

1.

$$\text{Since } \frac{\partial \sum_i e_i^2}{\partial \beta_j} = 0, \text{ each } \beta_j = \frac{\sum_i (x_{ji} - \bar{x}_j)(y_i - \bar{y})}{\sum_i (x_{ji} - \bar{x}_j)^2} = \frac{\sum_i (x_{ji} - \bar{x}_j)y_i}{\sum_i (x_{ji} - \bar{x}_j)^2} =: \frac{S_{x_j y}}{S_{x_j x_j}}, \text{ where } \bar{x}_j := \frac{\sum_{i=1}^n x_{ji}}{n}$$

$$\text{Then, } \text{Var}(\beta_j) = \frac{\sum_i (x_{ji} - \bar{x}_j)^2 \text{Var}(\varepsilon_i)}{(\sum_i (x_{ji} - \bar{x}_j)^2)^2} = \frac{\sigma^2}{\sum_i (x_{ji} - \bar{x}_j)^2}; E(\beta_j) = 0, \text{ since}$$

$$\text{Var}(y_i) = \text{Var}(\hat{y}_i) + \text{Var}(\varepsilon_i) = \text{Var}(\varepsilon_i) = \sigma^2 \text{ and } \hat{y}_i \text{ is a computed quantity.}$$

for each independent random variable y_i .

For unbiased estimator we need to evaluate the following ,

$$\begin{aligned} E\left(\sum_{i=1}^n e_i^2\right) &= \sum_{i=1}^n E\left(y_i - \frac{\sum_{l=1}^n y_l}{n} - \sum_{j=1}^K \beta_j (x_{ji} - \bar{x}_j)\right)^2 \\ &= \sum_{i=1}^n \left(\frac{1}{n}\right)^2 (E(y_1^2) + \dots + E(y_{i-1}^2) + E(y_{i+1}^2) + \dots + E(y_n^2)) + \left(1 - \frac{1}{n}\right)^2 E(y_i^2) - 2\left(\frac{n-1}{n}\right)^2 + \frac{(n-1)(n-2)}{n^2} E(y_i)^2 \\ &\quad + \sum_{j=1}^K E(\beta_j^2 \sum_{i=1}^n (x_{ji} - \bar{x}_j)^2 - 2\beta_j \sum_{i=1}^n (x_{ji} - \bar{x}_j)(y_i - \bar{y})) \\ &= n \frac{(n-1)}{n} (\sigma^2 + E(y_i)^2 - E(y_i)^2) + E\left(\sum_{j=1}^K \beta_j^2 S_{x_j x_j} - 2\beta_j S_{x_j y}\right) \\ &= (n-1)\sigma^2 - K\sigma^2 = (n-1-K)\sigma^2 \\ \text{Therefore, } \frac{E\left(\sum_{i=1}^n e_i^2\right)}{(n-1-K)} &= \text{Var}(\varepsilon_i) = \sigma^2 \end{aligned}$$

2.

For this problem the true optimal solution can be computed analytically: $x^* = 2.611$ years, giving an expected cost of \$11,586. This solution is obtained by minimizing the expected cost, which can be written as

$$2000x + \int_0^\infty 20000 I(y \leq 1) \frac{e^{-y/x}}{x} dx$$

where I is the indicator function.

3.

Let X represent a process $S-T$ such that S has a distribution that complements the exponential arrival times such that X is normally distributed with zero mean and standard deviation σ . Then:

$$\begin{aligned} W_{n+1} - w &= \phi(W_n - w) + X_n \\ W_{n+1} - w &= \phi^{n+1}(W_0 - w) + \sum_{j=0}^n \phi^{n-j} X_j \\ E(W_{n+1}) &= w + \phi^{n+1}(E(W_0) - w) \rightarrow w \text{ as } n \rightarrow \infty \\ \text{Var}(W_{n+1}) &= \phi^{2(n+1)} + \sigma^2 \sum_{j=0}^n \phi^{2j} \rightarrow \sigma^2 / (1 - \phi^2) \text{ as } n \rightarrow \infty \end{aligned}$$

The first two lines in the above can be obtained by induction.

