# MS-E2112 Multivariate Statistical Analysis (5cr) Lecture 7: Multiple Correspondence

Lecturer: Pauliina Ilmonen Slides: Ilmonen/Kantala

**Analysis** 

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# Multiple Correspondence Analysis

Multiple correspondence analysis (MCA) is an extension of bivariate correspondence analysis to more than 2 variables.

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## Frequency Tables

We consider a sample of size n described by P qualitative variables  $Y_1, ..., Y_P$ . The variable  $Y_p$  has  $K_p$  modalities (categories), and  $\sum_{p=1}^P K_p$  is the total number of the categories. The number of individuals having the modality I of the variable  $Y_p$  is denoted by  $n_{pl}$ . We set a variable  $x_{ipl} = 1$  if individual I has modality I of  $Y_p$ , and we set  $X_{ipl} = 0$  otherwise. Now

$$\sum_{l=1}^{K_{p}}n_{pl}=n,$$

and

$$\sum_{p=1}^{P} \sum_{l=1}^{K_p} n_{pl} = nP.$$

The table of  $K_p$  dummy variables associated with variable  $Y_p$ .

	1	2		$K_p$	
1	X <sub>1p1</sub>	X <sub>1p2</sub>		$X_{1pK_p}$	1
2	X <sub>2p1</sub>	$X_{2p2}$	• • •	$X_{2pK_p}$	1
:	:	:	:	:	:
n	X <sub>np1</sub>	$X_{np2}$		$X_{npK_p}$	1
	$n_{p1}$	$n_{p2}$		$n_{pK_p}$	n

Table: Table of dummy variables

Now we introduce the  $n \times K$  table/matrix  $X = [X_1, \dots, X_P]$ , called the complete disjunctive table.

		$X_1$				$X_P$		Column Profiles
	X <sub>11</sub>		$X_{1K_1}$		$X_{P1}$		$X_{PK_P}$	$\sum_{p=1}^{P} \sum_{l=1}^{K_p} x_{ipl}$
1	X <sub>111</sub>		X <sub>11K<sub>1</sub></sub>		X <sub>1P1</sub>		$X_{1PK_{P}}$	Politiple Correspondence
:	:	÷	:	:	:	÷	:	Analysis Graphical Presenta
i	<i>X</i> <sub>i11</sub>		$X_{i1K_1}$		X <sub>iP1</sub>		$X_{iPK_P}$	<b>P</b> Example
:	:	÷	:	:	:	÷	:	Some Remarks References
n	<i>X</i> <sub>n11</sub>		$X_{n1K_1}$		X <sub>nP1</sub>		$X_{nPK_P}$	P
$\sum_{i=1}^{n} x_{ipl}$	n <sub>11</sub>		n <sub>1 K1</sub>		n <sub>P1</sub>		$n_{PK_P}$	nP

Table: Complete disjunctive table

A group of kids were asked to select party snacks. Each kid chose one cookie, one milk shake and one salty snack. Here we have a sample of 4 individuals and 3 variables — n = 4, P = 3.

- Variable X<sub>1</sub> cookie has two options (modalities/categories)
   chocolate chip cookie (1) and oat cookie (2).
- Variable X<sub>2</sub> milk shake has three options vanilla (1), strawberry (2), and chocolate (3).
- Variable X<sub>3</sub> salty snack has two options pop corn (1), and potato chips (2).

Now 
$$K = K_1 + K_2 + K_3 = 2 + 3 + 2 = 7$$
.

We display the party snack data as a complete disjunctive table.

