PCML Cheat Sheet

1 Math Prerequisites

- Bayes rule

$$p(A, B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

- Gaussian distribution

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}))$$

Production of independent variables:

$$V(XY) = E(X^{2})E(Y^{2}) - [E(X)]^{2}[E(Y)]^{2}$$

- Log-properties

$$\log(mn) = \log(m) + \log(n)$$
$$\log(m^n) = n\log(m)$$

Covariance matrix of a data vector x

$$\Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - E(\mathbf{x})(\mathbf{x}_n - E(\mathbf{x})^T)$$

1.1 Convexity

- A function is convex when a line joining two points never interects with the function anywhere else.
- A function f(x) is convex, if for any $x_1, x_2 \in \mathbf{X}$ and for any $0 < \lambda < 1$, we have:

$$f(\lambda x_1 + (1 - \lambda)x_2) < \lambda f(x_1) + (1 - \lambda)f(x_2)$$

- A function is strictly convex if the inequality is strict.
- A convex function has only one global minimum.
- Sums of convex functions are also convex.
- The Hessian is related to the convexity of a function: a twice differentiable function is convex if and only if the Hessian is positive definite.
- The Hessian of a convex function is positive semi-definite and for a strictly-convex function it is positive definite.
- The Hessian matrix of a function

$$\mathbf{H}_{i,j} = \frac{d^2 f}{dx_i dx_j}$$

1.2 Linear Algebra

- Column $\mathbf{x} \in \mathbb{R}^n$, rows \mathbf{x}^T , matrix $\mathbf{A} \in \mathbb{R}^{m \times n}$
- $\mathbf{x}^T \mathbf{x}$ is a scalar, $\mathbf{x} \mathbf{x}^T$ is a matrix
- A^{-1} exist if A is full rank
- Condition number of a function measures how much the output value can change for a small change in the input argument. A matrix with a high condition number is said to be **ill-conditioned**. If **A** is normal $(A^TA = AA^T)$ then

$$k(\mathbf{A}) = \left| rac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

- A positive definite matrix is **symmetric** with all positive eigenvalues
- The real symmetric $N \times N$ matrix **V** is said to be **positive semidefinite** if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} \ge 0$$

for any real $N \times 1$ vector a.

- The real symmetric $N \times N$ matrix **V** is said to be **positive definite** if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

for any real $N \times 1$ vector a.

- Cost of matrix inversion: $O(n^3) \to O(n^{2.372})$
- Cost of determinant computation using LU decomposition: $O(n^3)$

2 Cost functions

- Cost functions are used to learn parameters that explain the data well.
- It is essential to make sure that a global minimum exist \rightarrow lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{\beta}) = \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is convex thus it has only one global minumum value.
- MSE is not good when outliers are present.

Mean Absolute Error (MAE):

$$MAE = \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

Huber loss

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \le \delta \\ \delta|z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

- Huber loss is convex, differentiable, and also robust to outliers
- hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Tukey's loss is non-convex, non-differentiable, but robust to outliers. Hinge loss

$$Hinge = [1 - y_n f(\mathbf{x}_n)]_+$$

Logistic loss

$$Logistic = \log(1 - \exp(y_n f(\mathbf{x}_n)))$$

3 Regression

- **Data** consists of N pairs (y_n, \mathbf{x}_n)
- 1. u_n the n'th output
- 2. \mathbf{x}_n is a vector of D inputs
- **Prediction**: predict the ouput for a new input vector.

- **Interpretation**: understand the effect of inputs on output.
- Outliers are data that are far away from most of the other examples.

3.1 Linear Regression

- Model that assume linear relationship between inputs and the ouput.

$$y_n \equiv f(\mathbf{x}_n)$$

$$:= \beta_0 + \beta_1 x_{n1} + \dots$$

$$= \mathbf{x}_n^T \boldsymbol{\beta}$$

with β the parameters of the model.

