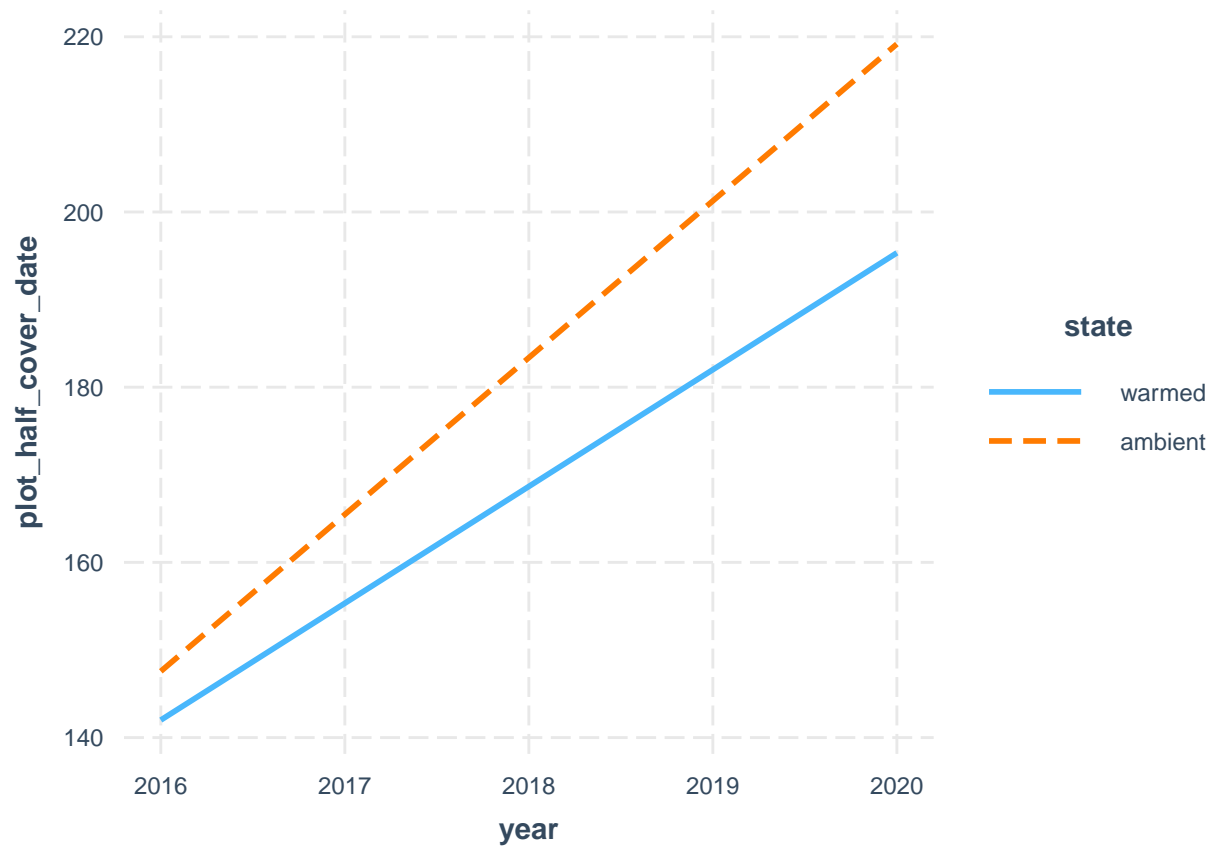
```
# Interaction plot (ignore for now the repeated measures with
# species); see:
# https://cran.r-project.org/web/packages/interactions/vignettes/interactions.html
# and: https://interactions.jacob-long.com/

# KBS
fit3 <- lm(plot_half_cover_date ~ state + year, data = green_kbsp)
interact_plot(fit3, pred = year, modx = state)
```

```
## Warning: year and state are not included in an interaction with one another in the
## model.
```

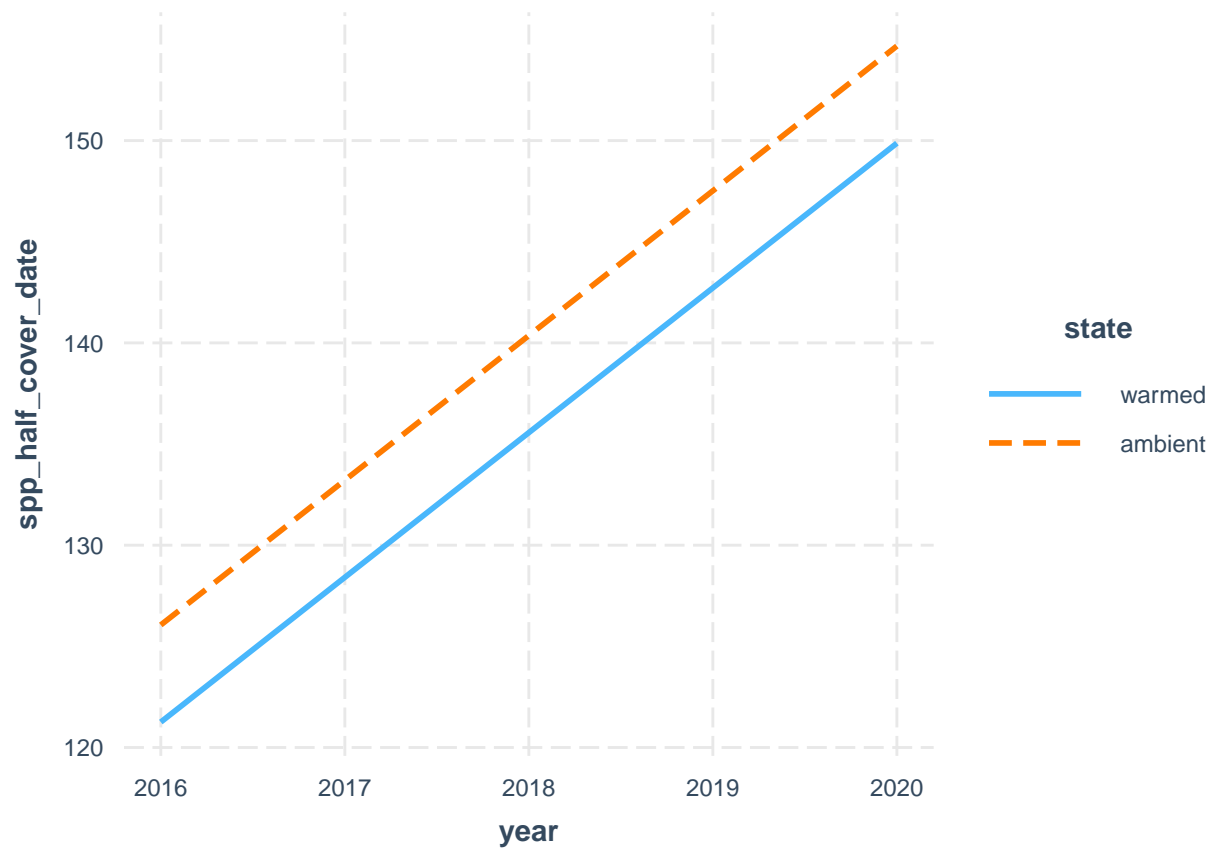


```
fit4 <- lm(plot_half_cover_date ~ state * year, data = green_kbsp)
interact_plot(fit4, pred = year, modx = state)
```

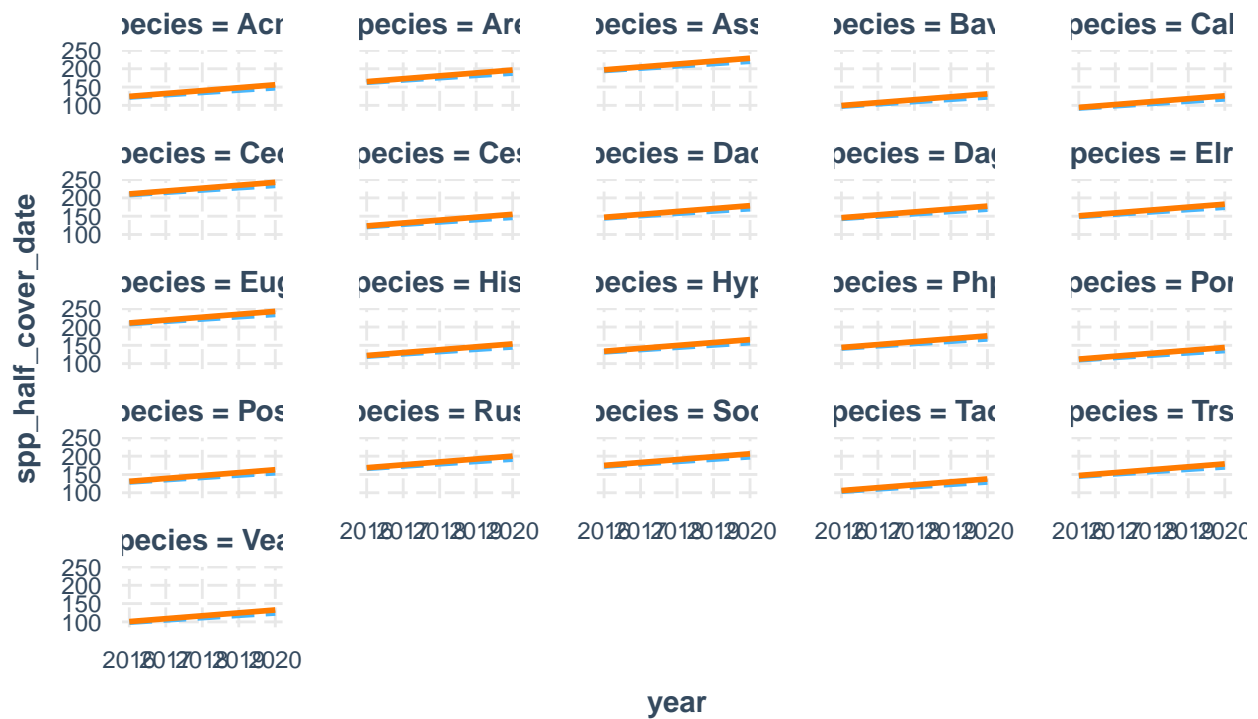
```
fit5 <- lm(spp_half_cover_date ~ state + year + species, data = green_kbs)
interact_plot(fit5, pred = year, modx = state)
```

```
## Warning: year and state are not included in an interaction with one another in the
## model.
```



```
fit6 <- lm(spp_half_cover_date ~ state * year + species, data = green_kbs)
interact_plot(fit6, pred = year, modx = state, mod2 = species)
```

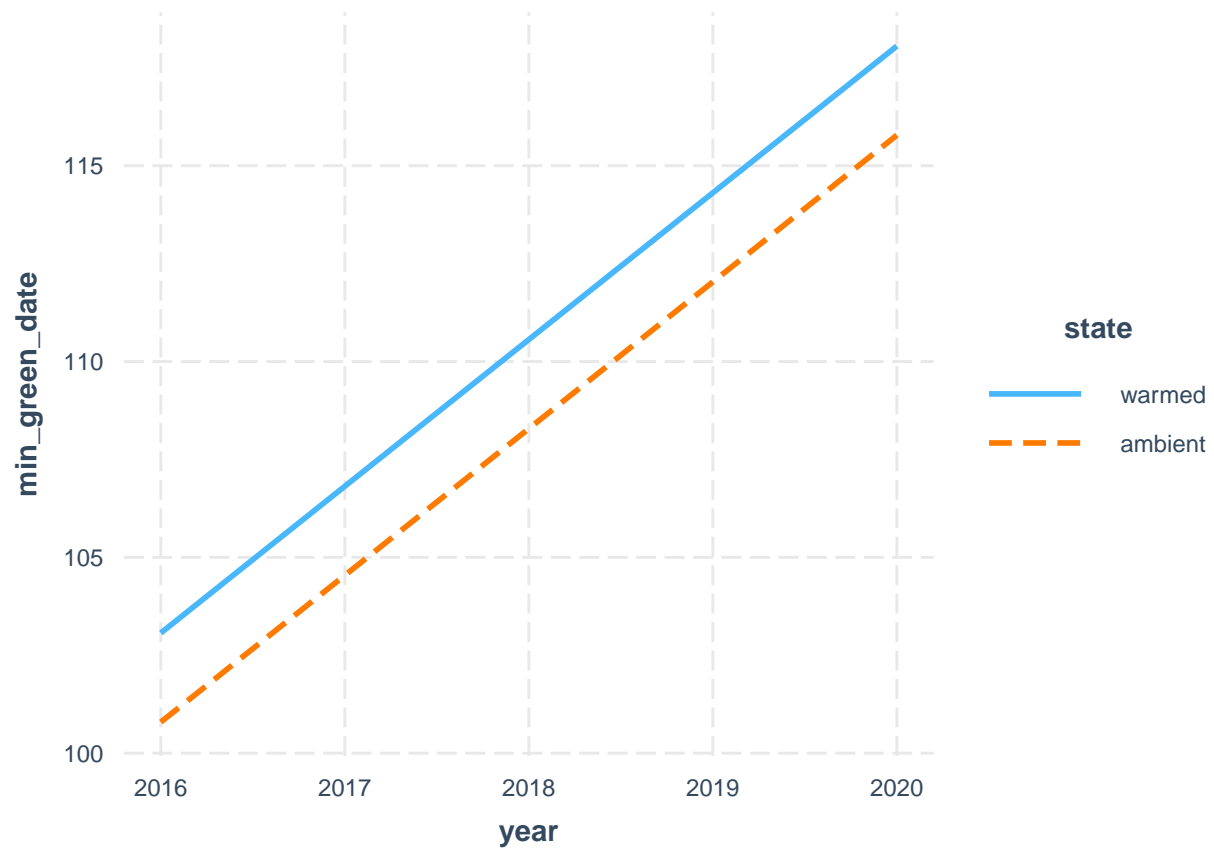
```
## Warning: year and state and species are not included in an interaction with one
## another in the model.
```



state — — — warmed ——— ambient

```
fit7 <- lm(min_green_date ~ state + year + species, data = green_kbs)
interact_plot(fit7, pred = year, modx = state)
```

```
## Warning: year and state are not included in an interaction with one another in the
## model.
```



```
fit8 <- lm(min_green_date ~ state * year + species, data = green_kbs)
interact_plot(fit8, pred = year, modx = state, mod2 = species)
```

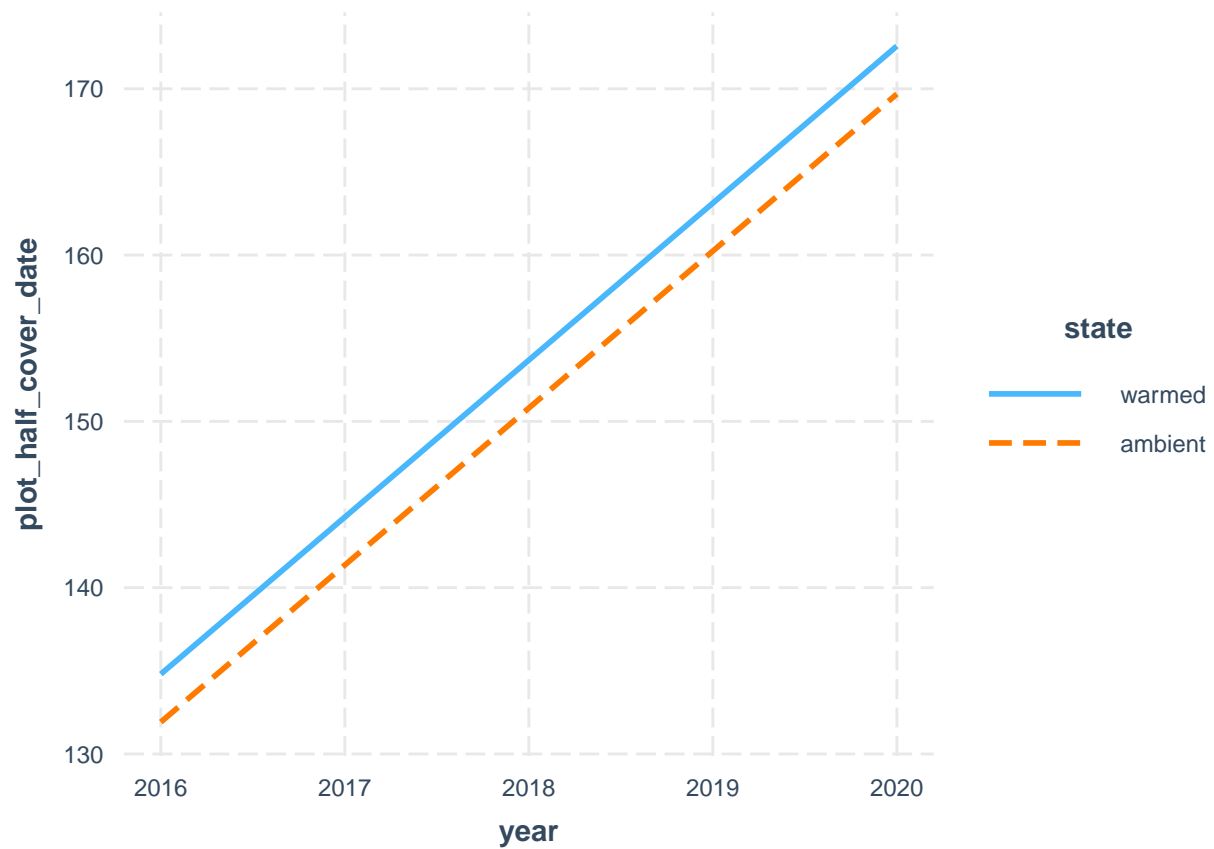
```
## Warning: year and state and species are not included in an interaction with one
## another in the model.
```



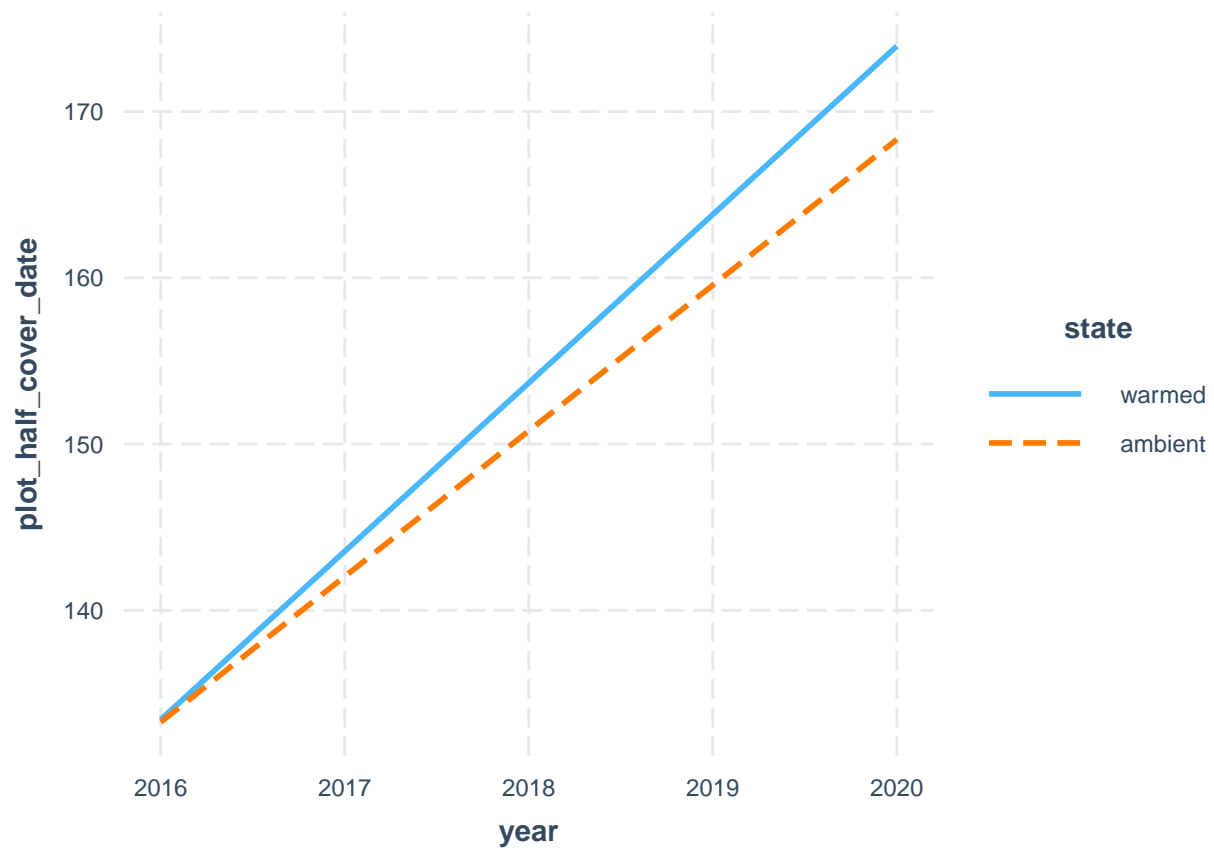
```
# UMBS
```

```
fit3u <- lm(plot_half_cover_date ~ state + year, data = green_umbsp)
interact_plot(fit3u, pred = year, modx = state)
```

```
## Warning: year and state are not included in an interaction with one another in the
## model.
```

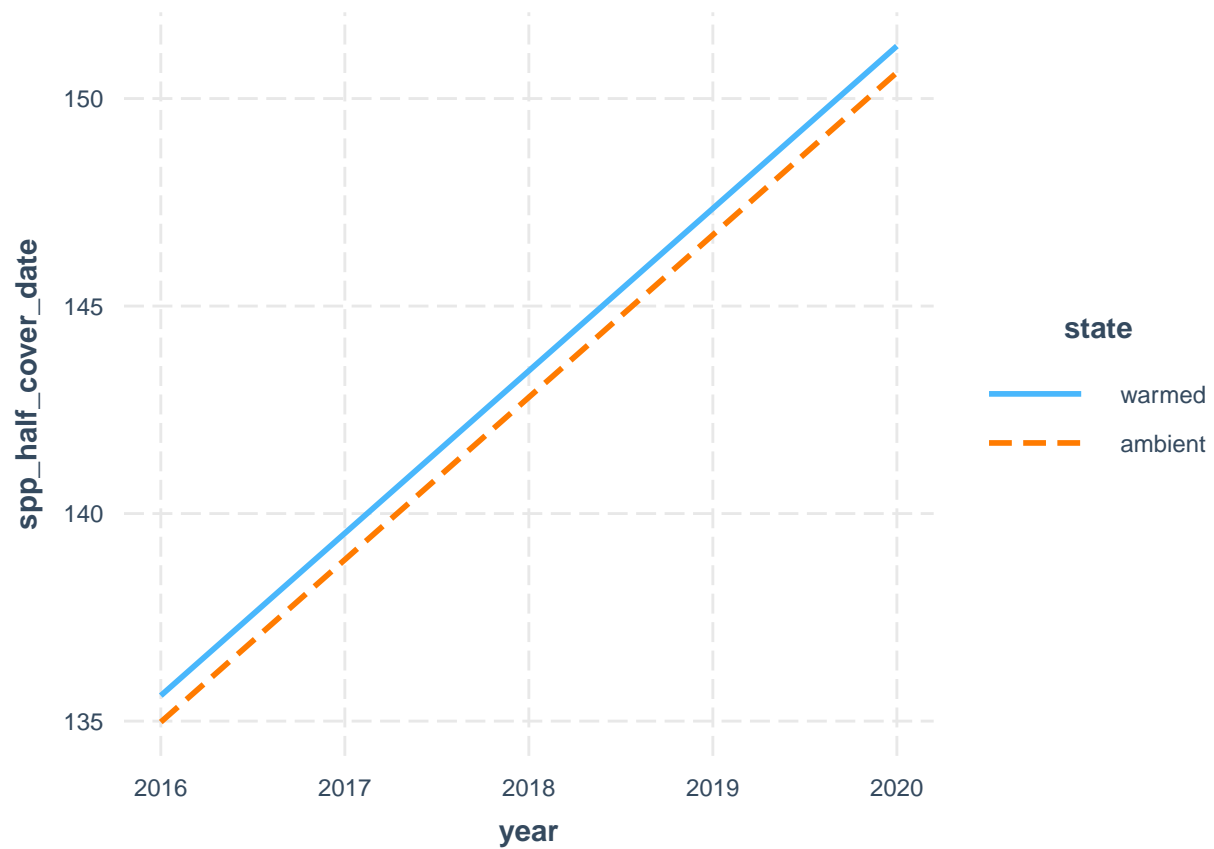


```
fit4u <- lm(plot_half_cover_date ~ state * year, data = green_umbsp)
interact_plot(fit4u, pred = year, modx = state)
```



