PCML Cheat Sheet

1 Math Prerequisites

- Bayes rule

$$p(A,B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

Gaussian distribution

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}))$$
Huber loss is converged to contain the property of the contains and the contains are set as

$$V(XY) = E(X^{2})E(Y^{2}) - [E(X)]^{2}[E(Y)]^{2}$$

- Log-properties

$$\log(mn) = \log(m) + \log(n)$$

$$\log(m^n) = n \log(m)$$

- Covariance matrix of a data vector x

$$\Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - E(\mathbf{x})(\mathbf{x}_n - E(\mathbf{x})^T)$$

- A function is convex when a line joining two points never interects with the function anywhere else
- A function f(x) is convex, if for any $x_1, x_2 \in \mathbf{X}$ and for any $0 < \lambda < 1$, we have :

$$f(\lambda x_1 + (1-\lambda)x_2) \leq \lambda f(x_1) + (1-\lambda)f(x_2)$$

- A function is strictly convex if the inequality is
- A convex function has only one global minimum. Sums of convex functions are also convex.
- The Hessian is related to the convexity of a function: a twice differentiable function is convex if and only if the Hessian is positive
- The Hessian of a convex function is positive semi-definite and for a strictly-convex function it is positive definite.
- The Hessian matrix of a function

$$\mathbf{H}_{i,j} = \frac{d^2f}{dx_i dx_j}$$

1.2 Linear Algebra

- Column $\mathbf{x} \in \mathbb{R}^n$, rows \mathbf{x}^T , matrix $\mathbf{A} \in \mathbb{R}^{m \times n}$
- $-\mathbf{x}^T\mathbf{x}$ is a scalar, $\mathbf{x}\mathbf{x}^T$ is a matrix
- \mathbf{A}^{-1} exist if \mathbf{A} is full rank $(\mathbf{A}\mathbf{B})^T = \mathbf{B}^T \mathbf{A}^T$
- Condition number of a function measures how much the output value can change for a small change in the input argument. A matrix with a high condition number is said to be

ill-conditioned. If **A** is normal $(A^T A = AA^T)$

$$k(\mathbf{A}) = \left| \frac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

- A positive definite matrix is symmetric with all positive eigenvalues
- The real symmetric $N \times N$ matrix **V** is said to be positive semidefinite if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} \ge 0$$
 for any real $N \times 1$ vector a .

The real symmetric $N \times N$ matrix **V** is said to be positive definite if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

for any real $N \times 1$ vector a.

- Cost of matrix inversion: $O(n^3) \rightarrow O(n^{2.372})$
- Cost of determinant computation using LU decomposition: $O(n^3)$

2 Cost functions

- Cost functions are used to learn parameters that explain the data well.
- It is essential to make sure that a global minimum exist → lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{\beta}) = \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is convex thus it has only one global minumum value.
- MSE is not good when outliers are present

Mean Absolute Error (MAE):

$$MAE = \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

Huber los

$$Huber = \left\{ \begin{array}{cc} \frac{1}{2}z^2 & , |z| \leq \delta \\ \delta|z| - \frac{1}{2}\delta^2 & , |z| > \delta \end{array} \right.$$

Huber loss is convex, differentiable, and also hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Tukey's loss is non-convex, non-differentiable, but robust to outliers.

 $Hinge = [1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$

Logistic loss

$$Logistic = \log(1 - \exp(y_n f(\mathbf{x}_n)))$$

3 Regression

- Data consists of N pairs (un. xn) 1. y_n the n'th output
- 2. \mathbf{x}_n is a vector of D inputs
- Prediction: predict the ouput for a new input
- Interpretation: understand the effect of inputs on output.
- Outliers are data that are far away from most of the other examples

3.1 Linear Regression

Model that assume linear relationship between inputs and the ouput

$$y_n \equiv f(\mathbf{x}_n)$$

$$:= \beta_0 + \beta_1 x_{n1} + \dots$$

$$= \beta_0 + \mathbf{x}_n^T \boldsymbol{\beta}$$

- with β the parameters of the model.
- Variance grows only linearly with dimensionality

3.2 Gradient Descent

- Gradient descent uses only first-order information and takes steps in the direction of
- Given a cost function $\mathcal{L}(\boldsymbol{\beta})$ we wish to find $\boldsymbol{\beta}$ that minimizes the cost:

$$\min_{\beta} \mathcal{L}(\beta)$$

3.2.1 Grid search

- Compute the cost over a grid of M points to find the minimum
- Exponential Complexity $O(NDM^D)$
- Hard to find a good range of values

3.3 Batch Gradient Descent

- Take steps in the opposite direction of the

$$\boldsymbol{\beta}^{(k+1)} \leftarrow \boldsymbol{\beta}^{(k)} - \alpha \frac{d\mathcal{L}(\boldsymbol{\beta}^{(k)})}{d\boldsymbol{\beta}}$$

with $\alpha > 0$ the learning rate.