- Variance grows only linearly with dimensionality

3.2 Gradient Descent

- Gradient descent uses only first-order information and takes steps in the direction of the gradient
- Given a cost function $\mathcal{L}(\beta)$ we wish to find β that minimizes the cost:

$$\min_{\boldsymbol{\beta}} \mathcal{L}(\boldsymbol{\beta})$$

3.2.1 Grid search

- Compute the cost over a grid of M points to find the minimum
- Exponential Complexity $O(NDM^D)$
- Hard to find a good range of values

3.3 Batch Gradient Descent

- Take steps in the opposite direction of the gradient

$$\boldsymbol{\beta}^{(k+1)} \leftarrow \boldsymbol{\beta}^{(k)} - \alpha \frac{d\mathcal{L}(\boldsymbol{\beta}^{(k)})}{d\boldsymbol{\beta}}$$

with $\alpha > 0$ the learning rate.

- With α too big, method might diverge. With α too small, convergence is slow.

3.4 Gradients for MSE

$$\mathbf{\tilde{X}} = [1 \ \mathbf{X}]$$

We define the error vector e:

$$e = y - \tilde{X}\beta$$

- and MSE as follows:

$$\mathcal{L}(\boldsymbol{\beta}) = \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

- then the gradient is given by

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = -\frac{1}{N}\tilde{\mathbf{X}}^T\mathbf{e}$$

- Optimality conditions:
- 1. necessary: gradient equal to zero: $\frac{d\mathcal{L}(\boldsymbol{\beta}^*)}{d\boldsymbol{\beta}} = 0$
- 2. sufficient: Hessian matrix is positive definite: $\mathbf{H}(\boldsymbol{\beta}^*) = \frac{d^2 \mathcal{L}(\boldsymbol{\beta}^*)}{d\beta d\beta^T}$
- Very sensitive to illconditioning. Therefore, always normalize your feature otherwise step-size selection is difficult since different directions might move at different speed.
- Complexity: O(NDI) with I the number of iterations

3.5 Least Squares

- In some cases, we can compute the minimum of the cost function analytically.
- use the first optimality conditions:

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = 0 \Rightarrow \tilde{\mathbf{X}}^T \mathbf{e} = \tilde{\mathbf{X}}^T (\mathbf{y} - \tilde{\mathbf{X}}\boldsymbol{\beta}) = 0$$

– When $\tilde{\mathbf{X}}^T \tilde{\mathbf{X}}$ is invertible, we have the closed-form expression

$$\boldsymbol{\beta}^* = (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

- thus we can predict values for a new \mathbf{x}_*

$$y_* = \tilde{\mathbf{x}}_*^{\mathbf{T}} \boldsymbol{\beta}^* = \tilde{\mathbf{x}}_*^{\mathbf{T}} (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

- The Gram matrix $\tilde{\mathbf{X}}^T\tilde{\mathbf{X}}$ is positive definite and is also invertible iff $\tilde{\mathbf{X}}$ has full column rank.
- Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$
- $\tilde{\mathbf{X}}$ can be rank deficient when D > N or when the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. In this case, the matrix is ill-conditioned, leading to numerical issues.

3.6 Maximum Likelihood

– Let define our mistakes $\epsilon_n \sim \mathcal{N}(0, \sigma^2)$.

$$\to y_n = \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} + \epsilon_n$$

- Another way of expressing this:

$$p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta}) = \prod_{n=1}^{N} p(y_n|\tilde{\mathbf{x}}_n, \boldsymbol{\beta})$$
$$= \prod_{n=1}^{N} \mathcal{N}(y_n|\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}, \sigma^2)$$

which defines the likelihood of observating ${\bf y}$ given $\tilde{{\bf X}}$ and ${\boldsymbol \beta}$

Define cost with log-likelihood

$$\mathcal{L}_{lik}(\boldsymbol{\beta}) = \log p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 + cnst$$

- Maximum likelihood estimator (MLE) gives another way to design cost functions

$$\underset{\boldsymbol{\beta}}{\operatorname{argmin}} \, \mathcal{L}_{mse}(\boldsymbol{\beta}) = \underset{\boldsymbol{\beta}}{\operatorname{argmax}} \, \mathcal{L}_{lik}(\boldsymbol{\beta})$$

