

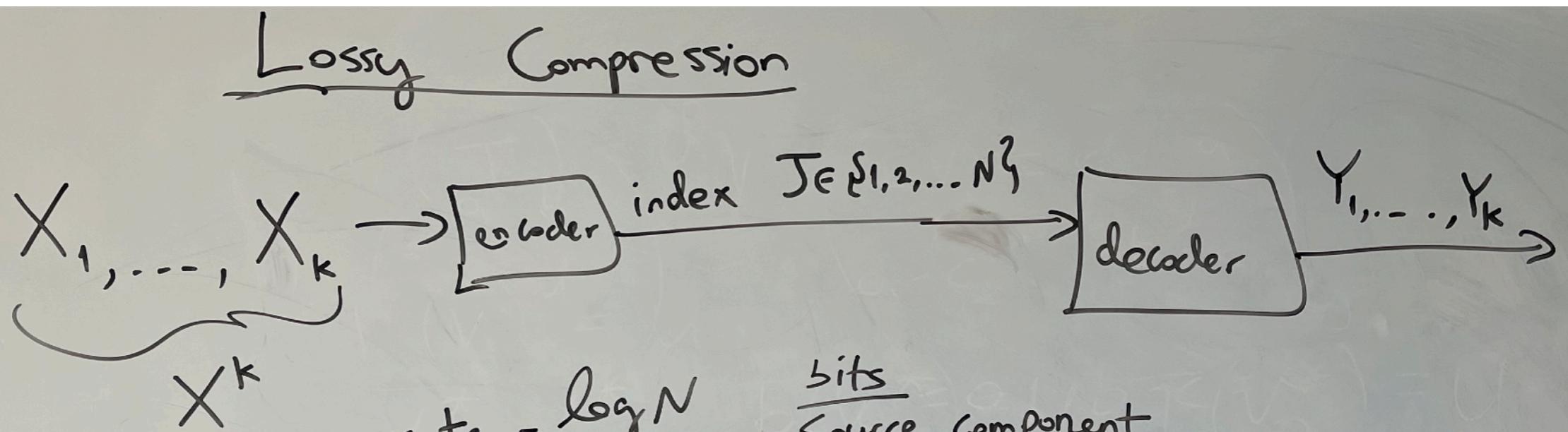
**EE 274:** Data Compression,  
Theory and Applications  
(Aut 22/23)



KEEP  
**CALM**  
AND  
**COMPRESS**  
**DATA**

water pouring characterization of  $R(D)$  for the Gaussian source

info source :



$$\text{rate} = \frac{\log N}{k} \quad \begin{matrix} \text{bits} \\ \text{Source Component} \end{matrix}$$

$$\text{distortion: } d(X^k, Y^k) = \frac{1}{k} \sum_{i=1}^k d(X_i, Y_i)$$

$R(D) = \min$  rate needed to achieve  $D$  distortion  
no more than  $D$

(optimizing  
across  $k$   
and encoders  
+ decoders)

Suppose  $X = (X_1, X_2, \dots)$  is a stationary source.

Shannon's theorem for lossy compression:

$$R(D) = \lim_{k \rightarrow \infty} \min_{\text{Ed}(X^k, Y^k)} \frac{1}{k} I(X^k; Y^k)$$

## $R(D)$ for a Gaussian Source: I

Denote  $R_G(\sigma^2, D) \triangleq \min_{E[(X-Y)^2] \leq D} I(X; Y)$ , when  $X \sim N(0, \sigma^2)$ .

For  $\vec{\sigma}^2 = \begin{pmatrix} \sigma_1^2 \\ \vdots \\ \sigma_n^2 \end{pmatrix}$  denote  $R_G(\vec{\sigma}^2, D) \triangleq \min_{E[\|X^n - Y^n\|^2] \leq D} \frac{1}{n} I(X^n; Y^n)$ ,