	<i>X</i> <sub>11</sub>	<i>X</i> <sub>12</sub>	X <sub>21</sub>	$X_{22}$	<i>X</i> <sub>23</sub>	<i>X</i> <sub>31</sub>	<i>X</i> <sub>32</sub>	$\sum_{p=1}^{7} \sum_{l=1}^{K_p}$	X <sub>ipl</sub>
1	0	1	1	0	0	1	0	3	Attraction R
2	0	1	1	0	0	0	1	3	
3	1	0	0	0	1	1	0	3	
4	0	1	0	1	0	0	1	3	
$\sum_{i=1}^{n} x_{ipl}$	1	3	2	1	1	2	2	12	Example

Table: Complete disjunctive table

- The first kid chose an oat cookie, vanilla milk shake and pop corn.
- The third kid chose a chocolate chip cookie, chocolate milk shake and pop corn.

# Multiple Correspondence Analysis

Bivariate correspondence analysis is now applied to the

complete disjunctive table!

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From the complete disjunctive table, one can compute the associated relative frequency table (F), where the elements of the complete disjunctive table are divided by the total sum nP leading to

$$f_{ipl} = \frac{x_{ipl}}{nP} \ (i=1,\ldots,n; p=1,\ldots,P; l=1,\ldots,K_p).$$

We have *P* variables and *n* individuals and  $f_{ipl} = \frac{1}{nP}$  or  $f_{ipl} = 0$ . Thus the marginal relative frequencies are computed as

$$f_{i..} = \sum_{p=1}^{P} \sum_{l=1}^{K_p} \frac{x_{ipl}}{nP} = P \frac{1}{nP} = \frac{1}{n}$$

and

$$f_{.pl} = \sum_{i=1}^{n} \frac{x_{ipl}}{nP} = \frac{n_{pl}}{nP}.$$

We display the party snacks data as a relative frequency table.

	<i>X</i> <sub>11</sub>	$X_{12}$	<i>X</i> <sub>21</sub>	$X_{22}$	$X_{23}$	<i>X</i> <sub>31</sub>	<i>X</i> <sub>32</sub>	$f_{i}$
1	0	1/2	1/2	0	0	1/12	0	1/4
2	0	12	12	0	0	Ö	<u>1</u> 12	$\frac{1}{4}$
3	1/12	Ö	Ö	0	<u>1</u> 12	1/12	Ö	1/4
4	Ö	$\frac{1}{12}$	0	1/2	Ö	Ö	1/12	$\frac{1}{4}$
$f_{.pl}$	1 12	3 12	<u>2</u> 12	12	1 12	<u>2</u> 12	12	1

Table: Relative frequency table

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**Row Profiles** 

The idea behind MCA, like in bivariate correspondence analysis, is to apply a PCA type approach on one hand to the row profiles, and on the other hand to the column profiles of the relative frequencies table F. The coordinate pl of the row profile  $l_i(1 \times K)$  associated with individual i is given as

$$(I_i)_{pl} = \frac{f_{ipl}}{f_{i..}} = \frac{\frac{x_{ipl}}{nP}}{\frac{1}{n}} = \frac{x_{ipl}}{nP} \frac{n}{1} = \frac{x_{ipl}}{P}, \qquad i = 1, ..., n.$$

As

$$\sum_{i=1}^{n} \frac{1}{n} (I_i)_{pl} = \sum_{i=1}^{n} \frac{1}{n} \frac{x_{ipl}}{P} = \frac{n_{pl}}{nP},$$

the n row profiles weighted by the marginal relative frequencies (1/n) compose a point cloud in  $\mathbb{R}^K$  with a center given by the relative marginal profile

$$G_I = (\frac{n_{11}}{nP}, \dots, \frac{n_{1K_1}}{nP}, \dots, \frac{n_{P1}}{nP}, \dots, \frac{n_{PK_P}}{nP}).$$

The row profiles of the party snacks data is given below.

