

# Greenup Analyses

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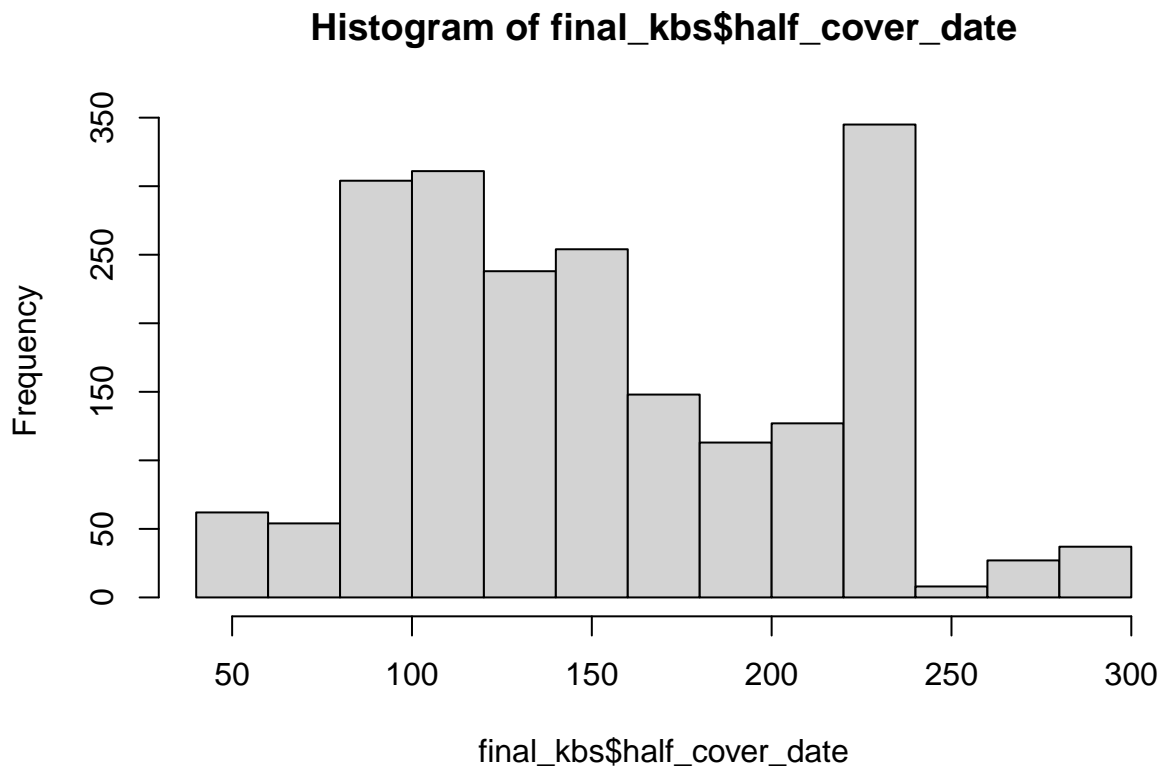
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DATA INPUT: Data imported as csv files from shared Google drive L1 plant comp folder  
PROJECT: warmXtrophic

## *Starting with KBS*

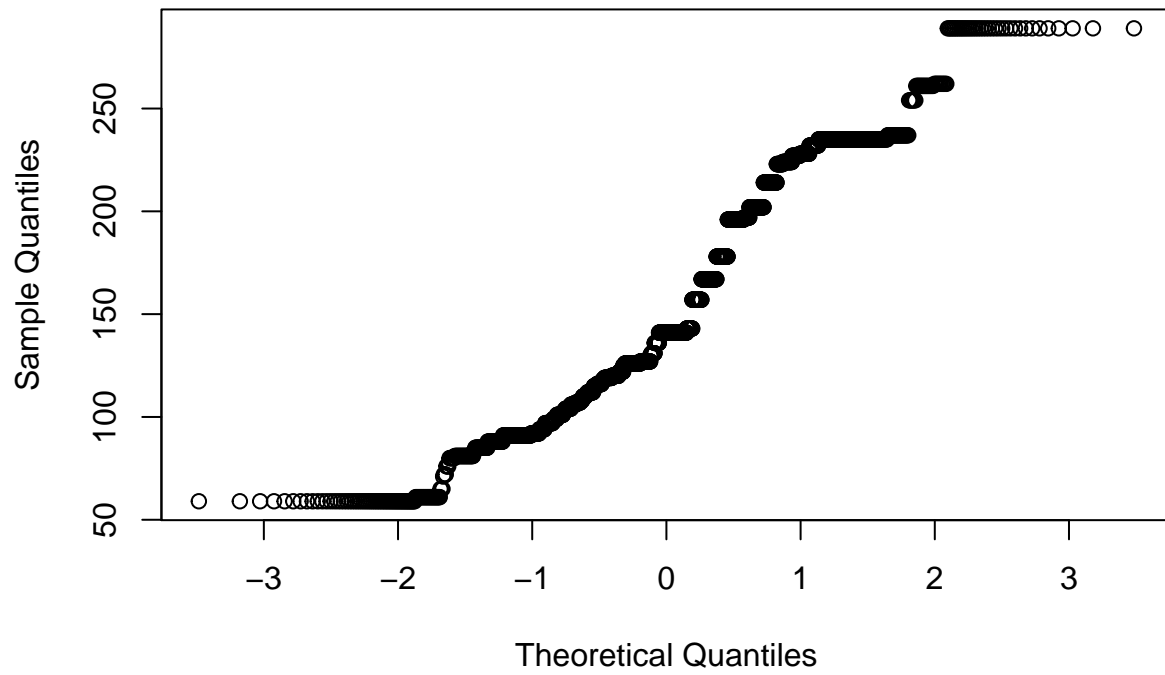
First, checking for normality

```
hist(final_kbs$half_cover_date)
```



```
qqnorm(final_kbs$half_cover_date)
```

