A Note on Estimating Elo

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1 Introduction

Suppose that two players are playing a game of chess. We want to make some inference on the difference in strength between the two players. Let l, d and w be the probabilities that player 1 loses, draws or wins respectively. It is not hard to see that player 1 is stronger than player 2 if and only if w > l. In fact, d does not matter at all, if we only want to know which of the two players is stronger.

Consider however the two following cases. In the first case, l=d=0 and w=1. And in the second case, l=0, d=0.99 and w=0.01. We see that player 1 is stronger than player 2 in both of these cases, but much more so in the first.

To quantify this strength difference we naturally look at the expected score of a game. If a loss is worth 0 points, a draw $\frac{1}{2}$ and a win 1, the expected score is given by $s=\frac{1}{2}d+w$. The closer s is to $\frac{1}{2}$ the more even the games are. And if s is close to 1 then player 1 wins always every game. We see that in the first case $s=\frac{1}{2}\cdot 0+1\cdot 1=1$, and in the second case $s=\frac{1}{2}\cdot 0.99+1\cdot 0.01=0.505$.

We see that player 1 is stronger than player 2 if and only if $s > \frac{1}{2}$. This is because l + d + w = 1 and we then have

$$s = \frac{1}{2}d + w = \frac{1}{2}(d+w) + \frac{1}{2}w = \frac{1}{2} + \frac{1}{2}(w-l).$$

The relationship between expected score and Elo difference [2] is given by

$$\Delta E = S^{-1}(s)$$

where

$$S(x) = \frac{1}{1 + 10^{-x/400}}.$$

The inverse of S is given by

$$S^{-1}(s) = -\frac{400}{\log 10} \log \left(\frac{1}{s} - 1\right).$$

2 Estimating Elo

We make the language a little more rigorous. Let $X_1, X_2, ...$ be independent random variables taking the values $0, \frac{1}{2}$ and 1 with probabilities l, d and w. Let \hat{s} be the random variable given by

$$\hat{s} = \frac{1}{N} \sum_{n=1}^{N} X_n.$$

The central limit theorem [1] then asserts that

$$\frac{\hat{s} - s}{\sigma / \sqrt{N}}$$

where σ^2 is the variance of the variables X_n , approaches N(0,1) in distribution. We want to construct a confidence interval for s of confidence grade $1-\alpha$. Let $\lambda_{\alpha/2}$ be the upper $\alpha/2$ quantile of N(0,1). Then

$$P\left(-\lambda_{\alpha/2} < \frac{\hat{s} - s}{\sigma/\sqrt{N}} < \lambda_{\alpha/2}\right) \approx 1 - \alpha$$

or equivalently

$$P\left(\hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} < s < \hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right) \approx 1 - \alpha.$$

Since S is an increasing function we easily derive the confidence interval

$$P\left(S^{-1}\left(\hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right) < \Delta E < S^{-1}\left(\hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right)\right) \approx 1 - \alpha$$

for $\Delta E = S^{-1}(s)$. Note however that unless $\hat{s} = 0$, this interval will not be centered around $\Delta \hat{E} = S^{-1}(\hat{s})$.

2.1 Centered Approximate Confidence Intervals

We will take two approaches. The first approach assumes that the Elo difference is small and that N is large. The second approach assumes only that N is large.

We begin with the first approach. We have $s = S(\Delta E)$. Expanding S in its first order Taylor polynomial around 0 implies

$$s = S(\Delta E) \approx S(0) + \Delta E S'(0) = \frac{1}{2} + \Delta E \frac{\log 10}{1600}.$$

Solving for ΔE then gives

$$\Delta E \approx \frac{1600}{\log 10} \left(s - \frac{1}{2} \right).$$

Starting with the confidence interval

$$P\left(\hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} < s < \hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right) \approx 1 - \alpha$$

from Section 2, we subtract $\frac{1}{2}$ and multiply by the constant $\beta = \frac{1600}{\log 10}$. This gives

$$P\left(\beta\left(\hat{s}-\frac{1}{2}\right)-\frac{\beta\lambda_{\alpha/2}\sigma}{\sqrt{N}}<\beta\left(s-\frac{1}{2}\right)<\beta\left(\hat{s}-\frac{1}{2}\right)+\frac{\beta\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right)\approx 1-\alpha.$$

The assumption that ΔE is small now gives

$$P\left(\Delta \hat{E} - \frac{\beta \lambda_{\alpha/2} \sigma}{\sqrt{N}} < \Delta E < \Delta \hat{E} + \frac{\beta \lambda_{\alpha/2} \sigma}{\sqrt{N}}\right) \approx 1 - \alpha.$$

For the second approach we start from

$$P\left(S^{-1}\left(\hat{s} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right) < \Delta E < S^{-1}\left(\hat{s} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right)\right) \approx 1 - \alpha.$$

We now instead expand S^{-1} in its first order Taylor polynomial around $s = \hat{s}$. We simply get

$$S^{-1}\left(\hat{s} \pm \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}}\right)$$

$$\approx S^{-1}\left(\hat{s}\right) + \left(\hat{s} \pm \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} - \hat{s}\right)S^{-1\prime}\left(\hat{s}\right)$$

$$= \Delta \hat{E} \pm \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \frac{1}{S'\left(\Delta \hat{E}\right)}.$$

This implies

$$P\left(\Delta \hat{E} - \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \frac{1}{S'\left(\Delta \hat{E}\right)} < \Delta E < \Delta \hat{E} + \frac{\lambda_{\alpha/2}\sigma}{\sqrt{N}} \frac{1}{S'\left(\Delta \hat{E}\right)}\right) \approx 1 - \alpha.$$

References

- [1] Wikipedia contributors. Central limit theorem Wikipedia, the free encyclopedia, 2023. https://en.wikipedia.org/w/index.php?title=Central_limit_theorem&oldid=1188038646.
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