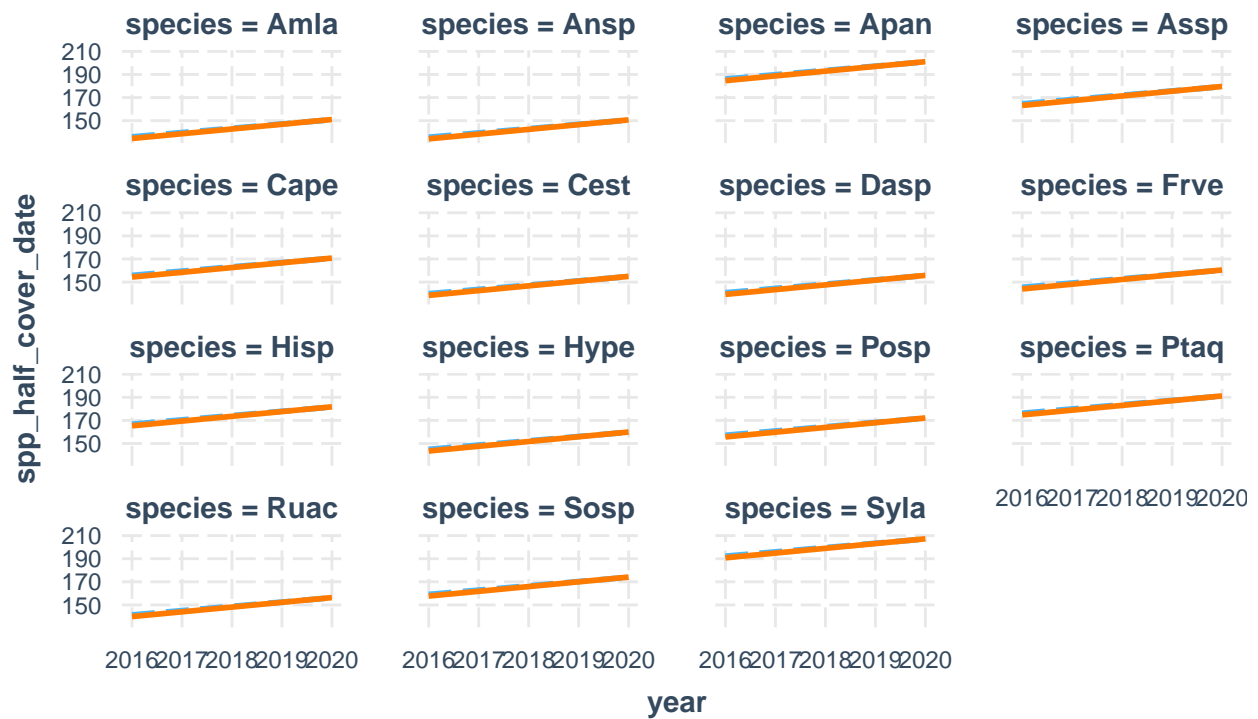
```
fit5u <- lm(spp_half_cover_date ~ state + year + species, data = green_umbs)
interact_plot(fit5u, pred = year, modx = state)
```

```
## Warning: year and state are not included in an interaction with one another in the
## model.
```



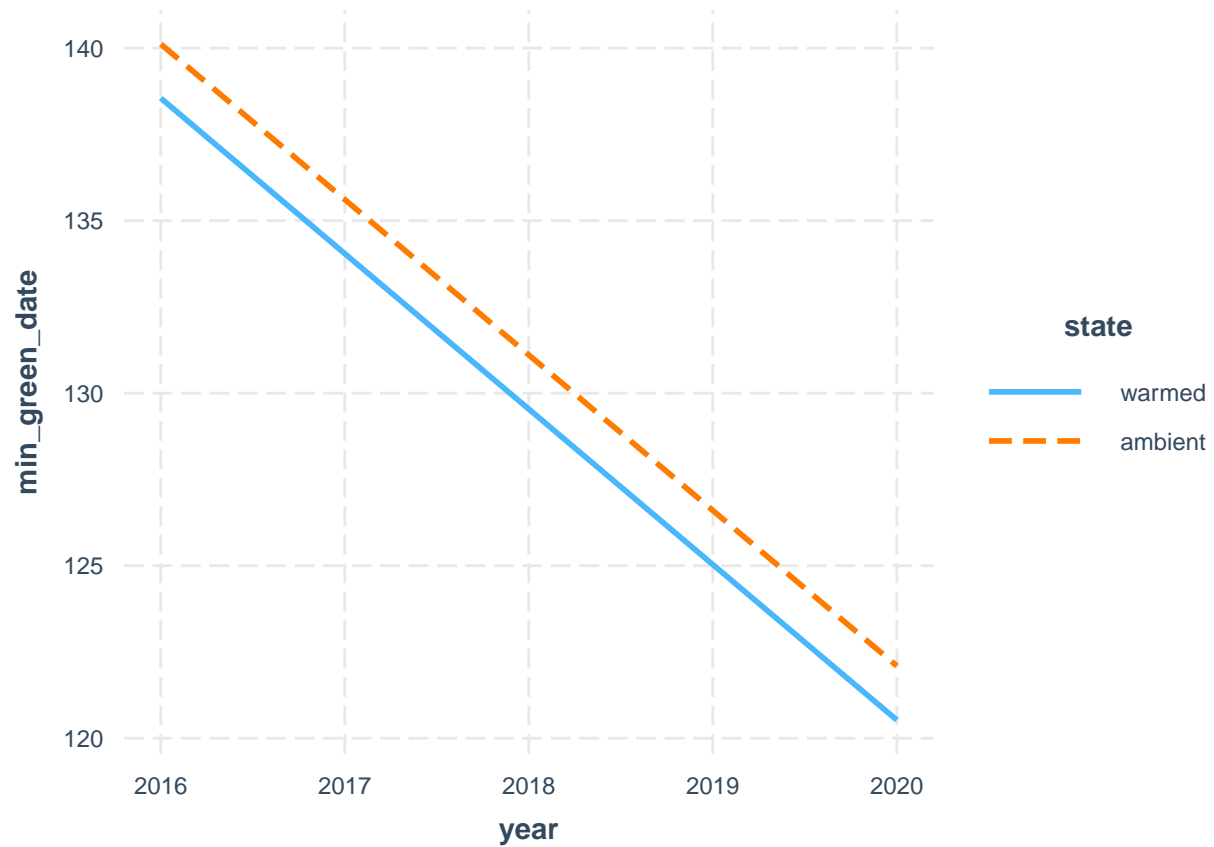
```
fit6u <- lm(spp_half_cover_date ~ state * year + species, data = green_umbs)
interact_plot(fit6u, pred = year, modx = state, mod2 = species)
```

```
## Warning: year and state and species are not included in an interaction with one
## another in the model.
```

```
fit7u <- lm(min_green_date ~ state + year + species, data = green_umbs)
interact_plot(fit7u, pred = year, modx = state)
```

```
## Warning: year and state are not included in an interaction with one another in the
## model.
```



```
fit8u <- lm(min_green_date ~ state * year + species, data = green_umbs)
interact_plot(fit8u, pred = year, modx = state, mod2 = species)
```

```
## Warning: year and state and species are not included in an interaction with one
## another in the model.
```



KBS Species-level Mixed Effects Models:

```
# Start by replicating (almost) what we did in the Decologia
# 2018 paper. The only difference here is that we have
# multiple years, so we are also including year as a fixed
# effect and as an interactive term. Our goal here is to find
# a model that is the best fit to the data. We also want to
# find a model that is the most parsimonious (one that has
# the fewest parameters).
```

```
# Do we need to include plot as a random effect with the KBS
# models?
```

```
mod1 <- lmer(spp_half_cover_date ~ state * year + insecticide *
  year_factor + (1 | species) + (1 | plot), green_kbs, REML = FALSE)
```

```
## fixed-effect model matrix is rank deficient so dropping 1 column / coefficient
```

```
## Warning: Some predictor variables are on very different scales: consider
## rescaling
```

```
## Warning: Some predictor variables are on very different scales: consider
## rescaling
```

```
mod2 <- lmer(spp_half_cover_date ~ state * year + insecticide *
  year_factor + (1 | species), green_kbs, REML = FALSE)
```

```
## fixed-effect model matrix is rank deficient so dropping 1 column / coefficient
```

```
## Warning: Some predictor variables are on very different scales: consider
## rescaling
```

```
## Warning: Some predictor variables are on very different scales: consider
## rescaling
```

```
# Run analysis of variance on each model (see this for more
# explanation on how anova on a linear mixed effects model is
# similar to an anova on a regular linear model:
# https://m-clark.github.io/docs/mixedModels/anovamixed.html)
# anova(mod1) anova(mod2)
```

```
# Run an ANOVA to test if 2 models to test whether the more
# complex model is significantly better at capturing the data
# than the simpler model. If the resulting p-value is
# sufficiently low (usually less than 0.05), we conclude that
# the more complex model is significantly better than the
# simpler model, and thus favor the more complex model. If
# the p-value is not sufficiently low (usually greater than
# 0.05), we should favor the simpler model.
```

```
# https://bookdown.org/ndphillips/YaRrr/comparing-regression-models-with-anova.html
```

```
anova(mod2, mod1) # They are different so plot as a random effect should stay in the model (we go with
```

```
## Data: green_kbs
```

```
## Models:
```

```
## mod2: spp_half_cover_date ~ state * year + insecticide * year_factor +
```

```
## mod2: (1 | species)
```

```
## mod1: spp_half_cover_date ~ state * year + insecticide * year_factor +
```

```
## mod1: (1 | species) + (1 | plot)
```

```
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
```

```
## mod2      8 13587 13628 -6785.6    13571
```

```
## mod1      9 13586 13632 -6784.0    13568 3.374  1    0.06623 .
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(mod1)
```

```
## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
```

```
## method [lmerModLmerTest]
```

```
## Formula: spp_half_cover_date ~ state * year + insecticide * year_factor +
```

```
## (1 | species) + (1 | plot)
```

```
## Data: green_kbs
```

```
##
```

```
##      AIC      BIC    logLik deviance df.resid
```

```
## 13585.9 13632.2 -6784.0 13567.9      1259
```

```
##
```

```
## Scaled residuals:
```

```
##      Min       1Q   Median       3Q      Max
```

```
## -2.1194 -0.7683 -0.2513  0.6857  3.2607
```

```
##
```

```
## Random effects:
```

```
## Groups   Name      Variance Std.Dev.
```

```
## plot    (Intercept)  32.77    5.725
```

```
## species (Intercept) 930.34   30.502
```

```
## Residual                2449.77  49.495
```

```
## Number of obs: 1268, groups: plot, 24; species, 21
```

```
##
```

```
## Fixed effects:
```

```

##               Estimate Std. Error      df t value
## (Intercept)    -1.113e+04  3.368e+03  1.236e+03  -3.304
## stateambient    -2.812e+03  3.903e+03  1.236e+03  -0.721
## year            5.589e+00  1.669e+00  1.236e+03   3.349
## insecticideno_insects -3.927e-01  6.678e+00  2.112e+02  -0.059
## stateambient:year  1.396e+00  1.934e+00  1.236e+03   0.722
## insecticideno_insects:year_factor  2.351e+00  1.940e+00  1.234e+03   1.212
##               Pr(>|t|)
## (Intercept)    0.000979 ***
## stateambient    0.471285
## year            0.000836 ***
## insecticideno_insects  0.953167
## stateambient:year  0.470576
## insecticideno_insects:year_factor  0.225879
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##      (Intr) sttmbn year  insct_ sttmb:
## stateambint -0.583
## year        -1.000  0.583
## insctcdn_ns -0.463 -0.036  0.462
## statmbnt:yr  0.583 -1.000 -0.583  0.036
## insctcdn:_   0.519  0.062 -0.519 -0.835 -0.062
## fit warnings:
## fixed-effect model matrix is rank deficient so dropping 1 column / coefficient
## Some predictor variables are on very different scales: consider rescaling
summary(mod2)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state * year + insecticide * year_factor +
## (1 | species)
## Data: green_kbs
##
##      AIC      BIC  logLik deviance df.resid
## 13587.3 13628.4 -6785.6 13571.3     1260
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.2600 -0.7665 -0.2441  0.6924  3.2567
##
## Random effects:
##  Groups   Name      Variance Std.Dev.
## species (Intercept)  951.2    30.84
## Residual          2479.5    49.79
## Number of obs: 1268, groups: species, 21
##
## Fixed effects:
##               Estimate Std. Error      df t value
## (Intercept)    -1.095e+04  3.381e+03  1.247e+03  -3.237
## stateambient    -2.851e+03  3.918e+03  1.247e+03  -0.728
## year            5.498e+00  1.676e+00  1.247e+03   3.282
## insecticideno_insects -1.336e-01  6.284e+00  1.246e+03  -0.021

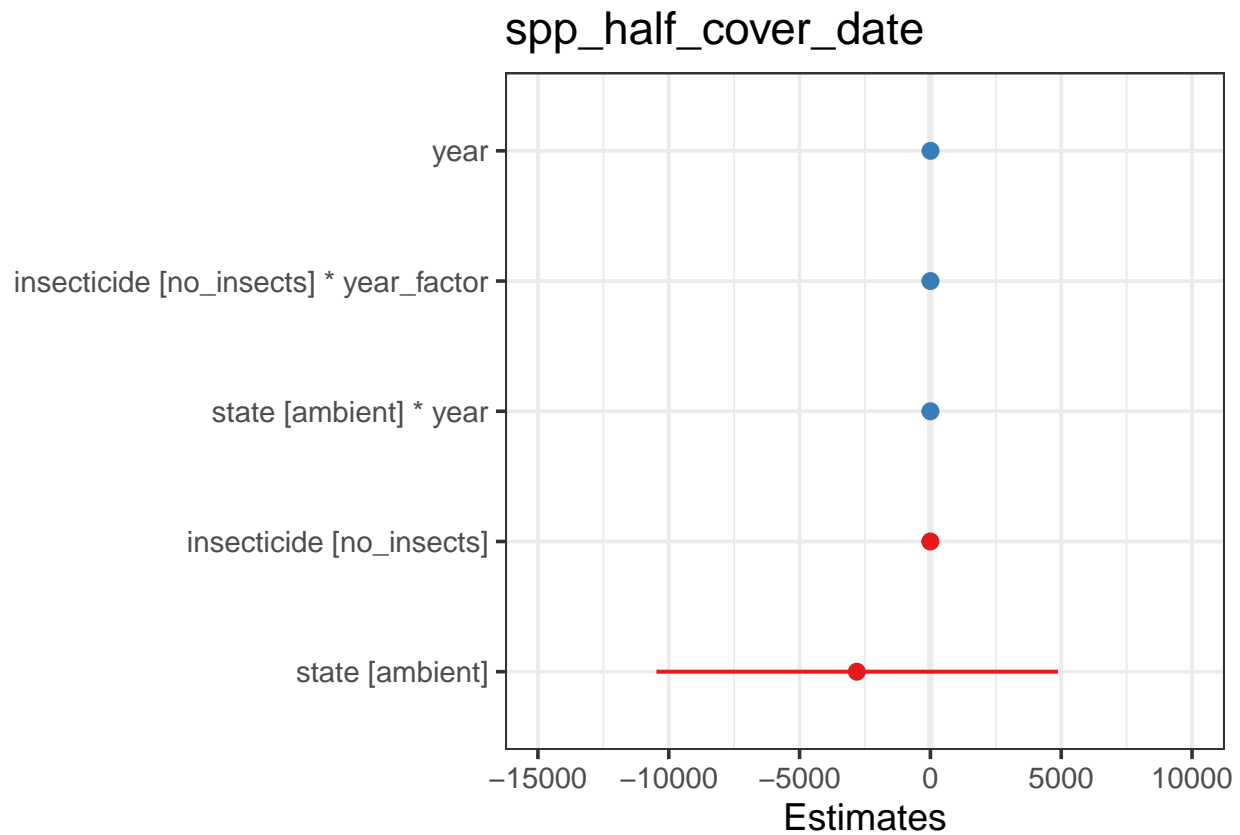
```

```

## stateambient:year          1.415e+00  1.942e+00  1.247e+03  0.729
## insecticideno_insects:year_factor  2.464e+00  1.949e+00  1.248e+03  1.264
##                               Pr(>|t|)
## (Intercept)                0.00124 **
## stateambient                0.46703
## year                        0.00106 **
## insecticideno_insects       0.98304
## stateambient:year           0.46635
## insecticideno_insects:year_factor  0.20647
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##      (Intr) sttmbn year  insct_ sttmb:
## stateambint -0.583
## year        -1.000  0.583
## insctcdn_ns -0.494 -0.039  0.494
## statmbnt:yr  0.583 -1.000 -0.583  0.039
## insctcdn:_   0.520  0.061 -0.520 -0.893 -0.061
## fit warnings:
## fixed-effect model matrix is rank deficient so dropping 1 column / coefficient
## Some predictor variables are on very different scales: consider rescaling

# Next, plot the model. There are multiple variables but
# here's one way to do it based on this package sjPlot:
# https://strengejacke.github.io/sjPlot/articles/plot\_model\_estimates.html
# Annoyingly, this package somehow overwrites the factor
# order in its plotting so we will have to modify the code to
# get warmed = red. I haven't figured this out yet. It does
# seem to work on some of the plots. hmm. ?plot_model Plot
# the fixed effects estimates for different models these are
# the fixed effects estimates from summary(mod5)
plot_model(mod1, sort.est = TRUE)

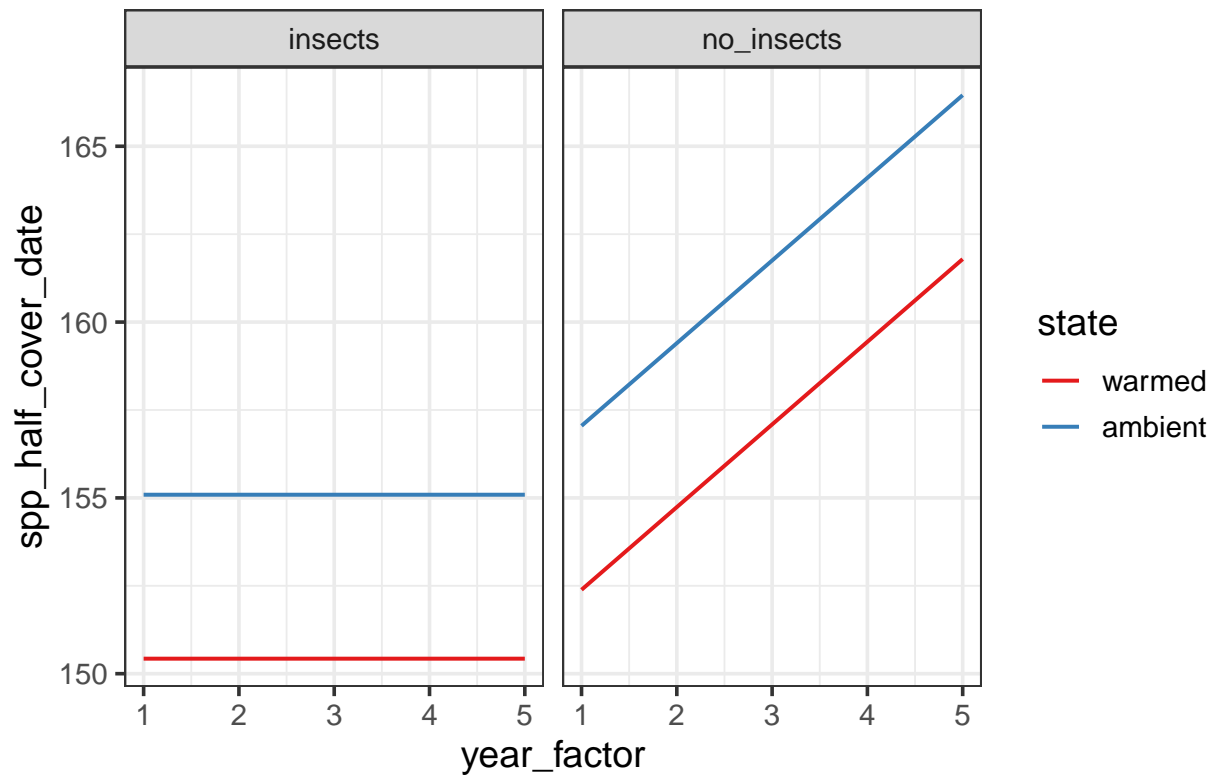
```



```
# these are the fixed predicted values:  
plot_model(mod1, type = "pred", terms = c("year_factor", "state",  
      "insecticide"))
```

```
## Error: Confidence intervals could not be computed.  
## * Reason: "non-conformable arguments"  
## * Source: mm %*% vcm
```

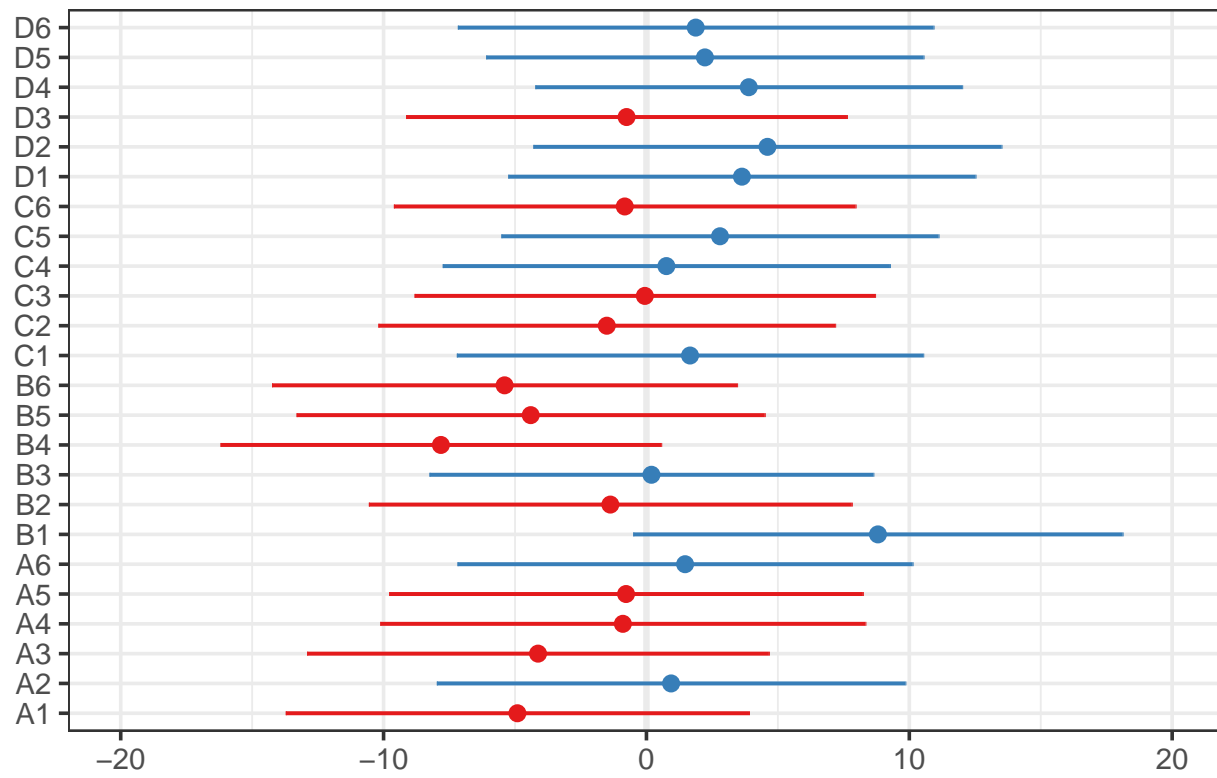
Predicted values of spp_half_cover_date



```
# these are the random effects estimates  
plot_model(mod1, type = "re", terms = c("species", "plot"))
```

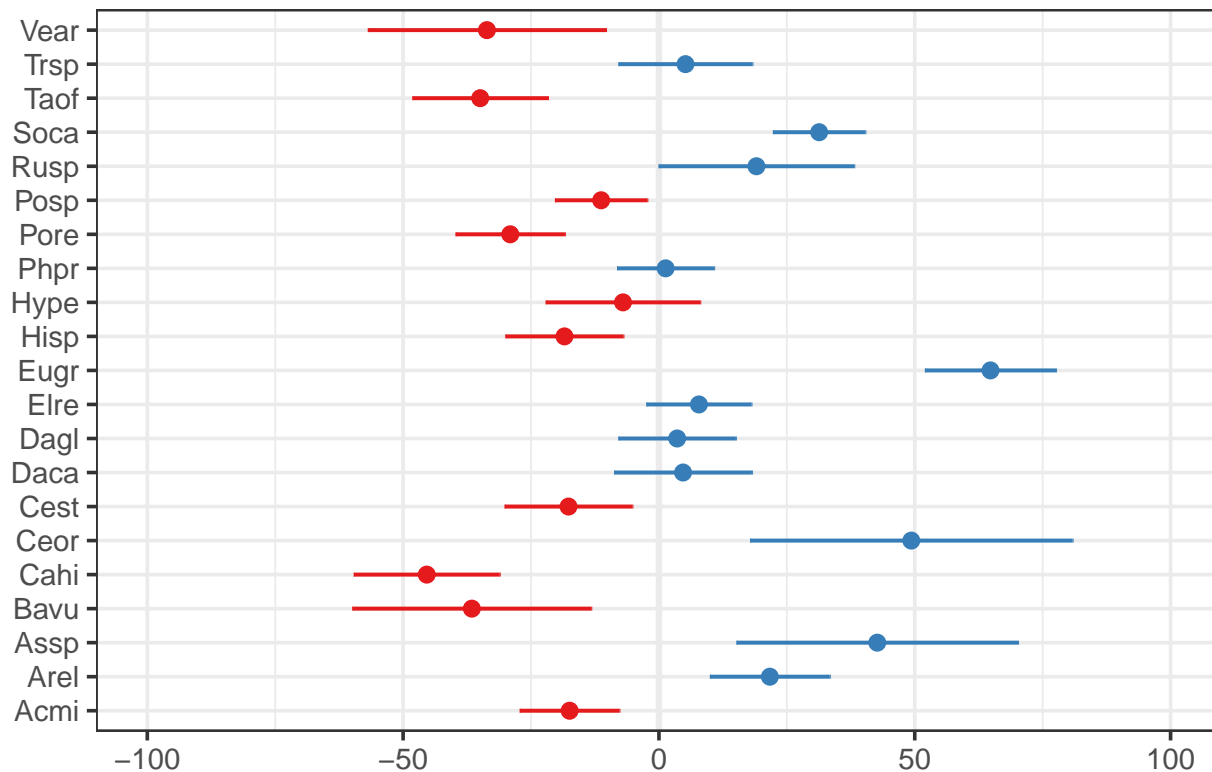
```
## [[1]]
```


Random effects



[[2]]

Random effects



Do we need to include insecticide?

```
mod3 <- lmer(spp_half_cover_date ~ state * year_factor + (1 |
  species), green_kbs, REML = FALSE)
anova(mod1, mod3)
```

Data: green_kbs

Models:

mod3: spp_half_cover_date ~ state * year_factor + (1 | species)

mod1: spp_half_cover_date ~ state * year + insecticide * year_factor +

mod1: (1 | species) + (1 | plot)

	npar	AIC	BIC	logLik	deviance	Chisq	Df	Pr(>Chisq)
## mod3	6	13591	13622	-6789.5	13579			
## mod1	9	13586	13632	-6784.0	13568	10.994	3	0.01176 *

mod3 6 13591 13622 -6789.5 13579

mod1 9 13586 13632 -6784.0 13568 10.994 3 0.01176 *

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
AICctab(mod1, mod3, weights = T)
```

	dAICc	df	weight
## mod1	0.0	9	0.921
## mod3	4.9	6	0.079

mod1 0.0 9 0.921

mod3 4.9 6 0.079

*# Looks like yes P<0.05, insecticide improves model fit so we
will continue to include it and stick with mod1*

Does year need to be interactive with insecticide?

