

PCML Cheat Sheet

1 Math Prerequisites

- Bayes rule

$$p(A, B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

- Gaussian distribution

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})\right)$$

- Production of independent variables:

$$V(XY) = E(X^2)E(Y^2) - [E(X)]^2[E(Y)]^2$$

- Log-properties

$$\log(mn) = \log(m) + \log(n)$$

$$\log(m^n) = n \log(m)$$

- Covariance matrix of a data vector \mathbf{x}_n

$$\mathbf{C} = \frac{1}{N} \sum_{n=1}^N (\mathbf{x}_n - E[\mathbf{x}])(\mathbf{x}_n - E[\mathbf{x}])^T$$

1.1 Convexity

- A function is convex when a line joining two points never intersects with the function anywhere else.
- A function $f(x)$ is convex, if for any $x_1, x_2 \in \mathbf{X}$ and for any $0 \leq \lambda \leq 1$, we have :

$$f(\lambda x_1 + (1-\lambda)x_2) \leq \lambda f(x_1) + (1-\lambda)f(x_2)$$

- A function is strictly convex if the inequality is strict.
- A convex function has only one global minimum.
- Sums of convex functions are also convex.
- The Hessian is related to the convexity of a function: a twice differentiable function is convex if and only if the Hessian is positive definite.
- The Hessian of a convex function is positive semi-definite and for a strictly-convex function it is positive definite.
- The Hessian matrix of a function

$$\mathbf{H}_{i,j} = \frac{d^2 f}{dx_i dx_j}$$

1.2 Linear Algebra

- Column $\mathbf{x} \in R^n$, rows \mathbf{x}^T , matrix $\mathbf{A} \in R^{m \times n}$
- $\mathbf{x}^T \mathbf{x}$ is a scalar, $\mathbf{x}\mathbf{x}^T$ is a matrix
- \mathbf{A}^{-1} exist if \mathbf{A} is full rank
- **Condition number** of a function measures how much the output value can change for a small change in the input argument. A matrix with a high condition number is said to be **ill-conditioned**. If \mathbf{A} is normal then

$$k(\mathbf{A}) = \left| \frac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

- A positive definite matrix is **symmetric** with all positive eigenvalues
- The real symmetric $N \times N$ matrix \mathbf{V} is said to be **positive semidefinite** if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} \geq 0$$

for any real $N \times 1$ vector \mathbf{a} .

- The real symmetric $N \times N$ matrix \mathbf{V} is said to be **positive definite** if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

for any real $N \times 1$ vector \mathbf{a} .

- Cost of matrix inversion: $O(n^3) \rightarrow O(n^{2.372})$
- Cost of determinant computation using LU decomposition: $O(n^3)$

2 Cost functions

- Cost functions are used to learn parameters that explain the data well.
- It is essential to make sure that a global minimum exist \rightarrow lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{\beta}) = \sum_{n=1}^N (y_n - f(\mathbf{x}_n))^2$$

- MSE is **convex** thus it has only one global minimum value.

- MSE is not good when outliers are present.

Mean Absolute Error (MAE):

$$MAE = \sum_{n=1}^N |y_n - f(\mathbf{x}_n)|$$

Huber loss

$$Huber = \begin{cases} \frac{1}{2} z^2 & , |z| \leq \delta \\ \delta |z| - \frac{1}{2} \delta^2 & , |z| > \delta \end{cases}$$

- Huber loss is convex, differentiable, and also robust to outliers
- hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 & , |z| < \delta \\ 0 & , |z| \geq \delta \end{cases}$$

Tukey's loss is non-convex, non-differentiable, but robust to outliers.

Hinge loss

$$Hinge = [1 - y_n f(\mathbf{x}_n)]_+$$

Logistic loss

$$Logistic = \log(1 - \exp(y_n f(\mathbf{x}_n)))$$

3 Regression

- **Data** consists of N pairs (y_n, \mathbf{x}_n)
 1. y_n the n 'th output
 2. \mathbf{x}_n is a vector of D inputs
- **Prediction**: predict the output for a new input vector.
- **Interpretation**: understand the effect of inputs on output.
- **Outliers** are data that are far away from most of the other examples.