- MLE can also be interpreted as finding the model under which the obeserved data is most likely to have been generated from.
- With Laplace distribution

$$p(y_n|\tilde{\mathbf{x}}_n, \boldsymbol{\beta}) = \frac{1}{2b} e^{-\frac{1}{b}|y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}|}$$

$$\sum_{n} \log p(y_n | \tilde{\mathbf{x}}_n, \boldsymbol{\beta}) = \sum_{n} |y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}| + cnst$$

3.7 Ridge Regression

- Linear models usually underfit. One way is to use nonlinear basis functions instead.

$$y_n = eta_0 + \sum_{j=1}^M eta_j \phi_j(\mathbf{x}_n) = ilde{oldsymbol{\phi}}(\mathbf{x}_n)^T oldsymbol{eta}$$

- This model is linear in β but nonlinear in \mathbf{x} . Note that the dimensionality is now M, not D.
- Polynomial basis

$$\phi(x_n) = [1, x_n, x_n^2, ..., x_n^M]$$

- The least square solution becomes

$$\boldsymbol{\beta}_{lso}^* = (\tilde{\boldsymbol{\Phi}}^T \tilde{\boldsymbol{\Phi}})^{-1} \tilde{\boldsymbol{\Phi}}^T \mathbf{y}$$

 Complex models overfit easily. Thus we can choose simpler models by adding a regularization term which penalizes complex models

$$\min_{oldsymbol{eta}} \left(\mathcal{L}(oldsymbol{eta}) + rac{\lambda}{2N} \sum_{j=1}^{M} eta_j^2
ight)$$

$$\boldsymbol{\beta}^* = \operatorname*{argmin}_{\boldsymbol{\beta}} \left(\frac{1}{2} (\mathbf{y} - \mathbf{X} \boldsymbol{\beta})^T (\mathbf{y} - \mathbf{X} \boldsymbol{\beta}) + \frac{\lambda}{2} \boldsymbol{\beta}^T \boldsymbol{\beta} \right)$$

- Note that β_0 is not penalized.
- By differentiating and setting to zero we get

$$oldsymbol{eta}_{ridge} = (\mathbf{ ilde{\Phi}}^T\mathbf{ ilde{\Phi}} + \mathbf{\Lambda})^{-1}\mathbf{ ilde{\Phi}}^T\mathbf{y}$$

$$\mathbf{\Lambda} = \left[\begin{array}{cc} 0 & \underline{0} \\ \underline{0} & \lambda I_m \end{array} \right]$$

- Ridge regression improves the condition number of the Gram matrix since the eigenvalues of $(\tilde{\Phi}^T \tilde{\Phi} + \lambda I_m)$ are at least λ
- Maximum-a-posteriori (MAP) estimator:
- Maximizes the product of the likelihood and the **prior**.

$$\boldsymbol{\beta}_{map} = \operatorname*{argmax}_{\boldsymbol{\beta}} \left(p(\mathbf{y}|\mathbf{X}, \boldsymbol{\Lambda}) p(\boldsymbol{\beta}|\boldsymbol{\Sigma}) \right)$$

- Assume $\beta_0 = 0$

$$m{eta}_{ridge} = \operatorname*{argmax}_{m{eta}} \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T m{eta}, \mathbf{\Lambda}) imes \mathcal{N}(m{eta} | 0, \mathbf{I})
ight]
ight)$$

- Lasso regularizer forces some β_i to be strictly 0 and therefore forces sparsity in the model.

$$\min_{\boldsymbol{\beta}} \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{\beta})^2, \quad \text{such that } \sum_{i=1}^{M} |\beta_i| \le \tau$$

3.8 Cross-Validation

– We should choose λ to minimize the mistakes that will be made in the future.

- We split the data into train and validation sets and we pretend that the validation set is the future data. We fit our model on the training set and compute a prediction-error on the validation set. This gives us an *estimate* of the *qeneralization error*.
- **K-fold cross validation** randomly partition the data into K groups. We train on K-1 groups and test on the remaining group. We repeat this until we have tested on all K sets. We then average the results.
- Cross-validation returns an unbiased estimate of the generalization error and its variance.

