

Improving the Replicability of Psychological Science Through Pedagogy

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Abstract

Replications are important to science, but who will do them? One proposal is that students can conduct direct replications as part of their training. As a proof-of-concept for this idea, here we report a series of 11 pre-registered, direct replications of findings from the 2015 volume of *Psychological Science*, all conducted as part of a graduate-level course. Congruent with previous studies, replications typically yielded smaller effects than originals: The modal outcome was partial support for the original claim. This work documents the challenges facing motivated students in reproducing previously published results on a first attempt. We describe the workflow and pedagogical methods that were used in the class and discuss challenges both for the adoption of this pedagogical model and for replication research more broadly.

Keywords: Replication; Reproducibility; Pedagogy; Experimental Methods

Improving the Replicability of Psychological Science Through Pedagogy

Replicability is a core value for empirical research and there is increasing concern throughout psychology that more independent replication is necessary (Open Science Collaboration, 2015; Wagenmakers, Wetzels, Borsboom, Maas, & Kievit, 2012). Yet under the current incentive structure for science, direct replication is not typically valued. One potential solution to this problem is to make direct replication an explicit part of pedagogy: that is, to teach students about experimental methods by asking them to run replication studies (Frank & Saxe, 2012; Grahe et al., 2012). Despite enthusiasm for this idea (Everett & Earp, 2015; M. King et al., 2016; LeBel, 2015; Standing, 2016), there is limited data beyond anecdotal reports or individual projects (Lakens, 2013; Phillips et al., 2015) to support its efficacy.

In the current article, we report the results of replication projects conducted in a single class (a graduate-level experimental methods course). As part of the required work of the course, students conducted high-power replications of published articles from the 2015 volume of the journal *Psychological Science*. These studies both give evidence about the state of replicability in the field and provide insight into the challenges and promise of pedagogical replication.

Past work has explicitly attempted to estimate the proportion of true effects in a particular literature (Ioannidis, 2005; Open Science Collaboration, 2015), but this goal is challenging – perhaps even impossible (Anderson et al., 2016; Gilbert, King, Pettigrew, & Wilson, 2016). Presupposing that this construct is even meaningful (cf. Gergen, 1973), assessing the truth of even an individual effect relies on delineating what constitutes an exact replication (e.g., Stroebe & Strack, 2014) and establishing a standard for replication success (e.g., Simonsohn, 2015). Estimating the proportion of true effects in a literature further requires a systematic sampling strategy to ensure

generality. Furthermore, no single replication attempt is definitive, so multiple iterations are typically necessary to increase the evidential value of replication attempts (Klein et al., 2014; M. L. Lewis & Frank, in press).

In the current work, we do not address these issues; instead, our goal is more modest. We assess the challenges facing a student in choosing an article of interest and – in a single attempt, within constraints of budget, expertise, and effort – reproducing the findings. We consider a number of criteria for replication success, including statistical significance, effect size, a Bayesian measure of evidence (Etz & Vandekerckhove, 2016), and a subjective assessment with respect to the original authors' interpretations. While each of these is imperfect, taken together these measures help provide a sense of how easy it is for a student to reproduce an effect to the degree that they could confidently build on it in future work. But, as we discuss below, the results of these replications may have different interpretations: While some are more direct, others differ on theoretically-relevant dimensions. Indeed, due to differences in population, timing, or method, some of our findings might be construed as assessing the boundary conditions of a particular effect rather than replicating the original exactly. Even in these cases, however, our work contributes to an understanding of the robustness, generalizability, and expected effect sizes of previously published findings.

We also describe our process for conducting well-powered, high-quality, direct replications as part of classroom pedagogy. Although mentorship in specific experimental methods is an important part of the standard advising relationship, the classroom context allows for elucidation of general principles of good research and discussion of how they can be modified to fit specific instances. And replication research specifically illustrates a number of important concepts – experimental design, power analysis, reporting standards, and preregistration, among others – in a way that more

open-ended projects, which require new conceptual development, often do not (see Frank & Saxe, 2012 for a more extended argument). There are significant limitations on what can be done in a single term, within the constraints of a course budget and the instructors' expertise, and we return to these in the Discussion. Nevertheless, were this approach implemented more widely, we believe the dividends paid to the field as a whole would be considerable.