	X <sub>11</sub>	<i>X</i> <sub>12</sub>	X <sub>21</sub>	$X_{22}$	$X_{23}$	<i>X</i> <sub>31</sub>	<i>X</i> <sub>32</sub>	
1	0	<u>1</u>	<u>1</u> 3	0	0	1/3	0	1
2	0	<u>1</u>	<u>1</u>	0	0	Ŏ	<u>1</u>	1
3	1/3	Ŏ	Ŏ	0	1/3	1/3	Ŏ	1
4	ŏ	<u>1</u>	0	<u>1</u>	ŏ	ŏ	<u>1</u>	1

Table: Row profiles

## **Row Profiles**

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Intuitively, the distance between two individuals is small if they have many modalities in common, and the distance between the individual i and the center increases as the modalities taking by the individual i becomes rare ( $x_{ipl} = 1$  for  $n_{pl}$  small).

$$d^{2}(l_{i_{1}}, l_{i_{2}}) = \sum_{p=1}^{P} \sum_{l=1}^{K_{P}} \frac{1}{f_{.pl}} ((l_{i_{1}})_{pl} - (l_{i_{2}})_{pl})^{2}$$
$$= \frac{n}{P} \sum_{p=1}^{P} \sum_{l=1}^{K_{P}} \frac{1}{n_{pl}} (x_{i_{1}pl} - x_{i_{2}pl})^{2}.$$

The distance between the first kid and the second kid in the party snacks data is

$$(\frac{4}{3}(1(0-0)^2 + \frac{1}{3}(1-1)^2 + \frac{1}{2}(1-1)^2 + 1(0-0)^2 + 1(0-0)^2 + \frac{1}{2}(1-0)^2 + \frac{1}{2}(0-1)^2))$$

$$= \frac{4}{3} \approx 1.33.$$

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### Column Profiles

The coordinate i of the column profile  $c_{pl}$   $(n \times 1)$  associated with the modality l of  $Y_p$  is given as

$$(c_{pl})_i = \frac{f_{ipl}}{f_{.pl}} = \frac{\frac{x_{ipl}}{nP}}{\frac{n_{pl}}{nP}} = \frac{x_{ipl}}{nP} \frac{nP}{n_{pl}} = \frac{x_{ipl}}{n_{pl}}, \qquad p = 1, \dots, P; l = 1, \dots, K_p.$$

As

$$\sum_{p=1}^{P} \sum_{l=1}^{K_p} f_{.pl}(c_{pl})_i = \sum_{p=1}^{P} \sum_{l=1}^{K_p} \frac{n_{pl}}{nP} \frac{x_{ipl}}{n_{pl}} = \sum_{p=1}^{P} \sum_{l=1}^{K_p} \frac{x_{ipl}}{nP} = \frac{P}{nP} = \frac{1}{n},$$

the K column profiles weighted by the marginal relative frequencies  $(\frac{n_{pl}}{nP})$  compose a point cloud in  $\mathbb{R}^K$  with the center given by the relative marginal profile  $G_c = (\frac{1}{n}, \dots, \frac{1}{n})$ .

The column profiles of the party snacks data are given below.

	<i>X</i> <sub>11</sub>	$X_{12}$	<i>X</i> <sub>21</sub>	$X_{22}$	$X_{23}$	<i>X</i> <sub>31</sub>	X <sub>32</sub>
1	0	1/3	1/2	0	0	1/2	0
2	0	<u>1</u>	1 2	0	0	Ō	$\frac{1}{2}$
3	1	Ŏ	Ō	0	1	1/2	ō
4	0	<u>1</u>	0	1	0	Ō	$\frac{1}{2}$
	1	1	1	1	1	1	1

Table: Column profiles

## Column Profiles

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Intuitively, the  $\chi^2$  distance between two modalities is small if the same individuals take these two modalities together, and the distance between the modality I of  $Y_p$  and the center increases as the modality becomes more rare ( $n_{pl}$  small).

