Lecture 11

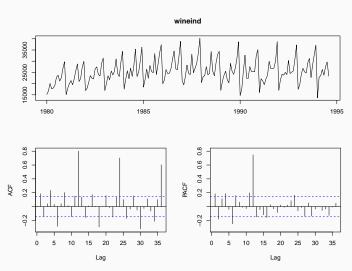
Seasonal ARIMA

Colin Rundel 02/22/2017

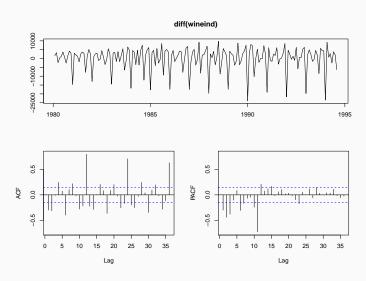
Seasonal Models

Australian Wine Sales Example (Lecture 6)

Australian total wine sales by wine makers in bottles <= 1 litre. Jan 1980 – Aug 1994.



Differencing



Seasonal ARIMA

We can extend the existing ARIMA model to handle these higher order lags (without having to include all of the intervening lags).

Seasonal ARIMA $(p,d,q) \times (P,D,Q)_s$:

$$\Phi_{P}(L^{s}) \, \phi_{P}(L) \, \Delta_{s}^{D} \, \Delta^{d} \, y_{t} = \delta + \Theta_{Q}(L^{s}) \, \theta_{q}(L) \, w_{t}$$

5

Seasonal ARIMA

We can extend the existing ARIMA model to handle these higher order lags (without having to include all of the intervening lags).

Seasonal ARIMA $(p, d, q) \times (P, D, Q)_s$:

$$\Phi_{P}(\mathbf{L}^{\mathrm{S}})\,\phi_{P}(\mathbf{L})\,\Delta_{\mathrm{S}}^{\mathrm{D}}\,\Delta^{\mathrm{d}}\,\mathbf{y}_{\mathrm{t}} = \delta + \Theta_{\mathrm{Q}}(\mathbf{L}^{\mathrm{S}})\,\theta_{\mathrm{q}}(\mathbf{L})\,\mathbf{w}_{\mathrm{t}}$$

where

$$\phi_p(L) = 1 - \phi_1 L - \phi_2 L^2 - \dots - \phi_p L^p$$

$$\theta_q(L) = 1 + \theta_1 L + \theta_2 L^2 + \dots + \theta_p L^q$$

$$\Delta^d = (1 - L)^d$$

$$\begin{aligned} \Phi_{P}(L^{s}) &= 1 - \Phi_{1}L^{s} - \Phi_{2}L^{2s} - \dots - \Phi_{P}L^{Ps} \\ \Theta_{Q}(L^{s}) &= 1 + \Theta_{1}L + \Theta_{2}L^{2s} + \dots + \theta_{P}L^{Qs} \\ \Delta_{s}^{D} &= (1 - L^{s})^{D} \end{aligned}$$

5

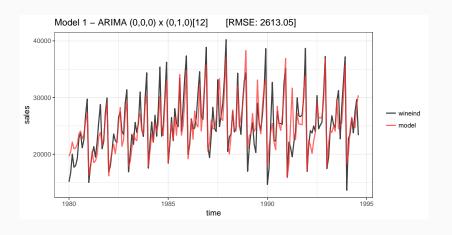
Seasonal ARIMA for wineind

Lets consider an ARIMA $(0,0,0) \times (1,0,0)_{12}$:

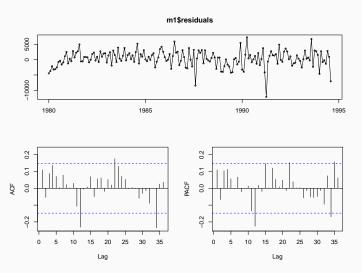
$$(1 - \Phi_1 L^{12}) y_t = \delta + w_t$$
$$y_t = \Phi_1 y_{t-12} + \delta + w_t$$

```
(m1 = Arima(wineind, seasonal=list(order=c(1,0,0), period=12)))
## Series: wineind
## ARIMA(0,0,0)(1,0,0)[12] with non-zero mean
##
## Coefficients:
## sar1 mean
## 0.8780 24489.243
## s.e. 0.0314 1154.487
##
## sigma^2 estimated as 6906536: log likelihood=-1643.39
## AIC=3292.78 AICc=3292.92 BIC=3302.29
```