```
mod4 <- lmer(spp_half_cover_date ~ state * year_factor + insecticide +
  (1 | species) + (1 | plot), green_kbs, REML = FALSE)
```

```
anova(mod1, mod4)
```

```

## Data: green_kbs
## Models:
## mod4: spp_half_cover_date ~ state * year_factor + insecticide + (1 |
## mod4:   species) + (1 | plot)
## mod1: spp_half_cover_date ~ state * year + insecticide * year_factor +
## mod1:   (1 | species) + (1 | plot)
##      npar   AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## mod4    8 13585 13626 -6784.7    13569
## mod1    9 13586 13632 -6784.0    13568 1.4664  1    0.2259

# No, P>0.05 so insecticide*year doesn't strongly improve
# model fit so we will shift to mod4
anova(mod3, mod4)

## Data: green_kbs
## Models:
## mod3: spp_half_cover_date ~ state * year_factor + (1 | species)
## mod4: spp_half_cover_date ~ state * year_factor + insecticide + (1 |
## mod4:   species) + (1 | plot)
##      npar   AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## mod3    6 13591 13622 -6789.5    13579
## mod4    8 13585 13626 -6784.7    13569 9.5277  2    0.008533 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

# Yes, P<0.05 so insecticide still improves model fit so we
# will stay with mod4

# Does year need to be interactive with state?
mod5 <- lmer(spp_half_cover_date ~ state + year_factor + insecticide +
  (1 | species) + (1 | plot), green_kbs, REML = FALSE)
anova(mod4, mod5)

## Data: green_kbs
## Models:
## mod5: spp_half_cover_date ~ state + year_factor + insecticide + (1 |
## mod5:   species) + (1 | plot)
## mod4: spp_half_cover_date ~ state * year_factor + insecticide + (1 |
## mod4:   species) + (1 | plot)
##      npar   AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## mod5    7 13584 13620 -6785.0    13570
## mod4    8 13585 13626 -6784.7    13569 0.6369  1    0.4249

AICcTab(mod4, mod5, weights = T)

##      dAICc df weight
## mod5 0.0    7  0.67
## mod4 1.4    8  0.33

# No, P>0.05 so state*year doesn't improve model fit so we
# could drop it and go with mod5, but note that the AIC
# values are super close. mod4 makes sense, with increased
# divergence between warmed and ambient.
summary(mod5)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]

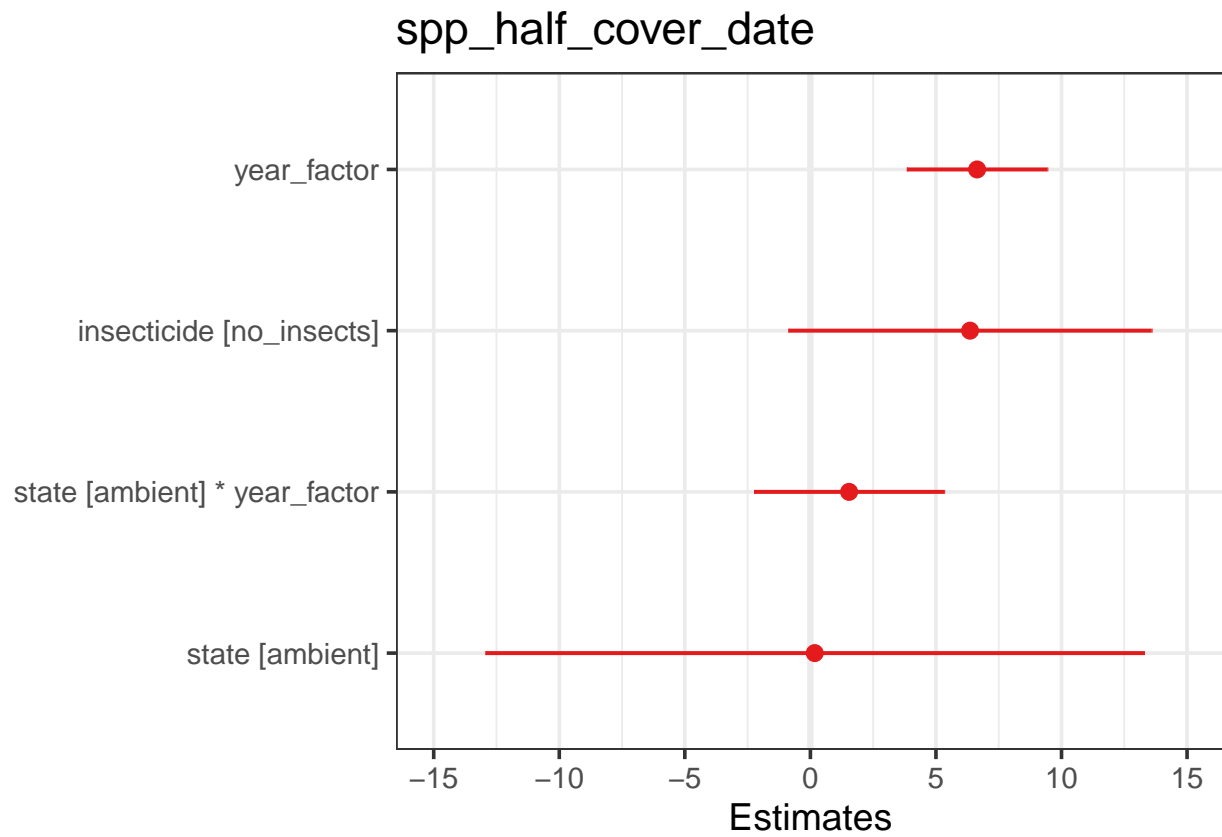
```

```

## Formula: spp_half_cover_date ~ state + year_factor + insecticide + (1 |
##   species) + (1 | plot)
## Data: green_kbs
##
##      AIC      BIC   logLik deviance df.resid
##  13584   13620   -6785   13570    1261
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.1128 -0.7727 -0.2469  0.6665  3.1808
##
## Random effects:
## Groups Name Variance Std.Dev.
## plot (Intercept) 33.61 5.797
## species (Intercept) 931.15 30.515
## Residual 2453.43 49.532
## Number of obs: 1268, groups: plot, 24; species, 21
##
## Fixed effects:
## Estimate Std. Error df t value Pr(>|t|)
## (Intercept) 128.0160 8.0014 33.2411 15.999 < 2e-16 ***
## stateambient 4.6225 3.7014 21.2778 1.249 0.2253
## year_factor 7.4606 0.9888 1240.6406 7.545 8.7e-14 ***
## insecticideno_insects 6.4410 3.6917 21.0085 1.745 0.0956 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
## (Intr) sttmbn yr_fct
## stateambint -0.221
## year_factor -0.358 -0.026
## insctcdn_ns -0.238 -0.038 0.059
anova(mod4)

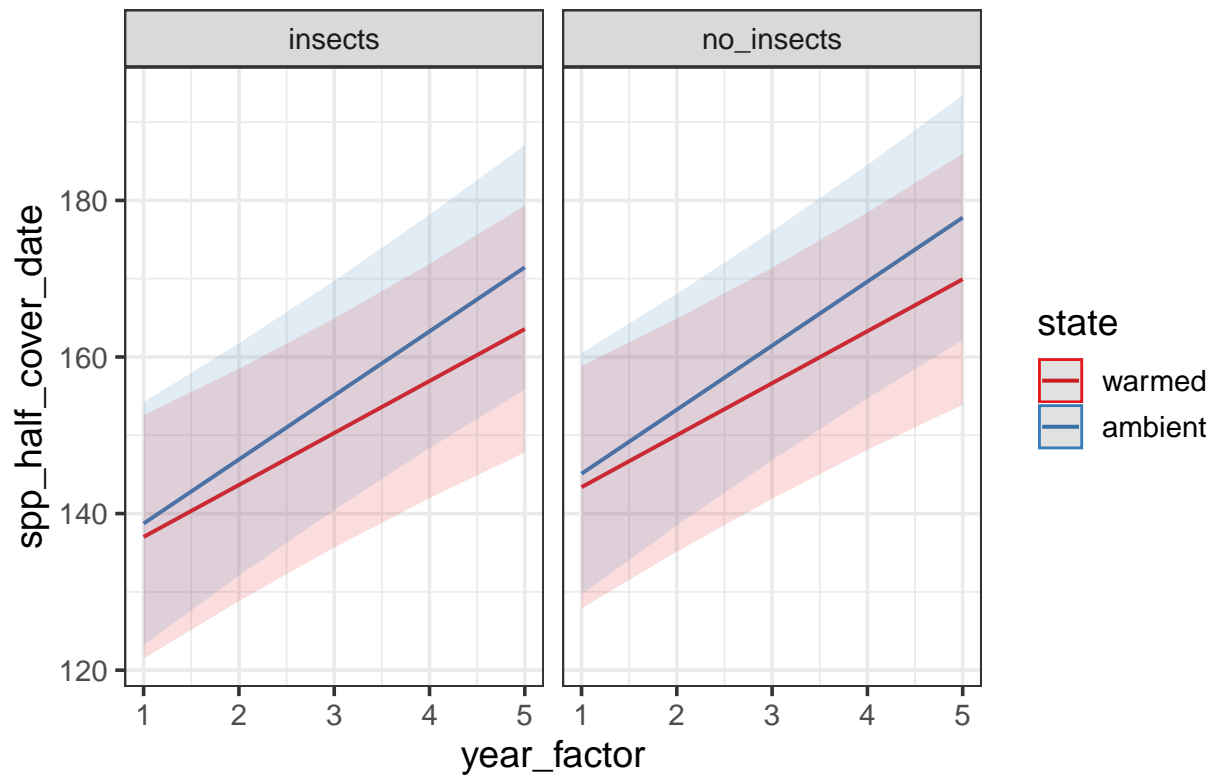
## Type III Analysis of Variance Table with Satterthwaite's method
## Sum Sq Mean Sq NumDF DenDF F value Pr(>F)
## state 2 2 1 208.20 0.0007 0.97938
## year_factor 137200 137200 1 1241.26 55.9464 1.404e-13 ***
## insecticide 7282 7282 1 21.05 2.9693 0.09952 .
## state:year_factor 1562 1562 1 1235.86 0.6371 0.42493
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
# these are the fixed effects estimates from summary(mod4)
plot_model(mod4, sort.est = TRUE)

```



```
# these are the fixed predicted values:  
plot_model(mod4, type = "pred", terms = c("year_factor", "state",  
      "insecticide"))
```

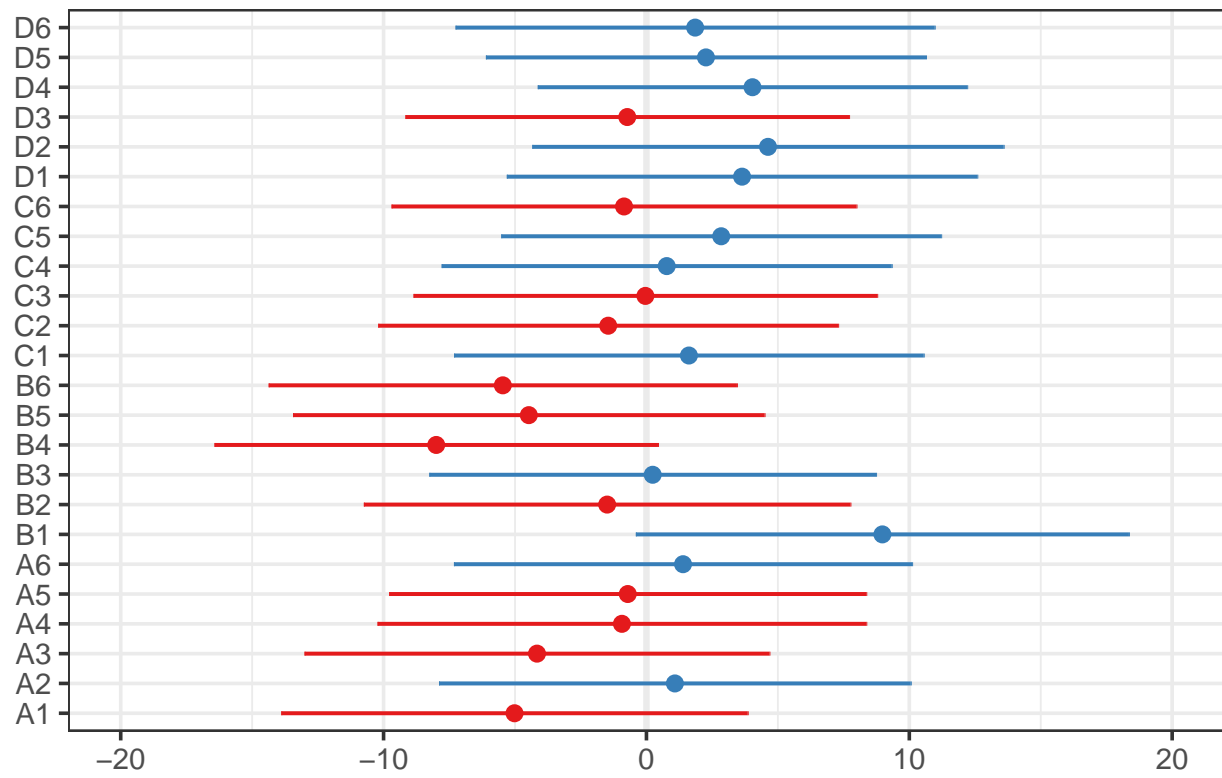
Predicted values of spp_half_cover_date



```
# these are the random effects estimates  
plot_model(mod4, type = "re", terms = c("species", "plot"))
```

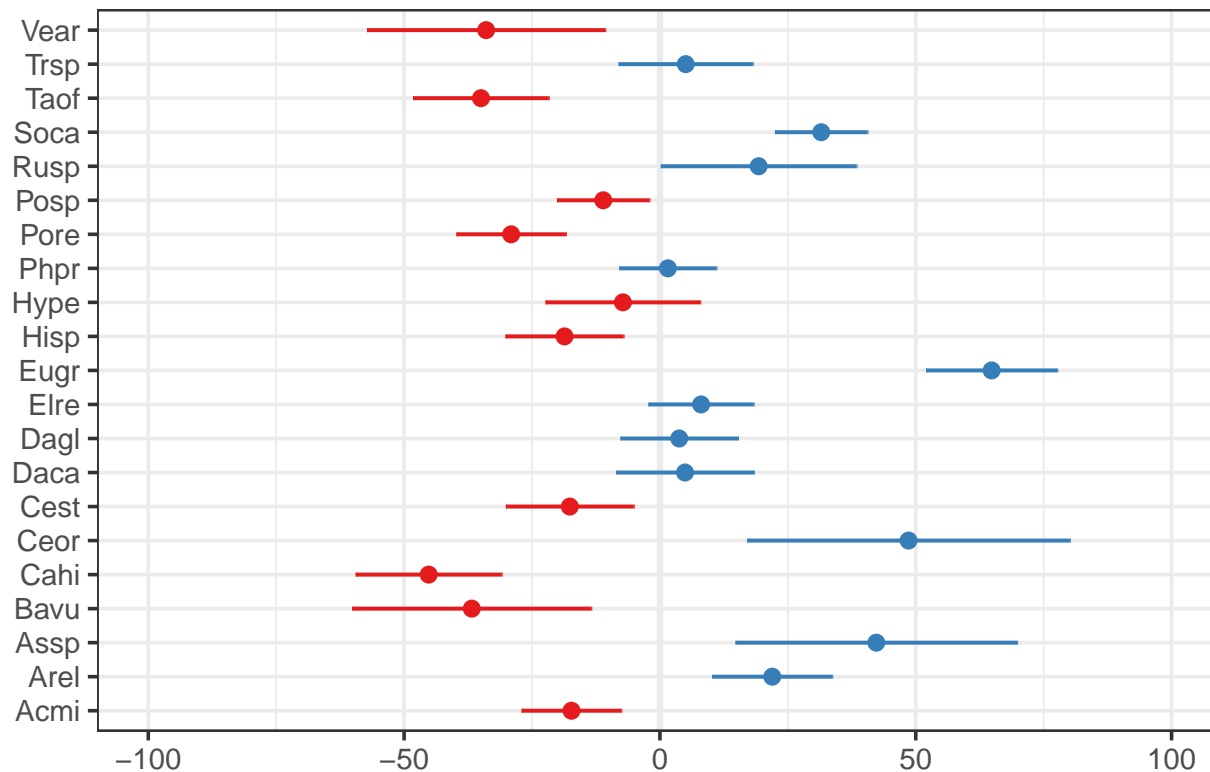
```
## [[1]]
```

Random effects



[[2]]

Random effects



*# If we wanted to include plots nested within year it would
look like this:*

```
mod6 <- lmer(spp_half_cover_date ~ state * year_factor + insecticide +
  (1 | species) + (1 + year | plot), green_kbs, REML = FALSE)
```

```
## boundary (singular) fit: see ?isSingular
```

```
## Warning: Model failed to converge with 1 negative eigenvalue: -1.6e+00
```

```
anova(mod4, mod6)
```

```
## Data: green_kbs
```

```
## Models:
```

```
## mod4: spp_half_cover_date ~ state * year_factor + insecticide + (1 |
```

```
## mod4:   species) + (1 | plot)
```

```
## mod6: spp_half_cover_date ~ state * year_factor + insecticide + (1 |
```

```
## mod6:   species) + (1 + year | plot)
```

```
##      npar   AIC   BIC logLik deviance Chisq Df Pr(>Chisq)
```

```
## mod4      8 13585 13626 -6784.7    13569
```

```
## mod6     10 13595 13647 -6787.6    13575      0  2          1
```

```
anova(mod6)
```

```
## Type III Analysis of Variance Table with Satterthwaite's method
```

```
##              Sum Sq Mean Sq NumDF    DenDF F value    Pr(>F)
```

```
## state              3         3      1    200.93  0.0013  0.9712
```

```
## year_factor    134914  134914      1   1238.02 55.6051 1.66e-13 ***
```

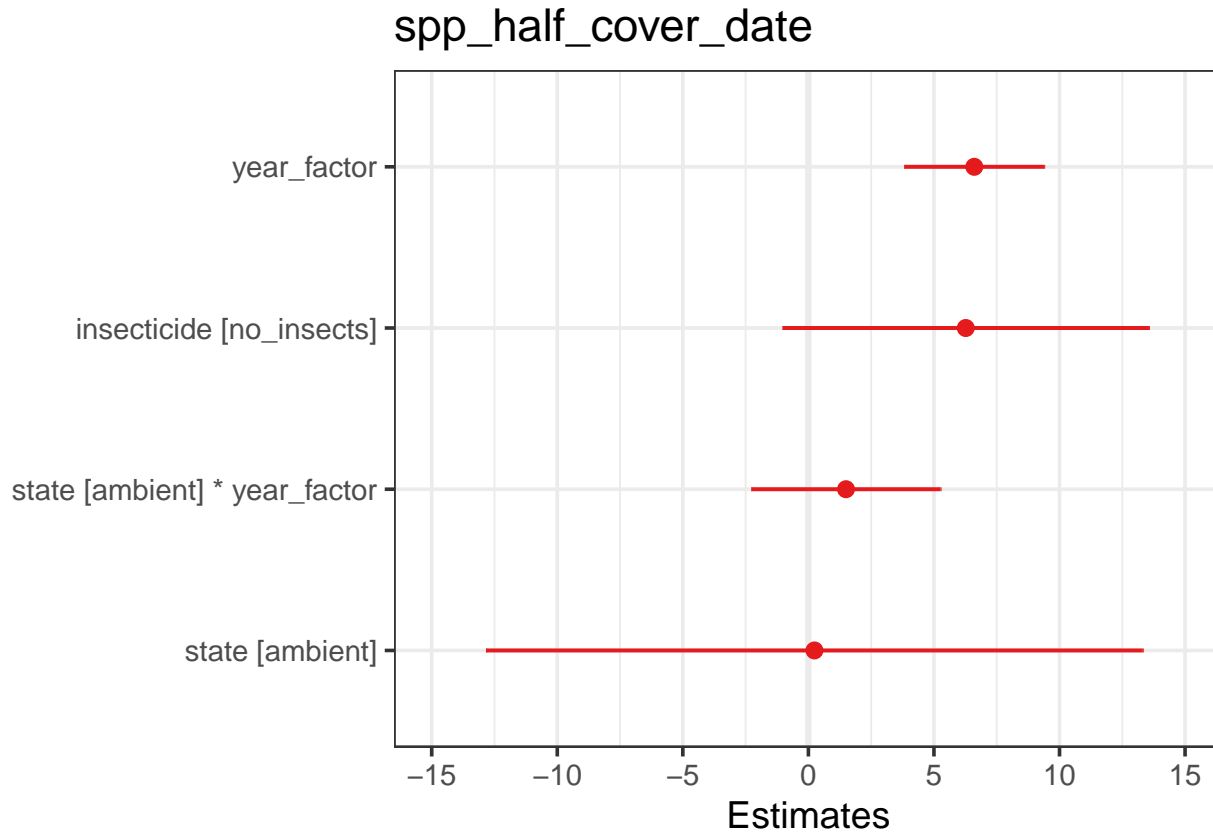
```
## insecticide      6882     6882      1     20.59  2.8364  0.1072
```

```
## state:year_factor  1474     1474      1   1231.93  0.6074  0.4359
```

```
## ---
```

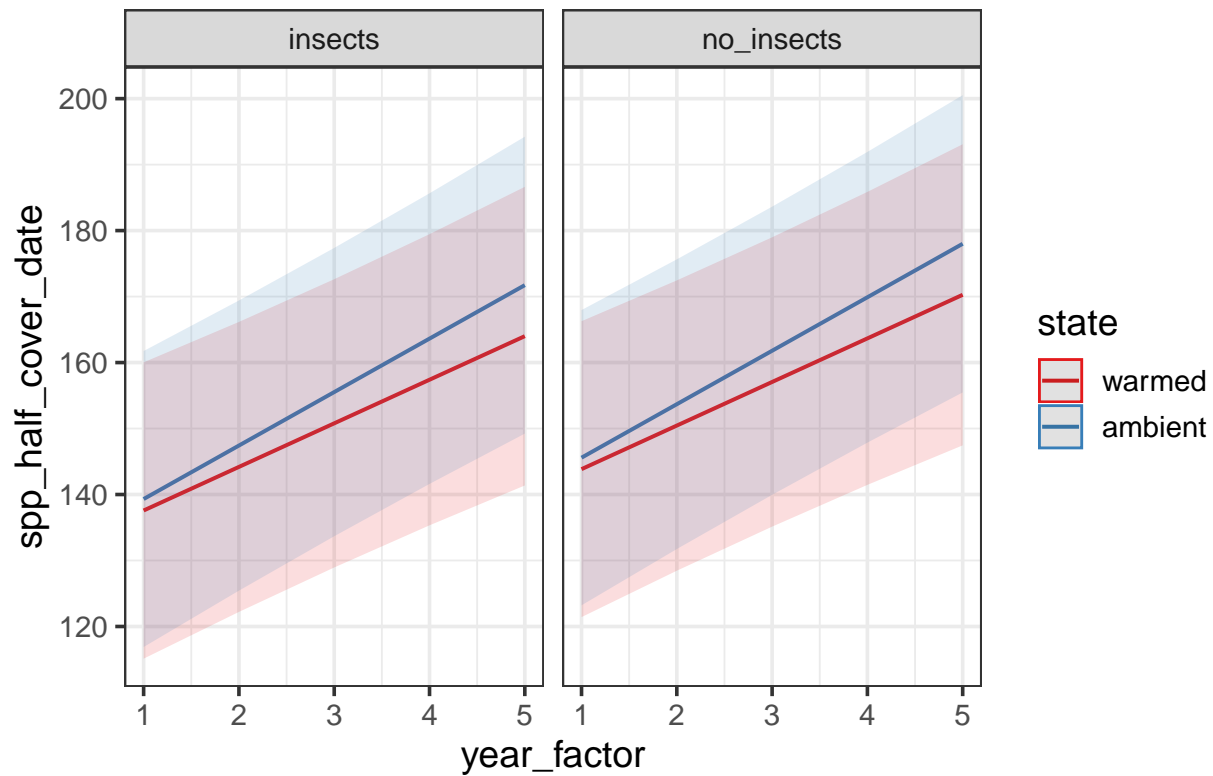


```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
# Yup, seems to matter but it is making this more complex,
# though not overly so because it's on the random effects
# structure only.
plot_model(mod6, sort.est = TRUE)
```