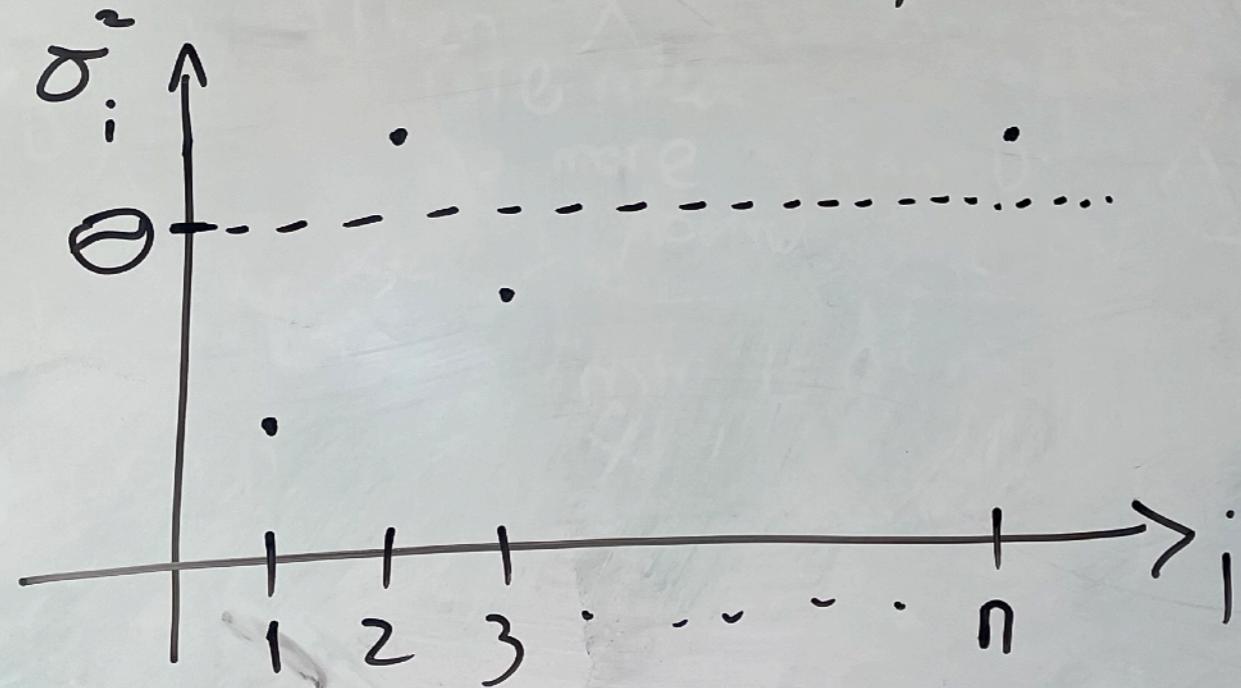
when  $X_1, \dots, X_n$  are independent with  $X_i \sim N(0, \sigma_i^2)$ .

## $R(D)$ for a Gaussian Source: II

Recall:  $R_G(\bar{\sigma}^2, D) = \min \frac{1}{n} \sum_{i=1}^n \left[ \frac{1}{2} \log \frac{\sigma_i^2}{D_i} \right]_+$   
 s.t.  $\frac{1}{n} \sum_{i=1}^n D_i \leq D$

and is given parametrically by the curve:

$$D_\theta = \frac{1}{n} \sum_{i=1}^n \min\{\theta, \sigma_i^2\}, \quad R_\theta = \frac{1}{n} \sum_{i=1}^n \left[ \frac{1}{2} \log \frac{\sigma_i^2}{\theta} \right]_+ \quad (\theta > 0)$$



## $R(D)$ for a Gaussian Source : III

$X^n$  Gaussian with covariance  $\Phi_{X^n}$

Then  $R(X^n, D)$  (under squared error)  
is given by  $R_G(\vec{\lambda}, D)$  where

$$\vec{\lambda} = \begin{pmatrix} \lambda_1 \\ \vdots \\ \lambda_n \end{pmatrix}$$

are the eigenvalues of  $\Phi_{X^n}$

## R(D) for a Gaussian Source: IV

When  $X^n$  are the first  $n$  components of a stationary Gaussian process  $X$  with covariance matrix  $\Phi_n = \{\phi_{|i-j|}\}_{\substack{1 \leq i \leq n \\ 1 \leq j \leq n}}$  where  $\phi_k = \text{Cov}(X_i, X_{i-k})$

$$R(X^n, D) = R_G(\vec{\lambda}^{(n)}, D)$$

where  $\vec{\lambda}^{(n)}$  is the vector of eigenvalues of  $\Phi_n$

## R(D) for a Gaussian Source: V

Theorem (Toeplitz distribution):