3.1 Linear Regression

- Model that assume linear relationship between inputs and the output.

$$y_n \equiv f(\mathbf{x}_n)$$

$$:= \beta_0 + \beta_1 x_{n1} + \dots$$

$$= \mathbf{x}_n^T \boldsymbol{\beta}$$

with $\boldsymbol{\beta}$ the parameters of the model.

- Variance grows only linearly with dimensionality

3.2 Gradient Descent

- Gradient descent uses only first-order information and takes steps in the direction of the gradient
- Given a cost function $\mathcal{L}(\boldsymbol{\beta})$ we wish to find $\boldsymbol{\beta}$ that minimizes the cost:

$$\min_{\boldsymbol{\beta}} \mathcal{L}(\boldsymbol{\beta})$$

3.2.1 Grid search

- Compute the cost over a grid of M points to find the minimum
- Exponential Complexity $O(ND^M D)$
- Hard to find a good range of values

3.3 Batch Gradient Descent

- Take steps in the opposite direction of the gradient

$$\boldsymbol{\beta}^{(k+1)} \leftarrow \boldsymbol{\beta}^{(k)} - \alpha \frac{d\mathcal{L}(\boldsymbol{\beta}^{(k)})}{d\boldsymbol{\beta}}$$

with $\alpha > 0$ the learning rate.

- With α too big, method might diverge. With α too small, convergence is slow.

3.4 Gradients for MSE

$$\tilde{\mathbf{X}} = [1 \quad \mathbf{X}]$$

- We define the error vector \mathbf{e} :

$$\mathbf{e} = \mathbf{y} - \tilde{\mathbf{X}}\boldsymbol{\beta}$$

- and MSE as follows:

$$\mathcal{L}(\boldsymbol{\beta}) = \frac{1}{2N} \sum_{n=1}^N (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

- then the gradient is given by

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = -\frac{1}{N} \tilde{\mathbf{X}}^T \mathbf{e}$$

- Optimality conditions:

1. *necessary*: gradient equal to zero: $\frac{d\mathcal{L}(\boldsymbol{\beta}^*)}{d\boldsymbol{\beta}} = 0$
2. *sufficient*: Hessian matrix is positive definite:

$$\mathbf{H}(\boldsymbol{\beta}^*) = \frac{d^2 \mathcal{L}(\boldsymbol{\beta}^*)}{d\boldsymbol{\beta} d\boldsymbol{\beta}^T}$$

- Very sensitive to illconditioning. Therefore, always normalize your feature otherwise step-size selection is difficult since different directions might move at different speed.
- *Complexity*: $O(NDI)$ with I the number of iterations

3.5 Least Squares

- In some cases, we can compute the minimum of the cost function analytically.
- use the first optimality conditions:

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = 0 \Rightarrow \tilde{\mathbf{X}}^T \mathbf{e} = \tilde{\mathbf{X}}^T (\mathbf{y} - \tilde{\mathbf{X}}\boldsymbol{\beta}) = 0$$

- When $\tilde{\mathbf{X}}^T \tilde{\mathbf{X}}$ is invertible, we have the closed-form expression

$$\boldsymbol{\beta}^* = (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

- thus we can predict values for a new \mathbf{x}_*

$$y_* = \tilde{\mathbf{x}}_*^T \boldsymbol{\beta}^* = \tilde{\mathbf{x}}_*^T (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

- The **Gram matrix** $\tilde{\mathbf{X}}^T \tilde{\mathbf{X}}$ is positive definite and is also invertible iff $\tilde{\mathbf{X}}$ has full column rank,
- *Complexity*: $O(ND^2 + D^3) \equiv O(ND^2)$
- $\tilde{\mathbf{X}}$ can be rank deficient when $D > N$ or when the columns $\tilde{\mathbf{x}}_j$ are nearly collinear. In this case, the matrix is ill-conditioned, leading to numerical issues.

