PCML Cheat Sheet

1 Math Prerequisites

- Baves rule

$$p(A,B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

- Gaussian distribution

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}))$$
Tu

- Production of independent variables

$$V(XY) = E(X^{2})E(Y^{2}) - [E(X)]^{2}[E(Y)]^{2}$$

Covariance matrix of a data vector x

$$\Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - E(\mathbf{x})(\mathbf{x}_n - E(\mathbf{x})^T)$$

- A function is convex when a line joining two points never interects with the function anywhere else.
- A function f(x) is convex, if for any $x_1, x_2 \in \mathbf{X}$ and for any $0 < \lambda < 1$, we have :

$$f(\lambda x_1 + (1 - \lambda)x_2) \le \lambda f(x_1) + (1 - \lambda)f(x_2)$$

- Strictly convex if the inequality is strict.
- A convex function has only one global minimum.
- Sums of convex functions are also convex. The Hessian is related to the convexity of a function: a twice differentiable function is convex if and only if the Hessian is positive
- The Hessian of a convex function is positive semi-definite and for a strictly-convex function it is positive definite.
- The Hessian matrix of a function

$$\mathbf{H}_{i,j} = \frac{d^2 f}{dx_i dx_j}$$

1.2 Linear Algebra

- Column $\mathbf{x} \in \mathbb{R}^n$, rows \mathbf{x}^T , matrix $\mathbf{A} \in \mathbb{R}^{m \times n}$
- x^Tx is a scalar, xx^T is a matrix
- \mathbf{A}^{-1} exist if \mathbf{A} is full rank $(\mathbf{A}\mathbf{B})^T = \mathbf{B}^T \mathbf{A}^T$
- Condition number of a function measures how much the output value can change for a small change in the input argument. A matrix with a high condition number is said to be

ill-conditioned. If **A** is normal $(A^T A = AA^T)$

$$k(\mathbf{A}) = \left| \frac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

- A positive definite matrix is symmetric with all positive eigenvalues
- The real symmetric $N \times N$ matrix **V** is said to be positive semidefinite if

for any real $N \times 1$ vector a.

The real symmetric $N \times N$ matrix \mathbf{V} is said to be positive definite if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

for any real $N \times 1$ vector a.

- Cost of matrix inversion: $O(n^3) \rightarrow O(n^{2.372})$
- Cost of determinant computation using LU decomposition: $O(n^3)$

2 Cost functions

- Cost functions are used to learn parameters that explain the data well.
- It is essential to make sure that a global minimum exist -> lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{\beta}) = \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is convex thus it has only one global minumum value.
- MSE is not good when outliers are present

Mean Absolute Error (MAE)

$$MAE = \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

Huber loss

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \leq \delta \\ \delta|z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

– Huber loss is convex, differentiable, and also robust to outliers but hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Tukey's loss is non-convex, non-differentiable, but robust to outliers.

Hinge loss

$$Hinge = [1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$$
 Logistic loss

$$Logistic = \log(1 - \exp(y_n f(\mathbf{x}_n)))$$

3 Regression

- Data consists of N pairs (yn, xn)
- 1. y_n the n'th output
- 2. \mathbf{x}_n is a vector of D inputs
- Prediction: predict the ouput for a new input
- Interpretation: understand the effect of inputs on output.
- Outliers are data that are far away from most of the other examples.

3.1 Linear Regression

Model that assume linear relationship between inputs and the ouput.

$$y_n \equiv f(\mathbf{x}_n) := \beta_0 + \beta_1 x_{n1} + \dots = \beta_0 + \mathbf{x}_n^T \boldsymbol{\beta}$$
 with $\boldsymbol{\beta}$ the parameters of the model.

Variance grows only linearly with dimensionality

3.2 Gradient Descent

- Gradient descent uses only first-order information and takes steps in the direction of the gradient
- Given a cost function $\mathcal{L}(\boldsymbol{\beta})$ we wish to find $\boldsymbol{\beta}$ that minimizes the cost:

$$\min_{\beta} \mathcal{L}(\beta)$$

3.2.1 Grid search

- Compute the cost over a grid of M points to find the minimum
- Exponential Complexity $O(NDM^D)$
- Hard to find a good range of values

3.3 Batch Gradient Descent

- Take steps in the opposite direction of the

$$\boldsymbol{\beta}^{(k+1)} \leftarrow \boldsymbol{\beta}^{(k)} - \alpha \frac{d\mathcal{L}(\boldsymbol{\beta}^{(k)})}{d\boldsymbol{\beta}}$$

with $\alpha > 0$ the learning rate.

