PCML Cheat Sheet

1 Math Prerequisites

- Bayes rule

$$p(A,B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

Gaussian distribution

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}))$$
 robust to outliers hard to set δ .

$$V(XY) = E(X^2)E(Y^2) - [E(X)]^2[E(Y)]^2$$

- Log-properties

$$\log(mn) = \log(m) + \log(n)$$

$$\log(m^n) = n \log(m)$$

- Covariance matrix of a data vector x

$$\Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - E(\mathbf{x})(\mathbf{x}_n - E(\mathbf{x})^T)$$

1.1 Convexity

- A function is convex when a line joining two points never interects with the function
- A function f(x) is convex, if for any $x_1, x_2 \in \mathbf{X}$ and for any $0 \le \lambda \le 1$, we have

$$f(\lambda x_1 + (1-\lambda)x_2) \leq \lambda f(x_1) + (1-\lambda)f(x_2)$$

- A function is strictly convex if the inequality is
- A convex function has only one global minimum.
- Sums of convex functions are also convex. The Hessian is related to the convexity of a
- function: a twice differentiable function is convex if and only if the Hessian is positive
- The Hessian of a convex function is positive semi-definite and for a strictly-convex function it is positive definite.
- The Hessian matrix of a function

$$\mathbf{H}_{i,j} = \frac{d^2f}{dx_i dx_j}$$

1.2 Linear Algebra

- Column $\mathbf{x} \in \mathbb{R}^n$, rows \mathbf{x}^T , matrix $\mathbf{A} \in \mathbb{R}^{m \times n}$
- $\mathbf{x}^T \mathbf{x}$ is a scalar, $\mathbf{x} \mathbf{x}^T$ is a matrix $= A^{-1}$ exist if A is full rank
- Condition number of a function measures how much the output value can change for a small change in the input argument. A matrix with a high condition number is said to be

ill-conditioned. If **A** is normal $(A^T A = AA^T)$

$$k(\mathbf{A}) = \left| \frac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

- A positive definite matrix is symmetric with all positive eigenvalues
- The real symmetric $N \times N$ matrix **V** is said to be positive semidefinite if

$$\mathbf{a}^{-} \mathbf{Va} \geq 0$$

for any real $N \times 1$ vector a .

The real symmetric $N \times N$ matrix **V** is said to be positive definite if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

for any real $N \times 1$ vector a.

- Cost of matrix inversion: $O(n^3) \rightarrow O(n^{2.372})$
- Cost of determinant computation using LU decomposition: $O(n^3)$

2 Cost functions

- Cost functions are used to learn parameters that explain the data well.
- It is essential to make sure that a global minimum exist → lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{\beta}) = \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is convex thus it has only one global minumum value.
- MSE is not good when outliers are present

Mean Absolute Error (MAE):

$$MAE = \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

Huber los

$$Huber = \left\{ \begin{array}{cc} \frac{1}{2}z^2 & , |z| \leq \delta \\ \delta|z| - \frac{1}{2}\delta^2 & , |z| > \delta \end{array} \right.$$

Huber loss is convex, differentiable, and also

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Tukey's loss is non-convex, non-differentiable, but robust to outliers.

Hinge loss

$$Hinge = [1 - y_n f(\mathbf{x}_n)]_+$$

Logistic loss

$$Logistic = \log(1 - \exp(y_n f(\mathbf{x}_n)))$$

3 Regression

- Data consists of N pairs (y_n, \mathbf{x}_n)
- 1. y_n the n'th output 2. \mathbf{x}_n is a vector of D inputs
- Prediction: predict the ouput for a new input
- Interpretation: understand the effect of inputs on output.
- Outliers are data that are far away from most of the other examples

3.1 Linear Regression

Model that assume linear relationship between inputs and the ouput

$$y_n \equiv f(\mathbf{x}_n)$$

$$:= \beta_0 + \beta_1 x_{n1} + \dots$$

$$= \mathbf{x}_n^T \boldsymbol{\beta}$$

- with β the parameters of the model.
- Variance grows only linearly with dimensionality

3.2 Gradient Descent

- Gradient descent uses only first-order information and takes steps in the direction of
- Given a cost function $\mathcal{L}(\beta)$ we wish to find β that minimizes the cost:

$$\min_{\boldsymbol{\beta}} \mathcal{L}(\boldsymbol{\beta})$$

3.2.1 Grid search

- Compute the cost over a grid of M points to find the minimum
- Exponential Complexity $O(NDM^D)$
- Hard to find a good range of values

3.3 Batch Gradient Descent

- Take steps in the opposite direction of the

$$\boldsymbol{\beta}^{(k+1)} \leftarrow \boldsymbol{\beta}^{(k)} - \alpha \frac{d\mathcal{L}(\boldsymbol{\beta}^{(k)})}{d\boldsymbol{\beta}}$$

with $\alpha > 0$ the learning rate.