4.

In the calculation below, substitute epsilon = 3 and t = 2.821 for 9 degrees of freedom and at 1% confidence level obtained as 0.04/(5-1). Also use S12, S23, S24, S25 in the calculations below.

\bar{Y}_i	1	2	3	4	5
	8.9464	6.5448	7.1387	9.5497	11.1213
S_{ij}^2	1	2	3	4	5
1		0.0251	0.0375	0.0330	0.0620
2			0.0028	0.0104	0.0226
3				0.0051	0.0104
4					0.0046

$$t_{0.04/5-1, (10-1)} = 2.821$$

$$\bar{Y}_1 = 8.9464 > \bar{Y}_2 + t \sqrt{S_{12}^2/10} = 6.6861$$

$$\bar{Y}_3 = 7.1387 > \bar{Y}_2 + t \sqrt{S_{13}^2/10} = 6.71755$$

$$\bar{Y}_4 = 9.5497 > \bar{Y}_2 + t \sqrt{S_{14}^2/10} = 6.70685$$

$$\bar{Y}_5 = 11.1213 > \bar{Y}_2 + t \sqrt{S_{15}^2/10} = 6.6724$$

Similar Calculations

Thus, there was adequate data to select the best, policy 2, with 96% confidence.

$$\hat{S}^2 = \max_{i \neq j} S_{ij}^2 = 0.0620. \text{ The second-stage sample size,}$$

$$R = \max \left(R_0, \left\lceil \frac{t^2 \hat{S}^2}{\epsilon^2} \right\rceil \right) = \max \left(10, \left\lceil \frac{(2.821)^2 (0.0620)}{3^2} \right\rceil \right) = 10$$

Thus, 10 replication is sufficient to make statistical comparisons.

Since $\min_{i=1}^5 \{\bar{Y}_i\} = \bar{Y}_2$, there was adequate data to conclude that policy 2 has the least expected cost per day with 96% confidence.

However, the conclusion remains as above. The answer does not change.

5. (a)

From the least squares, $\frac{\partial}{\partial \beta_1} \sum_{i=1}^n e_i^2 = 0$, that is, $\sum_{i=1}^n e_i \frac{\partial}{\partial \beta_1} e_i = 0$, that is, $\sum_{i=1}^n e_i (x_i - \bar{x}) = 0$.

$$\begin{aligned} \sum_{i=1}^n (y_i - \bar{y})^2 &= \sum_{i=1}^n (y_i - \hat{y}_i + \hat{y}_i - \bar{y})^2 = \sum_{i=1}^n e_i^2 + r_i^2 + 2\beta_1 e_i (x_i - \bar{x}) \\ &= \sum_{i=1}^n (e_i^2 + r_i^2) + 2\beta_1 \sum_{i=1}^n e_i (x_i - \bar{x}) = \sum_{i=1}^n e_i^2 + \sum_{i=1}^n r_i^2 \end{aligned}$$

5(b) $\left(n - \left(1 + 2 \sum_{k=1}^{n-1} \frac{n-k}{n} \rho_k \right) \right) \sigma^2$

6.

(a) Test for significance ($H_0 : \mu_d = 0$)

Letting $d_i = y_i - z_i$, $\bar{d} = 1.80$, $S_d = 3.60$

$$t_0 = 1.80 / (3.60 / \sqrt{4}) = 1.0$$

For $\alpha = 0.05$, $t_{3,0.025} = 3.18$

Since $|t_0| < 3.18$, do not reject the null hypothesis.

(b) Sample size needed for $\beta \leq 0.20$

$$\delta = 2 / 3.60 = 0.556$$

For $\alpha = 0.05$, $\beta \leq 0.20$ and $\delta = 0.556$

$n = 30$ observations.