## Normal Q-Q Plot

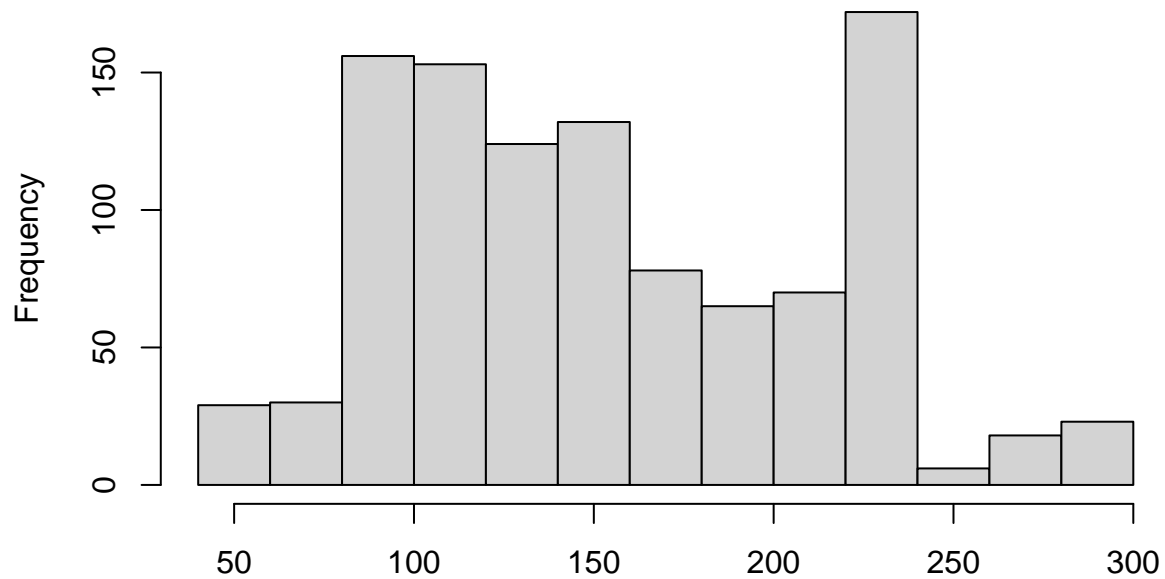


```
shapiro.test(final_kbs$half_cover_date)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data:  final_kbs$half_cover_date  
## W = 0.94273, p-value < 2.2e-16
```

```
hist(final_kbs$half_cover_date[final_kbs$state == "ambient"])
```

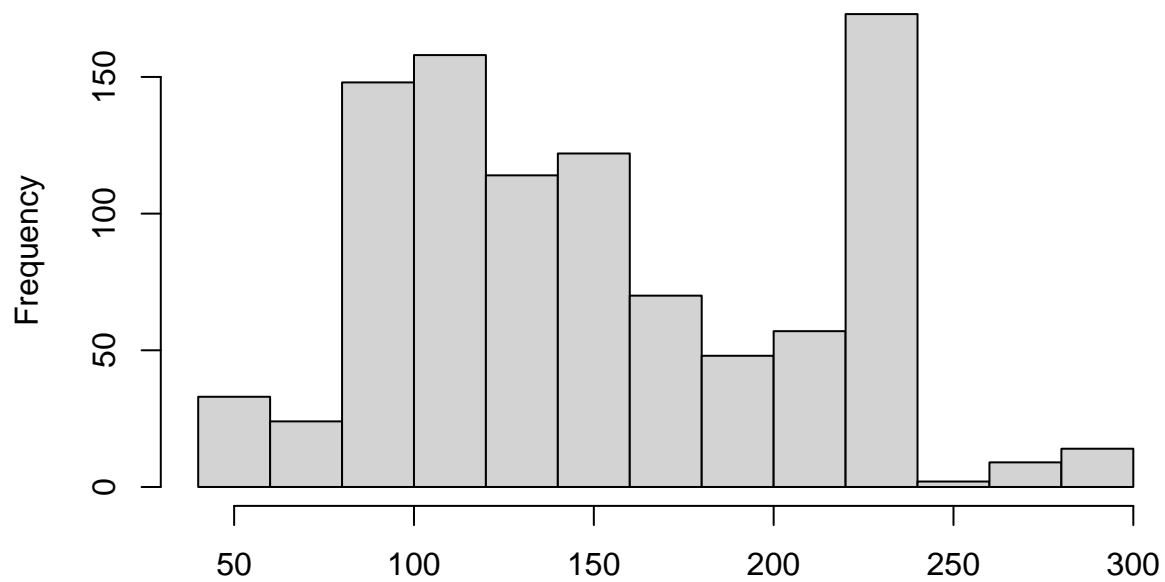
**Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$state == "ambient"]**



final\_kbs\$half\_cover\_date[final\_kbs\$state == "ambient"]

```
hist(final_kbs$half_cover_date[final_kbs$state == "warmed"])
```