With α too big, method might diverge. With α too small, convergence is slow.

3.4 Gradients for MSE

$$\tilde{\mathbf{X}} = \begin{bmatrix} 1 & \mathbf{X} \end{bmatrix}$$

We define the error vector e:

$$e = y - \tilde{X}\beta$$

and MSE as follows:

$$\mathcal{L}(\boldsymbol{\beta}) = \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = -\frac{1}{N}\tilde{\mathbf{X}}^T\mathbf{e}$$

- Optimality conditions:

1. necessary: gradient equal to zero:
$$\frac{d\mathcal{L}(\boldsymbol{\beta}^*)}{d\boldsymbol{\beta}} = 0$$

2. sufficient: Hessian matrix is positive definite: $\mathbf{H}(\boldsymbol{\beta}^*) = \frac{d^2 \mathcal{L}(\boldsymbol{\beta}^*)}{}$ dBdBT

- Very sensitive to illconditioning. Therefore, always normalize your feature otherwise step-size selection is difficult since different directions might move at different speed.
- Complexity: O(NDI) with I the number of

3.5 Least Squares

- In some cases, we can compute the minimum of the cost function analytically.
- use the first optimality conditions:

$$\frac{d\mathcal{L}}{d\boldsymbol{\beta}} = 0 \Rightarrow \tilde{\mathbf{X}}^T \mathbf{e} = \tilde{\mathbf{X}}^T (\mathbf{y} - \tilde{\mathbf{X}} \boldsymbol{\beta}) = 0$$

When $\tilde{\mathbf{X}}^T \tilde{\mathbf{X}}$ is invertible, we have the closed-form expression

$$\boldsymbol{\beta}^* = (\mathbf{\tilde{X}}^T \mathbf{\tilde{X}})^{-1} \mathbf{\tilde{X}}^T \mathbf{y}$$

- thus we can predict values for a new x*

$$y_* = \tilde{\mathbf{x}}_*^{\mathbf{T}} \boldsymbol{\beta}^* = \tilde{\mathbf{x}}_*^{\mathbf{T}} (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

- The Gram matrix $\tilde{\mathbf{X}}^T \tilde{\mathbf{X}}$ is positive definite and is also invertible iff $\tilde{\mathbf{X}}$ has full column rank.

Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$ $\tilde{\mathbf{X}}$ can be rank deficient when D > N or when

the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. In this case, the matrix is ill-conditioned, leading to numerical issues

3.6 Maximum Likelihood

Let define our mistakes ε_n ~ N(0, σ²).

$$\rightarrow y_n = \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} + \epsilon_n$$

- Another way of expressing this:

$$p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta}) = \prod_{n=1}^{N} p(y_n | \tilde{\mathbf{x}}_n, \boldsymbol{\beta})$$
$$= \prod_{n=1}^{N} \mathcal{N}(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}, \sigma^2)$$

which defines the likelihood of observating y given $\tilde{\mathbf{X}}$ and $\boldsymbol{\beta}$

Define cost with log-likelihood

$$\mathcal{L}_{lik}(\boldsymbol{\beta}) = \log p(\mathbf{y}|\tilde{\mathbf{X}}, \boldsymbol{\beta})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})^2 + cnst$$

another way to design cost functions

$$\underset{\boldsymbol{\beta}}{\operatorname{argmin}} \mathcal{L}_{mse}(\boldsymbol{\beta}) = \underset{\boldsymbol{\beta}}{\operatorname{argmax}} \mathcal{L}_{lik}(\boldsymbol{\beta})$$

MLE can also be interpreted as finding the model under which the obeserved data is most likely to have been generated from.