$$d^{2}(c_{p_{1}l_{1}}, c_{p_{2}l_{2}}) = \sum_{i=1}^{n} \frac{1}{f_{i..}} ((c_{p_{1}l_{1}})_{i} - (c_{p_{2}l_{2}})_{i})^{2}$$

$$= n \sum_{i=1}^{n} (\frac{x_{ip_{1}l_{1}}}{n_{p_{1}l_{1}}} - \frac{x_{ip_{2}l_{2}}}{n_{p_{2}l_{2}}})^{2}.$$

$$4((0-0)^2+(0-0)^2+(1-0)^2+(0-1)^2)=8$$

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Let  $n_{p_1l_1,p_2l_2}$  be the number of individuals having the modality  $l_1$  of the variable  $Y_{p_1}$  and the modality  $l_2$  of the variable  $Y_{p_2}$ . Now the attraction repulsion index  $d_{p_1l_1,p_2l_2}$  between the modality  $l_1$  of the variable  $Y_{p_1}$  and the modality  $l_2$  of the variable  $Y_{p_2}$  is given by

$$d_{p_1l_1,p_2l_2} = \frac{n_{p_1l_1,p_2l_2}/n}{n_{p_1l_1}/nn_{p_2l_2}/n} = \frac{n_{p_1l_1,p_2l_2}}{\frac{n_{p_1l_1}n_{p_2l_2}}{n}}.$$

# **Attraction Repulsion Indices**

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It is clear that if the attraction repulsion index is larger than one, the individuals are more inclined to take both modalities simultaneously than under the hypothesis of independence. And vice-versa, if the attraction repulsion index is smaller than one, the individuals are less inclined to take both modalities simultaneously than under the hypothesis of independence.

The attraction repulsion index  $d_{i,pl}$  between the individual i and the modality l of the variable  $Y_p$  is defined as follows.

$$d_{i,pl} = \frac{f_{ipl}}{f_{i..}f_{.pl}} = \frac{x_{ipl}}{n_{pl}/n}.$$

Now, clearly

$$d_{i,pl}=0,$$

if  $x_{inl} = 0$  and

$$d_{i,pl}=rac{n}{n_{pl}},$$

if  $x_{ipl} = 1$ . Thus, if the individual i does not have the modality l of the variable  $Y_p$ , then the attraction repulsion index  $d_{i,pl}$  is equal to 0, and if the individual i does have the modality l of  $Y_p$ , then the attraction repulsion index  $d_{i,pl}$  increases as the l of  $Y_p$  becomes rare.

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To maximize chi-square distances and to obtain a representation in lower dimension, PCA type transformation is applied on the two data clouds: the row profiles and the column profiles. A transformation of the profiles is necessary to center the variables, and to be able to base the maximization problem on euclidian distances instead of  $\chi^2$  distances directly:

$$(I_i^{\circ})_{pl} = \frac{(I_i)_{pl}}{\sqrt{f_{.pl}}} - \sqrt{f_{.pl}} \text{ and } (c_{pl}^{\circ})_i = \frac{(c_{pl})_i}{\sqrt{f_{i...}}} - \sqrt{f_{i...}}$$

$$V = T^T T$$
 and  $W = TT^T$  where the elements of  $T$  are given by  $\frac{x_{ipl} - n_{pl}/n}{\sqrt{Pn_{pl}}}$ .

Note that here also, the matrix V is a relative row frequency weighted covariance matrix of the scaled and shifted row profiles and the matrix W is a relative column frequency weighted covariance matrix of the scaled and shifted column profiles.

The MCA components for the individuals are derived from the eigenvectors of the matrix V, and the MCA components for the modalities from the eigenvectors of the matrix W.

Let H = rank(V) = rank(W). The scores of the individuals are given as

$$\phi_{h,i} = \sum_{k=1}^K u_{h,k}(I_i^\circ)_k \quad h = 1,\ldots,H,$$

where  $u_{h,k}$  is the kth element of the eigenvector associated with the hth largest eigenvalues of V.