**Histogram of final\_kbs\$half\_cover\_date[final\_kbs\$state == "warmed"]**

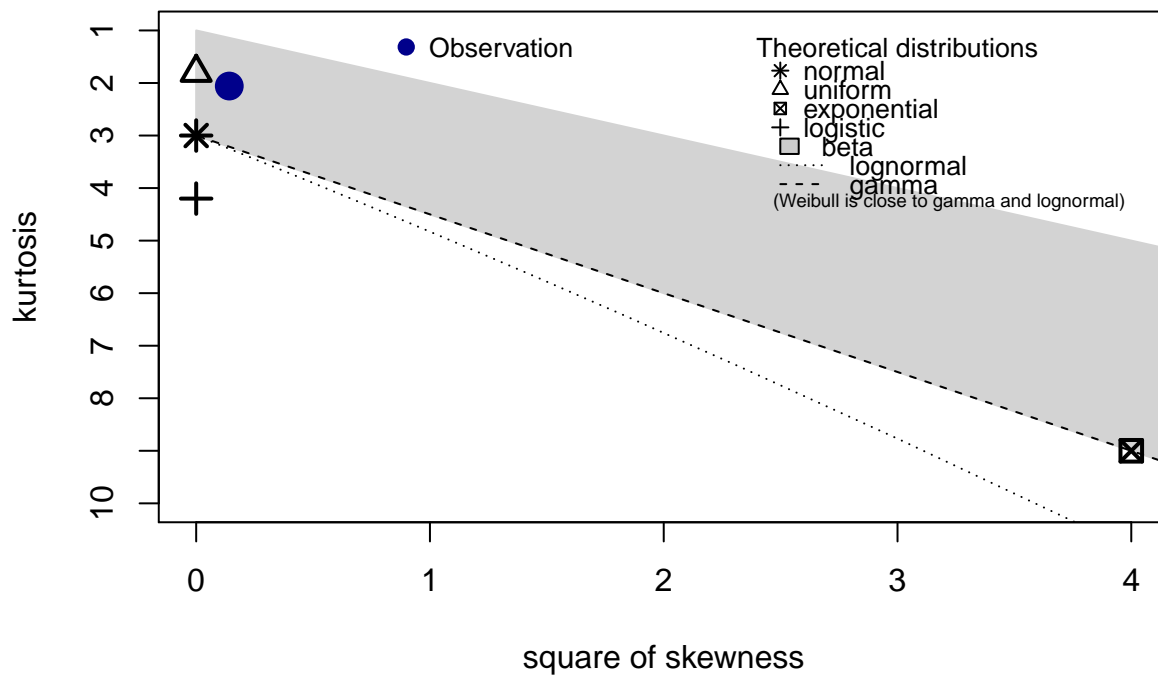


final\_kbs\$half\_cover\_date[final\_kbs\$state == "warmed"]

