Lecture 14

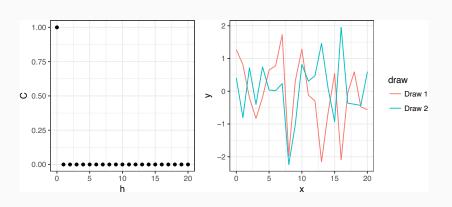
Covariance Functions

3/08/2018

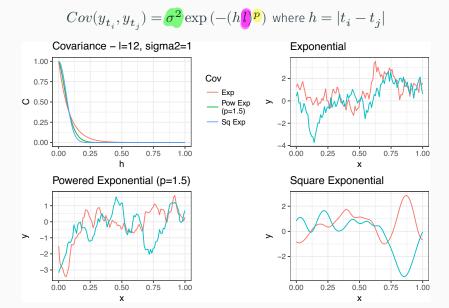
More on Covariance Functions

Nugget Covariance

$$Cov(\boldsymbol{y}_{t_i}, \boldsymbol{y}_{t_j}) = \sigma^2 \boldsymbol{1}_{\{h=0\}}$$
 where $h = |t_i - t_j|$



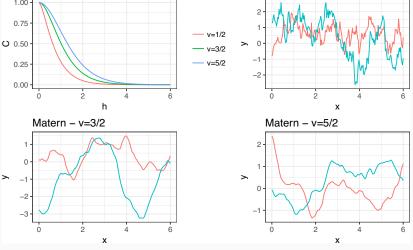
(- / Power / Square) Exponential Covariance



Matern Covariance

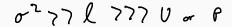
$$Cov(y_{t_i},y_{t_j}) = \sigma^2 \frac{2^{1-\nu}}{\Gamma(\nu)} \left(\sqrt{2\nu} \, h \cdot \boldsymbol{l}\right)^{\nu} \underline{K_{\nu}} \left(\sqrt{2\nu} \, h \cdot \boldsymbol{l}\right) \text{ where } h = |t_i - t_j|$$

$$\underbrace{\text{Covariance - I=2, sigma2=1}}_{0.75} \underline{\text{Matern - v=1/2}}_{2}$$

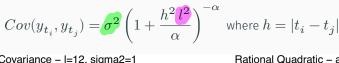


Matern Covariance

- \cdot $K_{
 u}$ is the modified Bessel function of the second kind.
- A Gaussian process with Matérn covariance has sample functions that are $\lceil \nu-1 \rceil$ times differentiable.
- When $\nu=1/2+p$ for $p\in\mathbb{N}^+$ then the Matern has a simplified form (product of an exponential and a polynomial of order p).
- · When $\nu=1/2$ the Matern is equivalent to the exponential covariance.
- As $\nu \to \infty$ the Matern converges to the square exponential covariance.
- . A Gaussian process with Matérii covariance has paths that are $|\nu|-1$ times differentiable.



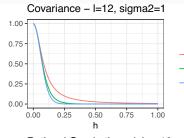
Rational Quadratic Covariance

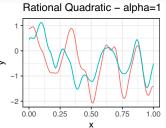


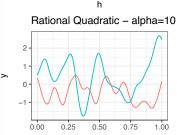
alpha=1

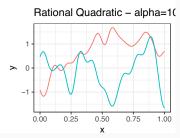
alpha=3

alpha=10







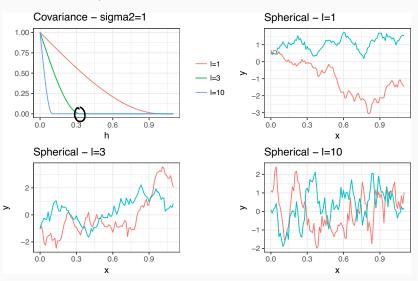


Rational Quadratic Covariance

- is a scaled mixture of squared exponential covariance functions with different characteristic length-scales (*l*).
- As $\alpha \to \infty$ the rational quadratic converges to the square exponential covariance.
- Has sample functions that are infinitely differentiable for any value of α