3.6 Maximum Likelihood

- Let define our mistakes $\epsilon_n \sim \mathcal{N}(0, \sigma^2)$.

$$\rightarrow y_n = \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} + \epsilon_n$$

- Another way of expressing this:

$$\begin{aligned} p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta}) &= \prod_{n=1}^N p(y_n|\tilde{\mathbf{x}}_n, \boldsymbol{\beta}) \\ &= \prod_{n=1}^N \mathcal{N}(y_n|\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}, \sigma^2) \end{aligned}$$

which defines the likelihood of observing \mathbf{y} given $\tilde{\mathbf{X}}$ and $\boldsymbol{\beta}$

- Define cost with log-likelihood

$$\mathcal{L}_{lik}(\boldsymbol{\beta}) = \log p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^N (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 + cns t$$

- Maximum likelihood estimator (MLE) gives another way to design cost functions

$$\operatorname{argmin}_{\boldsymbol{\beta}} \mathcal{L}_{mse}(\boldsymbol{\beta}) = \operatorname{argmax}_{\boldsymbol{\beta}} \mathcal{L}_{lik}(\boldsymbol{\beta})$$

- MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.
- With Laplace distribution

$$p(y_n|\tilde{\mathbf{x}}_n, \boldsymbol{\beta}) = \frac{1}{2b} e^{-\frac{1}{b}|y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}|}$$

$$\sum_n \log p(y_n|\tilde{\mathbf{x}}_n, \boldsymbol{\beta}) = \sum_n |y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}| + cns t$$

3.7 Ridge Regression

- Linear models usually underfit. One way is to use nonlinear basis functions instead.

$$y_n = \beta_0 + \sum_{j=1}^M \beta_j \phi_j(\mathbf{x}_n) = \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{\beta}$$

- This model is linear in $\boldsymbol{\beta}$ but nonlinear in \mathbf{x} . Note that the dimensionality is now M , not D .
- Polynomial basis

$$\boldsymbol{\phi}(x_n) = [1, x_n, x_n^2, \dots, x_n^M]$$

- The least square solution becomes

$$\boldsymbol{\beta}_{lse}^* = (\tilde{\boldsymbol{\Phi}}^T \tilde{\boldsymbol{\Phi}})^{-1} \tilde{\boldsymbol{\Phi}}^T \mathbf{y}$$

- Complex models overfit easily. Thus we can choose simpler models by adding a **regularization term** which penalizes complex models

$$\min_{\boldsymbol{\beta}} \left(\mathcal{L}(\boldsymbol{\beta}) + \frac{\lambda}{2N} \sum_{j=1}^M \beta_j^2 \right)$$

$$\boldsymbol{\beta}^* = \operatorname{argmin}_{\boldsymbol{\beta}} \left(\frac{1}{2} (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^T (\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) + \frac{\lambda}{2} \boldsymbol{\beta}^T \boldsymbol{\beta} \right)$$

- Note that β_0 is not penalized.

- By differentiating and setting to zero we get

$$\boldsymbol{\beta}_{ridge} = (\tilde{\boldsymbol{\Phi}}^T \tilde{\boldsymbol{\Phi}} + \boldsymbol{\Lambda})^{-1} \tilde{\boldsymbol{\Phi}}^T \mathbf{y}$$

$$\boldsymbol{\Lambda} = \begin{bmatrix} 0 & 0 \\ 0 & \lambda \mathbf{I}_m \end{bmatrix}$$

- Ridge regression improves the condition number of the Gram matrix since the eigenvalues of $(\tilde{\boldsymbol{\Phi}}^T \tilde{\boldsymbol{\Phi}} + \lambda \mathbf{I}_m)$ are at least λ

- **Maximum-a-posteriori (MAP) estimator**:

- Maximizes the product of the likelihood and the **prior**.

$$\boldsymbol{\beta}_{map} = \operatorname{argmax}_{\boldsymbol{\beta}} (p(\mathbf{y}|\mathbf{X}, \boldsymbol{\Lambda}) p(\boldsymbol{\beta}|\boldsymbol{\Sigma}))$$

- Assume $\beta_0 = 0$

$$\boldsymbol{\beta}_{ridge} = \operatorname{argmax}_{\boldsymbol{\beta}} \left(\log \left[\prod_{n=1}^N \mathcal{N}(y_n|\mathbf{x}_n^T \boldsymbol{\beta}, \boldsymbol{\Lambda}) \times \mathcal{N}(\boldsymbol{\beta}|0, \mathbf{I}) \right] \right)$$

- **Lasso regularizer** forces some β_i to be strictly 0 and therefore forces sparsity in the model.

$$\min_{\boldsymbol{\beta}} \frac{1}{2N} \sum_{n=1}^N (y_n - \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{\beta})^2, \quad \text{such that } \sum_{i=1}^M |\beta_i| \leq \tau$$

3.8 Cross-Validation

- We should choose λ to minimize the mistakes that will be made in the future.
- We split the data into train and validation sets and we pretend that the validation set is the future data. We fit our model on the training set and compute a prediction-error on the validation set. This gives us an *estimate of the generalization error*.
- **K-fold cross validation** randomly partition the data into K groups. We train on $K-1$ groups and test on the remaining group. We repeat this until we have tested on all K sets. We then average the results.
- Cross-validation returns an unbiased estimate of the generalization error and its variance.