3.9 Bias-Variance decomposition

- The expected test error can be expressed as the sum of two terms
- **Squared bias**: The average *shift* of the predictions
- Variance: measure how data points vary around their average.

expected loss =
$$(bias)^2 + variance + noise$$

- Both model bias and estimation bias are important
- Ridge regression increases estimation bias while reducing variance
- Increasing model complexity increases test error

Small $\lambda \to \text{low}$ bias but large variance

Large $\lambda \to \text{large bias but low variance}$

3.10 Logistic Regression

- Classification relates input variables ${\bf x}$ to discrete output variable y
- Binary classifier: we use y = 0 for C_1 and y = 1 for C_2 .
- Can use least-squares to predict \hat{y}_*

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(y_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$
$$p(y_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$

- The probabilistic model:

$$p(\mathbf{y}|\mathbf{X},\boldsymbol{\beta}) = \prod_{n=1}^{N} \sigma(\tilde{\mathbf{x}}_{n}^{T}\boldsymbol{\beta})^{y_{n}} (1 - \sigma(\tilde{\mathbf{x}}_{n}^{T}\boldsymbol{\beta})^{1 - y_{n}})$$

- The log-likelihood:

$$\mathcal{L}_{mle}(\boldsymbol{\beta}) = \sum_{n=1}^{N} \left(y_n \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} - \log(1 + \exp(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})) \right)$$

We can use the fact that

$$\frac{d}{dx}\log(1+\exp(x)) = \sigma(x)$$

- Gradient of the log-likelihood

$$\mathbf{g} = \frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \sum_{n=1}^{N} \left(\tilde{\mathbf{x}}_{n} y_{n} - \tilde{\mathbf{x}}_{n} \sigma(\tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta}) \right)$$
$$= -\tilde{\mathbf{X}}^{T} [\sigma(\tilde{\mathbf{X}} \boldsymbol{\beta}) - \mathbf{y}]$$

- The negative of the log-likelihood $-\mathcal{L}_{mle}(\beta)$ is convex
- **Hessian** of the log-likelihood
- We know that

$$\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t))$$

Hessian is the derivative of the gradient

$$\mathbf{H}(\boldsymbol{\beta}) = -\frac{d\mathbf{g}(\boldsymbol{\beta})}{d\boldsymbol{\beta}^T} = \sum_{n=1}^{N} \frac{d}{d\boldsymbol{\beta}^T} \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) \tilde{\mathbf{x}}_n$$
$$= \sum_{n=1}^{N} \tilde{\mathbf{x}}_n \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) (1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})) \tilde{\mathbf{x}}_n^T$$
$$= \tilde{\mathbf{X}}^T \mathbf{S} \tilde{\mathbf{X}}$$

where **S** is a $N \times N$ diagonal matrix with diagonals

$$S_{nn} = \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) (1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

- The negative of the log-likelihood is not strictly convex.

- Newton's Method

 Uses second-order information and takes steps in the direction that minimizes a quadratic approximation

$$\mathcal{L}(\boldsymbol{\beta}) = \mathcal{L}(\boldsymbol{\beta}^{(k)}) + \mathbf{g}_k^T (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)}) + (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})^T \mathbf{H}_k (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})$$

and it's minimum is at

$$\boldsymbol{\beta}^{k+1} = \boldsymbol{\beta}^{(k)} - \alpha_k \mathbf{H}_k^{-1} \mathbf{g}_k$$

where \mathbf{g}_k is the gradient and α_k the learning rate.

- Complexity: $O((ND^2 + D^3)I)$
- Penalized Logistic Regression

$$\min_{\boldsymbol{\beta}} \left(-\sum_{n=1}^{N} \log p(y_n | \mathbf{x}_n^T \boldsymbol{\beta}) + \lambda \sum_{d=1}^{D} \beta_d^2 \right)$$

4 Generalized Linear Model

- Exponential family distribution

$$p(\mathbf{y}|\boldsymbol{\eta}) = \frac{h(y)}{Z} \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

Bernoulli distribution

$$p(y|\mu) = \mu^y (1 - \mu)^{1-y}$$

= $\exp(y \log(\frac{\mu}{1 - \mu} + \log(1 - \mu)))$

- there is a relationship between η and μ throught the link function

$$\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$$

- Note that μ is the mean parameter of y
- Relationship between the mean μ and η is defined using a link function g

$$\eta = \mathbf{g}(\boldsymbol{\mu}) \Leftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta})$$