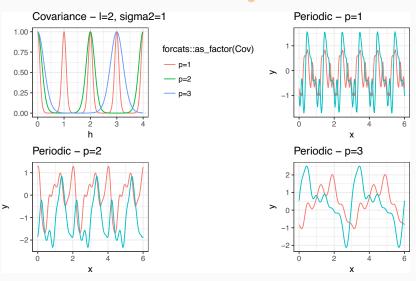
Spherical Covariance

$$Cov(\boldsymbol{y}_{t_i}, \boldsymbol{y}_{t_j}) = \begin{cases} \boldsymbol{\sigma^2} \big(1 - \frac{3}{2}h \cdot \boldsymbol{l} + \frac{1}{2}(h \cdot \boldsymbol{l})^3\big) \big) & \text{if } 0 < h < 1/l \\ 0 & \text{otherwise} \end{cases} \quad \text{where } h = |t_i - t_j|$$

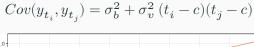


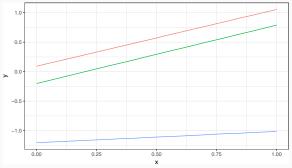
Periodic Covariance

$$Cov(y_{t_i},y_{t_j}) = \boxed{\sigma^2} \exp\left(-2 \boxed{t^2} \sin^2\left(\pi \frac{h}{p}\right)\right) \text{ where } h = |t_i - t_j|$$



Linear Covariance





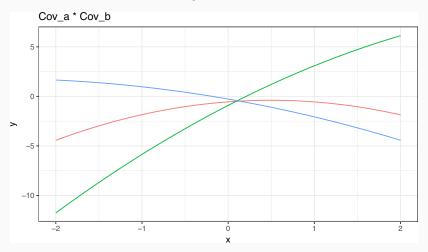
Combining Covariances

If we definite two valid covariance functions, $Cov_a(y_{t_i},y_{t_j})$ and $Cov_b(y_{t_i},y_{t_j})$ then the following are also valid covariance functions,

$$\begin{aligned} &Cov_a(y_{t_i}, y_{t_j}) + Cov_b(y_{t_i}, y_{t_j}) \\ &Cov_a(y_{t_i}, y_{t_j}) \times Cov_b(y_{t_i}, y_{t_j}) \end{aligned}$$

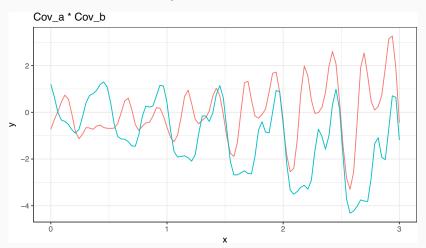
$Linear \times Linear \rightarrow Quadratic$

$$\begin{aligned} &Cov_a(y_{t_i}, y_{t_j}) = 1 + 2 \left(t_i \times t_j\right) \\ &Cov_b(y_{t_i}, y_{t_j}) = 2 + 1 \left(t_i \times t_j\right) \end{aligned}$$



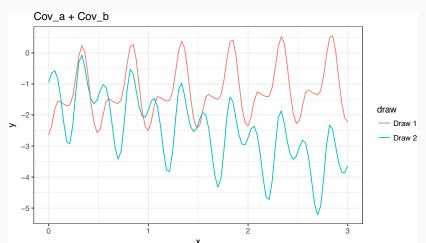
$Linear \times Periodic$

$$\begin{split} Cov_{a}(y_{t_{i}},y_{t_{j}}) &= 1 + 1 \; (t_{i} \times t_{j}) \\ Cov_{b}(y_{t_{i}},y_{t_{j}}) &= \exp\left(-2 \; \text{sin}^{2} \left(2 \pi \, h\right)\right) \end{split}$$