3.9 Bias-Variance decomposition

- The expected test error can be expressed as the sum of two terms
 - **Squared bias**: The average *shift* of the predictions
 - **Variance**: measure how data points vary around their average.

$$\text{expected loss} = (\text{bias})^2 + \text{variance} + \text{noise}$$

- Both model bias and estimation bias are important
- Ridge regression increases estimation bias while reducing variance
- Increasing model complexity increases test error

$$\text{Small } \lambda \rightarrow \text{low bias but large variance}$$

$$\text{Large } \lambda \rightarrow \text{large bias but low variance}$$

3.10 Logistic Regression

- **Classification** relates input variables \mathbf{x} to discrete output variable y
- **Binary classifier**: we use $y = 0$ for \mathbf{C}_1 and $y = 1$ for \mathbf{C}_2 .
- Can use least-squares to predict \hat{y}_*

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* \geq 0.5 \end{cases}$$

- **Logistic function**

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(y_n = \mathbf{C}_1|\mathbf{x}_n) = \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$

$$p(y_n = \mathbf{C}_2|\mathbf{x}_n) = 1 - \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$

- The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \boldsymbol{\beta}) = \prod_{n=1}^N \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^{y_n} (1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))^{1-y_n}$$

- The log-likelihood:

$$\mathcal{L}_{mle}(\boldsymbol{\beta}) = \sum_{n=1}^N (y_n \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} - \log(1 + \exp(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})))$$

- We can use the fact that

$$\frac{d}{dx} \log(1 + \exp(x)) = \sigma(x)$$

- Gradient of the log-likelihood

$$\begin{aligned}\mathbf{g} &= \frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \sum_{n=1}^N \left(\tilde{\mathbf{x}}_n y_n - \tilde{\mathbf{x}}_n \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) \right) \\ &= -\tilde{\mathbf{X}}^T [\sigma(\tilde{\mathbf{X}}\boldsymbol{\beta}) - \mathbf{y}]\end{aligned}$$

- The negative of the log-likelihood $-\mathcal{L}_{mle}(\boldsymbol{\beta})$ is convex
- **Hessian** of the log-likelihood

- We know that

$$\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t))$$

- Hessian is the derivative of the gradient

$$\begin{aligned}\mathbf{H}(\boldsymbol{\beta}) &= -\frac{d\mathbf{g}(\boldsymbol{\beta})}{d\boldsymbol{\beta}^T} = \sum_{n=1}^N \frac{d}{d\boldsymbol{\beta}^T} \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) \tilde{\mathbf{x}}_n \\ &= \sum_{n=1}^N \tilde{\mathbf{x}}_n \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) (1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})) \tilde{\mathbf{x}}_n^T \\ &= \tilde{\mathbf{X}}^T \mathbf{S} \tilde{\mathbf{X}}\end{aligned}$$

where \mathbf{S} is a $N \times N$ diagonal matrix with diagonals

$$S_{nn} = \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) (1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

- The negative of the log-likelihood is not strictly convex.
- **Newton's Method**

- Uses second-order information and takes steps in the direction that minimizes a quadratic approximation

$\mathcal{L}(\boldsymbol{\beta}) = \mathcal{L}(\boldsymbol{\beta}^{(k)}) + \mathbf{g}_k^T (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)}) + (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})^T \mathbf{H}_k (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})$ $\mathbf{X}\mathbf{X}^T$ is called the **kernel matrix**

$$\boldsymbol{\beta}^{k+1} = \boldsymbol{\beta}^{(k)} - \alpha_k \mathbf{H}_k^{-1} \mathbf{g}_k$$

where \mathbf{g}_k is the gradient and α_k the learning rate.