- First and second derivatives of $A(\eta)$ are related to the mean and the variance

$$\frac{dA(\eta)}{d\eta} = E[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = Var[\phi(\eta)]$$

- $A(\eta)$ is convex
- The generalized maximum likelihood cost to minimize is

$$\min_{\boldsymbol{\beta}} \mathcal{L}(\boldsymbol{\beta}) = -\sum_{n=1}^{N} \log(p(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

where $p(y_n|\tilde{\mathbf{x}}_n^T\boldsymbol{\beta})$ is an exponential family distribution

We obtain the solution

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \tilde{\mathbf{X}}^T [\mathbf{g}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\mathbf{y})]$$

5 k-Nearest Neighbor (k-NN)

– The k-NN prediction for \mathbf{x} is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$$

where $nbh_k(\mathbf{x})$ is the neighborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n in the training data

- Curse of dimensionality: Generalizing correctly becomes exponentially harder as the dimensionality grows.
- Gathering more inputs variables may be a bad thing

6 Kernel Ridge Regression

- The following is true for ridge regression

$$\boldsymbol{\beta} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y} \tag{1}$$

$$= \mathbf{X}^{T} (\mathbf{X} \mathbf{X}^{T} + \lambda \mathbf{I}_{N})^{-1} \mathbf{y}$$

$$= \mathbf{X}^{T} \boldsymbol{\alpha}$$
(2)

- Complexity of computing β
- 1 $O(D^2N + D^3)$
- 2. $O(N^2D + N^3)$

- Thus we have

$$\boldsymbol{\beta} = \sum_{n=1}^{N} \alpha_n \mathbf{x}_n, \ \mathbf{y} = \sum_{d=1}^{D} \beta_d \bar{\mathbf{x}}_d$$

with \mathbf{x}_n the rows of **X** and $\bar{\mathbf{x}}_d$ the columns of **X**

- The representer theorem allows us to write an equivalent optimization problem n terms of α .

$$\boldsymbol{\alpha} = \operatorname*{argmax}_{\boldsymbol{\alpha}} \left(-\frac{1}{2} \boldsymbol{\alpha} (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^T \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

- $\mathbf{X}\mathbf{X}^T$ is called the **kernel matrix**
- Kernelized rigde regression might be computationally more efficient in some cases.
- Kernel trick:
- We can work directly with $\mathbf{K} = \mathbf{X}\mathbf{X}^T$ and never have to worry about \mathbf{X}
- Using the basis function $\phi(\mathbf{x})$, we do not need to specify it explicitly, since we can work directly with $\mathbf{K} = \phi(\mathbf{x})$
- We will use a kernel function $k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$
- The evaluation of a kernel is usually faster with k than with ϕ
- Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Properties of a kernel to ensure the existence of a corresponding ϕ :
- **K** should be symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$
- K should be positive semidefinite.

7 Support Vector Machine

Hinge loss

$$[t]_{+} = \max(0, t)$$

- Solution to the dual problem is sparse and non-zero entries will be our support vectors.
- Assume $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$\min_{\boldsymbol{\beta}} \sum_{n=1}^{N} [1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}]_+ + \frac{\lambda}{2} \sum_{j=1}^{M} \beta_j^2$$

- The minimum doesn't change with a rescaling of β
- Duality:
- Hard to minimize $g(\beta)$ so we define

$$g(\boldsymbol{\beta}) = \max_{\boldsymbol{\alpha}} G(\boldsymbol{\beta}, \boldsymbol{\alpha})$$

- we use the property that

$$C[v_n]_+ = \max(0, Cv_n) = \max_{\alpha_n \in [0, C]} \alpha_n v_n$$

- We can rewrite the problem as

$$\min_{\boldsymbol{\beta}} \max_{\boldsymbol{\alpha} \in [0,C]^N} \sum_{n=1}^N \alpha_n (1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}) + \frac{1}{2} \sum_{i=1}^M \beta_j^2$$

- This is differentiable, convex in β and concave in α
- Minimax theorem:

$$\min_{\beta} \max_{\alpha} G(\beta, \alpha) = \max_{\alpha} \min_{\beta} G(\beta, \alpha)$$

because G is convex in β and concave in α .