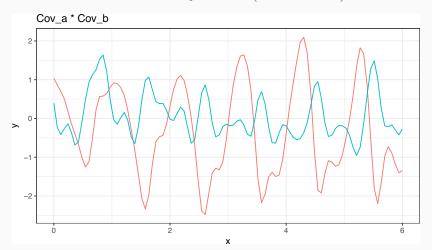
Linear + Periodic

$$\begin{split} Cov_{a}(y_{t_{i}},y_{t_{j}}) &= 1 + 1 \: (t_{i} \times t_{j}) \\ Cov_{b}(h = |t_{i} - t_{j}|) &= \exp\left(-2 \: \sin^{2}\left(2\pi \: h\right)\right) \end{split}$$



$\mathsf{Sq}\;\mathsf{Exp}\; imes\;\mathsf{Periodic} o \mathsf{Locally}\;\mathsf{Periodic}$

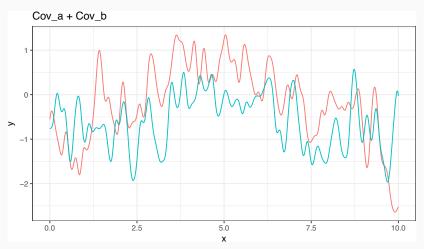
$$\begin{split} Cov_a(h = |t_i - t_j|) &= \exp(-(1/3)h^2) \\ Cov_b(h = |t_i - t_j|) &= \exp\left(-2\sin^2\left(\pi\,h\right)\right) \end{split}$$



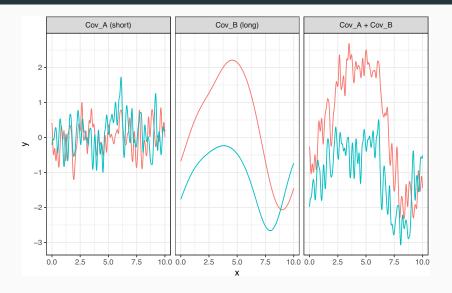
Sq Exp (short) + Sq Exp (long)

$$Cov_{a}(h = |t_{i} - t_{j}|) = (1/4) \exp(-4\sqrt{3}h^{2})$$

$$Cov_{b}(h = |t_{i} - t_{j}|) = \exp(-(\sqrt{3}/2)h^{2})$$



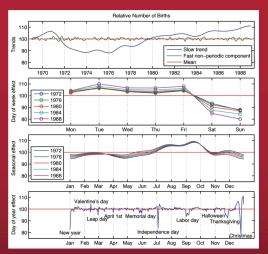
Sq Exp (short) + Sq Exp (long) (Seen another way)



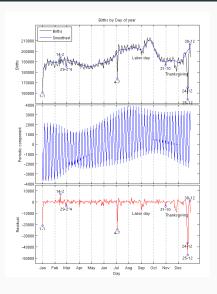
BDA3 example

Bayesian Data Analysis

Third Edition



Births (one year)



- 1. Smooth long term trend (sq exp cov)
- 2. Seven day periodic trend with decay ($periodic \times sq exp cov$)
- 3. Constant mean

Component Contributions

We can view our GP in the following ways,

$$\mathbf{y} \sim \mathcal{N}(\pmb{\mu},\; \pmb{\Sigma}_1 + \pmb{\Sigma}_2 + \sigma^2 \mathbf{I}\,)$$

but with appropriate conditioning we can also think of ${f y}$ as being the sum of multipe independent GPs

$$y = \mu + w_1(t) + w_2(t) + w_3(t)$$

where

$$\begin{split} w_1(\mathbf{t}) &\sim \mathcal{N}(0, \mathbf{\Sigma}_1) \\ w_2(\mathbf{t}) &\sim \mathcal{N}(0, \mathbf{\Sigma}_2) \\ w_3(\mathbf{t}) &\sim \mathcal{N}(0, \sigma^2 \mathbf{I}) \end{split}$$