With α too big, method might diverge. With α too small, convergence is slow.

3.4 Gradients for MSE

$$\tilde{\mathbf{X}} = \begin{bmatrix} 1 & \mathbf{X} \end{bmatrix}$$

We define the error vector e:

$$\mathbf{e} = \mathbf{y} - \mathbf{\tilde{X}} \boldsymbol{\beta}$$

and MSE as follows:

$$\mathcal{L}(\boldsymbol{\beta}) = \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = -\frac{1}{N}\tilde{\mathbf{X}}^T\mathbf{e}$$

- Optimality conditions:

1. necessary: gradient equal to zero:

2. sufficient: Hessian matrix is positive definite: $\mathbf{H}(\boldsymbol{\beta}^*) = \frac{d^2\mathcal{L}(\boldsymbol{\beta}^*)}{}$ dBdBT

- Very sensitive to illconditioning. Therefore, always normalize your feature otherwise step-size selection is difficult since different directions might move at different speed.
- Complexity: O(NDI) with I the number of

3.5 Least Squares

- In some cases, we can compute the minimum of the cost function analytically.
- use the first optimality conditions:

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = 0 \Rightarrow \tilde{\mathbf{X}}^T \mathbf{e} = \tilde{\mathbf{X}}^T (\mathbf{y} - \tilde{\mathbf{X}} \boldsymbol{\beta}) = 0$$

When $\mathbf{\tilde{X}}^T\mathbf{\tilde{X}}$ is invertible, we have the closed-form expression

$$\boldsymbol{\beta}^* = (\mathbf{\tilde{X}}^T \mathbf{\tilde{X}})^{-1} \mathbf{\tilde{X}}^T \mathbf{y}$$

- thus we can predict values for a new x_{*}

$$y_* = \tilde{\mathbf{x}}_*^{\mathbf{T}} \boldsymbol{\beta}^* = \tilde{\mathbf{x}}_*^{\mathbf{T}} (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

- The Gram matrix $\tilde{\mathbf{X}}^T \tilde{\mathbf{X}}$ is positive definite and is also invertible iff $\tilde{\mathbf{X}}$ has full column rank.
- Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$ $\tilde{\mathbf{X}}$ can be rank deficient when D > N or when the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. In this case, the matrix is ill-conditioned, leading to

3.6 Maximum Likelihood

numerical issues

Let define our mistakes ε_n ~ N(0, σ²).

$$\rightarrow y_n = \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} + \epsilon_n$$

- Another way of expressing this:

$$p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta}) = \prod_{n=1}^{N} p(y_n | \tilde{\mathbf{x}}_n, \boldsymbol{\beta})$$
$$= \prod_{n=1}^{N} \mathcal{N}(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}, \sigma^2)$$

which defines the likelihood of observating y given X and B

Define cost with log-likelihood

$$\mathcal{L}_{lik}(\boldsymbol{\beta}) = \log p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 + cnst$$

another way to design cost functions

$$\underset{\boldsymbol{\beta}}{\operatorname{argmin}} \mathcal{L}_{MSE}(\boldsymbol{\beta}) = \underset{\boldsymbol{\beta}}{\operatorname{argmax}} \mathcal{L}_{lik}(\boldsymbol{\beta})$$

- MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.
 - With Laplace distribution $p(y_n|\tilde{\mathbf{x}}_n,\boldsymbol{\beta}) = \frac{1}{2b}e^{-\frac{1}{b}|y_n - \tilde{\mathbf{x}}_n^T\boldsymbol{\beta}|}$

$$\sum \log p(y_n | \tilde{\mathbf{x}}_n, \boldsymbol{\beta}) = \sum |y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}| + cnst$$

3.7 Ridge Regression

- Linear models usually underfit. One way is to use nonlinear basis functions instead

$$y_n = \beta_0 + \sum_{j=1}^M \beta_j \phi_j(\mathbf{x}_n) = \tilde{\phi}(\mathbf{x}_n)^T \boldsymbol{\beta}$$

- This model is linear in β but nonlinear in x.
- Note that the dimensionality is now M, not D. Polynomial basis

$$m{\phi}(x_n) = [1, x_n, x_n^2, ..., x_n^M]$$
 – The least square solution becomes