- Complexity: $O((ND^2 + D^3)I)$
- **Penalized Logistic Regression**

$$\min_{\boldsymbol{\beta}} \left(-\sum_{n=1}^N \log p(y_n | \mathbf{x}_n^T \boldsymbol{\beta}) + \lambda \sum_{d=1}^D \beta_d^2 \right)$$

4 Generalized Linear Model

- **Exponential family distribution**

$$p(\mathbf{y} | \boldsymbol{\eta}) = \frac{h(\mathbf{y})}{Z} \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

- Bernoulli distribution

$$\begin{aligned}p(y | \mu) &= \mu^y (1 - \mu)^{1-y} \\ &= \exp(y \log(\frac{\mu}{1-\mu}) + \log(1 - \mu))\end{aligned}$$

- there is a relationship between $\boldsymbol{\eta}$ and μ through the **link function**

$$\boldsymbol{\eta} = \log\left(\frac{\mu}{1-\mu}\right) \leftrightarrow \mu = \frac{e^{\boldsymbol{\eta}}}{1 + e^{\boldsymbol{\eta}}}$$

- Note that μ is the mean parameter of y
- Relationship between the mean $\boldsymbol{\mu}$ and $\boldsymbol{\eta}$ is defined using a link function g

$$\boldsymbol{\eta} = \mathbf{g}(\boldsymbol{\mu}) \Leftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta})$$

- First and second derivatives of $A(\boldsymbol{\eta})$ are related to the mean and the variance

$$\frac{dA(\boldsymbol{\eta})}{d\boldsymbol{\eta}} = E[\boldsymbol{\phi}(\boldsymbol{\eta})], \quad \frac{d^2 A(\boldsymbol{\eta})}{d\boldsymbol{\eta}^2} = \text{Var}[\boldsymbol{\phi}(\boldsymbol{\eta})]$$

- $A(\boldsymbol{\eta})$ is convex
- The generalized maximum likelihood cost to minimize is

$$\min_{\boldsymbol{\beta}} \mathcal{L}(\boldsymbol{\beta}) = -\sum_{n=1}^N \log(p(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

- where $p(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})$ is an exponential family distribution
- We obtain the solution

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \tilde{\mathbf{X}}^T [\mathbf{g}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\mathbf{y})]$$

5 k-Nearest Neighbor (k-NN)

- The k-NN prediction for \mathbf{x} is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$$

where $nbh_k(\mathbf{x})$ is the neighborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n in the training data

- **Curse of dimensionality**: Generalizing correctly becomes exponentially harder as the dimensionality grows.
- Gathering more inputs variables may be a bad thing

6 Kernel Ridge Regression

- The following is true for ridge regression

$$\boldsymbol{\beta} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y} \quad (1)$$

$$= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} \quad (2)$$

$$= \mathbf{X}^T \boldsymbol{\alpha}$$

- Complexity of computing $\boldsymbol{\beta}$

1. $O(D^2 N + D^3)$
2. $O(N^2 D + N^3)$

- Thus we have

$$\boldsymbol{\beta} = \sum_{n=1}^N \alpha_n \mathbf{x}_n, \quad \mathbf{y} = \sum_{d=1}^D \beta_d \tilde{\mathbf{x}}_d$$

with \mathbf{x}_n the rows of \mathbf{X} and $\tilde{\mathbf{x}}_d$ the columns of \mathbf{X}

- The representer theorem allows us to write an equivalent optimization problem in terms of $\boldsymbol{\alpha}$.

$$\boldsymbol{\alpha} = \arg\max_{\boldsymbol{\alpha}} \left(-\frac{1}{2} \boldsymbol{\alpha} (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^T \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

- Kernelized ridge regression might be computationally more efficient in some cases.
- **Kernel trick**:

- We can work directly with $\mathbf{K} = \mathbf{X} \mathbf{X}^T$ and never have to worry about \mathbf{X}
- Using the basis function $\boldsymbol{\phi}(\mathbf{x})$, we do not need to specify it explicitly, since we can work directly with $\mathbf{K} = \boldsymbol{\phi}(\mathbf{x})$
- We will use a kernel function $k(\mathbf{x}, \mathbf{x}') = \boldsymbol{\phi}(\mathbf{x})^T \boldsymbol{\phi}(\mathbf{x}')$
- The evaluation of a kernel is usually faster with k than with $\boldsymbol{\phi}$

- **Radial Basis function kernel (RBF)**

$$k(\mathbf{x}, \mathbf{x}') = \exp\left(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T (\mathbf{x} - \mathbf{x}')\right)$$

- Properties of a kernel to ensure the existence of a corresponding $\boldsymbol{\phi}$:

- \mathbf{K} should be symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$
- \mathbf{K} should be positive semidefinite.

