Exercise 4

t-test

C.

d.

Homework exercise

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To be solved at home before the exercise session.
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a. Let x_1,\ldots,x_n be a random sample (iid) from some distribution F_{	heta} with the unknown parameter 	heta. Which of the three one-sample
   tests (t-test, sign test or signed rank test) would you use (and why!) to test whether the location (expected value/median) of the data
   is equal to 1 if we know for certain that the distribution F_{	heta} is
        i. an exponential distribution with unknown rate parameter \theta,
       ii. a normal distribution with variance 2 and unknown expected value 	heta,
      iii. a Laplace distirbution with known scale parameter 5 and unknown location parameter \theta,
```

iv. a Poisson distirbution with unknown parameter θ ?

In each case we want to use a test which makes the strictest assumptions (such that they are still satisfied). This gives us maximal power (lowest Type 2 error rate), as we "use more information" about the data. See the lecture examples of week # The assumptions the three tests make besides iid data are: # t-test: normality <- strictest # signed rank test: symmetric continuous distribution <- less strict

<- even less strict

sign test: continuous distribution # a. # Exponential distributions are not symmetric -> sign test. # b.

Laplace distributions are not normal but they are symmetric -> signed rank test.

regularly applied to discrete data as well (e.g. using the conventions of slide 3.7). b. The data set `airmiles` lists the passenger miles flown by commercial airlines in the United States for each year from 19 37 to 1960. To inspect whether the yearly passenger miles equal 10000 on average, a researcher performed a sign test to test the null hypothesis $med_x = 10000$ on significance level 5% with the results shown below and concluded that there is no evi dence against the null hypothesis. Do you agree with the researcher's conclusion?

Poisson distribution is neither continuous nor symmetric so, being strict, none of the tests apply. However, sign test is

airmiles

Time Series: ## Start = 1937## End = 1960## Frequency = 1 ## [1] 412 480 683 1052 1385 1418 1634 2178 3362 5948 6109 ## [12] 5981 6753 8003 10566 12528 14760 16769 19819 22362 25340 25343 ## [23] 29269 30514 # Sign test

binom.test(sum(airmiles > 10000), length(airmiles))

Exact binomial test ## data: sum(airmiles > 10000) and length(airmiles) ## number of successes = 10, number of trials = 24, p-value = 0.5413 ## alternative hypothesis: true probability of success is not equal to 0.5 ## 95 percent confidence interval: ## 0.2210969 0.6335694 ## sample estimates: ## probability of success 0.4166667 # The sign test assumes that the observations x1, x2, ..., xn form an iid. sample from some particular continuous distributi on. While it could be plausible to view the yearly passenger miles as realizations of identically distributed random variabl

hat they continue going up in the coming years as well (the technology develops etc.). This can be seen in the time series p lot of the data: plot(airmiles, type = "b")

es from a continuous distribution, they are certainly not independent. If the passenger miles go up one year, it is likely t

25000 airmiles 5000 0 1940 1945 1950 1955 1960

Class exercise

Time

In this kind of situation methods of time series analysis are needed (not covered in this course).

Note: all the needed data sets are either given below or available in base R.

က

7

##

ign test seems like the safest choice.

95 percent confidence interval:

the salaries differ).

a.

48000

46000

44000

Exact binomial test

d.

e.

1.2

in the line.

male <- (1:20)[line == "M"]

a.

Ls").

d. Conduct the two-sample rank test and draw conclusions.

s difference in scales is small enough that the test can still be used.

Find out how the package and the piping operator %>% work by going through an online tutorial.

data: sum(diff > 0) and length(diff)

women

a. Begin again by visualizing the data.

binom.test(sum(sleep_1 > 0), length(sleep_1))

To be solved at the exercise session.

interested in studying whether drug 1 helps in increasing the number of hours slept compared to placebo. a. Extract the increases in hours of sleep of the patients who received drug 1 (group == 1). b. Visualize the data.

c. Conduct an appropriate test to evaluate whether the location (expected value/median) of the increase in hours of sleep differs from 0 on significance level 5%.

1. The data set sleep shows the effect of two soporific drugs (increase in hours of sleep compared to control) on 10 patients. We are

d. Draw conclusions. # a. sleep_1 <- sleep[sleep\$group == 1, 1]</pre>

b. boxplot(sleep_1)

0 # C.

Exact binomial test ## data: sum(sleep_1 > 0) and length(sleep_1) ## number of successes = 5, number of trials = 10, p-value = 1 ## alternative hypothesis: true probability of success is not equal to 0.5

With so few observations it is difficult to say whether the data comes from a normal, or even a symmetric, distribution. S