Not normal, and previously attempted transformations don't help

Seeing what other distribution could fit

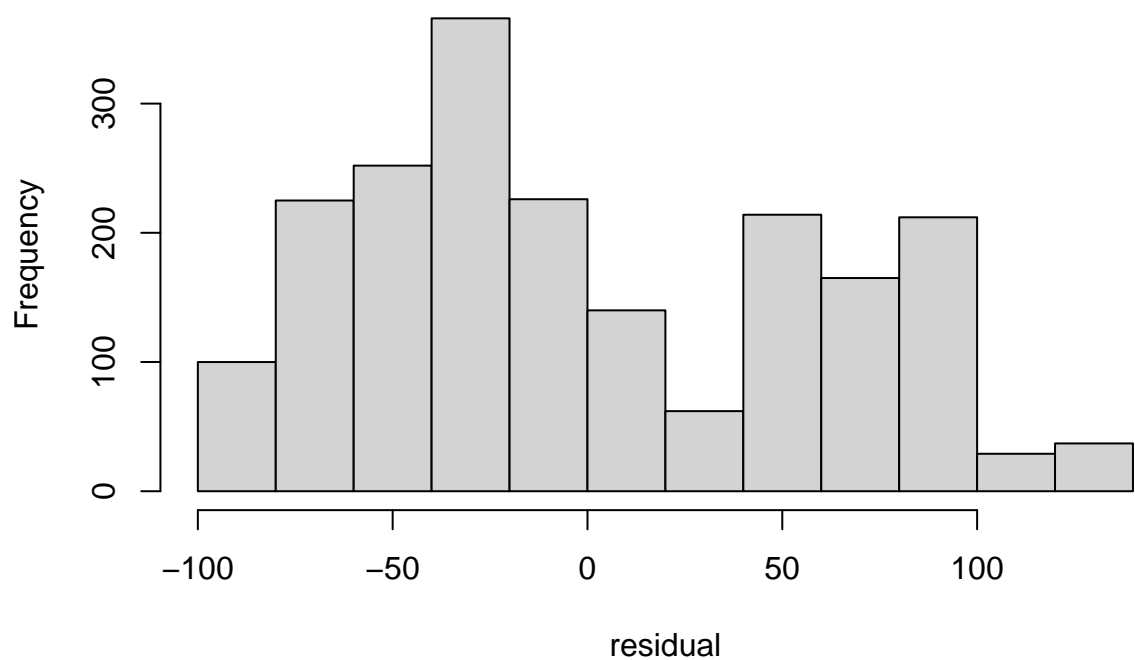
### Cullen and Frey graph



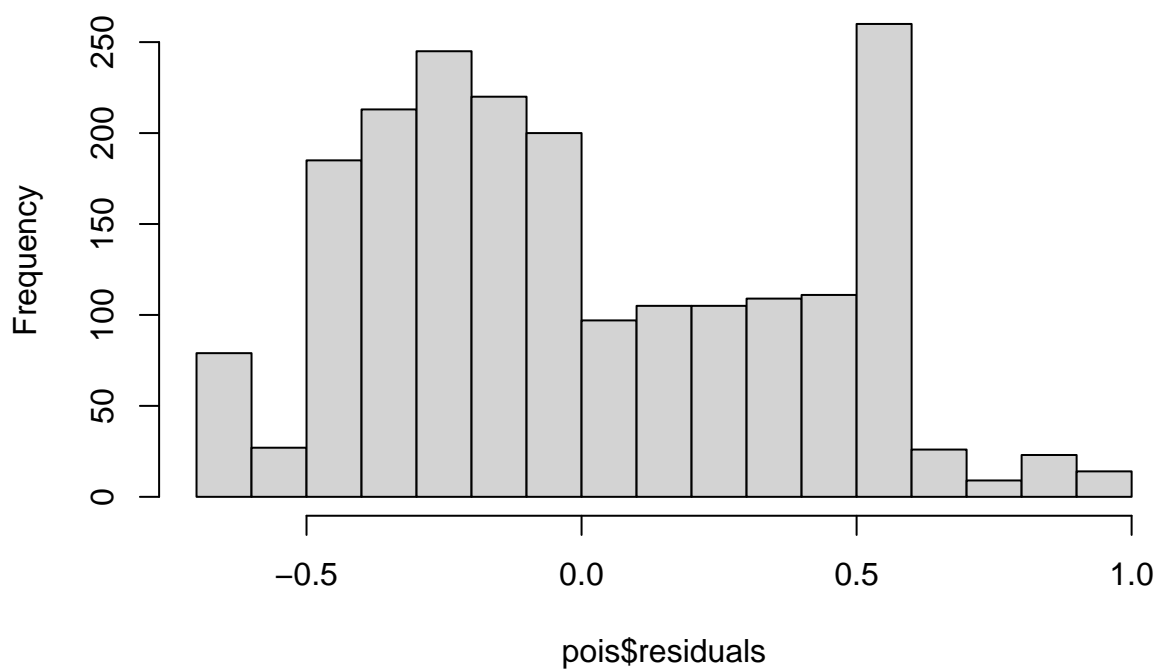
```
## summary statistics
## -----
## min: 59    max: 289
## median: 141
## mean: 152.8644
## estimated sd: 57.73216
## estimated skewness: 0.3763172
## estimated kurtosis: 2.058193
```

While uniform looks the closest, I'll try poisson

### Raw residuals



### Poisson glm residuals



Below I try a few different generalized linear models with poisson distribution:

An interaction between state and year, plus insecticide as a fixed effect and species and plot as random effects

```
moda <- glmer(half_cover_date ~ state*year + insecticide + (1|species) + (1|plot),
              data=final_kbs, family = poisson)
```

```
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unidentifiable:
## - Rescale variables?
```

```
summary(moda)
```

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]
## Family: poisson ( log )
## Formula: half_cover_date ~ state * year + insecticide + (1 | species) +
## (1 | plot)
## Data: final_kbs
##
##          AIC          BIC    logLik deviance df.resid
## 42924.6   42963.9 -21455.3  42910.6      2021
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -9.0161 -2.8352 -0.6425  2.3763 14.4656
##
## Random effects:
## Groups Name Variance Std.Dev.
## species (Intercept) 0.069954 0.26449
## plot (Intercept) 0.001848 0.04299
## Number of obs: 2028, groups: species, 55; plot, 24
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    5.020670   0.039631 126.684 < 2e-16 ***
## statewarmed     0.009389   0.019347   0.485  0.6275
## year            0.003600   0.001520   2.368  0.0179 *
## insecticideno_insects 0.028934   0.017937   1.613  0.1067
## statewarmed:year -0.009698   0.002116 -4.584 4.57e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sttwrm year  insct_
## statewarmed -0.243
## year        -0.130  0.249
## insctcdn_ns -0.227 -0.002  0.008
## statwrmd:yr  0.089 -0.375 -0.648  0.006
## convergence code: 0
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
```

No interaction between state and year, but with state and insecticide as fixed effects and species and plot as random effects