Citation (Psychological Science; 2015)	Expt.	Open			On MTurk?	N (orig)	Power Standard	N (rep)	Instructor Fidelity	Project link
		Data?	Materials?	Open						
Atir, Rosenzweig, & Dunning	1b	No	Yes	Yes	Yes	202	Other	50	6.67	https://osf.io/uegk5/
Ko, Sadler, & Galinsky	2	Yes	Some	No	No	40	Original	40	4.67	https://osf.io/g6k21/
Lewis & Oyserman	4	No	Yes	Yes	Yes	122	80% power	128	6.67	https://osf.io/rdw6e/
Liverence & Scholl	1	No	Some	No	No	18	Original*	19	5.33	https://osf.io/sfxrb/
Proudfoot, Kay, & Koval	1	No	No	Yes	Yes	80	80% power	84	6.67	https://osf.io/ne8pl/
Scopelliti, Loewenstein, & Vosgerau	3	Yes	Yes	Yes	Yes	550	Other	124	5.67	https://osf.io/32wsn/
Sofer et al.	1	No	Yes	No	No	48	Other	95	4.33	https://osf.io/5wmh1/
Storm & Stone	3	No	No	No	No	48	Original	61	4.00	https://osf.io/9cgpp/
Wang et al.	2	No	No	No	No	219	80% power	397	4.33	https://osf.io/ng96w/
Xu & Franconeri	1a	No	No	No	No	12	Other*	27	5.00	https://osf.io/cgvzs/
Zaval, Markowitz, & Weber	1	Yes	Yes	Yes	Yes	312	Original	321	5.00	https://osf.io/degwa/

Table 1

*Summary characteristics of original studies and our replications, along with links to project pages for individual projects. * marks projects where the number of trials was modified.*

Methods

We begin by describing the general method of the class, then discuss shared methods for the replications and the process of development for each project.

General Class Outline

All projects were completed as part of a graduate-level methodology class (syllabus and materials, as well as code and data for this paper, available at <https://osf.io/98ta4>). At the initiation of class, all students were told that they had the opportunity to contribute their individual class assignment to a group replication project. The requirements for joining the project were to conduct a pre-registered replication of a finding from the 2015 volume of *Psychological Science* and to contribute the code, data, and materials to the writeup.

Students selected projects based on interest, rather than via a systematic sampling strategy. After selecting a paper, students chose a particular study from within that paper based on their judgment of interest and feasibility. The modal study chosen was the first, but several students chose later studies in a set as well. Thus, all interpretation of our findings should be with respect to this non-probability sample. This sampling strategy may thus bias our results towards more exciting or novel findings, as well as those that were more feasible for students.

We chose to use Amazon Mechanical Turk (AMT) as a platform for our replication studies. This choice was made to facilitate the process of recruiting samples large enough to enable replications of between-subjects designs (which typically require large samples) as well as to help students learn about a valuable resource for recruitment whose use requires some specialized knowledge. Our general class structure does not depend on the use of AMT, however, and the use of this particular online sample should

be viewed as limiting the interpretation of some results.

A summary of the original experiments chosen for replication is given in Table 1. Students chose findings from a wide variety of domains for replication. One cluster of studies included investigations of memory and visual attention (e.g., Xu & Franconeri, 2015), for example in applied contexts like remembering “deleted” files (Storm & Stone, 2015) or tracking transition between locations on smartphone apps (Liverence & Scholl, 2015). A second cluster came from studies of social perception and judgment, including judgments of faces and voices (Ko & Galinsky, 2015; Sofer, Dotsch, Wigboldus, & Todorov, 2015), as well as studies of attributions of modesty, creativity, and expertise (Atir, Rosenzweig, & Dunning, 2015; Proudfoot, Kay, & Koval, 2015; Scopelliti, Loewenstein, & Vosgerau, 2015). Other studies investigated increasing retirement savings by orienting participants to the future (N. A. Lewis & Oyserman, 2015), legacy priming (Zaval, Markowitz, & Weber, 2015), and the effects of math anxiety on performance (Wang et al., 2015). Details of individual replications are available in SOM.