 $\boldsymbol{\beta}_{lse}^* = (\mathbf{\tilde{\Phi}}^T \mathbf{\tilde{\Phi}})^{-1} \mathbf{\tilde{\Phi}}^T \mathbf{v}$ Complex models overfit easily. Thus we can choose simpler models by adding a regularization term which penalizes complex

$$\min_{\boldsymbol{\beta}} \left(\mathcal{L}(\boldsymbol{\beta}) + \frac{\lambda}{2N} \sum_{j=1}^{M} \beta_j^2 \right)$$

$$\boldsymbol{\beta}^* = \operatorname*{argmin}_{\boldsymbol{\beta}} \left(\frac{1}{2} (\mathbf{y} - \mathbf{X} \boldsymbol{\beta})^T (\mathbf{y} - \mathbf{X} \boldsymbol{\beta}) + \frac{\lambda}{2} \boldsymbol{\beta}^T \boldsymbol{\beta} \right)$$

Note that β_0 is not penalized.

- By differentiating and setting to zero we get

$$oldsymbol{eta}_{ridge} = (\mathbf{\tilde{\Phi}}^T \mathbf{\tilde{\Phi}} + \mathbf{\Lambda})^{-1} \mathbf{\tilde{\Phi}}^T \mathbf{y}$$

$$\mathbf{\Lambda} = \begin{bmatrix} 0 & 0 \\ 0 & \lambda I_m \end{bmatrix}$$

- Ridge regression improves the condition number of the Gram matrix since the eigenvalues of $({ ilde \Phi}^T { ilde \Phi} + \lambda I_m)$ are at least λ
- Maximum-a-posteriori (MAP) estimator:
 - Maximizes the product of the likelihood and the prior.

$$\boldsymbol{\beta}_{MAP} = \operatorname*{argmax}_{\boldsymbol{\beta}} \left(p(\mathbf{y}|\mathbf{X}, \boldsymbol{\Lambda}) p(\boldsymbol{\beta}|\boldsymbol{\Sigma}) \right)$$

- Assume $\beta_0 = 0$

$$\beta_{ridge} = \underset{\beta}{\operatorname{argmax}} \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T \boldsymbol{\beta}, \boldsymbol{\Lambda}) \right] \right)$$

Lasso regularizer forces some β_i to be strictly 0 and therefore forces sparsity in the model.

$$\min_{\boldsymbol{\beta}} \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{\beta})^2, \quad \text{such that } \sum_{i=1}^{M} |\beta_i| \leq \tau$$

- We should choose λ to minimize the mistakes. that will be made in the future.
- We split the data into train and validation sets and we pretend that the validation set is the future data. We fit our model on the training set and compute a prediction-error on the validation set. This gives us an estimate of the generalization error.
- K-fold cross validation randomly partition the data into K groups. We train on K-1groups and test on the remaining group. We repeat this until we have tested on all K sets. We then average the results.
- Cross-validation returns an unbiased estimate of the generalization error and its variance.

- 3.9 Bias-Variance decomposition - The expected test error can be expressed as the
- sum of two terms - Squared bias: The average shift of the
- predictions - Variance: measure how data points vary around their average.

expected loss = $(bias)^2$ + variance + noise

- Both model bias and estimation bias are
- important Ridge regression increases estimation bias while
- reducing variance - Increasing model complexity increases test error

Small $\lambda \to \text{low}$ bias but large variance

Large $\lambda \rightarrow$ large bias but low variance

3.10 Logistic Regression

- Classification relates input variables x to
- discrete output variable y Binary classifier: we use y = 0 for C_1 and
 - y = 1 for C_2 . Can use least-squares to predict \hat{y}_* $\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* > 0.5 \end{cases}$

Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(y_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$

 $p(y_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$ - The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \boldsymbol{\beta}) = \prod_{n=1}^{N} \sigma(\tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta})^{y_{n}} (1 - \sigma(\tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta}))^{1-y_{n}}$$

- The log-likelihood:

$$\mathcal{L}_{MLE}(\boldsymbol{\beta}) = \sum_{n=1}^{N} \left(y_n \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} - \log(1 + \exp(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})) \right)$$

- We can use the fact that

$$\frac{d}{dx}\log(1+\exp(x)) = \sigma(x)$$

$$\begin{split} \mathbf{g} &= \frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \sum_{n=1}^{N} \left(\hat{\mathbf{x}}_{n} y_{n} - \hat{\mathbf{x}}_{n} \sigma(\hat{\mathbf{x}}_{n}^{T} \boldsymbol{\beta}) \right) \\ &= - \tilde{\mathbf{X}}^{T} [\sigma(\tilde{\mathbf{X}} \boldsymbol{\beta}) - \mathbf{y}] \end{split}$$