```
modb <- glmer(half_cover_date ~ state + year + insecticide + (1|species) + (1|plot),
              data=final_kbs, family = poisson)
```

```
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unidentifiable:
## - Rescale variables?
```

```
summary(modb)
```

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]
## Family: poisson ( log )
## Formula: half_cover_date ~ state + year + insecticide + (1 | species) +
## (1 | plot)
## Data: final_kbs
##
##      AIC      BIC    logLik deviance df.resid
## 42943.6 42977.3 -21465.8 42931.6      2022
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -9.0053 -2.8448 -0.6575  2.3929 14.2130
##
## Random effects:
## Groups Name          Variance Std.Dev.
## species (Intercept) 0.069954 0.26449
## plot (Intercept) 0.001847 0.04297
## Number of obs: 2028, groups: species, 55; plot, 24
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      5.0367372  0.0394723 127.602  <2e-16 ***
## statewarmed      -0.0238775  0.0179292  -1.332   0.183
## year             -0.0009128  0.0011579  -0.788   0.430
## insecticideno_insects 0.0294351  0.0179303   1.642   0.101
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sttwrm year
## statewarmed -0.227
## year        -0.096  0.007
## insctcdn_ns -0.228  0.001  0.015
## convergence code: 0
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
```

## State and insecticide as fixed effects & year, species and plot as random effects

```
modc <- glmer(half_cover_date ~ state + insecticide + (1|year) + (1|species) + (1|plot),
              data=final_kbs, family = poisson)
summary(modc)
```

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
##   Approximation) [glmerMod]
##   Family: poisson ( log )
## Formula: half_cover_date ~ state + insecticide + (1 | year) + (1 | species) +
##           (1 | plot)
##   Data: final_kbs
##
##           AIC          BIC    logLik deviance df.resid
##  41436.8   41470.5 -20712.4  41424.8      2022
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -9.3418 -2.8134 -0.5215  2.1783 13.1976
##
## Random effects:
##   Groups Name          Variance Std.Dev.
##   species (Intercept) 0.068407 0.26155
##   plot    (Intercept) 0.001820 0.04266
##   year    (Intercept) 0.005866 0.07659
## Number of obs: 2028, groups: species, 55; plot, 24; year, 6
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      5.01945    0.04991 100.579 <2e-16 ***
## statewarmed      -0.02683    0.01781  -1.507  0.1319
## insecticideno_insects 0.03093    0.01781   1.737  0.0824 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sttwrm
## statewarmed -0.178
## insctcdn_ns -0.178  0.001
```



Because no distributions seems to match well, I'll try a Friedman's test