Participants

All study participants were recruited on AMT. Individual sample sizes are given in Table 1. Each sample was recruited independently, using the same account but a different title. All experiments were approved by the Stanford University Institutional Review Board under protocol #23274, “Reproducibility of psychological science and instruction.”

Of the 11 studies included in the final sample, 5 (45%) were originally conducted on AMT. For the other studies in our sample, demographic differences (in terms of age, sex, socio-economic status, and in some cases, national origin) are an important factor in interpreting our findings; see SOM for more discussion of this issue in individual cases.

Repeat administration of experiments on AMT has been raised as an issue in some prior work (e.g., Chandler, Mueller, & Paolacci, 2014), but tracking participation in specific paradigms is an open challenge. We did not ask participants whether they had participated in similar research previously, however. First, half of the AMT population is estimated to change every six months (Stewart et al., 2015), and we expected that most of the studies in our sample that had been conducted on AMT had been performed at least a year previously (though see SOM for discussion of one particular paradigm that has been used more recently). Second, we suspected that asking this sort of question would lead to a large number of inaccurate responses.

Determining appropriate sample sizes is a major challenge in replication research (Button et al., 2013; Simonsohn, 2015). Post-hoc power analyses are problematic because of the likelihood of inflation of effect sizes due to the “winner’s curse” (Button et al., 2013; Hoenig & Heisey, 2001), but – especially within the constraints of a limited budget – it can be impossible to follow more conservative guidelines in all cases. For example, Simonsohn (2015) recommends 2.5x the original sample, which can be feasible for small or under-powered original studies but may lead to impractical or unnecessary recommendations for experiments that were initially large and/or adequately powered.

We eventually used a hybrid, case-by-case selection criterion that attempted to maximize the success of the project while staying within the constraints of our budget. For all criteria, experimenters recruited an additional 5% of participants to ensure we still achieved desired power if some participants skipped through or did not complete the experiment.¹ In three cases, we powered the replication attempt to 80% post-hoc

¹We initially were allocated \$1,500 but one student contributed personal research funds, leading to a total cost of \approx \$1,700 for all studies. We attempted to set payment for our studies at approximately \$6/hour based on timing estimated from Pilot A sessions.

power. In four cases, we used the original sample size. In one case, the original experiment was extremely over-powered in the post-hoc analysis, so we chose a smaller sample size that still would yield high power. In two others, we reduced the number of trials as part of the adaptation to Mechanical Turk (in one of these, changing the sample size to compensate). In the last case, the original study used a multi-step procedure where the first step consisted of stimulus generation; we decreased the N in the first stage of the procedure and focused on achieving high power at the second stage.

Materials and Methods

Of the 11 studies included in the final sample, 64% of studies had either some or all materials openly available. For 6 studies, however, students contacted the authors to request either materials or clarification of methods (using a template email that was customized by the students and reviewed by the instructors). Responses were received in all but one case, all within a matter of days.

Workflow

All materials and methods for our replication studies are available at <https://github.com/StanfordPsych254>. All experiments were coded in JavaScript, HTML, and CSS so as to be run in a standard web browser. All analyses were written using R, a free and open-source platform for statistical analysis. A schematic of our class workflow is shown in Figure 1.