With α too big, method might diverge. With α too small, convergence is slow.

3.4 Gradients for MSE

$$\tilde{\mathbf{X}} = \begin{bmatrix} 1 & \mathbf{X} \end{bmatrix}$$

We define the error vector e:

$$\mathcal{L}(\boldsymbol{\beta}) = \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = -\frac{1}{N}\tilde{\mathbf{X}}^T\mathbf{e}$$

- Optimality conditions
- 1. necessary: gradient equal zero: $\frac{d\mathcal{L}(\boldsymbol{\beta}^*)}{d\boldsymbol{\beta}} = 0$
- 2. sufficient: Hessian matrix is positive definite: $\mathbf{H}(\boldsymbol{\beta}^*) = \frac{d^2 \mathcal{L}(\boldsymbol{\beta}^*)}{d^2 \mathcal{L}(\boldsymbol{\beta}^*)}$
- $d\beta d\beta^T$ Very sensitive to illconditioning. Therefore, always normalize your feature otherwise step-size selection is difficult since different directions might move at different speed.
- Complexity: O(NDI) with I the number of

3.5 Least Squares

- In some cases, we can compute the minimum of the cost function analytically.
- use the first optimality conditions:

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = 0 \Rightarrow \tilde{\mathbf{X}}^T \mathbf{e} = \tilde{\mathbf{X}}^T (\mathbf{y} - \tilde{\mathbf{X}}\boldsymbol{\beta}) = 0$$

When $\mathbf{\tilde{X}}^T\mathbf{\tilde{X}}$ is invertible, we have the

$$\boldsymbol{\beta}^* = (\mathbf{\tilde{X}}^T \mathbf{\tilde{X}})^{-1} \mathbf{\tilde{X}}^T \mathbf{y}$$

- thus we can predict values for a new x*

$$y_* = \tilde{\mathbf{x}}_*^{\mathbf{T}} \boldsymbol{\beta}^* = \tilde{\mathbf{x}}_*^{\mathbf{T}} (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

- The Gram matrix X^TX is positive definite and is also invertible iff $\tilde{\mathbf{X}}$ has full column rank. Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$
- $\tilde{\mathbf{X}}$ can be rank deficient when D > N or when the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. In this case, the matrix is ill-conditioned, leading to numerical issues.

3.6 Maximum Likelihood

Let define our mistakes ε_n ~ N(0, σ²).

$$\rightarrow y_n = \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} + \epsilon_n$$

Another way of expressing this:

$$\begin{aligned} p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta}) &= \prod_{n=1}^{N} p(y_n|\tilde{\mathbf{x}}_n, \boldsymbol{\beta}) \\ &= \prod_{n=1}^{N} \mathcal{N}(y_n|\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}, \sigma^2) \end{aligned}$$

which defines the likelihood of observating y given X and B

Define cost with log-likelihood

$$\mathcal{L}_{lik}(\boldsymbol{\beta}) = \log p(\mathbf{y}|\tilde{\mathbf{X}},\boldsymbol{\beta})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 + cnst$$

Maximum likelihood estimator (MLE) gives another way to design cost functions

$$lpha_{MSE}(oldsymbol{eta}) = lpha_{max} \mathcal{L}_{lik}(oldsymbol{eta})$$

- MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.
- With Laplace distribution

$$p(y_n | \tilde{\mathbf{x}}_n, \boldsymbol{\beta}) = \frac{1}{2b} e^{-\frac{1}{b} |y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}|}$$

$$\sum_{n} \log p(y_{n} | \mathbf{\tilde{x}}_{n}, \boldsymbol{\beta}) = \sum_{n} |y_{n} - \mathbf{\tilde{x}}_{n}^{T} \boldsymbol{\beta}| + cnst$$

3.7 Ridge Regression

- Linear models usually underfit. One way is to use nonlinear basis functions instead.

$$y_n = \beta_0 + \sum_{i=1}^{M} \beta_j \phi_j(\mathbf{x}_n) = \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{\beta}$$

- This model is linear in β but nonlinear in x. Note that the dimensionality is now M, not D.
- Polynomial basis $\phi(x_n) = [1, x_n, x_n^2, ..., x_n^M] \label{eq:phi}$ – The least square solution becomes

equare solution becomes
$$\boldsymbol{\beta}_{lse}^* = (\mathbf{\tilde{\Phi}}^T \mathbf{\tilde{\Phi}})^{-1} \mathbf{\tilde{\Phi}}^T \mathbf{y}$$