7 Support Vector Machine

- **Hinge loss**

$$[t]_+ = \max(0, t)$$

- Solution to the dual problem is sparse and non-zero entries will be our **support vectors**.
- Assume $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$\min_{\boldsymbol{\beta}} \sum_{n=1}^N [1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}]_+ + \frac{\lambda}{2} \sum_{j=1}^M \beta_j^2$$

- The minimum doesn't change with a rescaling of $\boldsymbol{\beta}$

- **Duality**:

- Hard to minimize $g(\boldsymbol{\beta})$ so we define

$$g(\boldsymbol{\beta}) = \max_{\boldsymbol{\alpha}} G(\boldsymbol{\beta}, \boldsymbol{\alpha})$$

- we use the property that

$$C[v_n]_+ = \max(0, C v_n) = \max_{\alpha_n \in [0, C]} \alpha_n v_n$$

- We can rewrite the problem as

$$\min_{\boldsymbol{\beta}} \max_{\boldsymbol{\alpha} \in [0, C]^N} \sum_{n=1}^N \alpha_n (1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}) + \frac{1}{2} \sum_{j=1}^M \beta_j^2$$

- This is differentiable, convex in $\boldsymbol{\beta}$ and concave in $\boldsymbol{\alpha}$
- **Minimax theorem**:

$$\min_{\boldsymbol{\beta}} \max_{\boldsymbol{\alpha}} G(\boldsymbol{\beta}, \boldsymbol{\alpha}) = \max_{\boldsymbol{\alpha}} \min_{\boldsymbol{\beta}} G(\boldsymbol{\beta}, \boldsymbol{\alpha})$$

because G is convex in $\boldsymbol{\beta}$ and concave in $\boldsymbol{\alpha}$.

- Derivative w.r.t. $\boldsymbol{\beta}$:

$$\frac{dG}{d\boldsymbol{\beta}} = -\left(\sum_{n=1}^N \alpha_n y_n \tilde{\boldsymbol{\phi}}_n \right) + \begin{bmatrix} 0 \\ \boldsymbol{\beta}_{1:M} \end{bmatrix}$$

- Equating this to 0, we get:

$$\boldsymbol{\beta}_{1:M}^* = \sum_{n=1}^N \alpha_n y_n \boldsymbol{\phi}_n = \boldsymbol{\Phi}^T \text{diag}(\mathbf{y}) \boldsymbol{\alpha}$$

$$\boldsymbol{\alpha}^T \mathbf{y} = 0$$

- Plugging $\boldsymbol{\beta}^*$ back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0, C]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2} \boldsymbol{\alpha}^T \mathbf{Y} \boldsymbol{\Phi} \boldsymbol{\Phi}^T \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \boldsymbol{\Phi} \boldsymbol{\Phi}^T$
- The solution $\boldsymbol{\alpha}$ is sparse and is non-zero only for the training examples that are instrumental in determining the decision boundary.

8 K-means

- **Unsupervised learning**: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input patterns.
- **Cluster** are groups of points whose inter-point distances are small compared to the distances outside the cluster.

$$\min_{\mathbf{r}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{r}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|_2^2$$

such that $r_{nk} \in \{0, 1\}$ and $\sum_{k=1}^K r_{nk} = 1$

- K-means algorithm:

- Initialize $\boldsymbol{\mu}_k$, then iterate

1. For all n , compute \mathbf{r}_n given $\boldsymbol{\mu}$

$$r_{nk} = \begin{cases} 1 & \text{if } k = \arg\min_j \|\mathbf{x}_n - \boldsymbol{\mu}_j\|_2^2 \\ 0 & \text{otherwise} \end{cases}$$

2. For all k , compute $\boldsymbol{\mu}_k$ given \mathbf{r}

$$\boldsymbol{\mu}_k = \frac{\sum_{n=1}^N r_{nk} \mathbf{x}_n}{\sum_{n=1}^N r_{nk}}$$

- A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{r}, \boldsymbol{\mu}) = \prod_{n=1}^N \prod_{k=1}^K [\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \mathbf{I})]^{r_{nk}}$$

- Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

9 Gaussian Mixture Models

- Clusters can be spherical using a full covariance matrix instead of isotropic covariance.

$$p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{r}) = \prod_{n=1}^N \prod_{k=1}^K [\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)]^{r_{nk}}$$

- **Soft-clustering**: Points can belong to several cluster by defining r_n to be a random variable.

$$p(r_{nk} = 1) = \pi_k \text{ where } \pi_k > 0, \forall k$$

$$\sum_{k=1}^K \pi_k = 1$$

- Joint distribution of Gaussian mixture model

$$p(\mathbf{X}, \mathbf{r} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^N [p(\mathbf{x}_n | r_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(r_n | \boldsymbol{\pi})]$$

$$= \left[\prod_{k=1}^K [(\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k))^{r_{nk}}] \prod_{k=1}^K [\pi]^{r_{nk}} \right]$$

$$\text{Joint} = \text{Likelihood} \times \text{Prior}$$

- r_n are called *latent* unobserved variables
- Unknown parameters are given by $\boldsymbol{\theta} = \{\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}\}$
- We get the **marginal likelihood** by marginalizing r_n out

from the likelihood

$$\begin{aligned}p(\mathbf{x}_n | \boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, r_n = k | \boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(r_n = k | \boldsymbol{\theta}) p(\mathbf{x}_n | r_n = k, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)\end{aligned}$$

- Without a latent variable model, number of parameters grow at rate $O(N)$
- After marginalization, the growth is reduced to $O(D^2 K)$
- To get maximum likelihood estimate of $\boldsymbol{\theta}$, we maximize

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^N \log \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\sigma}_k)$$

10 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\boldsymbol{\theta}^{(1)}$ and iterate

1. *Expectation step*: Compute a lower bound to the cost such that it is tight at the previous $\boldsymbol{\theta}^{(i)}$

$$\log \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \geq \sum_{k=1}^K p_{kn} \log \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{p_{kn}}$$

with equality when,

$$p_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$

2. *Maximization step*: Update $\boldsymbol{\theta}$

$$\boldsymbol{\theta}^{(i+1)} = \arg\max_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(i)})$$

$$\boldsymbol{\mu}_k^{i+1} = \frac{\sum_{n=1}^N p_{kn}^{(i)} \mathbf{x}_n}{\sum_{n=1}^N p_{kn}^{(i)}}$$

$$\boldsymbol{\Sigma}_k^{(i+1)} = \frac{\sum_{n=1}^N p_{kn}^{(i)} (\mathbf{x}_n - \boldsymbol{\mu}_k^{(i+1)}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{(i+1)})^T}{\sum_{n=1}^N p_{kn}^{(i)}}$$

$$\pi_k^{(i+1)} = \frac{1}{N} \sum_{n=1}^N p_{kn}^{(i)}$$

- If the covariance is diagonal, then we have K-means.

11 Matrix factorization

- We have D movies and N users
- \mathbf{X} is a matrix $D \times N$ with x_{dn} the rating of n 'th user for d 'th movie.
- We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
 - \mathbf{Z} a $N \times M$ matrix that gives features for the users
 - \mathbf{W} a $D \times M$ matrix that gives features for the movies

$$x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$$

- We can add a regularizer and minimize the following cost:

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{n=1}^N \sum_{d=1}^D (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n)^2 + \frac{\lambda_w}{2} \sum_{d=1}^D \mathbf{w}_d^T \mathbf{w}_d + \frac{\lambda_z}{2} \sum_{n=1}^N \mathbf{z}_n^T \mathbf{z}_n$$

- We can use coordinate descent algorithm, by first minimizing w.r.t. \mathbf{Z} given \mathbf{W} and then minimizing \mathbf{W} given \mathbf{Z} . This is called **Alternating least-squares (ALS)**:

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_M)^{-1} \mathbf{W}^T \mathbf{X}$$

$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_M)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- *Complexity*: $O(DNM^2 + NM^3) \rightarrow O(DNM^2)$
- Probabilistic model

$$\prod_{n=1}^N \prod_{d \in \mathcal{O}_n} \mathcal{N}(x_{dn} | \mathbf{w}_d^T \mathbf{z}_n, I) \times \prod_{n=1}^N \mathcal{N}(\mathbf{z}_n | 0, \frac{1}{\lambda_z} I) \times \prod_{d=1}^D \mathcal{N}(\mathbf{w}_d | 0, \frac{1}{\lambda_w} I)$$