- The negative of the log-likelihood $-\mathcal{L}_{mle}(\boldsymbol{\beta})$ is
- Hessian of the log-likelihood

- We know that

$$\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t))$$

Assume $\beta_0 = 0$ $\beta_{ridge} = \underset{\beta}{\operatorname{argmax}} \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T \boldsymbol{\beta}, \boldsymbol{\Lambda}) \times \mathcal{N}(\boldsymbol{\beta} | \mathbf{0}, \mathbf{I}) \right] \right) + (\boldsymbol{\beta}) = -\frac{d\mathbf{g}(\boldsymbol{\beta})}{d\boldsymbol{\beta}^T} = \sum_{n=1}^{N} \frac{d}{d\boldsymbol{\beta}^T} \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) \tilde{\mathbf{x}}_n$

$$= \sum_{n=1}^{N} \bar{\mathbf{x}}_n \sigma(\bar{\mathbf{x}}_n^T \boldsymbol{\beta}) (1 - \sigma(\bar{\mathbf{x}}_n^T \boldsymbol{\beta})) \bar{\mathbf{x}}_n^T$$

where **S** is a $N \times N$ diagonal matrix with

$$S_{nn} = \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})(1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$
- The negative of the log-likelihood is not

strictly convex.

Newton's Method Uses second-order information and takes steps in the direction that minimizes a

quadratic approximation
$$\mathcal{L}(\boldsymbol{\beta}) = \mathcal{L}(\boldsymbol{\beta}^{(k)}) + \mathbf{g}_k^T(\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)}) \\ + (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})^T \mathbf{H}_k(\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})$$

and it's minimum is at

 $\boldsymbol{\beta}^{k+1} = \boldsymbol{\beta}^{(k)} - \alpha_k \mathbf{H}_{l}^{-1} \mathbf{g}_{k}$ where \mathbf{g}_k is the gradient and α_k the

learning rate.

$$\begin{aligned} & - \text{ Complexity: } O((ND^2 + D^3)I) \\ & \textbf{Penalized Logistic Regression} \\ & \min_{\pmb{\beta}} \left(- \sum_{n=1}^{N} \log p(y_n | \mathbf{x}_n^T \pmb{\beta}) + \lambda \sum_{d=1}^{D} \beta_d^2 \right) \end{aligned}$$

4 Generalized Linear Model

- Exponential family distribution
$$p(\mathbf{y}|\boldsymbol{\eta}) = \frac{h(y)}{-} \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

- Bernoulli distribution
$$p(y|\mu) = \mu^{y} (1 - \mu)^{1-y}$$
$$= \exp(y \log(\frac{\mu}{1-\mu} + \log(1 - \mu)))$$

- there is a relationship between η and μ throught the link function

ught the link function
$$\eta = \log(\frac{\mu}{1 - \mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1 + e^{\eta}}$$

 Note that μ is the mean parameter of y Relationship between the mean μ and η is defined using a link function q

$$\begin{split} \eta &= \mathbf{g}(\boldsymbol{\mu}) \Leftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta}) \\ \text{First and second derivatives of } A(\boldsymbol{\eta}) \text{ are related to the mean and the variance} \\ \frac{dA(\boldsymbol{\eta})}{d\boldsymbol{\eta}} &= E[\boldsymbol{\phi}(\boldsymbol{\eta})], \quad \frac{d^2A(\boldsymbol{\eta})}{d\boldsymbol{\eta}^2} = Var[\boldsymbol{\phi}(\boldsymbol{\eta})] \end{split}$$

- $A(\eta)$ is convex - The generalized maximum likelihood cost to

$$\min_{\boldsymbol{\beta}} \mathcal{L}(\boldsymbol{\beta}) = -\sum_{n=1}^{N} \log(p(y_n | \bar{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

where $p(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})$ is an exponential family

- We obtain the solution

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \tilde{\mathbf{X}}^T [\mathbf{g}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\mathbf{y})]$$

5 k-Nearest Neighbor (k-NN)

- The k-NN prediction for x is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$$

where $nbh_k(\mathbf{x})$ is the neightborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n in the training data

- Curse of dimensionality: Generalizing correctly becomes exponentially harder as the dimensionality grows.
- Gathering more inputs variables may be a bad

6 Kernel Ridge Regression

- The following is true for ridge regression

$$\begin{split} \boldsymbol{\beta} &= (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y} \\ &= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \boldsymbol{\alpha} \end{split} \tag{1}$$