```
#friedman_kbs <- final_kbs %>%  
#  friedman_test(half_cover_date ~ state)
```

I get this - Error: Must extract column with a single valid subscript. x Subscript var can't be NA

Can't figure out what this means

If I include the blocks portion of the formula (from the documentation) I get this error

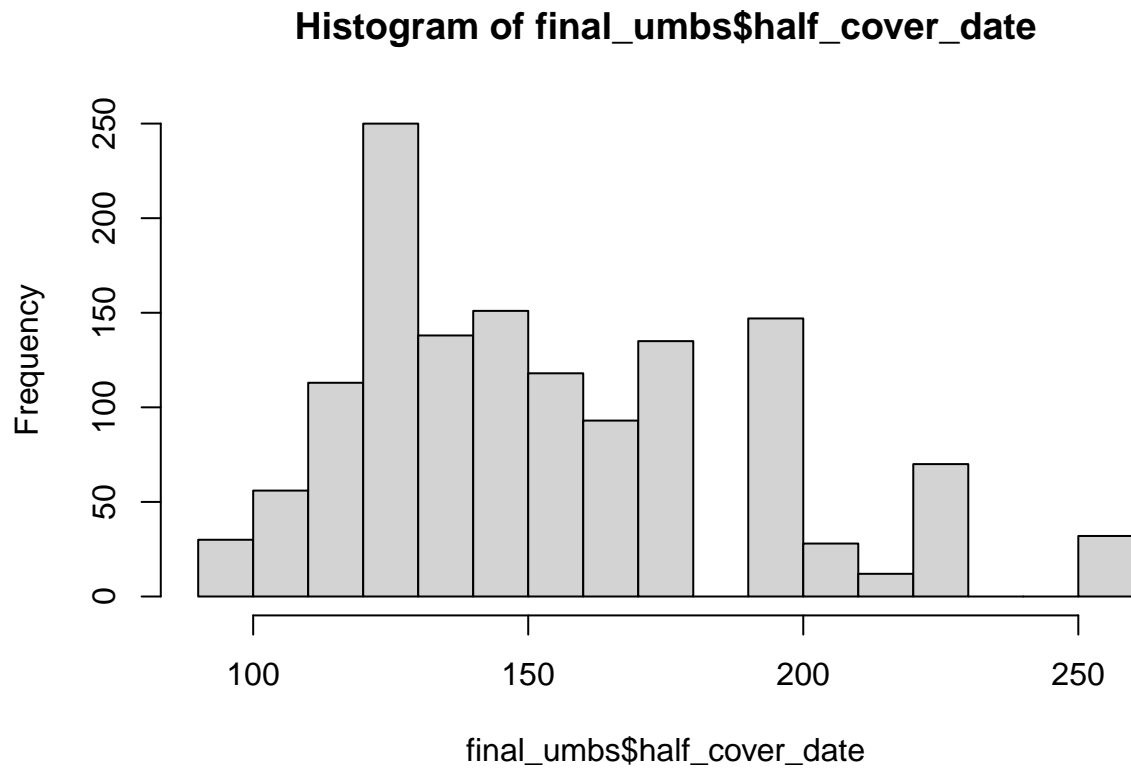
```
#friedman_kbs <- final_kbs %>%  
#  friedman_test(half_cover_date ~ state | plot)
```

Error in friedman.test.default(c(141L, 202L, 122L, 101L, 127L, 120L, 197L, : not an unreplicated complete block design

## UMBS

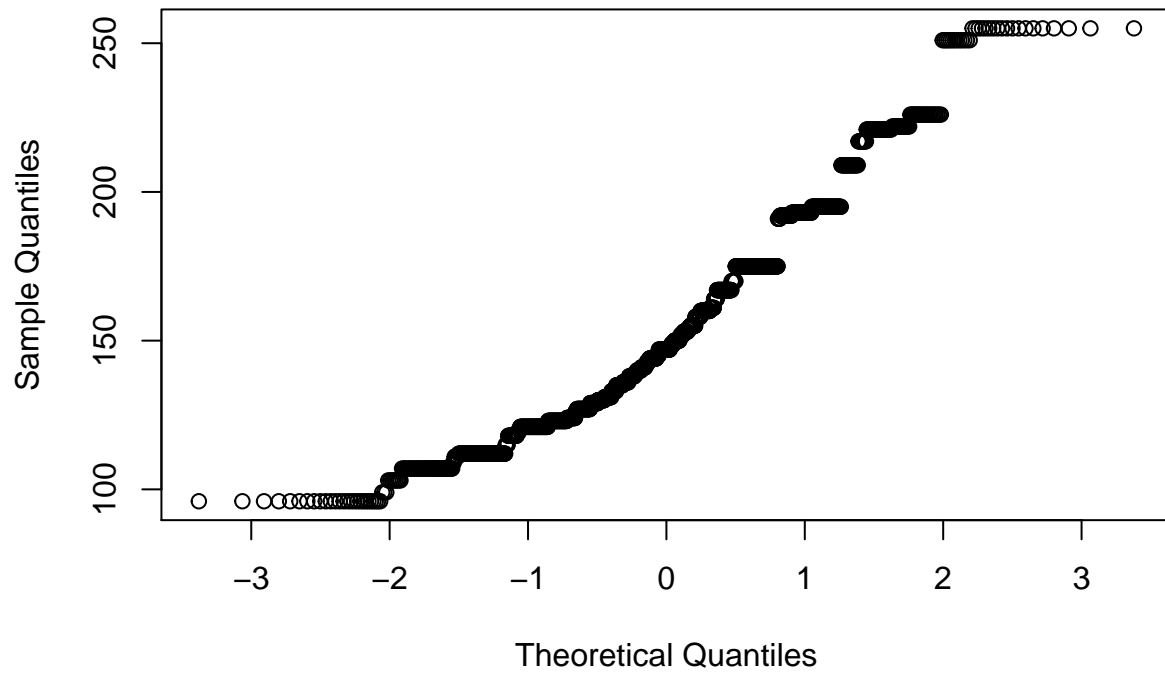
### Checking for normality

```
hist(final_umbs$half_cover_date)
```



```
qqnorm(final_umbs$half_cover_date)
```

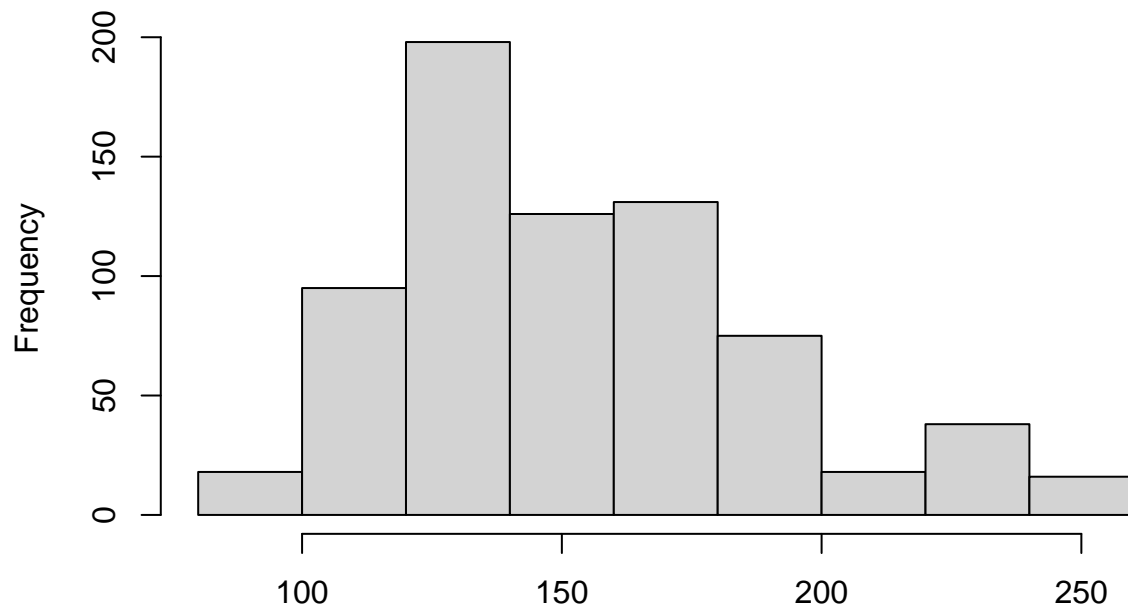
## Normal Q-Q Plot