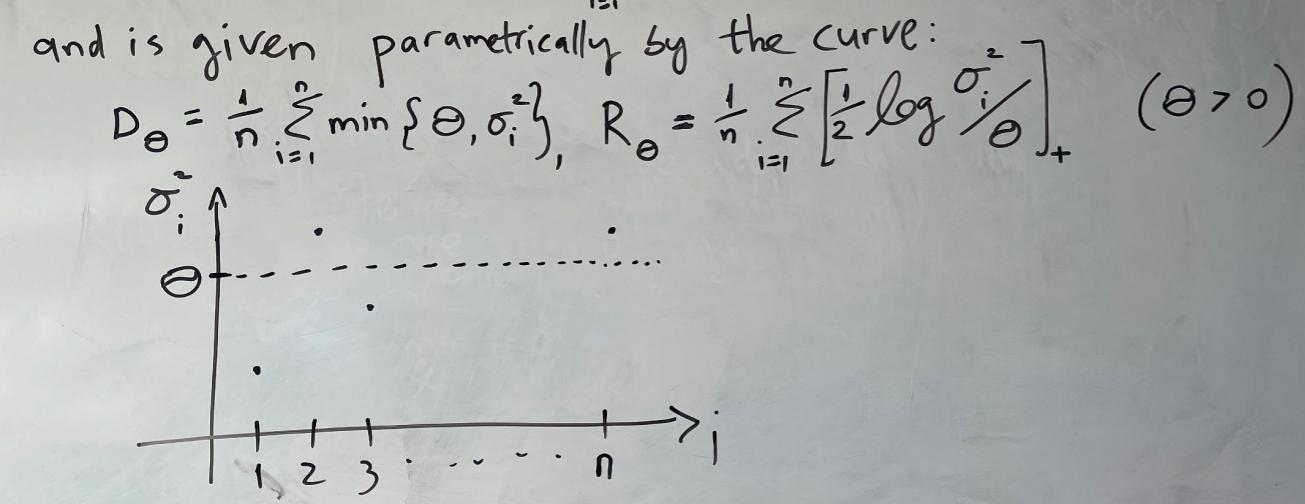
Let  $S(\omega) \triangleq \sum_{k=-\infty}^{\infty} \phi_k e^{-j\omega k}$  be the spectral density of  $X$  and  $G(\cdot)$  a continuous function.

Then  $\frac{1}{n} \sum_{i=1}^n G(\lambda_i^{(n)}) \xrightarrow{n \rightarrow \infty} \frac{1}{2\pi} \int_{-\pi}^{\pi} G(S(\omega)) d\omega$

# recap

## $R(D)$ for a Gaussian Source: II

Recall:  $R_G(\vec{\sigma}^2, D) = \min_{\text{s.t. } \frac{1}{n} \sum_{i=1}^n D_i \leq D} \frac{1}{n} \sum_{i=1}^n \left[ \frac{1}{2} \log \frac{\sigma_i^2}{D_i} \right]_+$



## $R(D)$ for a Gaussian Source: III

$X^n$  Gaussian with covariance  $\Phi_{X^n}$

Then  $R(X^n, D)$  (under squared error) is given by  $R_G(\vec{\lambda}, D)$  where  $\vec{\lambda} = \begin{pmatrix} \lambda_1 \\ \vdots \\ \lambda_n \end{pmatrix}$  are the eigenvalues of  $\Phi_{X^n}$

## $R(D)$ for a Gaussian Source: IV

When  $X^n$  are the first  $n$  components of a stationary Gaussian process  $X$  with covariance matrix  $\Phi_n = \{\phi_{|i-j|}\}_{1 \leq i \leq n, 1 \leq j \leq n}$  where  $\phi_k = \text{Cov}(X_i, X_{i-k})$

$$R(X^n, D) = R_G(\vec{\lambda}^{(n)}, D)$$

where  $\vec{\lambda}^{(n)}$  is the vector of eigenvalues of  $\Phi_n$

## $R(D)$ for a Gaussian Source: V

Theorem (Toeplitz distribution):