- Complexity of computing \$\beta\$
- (1) $O(D^2N + D^3)$
- (2) $O(DN^2 + N^3)$

$$\boldsymbol{\beta} = \sum_{n=1}^{N} \alpha_n \mathbf{x}_n, \quad \mathbf{y} = \sum_{d=1}^{D} \beta_d \bar{\mathbf{x}}_d$$

with \mathbf{x}_n the rows of \mathbf{X} and $\mathbf{\bar{x}}_d$ the columns of \mathbf{X} The representer theorem allows us to write an equivalent optimization problem n terms of α .

$$\boldsymbol{\alpha} = \operatorname*{argmax}_{\boldsymbol{\alpha}} \left(-\frac{1}{2} \boldsymbol{\alpha} (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^T \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

- $\mathbf{K} = \mathbf{X}\mathbf{X}^T$ is called the **kernel matrix** or Gram matrix.
- If K is positive definite, then it's called a
- Mercer Kernel.
- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$
- If the kernel is Mercer, then there exists a function $\phi(\mathbf{x})$ s.t.

$$k(\mathbf{x}, \mathbf{x'}) = \phi(\mathbf{x})^T \phi(\mathbf{x'})$$

Kernel trick:

- We can work directly with K and never have to worry about X
- Replace $\langle \mathbf{x}, \mathbf{x}' \rangle$ with $k(\mathbf{x}, \mathbf{x}')$ - Kernel function can be interpreted as a
- measure of similarity The evaluation of a kernel is usually faster
- with k than with ϕ
- Kernelized rigde regression might be
- computationally more efficient in some cases.

Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Properties of a kernel to ensure the existance of a corresponding ϕ :
- K should be symmetric: k(x, x') = k(x', x) - K should be positive semidefinite.

$$\mathbf{y} = \boldsymbol{\beta}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

7 Support Vector Machine

- Combination of the kernel trick plus a modified loss function (Hinge loss)
- Solution to the dual problem is sparse and
- non-zero entries will be our support vectors. Kernelised feature vector where μ_k are

centroids
$$\phi(\mathbf{x}) = [k(\mathbf{x}, \mu_1), ..., k(\mathbf{x}, \mu_K)]$$

- In practice we'll take a subset of data points to be prototype -> sparse vector machine.
- Assume $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$g(\boldsymbol{\beta}) = \min_{\boldsymbol{\beta}} \sum_{n=1}^{N} [1 - y_n \, \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}]_+ + \frac{\lambda}{2} \sum_{j=1}^{M} \beta_j^2$$

- The minimum doesn't change with a rescaling
- Duality

Hard to minimize g(β) so we define

$$g(\boldsymbol{\beta}) = \max_{\boldsymbol{\alpha}} G(\boldsymbol{\beta}, \boldsymbol{\alpha})$$

- we use the property that

$$C[v_n]_+ = \max(0, Cv_n) = \max_{\alpha_n \in [0, C]} \alpha_n v_n$$

$$\min_{\boldsymbol{\beta}} \max_{\boldsymbol{\alpha} \in [0,C]^N} \sum_{n=1}^N \alpha_n (1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}) + \frac{1}{2} \sum_{j=1}^M \hat{\boldsymbol{\beta}}_j^2 - \frac{1}{2} \sum_{j=1}^M \hat{\boldsymbol$$

- This is differentiable, convex in β and concave in α
- Minimax theorem:

$$\min_{\beta} \max_{\alpha} G(\beta, \alpha) = \max_{\alpha} \min_{\beta} G(\beta, \alpha)$$

- because G is convex in β and concave in α .

$$rac{dG}{doldsymbol{eta}} = -\left(\sum_{n=1}^{N} lpha y_n \, ilde{\phi_n}
ight) + \left[egin{array}{c} 0 \ oldsymbol{eta}_{1:M} \end{array}
ight]$$

$$\boldsymbol{\beta}_{1:M}^* = \sum_{n=1}^N \alpha_n y_n \phi_n = \boldsymbol{\Phi}^T \operatorname{diag}(\mathbf{y}) \boldsymbol{\alpha}$$

$$\alpha^T y = 0$$

- Plugging $\boldsymbol{\beta}^*$ back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,C]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2} \boldsymbol{\alpha}^T \mathbf{Y} \boldsymbol{\Phi} \boldsymbol{\Phi} \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{\Phi} \mathbf{\Phi}^T$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision boundary.