- Derivative w.r.t. β :

$$\frac{dG}{d\beta} = -\left(\sum_{n=1}^{N} \alpha y_n \tilde{\phi_n}\right) + \begin{bmatrix} 0 \\ \beta_{1:M} \end{bmatrix}$$

- Equating this to 0, we get:

$$eta_{1:M}^* = \sum_{n=1}^N lpha_n y_n \phi_n = \mathbf{\Phi}^T \mathrm{diag}(\mathbf{y}) \boldsymbol{\alpha}$$

 $\mathbf{\alpha}^T \mathbf{y} = 0$

- Plugging β^* back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,C]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2} \boldsymbol{\alpha}^T \mathbf{Y} \boldsymbol{\Phi} \boldsymbol{\Phi} \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{\Phi} \mathbf{\Phi}^T$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision boundary.

8 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances outside
 the cluster.

$$\min_{\mathbf{r}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{r}, \boldsymbol{\mu}) = \sum_{k=1}^{K} \sum_{n=1}^{N} r_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

such that $r_{nk} \in \{0,1\}$ and $\sum_{k=1}^{K} r_{nk} = 1$

- K-means algorithm:

Initialize μ_k , then iterate

1. For all n, compute \mathbf{r}_n given $\boldsymbol{\mu}$

$$r_{nk} = \begin{cases} 1 & \text{if } k = \operatorname{argmin}_{j} ||\mathbf{x}_{n} - \boldsymbol{\mu}||_{2}^{2} \\ 0 & \text{otherwise} \end{cases}$$

2. For all k, compute μ_k given r

$$\mu_k = \frac{\sum_{n=1}^{N} r_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} r_{nk}}$$

 A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points. Probabilistic model

$$p(\mathbf{r}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{r_{nk}}$$

 Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

9 Gaussian Mixture Models

- Clusters can be spherical using a full covariance matrix instead of isotropic covariance.

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{r}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{r_{nk}}$$

- Soft-clustering: Points can belong to several cluster by defining r_n to be a random variable.

$$p(r_{nk} = 1) = \pi_k$$
 where $\pi_k > 0, \forall k$
$$\sum_{k=1}^K \pi_k = 1$$

- Joint distribution of Gaussian mixture model

$$p(\mathbf{X}, \mathbf{r} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^{N} \left[p(\mathbf{x}_n | r_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(r_n | \boldsymbol{\pi}) \right]$$
$$= \left[\prod_{k=1}^{K} \left[(\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k))^{r_{nk}} \right] \prod_{k=1}^{K} [\boldsymbol{\pi}]^{r_{nk}} \right]$$

 $Joint = Likelihood \times Prior$

- $-r_n$ are called *latent* unobserved variables
- Unknown parameters are given by $\theta = \{\mu, \Sigma, \pi\}$
- We get the **marginal likelihood** by marginalizing r_n out from the likelihood

$$p(\mathbf{x}_n|\boldsymbol{\theta}) = \sum_{k=1}^K p(\mathbf{x}_n, r_n = k|\boldsymbol{\theta})$$
$$= \sum_{k=1}^K p(r_n = k|\boldsymbol{\theta})p(\mathbf{x}_n|r_n = k, \boldsymbol{\theta})$$
$$= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to $O(D^2K)$
- To get maximum likelihood estimate of θ , we maximize

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\sigma}_{k})$$

10 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\boldsymbol{\theta}^{(1)}$ and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\theta^{(i)}$

$$\log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \ge \sum_{k=1}^{K} p_{kn} \log \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{p_{kn}}$$

with equality when,

$$p_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n, \mathbf{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \mathbf{\Sigma}_k)}$$

2. Maximization step: Update θ

$$\boldsymbol{\theta}^{(i+1)} = \operatorname*{argmax}_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(i)})$$

$$\mu_k^{i+1} = \frac{\sum_{n=1}^{N} p_{kn}^{(i)} \mathbf{x}_n}{\sum_{n=1}^{N} p_{kn}^{(i)}}$$

$$\boldsymbol{\Sigma}_{k}^{(i+1)} = \frac{\sum_{n=1}^{N} p_{kn}^{(i)} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})^{T}}{\sum_{n=1}^{N} p_{kn}^{(i)}}$$

$$\pi_k^{(i+1)} = \frac{1}{N} \sum_{n=1}^{N} p_{kn}^{(i)}$$

- If the covariance is diagonal, then we have K-means.