```
# these are the fixed predicted values:
plot_model(mod6, type = "pred", terms = c("year_factor", "state",
      "insecticide"))
```

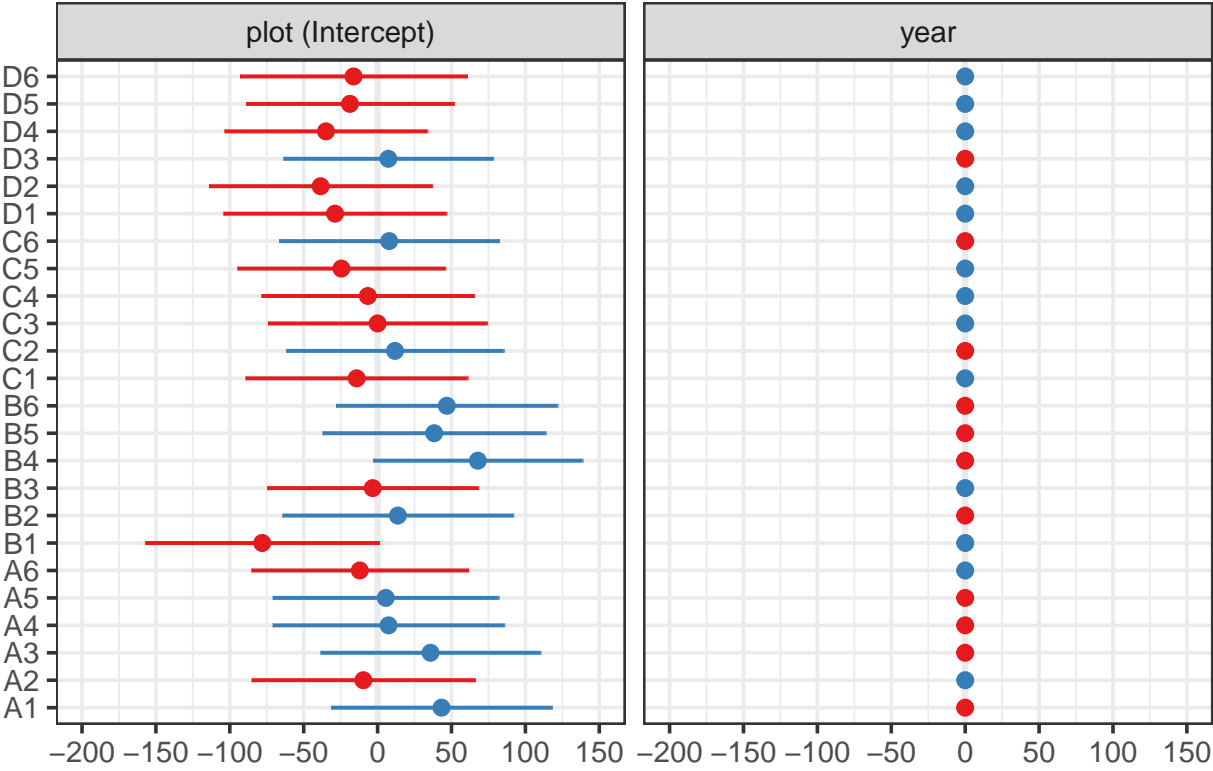
Predicted values of spp_half_cover_date



```
# these are the random effects estimates  
plot_model(mod6, type = "re", terms = c("species", "plot"))
```

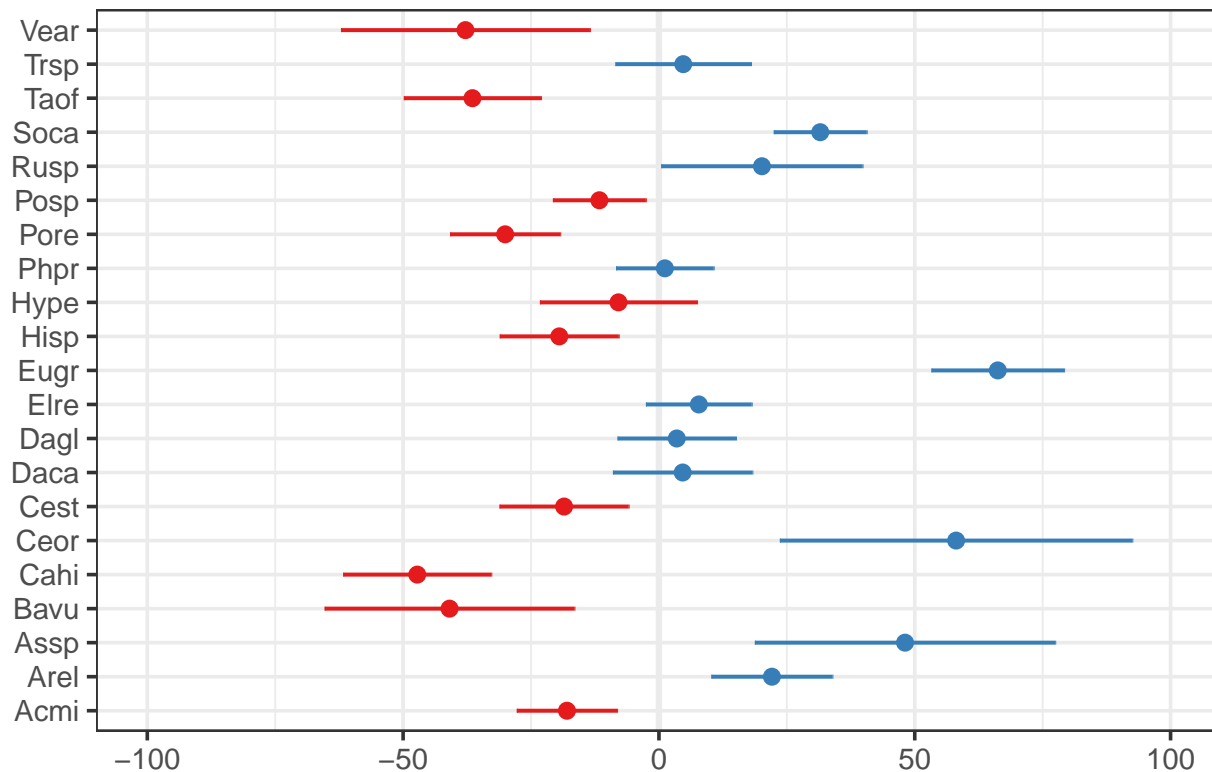
```
## [[1]]
```

Random effects



[[2]]

Random effects



mod4 (and mod6) are pretty complex in terms of interpretation (they actually don't have many parameters though). We could consider an alternative model that's simpler to understand and also one that provides more insight about the species. That would be something like this:

```
mod7 <- lmer(spp_half_cover_date ~ state + species + (1 + year_factor | plot), green_kbs, REML = FALSE)
```

boundary (singular) fit: see ?isSingular

```
mod7a <- lmer(spp_half_cover_date ~ state + species + year_factor + (1 | plot), green_kbs, REML = FALSE)
```

```
mod7b <- lmer(spp_half_cover_date ~ state * year_factor + species + (1 | plot), green_kbs, REML = FALSE)
```

```
mod7c <- lmer(spp_half_cover_date ~ state + species + year_factor + insecticide + (1 | plot), green_kbs, REML = FALSE)
```

```
anova(mod6, mod7) # model 7 is a better fit to data
```

Data: green_kbs

Models:

```
## mod6: spp_half_cover_date ~ state * year_factor + insecticide + (1 | species) + (1 + year | plot)
```

```
## mod7: spp_half_cover_date ~ state + species + (1 + year_factor | plot)
```

```
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
```

```
## mod6    10 13595 13647 -6787.6    13575
```

```
## mod7    26 13555 13688 -6751.3    13503 72.591 16 3.495e-09 ***
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

anova(mod7, mod7a) #mod 7a

## Data: green_kbs
## Models:
## mod7a: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## mod7: spp_half_cover_date ~ state + species + (1 + year_factor | plot)
##      npar   AIC   BIC logLik deviance Chisq Df Pr(>Chisq)
## mod7a   25 13539 13668 -6744.6    13489
## mod7    26 13555 13688 -6751.3    13503      0  1          1

anova(mod7a, mod7b) #mod 7a.2 - interaction between state and year doesn't improve model

## Data: green_kbs
## Models:
## mod7a: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## mod7b: spp_half_cover_date ~ state * year_factor + species + (1 | plot)
##      npar   AIC   BIC logLik deviance Chisq Df Pr(>Chisq)
## mod7a   25 13539 13668 -6744.6    13489
## mod7b   26 13540 13674 -6744.2    13488 0.6487  1    0.4206

anova(mod7a, mod7c) #mod 7a.2 slightly better

## Data: green_kbs
## Models:
## mod7a: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## mod7c: spp_half_cover_date ~ state + species + year_factor + insecticide +
## mod7c:      (1 | plot)
##      npar   AIC   BIC logLik deviance Chisq Df Pr(>Chisq)
## mod7a   25 13539 13668 -6744.6    13489
## mod7c   26 13538 13672 -6743.1    13486 2.9664  1    0.08501 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

summary(mod7a)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## Data: green_kbs
##
##      AIC      BIC logLik deviance df.resid
## 13539.1 13667.8 -6744.6 13489.1      1243
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.1388 -0.7706 -0.2263  0.6478  3.2151
##
## Random effects:
## Groups Name Variance Std.Dev.
## plot (Intercept) 40.94 6.398
## Residual 2411.99 49.112
## Number of obs: 1268, groups: plot, 24
##
## Fixed effects:
## Estimate Std. Error df t value Pr(>|t|)
```

```
## (Intercept) 113.9583 6.0649 450.7559 18.790 < 2e-16 ***
## stateambient 4.7486 3.8464 23.6837 1.235 0.229105
## speciesArel 40.4606 7.6944 1258.0613 5.258 1.71e-07 ***
## speciesAssp 71.3590 16.4037 1267.3877 4.350 1.47e-05 ***
## speciesBavu -25.3647 13.7051 1267.2627 -1.851 0.064438 .
## speciesCahi -30.3892 8.8440 1255.5956 -3.436 0.000609 ***
## speciesCeor 84.6039 19.4064 1264.4777 4.360 1.41e-05 ***
## speciesCest -0.9968 8.0302 1251.8628 -0.124 0.901234
## speciesDaca 22.8735 8.4537 1254.0641 2.706 0.006907 **
## speciesDagl 21.9625 7.5879 1257.1006 2.894 0.003864 **
## speciesElre 26.1058 7.1253 1250.8183 3.664 0.000259 ***
## speciesEugr 85.6023 8.1688 1262.0879 10.479 < 2e-16 ***
## speciesHisp -1.9860 7.5978 1248.6621 -0.261 0.793833
## speciesHype 10.0686 9.2223 1264.8458 1.092 0.275143
## speciesPhpr 19.4282 6.7908 1248.0698 2.861 0.004294 **
## speciesPore -12.4161 7.2532 1254.2066 -1.712 0.087177 .
## speciesPosp 6.3858 6.6222 1245.8443 0.964 0.335077
## speciesRusp 39.8126 11.2911 1264.8911 3.526 0.000437 ***
## speciesSoca 49.9275 6.6222 1245.8443 7.539 9.05e-14 ***
## speciesTaof -19.2468 8.3548 1263.4009 -2.304 0.021402 *
## speciesTrsp 22.6301 8.2858 1252.1765 2.731 0.006399 **
## speciesVear -22.9444 13.6644 1262.9713 -1.679 0.093372 .
## year_factor 7.2772 0.9805 1259.8102 7.422 2.12e-13 ***
## ---
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
```

```
## Correlation matrix not shown by default, as p = 23 > 12.
```

```
## Use print(x, correlation=TRUE) or
```

```
## vcov(x) if you need it
```

```
anova(mod7a) # investigates whether at least one of the levels within each factor is significantly dif
```

```
## Type III Analysis of Variance Table with Satterthwaite's method
```

```
## Sum Sq Mean Sq NumDF DenDF F value Pr(>F)
## state 3676 3676 1 23.68 1.5241 0.2291
## species 890400 44520 20 1259.23 18.4578 < 2.2e-16 ***
## year_factor 132866 132866 1 1259.81 55.0857 2.116e-13 ***
## ---
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
# Yes, at least one of the species is different (they do not
# all have the same half cover dates).
```

```
emmeans(mod7a, list(pairwise ~ year_factor), adjust = "tukey")
```

```
## $`emmeans of year_factor`
```

```
## year_factor emmean SE df lower.CL upper.CL
```

```
## 2.9 156 2.42 55.1 151 161
```

```
##
```

```
## Results are averaged over the levels of: state, species
```

```
## Degrees-of-freedom method: kenward-roger
```

```
## Confidence level used: 0.95
```

```
##
```

```
## $` of year_factor`
```

```
## 1 estimate SE df z.ratio p.value
```

```
## (nothing) nonEst NA NA NA NA
```

```
##
## Results are averaged over the levels of: state, species
## Degrees-of-freedom method: kenward-roger
```

```
emmeans(mod7a, list(pairwise ~ species), adjust = "tukey")
```

```
## $`emmeans of species`
## species emmean SE df lower.CL upper.CL
## Acmi 137 5.12 730 127.4 147
## Arel 178 6.19 931 165.8 190
## Assp 209 15.89 1284 177.6 240
## Bavu 112 13.03 1263 86.5 138
## Cahi 107 7.60 1135 92.1 122
## Ceor 222 19.05 1275 184.7 259
## Cest 136 6.63 1036 123.4 149
## Daca 160 7.14 1099 146.3 174
## Dagl 159 6.05 920 147.5 171
## Elre 164 5.42 812 152.9 174
## Eugr 223 6.78 1005 209.7 236
## Hisp 135 6.07 952 123.6 147
## Hype 148 8.04 1126 131.7 163
## Phpr 157 4.98 691 147.1 167
## Pore 125 5.62 837 114.0 136
## Posp 144 4.74 625 134.5 153
## Rusp 177 10.40 1191 156.8 198
## Soca 187 4.74 625 178.1 197
## Taof 118 7.00 1047 104.4 132
## Trsp 160 6.92 1079 146.5 174
## Vear 114 12.98 1280 89.0 140
```

```
##
## Results are averaged over the levels of: state
## Degrees-of-freedom method: kenward-roger
## Confidence level used: 0.95
```

```
## $`pairwise differences of species`
## 1 estimate SE df t.ratio p.value
## Acmi - Arel -40.461 7.77 1280 -5.210 <.0001
## Acmi - Assp -71.359 16.58 1290 -4.305 0.0032
## Acmi - Bavu 25.365 13.85 1290 1.832 0.9634
## Acmi - Cahi 30.389 8.92 1277 3.405 0.0854
## Acmi - Ceor -84.604 19.63 1290 -4.309 0.0032
## Acmi - Cest 0.997 8.10 1274 0.123 1.0000
## Acmi - Daca -22.874 8.53 1276 -2.682 0.4662
## Acmi - Dagl -21.962 7.66 1279 -2.868 0.3307
## Acmi - Elre -26.106 7.19 1272 -3.632 0.0417
## Acmi - Eugr -85.602 8.25 1284 -10.379 <.0001
## Acmi - Hisp 1.986 7.66 1270 0.259 1.0000
## Acmi - Hype -10.069 9.31 1287 -1.081 1.0000
## Acmi - Phpr -19.428 6.85 1270 -2.836 0.3522
## Acmi - Pore 12.416 7.32 1276 1.696 0.9839
## Acmi - Posp -6.386 6.68 1267 -0.956 1.0000
## Acmi - Rusp -39.813 11.42 1290 -3.485 0.0669
## Acmi - Soca -49.928 6.68 1267 -7.475 <.0001
## Acmi - Taof 19.247 8.44 1286 2.281 0.7696
## Acmi - Trsp -22.630 8.36 1274 -2.707 0.4466
```

##	Acmi - Vear	22.944	13.80	1285	1.663	0.9872
##	Arel - Assp	-30.898	16.95	1291	-1.823	0.9652
##	Arel - Bavu	65.825	14.26	1290	4.616	0.0008
##	Arel - Cahi	70.850	9.63	1286	7.359	<.0001
##	Arel - Ceor	-44.143	19.94	1290	-2.214	0.8120
##	Arel - Cest	41.457	8.87	1286	4.675	0.0006
##	Arel - Daca	17.587	9.23	1280	1.905	0.9460
##	Arel - Dagl	18.498	8.38	1280	2.208	0.8156
##	Arel - Elre	14.355	7.95	1281	1.806	0.9683
##	Arel - Eugr	-45.142	8.94	1289	-5.051	0.0001
##	Arel - Hisp	42.447	8.42	1279	5.043	0.0001
##	Arel - Hype	30.392	9.92	1288	3.063	0.2147
##	Arel - Phpr	21.032	7.68	1281	2.740	0.4219
##	Arel - Pore	52.877	8.10	1283	6.528	<.0001
##	Arel - Posp	34.075	7.52	1281	4.530	0.0012
##	Arel - Rusp	0.648	11.90	1291	0.054	1.0000
##	Arel - Soca	-9.467	7.52	1281	-1.259	0.9997
##	Arel - Taof	59.707	9.12	1286	6.550	<.0001
##	Arel - Trsp	17.831	9.06	1280	1.967	0.9275
##	Arel - Vear	63.405	14.25	1285	4.450	0.0017
##	Assp - Bavu	96.724	20.45	1291	4.730	0.0005
##	Assp - Cahi	101.748	17.45	1287	5.830	<.0001
##	Assp - Ceor	-13.245	24.74	1291	-0.535	1.0000
##	Assp - Cest	72.356	17.11	1290	4.230	0.0044
##	Assp - Daca	48.485	17.26	1286	2.808	0.3717
##	Assp - Dagl	49.397	16.88	1290	2.926	0.2931
##	Assp - Elre	45.253	16.67	1290	2.714	0.4411
##	Assp - Eugr	-14.243	17.11	1287	-0.832	1.0000
##	Assp - Hisp	73.345	16.88	1289	4.345	0.0027
##	Assp - Hype	61.290	17.70	1291	3.463	0.0717
##	Assp - Phpr	51.931	16.53	1290	3.142	0.1767
##	Assp - Pore	83.775	16.72	1288	5.012	0.0001
##	Assp - Posp	64.973	16.46	1290	3.946	0.0136
##	Assp - Rusp	31.546	18.83	1289	1.675	0.9860
##	Assp - Soca	21.431	16.46	1290	1.302	0.9995
##	Assp - Taof	90.606	17.22	1288	5.261	<.0001
##	Assp - Trsp	48.729	17.18	1287	2.837	0.3516
##	Assp - Vear	94.303	20.30	1278	4.645	0.0007
##	Bavu - Cahi	5.024	14.91	1287	0.337	1.0000
##	Bavu - Ceor	-109.969	23.06	1288	-4.769	0.0004
##	Bavu - Cest	-24.368	14.52	1291	-1.678	0.9857
##	Bavu - Daca	-48.238	14.74	1290	-3.273	0.1248
##	Bavu - Dagl	-47.327	14.23	1291	-3.327	0.1074
##	Bavu - Elre	-51.471	13.96	1291	-3.688	0.0345
##	Bavu - Eugr	-110.967	14.50	1288	-7.655	<.0001
##	Bavu - Hisp	-23.379	14.24	1291	-1.642	0.9889
##	Bavu - Hype	-35.433	15.17	1291	-2.336	0.7322
##	Bavu - Phpr	-44.793	13.80	1291	-3.246	0.1344
##	Bavu - Pore	-12.949	14.08	1292	-0.920	1.0000
##	Bavu - Posp	-31.751	13.72	1291	-2.314	0.7472
##	Bavu - Rusp	-65.177	16.56	1291	-3.935	0.0142
##	Bavu - Soca	-75.292	13.72	1291	-5.488	<.0001
##	Bavu - Taof	-6.118	14.68	1292	-0.417	1.0000
##	Bavu - Trsp	-47.995	14.59	1289	-3.289	0.1193

##	Bavu - Vear	-2.420	18.27	1289	-0.132	1.0000
##	Cahi - Ceor	-114.993	20.39	1292	-5.641	<.0001
##	Cahi - Cest	-29.392	9.84	1280	-2.987	0.2560
##	Cahi - Dac	-53.263	10.19	1278	-5.227	<.0001
##	Cahi - Dagl	-52.352	9.53	1286	-5.492	<.0001
##	Cahi - Elre	-56.495	9.15	1282	-6.172	<.0001
##	Cahi - Eugr	-115.991	9.97	1278	-11.636	<.0001
##	Cahi - Hisp	-28.403	9.51	1280	-2.987	0.2560
##	Cahi - Hype	-40.458	10.86	1283	-3.726	0.0303
##	Cahi - Phpr	-49.817	8.86	1280	-5.623	<.0001
##	Cahi - Pore	-17.973	9.25	1285	-1.943	0.9351
##	Cahi - Posp	-36.775	8.74	1280	-4.210	0.0048
##	Cahi - Rusp	-70.202	12.73	1291	-5.513	<.0001
##	Cahi - Soca	-80.317	8.74	1280	-9.194	<.0001
##	Cahi - Taof	-11.142	10.13	1285	-1.100	1.0000
##	Cahi - Trsp	-53.019	10.02	1272	-5.289	<.0001
##	Cahi - Vear	-7.445	14.83	1279	-0.502	1.0000
##	Ceor - Cest	85.601	20.07	1290	4.266	0.0038
##	Ceor - Dac	61.730	20.25	1290	3.049	0.2221
##	Ceor - Dagl	62.641	19.91	1289	3.146	0.1750
##	Ceor - Elre	58.498	19.70	1291	2.969	0.2670
##	Ceor - Eugr	-0.998	20.14	1289	-0.050	1.0000
##	Ceor - Hisp	86.590	19.89	1291	4.354	0.0026
##	Ceor - Hype	74.535	20.64	1287	3.612	0.0445
##	Ceor - Phpr	65.176	19.61	1289	3.324	0.1083
##	Ceor - Pore	97.020	19.79	1288	4.903	0.0002
##	Ceor - Posp	78.218	19.53	1290	4.005	0.0109
##	Ceor - Rusp	44.791	21.55	1291	2.078	0.8839
##	Ceor - Soca	34.676	19.53	1290	1.775	0.9735
##	Ceor - Taof	103.851	20.17	1291	5.148	0.0001
##	Ceor - Trsp	61.974	20.15	1291	3.075	0.2087
##	Ceor - Vear	107.548	22.98	1291	4.680	0.0006
##	Cest - Dac	-23.870	9.46	1277	-2.523	0.5910
##	Cest - Dagl	-22.959	8.76	1284	-2.622	0.5128
##	Cest - Elre	-27.103	8.37	1277	-3.236	0.1380
##	Cest - Eugr	-86.599	9.29	1287	-9.320	<.0001
##	Cest - Hisp	0.989	8.73	1272	0.113	1.0000
##	Cest - Hype	-11.065	10.24	1287	-1.081	1.0000
##	Cest - Phpr	-20.425	8.05	1277	-2.538	0.5791
##	Cest - Pore	11.419	8.42	1274	1.357	0.9991
##	Cest - Posp	-7.383	7.90	1275	-0.934	1.0000
##	Cest - Rusp	-40.809	12.19	1289	-3.346	0.1014
##	Cest - Soca	-50.924	7.90	1275	-6.442	<.0001
##	Cest - Taof	18.250	9.41	1284	1.938	0.9366
##	Cest - Trsp	-23.627	9.34	1276	-2.529	0.5864
##	Cest - Vear	21.948	14.41	1286	1.523	0.9956
##	Dac - Dagl	0.911	9.12	1276	0.100	1.0000
##	Dac - Elre	-3.232	8.77	1277	-0.369	1.0000
##	Dac - Eugr	-62.729	9.64	1284	-6.505	<.0001
##	Dac - Hisp	24.860	9.12	1274	2.725	0.4333
##	Dac - Hype	12.805	10.57	1287	1.211	0.9998
##	Dac - Phpr	3.445	8.47	1276	0.407	1.0000
##	Dac - Pore	35.290	8.83	1275	3.998	0.0112
##	Dac - Posp	16.488	8.34	1277	1.978	0.9239

##	Daca	-	Rusp	-16.939	12.46	1291	-1.360	0.9990
##	Daca	-	Soca	-27.054	8.34	1277	-3.245	0.1348
##	Daca	-	Taof	42.120	9.79	1285	4.304	0.0033
##	Daca	-	Trsp	0.243	9.70	1273	0.025	1.0000
##	Daca	-	Vear	45.818	14.63	1282	3.132	0.1813
##	Dagl	-	Elre	-4.143	7.84	1275	-0.528	1.0000
##	Dagl	-	Eugr	-63.640	8.86	1289	-7.187	<.0001
##	Dagl	-	Hisp	23.948	8.32	1279	2.879	0.3233
##	Dagl	-	Hype	11.894	9.85	1289	1.208	0.9998
##	Dagl	-	Phpr	2.534	7.56	1274	0.335	1.0000
##	Dagl	-	Pore	34.379	7.99	1279	4.305	0.0032
##	Dagl	-	Posp	15.577	7.41	1277	2.102	0.8725
##	Dagl	-	Rusp	-17.850	11.84	1291	-1.508	0.9961
##	Dagl	-	Soca	-27.965	7.41	1277	-3.775	0.0255
##	Dagl	-	Taof	41.209	9.02	1284	4.570	0.0010
##	Dagl	-	Trsp	-0.668	8.97	1281	-0.074	1.0000
##	Dagl	-	Vear	44.907	14.18	1285	3.167	0.1659
##	Elre	-	Eugr	-59.496	8.42	1286	-7.066	<.0001
##	Elre	-	Hisp	28.092	7.89	1274	3.561	0.0526
##	Elre	-	Hype	16.037	9.48	1290	1.691	0.9844
##	Elre	-	Phpr	6.678	7.07	1269	0.945	1.0000
##	Elre	-	Pore	38.522	7.55	1275	5.103	0.0001
##	Elre	-	Posp	19.720	6.90	1268	2.857	0.3382
##	Elre	-	Rusp	-13.707	11.53	1292	-1.188	0.9999
##	Elre	-	Soca	-23.822	6.90	1268	-3.451	0.0743
##	Elre	-	Taof	45.353	8.62	1281	5.261	<.0001
##	Elre	-	Trsp	3.476	8.57	1276	0.406	1.0000
##	Elre	-	Vear	49.050	13.96	1287	3.515	0.0610
##	Eugr	-	Hisp	87.588	8.87	1286	9.871	<.0001
##	Eugr	-	Hype	75.534	10.30	1290	7.330	<.0001
##	Eugr	-	Phpr	66.174	8.16	1285	8.112	<.0001
##	Eugr	-	Pore	98.018	8.58	1288	11.430	<.0001
##	Eugr	-	Posp	79.216	8.02	1286	9.881	<.0001
##	Eugr	-	Rusp	45.790	12.19	1291	3.755	0.0273
##	Eugr	-	Soca	35.675	8.02	1286	4.450	0.0017
##	Eugr	-	Taof	104.849	9.51	1284	11.024	<.0001
##	Eugr	-	Trsp	62.972	9.46	1280	6.654	<.0001
##	Eugr	-	Vear	108.547	14.49	1284	7.489	<.0001
##	Hisp	-	Hype	-12.055	9.87	1288	-1.221	0.9998
##	Hisp	-	Phpr	-21.414	7.59	1274	-2.822	0.3619
##	Hisp	-	Pore	10.430	8.00	1273	1.304	0.9995
##	Hisp	-	Posp	-8.372	7.43	1271	-1.127	0.9999
##	Hisp	-	Rusp	-41.799	11.89	1290	-3.517	0.0606
##	Hisp	-	Soca	-51.913	7.43	1271	-6.987	<.0001
##	Hisp	-	Taof	17.261	9.03	1280	1.912	0.9442
##	Hisp	-	Trsp	-24.616	8.96	1272	-2.746	0.4173
##	Hisp	-	Vear	20.958	14.17	1284	1.479	0.9970
##	Hype	-	Phpr	-9.360	9.23	1287	-1.014	1.0000
##	Hype	-	Pore	22.485	9.61	1290	2.340	0.7295
##	Hype	-	Posp	3.683	9.11	1288	0.404	1.0000
##	Hype	-	Rusp	-29.744	12.99	1291	-2.289	0.7642
##	Hype	-	Soca	-39.859	9.11	1288	-4.373	0.0024
##	Hype	-	Taof	29.315	10.48	1291	2.797	0.3800
##	Hype	-	Trsp	-12.562	10.41	1286	-1.206	0.9998

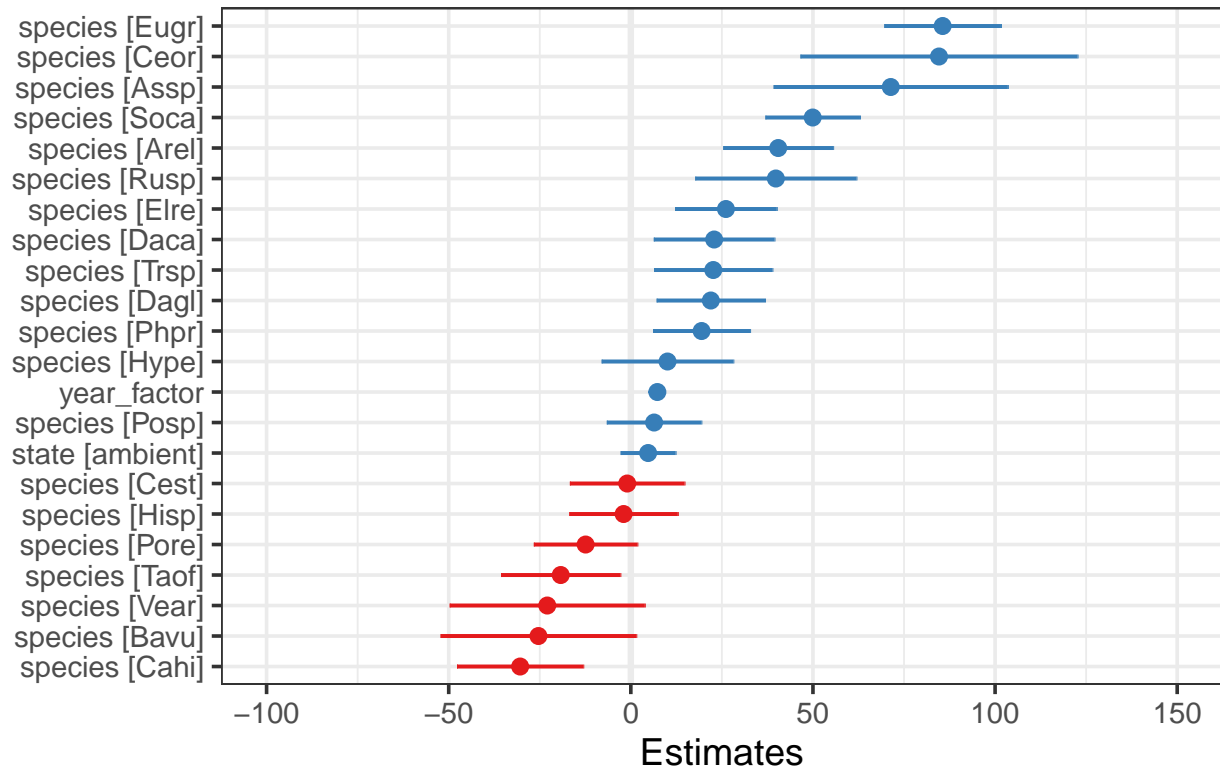
```