8 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances

$$\min_{\mathbf{r}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{r}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N r_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

- such that $r_{n\,k} \in \{0,1\}$ and $\sum_{k=1}^K r_{n\,k} = 1$ K-means algorithm:
- Initialize μ_k , then iterate
- 1. For all n, compute \mathbf{r}_n given $\boldsymbol{\mu}$
 - $r_{nk} = \left\{ \begin{array}{ll} 1 & \text{if } k = \operatornamewithlimits{argmin}_{j} ||\mathbf{x}_{n} \boldsymbol{\mu}||_{2}^{2} \\ 0 & \text{otherwise} \end{array} \right.$

2. For all k, compute μ_k given \mathbf{r}

$$\mu_k = \frac{\sum_{n=1}^{N} r_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} r_{nk}}$$

- A good initialization procedure is to choose the prototypes to be equal to a random subset of Kdata points.
- Probabilistic model

$$p(\mathbf{r}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{r_{nk}}$$

Computation can be heavy, each example can belong to only on cluster and clusters have to he spherical

9 Gaussian Mixture Models

Clusters can be spherical using a full covariance matrix instead of isotropic covariance.

$$p(\mathbf{X}|\boldsymbol{\mu},\boldsymbol{\Sigma},\mathbf{r}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k},\boldsymbol{\Sigma}_{k}) \right]^{r_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining r_n to be a random variable $p(r_{nk} = 1) = \pi_k$ where $\pi_k > 0, \forall k$

$$\sum_{k=1}^{K} \pi_k = 1$$

Joint distribution of Gaussian mixture model

$$\begin{aligned} p(\mathbf{X}, \mathbf{r} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) &= \prod_{n=1}^{N} \left[p(\mathbf{x}_{n} | \mathbf{r}_{n}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{r}_{n} | \boldsymbol{\pi}) \right] & \mathbf{Z}^{T} \leftarrow (\mathbf{W}^{T} \mathbf{W} + \lambda_{z} \mathbf{I}_{M})^{-1} \mathbf{W}^{T} \mathbf{X} \\ \mathbf{W}^{T} \leftarrow (\mathbf{Z}^{T} \mathbf{Z} + \lambda_{w} \mathbf{I}_{M})^{-1} \mathbf{Z}^{T} \mathbf{X}^{T} \\ &= \begin{bmatrix} K \\ \prod_{k=1}^{K} \left[(\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}))^{T} n k \right] \prod_{k=1}^{K} \prod_{j=1}^{T} Probabilistic model \\ N \end{bmatrix} \begin{bmatrix} \mathbf{Z}^{T} \leftarrow (\mathbf{W}^{T} \mathbf{W} + \lambda_{z} \mathbf{I}_{M})^{-1} \mathbf{W}^{T} \mathbf{X} \\ \mathbf{W}^{T} \leftarrow (\mathbf{Z}^{T} \mathbf{Z} + \lambda_{w} \mathbf{I}_{M})^{-1} \mathbf{Z}^{T} \mathbf{X}^{T} \end{bmatrix} \end{aligned}$$

- rn are called latent unobserved variables Unknown parameters are given by $\theta = \{\mu, \Sigma, \pi\}$

- We get the marginal likelihood by marginalizing r_n out from the likelihood

$$\begin{split} p(\mathbf{x}_n|\boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, r_{nk} = 1|\boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(r_{nk} = 1|\boldsymbol{\theta}) p(\mathbf{x}_n|r_{nk} = 1, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{split}$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to $O(D^2K)$
- To get maximum likelihood estimate of θ , we

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

10 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate

1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\theta^{(i)}$ with equality when,

$$\gamma(r_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$

2. Maximization step: Update θ

$$\boldsymbol{\theta}^{\left(i+1\right)} = \operatorname*{argmax}_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{\left(i\right)})$$

$$\mu_k^{(i+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk}) \mathbf{x}_n}{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})}$$

$$\boldsymbol{\Sigma}_{k}^{(i+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})(\mathbf{x}_{n}}{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk})} - \frac{\mathbf{X} \approx \hat{\mathbf{X}} = \mathbf{WZ}^{1}}{\mathbf{E}_{k}^{(i+1)} + \mathbf{X}} \times \hat{\mathbf{X}} = \mathbf{WZ}^{1}$$

$$- \mathbf{H}_{k}^{(i+1)} + \mathbf{X} \times \hat{\mathbf{X}} = \mathbf{WZ}^{1}$$

$$- \mathbf{H}_{k}^{(i+1)} + \mathbf{H}_{k}$$

$$\pi_k^{(i+1)} = \frac{1}{N} \sum_{n=1}^N \gamma^{\left(i\right)}(r_{nk})$$

If the covariance is diagonal, then we have

11 Matrix factorization

- We have D movies and N users
- ${f X}$ is a matrix $D \times N$ with x_{dn} the rating of n'th user for d'th movie.
- We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
- Z a N × M matrix that gives features for the users
- W a D × M matrix that gives features for the movies