With Laplace distribution $p(y_n|\tilde{\mathbf{x}}_n,\boldsymbol{\beta}) = \frac{1}{2b}e^{-\frac{1}{b}|y_n - \tilde{\mathbf{x}}_n^T\boldsymbol{\beta}|}$

$$\sum \log p(y_n | \mathbf{\tilde{x}}_n, \boldsymbol{\beta}) = \sum |y_n - \mathbf{\tilde{x}}_n^T \boldsymbol{\beta}| + cnst$$

3.7 Ridge Regression

- Linear models usually underfit. One way is to use nonlinear basis functions instead

$$y_n = \beta_0 + \sum_{j=1}^{M} \beta_j \phi_j(\mathbf{x}_n) = \tilde{\phi}(\mathbf{x}_n)^T \boldsymbol{\beta}$$

This model is linear in \$\beta\$ but nonlinear in \$\boldsymbol{x}\$.

Note that the dimensionality is now M, not D. Polynomial basis

 $\phi(x_n) = [1, x_n, x_n^2, ..., x_n^M]$ - The least square solution becomes

 $\boldsymbol{\beta}_{lse}^* = (\mathbf{\tilde{\Phi}}^T \mathbf{\tilde{\Phi}})^{-1} \mathbf{\tilde{\Phi}}^T \mathbf{v}$ Complex models overfit easily. Thus we can choose simpler models by adding a regularization term which penalizes complex

$$\min_{\boldsymbol{\beta}} \left(\mathcal{L}(\boldsymbol{\beta}) + \frac{\lambda}{2N} \sum_{j=1}^{M} \beta_j^2 \right)$$

$$\boldsymbol{\beta}^* = \underset{\boldsymbol{\beta}}{\operatorname{argmin}} \left(\frac{1}{2} (\mathbf{y} - \mathbf{X} \boldsymbol{\beta})^T (\mathbf{y} - \mathbf{X} \boldsymbol{\beta}) + \frac{\lambda}{2} \boldsymbol{\beta}^T \boldsymbol{\beta} \right)$$

Note that β_0 is not penalized.

- By differentiating and setting to zero we get

$$oldsymbol{eta}_{ridge} = (\mathbf{ ilde{\Phi}}^T\mathbf{ ilde{\Phi}} + \mathbf{\Lambda})^{-1}\mathbf{ ilde{\Phi}}^T\mathbf{y}$$
 $\mathbf{\Lambda} = \left[egin{array}{ccc} 0 & 0 & 0 \ 0 & \lambda I_m \end{array}
ight]$

Ridge regression improves the condition number of the Gram matrix since the eigenvalues of $({ ilde \Phi}^T { ilde \Phi} + \lambda I_m)$ are at least λ

Maximum-a-posteriori (MAP) estimator:

- Maximizes the product of the likelihood and the prior.

$$\boldsymbol{\beta}{map} = \operatorname*{argmax}_{\boldsymbol{\beta}} \left(p(\mathbf{y}|\mathbf{X}, \boldsymbol{\Lambda}) p(\boldsymbol{\beta}|\boldsymbol{\Sigma}) \right)$$

- Assume $\beta_0 = 0$

Assume
$$\beta_0 = 0$$

$$\beta_{ridge} = \underset{\boldsymbol{\beta}}{\operatorname{argmax}} \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T \boldsymbol{\beta}, \boldsymbol{\Lambda}) \times \mathcal{N}(\boldsymbol{\beta} | 0, \mathbf{I}) \right] \right) = -\frac{d\mathbf{g}(\boldsymbol{\beta})}{d\boldsymbol{\beta}^T} = \sum_{n=1}^{N} \frac{d}{d\boldsymbol{\beta}^T} \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) \tilde{\mathbf{x}}_n$$

Lasso regularizer forces some β_i to be strictly 0 and therefore forces sparsity in the model.

$$\min_{\pmb{\beta}} \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\pmb{\phi}}(\mathbf{x}_n)^T \pmb{\beta})^2, \quad \text{ such that } \sum_{i=1}^{M} |\beta_i| \leq \tau$$

- We should choose λ to minimize the mistakes. that will be made in the future.
- We split the data into train and validation sets and we pretend that the validation set is the future data. We fit our model on the training set and compute a prediction-error on the validation set. This gives us an estimate of the generalization error.
- K-fold cross validation randomly partition the data into K groups. We train on K-1groups and test on the remaining group. We repeat this until we have tested on all K sets. We then average the results.
- Cross-validation returns an unbiased estimate of the generalization error and its variance.