## Hype - Vear 33.013 15.11 1283 2.185 0.8292
## Phpr - Pore 31.844 7.23 1273 4.406 0.0021
## Phpr - Posp 13.042 6.57 1267 1.984 0.9215
## Phpr - Rusp -20.384 11.35 1291 -1.796 0.9700
## Phpr - Soca -30.499 6.57 1267 -4.640 0.0007
## Phpr - Taof 38.675 8.36 1285 4.629 0.0008
## Phpr - Trsp -3.202 8.28 1275 -0.387 1.0000
## Phpr - Vear 42.373 13.75 1285 3.081 0.2058
## Pore - Posp -18.802 7.07 1275 -2.659 0.4835
## Pore - Rusp -52.229 11.63 1291 -4.491 0.0014
## Pore - Soca -62.344 7.07 1275 -8.817 <.0001
## Pore - Taof 6.831 8.74 1284 0.782 1.0000
## Pore - Trsp -35.046 8.68 1276 -4.038 0.0096
## Pore - Vear 10.528 13.99 1284 0.753 1.0000
## Posp - Rusp -33.427 11.25 1291 -2.971 0.2656
## Posp - Soca -43.542 6.39 1265 -6.810 <.0001
## Posp - Taof 25.633 8.21 1284 3.120 0.1867
## Posp - Trsp -16.244 8.15 1275 -1.994 0.9181
## Posp - Vear 29.330 13.68 1286 2.144 0.8516
## Rusp - Soca -10.115 11.25 1291 -0.899 1.0000
## Rusp - Taof 59.059 12.32 1290 4.795 0.0004
## Rusp - Trsp 17.183 12.32 1291 1.395 0.9986
## Rusp - Vear 62.757 16.50 1290 3.802 0.0231
## Soca - Taof 69.174 8.21 1284 8.421 <.0001
## Soca - Trsp 27.297 8.15 1275 3.351 0.1002
## Soca - Vear 72.872 13.68 1286 5.327 <.0001
## Taof - Trsp -41.877 9.63 1282 -4.351 0.0027
## Taof - Vear 3.698 14.60 1286 0.253 1.0000
## Trsp - Vear 45.575 14.53 1281 3.138 0.1787
##
## Results are averaged over the levels of: state
## Degrees-of-freedom method: kenward-roger
## P value adjustment: tukey method for comparing a family of 21 estimates
# using model 7a for overall greenup model #

# Take a look at the estimates for each fixed effect. These
# are the estimates from summary(mod7a). You'll see that
# species vary a lot - and many of them are different from
# zero (meaning their half cover date is significantly
# different from zero).
plot_model(mod7a, sort.est = TRUE)

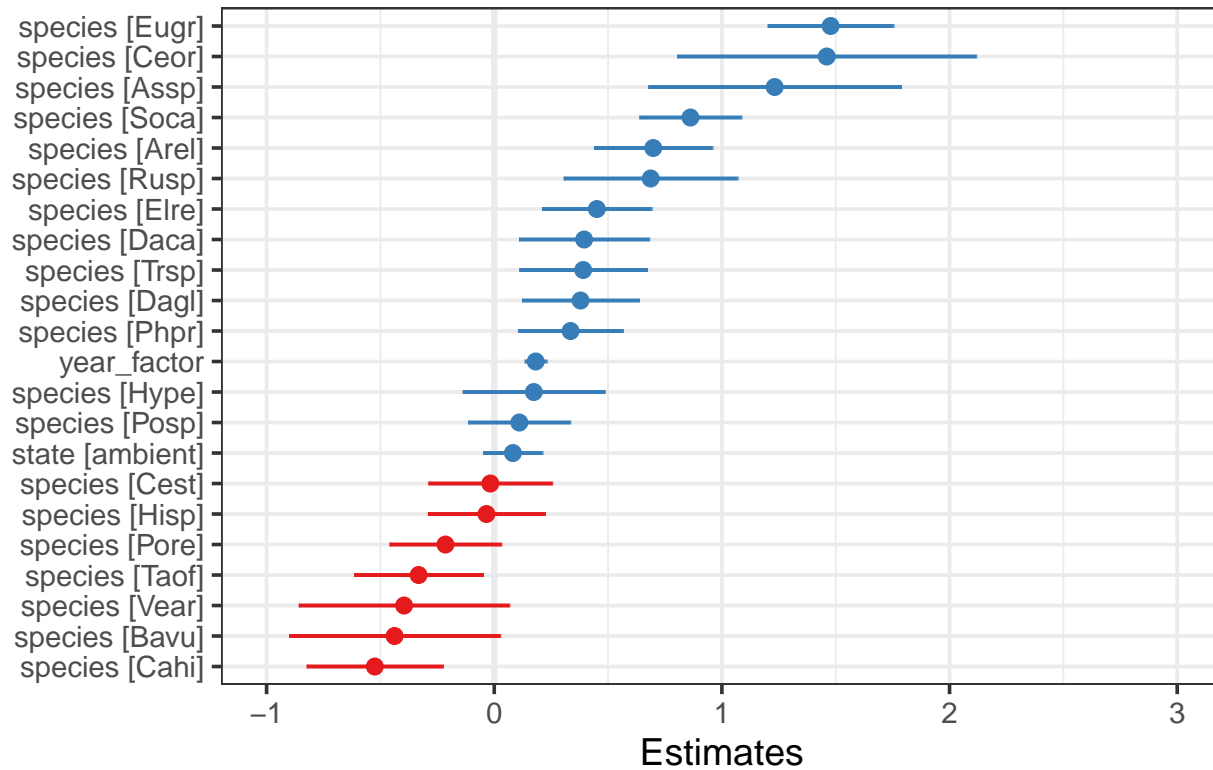
```

spp_half_cover_date



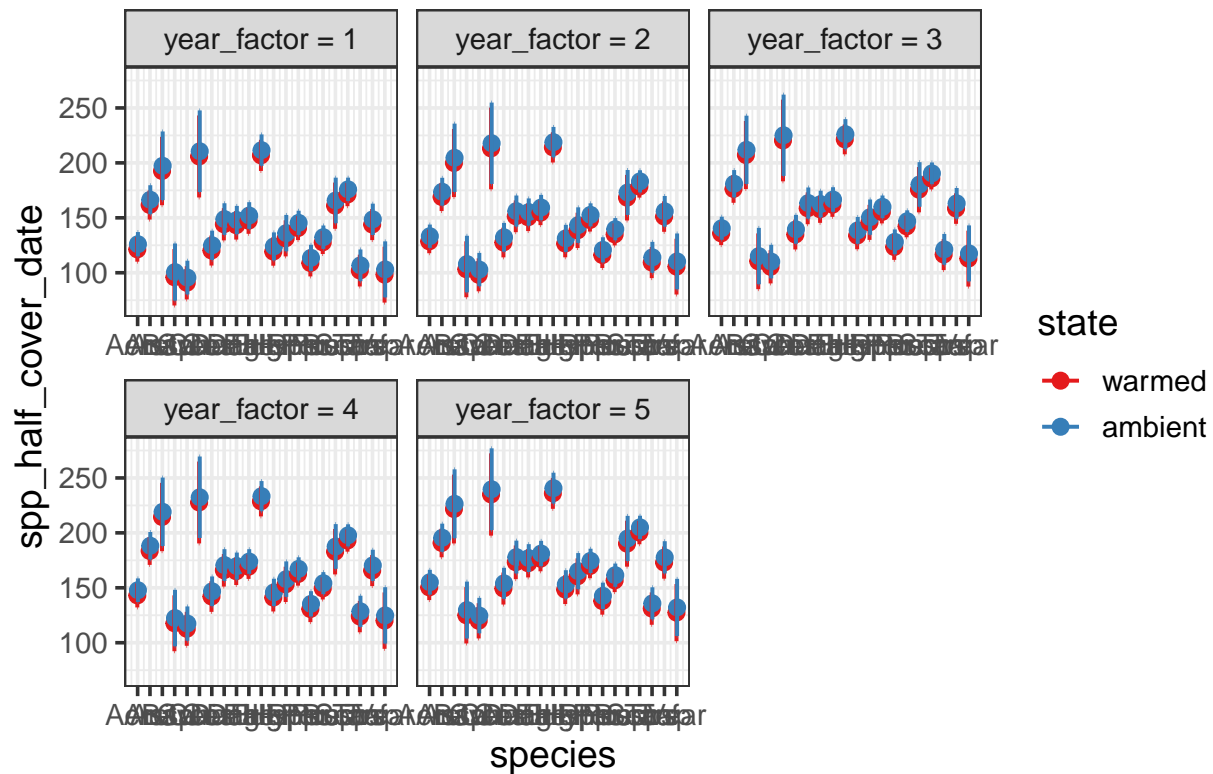
```
# if you want to standardize the estimates:
plot_model(mod7a, sort.est = TRUE, type = "std")
```

spp_half_cover_date



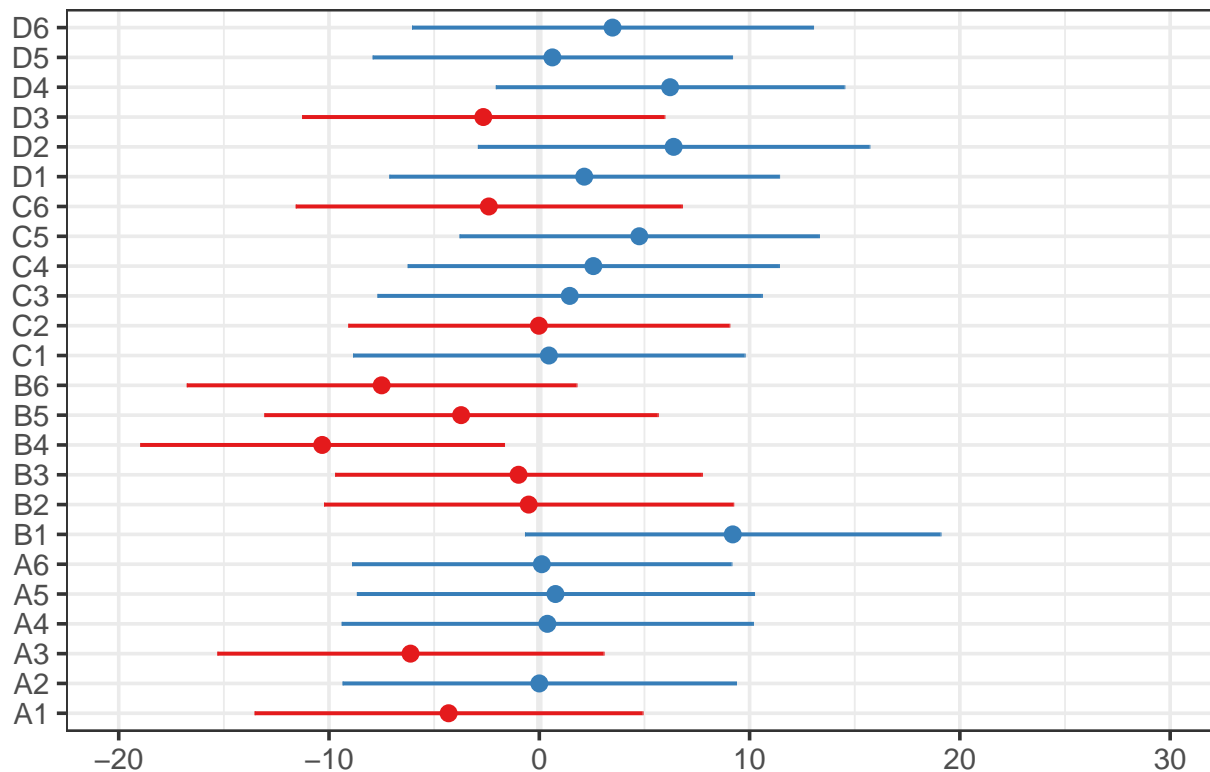
```
# these are the fixed predicted values: - note this is a new
# plot
plot_model(mod7a, type = "pred", terms = c("species", "state",
      "year_factor"))
```

Predicted values of spp_half_cover_date



```
# these are the random effects estimates
plot_model(mod7a, type = "re")
```

Random effects



```
# including native vs. exotic
green_kbs <- within(green_kbs, origin <- relevel(factor(origin),
  ref = "Native")) # releveling so native is the reference
mod8 <- lmer(spp_half_cover_date ~ state * origin + (1 + year_factor |
  plot), green_kbs, REML = FALSE)
```

```
## boundary (singular) fit: see ?isSingular
```

```
## Warning: Model failed to converge with 1 negative eigenvalue: -4.3e+02
```

```
mod9 <- lmer(spp_half_cover_date ~ state + origin + (1 + year_factor |
  plot), green_kbs, REML = FALSE)
```

```
## boundary (singular) fit: see ?isSingular
```

```
## Warning: Model failed to converge with 1 negative eigenvalue: -4.4e+02
```

```
mod9a <- lmer(spp_half_cover_date ~ state + origin + year_factor +
  (1 | plot), green_kbs, REML = FALSE)
anova(mod8, mod9) # model 9 is a better fit to data
```

```
## Data: green_kbs
```

```
## Models:
```

```
## mod9: spp_half_cover_date ~ state + origin + (1 + year_factor | plot)
```

```
## mod8: spp_half_cover_date ~ state * origin + (1 + year_factor | plot)
```

```
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
```

```
## mod9      9 13805 13851 -6893.3   13787
```

```
## mod8     12 13806 13868 -6891.1   13782 4.4194  3    0.2196
```

```
anova(mod9, mod9a) # mod 9a?
```

```

## Data: green_kbs
## Models:
## mod9a: spp_half_cover_date ~ state + origin + year_factor + (1 | plot)
## mod9: spp_half_cover_date ~ state + origin + (1 + year_factor | plot)
##      npar   AIC   BIC logLik deviance Chisq Df Pr(>Chisq)
## mod9a    8 13768 13810 -6876.2   13752
## mod9     9 13805 13851 -6893.3   13787    0  1          1
summary(mod9a)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state + origin + year_factor + (1 | plot)
## Data: green_kbs
##
##      AIC      BIC   logLik deviance df.resid
## 13768.3 13809.5 -6876.2 13752.3    1260
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.0457 -0.7656 -0.3324  0.8431  2.4332
##
## Random effects:
## Groups Name Variance Std.Dev.
## plot (Intercept) 51.94 7.207
## Residual 2967.56 54.475
## Number of obs: 1268, groups: plot, 24
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept) 147.418      5.212 203.427 28.284 < 2e-16 ***
## stateambient  4.139      4.267 24.778 0.970 0.341
## origin -32.633      5.951 1248.129 -5.484 5.03e-08 ***
## originBoth -21.481      5.436 1259.517 -3.952 8.20e-05 ***
## originExotic -29.216      3.859 1255.517 -7.570 7.19e-14 ***
## year_factor  9.232      1.059 1256.864 8.714 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sttmbn origin orgnBt orgnEx
## stateambint -0.407
## origin -0.342 0.006
## originBoth -0.392 -0.013 0.329
## originExotc -0.539 -0.010 0.464 0.511
## year_factor -0.595 -0.015 -0.008 0.037 0.022
anova(mod9)

## Type III Analysis of Variance Table with Satterthwaite's method
##      Sum Sq Mean Sq NumDF DenDF F value Pr(>F)
## state 231 231 1 169.36 0.0783 0.78
## origin 186320 62107 3 1231.80 21.0643 2.61e-13 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```



```
emmeans(mod9a, list(pairwise ~ state + origin), adjust = "tukey")
```

```
## $`emmeans of state, origin`
## state origin emmean SE df lower.CL upper.CL
## warmed Native 174 4.29 93.4 166 183
## ambient Native 178 4.25 92.2 170 187
## warmed 142 5.69 274.0 130 153
## ambient 146 5.69 278.5 135 157
## warmed Both 153 5.20 189.3 142 163
## ambient Both 157 5.11 180.9 147 167
## warmed Exotic 145 3.48 39.3 138 152
## ambient Exotic 149 3.39 35.5 142 156
##
## Degrees-of-freedom method: kenward-roger
## Confidence level used: 0.95
##
## $`pairwise differences of state, origin`
## 1 estimate SE df t.ratio p.value
## warmed Native - ambient Native -4.139 4.46 25.9 -0.929 0.9802
## warmed Native - warmed 32.633 5.96 1251.5 5.474 <.0001
## warmed Native - ambient 28.494 7.47 209.2 3.817 0.0043
## warmed Native - warmed Both 21.481 5.45 1263.6 3.941 0.0022
## warmed Native - ambient Both 17.341 7.00 161.4 2.478 0.2126
## warmed Native - warmed Exotic 29.216 3.87 1259.3 7.553 <.0001
## warmed Native - ambient Exotic 25.077 5.87 81.8 4.269 0.0013
## ambient Native - warmed 36.772 7.42 205.4 4.955 <.0001
## ambient Native - ambient 32.633 5.96 1251.5 5.474 <.0001
## ambient Native - warmed Both 25.620 7.08 169.2 3.618 0.0092
## ambient Native - ambient Both 21.481 5.45 1263.6 3.941 0.0022
## ambient Native - warmed Exotic 33.355 5.93 85.8 5.627 <.0001
## ambient Native - ambient Exotic 29.216 3.87 1259.3 7.553 <.0001
## warmed - ambient -4.139 4.46 25.9 -0.929 0.9802
## warmed - warmed Both -11.152 6.63 1261.0 -1.683 0.6985
## warmed - ambient Both -15.291 7.93 254.8 -1.929 0.5326
## warmed - warmed Exotic -3.417 5.39 1253.2 -0.633 0.9984
## warmed - ambient Exotic -7.556 6.95 159.8 -1.087 0.9587
## ambient - warmed Both -7.013 8.04 267.6 -0.872 0.9883
## ambient - ambient Both -11.152 6.63 1261.0 -1.683 0.6985
## ambient - warmed Exotic 0.722 7.04 168.9 0.103 1.0000
## ambient - ambient Exotic -3.417 5.39 1253.2 -0.633 0.9984
## warmed Both - ambient Both -4.139 4.46 25.9 -0.929 0.9802
## warmed Both - warmed Exotic 7.735 4.81 1262.8 1.609 0.7451
## warmed Both - ambient Exotic 3.596 6.58 127.5 0.547 0.9994
## ambient Both - warmed Exotic 11.874 6.54 125.3 1.817 0.6100
## ambient Both - ambient Exotic 7.735 4.81 1262.8 1.609 0.7451
## warmed Exotic - ambient Exotic -4.139 4.46 25.9 -0.929 0.9802
##
## Degrees-of-freedom method: kenward-roger
## P value adjustment: tukey method for comparing a family of 8 estimates
# including growth form - first with interaction term
green_kbs <- within(green_kbs, growth_habit <- relevel(factor(growth_habit),
  ref = "Forb")) # releveling so forb is the reference
mod10 <- lmer(spp_half_cover_date ~ state * growth_habit + (1 +
```

```

    year_factor | plot), green_kbs, REML = FALSE)

## boundary (singular) fit: see ?isSingular
mod11 <- lmer(spp_half_cover_date ~ state + growth_habit + (1 +
  year_factor | plot), green_kbs, REML = FALSE)

## boundary (singular) fit: see ?isSingular
## Warning: Model failed to converge with 1 negative eigenvalue: -4.8e+02
mod11a <- lmer(spp_half_cover_date ~ state + growth_habit + year_factor +
  (1 | plot), green_kbs, REML = FALSE)
anova(mod10, mod11) # model 11 is a better fit to data

## Data: green_kbs
## Models:
## mod11: spp_half_cover_date ~ state + growth_habit + (1 + year_factor |
## mod11:      plot)
## mod10: spp_half_cover_date ~ state * growth_habit + (1 + year_factor |
## mod10:      plot)
##      npar   AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## mod11     9 13851 13898 -6916.6    13833
## mod10    12 13840 13902 -6908.3    13816 16.764   3 0.0007904 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

anova(mod11, mod11a)

## Data: green_kbs
## Models:
## mod11a: spp_half_cover_date ~ state + growth_habit + year_factor + (1 |
## mod11a:      plot)
## mod11: spp_half_cover_date ~ state + growth_habit + (1 + year_factor |
## mod11:      plot)
##      npar   AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## mod11a     8 13815 13856 -6899.4    13799
## mod11     9 13851 13898 -6916.6    13833      0  1          1

summary(mod11a)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state + growth_habit + year_factor + (1 |
##      plot)
##      Data: green_kbs
##
##      AIC      BIC   logLik deviance df.resid
## 13814.9 13856.0 -6899.4 13798.9      1260
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -1.8570 -0.8206 -0.3565  0.9522  2.3273
##
## Random effects:
##      Groups      Name      Variance Std.Dev.
##      plot      (Intercept)  52.34    7.235