$$x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$$

We can add a regularizer and minimize the

$$\begin{split} \mathcal{L}(\mathbf{W}, \mathbf{Z}) &= \frac{1}{2} \sum_{n=1}^{N} \sum_{d=1}^{D} (x_{dn} - \mathbf{w}_{d}^{T} \mathbf{z}_{n})^{2} \\ &+ \frac{\lambda_{w}}{2} \sum_{n=1}^{D} \mathbf{w}_{d}^{T} \mathbf{w}_{d} + \frac{\lambda_{z}}{2} \sum_{n=1}^{N} \mathbf{z}_{n}^{T} \mathbf{z}_{n} \end{split}$$

We can use coordinate descent algorithm, by first minimizing w.r.t. Z given W and then minimizing W given Z. This is called

Alternating least-squares (ALS):

$$\mathbf{z}^T \leftarrow (\mathbf{w}^T \mathbf{w} + \lambda_z \mathbf{I}_M)^{-1} \mathbf{w}^T \mathbf{x}$$

 $\mathbf{w}^T \leftarrow (\mathbf{z}^T \mathbf{z} + \lambda_m \mathbf{I}_M)^{-1} \mathbf{z}^T \mathbf{x}^T$

Propabilistic model
$$\prod_{n=1}^{N} \prod_{d \in O_n} \mathcal{N}(\mathbf{x}_{dn} | \mathbf{w}_d^T \mathbf{z}_n, I) \times \prod_{n=1}^{N} \mathcal{N}(\mathbf{z}_n | 0, \frac{1}{\lambda_z} I) - \begin{cases} \mathbf{a}_n^{(k)} = \mathbf{B}^{(k)} \mathbf{z}_n^{(k-1)}, \quad \mathbf{z}_n^{(k)} = h(\mathbf{a}_n^{(k)}) \\ \text{thus we have the input-output relationship} \end{cases}$$

$$\times \prod_{d=1}^{D} \mathcal{N}(\mathbf{w}_d | 0, \frac{1}{\lambda_w} I) - \begin{cases} \mathbf{a}_n^{(k)} = \mathbf{B}^{(k)} \mathbf{z}_n^{(k-1)}, \quad \mathbf{z}_n^{(k)} = h(\mathbf{a}_n^{(k)}) \\ \text{thus we have the input-output relationship} \end{cases}$$

$$\hat{y}_n = g((\boldsymbol{\beta}^{(K-1)})^T * h(\mathbf{B}^{(K-2)} * h(*\dots * w)) + h(*\dots * w)$$

$$\text{with } g(t) \text{ the link function}$$

$$\text{We learn parameters } \mathbf{B} \text{ using stochastic}$$

Since many ratings are missing we cannot normalize the data. A solution is to add offset

$$\frac{1}{2} \sum_{n=1}^{N} \sum_{d \in O_n} (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu)^2$$

12 Singular Value Decomposition

- Matrix factorization method

$$\mathbf{X} = \mathbf{U}\mathbf{S}\mathbf{V}^T$$

- \mathbf{U} is an $D \times D$ matrix
- \mathbf{V} is an $N \times N$ matrix - **S** is a non-negative diagonal matrix of size $D \times N$ which are called singular values
- appearing in a descending order. Columns of U and V are the left and right
- singular vectors respectively. Assuming D < N we have

$$\mathbf{X} = \sum_{l=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information.

Let's now truncate these matrices

$$\mathbf{X} pprox \mathbf{U}_{tr} \mathbf{S}_{tr} \mathbf{V}_{tr}^*$$
 $\mathbf{T}_{tr} pprox \mathbf{U}_{tr} \mathbf{S}_{tr}$
 $\mathbf{T}_{te} pprox \mathbf{X}_{te} \mathbf{V}_{tr}$

with T_{tr} the reduced feature set of X and T_{te} the reduced feature set of \mathbf{X}_{te}