- Since many ratings are missing we cannot normalize the

data. A solution is to add offset terms:

$$\frac{1}{2} \sum_{n=1}^N \sum_{d \in O_n} (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu)^2$$

12 Singular Value Decomposition

- Matrix factorization method

$$\mathbf{X} = \mathbf{U} \mathbf{S} \mathbf{V}^T$$

- \mathbf{U} is an $D \times D$ matrix
- \mathbf{V} is an $N \times N$ matrix
- \mathbf{S} is a non-negative diagonal matrix of size $D \times N$ which are called **singular values** appearing in a descending order.
- Columns of \mathbf{U} and \mathbf{V} are the left and right **singular vectors** respectively.
- Assuming $D < N$ we have

$$\mathbf{X} = \sum_{d=1}^D s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of \mathbf{X} where higher singular vectors contain the *low-frequency information* and lower singular values contain the *high-frequency information*.

13 Principal Component Analysis

- PCA is a dimensionality reduction method and a method to decorrelate the data

$$\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W} \mathbf{Z}^T$$

such that columns of \mathbf{W} are orthogonal.

- If the data is zero mean

$$\mathbf{C} = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$

$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Thus the columns of matrix \mathbf{U} are called the **principal components** and they decorrelate the covariance matrix.
- Using SVD, we can compute the matrices in the following way

$$\mathbf{W} = \mathbf{U} \mathbf{S}^{1/2}$$

$$\mathbf{Z} = \mathbf{V} \mathbf{S}^{1/2}$$

14 Multi-Layer Perceptron (MLP)

- Known as **feed-forward neural network**

- $\mathbf{z}_n^{(k)}$ is the k 'th hidden vector

- $\mathbf{a}_n^{(k)}$ is the corresponding activation

- There are a total of K layers

$$a_{mn}^{(k)} = (\beta_m^{(k)})^T z_n^{(k-1)}, \quad z_{mn}^{(k)} = h(a_{mn}^{(k)})$$

- For the first layer we have $\mathbf{z}_n^{(0)} = \mathbf{x}_n$

- For the last layer, we use a link function to map $\mathbf{z}_n^{(K-1)}$ to the output \mathbf{y}_n

- A 1-layer MLP is simply a generalization of linear/logistic regression

- $\mathbf{B}^{(k)}$ a matrix with rows $(\beta_m^{(k)})^T$

$$\mathbf{a}_n^{(k)} = \mathbf{B}^{(k)} \mathbf{z}_n^{(k-1)}, \quad \mathbf{z}_n^{(k)} = h(\mathbf{a}_n^{(k)})$$

- thus we have the input-output relationship

$$\hat{\mathbf{y}}_n = g((\beta^{(K-1)})^T * h(\mathbf{B}^{(K-2)} * h(* \dots * h(\mathbf{B}^{(1)} * \mathbf{x}_n)))$$

with $g()$ the link function

- We learn parameters \mathbf{B} using stochastic gradient-descent
- Frequently used transfer function

$$g(a) = \tanh(a) = \frac{e^a - e^{-a}}{e^a + e^{-a}}$$

- Backpropagation** is a technic to compute the gradient in time linear in the number of training points and the number of weights.

15 Gaussian Process (GP)

- Let us place a probabilistic prior shape on the approximation of a function.

- A GP process defines a prior over function f

$$p(f|X) = \mathcal{N}(f|0, K(X))$$

- $K(X)$ defines shape and prior knowledge about our problem
- RBF kernel

$$k(\mathbf{x}_n, \mathbf{x}_m) = e^{-||\mathbf{x}_n - \mathbf{x}_m||^2 / L^2}$$

- Quadratic kernel

$$k(\mathbf{x}_n, \mathbf{x}_m) = (1 + \mathbf{x}_n^T \mathbf{x}_m)^2$$

16 Decision Trees

- Fast to train and fast to make predictions
- Efficient for very high dimensional feature spaces and very large amounts of training data
- Lack of smoothness and high variance
- Goal: find a split (k, τ) that minimizes an impurity measure at the leaves
-

17 Random Forests