```

```

## Residual          3079.28  55.491
## Number of obs: 1268, groups: plot, 24
##
## Fixed effects:
##               Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)    123.0075    4.5507  118.9402  27.031 < 2e-16 ***
## stateambient     3.9039    4.3187   24.9788   0.904 0.374654
## growth_habit    -0.1793    5.0509 1255.9654  -0.035 0.971689
## growth_habitGraminoid  6.4752    3.4201 1263.2381   1.893 0.058551 .
## growth_habitVine  71.0413   21.3388 1263.6544   3.329 0.000896 ***
## year_factor      9.1144    1.0870 1258.6830   8.385 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##      (Intr) sttmbn grwth_ grwt_G grwt_V
## stateambint -0.481
## growth_habt -0.198 -0.008
## grwth_hbtGr -0.239  0.026  0.275
## grwth_hbtVn -0.044 -0.019  0.049  0.066
## year_factor -0.648 -0.018 -0.006 -0.124  0.003
anova(mod11a)

## Type III Analysis of Variance Table with Satterthwaite's method
##               Sum Sq Mean Sq NumDF   DenDF F value    Pr(>F)
## state           2516     2516     1    24.98  0.8171  0.374654
## growth_habit    44231    14744     3 1260.85  4.7881  0.002538 **
## year_factor    216501   216501     1 1258.68 70.3090 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
emmeans(mod11a, list(pairwise ~ state + growth_habit), adjust = "tukey")

## $`emmeans of state, growth_habit`
##   state growth_habit emmean    SE      df lower.CL upper.CL
## warmed Forb          149  3.60   42.4      142      157
## ambient Forb          153  3.49   38.1      146      160
## warmed              149  5.41  211.8      139      160
## ambient              153  5.30  200.3      143      164
## warmed Graminoid      156  3.81   55.1      148      164
## ambient Graminoid      160  3.81   55.1      152      167
## warmed Vine           220 21.55 1233.0      178      263
## ambient Vine           224 21.45 1233.0      182      266
##
## Degrees-of-freedom method: kenward-roger
## Confidence level used: 0.95
##
## $`pairwise differences of state, growth_habit`
##   1 estimate    SE      df t.ratio p.value
## warmed Forb - ambient Forb      -3.904  4.51   25.9 -0.865  0.9868
## warmed Forb - warmed              0.179  5.06 1259.7  0.035  1.0000
## warmed Forb - ambient              -3.725  6.76  135.5 -0.551  0.9993
## warmed Forb - warmed Graminoid     -6.475  3.43 1267.8 -1.888  0.5596
## warmed Forb - ambient Graminoid    -10.379  5.74   69.7 -1.809  0.6162

```

```
## warmed Forb - warmed Vine          -71.041 21.44 1271.2 -3.313 0.0213
## warmed Forb - ambient Vine          -74.945 21.83 1154.7 -3.433 0.0143
## ambient Forb - warmed               4.083  6.80  140.6  0.600 0.9988
## ambient Forb - ambient              0.179  5.06 1259.7  0.035 1.0000
## ambient Forb - warmed Graminoid     -2.571  5.60  64.7 -0.459 0.9998
## ambient Forb - ambient Graminoid    -6.475  3.43 1267.8 -1.888 0.5596
## ambient Forb - warmed Vine          -67.137 22.00 1157.7 -3.052 0.0477
## ambient Forb - ambient Vine         -71.041 21.44 1271.2 -3.313 0.0213
## warmed - ambient                   -3.904  4.51  25.9 -0.865 0.9868
## warmed - warmed Graminoid          -6.654  5.28 1264.2 -1.261 0.9130
## warmed - ambient Graminoid         -10.558  7.02  155.4 -1.503 0.8047
## warmed - warmed Vine               -71.221 21.79 1273.3 -3.269 0.0245
## warmed - ambient Vine              -75.125 22.17 1174.9 -3.388 0.0166
## ambient - warmed Graminoid         -2.751  6.86  144.4 -0.401 0.9999
## ambient - ambient Graminoid        -6.654  5.28 1264.2 -1.261 0.9130
## ambient - warmed Vine              -67.317 22.32 1176.3 -3.016 0.0531
## ambient - ambient Vine             -71.221 21.79 1273.3 -3.269 0.0245
## warmed Graminoid - ambient Graminoid -3.904  4.51  25.9 -0.865 0.9868
## warmed Graminoid - warmed Vine     -64.566 21.49 1271.7 -3.004 0.0548
## warmed Graminoid - ambient Vine    -68.470 21.86 1159.0 -3.132 0.0376
## ambient Graminoid - warmed Vine    -60.662 22.06 1160.7 -2.750 0.1091
## ambient Graminoid - ambient Vine    -64.566 21.49 1271.7 -3.004 0.0548
## warmed Vine - ambient Vine         -3.904  4.51  25.9 -0.865 0.9868
##
```

```
## Degrees-of-freedom method: kenward-roger
```

```
## P value adjustment: tukey method for comparing a family of 8 estimates
```

```
# You could now run some post hoc tests on these (see:
```

```
# https://stats.stackexchange.com/questions/169543/output-of-fixed-effects-summary-in-lmertest-in-r-and
```

```
# Here are some other options for plotting these plots above:
```

```
# https://stackoverflow.com/questions/31075407/plot-mixed-effects-model-in-ggplot
```

```
# Here's another approach:
```

```
# https://stats.stackexchange.com/questions/98958/plots-to-illustrate-results-of-linear-mixed-effect-mo
```

```
# Not quite working yet: newdat <-
```

```
# expand.grid(state=unique(green_kbs$state),
```

```
# year=c(min(green_kbs$year), max(green_kbs$year)),
```

```
# insecticide=unique(green_kbs$insecticide)) p <-
```

```
# ggplot(green_kbs, aes(x=year, y=spp_half_cover_date,
```

```
# colour=state, shape=insecticide)) + geom_point(size=3) +
```

```
# geom_line(aes(y=predict(mod5), group=species,
```

```
# size='species')) + geom_line(data=newdat,
```

```
# aes(y=predict(mod5, level=0, newdata=newdat),
```

```
# size='Population')) + scale_size_manual(name='Predictions',
```

```
# values=c('species'=0.5, 'Population'=3)) +
```

```
# #facet_wrap(~insecticide) + theme_bw(base_size=22) print(p)
```

```
# New version of our model incorporating interaction term and  
# species within year so that there is a separate intercept  
# and slope for each species. The issue here is that there  
# are some species that are not found each year. Easiest to  
# remove those from another version of this dataframe before  
# running below. Otherwise, it's not a balanced design.
```

```

# updated mod4
mod12 <- lmer(spp_half_cover_date ~ state * year + (1 + year |
  species), green_kbs)

## Warning: Some predictor variables are on very different scales: consider
## rescaling

## boundary (singular) fit: see ?isSingular

## Warning: Some predictor variables are on very different scales: consider
## rescaling

# So another version of this model would include the
# interaction but not include the nesting (and thus would
# assume that species aren't observed ea yr) updated mod5
mod13 <- lmer(spp_half_cover_date ~ state * year + (1 | species),
  green_kbs)

## Warning: Some predictor variables are on very different scales: consider
## rescaling

## Warning: Some predictor variables are on very different scales: consider
## rescaling

```

KBS Plot-level Mixed Effects Models:

```

mod1p <- lmer(plot_half_cover_date ~ state + (1 | plot), green_kbsp,
  REML = FALSE)

## boundary (singular) fit: see ?isSingular

mod2p <- lmer(plot_half_cover_date ~ state + year_factor + (1 |
  plot), green_kbsp, REML = FALSE)

## boundary (singular) fit: see ?isSingular

mod3p <- lmer(plot_half_cover_date ~ state * year_factor + (1 |
  plot), green_kbsp, REML = FALSE)

## boundary (singular) fit: see ?isSingular

anova(mod1p, mod2p, mod3p) #mod2p

## Data: green_kbsp
## Models:
## mod1p: plot_half_cover_date ~ state + (1 | plot)
## mod2p: plot_half_cover_date ~ state + year_factor + (1 | plot)
## mod3p: plot_half_cover_date ~ state * year_factor + (1 | plot)
##      npar    AIC    BIC logLik deviance   Chisq Df Pr(>Chisq)
## mod1p     4 1308.6 1319.8 -650.32   1300.6
## mod2p     5 1289.2 1303.1 -639.60   1279.2 21.4246   1 3.68e-06 ***
## mod3p     6 1290.7 1307.4 -639.35   1278.7  0.5025   1  0.4784
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

summary(mod2p)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's

```

```
## method [lmerModLmerTest]
## Formula: plot_half_cover_date ~ state + year_factor + (1 | plot)
## Data: green_kbsp
##
##      AIC      BIC   logLik deviance df.resid
## 1289.2   1303.1   -639.6   1279.2     115
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -1.9138 -0.8568  0.1682  1.0336  1.8931
##
## Random effects:
## Groups Name Variance Std.Dev.
## plot (Intercept) 0 0.00
## Residual 2495 49.95
## Number of obs: 120, groups: plot, 24
##
## Fixed effects:
## Estimate Std. Error df t value Pr(>|t|)
## (Intercept) 121.817 11.626 120.000 10.478 < 2e-16 ***
## stateambient 14.700 9.120 120.000 1.612 0.11
## year_factor 15.617 3.224 120.000 4.843 3.85e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
## (Intr) sttmbn
## stateambint -0.392
## year_factor -0.832 0.000
## convergence code: 0
## boundary (singular) fit: see ?isSingular

mod2p.2 <- lmer(plot_half_cover_date ~ state + year_factor +
  insecticide + (1 | plot), green_kbsp, REML = FALSE)

## boundary (singular) fit: see ?isSingular

anova(mod2p, mod2p.2) #mod2p

## Data: green_kbsp
## Models:
## mod2p: plot_half_cover_date ~ state + year_factor + (1 | plot)
## mod2p.2: plot_half_cover_date ~ state + year_factor + insecticide + (1 |
## mod2p.2: plot)
## npar AIC BIC logLik deviance Chisq Df Pr(>Chisq)
## mod2p 5 1289.2 1303.1 -639.60 1279.2
## mod2p.2 6 1291.0 1307.8 -639.52 1279.0 0.1647 1 0.6849
```

UMBS Mixed Effects Models

```
# umod4 (and umod6) are pretty complex in terms of
# interpretation (they actually don't have many parameters
# though). We could consider an alternative umodel that's
# simpler to understand and also one that provides more
```

```

# insight about the species. That would be something like
# this:
umod7 <- lmer(spp_half_cover_date ~ state + species + (1 + year_factor |
  plot), green_umbs, REML = FALSE)

## boundary (singular) fit: see ?isSingular

umod7a <- lmer(spp_half_cover_date ~ state + species + year_factor +
  (1 | plot), green_umbs, REML = FALSE)
umod7b <- lmer(spp_half_cover_date ~ state * year_factor + species +
  (1 | plot), green_umbs, REML = FALSE)
umod7c <- lmer(spp_half_cover_date ~ state + species + year_factor +
  insecticide + (1 | plot), green_umbs, REML = FALSE)
# anova(umod6, umod7) # umodel 7 is a better fit to data
anova(umod7, umod7a) #umod 7a

## Data: green_umbs
## Models:
## umod7a: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## umod7: spp_half_cover_date ~ state + species + (1 + year_factor | plot)
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## umod7a   19 7497.7 7585.7 -3729.8   7459.7
## umod7    20 7509.8 7602.4 -3734.9   7469.8      0  1          1
anova(umod7a, umod7b) #umod 7a

## Data: green_umbs
## Models:
## umod7a: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## umod7b: spp_half_cover_date ~ state * year_factor + species + (1 | plot)
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## umod7a   19 7497.7 7585.7 -3729.8   7459.7
## umod7b   20 7499.6 7592.2 -3729.8   7459.6 0.0588  1    0.8085
anova(umod7a, umod7c) #umod 7a

## Data: green_umbs
## Models:
## umod7a: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## umod7c: spp_half_cover_date ~ state + species + year_factor + insecticide +
## umod7c: (1 | plot)
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## umod7a   19 7497.7 7585.7 -3729.8   7459.7
## umod7c   20 7499.7 7592.3 -3729.8   7459.7 0.0208  1    0.8854
summary(umod7a)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state + species + year_factor + (1 | plot)
## Data: green_umbs
##
##      AIC      BIC  logLik deviance df.resid
##  7497.7   7585.7 -3729.8   7459.7      739
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max

```

```

## -2.5835 -0.6079 -0.2494  0.4752  3.1253
##
## Random effects:
##   Groups   Name      Variance Std.Dev.
## plot      (Intercept)    9.4    3.066
## Residual                1092.1   33.046
## Number of obs: 758, groups: plot, 24
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)  131.17082   13.13233  698.99506    9.988 < 2e-16 ***
## stateambient  -0.51847    2.73549   18.95368   -0.190  0.851689
## speciesAnsp    0.08818   15.69454  746.26222    0.006  0.995519
## speciesApan   51.11999   17.76623  756.25039    2.877  0.004123 **
## speciesAssp   29.71719   14.22495  719.48551    2.089  0.037050 *
## speciesCape   20.04356   13.15689  739.06765    1.523  0.128080
## speciesCest    4.47430   12.96925  743.11707    0.345  0.730197
## speciesDasp    5.23558   13.02731  743.48921    0.402  0.687878
## speciesFrve   10.38974   14.50633  724.31057    0.716  0.474086
## speciesHisp   31.07700   14.95040  757.06388    2.079  0.037984 *
## speciesHype    9.44938   13.35464  750.28371    0.708  0.479430
## speciesPosp   21.66218   12.98183  744.17278    1.669  0.095607 .
## speciesPtaq   40.75255   13.15100  749.25912    3.099  0.002016 **
## speciesRuac    5.98322   13.08141  748.10341    0.457  0.647528
## speciesSosp   23.83121   14.75435  753.60933    1.615  0.106686
## speciesSyla   56.42041   16.75388  757.29228    3.368  0.000797 ***
## year_factor    3.89359    0.86526  737.50628    4.500  7.9e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation matrix not shown by default, as p = 17 > 12.
## Use print(x, correlation=TRUE) or
##      vcov(x)          if you need it
summary(umod7b)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state * year_factor + species + (1 | plot)
## Data: green_umbs
##
##      AIC      BIC   logLik deviance df.resid
## 7499.6   7592.2  -3729.8   7459.6      738
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.5719 -0.6104 -0.2469  0.4666  3.1257
##
## Random effects:
##   Groups   Name      Variance Std.Dev.
## plot      (Intercept)    9.434   3.072
## Residual                1091.945  33.045
## Number of obs: 758, groups: plot, 24
##

```



```

## Fixed effects:
##
##           Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)    131.8256    13.4082  703.4460   9.832 < 2e-16 ***
## stateambient     -1.7775     5.8711  307.8369  -0.303 0.762288
## year_factor       3.6822     1.2283  735.0831   2.998 0.002811 **
## speciesAnsp       0.1050    15.6942  746.3583   0.007 0.994665
## speciesApan      51.0666    17.7671  756.2431   2.874 0.004164 **
## speciesAssp      29.6707    14.2261  719.5153   2.086 0.037361 *
## speciesCape      20.0150    13.1571  739.0296   1.521 0.128628
## speciesCest       4.4662    12.9689  743.1594   0.344 0.730661
## speciesDasp       5.2264    13.0270  743.5325   0.401 0.688392
## speciesFrve      10.3662    14.5064  724.3909   0.715 0.475088
## speciesHisp      31.0973    14.9501  757.0795   2.080 0.037854 *
## speciesHype       9.4365    13.3543  750.3046   0.707 0.480019
## speciesPosp      21.6472    12.9816  744.1937   1.668 0.095831 .
## speciesPtaq      40.7584    13.1506  749.3119   3.099 0.002012 **
## speciesRuac       5.9792    13.0810  748.1476   0.457 0.647738
## speciesSosp      23.8092    14.7542  753.6366   1.614 0.107005
## speciesSyla      56.4880    16.7555  757.2835   3.371 0.000786 ***
## stateambient:year_factor  0.4114     1.6970  737.1355   0.242 0.808507
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##
## Correlation matrix not shown by default, as p = 18 > 12.
## Use print(x, correlation=TRUE) or
##     vcov(x)           if you need it
summary(umod7c)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state + species + year_factor + insecticide +
## (1 | plot)
## Data: green_umbs
##
##      AIC      BIC  logLik deviance df.resid
##  7499.7   7592.3 -3729.8   7459.7     738
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.5860 -0.6093 -0.2500  0.4756  3.1306
##
## Random effects:
## Groups   Name      Variance Std.Dev.
## plot     (Intercept)    9.418  3.069
## Residual                1092.013 33.046
## Number of obs: 758, groups: plot, 24
##
## Fixed effects:
##
##           Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)    130.99606    13.18765  668.26073   9.933 < 2e-16 ***
## stateambient     -0.51887     2.73602  19.01321  -0.190 0.851599
## speciesAnsp       0.05593    15.69604  746.83227   0.004 0.997158
## speciesApan      50.95022    17.80572  757.89874   2.861 0.004333 **

```

```
## speciesAssp      29.77413    14.23004 720.04210    2.092 0.036757 *
## speciesCape      20.02311    13.15757 739.59718    1.522 0.128488
## speciesCest       4.45599    12.96981 743.56984    0.344 0.731270
## speciesDasp       5.21894    13.02774 743.91429    0.401 0.688828
## speciesFrve      10.36379    14.50747 725.71680    0.714 0.475224
## speciesHisp      30.99767    14.96048 757.70548    2.072 0.038606 *
## speciesHype       9.41831    13.35633 750.89419    0.705 0.480931
## speciesPosp      21.64688    12.98219 744.53702    1.667 0.095850 .
## speciesPtaq      40.72289    13.15258 749.89255    3.096 0.002033 **
## speciesRuac       5.95334    13.08303 748.78234    0.455 0.649210
## speciesSosp      23.73807    14.76866 755.35081    1.607 0.108401
## speciesSyla      56.38543    16.75547 757.42449    3.365 0.000803 ***
## year_factor       3.89196     0.86531 737.13965    4.498 7.98e-06 ***
## insecticideno_insects 0.39630     2.75007 19.48138    0.144 0.886897
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
## Correlation matrix not shown by default, as p = 18 > 12.
## Use print(x, correlation=TRUE) or
##     vcov(x)           if you need it
```

```
anova(umod7a) # investigates whether at least one of the levels within each factor is significantly different
```

```
## Type III Analysis of Variance Table with Satterthwaite's method
##           Sum Sq Mean Sq NumDF   DenDF F value    Pr(>F)
## state           39    39.2      1   18.95  0.0359    0.8517
## species      129594  9256.7     14  743.23  8.4764 < 2.2e-16 ***
## year_factor   22113 22113.4      1  737.51 20.2493 7.896e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
# Yes, at least one of the species is different (they do not
# all have the same half cover dates).
```

```
emmeans(umod7a, list(pairwise ~ year_factor), adjust = "tukey")
```

```
## $`emmeans of year_factor`
##   year_factor emmean   SE    df lower.CL upper.CL
##           3.07    164 2.14 95.8      159      168
##
## Results are averaged over the levels of: state, species
## Degrees-of-freedom method: kenward-roger
## Confidence level used: 0.95
##
```

```
## $` of year_factor`
## 1      estimate SE df z.ratio p.value
## (nothing) nonEst NA NA NA      NA
##
## Results are averaged over the levels of: state, species
## Degrees-of-freedom method: kenward-roger
```

```
emmeans(umod7a, list(pairwise ~ species), adjust = "tukey")
```

```
## $`emmeans of species`
##   species emmean   SE df lower.CL upper.CL
## Amla      143 12.86 756      118      168
## Ansp      143  9.40 774      124      161
```

```

## Apan      194 12.86 764      169      219
## Assp      173  6.72 696      159      186
## Cape      163  3.91 574      155      171
## Cest      147  3.15 446      141      154
## Dasp      148  3.36 493      141      155
## Frve      153  7.31 680      139      168
## Hisp      174  8.20 771      158      190
## Hype      152  4.62 647      143      161
## Posp      165  3.22 466      158      171
## Ptaq      184  3.89 554      176      191
## Ruac      149  3.64 521      142      156
## Sosp      167  7.87 694      151      182
## Syla      199 11.39 730      177      222
##
## Results are averaged over the levels of: state
## Degrees-of-freedom method: kenward-roger
## Confidence level used: 0.95
##
## $`pairwise differences of species`
## 1      estimate      SE df t.ratio p.value
## Amla - Ansp -0.0882 15.96 772 -0.006 1.0000
## Amla - Apan -51.1200 18.04 776 -2.834 0.2311
## Amla - Assp -29.7172 14.50 756 -2.049 0.7680
## Amla - Cape -20.0436 13.39 768 -1.497 0.9766
## Amla - Cest -4.4743 13.20 770 -0.339 1.0000
## Amla - Dasp -5.2356 13.26 770 -0.395 1.0000
## Amla - Frve -10.3897 14.78 759 -0.703 1.0000
## Amla - Hisp -31.0770 15.18 776 -2.048 0.7690
## Amla - Hype -9.4494 13.58 774 -0.696 1.0000
## Amla - Posp -21.6622 13.21 771 -1.640 0.9493
## Amla - Ptaq -40.7525 13.37 773 -3.048 0.1379
## Amla - Ruac -5.9832 13.30 773 -0.450 1.0000
## Amla - Sosp -23.8312 14.99 775 -1.590 0.9607
## Amla - Syla -56.4204 17.01 776 -3.317 0.0649
## Ansp - Apan -51.0318 15.91 775 -3.208 0.0893
## Ansp - Assp -29.6290 11.49 775 -2.579 0.3855
## Ansp - Cape -19.9554 10.15 773 -1.966 0.8177
## Ansp - Cest -4.3861  9.84 771 -0.446 1.0000
## Ansp - Dasp -5.1474  9.90 770 -0.520 1.0000
## Ansp - Frve -10.3016 11.80 775 -0.873 0.9999
## Ansp - Hisp -30.9888 12.39 770 -2.501 0.4407
## Ansp - Hype -9.3612 10.44 770 -0.896 0.9999
## Ansp - Posp -21.5740  9.88 771 -2.183 0.6769
## Ansp - Ptaq -40.6644 10.12 774 -4.019 0.0057
## Ansp - Ruac -5.8950 10.00 767 -0.589 1.0000
## Ansp - Sosp -23.7430 12.20 772 -1.947 0.8283
## Ansp - Syla -56.3322 14.76 766 -3.817 0.0122
## Apan - Assp  21.4028 14.48 764  1.478 0.9790
## Apan - Cape  31.0764 13.38 772  2.323 0.5736
## Apan - Cest  46.6457 13.19 774  3.536 0.0326
## Apan - Dasp  45.8844 13.25 774  3.463 0.0413
## Apan - Frve  40.7303 14.80 759  2.752 0.2761
## Apan - Hisp  20.0430 15.20 776  1.319 0.9928
## Apan - Hype  41.6706 13.56 775  3.072 0.1294