3.9 Bias-Variance decomposition

- The expected test error can be expressed as the
- sum of two terms - Squared bias: The average shift of the
- predictions - Variance: measure how data points vary around their average.

expected loss =
$$(bias)^2 + variance + noise$$

- Both model bias and estimation bias are
- important Ridge regression increases estimation bias while
- reducing variance - Increasing model complexity increases test error

Small $\lambda \to \text{low}$ bias but large variance

Large $\lambda \rightarrow$ large bias but low variance

3.10 Logistic Regression

- Classification relates input variables x to
- discrete output variable y Binary classifier: we use y = 0 for C_1 and
 - Can use least-squares to predict \hat{y}_*

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* > 0.5 \end{cases}$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(y_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$
$$p(y_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\tilde{\mathbf{x}}^T \boldsymbol{\beta})$$

$$p(\mathbf{y}|\mathbf{X}, \boldsymbol{\beta}) = \prod_{n=1}^{N} \sigma(\tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta})^{y_{n}} (1 - \sigma(\tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta})^{1 - y_{n}}$$

- The log-likelihood:

$$\mathcal{L}_{mle}(\boldsymbol{\beta}) = \sum_{n=1}^{N} \left(y_n \tilde{\mathbf{x}}_n^T \boldsymbol{\beta} - \log(1 + \exp(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})) \right)$$

$$\frac{d}{dx}\log(1+\exp(x)) = \sigma(x)$$

$$\begin{split} \mathbf{g} &= \frac{d\mathcal{L}}{d\boldsymbol{\beta}} = \sum_{n=1}^{N} \left(\hat{\mathbf{x}}_{n} y_{n} - \hat{\mathbf{x}}_{n} \sigma(\hat{\mathbf{x}}_{n}^{T} \boldsymbol{\beta}) \right) \\ &= -\hat{\mathbf{X}}^{T} [\sigma(\hat{\mathbf{X}} \boldsymbol{\beta}) - \mathbf{y}] \end{split}$$

The negative of the log-likelihood -L_{mle}(β) is

Hessian of the log-likelihood

- We know that

$$\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t))$$

$$\mathbf{H}(\boldsymbol{\beta}) = -\frac{d\mathbf{g}(\boldsymbol{\beta})}{d\boldsymbol{\beta}^T} = \sum_{n=1}^{N} \frac{d}{d\boldsymbol{\beta}^T} \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}) \tilde{\mathbf{x}}_n$$

 $= \sum_{n=1}^{N} \tilde{\mathbf{x}}_{n} \sigma(\tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta}) (1 - \sigma(\tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta})) \tilde{\mathbf{x}}_{n}^{T} \boldsymbol{\beta}$

where S is a $N \times N$ diagonal matrix with diagonals

$$S_{nn} = \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta})(1 - \sigma(\tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$
- The negative of the log-likelihood is not

strictly convex. Newton's Method Uses second-order information and takes

steps in the direction that minimizes a quadratic approximation
$$\mathcal{L}(\boldsymbol{\beta}) = \mathcal{L}(\boldsymbol{\beta}^{(k)}) + \mathbf{g}_k^T (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)}) + (\boldsymbol{\beta} - \boldsymbol{\beta}^{(k)})^T \mathbf{F}$$

and it's minimum is at
$$oldsymbol{eta}^{k+1} = oldsymbol{eta}^{(k)} - lpha_k \mathbf{H}_k^{-1} \mathbf{g}_k$$

where \mathbf{g}_k is the gradient and α_k the learning rate.

$$\begin{aligned} & - \text{ Complexity: } O((ND^2 + D^3)I) \\ & \textbf{Penalized Logistic Regression} \\ & \min_{\pmb{\beta}} \left(-\sum_{n=1}^{N} \log p(y_n | \mathbf{x}_n^T \pmb{\beta}) + \lambda \sum_{d=1}^{D} \boldsymbol{\beta}_d^2 \right) \end{aligned}$$

4 Generalized Linear Model

defined using a link function q

Exponential family distribution
$$p(\mathbf{y}|\boldsymbol{\eta}) = \frac{h(y)}{\pi} \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