```
shapiro.test(final_umbs$half_cover_date)
```

```
##  
##  Shapiro-Wilk normality test  
##  
## data:  final_umbs$half_cover_date  
## W = 0.94754, p-value < 2.2e-16  
hist(final_umbs$half_cover_date[final_kbs$state == "ambient"])
```

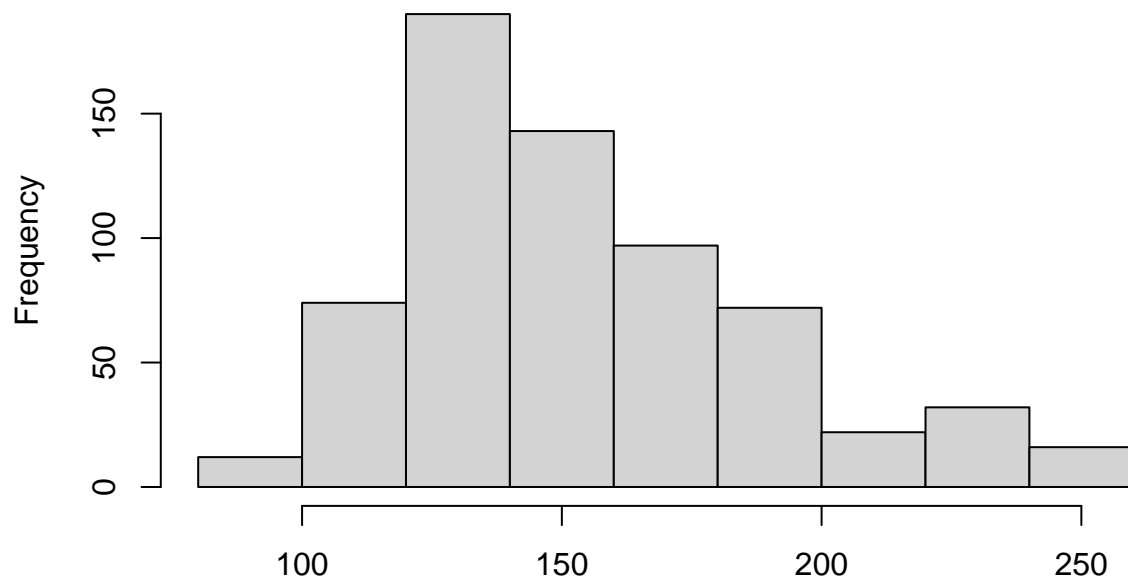
**Histogram of final\_umbs\$half\_cover\_date[final\_kbs\$state == "ambier**



final\_umbs\$half\_cover\_date[final\_kbs\$state == "ambient"]

```
hist(final_umbs$half_cover_date[final_kbs$state == "warmed"])
```

**Histogram of final\_umbs\$half\_cover\_date[final\_kbs\$state == "warmed**

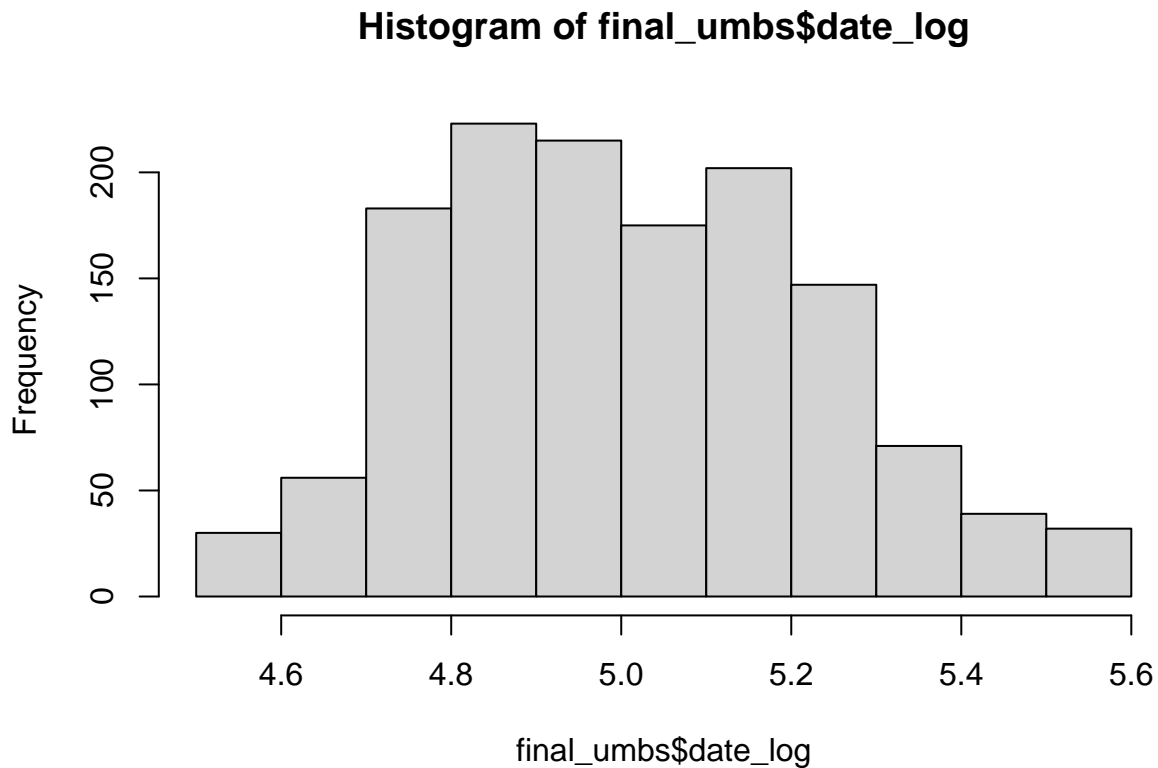


final\_umbs\$half\_cover\_date[final\_kbs\$state == "warmed"]

These look pretty good

## Trying log transformation