13 Principal Componement Analysis

- PCA is a dimensionality reduction method and a method to decorrelate the data

$$\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$$

$$\mathbf{\Sigma} = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$
$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Thus the columns of matrix U are called the principal components and they decorrelate
- the covariance matrix. Using SVD, we can compute the matrices in the

$$\mathbf{W} = \mathbf{U}\mathbf{S}^{1/2}$$
$$\mathbf{Z} = \mathbf{V}\mathbf{S}^{1/2}$$

following way

- 14 Belief Propagation - the goal is to learn inference of the latent variables
- using belief propagation Given a directed acyclic graph G and parameters θ , a Bayesian network defines the joint distribution as follows

$$p_{\theta}(\mathbf{x}) = \prod_{k=1}^{K} p_{\theta}(x_k | \text{parents}_k)$$

- We can reduce the complexity by using the structure of the problem and the bayesian rule.
- 15 Multi-Layer Perceptron (MLP)
- Known as feed-forward neural network - There are a total of K layers
- $-\mathbf{z}_n^{(k)}$ is the k'th hidden vector $-\mathbf{z}_n^{(k)}$ is the corresponding activation

$$a_{mn}^{(k)} = (\beta_m^{(k)})^T \mathbf{z}_n^{(k-1)}, \quad z_{mn}^{(k)} = h(a_{mn}^{(k)})$$

- For the first layer we have $\mathbf{z}_n^{(0)} = \mathbf{x}_n$ - For the last layer, we use a link function to map

 $\mathbf{z}_n^{(K-1)}$ to the output \mathbf{y}_n A 1-layer MLP is simply a generalization of linear/logistic regression

 $\mathbf{B}^{(k)}$ a matrix with rows $(\boldsymbol{\beta}_{m}^{(k)})^{T}$

$$\mathbf{a}_n^{(k)} = \mathbf{B}^{(k)} \mathbf{z}_n^{(k-1)}, \quad \mathbf{z}_n^{(k)} = h(\mathbf{a}_n^{(k)}$$

$$\hat{y}_n = g((\boldsymbol{\beta}^{(K-1)})^T * h(\mathbf{B}^{(K-2)}) * h(*...*h(\mathbf{B}^{(1)}) *$$
 with $g()$ the link function

We learn parameters B using stochastic gradient-descent

Frequently used transfer function

$$g(a) = \tanh(a) = \frac{e^a - e^{-a}}{e^a + e^{-a}}$$

Backpropagation is a technic to compute the gradient in time linear in the number of training points and the number of weights.

16 Gaussian Process (GP)

- GP are a family of statistical distributions in which time plays a role and for which any finite linear combination of samples has a joint Gaussian distribution.
- They can be completely defined by their second-order statistics which means that if they have mean zero, then defining the covariance function completely defines the process behaviour.
- Let us place a probabilistic prior shape on the approximation of a function.
- A GP process defines a prior over function f

$$p(\mathbf{f}|X) = \mathcal{N}(\mathbf{f}|\mathbf{0}, K(X))$$

K(X) defines shape and prior knowledge about

$$\begin{split} p(\mathbf{f}|\mathbf{y}, \mathbf{X}_*, \mathbf{X}) &\sim \mathcal{N}(\mu', \sigma') \\ \mu' &= K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]^{-1} \mathbf{y} \\ \sigma' &= K(\mathbf{X}_*, \mathbf{X}_*) - K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]K(\mathbf{X}, \mathbf{X}) \end{split}$$

with σ_n the variance of the noise

$$k(\mathbf{x}_n, \mathbf{x}_m) = e^{-||\mathbf{x}_n - \mathbf{x}_m||^2/L^2}$$

- Quadratic kernel

$$k(\mathbf{x}_n, \mathbf{x}_m) = (1 + \mathbf{x}^T \mathbf{x}_m)^2$$

- 17 Decision Trees (DT) - A decision tree is a structure in which each internal node represents a test on an attribute and each branch represents the outcome of the
- test and each leaf represents a class label. Fast to train and fast to make predictions Efficient for very high dimensional feature
- spaces and very large amounts of training data Lack of smoothness and high variance (overfitting)
- Goal: find a split (k, τ) that minimizes an impurity measure at the leaves

- Variance of the model averaging:

- Find best feature to split on

- Find best threshold

18 Random Forests (RF)

- RF correct the overfitting bad "habit" of DTs Training: Learn M trees on different subsets

(random) of training data - Prediction: Average of prediction of each tree

$$z_1 = f_1 \quad \Rightarrow \quad z_M = \frac{1}{M} \sum_{i=1}^{M} f_i$$

$$V(z_1) = \sigma^2 \quad \Rightarrow \quad V(z_M) = \frac{1}{M}\sigma^2 + \rho \frac{M-1}{M}\sigma^2$$

Variance reduction ratio:

$$\frac{V(z_1)}{V(z_M)} = \frac{M}{1 + \rho(M-1)}$$

- 2 techniques for decorrelating trees
- Bagging: Randomize training data
- Randomized feature selection