- Bernoulli distribution
$$p(y|\mu) = \mu^{y}(1-\mu)^{1-y}$$

$$= \exp(y \log(\frac{\mu}{1-\mu} + \log(1-\mu)))$$

- there is a relationship between η and μ throught the link function

ght the link function
$$\eta = \log(\frac{\mu}{1 - \mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1 + e^{\eta}}$$

 Note that μ is the mean parameter of y Relationship between the mean μ and η is

$$\eta = g(\mu) \Leftrightarrow \mu = g^{-1}(\eta)$$

First and second derivatives of $A(\eta)$ are related to the mean and the variance

 $\frac{dA(\eta)}{d\eta} = E[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = Var[\phi(\eta)]$

- $A(\eta)$ is convex - The generalized maximum likelihood cost to

mize is
$$\min_{\boldsymbol{\beta}} \mathcal{L}(\boldsymbol{\beta}) = -\sum_{n=1}^{N} \log(p(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta}))$$

where $p(y_n | \tilde{\mathbf{x}}_n^T \boldsymbol{\beta})$ is an exponential family

- We obtain the solution

$$\frac{d\mathcal{L}}{d\theta} = \tilde{\mathbf{X}}^T [\mathbf{g}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\mathbf{y})]$$

5 k-Nearest Neighbor (k-NN)

- The k-NN prediction for x is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$$

where $nbh_k(\mathbf{x})$ is the neightborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n in the training data

- Curse of dimensionality: Generalizing correctly becomes exponentially harder as the dimensionality grows.
- Gathering more inputs variables may be a bad

6 Kernel Ridge Regression

- The following is true for ridge regression

$$\beta = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y}$$
(1)
= $\mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y}$ (2)

$$= \mathbf{X}^T \boldsymbol{\alpha}$$

Complexity of computing β

- 1. $O(D^2N + D^3)$
- 2. $O(N^2D + N^3)$ Thus we have

$$\boldsymbol{\beta} = \sum_{n=1}^{N} \alpha_n \mathbf{x}_n, \quad \mathbf{y} = \sum_{d=1}^{D} \beta_d \bar{\mathbf{x}}_d$$

The representer theorem allows us to write an equivalent optimization problem n terms of α .

$$\boldsymbol{\alpha} = \operatorname*{argmax}_{\boldsymbol{\alpha}} \left(-\frac{1}{2} \boldsymbol{\alpha} (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^T \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

- XX^T is called the kernel matrix
- Kernelized rigde regression might be
- computationally more efficient in some cases.

Kernel trick:

- We can work directly with $\mathbf{K} = \mathbf{X}\mathbf{X}^T$ and never have to worry about X
- Using the basis function $\phi(\mathbf{x})$, we do not need to specify it explicitly, since we can work directly with $\mathbf{K} = \phi(\mathbf{x})$
- We will use a kernel function $k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$
- The evaluation of a kernel is usually faster with k than with d

Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Properties of a kernel to ensure the existance of a corresponding φ:

K should be symmetric: k(x, x') = k(x', x)

- K should be positive semidefinite.

7 Support Vector Machine

- Hinge loss

$$[t]_{+} = \max(0, t)$$

$$[t]_+ = \max(0, t)$$

- Solution to the dual problem is sparse and non-zero entries will be our support vectors

Assume $y_n \in \{-1, 1\}$

- SVM optimizes the following cost

$$\min_{\boldsymbol{\beta}} \sum_{n=1}^{N} [1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}]_+ + \frac{\lambda}{2} \sum_{j=1}^{M} \beta_j^2$$

- The minimum doesn't change with a rescaling of β

- Duality:

– Hard to minimize $g(\boldsymbol{\beta})$ so we define

$$g(\beta) = \max_{\alpha} G(\beta, \alpha)$$

we use the property that

$$C[v_n]_+ = \max(0, Cv_n) = \max_{\alpha_n \in [0, C]} \alpha_n v_n$$

- We can rewrite the problem as

$$\min_{\boldsymbol{\beta}} \max_{\boldsymbol{\alpha} \in [0,C]^N} \sum_{n=1}^N \alpha_n (1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{\beta}) + \frac{1}{2} \sum_{j=1}^M \beta_j^2$$