11 Matrix factorization

- We have D movies and N users
- **X** is a matrix $D \times N$ with x_{dn} the rating of n'th user for d'th movie.
- We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
- **Z** a $N \times M$ matrix that gives features for the users
- W a $D \times M$ matrix that gives features for the movies

$$x_{dn} pprox \mathbf{w}_d^T \mathbf{z}_n$$

- We can add a regularizer and minimize the following cost:

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{n=1}^{N} \sum_{d=1}^{D} (x_{dn} - \mathbf{w}_{d}^{T} \mathbf{z}_{n})^{2} + \frac{\lambda_{w}}{2} \sum_{d=1}^{D} \mathbf{w}_{d}^{T} \mathbf{w}_{d} + \frac{\lambda_{z}}{2} \sum_{n=1}^{N} \mathbf{z}_{n}^{T} \mathbf{z}_{n}$$

We can use coordinate descent algorithm, by first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_M)^{-1} \mathbf{W}^T \mathbf{X}$$
$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_M)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- Complexity: $O(DNM^2 + NM^3) \rightarrow O(DNM^2)$

Probabilistic model

$$\prod_{n=1}^{N} \prod_{d \in O_n} \mathcal{N}(x_{dn} | \mathbf{w}_d^T \mathbf{z}_n, I) \times \prod_{n=1}^{N} \mathcal{N}(\mathbf{z}_n | 0, \frac{1}{\lambda_z} I) \times \prod_{d=1}^{D} \mathcal{N}(\mathbf{w}_d | 0, \frac{1}{\lambda_w} I)$$

- Since many ratings are missing we cannot normalize the data. A solution is to add offset terms:

$$\frac{1}{2} \sum_{n=1}^{N} \sum_{d \in O_n} (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu)^2$$

Singular Value Decomposition

Matrix factorization method

$$\mathbf{X} = \mathbf{IISV}^T$$

- **U** is an $D \times D$ matrix
- **V** is an $N \times N$ matrix
- S is a non-negative diagonal matrix of size $D \times N$ which are called **singular values** appearing in a descending order.
- Columns of **U** and **V** are the left and right **singular vectors** respectively.
- Assuming D < N we have

$$\mathbf{X} = \sum_{d=i}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information.

- Let's now truncate these matrices

$$\mathbf{X} pprox \mathbf{U}_{tr} \mathbf{S}_{tr} \mathbf{V}_{tr}^* \ \mathbf{T}_{tr} pprox \mathbf{U}_{tr} \mathbf{S}_{tr} \ \mathbf{T}_{te} pprox \mathbf{X}_{te} \mathbf{V}_{tr}$$

with \mathbf{T}_{tr} the reduced feature set of \mathbf{X} and \mathbf{T}_{te} the reduced feature set of \mathbf{X}_{te}

Principal Componement Analysis

- PCA is a dimensionality reduction method and a method to decorrelate the data

$$\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$$

such that columns of **W** are orthogonal.

If the data is zero mean

$$\mathbf{C} = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$
$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Thus the columns of matrix **U** are called the **principal components** and they decorrelate the covariance matrix.
- Using SVD, we can compute the matrices in the following way

$$\mathbf{W} = \mathbf{U}\mathbf{S}^{1/2}$$
$$\mathbf{Z} = \mathbf{V}\mathbf{S}^{1/2}$$

14 Belief Propagation

- the goal is to learn inference of the latent variables using belief propagation
- Given a directed acyclic graph G and parameters θ , a **Bayesian network** defines the joint distribution as follows

$$p_{\theta}(\mathbf{x}) = \prod_{k=1}^{K} p_{\theta}(x_k | \text{parents}_k)$$

We can reduce the complexity by using the structure of the problem and the bayesian rule.