The scores for the modalities are given as

$$\psi_{h,pl} = \sum_{i=1}^n v_{h,i}(c_{pl}^\circ)_i \quad h = 1,\ldots,H.$$

$$C(pl,h) = \frac{f_{,pl}\psi_{h,pl}^2}{\lambda_h} = \frac{n_{pl}\psi_{h,pl}^2}{nP\lambda_h}.$$

Global contribution of the variable  $Y_p$  is given by

$$C(p,h) = \sum_{l=1}^{K_p} C(pl,h).$$

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# **Graphical Presentation**

The attraction repulsion index

$$d_{p_1l_1,p_2l_2} = 1 + \sum_{h=1}^{H} \psi_{h,p_1l_1} \psi_{h,p_2l_2}.$$

The graphical output of MCA is the approximation of the previous formula using few dimensions. Suppose that the modalities are well represented in two dimensions. Then we can plot the two first MCA components and interpret the proximity between the points on the first principal plan with the following approximation

$$d_{p_1l_1,p_2l_2} \approx 1 + \sum_{h=1}^2 \psi_{h,p_1l_1} \psi_{h,p_2l_2}.$$

$$d_{i_1,i_2} = 1 + \sum_{h=1}^{H} \phi_{h,i_1} \phi_{h,i_2}.$$

Two individuals are close if they have in general the same modalities.

Now  $d_{i_1,i_2}$  can be approximated by

$$d_{i_1,i_2} \approx 1 + \sum_{h=1}^2 \phi_{h,i_1} \phi_{h,i_2}.$$

The attraction repulsion index

$$d_{i,pl} = 1 + \sum_{h=1}^{H} \frac{1}{\sqrt{\lambda_h}} \phi_{h,i} \psi_{h,pl},$$

and thus again

$$d_{i,pl} \approx 1 + \sum_{h=1}^{2} \frac{1}{\sqrt{\lambda_h}} \phi_{h,i} \psi_{h,pl}.$$

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$$\hat{\phi}_{1,j} = \frac{1}{\sqrt{\lambda_1}} \phi_{1,j}$$

and

$$\hat{\phi}_{2,j} = \frac{1}{\sqrt{\lambda_2}} \phi_{2,j}.$$

Then

$$d_{i,pl} \approx 1 + \sum_{h=1}^{2} \hat{\phi}_{h,i} \psi_{h,pl},$$

and the final graphical representation can be given simultaneously as a double biplot.

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Disclaimer: This example data set is randomly generated. Please do not draw real life conclusions from it.

	<i>X</i> <sub>11</sub>	<i>X</i> <sub>12</sub>	X <sub>21</sub>	$X_{22}$	$X_{23}$	X <sub>31</sub>	<i>X</i> <sub>32</sub>	$\int_{p=1}^{7} \sum_{l=1}^{K_p} K_p$	1 X <sub>ipl</sub>
1	0	1	1	0	0	1	0	3	Analysis
2	0	1	1	0	0	0	1	3	
:	:	÷	:	:	:	:	:	:	
25	1	0	0	0	1	0	1	3	
$\sum_{i=1}^{n} x_{ipl}$	16	9	9	6	10	14	11		

Table: Complete disjunctive table

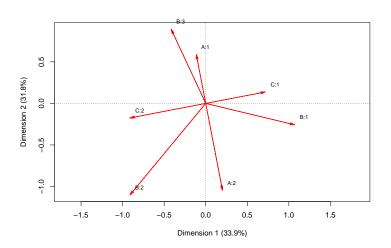


Figure: Result of MCA (A1=chocolate chip cookie, A2=oat cookie, B1=vanilla milk shake, B2=strawberry milk shake, B3=chocolate milk shake, C1=pop corn, C2=potato chips.) It seems that kids that like chocolate chip cookies like chocolate milk shake as well.

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When performing MCA, it is better to take into account variables that have more or less the same number of modalities. (The number of modalities has an effect on the analysis.) It is also advised to avoid having very rare modalities. (Rare modalities have a big impact on analysis, and that makes MCA quite nonrobust method.) One can preprocess the data by grouping modalities if necessary.

# **Next Week**

Next week we will talk about canonical correlation analysis.

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