– This is differentiable, convex in β and

- Minimax theorem:

$$\min_{\pmb{\beta}} \max_{\pmb{\alpha}} G(\pmb{\beta}, \pmb{\alpha}) = \max_{\pmb{\alpha}} \min_{\pmb{\beta}} G(\pmb{\beta}, \pmb{\alpha})$$

because G is convex in $oldsymbol{eta}$ and concave in $oldsymbol{lpha}$

$$\frac{dG}{d\boldsymbol{\beta}} = -\left(\sum_{n=1}^{N} \alpha y_n \tilde{\phi_n}\right) + \begin{bmatrix} 0 \\ \boldsymbol{\beta}_{1:M} \end{bmatrix}$$

$$\boldsymbol{\beta}_{1:M}^* = \sum_{n=1}^N \alpha_n y_n \phi_n = \boldsymbol{\Phi}^T \mathrm{diag}(\mathbf{y}) \boldsymbol{\alpha}$$

Plugging \(\beta^*\) back in the dual problem

$$\max_{oldsymbol{lpha} \in [0,C]^N} oldsymbol{lpha}^T \mathbf{1} - rac{1}{2} oldsymbol{lpha}^T \mathbf{Y} oldsymbol{\Phi} \mathbf{Y} oldsymbol{lpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{\Phi} \mathbf{\Phi}^T$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision

8 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of
- Cluster are groups of points whose inter-point distances are small compared to the distances

$$\min_{\mathbf{r}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{r}, \boldsymbol{\mu}) = \sum_{k=1}^{K} \sum_{n=1}^{N} r_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

such that $r_{nk} \in \{0,1\}$ and $\sum_{k=1}^{K} r_{nk} = 1$ K-means algorithm:

Initialize μ_k , then iterate

1. For all n, compute \mathbf{r}_n given $\boldsymbol{\mu}$

$$r_{nk} = \left\{ \begin{array}{ll} 1 & \text{if } k = \mathop{\mathrm{argmin}}_j ||\mathbf{x}_n - \boldsymbol{\mu}||_2^2 \\ 0 & \text{otherwise} \end{array} \right.$$

2. For all k, compute μ_k given ${\bf r}$

$$\mu_k = \frac{\sum_{n=1}^{N} r_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} r_{nk}}$$

 A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.

Probabilistic model

$$p(\mathbf{r}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{r_{nk}}$$

Computation can be heavy, each example can belong to only on cluster and clusters have to

9 Gaussian Mixture Models

Clusters can be spherical using a full covariance matrix instead of isotropic covariance.

$$p(\mathbf{X}|\boldsymbol{\mu},\boldsymbol{\Sigma},\mathbf{r}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k},\boldsymbol{\Sigma}_{k}) \right]^{r_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining r_n to be a random variable

$$p(r_{nk} = 1) = \pi_k$$
 where $\pi_k > 0, \forall k$

$$\sum_{k=1}^{K} \pi_k = 1$$

Joint distribution of Gaussian mixture model

 r_n are called *latent* unobserved variables Unknown parameters are given by

We get the marginal likelihood by marginalizing r_n out from the likelihood

$$\begin{split} p(\mathbf{x}_n|\boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, r_n = k|\boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(r_n = k|\boldsymbol{\theta}) p(\mathbf{x}_n|r_n = k, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{split}$$

- parameters grow at rate O(N)
- After marginalization, the growth is reduced to
- To get maximum likelihood estimate of θ , we

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\sigma}_k)$$

10 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate

1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous

$$\log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \ge \sum_{k=1}^{K} p_{kn} \log \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{p_{k\text{With } \mathbf{T}_{t_n}}}$$

$$p_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n, \mathbf{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \mathbf{\Sigma}_k)}$$
 2. Maximization step: Update $\boldsymbol{\theta}$

$$\theta^{(i+1)} = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(i)})$$

$$\mu_k^{i+1} = \frac{\sum_{n=1}^{N} p_{kn}^{(i)} \mathbf{x}_n}{\sum_{n=1}^{N} p_{kn}^{(i)}}$$

$$\boldsymbol{\Sigma}_{k}^{(i+1)} = \frac{\sum_{n=1}^{N} p_{kn}^{(i)}(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(i+1)})}{\sum_{n=1}^{N} p_{kn}^{(i)}} \text{Thus the columns of matrix } \boldsymbol{U} \text{ are called the principal components and they decorrelate the covariance matrix.}$$

$$\pi_k^{(i+1)} = \frac{1}{N} \sum_{n=1}^{N} p_{kn}^{(i)}$$

If the covariance is diagonal, then we have

11 Matrix factorization

- We have D movies and N users

 X is a matrix D × N with x_{dn} the rating of n'th user for d'th movie.