Multi-Laver Perceptron (MLP)

- Known as feed-forward neural network
- $-\mathbf{z}_{n}^{(k)}$ is the k'th hidden vector $-\mathbf{a}_{n}^{(k)}$ is the corresponding activation
- There are a total of K layers

$$a_{mn}^{(k)} = (\boldsymbol{\beta}_m^{(k)})^T \mathbf{z}_n^{(k-1)}, \quad z_{mn}^{(k)} = h(a_{mn}^{(k)})$$

- For the first layer we have $\mathbf{z}_n^{(0)} = \mathbf{x}_n$
- For the last layer, we use a link function to map $\mathbf{z}_n^{(K-1)}$ to the output \mathbf{y}_n
- A 1-layer MLP is simply a generalization of linear/logistic regression
- $\mathbf{B}^{(k)}$ a matrix with rows $(\boldsymbol{\beta}_m^{(k)})^T$

$$\mathbf{a}_{n}^{(k)} = \mathbf{B}^{(k)} \mathbf{z}_{n}^{(k-1)}, \quad \mathbf{z}_{n}^{(k)} = h(\mathbf{a}_{n}^{(k)})$$

thus we have the input-output relationship.

$$\hat{y}_n = g((\boldsymbol{\beta}^{(K-1)})^T * h(\mathbf{B}^{(K-2)} * h(*...*h(\mathbf{B}^{(1)} * \mathbf{x}_n)))$$

with q() the link function

- We learn parameters B using stochastic gradient-descent
- Frequently used transfer function

$$g(a) = \tanh(a) = \frac{e^a - e^{-a}}{e^a + e^{-a}}$$

Backpropagation is a technic to compute the gradient in time linear in the number of training points and the number of weights.

Gaussian Process (GP)

- GP are a family of statistical distributions in which time plays a role and for which any finite linear combination of samples has a joint Gaussian distribution.
- They can be completely defined by their second-order statistics which means that if they have mean zero, then defining the covariance function completely defines the process behaviour.
- Let us place a probabilistic prior shape on the approximation of a function.
- A GP process defines a prior over function f

$$p(\mathbf{f}|X) = \mathcal{N}(\mathbf{f}|\mathbf{0}, K(X))$$

-K(X) defines shape and prior knowledge about our problem

$$p(\mathbf{f}|\mathbf{y}, \mathbf{X}_*, \mathbf{X}) \sim \mathcal{N}(\mu', \sigma')$$

$$\mu' = K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]^{-1}\mathbf{y}$$

$$\sigma' = K(\mathbf{X}_*, \mathbf{X}_*) - K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]K(\mathbf{X}, \mathbf{X}_*)$$

with σ_n the variance of the noise

RBF kernel

$$k(\mathbf{x}_n, \mathbf{x}_m) = e^{-||\mathbf{x}_n - \mathbf{x}_m||^2/L^2}$$

- Quadratic kernel

$$k(\mathbf{x}_n, \mathbf{x}_m) = (1 + \mathbf{x}_n^T \mathbf{x}_m)^2$$

17 Decision Trees (DT)

- A decision tree is a structure in which each internal node represents a test on an attribute and each branch represents the outcome of the test and each leaf represents a class label.
- Fast to train and fast to make predictions
- Efficient for very high dimensional feature spaces and very large amounts of training data
- Lack of smoothness and high variance (overfitting)
- Goal: find a split (k, τ) that minimizes an impurity measure at the leaves
- Find best feature to split on
- Find best threshold

18 Random Forests (RF)

- RF correct the overfitting bad "habit" of DTs.
- Training: Learn M trees on different subsets (random) of training data
- Prediction: Average of prediction of each tree
- Variance of the model averaging:

$$z_1 = f_1 \implies z_M = \frac{1}{M} \sum_{i=1}^M f_i$$

$$V(z_1) = \sigma^2 \implies V(z_M) = \frac{1}{M} \sigma^2 + \rho \frac{M-1}{M} \sigma^2$$

Variance reduction ratio:

$$\frac{V(z_1)}{V(z_M)} = \frac{M}{1 + \rho(M - 1)}$$

- 2 techniques for decorrelating trees:
- **Bagging**: Randomize training data
- Randomized feature selection