Complex models overfit easily. Thus we can choose simpler models by adding a regularization term which penalizes complex

$$\min_{\boldsymbol{\beta}} \left(\mathcal{L}(\boldsymbol{\beta}) + \frac{\lambda}{2N} \sum_{i=1}^{M} \beta_j^2 \right)$$

$$\boldsymbol{\beta}^* = \underset{\boldsymbol{\beta}}{\operatorname{argmin}} \left(\frac{1}{2} (\mathbf{y} - \mathbf{X} \boldsymbol{\beta})^T (\mathbf{y} - \mathbf{X} \boldsymbol{\beta}) + \frac{\lambda}{2} \boldsymbol{\beta}^T \boldsymbol{\beta} \right)$$

- Note that β₀ is not penalized. - By differentiating and setting to zero we get
 - $\boldsymbol{\beta}_{ridge} = (\tilde{\boldsymbol{\Phi}}^T \tilde{\boldsymbol{\Phi}} + \boldsymbol{\Lambda})^{-1} \tilde{\boldsymbol{\Phi}}^T \mathbf{y}$
- Ridge regression improves the condition number

of the Gram matrix since the eigenvalues of $(\mathbf{\tilde{\Phi}}^T\mathbf{\tilde{\Phi}} + \lambda I_m)$ are at least λ

- Maximum-a-posteriori (MAP) estimator:
- Maximizes the product of the likelihood and

$$\boldsymbol{\beta}_{MAP} = \underset{\boldsymbol{\beta}}{\operatorname{argmax}} \left(p(\mathbf{y}|\mathbf{X}, \boldsymbol{\Lambda}) p(\boldsymbol{\beta}|\boldsymbol{\Sigma}) \right)$$

- Assume $\beta_0 = 0$

$$oldsymbol{eta}_{ridge} = \mathop{\mathrm{argmax}}_{oldsymbol{eta}} \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T oldsymbol{eta}, oldsymbol{\Lambda}) \times \right] \right)$$

Lasso regularizer forces some β_i to be strictly 0 and therefore forces sparsity in the model.

$$\min_{\boldsymbol{\beta}} \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{\beta})^2, \quad \text{such that } \sum_{i=1}^{M} |\beta_i| \leq \text{where } \mathbf{S} \text{ is a } N \times N \text{ diagonal matrix with diagonals}$$

- We should choose λ to minimize the mistakes that will be made in the future.
- We split the data into train and validation sets and we pretend that the validation set is the future data. We fit our model on the training set and compute a prediction-error on the validation set. This gives us an estimate of the generalization error.
- K-fold cross validation randomly partition the data into K groups. We train on K-1groups and test on the remaining group. We repeat this until we have tested on all K sets. We then average the results.
- Cross-validation returns an unbiased estimate of the generalization error and its variance.

3.9 Bias-Variance decomposition

- The expected test error can be expressed as the sum of two terms
- Squared bias: The average shift of the predictions
- Variance: measure how data points vary around their average.

expected loss =
$$(bias)^2 + variance + noise$$

- Both model bias and estimation bias are
- Ridge regression increases estimation bias while
- reducing variance Increasing model complexity increases test error
- Small λ → low bias but large variance

- Large
$$\lambda \to \text{large bias but low variance}$$

 $err = \sigma^2 + E[f_{lse} - E[f_{lse}]]^2 + [f_{true} - E[f_{lse}]]^2$

3.10 Logistic Regression

Logistic function

- Classification relates input variables x to
- discrete output variable y Binary classifier: we use y = 0 for C_1 and
- Can use least-squares to predict ŷ*

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(y_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$

 $p(y_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$

- The probabilistic model:
$$p(\mathbf{y}|\mathbf{X},\boldsymbol{\beta}) = \prod_{j=1}^{N} \sigma(\tilde{\mathbf{x}}_{n}^{T}\boldsymbol{\beta})^{y_{n}} (1 - \sigma(\tilde{\mathbf{x}}_{n}^{T}\boldsymbol{\beta}))^{1-y_{n}}$$

- The log-likelihood:

$$\mathcal{L}_{MLE}(\boldsymbol{\beta}) = \sum_{n=1}^{N} \left(y_n \bar{\mathbf{x}}_n^T \boldsymbol{\beta} - \log(1 + \exp(\bar{\mathbf{x}}_n^T \boldsymbol{\beta})) \right)$$