```
## 0.187086 0.812914
## sample estimates:
## probability of success
                        0.5
# p-value = 1
# The highest possible p-value -> no evidence against H0 -> the drug 1 is no better than placebo.
  2. The data set below contains the annual salaries (in dollars) of 8 American women and 8 American men (recall exercise 3.2). The
    observations are paired such that each woman is matched with a man having similar background (age, occupation, level of education, etc).
    We are interested in studying whether the locations of the salaries of women and men differ (recall that last time paired t-test concluded that
```

b. Which two non-parametric tests are appropriate in studying our question of interest? c. State the hypotheses of the tests and conduct them on the significance level 10%. d. What are the conclusions of the tests? e. What assumptions did the test in part c make? Are they justifiable? salary \leftarrow data.frame(women = c(42600, 43600, 49300, 42300, 46200, 45900, 47500, 41300), men = c(46200, 44700, 48400, 41700, 48600, 49300, 48300, 44300))

men = c(46200, 44700, 48400, 41700, 48600, 49300, 48300, 44300))# Alternative visualization to the last time plot(women ~ men, data = salary) abline(a = 0, b = 1)

0

42000 0 42000 44000 46000 48000 men # Most points are below the y=x -line, meaning that the salary of the man in a pair is more often larger than that of the wo man # b. # The data is paired (and the pairs are not independent), making paired sign test and paired signed rank test appropriate ch oices. # Note that using a two-sample rank test is not justified as it assumes the independence of the two samples. # C. # Both tests have the same hypotheses # H0: med_(women - men) == 0 # H1: med_(women - men) != 0 diff <- salary\$women - salary\$men</pre> # Paired sign test binom.test(sum(diff > 0), length(diff))

Paired sign test does not reject the null but the paired signed rank test does.

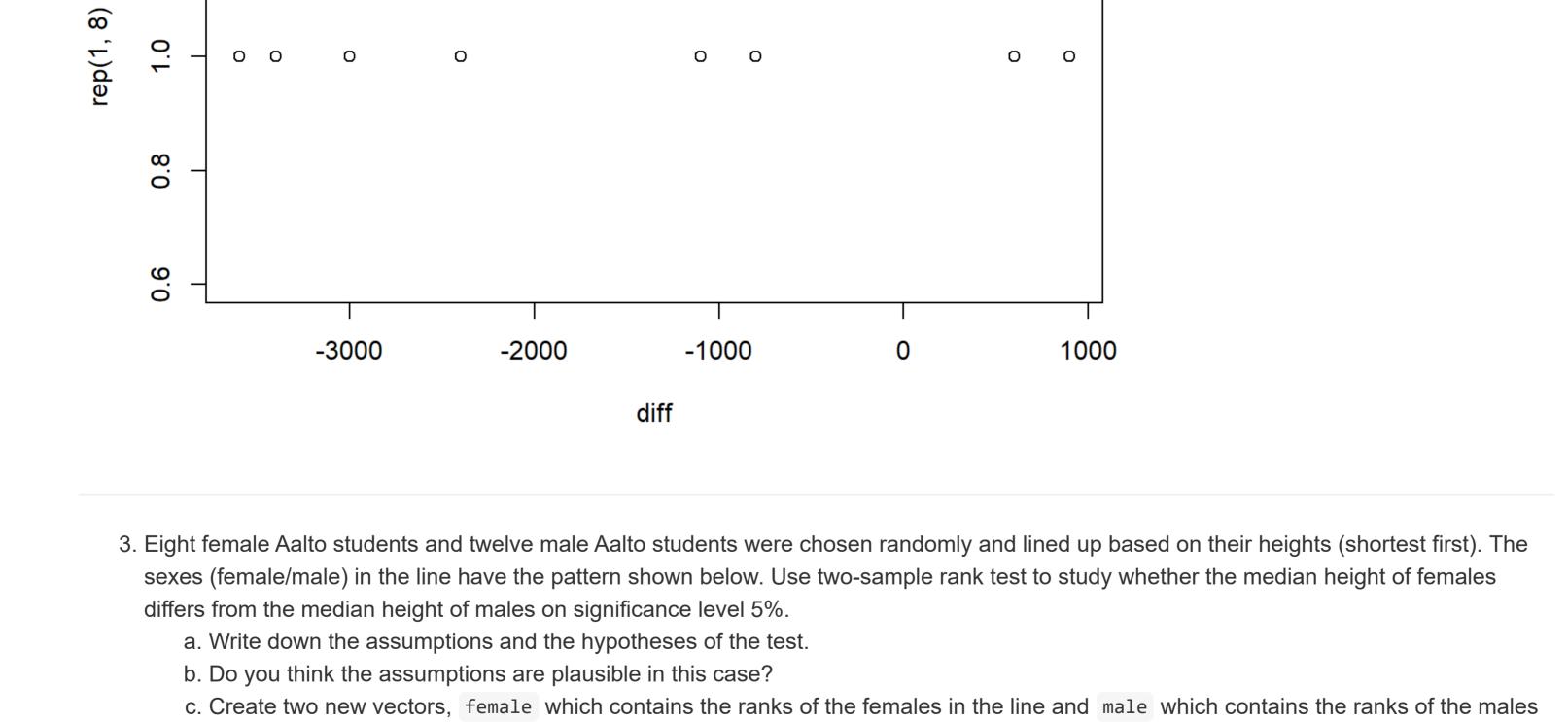
salary \leftarrow data.frame(women = c(42600, 43600, 49300, 42300, 46200, 45900, 47500, 41300),