Let  $S(w) \triangleq \sum_{k=-\infty}^{\infty} \phi_k e^{-jwk}$  be the spectral density of  $X$  and  $G(\cdot)$  a continuous function.

$$\text{Then } \frac{1}{n} \sum_{i=1}^n G(\lambda_i^{(n)}) \xrightarrow{n \rightarrow \infty} \frac{1}{2\pi} \int_{-\pi}^{\pi} G(S(w)) dw$$

## R(D) for a Gaussian Source: VI

Specializing to  $G(\lambda) = \min\{\Theta, \lambda^2\}$  and to  $G(\lambda) = \left[\frac{1}{2} \log \frac{\lambda}{\Theta}\right]_+$   
we get:

The rate distortion function of a stationary Gaussian process with spectral density  $S(\omega)$   
is given parametrically by:

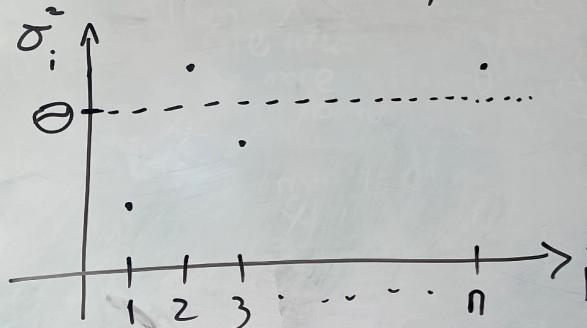
$$D_\Theta = \frac{1}{2\pi} \int_{-\pi}^{\pi} \min\{\Theta, S(\omega)\} d\omega, \quad R_\Theta = \frac{1}{4\pi} \int_{-\pi}^{\pi} \left[ \log \frac{S(\omega)}{\Theta} \right]_+ d\omega$$

## R(D) for a Gaussian Source: II

$$\text{Recall: } R_G(\vec{\sigma}^2, D) = \min_{\text{s.t. } \frac{1}{n} \sum_{i=1}^n D_i \leq D} \frac{1}{n} \sum_{i=1}^n \left[ \frac{1}{2} \log \frac{\sigma_i^2}{D_i} \right]_+$$

and is given parametrically by the curve:

$$D_\Theta = \frac{1}{n} \sum_{i=1}^n \min\{\Theta, \sigma_i^2\}, \quad R_\Theta = \frac{1}{n} \sum_{i=1}^n \left[ \frac{1}{2} \log \frac{\sigma_i^2}{\Theta} \right]_+ \quad (\Theta > 0)$$



## R(D) for a Gaussian Source: IV

When  $X^n$  are the first  $n$  components of a stationary Gaussian process  $X$  with covariance matrix  $\Phi_n = \{\Phi_{|i-j|}\}_{1 \leq i \leq n, 1 \leq j \leq n}$  where  $\Phi_k = \text{Cov}(X_i, X_{i-k})$

$$R(X^n, D) = R_G(\vec{\lambda}^{(n)}, D)$$

where  $\vec{\lambda}^{(n)}$  is the vector of eigenvalues of  $\Phi_n$

## R(D) for a Gaussian Source: III

$X^n$  Gaussian with covariance  $\Phi_{X^n}$

Then  $R(X^n, D)$  (under squared error) is given by  $R_G(\vec{\lambda}, D)$  where  $\vec{\lambda} = \begin{pmatrix} \lambda_1 \\ \vdots \\ \lambda_n \end{pmatrix}$  are the eigenvalues of  $\Phi_{X^n}$

## R(D) for a Gaussian Source: V

Theorem (Toeplitz distribution):

Let  $S(w) \triangleq \sum_{k=-\infty}^{\infty} \phi_k e^{-jwk}$  be the spectral density of  $X$  and  $G(\cdot)$  a continuous function.

$$\text{Then } \frac{1}{n} \sum_{i=1}^n G(\lambda_i^{(n)}) \xrightarrow{n \rightarrow \infty} \frac{1}{2\pi} \int_{-\pi}^{\pi} G(S(w)) dw$$