- We can use the fact that $\frac{d}{dx}\log(1+\exp(x)) = \sigma(x)$

$$\begin{split} \mathbf{g} &= \frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \sum_{n=1}^{N} \left(\bar{\mathbf{x}}_n y_n - \bar{\mathbf{x}}_n \sigma(\bar{\mathbf{x}}_n^T \boldsymbol{\beta}) \right) \\ &= -\bar{\mathbf{X}}^T [\sigma(\bar{\mathbf{X}}\boldsymbol{\beta}) - \mathbf{y}] \end{split}$$

- The negative of the log-likelihood $-\mathcal{L}_{mle}(\boldsymbol{\beta})$ is
- Hessian of the log-likelihood

– We know that
$$\frac{d\sigma(t)}{dt} = \sigma(t)(1-\sigma(t))$$

Assume
$$\beta_0 = 0$$

$$\beta_{ridge} = \underset{\beta}{\operatorname{argmax}} \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T \boldsymbol{\beta}, \boldsymbol{\Lambda}) \times \mathcal{N}(\boldsymbol{\beta} | \mathbf{0}, \mathbf{1}) \right] \mathbf{F}(\boldsymbol{\beta}) = -\frac{d\mathbf{g}(\boldsymbol{\beta})}{d\boldsymbol{\beta}^T} = \sum_{n=1}^{N} \frac{d}{d\boldsymbol{\beta}^T} \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) \tilde{\mathbf{x}}_n$$
asso regularizer forces some β_i to be strictly and therefore forces sparsity in the model.

$$S_{nn} = \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})(1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

- The negative of the log-likelihood is not

Newton's Method

 Uses second-order information and takes steps in the direction that minimizes a quadratic approximation

$$\mathcal{L}(\boldsymbol{\beta}) = \mathcal{L}(\boldsymbol{\beta}^{(k)}) + \mathbf{g}_k^T(\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)}) + (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})^T \mathbf{H}_k(\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})$$

and it's minimum is at

$$\boldsymbol{\beta}^{k+1} = \boldsymbol{\beta}^{(k)} - \alpha_k \mathbf{H}_{\star}^{-1} \mathbf{g}_k$$

where \mathbf{g}_k is the gradient and α_k the learning rate.

- Complexity:
$$O((ND^2 + D^3)I)$$

Penalized Logistic Regression

$$\min_{\boldsymbol{\beta}} \left(-\sum_{n=1}^{N} \log p(y_n | \mathbf{x}_n^T \boldsymbol{\beta}) + \lambda \sum_{d=1}^{D} \beta_d^2 \right)$$

4 Generalized Linear Model

- Exponential family distribution
$$p(\mathbf{y}|\boldsymbol{\eta}) = \frac{h(y)}{z} \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

- Bernoulli distribution
$$p(y|\mu) = \mu^{y} (1-\mu)^{1-y}$$
$$= \exp(y \log(\frac{\mu}{1-\mu} + \log(1-\mu)))$$

there is a relationship between η and μ

throught the link function
$$\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$$

 Note that μ is the mean parameter of y Relationship between the mean μ and η is

defined using a link function
$$g$$

 $\eta = \mathbf{g}(\boldsymbol{\mu}) \Leftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta})$ First and second derivatives of $A(\eta)$ are related

to the mean and the variance
$$\frac{dA(\eta)}{d\eta} = E[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = Var[\phi(\eta)]$$

- $A(\eta)$ is convex - The generalized maximum likelihood cost to

mize is
$$\min_{\boldsymbol{\Omega}} \mathcal{L}(\boldsymbol{\beta}) = -\sum_{n=1}^{N} \log(p(y_n | \bar{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

where $p(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})$ is an exponential family

distribution - We obtain the solution

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \tilde{\mathbf{X}}^T [\mathbf{g}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\mathbf{y})]$$

5 k-Nearest Neighbor (k-NN) - The k-NN prediction for x is

 $f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$ where $nbh_k(\mathbf{x})$ is the neightborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n in the

Curse of dimensionality: Generalizing

correctly becomes exponentially harder as the

- dimensionality grows.
- Gathering more inputs variables may be a bad thing

6 Kernel Ridge Regression

- The following is true for ridge regression

$$\begin{split} \boldsymbol{\beta} &= (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y} \\ &= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \boldsymbol{\alpha} \end{split} \tag{1}$$