Fitted model



Residuals



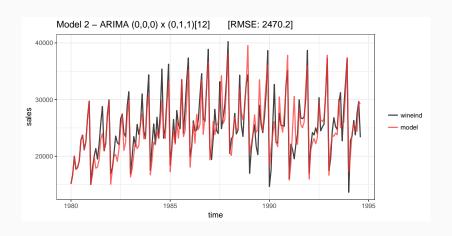
```
ARIMA(0,0,0) \times (0,1,1)_{12}:
```

$$(1 - L^{12})y_t = \delta + (1 + \Theta_1 L^{12})w_t$$

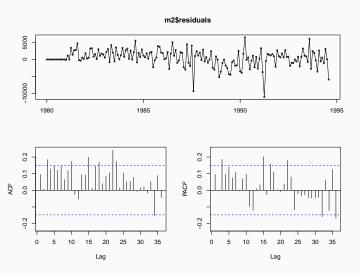
$$y_t - y_{t-12} = \delta + w_t + \Theta_1 w_{t-12}$$

$$y_t = \delta + y_{t-12} + w_t + \Theta_1 w_{t-12}$$

Fitted model

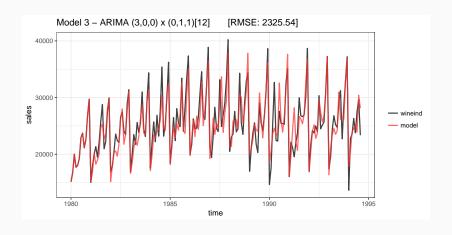


Residuals

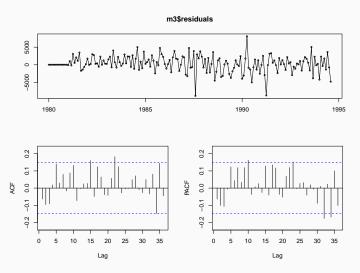


$$\begin{aligned} \text{ARIMA}(3,0,0) \times (0,1,1)_{12} \\ & (1-\phi_1 L-\phi_2 L^2-\phi_3 L^3) \left(1-L^{12}\right) y_t = \delta + (1+\Theta_1 L) w_t \\ & (1-\phi_1 L-\phi_2 L^2-\phi_3 L^3) \left(y_t-y_{t-12}\right) = \delta + w_t + w_{t-12} \\ & y_t = \delta + \sum_{i=1}^3 \phi_i y_{t-1} + y_{t-12} - \sum_{i=1}^3 \phi_i y_{t-12-i} + w_t + w_{t-12} \\ & (\text{m3 = Arima}(\text{wineind, order=c}(3,0,0), \\ & \text{seasonal=list}(\text{order=c}(0,1,1), \text{period=12}))) \\ \text{\## Series: wineind} \\ \text{\## ARIMA}(3,0,0)(0,1,1)[12] \\ \text{\##} \\ \text{\## Coefficients:} \\ \text{\## ar1 ar2 ar3 sma1} \\ \text{\## 0.1402 0.0806 0.3040 -0.5790} \\ \text{\## s.e. 0.0755 0.0813 0.0823 0.1023} \\ \text{\## sigma^2 estimated as 5948935: log likelihood=-1512.38} \\ \text{\## AIC=3034.77 AICc=3035.15 BIC=3050.27} \end{aligned}$$

Fitted model



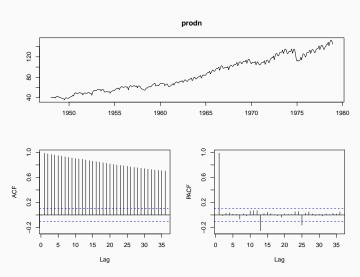
Model - Residuals



Federal Reserve Board Production Index

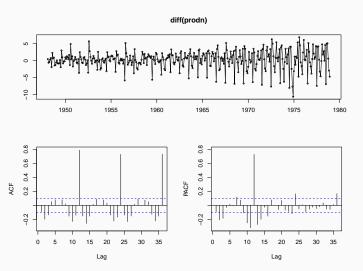
prodn from the astsa package

Monthly Federal Reserve Board Production Index (1948-1978)



Differencing

Based on the ACF it seems like standard differencing may be required

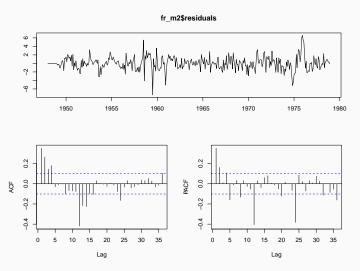


Differencing + Seasonal Differencing

Additional seasonal differencing also seems warranted

```
(fr_m1 = Arima(prodn, order = c(0,1,0),
            seasonal = list(order=c(0,0,0), period=12)))
## Series: prodn
## ARIMA(0.1.0)
##
## sigma^2 estimated as 7.147: log likelihood=-891.26
## AIC=1784.51 AICc=1784.52 BIC=1788.43
(fr_m2 = Arima(prodn, order = c(0,1,0),
           seasonal = list(order=c(0,1,0), period=12)))
## Series: prodn
## ARIMA(0,1,0)(0,1,0)[12]
##
## sigma^2 estimated as 2.52: log likelihood=-675.29
## ATC=1352.58 ATCc=1352.59 BTC=1356.46
```