Decomposition of Covariance Components

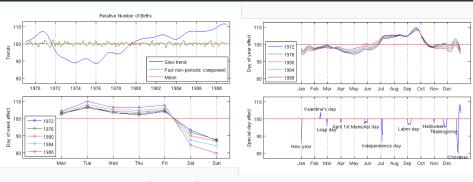
$$\begin{bmatrix} y \\ w_1 \\ w_2 \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \pmb{\mu} \\ 0 \\ 0 \end{bmatrix}, \ \begin{bmatrix} \Sigma_1 + \Sigma_2 + \sigma^2 \mathbf{I} & \Sigma_1 & \Sigma_2 \\ \Sigma_1 & \Sigma_1 & 0 \\ \Sigma_2 & 0 & \Sigma_2 \end{bmatrix} \right)$$

therefore

$$w_1 \mid \mathbf{y}, \boldsymbol{\mu}, \boldsymbol{\theta} \sim \mathcal{N}(\boldsymbol{\mu}_{cond}, \ \boldsymbol{\Sigma}_{cond})$$

$$\begin{split} \boldsymbol{\mu_{cond}} &= 0 + \boldsymbol{\Sigma_1} \; (\boldsymbol{\Sigma_1} + \boldsymbol{\Sigma_2} + \boldsymbol{\sigma^2} \boldsymbol{I})^{-1} (\mathbf{y} - \boldsymbol{\mu}) \\ \boldsymbol{\Sigma_{cond}} &= \boldsymbol{\Sigma_1} - \boldsymbol{\Sigma_1} (\boldsymbol{\Sigma_1} + \boldsymbol{\Sigma_2} + \boldsymbol{\sigma^2} \mathbf{I})^{-1} \boldsymbol{\Sigma_1}^t \end{split}$$

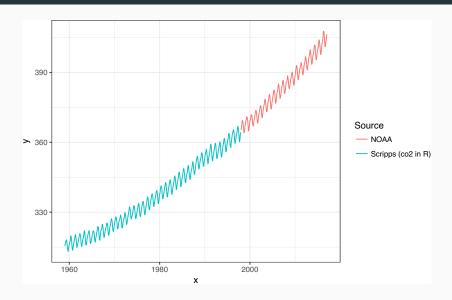
Births (multiple years)



- 1. slowly changing trend (sq exp cov)
- 2. small time scale correlating noise (sq exp cov)
- 3. 7 day periodical component capturing day of week effect (periodic \times sq exp cov)
- 4. 365.25 day periodical component capturing day of year effect (periodic × sq exp cov)
- component to take into account the special days and interaction with weekends (linear cov)
- 6. independent Gaussian noise (nugget cov)
- 7. constant mean



${\it Atmospheric}~{\it CO}_2$



Based on Rasmussen 5.4.3 (we are using slightly different data and parameterization)

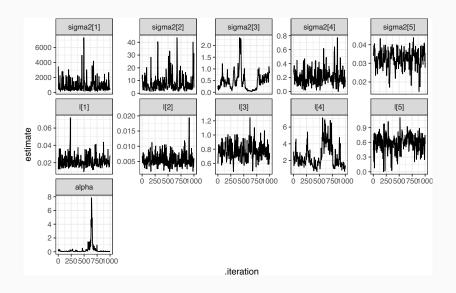
$$\mathbf{y} \sim \mathcal{N}(\pmb{\mu},\; \pmb{\Sigma}_1 + \pmb{\Sigma}_2 + \pmb{\Sigma}_3 + \pmb{\Sigma}_4 + \sigma^2 \mathbf{I})$$