```

##	Apan - Posp	29.4578	13.21	773	2.230	0.6429
##	Apan - Ptaq	10.3674	13.36	776	0.776	1.0000
##	Apan - Ruac	45.1368	13.30	775	3.394	0.0513
##	Apan - Sosp	27.2888	15.06	772	1.812	0.8929
##	Apan - Syla	-5.3004	17.13	768	-0.309	1.0000
##	Assp - Cape	9.6736	7.70	767	1.256	0.9956
##	Assp - Cest	25.2429	7.34	773	3.440	0.0444
##	Assp - Dasp	24.4816	7.42	775	3.298	0.0687
##	Assp - Frve	19.3274	9.82	768	1.967	0.8170
##	Assp - Hisp	-1.3598	10.51	776	-0.129	1.0000
##	Assp - Hype	20.2678	8.07	775	2.512	0.4328
##	Assp - Posp	8.0550	7.37	772	1.093	0.9990
##	Assp - Ptaq	-11.0354	7.68	775	-1.438	0.9837
##	Assp - Ruac	23.7340	7.55	773	3.142	0.1074
##	Assp - Sosp	5.8860	10.29	763	0.572	1.0000
##	Assp - Syla	-26.7032	13.12	768	-2.035	0.7769
##	Cape - Cest	15.5693	4.90	763	3.179	0.0970
##	Cape - Dasp	14.8080	5.05	768	2.933	0.1837
##	Cape - Frve	9.6538	8.22	765	1.175	0.9978
##	Cape - Hisp	-11.0334	9.03	774	-1.222	0.9967
##	Cape - Hype	10.5942	5.94	775	1.782	0.9048
##	Cape - Posp	-1.6186	4.94	762	-0.327	1.0000
##	Cape - Ptaq	-20.7090	5.41	773	-3.830	0.0116
##	Cape - Ruac	14.0603	5.25	776	2.679	0.3196
##	Cape - Sosp	-3.7876	8.74	761	-0.433	1.0000
##	Cape - Syla	-36.3768	12.01	751	-3.030	0.1445
##	Cest - Dasp	-0.7613	4.46	753	-0.171	1.0000
##	Cest - Frve	-5.9154	7.88	762	-0.751	1.0000
##	Cest - Hisp	-26.6027	8.71	771	-3.054	0.1356
##	Cest - Hype	-4.9751	5.49	771	-0.907	0.9999
##	Cest - Posp	-17.1879	4.37	749	-3.937	0.0078
##	Cest - Ptaq	-36.2782	4.88	768	-7.441	<.0001
##	Cest - Ruac	-1.5089	4.68	764	-0.322	1.0000
##	Cest - Sosp	-19.3569	8.40	762	-2.305	0.5872
##	Cest - Syla	-51.9461	11.76	758	-4.416	0.0011
##	Dasp - Frve	-5.1541	7.96	761	-0.647	1.0000
##	Dasp - Hisp	-25.8414	8.78	769	-2.944	0.1789
##	Dasp - Hype	-4.2138	5.61	769	-0.751	1.0000
##	Dasp - Posp	-16.4266	4.52	753	-3.632	0.0236
##	Dasp - Ptaq	-35.5170	5.01	767	-7.088	<.0001
##	Dasp - Ruac	-0.7476	4.82	762	-0.155	1.0000
##	Dasp - Sosp	-18.5956	8.47	761	-2.194	0.6690
##	Dasp - Syla	-51.1848	11.82	759	-4.330	0.0016
##	Frve - Hisp	-20.6873	10.85	774	-1.907	0.8495
##	Frve - Hype	0.9404	8.58	766	0.110	1.0000
##	Frve - Posp	-11.2724	7.91	764	-1.425	0.9849
##	Frve - Ptaq	-30.3628	8.21	757	-3.699	0.0186
##	Frve - Ruac	4.4065	8.08	764	0.545	1.0000
##	Frve - Sosp	-13.4415	10.63	765	-1.265	0.9953
##	Frve - Syla	-46.0307	13.40	772	-3.435	0.0451
##	Hisp - Hype	21.6276	9.33	763	2.319	0.5768
##	Hisp - Posp	9.4148	8.74	771	1.077	0.9991
##	Hisp - Ptaq	-9.6755	9.00	770	-1.075	0.9992
##	Hisp - Ruac	25.0938	8.88	766	2.825	0.2354

```

## Hisp - Sosp 7.2458 11.27 776 0.643 1.0000
## Hisp - Syla -25.3434 13.91 776 -1.821 0.8891
## Hype - Posp -12.2128 5.52 772 -2.211 0.6571
## Hype - Ptaq -31.3032 5.93 773 -5.275 <.0001
## Hype - Ruac 3.4662 5.76 768 0.601 1.0000
## Hype - Sosp -14.3818 9.09 759 -1.583 0.9621
## Hype - Syla -46.9710 12.22 765 -3.843 0.0111
## Posp - Ptaq -19.0904 4.93 769 -3.874 0.0099
## Posp - Ruac 15.6790 4.74 766 3.310 0.0664
## Posp - Sosp -2.1690 8.43 763 -0.257 1.0000
## Posp - Syla -34.7582 11.78 758 -2.950 0.1762
## Ptaq - Ruac 34.7693 5.21 772 6.677 <.0001
## Ptaq - Sosp 16.9213 8.69 764 1.947 0.8281
## Ptaq - Syla -15.6679 11.95 766 -1.311 0.9933
## Ruac - Sosp -17.8480 8.59 764 -2.077 0.7500
## Ruac - Syla -50.4372 11.89 763 -4.243 0.0023
## Sosp - Syla -32.5892 13.61 773 -2.394 0.5200
##
## Results are averaged over the levels of: state
## Degrees-of-freedom method: kenward-roger
## P value adjustment: tukey method for comparing a family of 15 estimates
# including native vs. exotic - first with interaction term
green_umbs <- within(green_umbs, origin <- relevel(factor(origin),
  ref = "Native")) # releveling so native is the reference
umod8 <- lmer(spp_half_cover_date ~ state * origin + (1 + year_factor |
  plot), green_umbs, REML = FALSE)

## boundary (singular) fit: see ?isSingular
umod9 <- lmer(spp_half_cover_date ~ state + origin + (1 + year_factor |
  plot), green_umbs, REML = FALSE)

## boundary (singular) fit: see ?isSingular
umod9a <- lmer(spp_half_cover_date ~ state + origin + year_factor +
  (1 | plot), green_umbs, REML = FALSE)
anova(umod8, umod9) # umodel 9 is a better fit to data

## Data: green_umbs
## Models:
## umod9: spp_half_cover_date ~ state + origin + (1 + year_factor | plot)
## umod8: spp_half_cover_date ~ state * origin + (1 + year_factor | plot)
##      npar      AIC      BIC logLik deviance Chisq Df Pr(>Chisq)
## umod9      9 7566.6 7608.3 -3774.3 7548.6
## umod8     12 7567.1 7622.7 -3771.5 7543.1 5.4874 3 0.1394
anova(umod9, umod9a) # umod 9a?

## Data: green_umbs
## Models:
## umod9a: spp_half_cover_date ~ state + origin + year_factor + (1 | plot)
## umod9: spp_half_cover_date ~ state + origin + (1 + year_factor | plot)
##      npar      AIC      BIC logLik deviance Chisq Df Pr(>Chisq)
## umod9a      8 7549.8 7586.9 -3766.9 7533.8
## umod9      9 7566.6 7608.3 -3774.3 7548.6 0 1 1

```

```
summary(umod9a)
```

```
## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state + origin + year_factor + (1 | plot)
## Data: green_umbs
##
##      AIC      BIC    logLik deviance df.resid
##  7549.8   7586.9  -3766.9   7533.8     750
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.4628 -0.6648 -0.2797  0.6304  2.9794
##
## Random effects:
## Groups Name Variance Std.Dev.
## plot (Intercept) 6.849 2.617
## Residual 1207.175 34.744
## Number of obs: 758, groups: plot, 24
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept) 149.9167    3.5890 192.7399  41.771 < 2e-16 ***
## stateambient -0.7694    2.7599  19.4947  -0.279  0.7834
## origin       1.4282    3.7694 745.6043   0.379  0.7049
## originBoth    9.9785    5.6683 756.3062   1.760  0.0787 .
## originExotic -14.2582    2.8640 755.5438  -4.978 7.95e-07 ***
## year_factor   4.4146    0.8912 738.1225   4.954 9.04e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sttmbn origin orgnBt orgnEx
## stateambint -0.378
## origin      -0.250  0.006
## originBoth  -0.147 -0.078  0.167
## originExotc -0.320  0.004  0.335  0.221
## year_factor -0.741 -0.010 -0.025  0.002 -0.043
```

```
anova(umod9)
```

```
## Type III Analysis of Variance Table with Satterthwaite's method
##      Sum Sq Mean Sq NumDF DenDF F value    Pr(>F)
## state    126   126.1      1  51.70  0.1051    0.7471
## origin 43190 14396.8      3 737.23 11.9967 1.125e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
emmeans(umod9a, list(pairwise ~ state + origin), adjust = "tukey")
```

```
## $`emmeans of state, origin`
## state origin emmean SE      df lower.CL upper.CL
## warmed Native    163 2.51   59.6     158     168
## ambient Native    163 2.51   58.9     158     168
## warmed          165 3.64  250.2     158     172
```

```

## ambient          164 3.65 256.2      157      171
## warmed Both      173 5.74 531.3      162      185
## ambient Both      173 5.53 472.5      162      184
## warmed Exotic     149 2.69  80.4      144      155
## ambient Exotic     148 2.70  82.4      143      154
##
## Degrees-of-freedom method: kenward-roger
## Confidence level used: 0.95
##
## $`pairwise differences of state, origin`
## 1 estimate SE df t.ratio p.value
## warmed Native - ambient Native 0.769 2.92 26.7 0.263 1.0000
## warmed Native - warmed -1.428 3.78 751.1 -0.377 0.9999
## warmed Native - ambient -0.659 4.79 192.2 -0.137 1.0000
## warmed Native - warmed Both -9.978 5.71 764.1 -1.747 0.6562
## warmed Native - ambient Both -9.209 6.23 332.3 -1.478 0.8185
## warmed Native - warmed Exotic 14.258 2.88 760.6 4.952 <.0001
## warmed Native - ambient Exotic 15.028 4.11 107.6 3.657 0.0091
## ambient Native - warmed -2.198 4.77 185.1 -0.461 0.9998
## ambient Native - ambient -1.428 3.78 751.1 -0.377 0.9999
## ambient Native - warmed Both -10.748 6.60 415.7 -1.629 0.7325
## ambient Native - ambient Both -9.978 5.71 764.1 -1.747 0.6562
## ambient Native - warmed Exotic 13.489 4.10 104.2 3.293 0.0285
## ambient Native - ambient Exotic 14.258 2.88 760.6 4.952 <.0001
## warmed - ambient 0.769 2.92 26.7 0.263 1.0000
## warmed - warmed Both -8.550 6.31 764.2 -1.355 0.8771
## warmed - ambient Both -7.781 6.77 394.9 -1.149 0.9454
## warmed - warmed Exotic 15.686 3.91 748.3 4.012 0.0017
## warmed - ambient Exotic 16.456 4.88 202.8 3.375 0.0196
## ambient - warmed Both -9.320 7.13 476.3 -1.307 0.8961
## ambient - ambient Both -8.550 6.31 764.2 -1.355 0.8771
## ambient - warmed Exotic 14.917 4.89 205.1 3.052 0.0515
## ambient - ambient Exotic 15.686 3.91 748.3 4.012 0.0017
## warmed Both - ambient Both 0.769 2.92 26.7 0.263 1.0000
## warmed Both - warmed Exotic 24.237 5.80 764.3 4.177 0.0009
## warmed Both - ambient Exotic 25.006 6.68 429.2 3.743 0.0050
## ambient Both - warmed Exotic 23.467 6.31 341.7 3.720 0.0056
## ambient Both - ambient Exotic 24.237 5.80 764.3 4.177 0.0009
## warmed Exotic - ambient Exotic 0.769 2.92 26.7 0.263 1.0000
##
## Degrees-of-freedom method: kenward-roger
## P value adjustment: tukey method for comparing a family of 8 estimates
# including growth form - first with interaction term
green_umbs <- within(green_umbs, growth_habit <- relevel(factor(growth_habit),
  ref = "Forb")) # releveling so forb is the reference
umod10 <- lmer(spp_half_cover_date ~ state * growth_habit + (1 +
  year_factor | plot), green_umbs, REML = FALSE)

## boundary (singular) fit: see ?isSingular
umod11 <- lmer(spp_half_cover_date ~ state + growth_habit + (1 +
  year_factor | plot), green_umbs, REML = FALSE)

## boundary (singular) fit: see ?isSingular

```

```

umod11a <- lmer(spp_half_cover_date ~ state + growth_habit +
  year_factor + (1 | plot), green_umbs, REML = FALSE)
anova(umod10, umod11) # umodel 11 is a better fit to data

## Data: green_umbs
## Models:
## umod11: spp_half_cover_date ~ state + growth_habit + (1 + year_factor |
## umod11: plot)
## umod10: spp_half_cover_date ~ state * growth_habit + (1 + year_factor |
## umod10: plot)
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## umod11     9 7596.5 7638.2 -3789.3   7578.5
## umod10    12 7599.7 7655.3 -3787.9   7575.7 2.7863  3    0.4258

anova(umod11, umod11a)

## Data: green_umbs
## Models:
## umod11a: spp_half_cover_date ~ state + growth_habit + year_factor + (1 |
## umod11a: plot)
## umod11: spp_half_cover_date ~ state + growth_habit + (1 + year_factor |
## umod11: plot)
##      npar    AIC    BIC logLik deviance Chisq Df Pr(>Chisq)
## umod11a     8 7581.2 7618.3 -3782.6   7565.2
## umod11     9 7596.5 7638.2 -3789.3   7578.5    0  1          1

summary(umod11a)

## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: spp_half_cover_date ~ state + growth_habit + year_factor + (1 |
## plot)
## Data: green_umbs
##
##      AIC      BIC logLik deviance df.resid
## 7581.2   7618.3 -3782.6   7565.2      750
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.4212 -0.7012 -0.3584  0.7439  2.6628
##
## Random effects:
## Groups   Name                Variance Std.Dev.
## plot     (Intercept)         9.641    3.105
## Residual                    1256.137  35.442
## Number of obs: 758, groups: plot, 24
##
## Fixed effects:
##              Estimate Std. Error    df t value Pr(>|t|)
## (Intercept)   145.8170    3.6139 172.1373  40.349 < 2e-16 ***
## stateambient   -0.2661    2.8843  19.8044  -0.092  0.9274
## growth_habit   15.0043    8.7901 751.3840   1.707  0.0882 .
## growth_habitGraminoid -0.6708    2.6768 757.9303  -0.251  0.8022
## growth_habitTree -16.9489   13.6146 744.4640  -1.245  0.2136
## year_factor     4.3214    0.9101 738.3063   4.748 2.46e-06 ***

```



```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##          (Intr) sttmbn grwth_ grwt_G grwt_T
## stateambint -0.397
## growth_habt -0.097 -0.037
## grwth_hbtGr -0.293  0.006  0.119
## grwth_hbtTr -0.010 -0.012  0.024  0.078
## year_factor -0.764 -0.009  0.030 -0.013 -0.055

anova(umod11)

## Type III Analysis of Variance Table with Satterthwaite's method
##              Sum Sq Mean Sq NumDF  DenDF F value Pr(>F)
## state           36.6   36.62     1  46.04  0.0293 0.8648
## growth_habit 6429.6 2143.21     3 738.29  1.7166 0.1621

emmeans(umod11a, list(pairwise ~ state + growth_habit), adjust = "tukey")

## $`emmeans of state, growth_habit`
##   state growth_habit emmean    SE    df lower.CL upper.CL
## warmed  Forb         159  2.44  43.7     154     164
## ambient  Forb         159  2.43  41.0     154     164
## warmed           174  8.91 730.0     157     192
## ambient           174  8.80 726.5     157     191
## warmed  Graminoid     158  2.69  67.9     153     164
## ambient  Graminoid     158  2.69  69.1     153     163
## warmed    Tree       142 13.78 726.4     115     169
## ambient    Tree       142 13.74 729.3     115     169
##
## Degrees-of-freedom method: kenward-roger
## Confidence level used: 0.95
##
## $`pairwise differences of state, growth_habit`
##   1 estimate    SE    df t.ratio p.value
## warmed Forb - ambient Forb      0.266  3.05  26.3  0.087  1.0000
## warmed Forb - warmed           -15.004  8.83 756.0 -1.699  0.6880
## warmed Forb - ambient          -14.738  9.24 644.9 -1.595  0.7536
## warmed Forb - warmed Graminoid    0.671  2.69 763.7  0.249  1.0000
## warmed Forb - ambient Graminoid    0.937  4.08  87.6  0.230  1.0000
## warmed Forb - warmed Tree      16.949 13.75 758.5  1.233  0.9221
## warmed Forb - ambient Tree      17.215 14.04 700.8  1.226  0.9242
## ambient Forb - warmed          -15.270  9.44 654.0 -1.617  0.7401
## ambient Forb - ambient          -15.004  8.83 756.0 -1.699  0.6880
## ambient Forb - warmed Graminoid    0.405  4.06  81.8  0.100  1.0000
## ambient Forb - ambient Graminoid    0.671  2.69 763.7  0.249  1.0000
## ambient Forb - warmed Tree      16.683 14.13 692.4  1.181  0.9372
## ambient Forb - ambient Tree      16.949 13.75 758.5  1.233  0.9221
## warmed  - ambient      0.266  3.05  26.3  0.087  1.0000
## warmed  - warmed Graminoid  15.675  8.93 760.8  1.756  0.6503
## warmed  - ambient Graminoid  15.941  9.54 646.7  1.671  0.7059
## warmed  - warmed Tree     31.953 16.15 764.4  1.979  0.4972
## warmed  - ambient Tree     32.219 16.46 745.2  1.958  0.5116
## ambient  - warmed Graminoid  15.409  9.33 630.3  1.651  0.7187
```

```
## ambient - ambient Graminoid      15.675  8.93 760.8  1.756  0.6503
## ambient - warmed Tree             31.687 16.41 740.5  1.931  0.5302
## ambient - ambient Tree            31.953 16.15 764.4  1.979  0.4972
## warmed Graminoid - ambient Graminoid  0.266  3.05  26.3  0.087  1.0000
## warmed Graminoid - warmed Tree       16.278 13.81 756.0  1.179  0.9379
## warmed Graminoid - ambient Tree      16.544 14.10 693.4  1.173  0.9394
## ambient Graminoid - warmed Tree      16.012 14.19 687.5  1.129  0.9505
## ambient Graminoid - ambient Tree     16.278 13.81 756.0  1.179  0.9379
## warmed Tree - ambient Tree           0.266  3.05  26.3  0.087  1.0000
##
## Degrees-of-freedom method: kenward-roger
## P value adjustment: tukey method for comparing a family of 8 estimates
```

UMBS Plot-level Mixed Effects Models:

```
mod1pu <- lmer(plot_half_cover_date ~ state + (1 | plot), green_umbsp,
  REML = FALSE)
mod2pu <- lmer(plot_half_cover_date ~ state + year_factor + (1 |
  plot), green_umbsp, REML = FALSE)
mod3pu <- lmer(plot_half_cover_date ~ state * year_factor + (1 |
  plot), green_umbsp, REML = FALSE)
anova(mod1pu, mod2pu, mod3pu) #mod2pu
```

```
## Data: green_umbsp
## Models:
## mod1pu: plot_half_cover_date ~ state + (1 | plot)
## mod2pu: plot_half_cover_date ~ state + year_factor + (1 | plot)
## mod3pu: plot_half_cover_date ~ state * year_factor + (1 | plot)
##      npar    AIC    BIC logLik deviance  Chisq Df Pr(>Chisq)
## mod1pu    4 1157.9 1169.0 -574.93   1149.9
## mod2pu    5 1129.5 1143.5 -559.78   1119.5 30.3054  1  3.691e-08 ***
## mod3pu    6 1131.4 1148.1 -559.68   1119.4  0.1868  1    0.6656
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(mod2pu)
```

```
## Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's
## method [lmerModLmerTest]
## Formula: plot_half_cover_date ~ state + year_factor + (1 | plot)
## Data: green_umbsp
##
##      AIC      BIC    logLik deviance df.resid
## 1129.6   1143.5   -559.8   1119.6      115
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -1.4880 -0.5653 -0.2716  0.3457  3.8657
##
## Random effects:
## Groups   Name                Variance Std.Dev.
## plot     (Intercept)         72.17     8.495
## Residual                        600.40    24.503
## Number of obs: 120, groups: plot, 24
```

```
##
## Fixed effects:
##           Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)  125.358      6.208  92.959  20.194 < 2e-16 ***
## stateambient   -2.883      5.661  24.000  -0.509   0.615
## year_factor     9.442      1.582  96.000   5.969 3.99e-08 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##           (Intr) sttmbn
## stateambint -0.456
## year_factor -0.764  0.000
outlierTest(mod2pu) # remove these outliers?

##      rstudent unadjusted p-value Bonferroni p
## 118 4.042307      9.6488e-05    0.011579
## 158 3.757143      2.7250e-04    0.032700
```

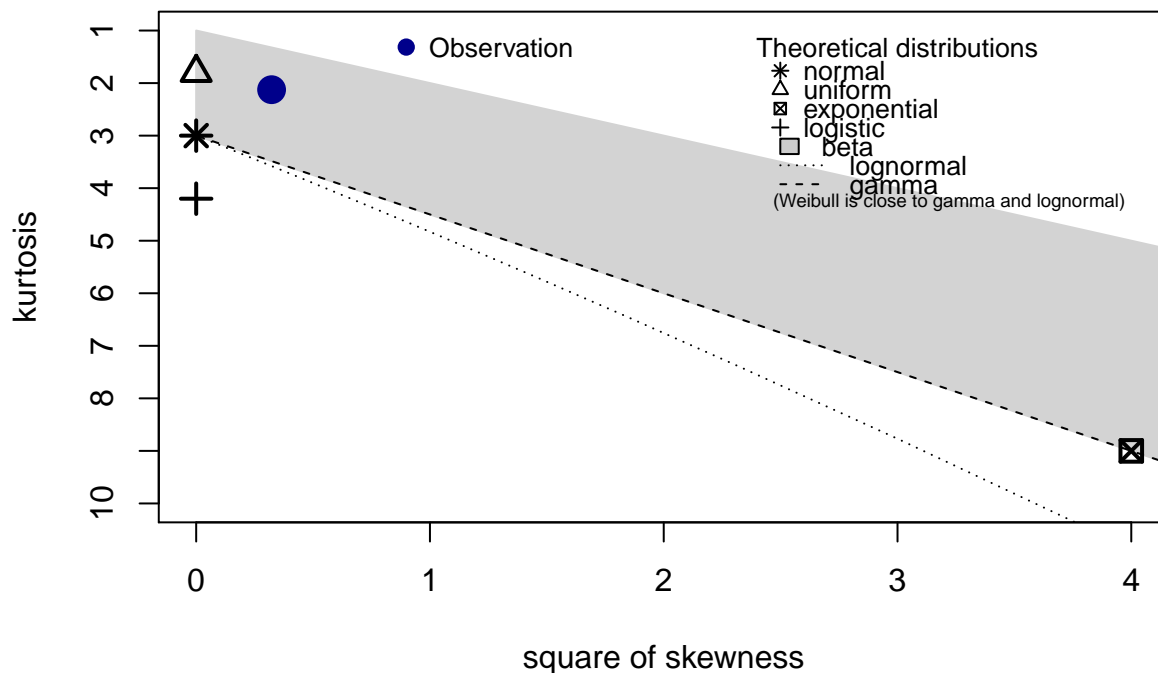
ORIGINAL CODE BELOW; not edited by Phoebe

can pretty much ignore everything below!