We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$

We have now 2 latent variables:

- \mathbf{Z} a $N \times M$ matrix that gives features for

the users - W a D × M matrix that gives features for the movies

$$x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$$

We can add a regularizer and minimize the

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{n=1}^{N} \sum_{d=1}^{D} (x_{dn} - \mathbf{w}_{d}^{T} \mathbf{z}_{n})^{2} + \frac{\lambda_{w}}{2} \sum_{d=1}^{D} \sqrt{\frac{k_{d}^{N}}{m_{d}^{N}}} \sqrt{\frac{k_{d}^{N}}{m_{d}^{N}}}} \sqrt{\frac{k_{d}^{N}}{m_{d}^{N}}} \sqrt{\frac{k_{d}^{N}}{m_{d}^{N}}} \sqrt{\frac{k_{d}^{N}}{m_{d}^{N}}}} \sqrt{\frac{k_{d}$$

We can use coordinate descent algorithm, by first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T\mathbf{W} + \lambda_z \mathbf{I}_M)^{-1} \mathbf{W}^T \mathbf{X}$$

$$\mathbf{w}^T \leftarrow (\mathbf{z}^T \mathbf{z} + \lambda_w \mathbf{I}_M)^{-1} \mathbf{z}^T \mathbf{x}^T$$
Complexity, $O(DNM^2 + NM^3) \rightarrow O(DNM^2$

 $\prod_{n=1}^{N} \prod_{d \in O_{n}} \mathcal{N}(x_{dn} | \mathbf{w}_{d}^{T} \mathbf{z}_{n}, I) \times \prod_{n=1}^{N} \mathcal{N}(\mathbf{z}_{n} | 0, \frac{1}{\lambda_{z}} I) \times \prod_{d \in \mathcal{D}_{n}} \mathcal{N}(\mathbf{w}_{d}^{H} | 0, \frac{1}{K_{\lambda_{d}^{H}}}) \xrightarrow{T} \sum_{n=1}^{N} \mathcal{N}(\mathbf{w}_{d}^{H} | 0, \frac{1}{K_{\lambda_{d}^{H}}}) \xrightarrow{T} \sum_{n=1}^{N} \mathcal{N}(\mathbf{z}_{n} | 0, \frac{1}{K_{\lambda_{d}^{H}}}) \times \prod_{n=1}^{N} \mathcal{N}(\mathbf{$

normalize the data. A solution is to add offset

$$\frac{1}{2} \sum_{n=1}^{N} \sum_{d \in O_n} (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu)^2$$

12 Singular Value Decomposition

Matrix factorization method

$$\mathbf{x} = \mathbf{u}\mathbf{s}\mathbf{v}^T$$

- \mathbf{U} is an $D \times D$ matrix
- **V** is an $N \times N$ matrix
- S is a non-negative diagonal matrix of size $D \times N$ which are called singular values appearing in a descending order
- Columns of U and V are the left and right singular vectors respectively.
- Assuming D < N we have

$$\mathbf{X} = \sum_{l=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information.

Let's now truncate these matrices

$$\mathbf{X} pprox \mathbf{U}_{tr} \mathbf{S}_{tr} \mathbf{V}_{tr}^*$$
 $\mathbf{T}_{tr} pprox \mathbf{U}_{tr} \mathbf{S}_{tr}$
 $\mathbf{T}_{te} pprox \mathbf{X}_{te} \mathbf{V}_{tr}$

 $p_{k, \mathbf{W}}$ ith \mathbf{T}_{tr} the reduced feature set of \mathbf{X} and \mathbf{T}_{te} the reduced feature set of \mathbf{X}_{te}

13 Principal Componement Analysis

- PCA is a dimensionality reduction method and a method to decorrelate the data

$$\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$$

such that columns of W are orthogonal.

$$\mathbf{C} = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$

$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

the covariance matrix.