$$\{\boldsymbol{\mu}\}_i = \bar{y}$$

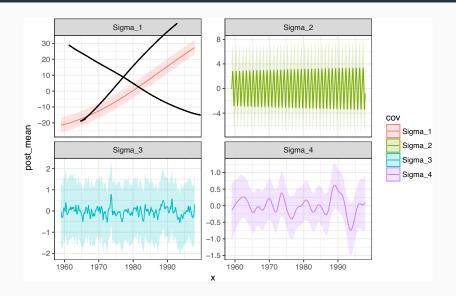
$$\begin{split} \left\{ \boldsymbol{\Sigma}_1 \right\}_{ij} &= \sigma_1^2 \exp \left(-(l_1 \cdot d_{ij})^2 \right) \\ \left\{ \boldsymbol{\Sigma}_2 \right\}_{ij} &= \sigma_2^2 \exp \left(-(l_2 \cdot d_{ij})^2 \right) \exp \left(-2 \left(l_3 \right)^2 \sin^2 (\pi \, d_{ij}/p) \right) \\ \left\{ \boldsymbol{\Sigma}_3 \right\}_{ij} &= \sigma_3^2 \left(1 + \frac{(l_4 \cdot d_{ij})^2}{\alpha} \right)^{-\alpha} \\ \left\{ \boldsymbol{\Sigma}_4 \right\}_{ij} &= \sigma_4^2 \exp \left(-(l_5 \cdot d_{ij})^2 \right) \end{split}$$

```
ml model = "model{
        v ~ dmnorm(mu, inverse(Sigma))
        for (i in 1:(length(v)-1)) {
                 for (i in (i+1):length(v)) {
                         k1[i,j] \leftarrow sigma2[1] * exp(-pow(l[1] * d[i,j],2))
                         k2[i,j] < sigma2[2] * exp(-pow(l[2] * d[i,j],2) - 2 * pow(l[3] * sin(pi*d[i,j],2) - 2 * pow(l[3] *
                         k3[i,j] < - sigma2[3] * pow(1+pow(l[4] * d[i,j],2)/alpha, -alpha)
                         k4[i,j] <- sigma2[4] * exp(- pow(l[5] * d[i,j],2))
                         Sigma[i,j] \leftarrow k1[i,j] + k2[i,j] + k3[i,j] + k4[i,j]
                        Sigma[j,i] <- Sigma[i,j]</pre>
        for (i in 1:length(v)) {
                 Sigma[i,i] \leftarrow Sigma2[1] + Sigma2[2] + Sigma2[3] + Sigma2[4] + Sigma2[5]
        for(i in 1:5){
                 sigma2[i] \sim dt(0, 2.5, 1) T(0,)
                l[i] \sim dt(0. 2.5. 1) T(0.)
        alpha \sim dt(0, 2.5, 1) T(0,)
```

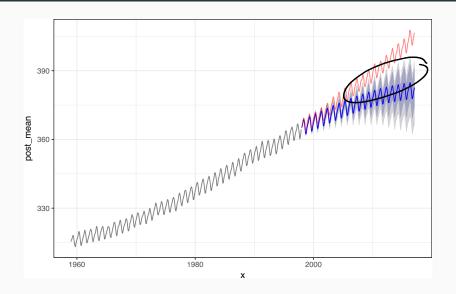
Diagnostics



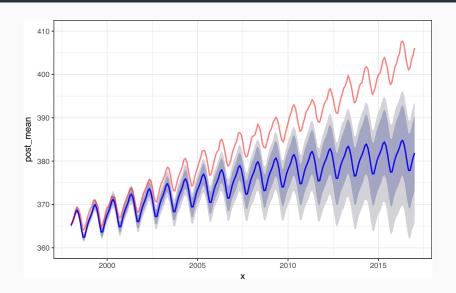
Fit Components



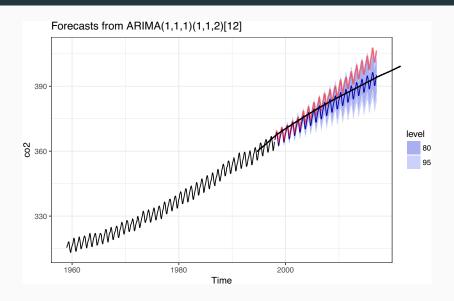
Forecasting



Forecasting (zoom)



Forecasting ARIMA (auto)



Forecasting RMSE

dates	RMSE (arima)	RMSE (gp)
Jan 1998 - Jan 2003	1.103	1.911
Jan 1998 - Jan 2008	2.506	4.575
Jan 1998 - Jan 2013	3.824	7.706
Jan 1998 - Mar 2017	5.461	11.395

Forecasting Components