- Complexity of computing β : (1) $O(D^2N + D^3)$ (2) $O(DN^2 + N^3)$

$$\boldsymbol{\beta} = \sum_{n=1}^{N} \alpha_n \mathbf{x}_n, \ \mathbf{y} = \sum_{d=1}^{D} \beta_d \bar{\mathbf{x}}_d$$

with \mathbf{x}_n the rows of \mathbf{X} and $\bar{\mathbf{x}}_d$ the columns of \mathbf{X} The representer theorem allows us to write an equivalent optimization problem n terms of α .

$$\boldsymbol{\alpha} = \operatorname*{argmax}_{\boldsymbol{\alpha}} \left(-\frac{1}{2} \boldsymbol{\alpha} (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^T \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

- $\mathbf{K} = \mathbf{X}\mathbf{X}^T$ is called the **kernel matrix** or Gram matrix.
- If K is positive definite, then it's called a Mercer Kernel.
- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$ If the kernel is Mercer, then there exists a function $\phi(\mathbf{x})$ s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

- Kernel trick

- We can work directly with ${\bf K}$ and never have to worry about ${\bf X}$
- Replace (x, x') with k(x, x').
- Kernel function can be interpreted as a measure of similarity
- The evaluation of a kernel is usually faster with k than with ϕ
- Kernelized rigde regression might be
- computationally more efficient in some cases.
- Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Properties of a kernel to ensure the existance of a corresponding d:
- K should be symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$
- K should be positive semidefinite.

$$\mathbf{y} = \boldsymbol{\beta}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

7 Support Vector Machine

- Combination of the kernel trick plus a modified loss function (Hinge loss)
- Solution to the dual problem is sparse and non-zero entries will be our support vectors.
- Kernelised feature vector where μ_L are centroids

$$\phi(\mathbf{x}) = [k(\mathbf{x}, \boldsymbol{\mu}_1), ..., k(\mathbf{x}, \boldsymbol{\mu}_K)]$$

- In practice we'll take a subset of data points to be prototype → sparse vector machine.
- Assume $y_n \in \{-1, 1\}$ - SVM optimizes the following cost

$$g(\boldsymbol{\beta}) = \min_{\boldsymbol{\beta}} \, \sum_{n=1}^{N} [1 - y_n \, \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}]_+ + \frac{\lambda}{2} \, \sum_{j=1}^{M} \beta_j^2$$

- Minimum doesn't change with a rescaling of B choose the hyperplane so that the distance from it to the nearest data point on each side is maximized
- Duality
- Hard to minimize $g(\boldsymbol{\beta})$ so we define $g(\boldsymbol{\beta}) = \max_{\boldsymbol{\alpha}} G(\boldsymbol{\beta}, \boldsymbol{\alpha})$
- we use the property that
- $C[v_n]_+ = \max(0, Cv_n) = \max_{\alpha_n \in [0, C]} \alpha_n v_n$
- We can rewrite the problem as

$$\min_{\boldsymbol{\beta}} \max_{\boldsymbol{\alpha} \in [0,C]} \sum_{n=1}^{N} \alpha_n (1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}) + \frac{1}{2} \sum_{j=1}^{M} \hat{\boldsymbol{\beta}}_{j-1}^2$$

- This is differentiable, convex in \$\beta\$ and

- Minimax theorem:

$$\min_{\pmb{\beta}} \max_{\pmb{\alpha}} G(\pmb{\beta}, \pmb{\alpha}) = \max_{\pmb{\alpha}} \min_{\pmb{\beta}} G(\pmb{\beta}, \pmb{\alpha})$$

because G is convex in $oldsymbol{eta}$ and concave in $oldsymbol{lpha}$

$$\frac{dG}{d\boldsymbol{\beta}} = -\left(\sum_{n=1}^{N} \alpha y_n \tilde{\phi_n}\right) + \begin{bmatrix}0\\\boldsymbol{\beta}_{1:M}\end{bmatrix}$$

$$\boldsymbol{\beta}_{1:M}^* = \sum_{n=1}^N \alpha_n y_n \phi_n = \boldsymbol{\Phi}^T \operatorname{diag}(\mathbf{y}) \boldsymbol{\alpha}$$

- Plugging $\boldsymbol{\beta}^*$ back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,C]^N} {}^{N} \mathbf{1} - \frac{1}{2} \boldsymbol{\alpha}^T \mathbf{Y} \boldsymbol{\Phi} \boldsymbol{\Phi} \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{\Phi} \mathbf{\Phi}^T$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision

8 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances outside the cluster

$$\min_{\mathbf{r}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{r}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N r_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

- such that $r_{n\,k} \in \{0,1\}$ and $\sum_{k=1}^K r_{n\,k} = 1$ K-means algorithm:
- Initialize μ_k , then iterate
- 1. For all n, compute \mathbf{r}_n given μ

$$r_{nk} = \left\{ \begin{array}{ll} 1 & \text{ if } k = \mathop{\mathrm{argmin}}_{j} ||\mathbf{x}_{n} - \boldsymbol{\mu}||_{2}^{2} \\ 0 & \text{ otherwise} \end{array} \right.$$

$$\mu_k = \frac{\sum_{n=1}^{N} r_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} r_{nk}}$$

- A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{r}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{r_{nk}}$$

Computation can be heavy, each example can belong to only on cluster and clusters have to he spherical

Gaussian Mixture Models

Clusters can be spherical using a full covariance

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{r}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{r_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining r_n to be a random variable

$$p(r_{nk}=1)=\pi_k \text{ where } \pi_k>0, \forall k, \sum_{k=1}^K \pi_k=1$$

- Joint distribution of Gaussian mixture model

$$\begin{split} p(\mathbf{X}, \mathbf{r} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) &= \prod_{n=1}^{N} \left[p(\mathbf{x}_n | \mathbf{r}_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{r}_n | \boldsymbol{\pi}) \right] \\ &= \left[\prod_{k=1}^{K} \left[\left(\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right)^r n k \right] \prod_{k=1}^{K} \right] \end{split}$$

 r_n are called *latent* unobserved variables Unknown parameters are given by

We get the marginal likelihood by

marginalizing r_n out from the likelihood

$$\begin{split} p(\mathbf{x}_n|\boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, r_{nk} = 1|\boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(r_{nk} = 1|\boldsymbol{\theta}) p(\mathbf{x}_n|r_{nk} = 1, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{split}$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to $O(D^2K)$
- To get maximum likelihood estimate of θ , we

$$\max_{\pmb{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \pmb{\mu}_k, \pmb{\Sigma}_k)$$

10 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\boldsymbol{\theta}^{(i)}$ with equality when,

$$\gamma(r_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$

2. Maximization step: Update
$$\theta$$

$$\theta^{(i+1)} = \operatorname*{argmax}_{\theta} \mathcal{L}(\theta, \theta^{(i)})$$

$$\boldsymbol{\mu}_k^{(i+1)} = \frac{\sum_{n=1}^N \boldsymbol{\gamma}^{(i)}(\boldsymbol{r}_{nk}) \mathbf{x}_n}{\sum_{n=1}^N \boldsymbol{\gamma}^{(i)}(\boldsymbol{r}_{nk})}$$

$$\mathbf{\Sigma}_{k}^{(i+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i)})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^$$

$$\pi_k^{(i+1)} = \frac{1}{N} \sum_{n=1}^{N} \gamma^{(i)}(r_{nk})$$

If the covariance is diagonal, then we have K-means.

11 Matrix factorization

- We have D movies and N users
- ${\bf X}$ is a matrix $D \times N$ with x_{dn} the rating of n'th user for d'th movie.
- We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
 - **Z** a $N \times M$ matrix that gives features for the users
- W a $D \times M$ matrix that gives features for the movies

$$x_{dn} \approx \mathbf{w}_{d}^T \mathbf{z}_n$$

We can add a regularizer and minimize the

$$\begin{split} \mathcal{L}(\mathbf{W}, \mathbf{Z}) &= \frac{1}{2} \sum_{n=1}^{N} \sum_{d=1}^{D} (x_{dn} - \mathbf{w}_{d}^{T} \mathbf{z}_{n})^{2} \\ &+ \frac{\lambda w}{2} \sum_{d=1}^{D} \mathbf{w}_{d}^{T} \mathbf{w}_{d} + \frac{\lambda z}{2} \sum_{n=1}^{N} \mathbf{z}_{n}^{T} \mathbf{z}_{n} \end{split}$$

first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{z}^T \leftarrow (\mathbf{w}^T \mathbf{w} + \lambda_z \mathbf{I}_M)^{-1} \mathbf{w}^T \mathbf{x}$$

