

# A Note on Estimating Elo

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## 1 Introduction

Suppose that two players are playing a game of chess. We want to make some inference on the difference in strength between the two players. Let  $l$ ,  $d$  and  $w$  be the probabilities that player 1 loses, draws or wins respectively. It is not hard to see that player 1 is stronger than player 2 if and only if  $w > l$ . In fact,  $d$  does not matter at all, if we only want to know which of the two players is stronger.

Consider however the two following cases. In the first case,  $l = d = 0$  and  $w = 1$ . And in the second case,  $l = 0$ ,  $d = 0.99$  and  $w = 0.01$ . We see that player 1 is stronger than player 2 in both of these cases, but much more so in the first.

To quantify this strength difference we naturally look at the expected score of a game. If a loss is worth 0 points, a draw  $\frac{1}{2}$  and a win 1, the expected score is given by  $s = \frac{1}{2}d + w$ . The closer  $s$  is to  $\frac{1}{2}$  the more even the games are. And if  $s$  is close to 1 then player 1 wins always every game. We see that in the first case  $s = \frac{1}{2} \cdot 0 + 1 \cdot 1 = 1$ , and in the second case  $s = \frac{1}{2} \cdot 0.99 + 1 \cdot 0.01 = 0.505$ .

We see that player 1 is stronger than player 2 if and only if  $s > \frac{1}{2}$ . This is because  $l + d + w = 1$  and we then have

$$s = \frac{1}{2}d + w = \frac{1}{2}(d + w) + \frac{1}{2}w = \frac{1}{2} + \frac{1}{2}(w - l).$$

The relationship between expected score and Elo difference [2] is given by

$$\Delta E = S^{-1}(s)$$

where

$$S(x) = \frac{1}{1 + 10^{-x/400}}.$$

The inverse of  $S$  is given by

$$S^{-1}(s) = -\frac{400}{\log 10} \log \left( \frac{1}{s} - 1 \right).$$

## 2 Estimating Elo

We make the language a little more rigorous. Let  $X$  be a random variable taking the values 0,  $\frac{1}{2}$  and 1 with probabilities  $l$ ,  $d$  and  $w$ . Suppose that  $x_1, \dots, x_N$  are independent observations of  $X$ . The maximum likelihood estimator  $\hat{s}$  of  $s = E(X) = \frac{1}{2}d + w$  is given by

$$\hat{s} = \frac{1}{N} \sum_{j=1}^N x_j.$$

If  $\sigma^2$  is the variance of  $X$ , then by the Central Limit Theorem [1], we have

$$\frac{\hat{s} - s}{\sigma/\sqrt{N}} \rightarrow N(0, 1).$$

We want to construct a confidence interval for  $s$  of confidence grade  $1 - \alpha$ . Let  $\lambda_{\alpha/2}$  be the upper  $\alpha/2$  quantile of  $N(0, 1)$ . Then

$$\frac{\hat{s} - s}{\sigma/\sqrt{N}} \in (-\lambda_{\alpha/2}, \lambda_{\alpha/2})$$

with probability  $1 - \alpha$ . Equivalently

$$s \in \left( \hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}, \hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \right)$$

with the same probability. Since  $S$  is an increasing function we easily derive the confidence interval

$$\Delta E \in \left( S^{-1} \left( \hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \right), S^{-1} \left( \hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \right) \right)$$

for  $\Delta E = S^{-1}(s)$ . Note however that unless  $\hat{s} = 0$ ,  $\Delta \hat{E}$  will not lie in the middle of this interval.

### 2.1 Centered Approximate Confidence Intervals

We will take two approaches. The first approach assumes that the Elo difference is small and that  $N$  is large. The second approach only assumes that  $N$  is large.

We begin with the first approach. We have  $s = S(\Delta E)$ . We expand  $S$  in its first order Taylor polynomial around 0. We get

$$s = S(\Delta E) = S(0) + \Delta E S'(0) = \frac{1}{2} + \Delta E \frac{\log 10}{1600}.$$

Solving for  $\Delta E$  gives

$$\Delta E = \frac{1600}{\log 10} \left( s - \frac{1}{2} \right).$$

Starting with the confidence interval

$$s \in \left( \hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}, \hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \right)$$

from Section 2, we subtract  $\frac{1}{2}$  and multiply by the constant  $\beta = \frac{1600}{\log 10}$ . This gives

$$\beta \left( s - \frac{1}{2} \right) \in \left( \beta \left( \hat{s} - \frac{1}{2} \right) - \frac{\beta \lambda_{\alpha/2}\sigma}{\sqrt{N}}, \beta \left( \hat{s} - \frac{1}{2} \right) + \frac{\beta \lambda_{\alpha/2}\sigma}{\sqrt{N}} \right).$$

The assumption that  $\Delta E$  is small now gives

$$\Delta E \in \left( \Delta \hat{E} - \frac{\beta \lambda_{\alpha/2}\sigma}{\sqrt{N}}, \Delta \hat{E} + \frac{\beta \lambda_{\alpha/2}\sigma}{\sqrt{N}} \right).$$

For the second approach we start from

$$\Delta E \in \left( S^{-1} \left( \hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \right), S^{-1} \left( \hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \right) \right).$$

We now instead expand  $S^{-1}$  in its first order Taylor polynomial around  $s = \hat{s}$ . We simply get

$$\begin{aligned} S^{-1} \left( \hat{s} \pm \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \right) &= S^{-1}(\hat{s}) + \left( \hat{s} \pm \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} - \hat{s} \right) S^{-1'}(\hat{s}) = \\ &\Delta \hat{E} \pm \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \frac{1}{S'(\Delta \hat{E})}. \end{aligned}$$

This implies that

$$\Delta E \in \left( \Delta \hat{E} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \frac{1}{S'(\Delta \hat{E})}, \Delta \hat{E} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \frac{1}{S'(\Delta \hat{E})} \right).$$

## References

- [1] Wikipedia contributors. Central limit theorem — Wikipedia, the free encyclopedia, 2023. [https://en.wikipedia.org/w/index.php?title=Central\\_limit\\_theorem&oldid=1188038646](https://en.wikipedia.org/w/index.php?title=Central_limit_theorem&oldid=1188038646).
- [2] Wikipedia contributors. Elo rating system — Wikipedia, the free encyclopedia, 2023. [https://en.wikipedia.org/w/index.php?title=Elo\\_rating\\_system&oldid=1185663969](https://en.wikipedia.org/w/index.php?title=Elo_rating_system&oldid=1185663969).