Residuals

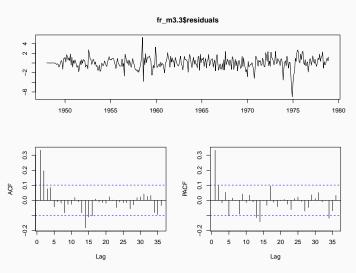


Adding Seasonal MA

```
(fr m3.1 = Arima(prodn, order = c(0,1,0),
           seasonal = list(order=c(0,1,1), period=12)))
## Series: prodn
## ARIMA(0,1,0)(0,1,1)[12]
##
## Coefficients:
##
           sma1
## -0.7151
## s.e. 0.0317
##
## sigma^2 estimated as 1.616: log likelihood=-599.29
## AIC=1202.57 AICc=1202.61 BIC=1210.34
(fr m3.2 = Arima(prodn, order = c(0,1,0),
           seasonal = list(order=c(0,1,2), period=12)))
## Series: prodn
## ARIMA(0,1,0)(0,1,2)[12]
##
## Coefficients:
##
           sma1 sma2
## -0.7624 0.0520
## s.e. 0.0689 0.0666
##
## sigma^2 estimated as 1.615: log likelihood=-598.98
## AIC=1203.96 AICc=1204.02 BIC=1215.61
```

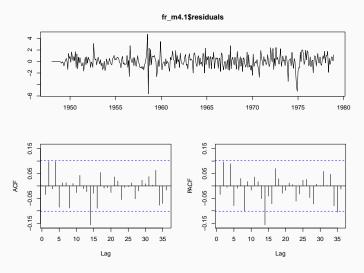
Adding Seasonal MA (cont.)

Residuals - Model 3.3

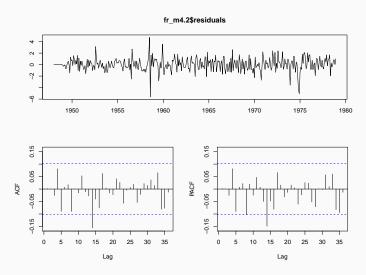


```
(fr m4.1 = Arima(prodn, order = c(1,1,0),
           seasonal = list(order=c(0,1,3), period=12)))
## Series: prodn
## ARIMA(1,1,0)(0,1,3)[12]
##
## Coefficients:
##
          ar1 sma1 sma2 sma3
## 0.3393 -0.7619 -0.1222 0.2756
## s.e. 0.0500 0.0527 0.0646 0.0525
##
## sigma^2 estimated as 1.341: log likelihood=-565.98
## AIC=1141.95 AICc=1142.12 BIC=1161.37
(fr m4.2 = Arima(prodn, order = c(2,1,0),
           seasonal = list(order=c(0,1,3), period=12)))
## Series: prodn
## ARIMA(2,1,0)(0,1,3)[12]
##
## Coefficients:
##
          ar1 ar2 sma1
                                 sma2 sma3
## 0.3038 0.1077 -0.7393 -0.1445 0.2815
## s.e. 0.0526 0.0538 0.0539 0.0653 0.0526
##
## sigma^2 estimated as 1.331: log likelihood=-563.98
## AIC=1139.97 AICc=1140.2 BIC=1163.26
```

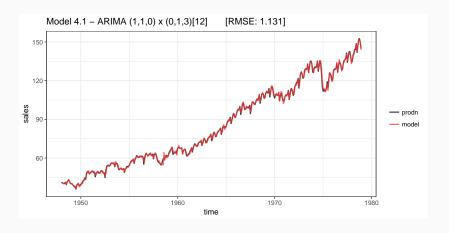
Residuals - Model 4.1



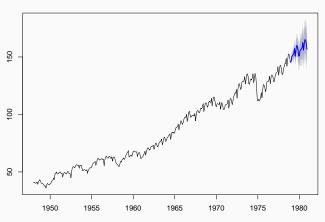
Residuals - Model 4.2



Model Fit

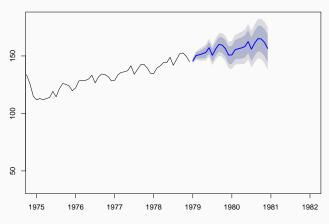






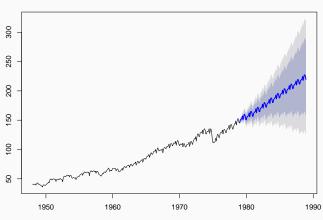
Model Forecast (cont.)

Forecasts from ARIMA(1,1,0)(0,1,3)[12]



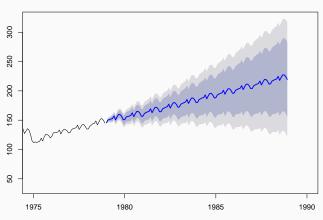
Model Forecast (cont.)





Model Forecast (cont.)

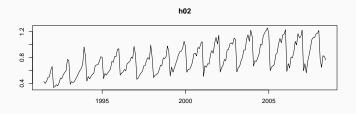


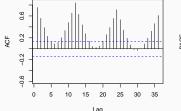


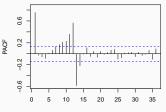
Exercise - Cortecosteroid Drug Sales

Monthly cortecosteroid drug sales in Australia from 1992 to 2008.

library(fpp)
tsdisplay(h02,points=FALSE)







Hint

ts.intersect(h02, log(h02)) %>% plot()