 $\mathbf{w}^T \leftarrow (\mathbf{z}^T \mathbf{z} + \lambda_w \mathbf{I}_M)^{-1} \mathbf{z}^T \mathbf{x}^T$

- Complexity: $O(DNM^2 + NM^3) \rightarrow O(DNM^2)$ - Probabilistic model

$$\begin{aligned} &) & = \prod_{n=1}^{N} \left[p(\mathbf{x}_{n} | \mathbf{r}_{n}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{r}_{n} | \boldsymbol{\pi}) \right] \\ & = \left[\prod_{k=1}^{K} \left[\left(\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right)^{r_{n} k} \right] \prod_{k=1}^{K} \prod_{d \in O_{n}}^{N} \mathcal{N}(x_{dn} | \mathbf{w}_{d}^{T} \mathbf{z}_{n}, I) \times \prod_{n=1}^{N} \mathcal{N}(\mathbf{z}_{n} | 0, \frac{1}{\lambda_{z}} I) \right] \\ & \text{Statest unobserved variables} \end{aligned}$$

Since many ratings are missing we cannot normalize the data. A solution is to add offset

$$\frac{1}{2} \sum_{n=1}^{N} \sum_{d \in O_n} (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu)^2$$

12 Singular Value Decomposition

- Matrix factorization method

$$\mathbf{X} = \mathbf{U}\mathbf{S}\mathbf{V}^T$$

- \mathbf{U} is an $D \times D$ matrix
- **V** is an $N \times N$ matrix
- S is a non-negative diagonal matrix of size $D \times N$ which are called singular values appearing in a descending order.
- Columns of U and V are the left and right singular vectors respectively.
- Assuming D < N we have

$$\mathbf{X} = \sum_{d=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information

Let's now truncate these matrices

$$\mathbf{X} pprox \mathbf{U}_{tr} \mathbf{S}_{tr} \mathbf{V}_{tr}^*$$
 $\mathbf{T}_{tr} pprox \mathbf{U}_{tr} \mathbf{S}_{tr}, \mathbf{T}_{te} \qquad pprox \mathbf{X}_{te} \mathbf{V}_{tr}$

with \mathbf{T}_{tr} the reduced feature set of \mathbf{X} and \mathbf{T}_{te} the reduced feature set of \mathbf{X}_{te}

13 Principal Componement Analysis

- PCA is a dimensionality reduction method and a method to decorrelate the data $\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$ such that columns of \mathbf{W} are
- orthogonal. If the data is zero mean

$$\boldsymbol{\Sigma}_{k}^{(i+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})}{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})^{T}} = \frac{1}{N} \mathbf{X} \mathbf{X}^{T} \Rightarrow \mathbf{X} \mathbf{X}^{T} = \mathbf{U} \mathbf{S}^{2} \mathbf{U}^{T}$$

$$\boldsymbol{\Sigma}_{k}^{(i+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})^{T}}{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})} \Rightarrow \mathbf{U}^{T} \mathbf{X} \mathbf{X}^{T} \mathbf{U} = \mathbf{U}^{T} \mathbf{U} \mathbf{S}^{2} \mathbf{U}^{T} \mathbf{U} = \mathbf{S}^{2}$$

- Thus the columns of matrix U are called the principal components and they decorrelate the covariance matrix.
- Using SVD, we can compute the matrices in the following way

$$W = US^{1/2}, Z = VS^{1/2}$$

14 Belief Propagation

- the goal is to learn inference of the latent variables using belief propagation
- Given a directed acyclic graph G and parameters θ , a Bayesian network defines the joint distribution as follows

$$p_{\theta}(\mathbf{x}) = \prod_{k=1}^{K} p_{\theta}(x_k | \text{parents}_k)$$

We can reduce the complexity by using the structure of the problem and the bayesian rule

15 Multi-Layer Perceptron (MLP)

- There are nonlinear function classes, whose conveninent layered structure leads to efficient
- computation of gradients. Known as feed-forward neural network
- $-\mathbf{z}_{n}^{(k)}$ is the k'th hidden vector $-\mathbf{a}_{n}^{(k)}$ is the corresponding activation

There are a total of K layers
$$a_{N}^{(k)} = (\boldsymbol{\beta}_{N}^{(k)})^{T} \mathbf{z}_{N}^{(k-1)}, \quad z_{N}^{(k)} = h(a_{N}^{(k)})$$