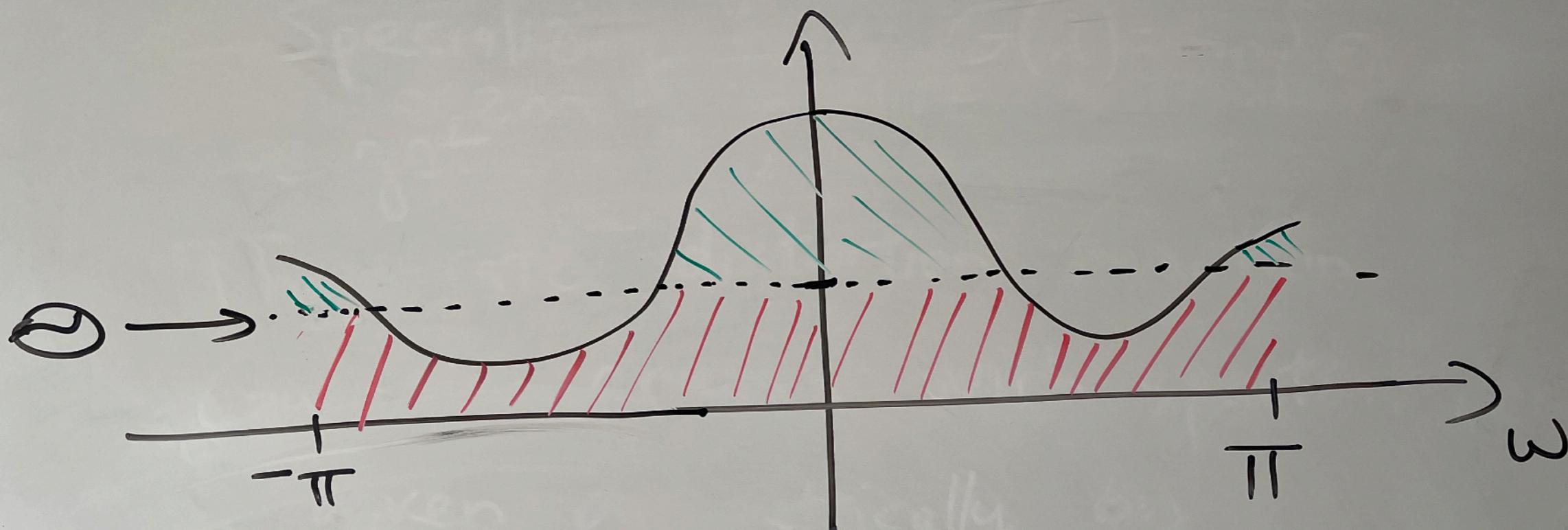
## R(D) for a Gaussian Source: VI

Specializing to  $G(\lambda) = \min\{\Theta, \lambda^2\}$  and to  $G(\lambda) = \left[ \frac{1}{2} \log \frac{\lambda^2}{\Theta} \right]_+$  we get:

The rate distortion function of a stationary Gaussian process with spectral density  $S(w)$  is given parametrically by:

$$D_\Theta = \frac{1}{2\pi} \int_{-\pi}^{\pi} \min\{\Theta, S(w)\} dw, \quad R_\Theta = \frac{1}{4\pi} \int_{-\pi}^{\pi} \left[ \log \frac{S(w)}{\Theta} \right]_+ dw$$

# R(D) for a Gaussian Source: VII



$$D_\Theta = \text{Area}(\text{---})$$

$$R_\Theta = \text{Area in log scale} (\text{---})$$

## R(D) for a Gaussian Source: VI

Specializing to  $G(2) = \min\{\Theta, 2\}$  and to  $G(2) = \left[\frac{1}{2} \log \frac{2}{\Theta}\right]_+$   
we get:

The rate distortion function of a stationary Gaussian process with spectral density  $S(\omega)$   
is given parametrically by:

$$D_\Theta = \frac{1}{2\pi} \int_{-\pi}^{\pi} \min\{\Theta, S(\omega)\} d\omega, \quad R_\Theta = \frac{1}{4\pi} \int_{-\pi}^{\pi} \left[ \log \frac{S(\omega)}{\Theta} \right]_+ d\omega$$

## R(D) for a Gaussian Source: VIII

For  $D \leq \min_w S(w)$  can verify that

$$R(D) = \left[ \frac{1}{2} \log \frac{\sigma^2}{D} \right]_+$$

where  $\sigma^2$  is the variance of the innovations of  $X$

("kind-of" justification/inspiration for  
predictive coding)