

Mid-term solutions
ECE 271A
Electrical and Computer Engineering
University of California San Diego

Nuno Vasconcelos

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1.a) The log-likelihood function is

$$\begin{aligned}\mathcal{L}(\mathcal{D})(\theta) &= \sum_{i=1}^N \left\{ (k-1) \log x_i - \frac{x_i}{\theta} - k \log(\theta) - \log((k-1)!) \right\} \\ &= (k-1) \sum_{i=1}^N \log x_i - \sum_{i=1}^N \frac{x_i}{\theta} - kN \log(\theta) - N \log((k-1)!)\end{aligned}$$

The derivative with respect to θ is zero when

$$\frac{\partial \mathcal{L}(\mathcal{D})}{\partial \theta}(\theta) = \frac{1}{\theta^2} \sum_{i=1}^N x_i - \frac{kN}{\theta} = 0$$

or

$$\theta^* = \frac{1}{kN} \sum_{i=1}^N x_i.$$

The second derivative is

$$\frac{\partial^2 \mathcal{L}(\mathcal{D})}{\partial \theta^2}(\theta) = -\frac{2}{\theta^3} \sum_{i=1}^N x_i + \frac{kN}{\theta^2} = \frac{1}{\theta^2} \left(-\frac{2}{\theta} \sum_{i=1}^N x_i + kN \right)$$

At the critical point this is

$$\frac{\partial^2 \mathcal{L}(\mathcal{D})}{\partial \theta^2}(\theta^*) = \frac{1}{(\theta^*)^2} (-2kN + kN) = -\frac{kN}{(\theta^*)^2} < 0$$

and θ^* is an ML estimate.

b) Using the properties of the Gamma distribution,

$$\sum_{i=1}^N X_i \sim \text{Gamma}(x; kN, \theta)$$

and

$$\theta^* \sim \text{Gamma}\left(x; kN, \frac{\theta}{kN}\right).$$

It follows that

$$E_{X_1, \dots, X_N}[\theta^*] = kN \frac{\theta}{kN} = \theta$$

and the estimator is unbiased. Also

$$\text{var}_{X_1, \dots, X_N}[\theta^*] = kN \left(\frac{\theta}{kN} \right)^2 = \frac{\theta^2}{kN}.$$

Finally, since there is no bias, the MSE is the same as the variance.

c) In this case, the log-likelihood is a function of both k and θ

$$\mathcal{L}(\mathcal{D})(k, \theta) = (k-1) \sum_{i=1}^N \log x_i - \sum_{i=1}^N \frac{x_i}{\theta} - kN \log(\theta) - N \log((k-1)!).$$

We have seen that, for any k , this is maximized when

$$\theta = \frac{1}{kN} \sum_{i=1}^N x_i.$$

Plugging this into the log-likelihood, we obtain

$$\begin{aligned} \mathcal{L}(\mathcal{D})(k) &= (k-1) \sum_{i=1}^N \log x_i - kN - kN \log \left(\sum_{i=1}^N x_i \right) + kN \log(kN) - N \log((k-1)!) \\ &= k \left(\sum_{i=1}^N \log x_i - N - N \log \left(\sum_{i=1}^N x_i \right) \right) + kN \log(kN) - N \log((k-1)!) \\ &= ak + N \log \frac{(kN)^k}{(k-1)!} \end{aligned}$$

with

$$a = \sum_{i=1}^N \log x_i - N - N \log \left(\sum_{i=1}^N x_i \right).$$

Since k is an integer, the following procedure will find the ML solution

1. using the sample \mathcal{D} compute a .
2. starting with $k = 1$, search for the k^* that maximizes

$$ak + N \log \frac{(kN)^k}{(k-1)!}$$

This can be done by simply trying all the values of k in $1 < k < k_{max}$, where k_{max} is some large integer, and picking the one that leads to the largest value of the function.

3. set

$$\theta^* = \frac{1}{k^*N} \sum_{i=1}^N x_i.$$

4. return k^* and θ^* .

2. a) The BDR is

$$i^* = \arg \min_i \|\mathbf{x} - \mu_i\|^2.$$

The decision boundary is given by

$$\begin{aligned} \|\mathbf{x} - \mu_1\|^2 &= \|\mathbf{x} - \mu_{-1}\|^2 \\ -2\mathbf{x}^T \mu_1 + \|\mu_1\|^2 &= -2\mathbf{x}^T \mu_{-1} + \|\mu_{-1}\|^2 \\ (\mu_1 - \mu_{-1})^T \mathbf{x} - \frac{1}{2}(\|\mu_1\|^2 - \|\mu_{-1}\|^2) &= 0 \end{aligned}$$