Seeing what other distribution could fit

```
descdist(green_kbs$spp_half_cover_date, discrete = FALSE)
```

Cullen and Frey graph

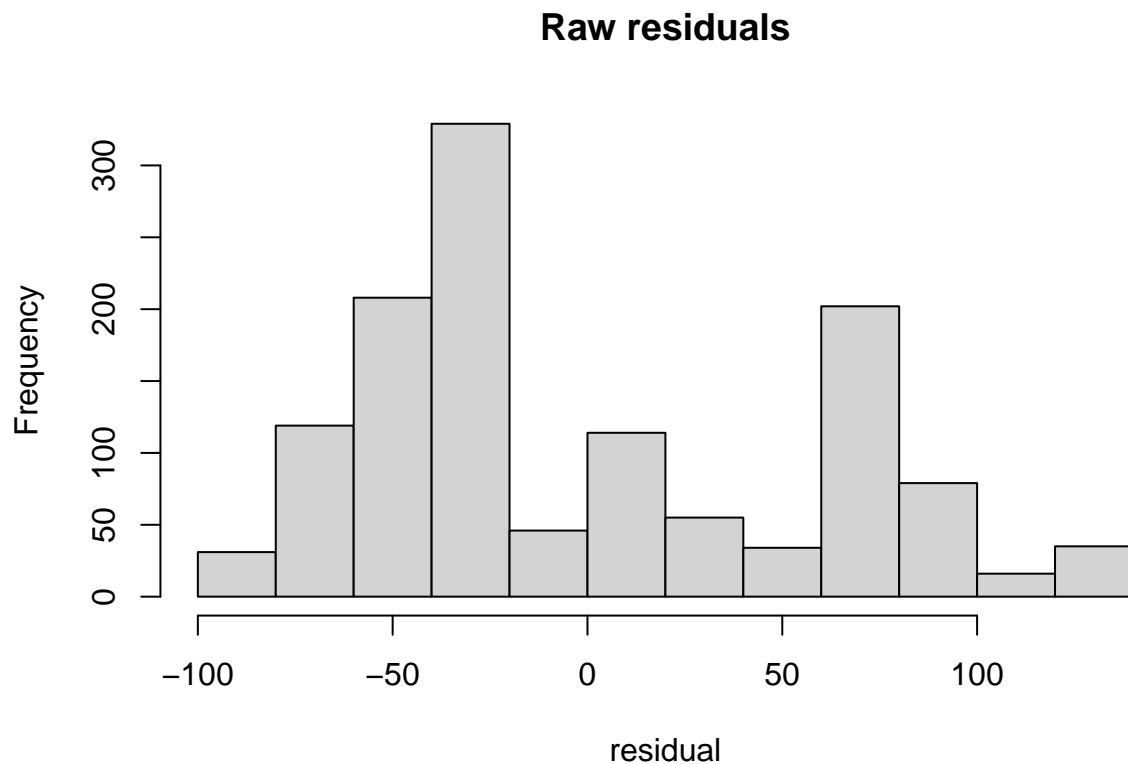


```
## summary statistics
```

```
## -----  
## min: 59   max: 289  
## median: 127  
## mean: 154.2169  
## estimated sd: 57.9311  
## estimated skewness: 0.5680173  
## estimated kurtosis: 2.125259
```

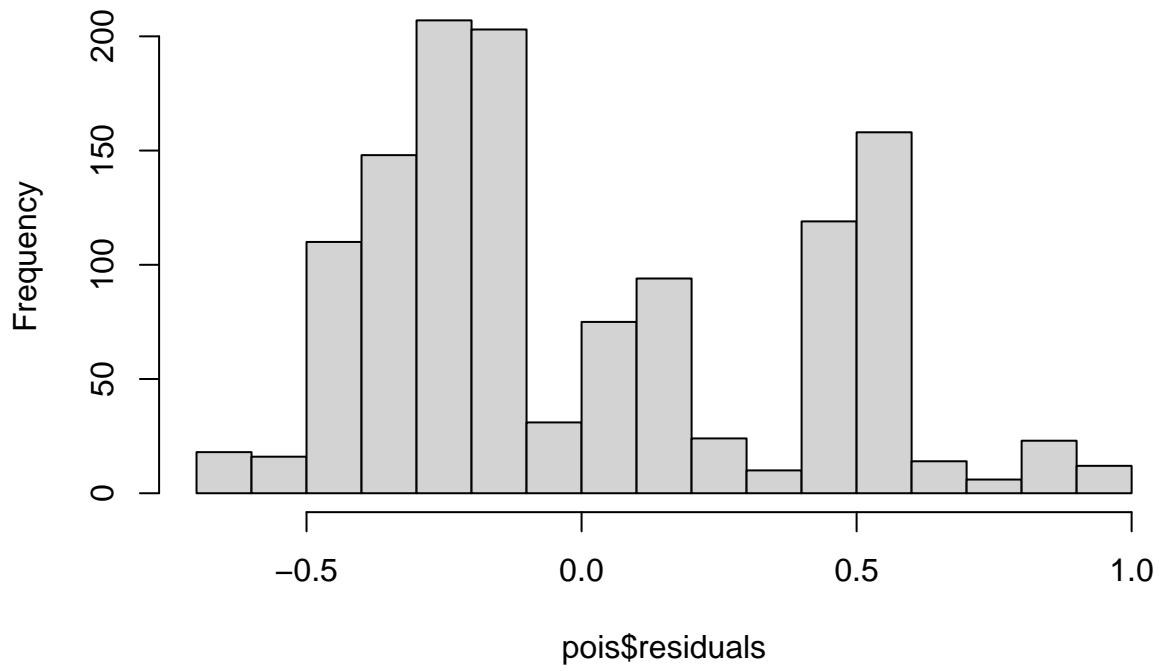
While uniform looks the closest, I'll try poisson

```
fit <- lm(spp_half_cover_date ~ state, data = green_kbs)  
residual <- fit$residuals  
hist(residual, main = "Raw residuals")
```



```
pois <- glm(spp_half_cover_date ~ state, data = green_kbs, family = "poisson")  
hist(pois$residuals, main = "Poisson glm residuals")
```

Poisson glm residuals



Below I try a few different generalized linear models with poisson distribution:

An interaction between state and year, plus insecticide as a fixed effect and species and plot as random effects

```
moda <- glmer(spp_half_cover_date ~ state * year + insecticide +
  (1 | species) + (1 | plot), data = green_kbs, family = poisson)

## Warning: Some predictor variables are on very different scales: consider
## rescaling

## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.0232094 (tol = 0.002, component 1)

## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unidentifiable:
## - Rescale variables?;Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?

summary(moda)

## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]
## Family: poisson ( log )
## Formula: spp_half_cover_date ~ state * year + insecticide + (1 | species) +
## (1 | plot)
## Data: green_kbs
##
##      AIC      BIC   logLik deviance df.resid
## 27686.8 27722.8 -13836.4 27672.8    1261
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -8.3234 -3.0092 -0.9901  2.4141 14.4417
##
## Random effects:
## Groups Name Variance Std.Dev.
## plot (Intercept) 0.003601 0.06001
## species (Intercept) 0.043398 0.20832
## Number of obs: 1268, groups: plot, 24; species, 21
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -82.699485   4.741076 -17.443 <2e-16 ***
## stateambient    -15.581997   6.330178  -2.462  0.0138 *
## year              0.043454   0.002349  18.497 <2e-16 ***
## insecticideno_insects  0.036029   0.024939   1.445  0.1485
## stateambient:year   0.007736   0.003137   2.466  0.0137 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sttmbn year  insct_
## stateambint -0.727
## year        -1.000  0.727
```

```
## insctcdn_ns -0.017  0.007  0.015
## statmbnt:yr  0.727 -1.000 -0.727 -0.007
## fit warnings:
## Some predictor variables are on very different scales: consider rescaling
## convergence code: 0
## Model failed to converge with max|grad| = 0.0232094 (tol = 0.002, component 1)
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
## Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
```

No interaction between state and year, but with state and insecticide as fixed effects and species and plot as random effects

```
modb <- glmer(spp_half_cover_date ~ state + year + insecticide +
  (1 | species) + (1 | plot), data = green_kbs, family = poisson)
```

```
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.00310689 (tol = 0.002, component 1)

## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unidentifiable:
## - Rescale variables?;Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
```

```
summary(modb)
```

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]
## Family: poisson ( log )
## Formula: spp_half_cover_date ~ state + year + insecticide + (1 | species) +
## (1 | plot)
## Data: green_kbs
##
##          AIC          BIC    logLik deviance df.resid
## 27690.8 27721.7 -13839.4 27678.8      1262
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -8.3309 -3.0222 -0.9997  2.3954 14.2765
##
## Random effects:
## Groups Name Variance Std.Dev.
## plot (Intercept) 0.003606 0.06005
## species (Intercept) 0.043488 0.20854
## Number of obs: 1268, groups: plot, 24; species, 21
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -91.116751   3.172907 -28.717 <2e-16 ***
## stateambient      0.030236   0.024956  1.212 0.226
## year             0.047625   0.001572 30.293 <2e-16 ***
## insecticideno_insects 0.036450   0.024954  1.461 0.144
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sttmbn year
## stateambint  0.002
## year        -1.000 -0.006
## insctcdn_ns -0.018 -0.002  0.014
## convergence code: 0
## Model failed to converge with max|grad| = 0.00310689 (tol = 0.002, component 1)
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
## Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
```


State and insecticide as fixed effects & year, species and plot as random effects

```
modc <- glmer(spp_half_cover_date ~ state + insecticide + (1 |  
  year) + (1 | species) + (1 | plot), data = green_kbs, family = poisson)  
summary(modc)
```

Because no distributions seems to match well, I'll try a Friedman's test

```
# friedman_kbs <- green_kbs %>%  
# friedman_test(spp_half_cover_date ~ state)
```

Error: Must extract column with a single valid subscript. x Subscript var can't be NA

Can't figure out what this means

If I include the blocks portion of the formula (from the documentation) I get this error

```
# friedman_kbs <- green_kbs %>%  
# friedman_test(spp_half_cover_date ~ state | plot)
```

Error in friedman.test.default(c(141L, 202L, 122L, 101L, 127L, 120L, 197L, : not an unreplicated complete block design

Permanova?

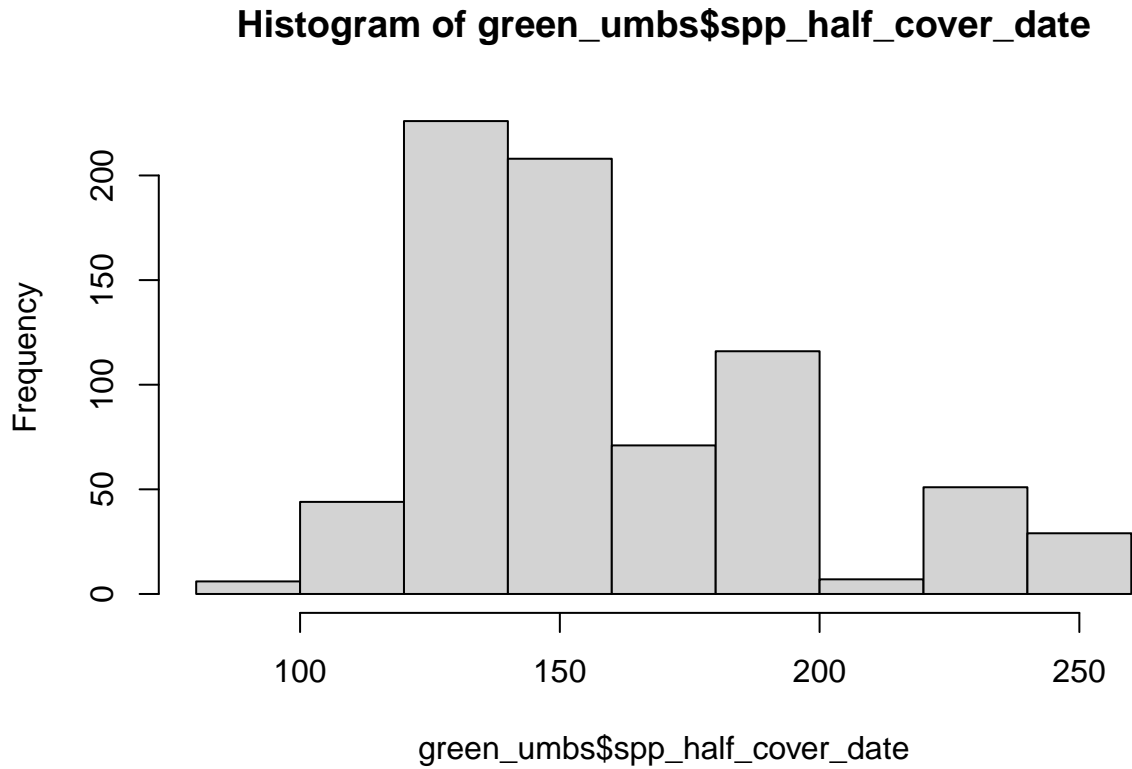
```
per1 <- adonis2(green_kbs$spp_half_cover_date ~ state * year +  
  insecticide, data = green_kbs)  
per1  
per2 <- adonis(formula = green_kbs$spp_half_cover_date ~ state *  
  year + insecticide, strata = green_kbs$plot, data = green_kbs)  
per2
```

With per2, when controlling for “plot”, there is a difference btwn treatments

UMBS

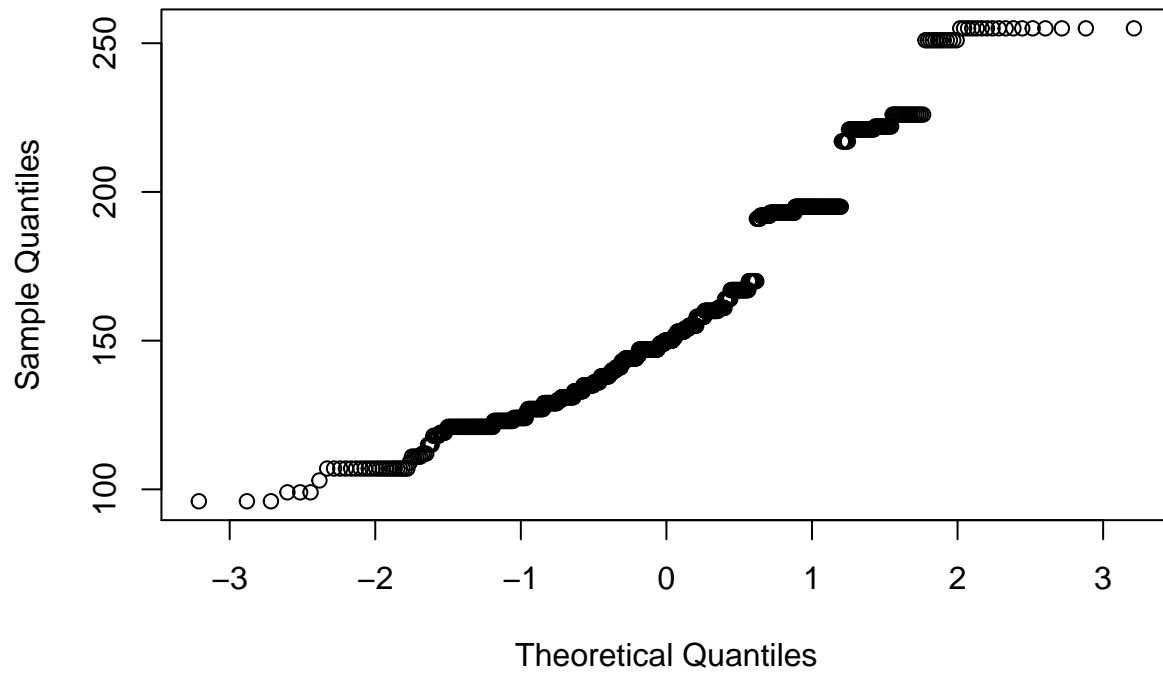
Checking for normality

```
hist(green_umbs$spp_half_cover_date)
```



```
qqnorm(green_umbs$spp_half_cover_date)
```

Normal Q-Q Plot

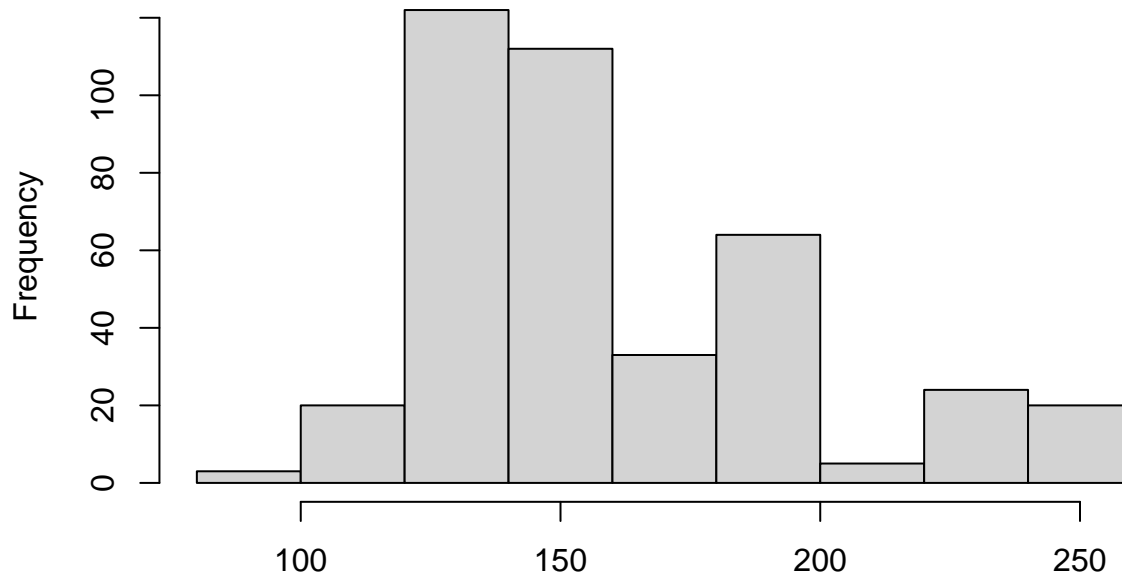


```
shapiro.test(green_umbs$spp_half_cover_date)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data: green_umbs$spp_half_cover_date  
## W = 0.92247, p-value < 2.2e-16
```

```
hist(green_umbs$spp_half_cover_date[green_kbs$state == "ambient"])
```

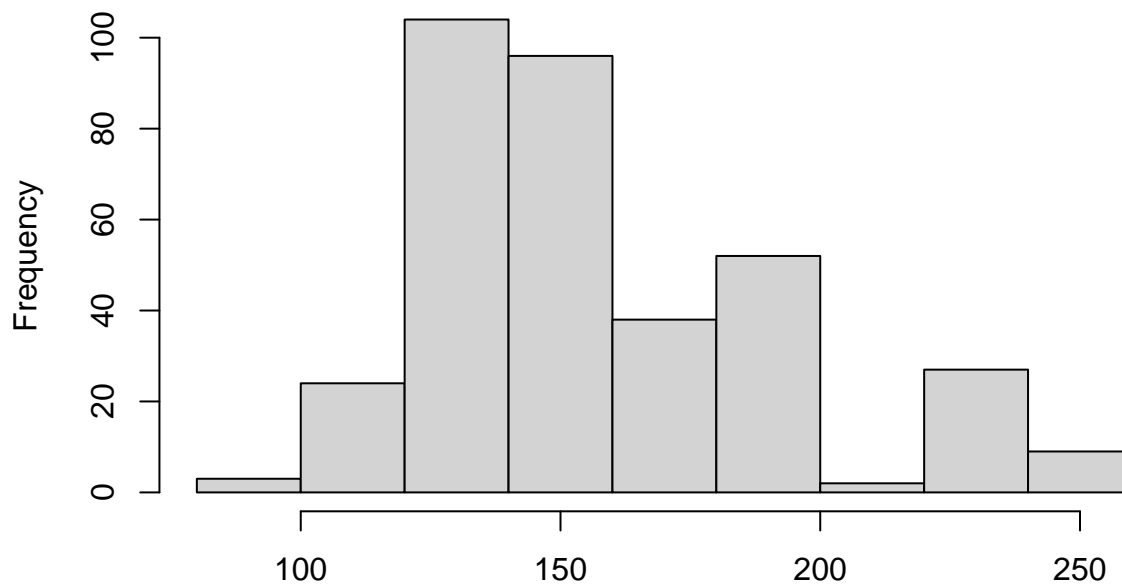
istogram of green_umbs\$spp_half_cover_date[green_kbs\$state == "am



green_umbs\$spp_half_cover_date[green_kbs\$state == "ambient"]

```
hist(green_umbs$spp_half_cover_date[green_kbs$state == "warmed"])
```

istogram of green_umbs\$spp_half_cover_date[green_kbs\$state == "wa

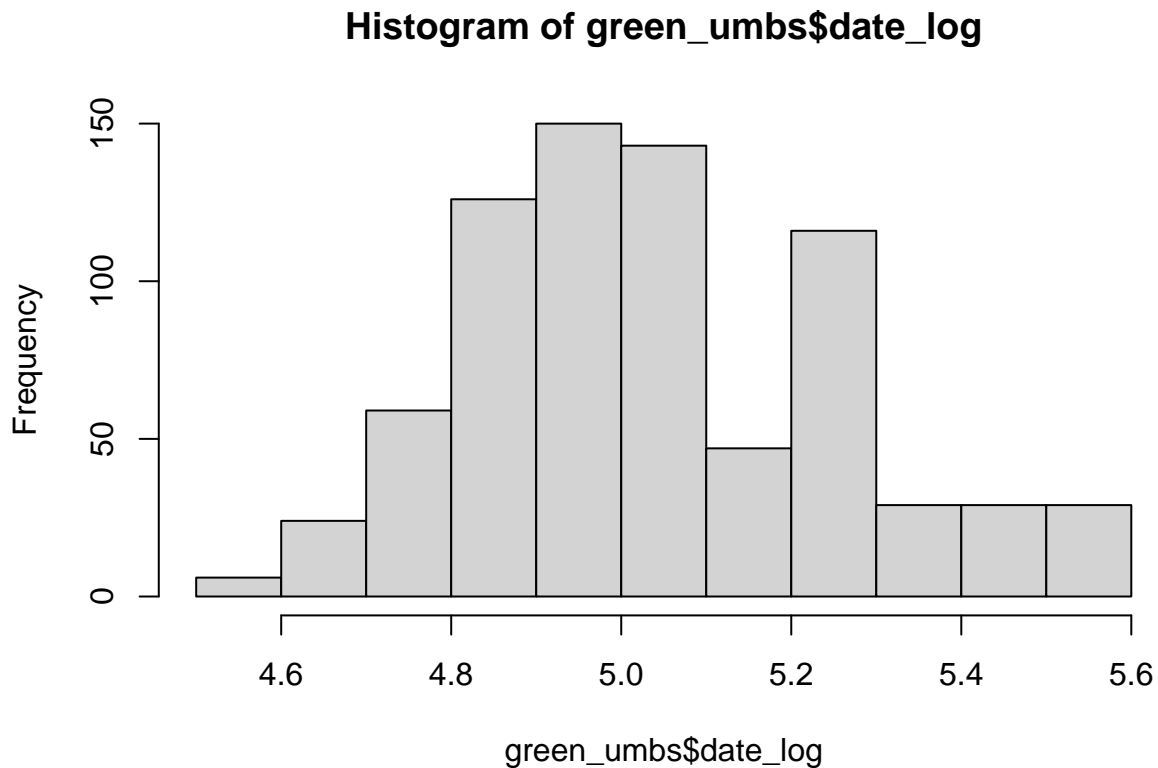


green_umbs\$spp_half_cover_date[green_kbs\$state == "warmed"]

These look pretty good

Trying log transformation

```
green_umbs$date_log <- log(green_umbs$spp_half_cover_date)
hist(green_umbs$date_log)
```



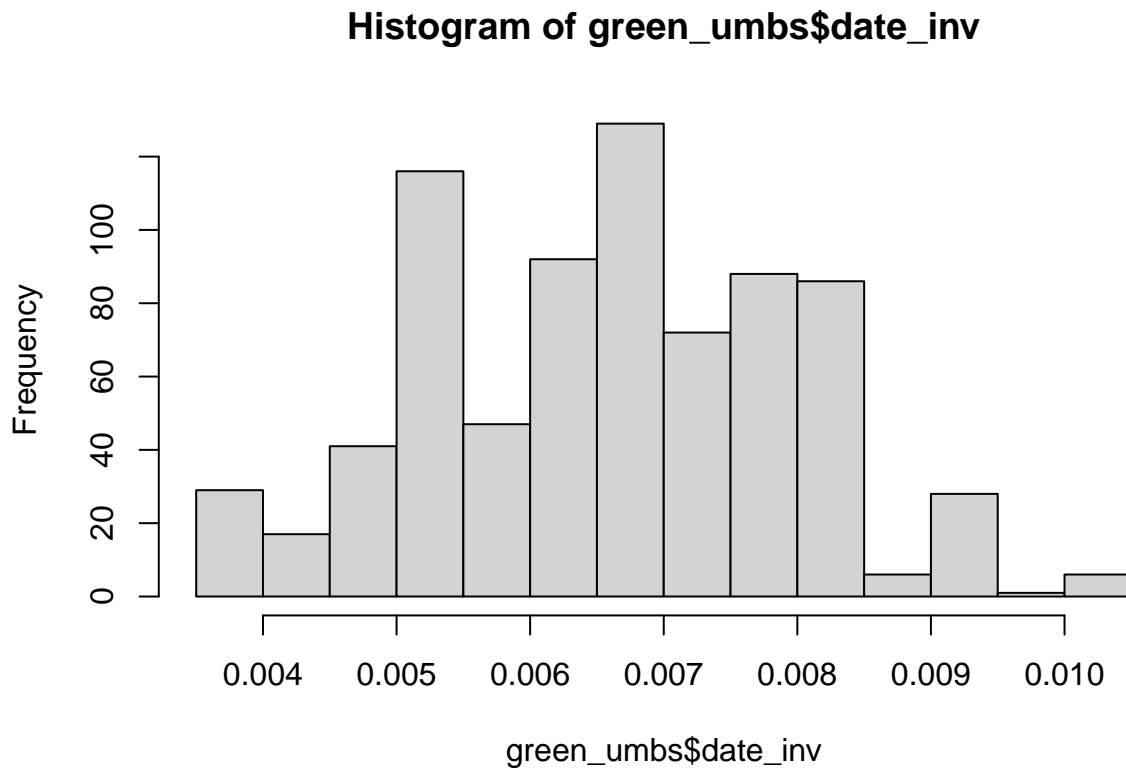
```
shapiro.test(green_umbs$date_log)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  green_umbs$date_log
## W = 0.96356, p-value = 8.516e-13
```

I think this looks good but shapiro-wilk is lower than 0.05

Trying inverse tranformation

```
green_umbs$date_inv <- 1/(green_umbs$spp_half_cover_date)
hist(green_umbs$date_inv)
```



```
shapiro.test(green_umbs$date_inv)

##
##  Shapiro-Wilk normality test
##
## data:  green_umbs$date_inv
## W = 0.97928, p-value = 6.952e-09
```

This also looks good but is also still low for shapiro-wilk