number of successes = 2, number of trials = 8, p-value = 0.2891 ## alternative hypothesis: true probability of success is not equal to 0.5 ## 95 percent confidence interval: ## 0.03185403 0.65085579 ## sample estimates: ## probability of success 0.25 # Paired signed rank test wilcox.test(diff, mu = 0) ## Wilcoxon signed rank test ## data: diff ## V = 4, p-value = 0.05469 ## alternative hypothesis: true location is not equal to 0

g distribution is symmetric based on a so small sample so the paired sign test might be a better choice -> no difference in salaries. plot(diff, rep(1, 8))

Both tests assume that the differences d1, d2, ..., d8 form an iid. sample from some particular continuous distribution. P

aired signed rank test furthermore assumes that this distribution is symmetric. It is difficult to say whether the underlyin



ative hypothesis is the opposite of that). # b. # As the samples where chosen randomly, it is plausible that the samples are independent and iid. Also, googling "female vs. male height distribution" shows that the male distribution of heights is slightly wider than for females. We assume that thi

The null hypothesis is that the medians of Fx and Fy (and consequently the distributions itself) are equal (and the altern

The test assumes that the female and male samples are mutually independent iid samples from the continuous distributions F

x and Fy, respectively. Moreover, the distributions Fx and Fy are assumed to be equal up to location shift ("same-shaped hil

C. female <- (1:20)[line == "F"]

d. wilcox.test(female, male)

Wilcoxon rank sum test ## data: female and male ## W = 25, p-value = 0.0825 ## alternative hypothesis: true location shift is not equal to 0

at there should be a difference, then either the sample size was too small, the result was caused by randomness or the assum ptions weren't justified.

p-value = 0.0825 -> not enough evidence to reject H0 on significance level 5% -> no difference in medians. As we "know" th

achieved more transparently using the package dplyr as follows. # install.packages("dplyr") library(dplyr)

4. (Optional) Data manipulation using just functions in base R does not always produce the most readable code. The task in 1a. can be

sleep_1 <- sleep %>% filter(group == 1) %>% select(extra)