```
final_umbs$date_log <- log(final_umbs$half_cover_date)
hist(final_umbs$date_log)
```



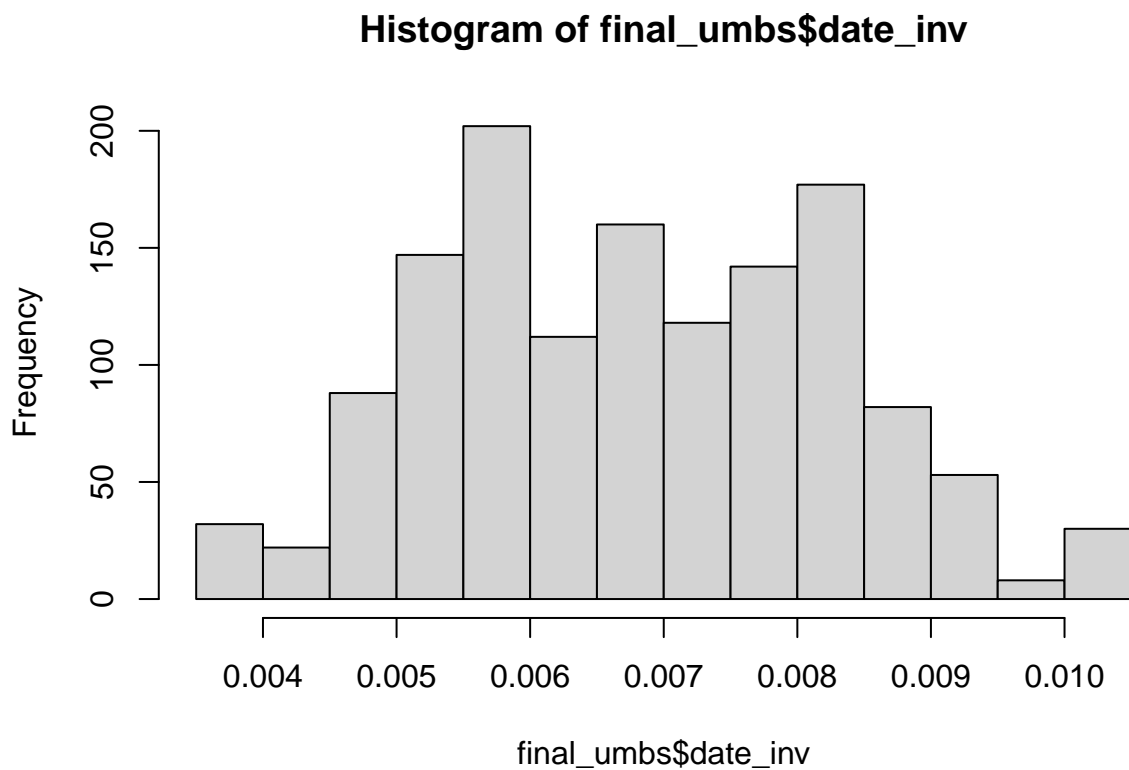
```
shapiro.test(final_umbs$date_log)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  final_umbs$date_log
## W = 0.97728, p-value = 6.765e-14
```

I think this looks good but shapiro-wilk is lower than 0.05

## Trying inverse tranformation

```
final_umbs$date_inv <- 1/(final_umbs$half_cover_date)
hist(final_umbs$date_inv)
```



```
shapiro.test(final_umbs$date_inv)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  final_umbs$date_inv
## W = 0.97993, p-value = 6.679e-13
```

This also looks good but is also still low for shapiro-wilk