– Using SVD, we can compute the matrices in the following way
$$W = U S^{1/2} \label{eq:W}$$

$$z = vs^{1/2}$$

14 Belief Propagation

following way

- the goal is to learn inference of the latent variables
- using belief propagation Given a directed acyclic graph G and parameters θ , a Bayesian network defines the joint distribution as follows

$$p_{\theta}(\mathbf{x}) = \prod_{k=1}^{K} p_{\theta}(x_k | \text{parents}_k)$$

We can reduce the complexity by using the structure of the problem and the bayesian rule.

15 Multi-Layer Perceptron (MLP)

- Known as feed-forward neural network

$$-\mathbf{z}_{n}^{(k)}$$
 is the k'th hidden vector $\mathbf{z}_{n}^{T}\mathbf{z}_{n}^{(k)}\mathbf{z}_{n}^{N}\mathbf{z}_{$

 $a_{mn}^{(k)} = (\beta_m^{(k)})^T \mathbf{z}_n^{(k-1)}, \quad z_{mn}^{(k)} = h(a_{mn}^{(k)})$

- For the first layer we have $\mathbf{z}_n^{(0)} = \mathbf{x}_n$ - For the last layer, we use a link function to map

 $\mathbf{z}^{(K-1)}$ to the output \mathbf{y}_n A 1-layer MLP is simply a generalization of linear/logistic regression

- $\mathbf{B}^{(k)}$ a matrix with rows $(\boldsymbol{\beta}_m^{(k)})^T$

$$\mathbf{a}_n^{(k)} = \mathbf{B}^{(k)} \mathbf{z}_n^{(k-1)}, \quad \mathbf{z}_n^{(k)} = h(\mathbf{a}_n^{(k)})$$

with g() the link function

We learn parameters B using stochastic gradient-descent

Frequently used transfer function

$$g(a) = \tanh(a) = \frac{e^a - e^{-a}}{e^a + e^{-a}}$$

- Backpropagation is a technic to compute the gradient in time linear in the number of training points and the number of weights.

16 Gaussian Process (GP)

GP are a family of statistical distributions in which time plays a role and for which any finite linear combination of samples has a joint Gaussian distribution.

They can be completely defined by their second-order statistics which means that if they have mean zero, then defining the covariance function completely defines the process behaviour.

Let us place a probabilistic prior shape on the approximation of a function.

A GP process defines a prior over function f

$$p(\mathbf{f}|X) = \mathcal{N}(\mathbf{f}|\mathbf{0}, K(X))$$

- K(X) defines shape and prior knowledge about our problem

$$\begin{split} p(\mathbf{f}|\mathbf{y}, \mathbf{X}_*, \mathbf{X}) &\sim \mathcal{N}(\mu', \sigma') \\ \mu' &= K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]^{-1} \mathbf{y} \\ \sigma' &= K(\mathbf{X}_*, \mathbf{X}_*) - K(\mathbf{X}_*, \mathbf{X})[K(\mathbf{X}, \mathbf{X}) + \sigma_n^2 I]K(\mathbf{X}, \mathbf{X}) \end{split}$$

with σ_n the variance of the noise

$$k(\mathbf{x}_n, \mathbf{x}_m) = e^{-||\mathbf{x}_n - \mathbf{x}_m||^2/L^2}$$

$$k(\mathbf{x}_n, \mathbf{x}_m) = (1 + \mathbf{x}_n^T \mathbf{x}_m)^2$$

17 Decision Trees (DT)

- A decision tree is a structure in which each internal node represents a test on an attribute and each branch represents the outcome of the

test and each leaf represents a class label. Fast to train and fast to make predictions

Efficient for very high dimensional feature spaces and very large amounts of training data

Lack of smoothness and high variance (overfitting) Goal: find a split (k, τ) that minimizes an

impurity measure at the leaves

- Find best feature to split on

Find best threshold

18 Random Forests (RF)

- RF correct the overfitting bad "habit" of DTs - Training: Learn M trees on different subsets

(random) of training data - Prediction: Average of prediction of each tree

$$\begin{split} z_1 &= f_1 \quad \Rightarrow \quad z_M = \frac{1}{M} \sum_{i=1}^M f_i \\ V(z_1) &= \sigma^2 \quad \Rightarrow \quad V(z_M) = \frac{1}{1} \sigma^2 + \rho \frac{M-1}{1} \sigma^2 \end{split}$$

- Variance reduction ratio:

$$\frac{V(z_1)}{} = \frac{M}{}$$

- 2 techniques for decorrelating trees

- Bagging: Randomize training data

- Randomized feature selection