Hence, the decision boundary as the desired form with

$$\begin{aligned} \mathbf{w} &= \mu_1 - \mu_{-1} \\ b &= -\frac{1}{2}(\|\mu_1\|^2 - \|\mu_{-1}\|^2). \end{aligned}$$

b) The ML estimate for the mean is $\hat{\mu} = \frac{1}{n} \sum_i \mathbf{x}_i$. Hence, we have

$$\begin{aligned} \hat{\mathbf{w}} &= \left(\frac{1}{n_1} \sum_{i|y_i=1} \mathbf{x}_i - \frac{1}{n_{-1}} \sum_{i|y_i=-1} \mathbf{x}_i \right) \\ \hat{b} &= - \left(\frac{1}{2n_1^2} \sum_{i,j|y_i=y_j=1} \mathbf{x}_i^T \mathbf{x}_j - \frac{1}{2n_{-1}^2} \sum_{i,j|y_i=y_j=-1} \mathbf{x}_i^T \mathbf{x}_j \right). \end{aligned}$$

where

$$\begin{aligned} n_1 &= \sum_i \left(\frac{1+y_i}{2} \right) \\ n_{-1} &= \sum_i \left(\frac{1-y_i}{2} \right) \end{aligned}$$

are the numbers of positive and negative examples in Joe's sample.

c) The expressions above can be written as

$$\begin{aligned} \hat{\mathbf{w}} &= \sum_i \alpha_i \mathbf{x}_i \\ \hat{b} &= -\frac{1}{2} \sum_{i,j|y_i=y_j} y_i \alpha_i^2 \mathbf{x}_i^T \mathbf{x}_j. \end{aligned}$$

where

$$\alpha_i = \begin{cases} 1/n_1, & \text{if } y_i = 1 \\ -1/n_{-1}, & \text{if } y_i = -1 \end{cases} \quad (1)$$

d) This problem is less trivial than it may seem initially. The reason is that the number of points in each class is itself a random variable (N_1 positives and N_1 negatives) since we don't know what the values of n_1 and n_{-1} are. We need to account for this when we consider the expected value of $\hat{\mathbf{w}}$. As usual, we have

$$\mu_{\mathbf{w}} = E_{\mathbf{X}_1, \dots, \mathbf{X}_n, Y_1, \dots, Y_n} [\hat{\mathbf{w}}] = E_{\mathbf{X}_1, \dots, \mathbf{X}_n, Y_1, \dots, Y_n} \left[\sum_i \alpha_i \mathbf{x}_i \right]$$

Note, however, that α_i depends on the whole sample Y_1, \dots, Y_n and we cannot just get rid of the Y_i . We can use the fact, however, that α_i does not depend on the Y_i exactly, just on the number of positive examples N_1 . This leads to

$$\begin{aligned}
\mu_{\mathbf{w}} &= \sum_i E_{\mathbf{X}_i, Y_1, \dots, Y_n} [\alpha_i \mathbf{X}_i] = n E_{\mathbf{X}, Y, N_1} [\alpha \mathbf{X}] = n E_{N_1} [E_{\mathbf{X}, Y | N_1} [\alpha \mathbf{X} | N_1 = k]] \\
&= n E_{N_1} [P_{Y|N_1}(1|k) E_{\mathbf{X}|N_1, Y} [\alpha \mathbf{X} | N_1 = k, Y = 1] + P_{Y|N_1}(-1|k) E_{\mathbf{X}|N_1, Y} [\alpha \mathbf{X} | N_1 = k, Y = -1]] \\
&= n E_{N_1} \left[P_{Y|N_1}(1|k) \frac{1}{k} \mu_1 - P_{Y|N_1}(-1|k) \frac{1}{n-k} \mu_{-1} \right] \\
&= E_{N_1} \left[P_{Y|N_1}(1|k) \frac{n}{k} \right] \mu_1 - E_{N_1} \left[P_{Y|N_1}(-1|k) \frac{n}{n-k} \right] \mu_{-1} \\
&= E_{N_1} \left[\frac{P_{N_1|Y}(k|1) P_Y(1) n}{P_{N_1}(k)} \frac{1}{k} \right] \mu_1 - E_{N_1} \left[\frac{P_{N_1|Y}(k|-1) P_Y(-1) n}{P_{N_1}(k)} \frac{1}{n-k} \right] \mu_{-1} \\
&= E_{N_1} \left[\frac{\binom{n-1}{k-1} P_Y(1)^{k-1} P_Y(-1)^{n-k} P_Y(1) n}{\binom{n}{k} P_Y(1)^k P_Y(-1)^{n-k}} \frac{1}{k} \right] \mu_1 - E_{N_1} \left[\frac{\binom{n-1}{k} P_Y(1)^k P_Y(-1)^{n-1-k} P_Y(-1) n}{\binom{n}{k} P_Y(1)^k P_Y(-1)^{n-k}} \frac{1}{n-k} \right] \mu_{-1} \\
&= E_{N_1} \left[\frac{k}{n} \frac{n}{k} \right] \mu_1 - E_{N_1} \left[\frac{n-k}{n} \frac{n}{n-k} \right] \mu_{-1} \\
&= \mu_1 - \mu_{-1} \\
&= \mathbf{w}
\end{aligned}$$