- For the first layer we have $\mathbf{z}_n^{(0)} = \mathbf{x}_n$ For the last layer, we use a link function to map
- $\mathbf{z}_{n}^{(K-1)}$ to the output \mathbf{y}_{n} A 1-layer MLP is simply a generalization of
- linear/logistic regression $\mathbf{B}^{(k)}$ a matrix with rows $(\boldsymbol{\beta}_m^{(k)})^T$

$$\mathbf{a}_n^{(k)} = \mathbf{B}^{(k)} \mathbf{z}_n^{(k-1)}, \quad \mathbf{z}_n^{(k)} = h(\mathbf{a}_n^{(k)})$$
- thus we have the input-output relationship

- $\hat{y}_n = g((\boldsymbol{\beta}^{(K-1)})^T * h(\mathbf{B}^{(K-2)}) * h(*\dots * h(\mathbf{B}^{(1)})) = \mathbf{z}_n^2$ techniques for decorrelating trees: with g() the link function with g() the link function
- We learn parameters B using stochastic

- Frequently used transfer function

$$g(a) = \tanh(a) = \frac{e^a - e^-}{a + a^-}$$

- Backpropagation is a technic to compute the gradient in time linear in the number of training points and the number of weights and is nothing else than a chain rule of differential calculus
- The key-idea is to express the derivatives in terms of activations and hidden variables.
- 1. Forward: compute $\mathbf{a}_n^{(k)}$ and $\mathbf{z}_n^{(k)}$
- 2. Backward: compute $\delta_n^{(k)} = d\mathcal{L}/d\mathbf{a}_n^{(k)}$ using $\boldsymbol{\delta}_{n}^{(k-1)} = diag[(\mathbf{h}'(a_{n}^{(k)})](\mathbf{B}^{(k)})^{T} \boldsymbol{\delta}_{n}^{(k)}$
- 3. Compute $d\mathcal{L}/d\mathbf{B}^{(k)}$ using

$$\frac{d\mathcal{L}}{d\mathbf{B}^{(k)}} = \sum_{n} \delta_{n}^{(k)} (\mathbf{z}_{n}^{(k)})^{T}$$

16 Gaussian Process (GP)

- GP are a family of statistical distributions in which time plays a role and for which any finite linear combination of samples has a joint Gaussian distribution
- non-parametric method (rather uses latent variables) that compute a probability dist. over predictions
- They can be completely defined by their second-order statistics which means that if they have mean zero, then defining the covariance function completely defines the process hehaviour
- Let us place a probabilistic prior shape on the approximation of a function.
- A GP process defines a prior over function f

$$p(\mathbf{f}|X) = \mathcal{N}(\mathbf{f}|\mathbf{0}, K(X))$$

K(X) defines shape and prior knowledge about our problem

$$\begin{split} p(\mathbf{f}|\mathbf{y}, \mathbf{X}_*, \mathbf{X}) &\sim \mathcal{N}(\mu', \sigma') \\ \mu' &= K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]^{-1} \mathbf{y} \\ \sigma' &= K(\mathbf{X}_*, \mathbf{X}_*) - K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]K(\mathbf{X}, \mathbf{X}) \end{split}$$

with σ_n the variance of the noise

Quadratic kernel
$$k(\mathbf{x}_n, \mathbf{x}_m) = (1 + \mathbf{x}_n^T \mathbf{x}_m)^2$$

17 Decision Trees (DT)

- A decision tree is a structure in which each internal node represents a test on an attribute and each branch represents the outcome of the test and each leaf represents a class label.
- Fast to train and fast to make predictions Efficient for very high dimensional feature
- spaces and very large amounts of training data Lack of smoothness and high variance (overfitting)
 - Goal: find a split (k, τ) that minimizes an impurity measure at the leaves - Find best feature to split on

Variance of the model averaging:

- Find best threshold

- 18 Random Forests (RF)
- RF correct the overfitting bad "habit" of DTs. Training: Learn M trees on different subsets (random) of training data
- Prediction: Average of prediction of each tree

$$z_1 = f_1 \quad \Rightarrow \quad z_M = \frac{1}{M} \sum_{i=1}^{M} f_i$$

$$V(z_1) = \sigma^2 \quad \Rightarrow \quad V(z_M) = \frac{1}{100}\sigma^2 + \rho \frac{M-1}{100}\sigma^2$$

Variance reduction ratio:

$$\frac{V(z_1)}{V(z_M)} = \frac{M}{1 + \rho(M-1)}$$

- - Randomized feature selection