Students wrote their final reports using R Markdown, a “literate programming” environment in which statistical analysis code can be interspersed with text, figures, and tables. This part of our workflow was extremely useful both pedagogically and for encouraging reproducible research practices. Using the template developed by Open

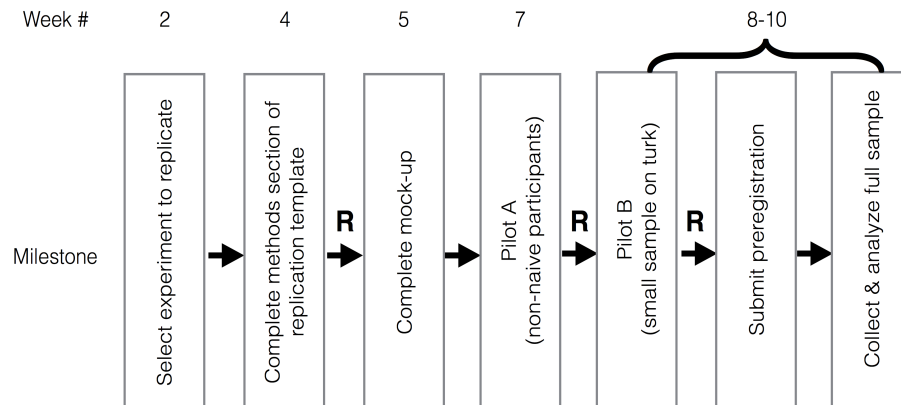


Figure 1. A schematic view of our class timeline for replication projects. R indicates the approximate timing of instructor team reviews of student materials.

Science Collaboration (2015) for replication reports, students created dynamic documents that included their proposed replication study, the code necessary to analyze their data, and the outputs of their data analysis (e.g., statistical tests and figures). This method of writing allows students to share a single compiled document via a hyperlink, facilitating review of writing, results, and code together in a single platform.² This writing method is also likely to reduce the frequency of statistical reporting errors (which appear to be regrettably common; Nuijten, Hartgerink, Assen, Epskamp, & Wicherts, 2015), given that errors are often introduced by transferring results between statistics and word-processing software packages.

Preregistration and Review

In addition to as-needed guidance on projects, students went through a process in which each student's work was reviewed several times by both the instructor (MCF) and the teaching assistants (ENS and RXDH). Each student collected two pilot samples and

²The current manuscript is written in this fashion as well.

review was coordinated with these. The first, “Pilot A,” consisted of a handful of non-naive participants (e.g., the experimenter, other students). The goal of this pilot was to ensure that all needed data were being logged by the experiment software and that analytic scripts for the confirmatory analyses functioned appropriately. After Pilot A was completed, the instructor or a TA critiqued and reviewed the student’s experimental script, analytic code, and resulting output.

Once requested changes were made, the student conducted “Pilot B,” using a handful of naive participants recruited from Mechanical Turk. The goal of this pilot was to ensure that participants were able to complete the experiment and did not report any substantial issues with instructions or technical details of the experiment. (All students were instructed to give participants a way to leave free-form comments at the end of the experiment but prior to debriefing). At the conclusion of Pilot B, both the instructor and a TA reviewed the student’s analytic code and its outputs on the data for all confirmatory analyses. The goal of this code review was to ensure that all planned analyses were specified correctly and with sufficient detail to permit inclusion in the broader analysis.

After Pilot B review was completed, students were given authorization to collect the full sample of participants, contingent on having made any requested changes. Prior to data collection, the analytic script containing all confirmatory analyses was pre-registered using the Open Science Framework. In addition, prior to data collection, we pre-registered the confirmatory analyses reported in the current paper (<https://osf.io/rxz8m/>).

Statistical approach

Despite the general emphasis in meta-analytic and reproducibility efforts on aggregating a single effect size of interest (Lipsey & Wilson, 2001; Open Science Collaboration, 2015), it was often challenging for us to identify a single key statistical test for each study. Usually a substantial number of statistical tests were conducted, often within one or several regression models with multiple specifications. Thus, we interpret “key statistical test” results with caution. Despite this interpretive difficulty, following previous work, we attempted to use a single statistical test as our target for experiment planning and analysis.

Results

Pedagogical assessment

Of the 15 students in the class, all completed projects (two with extensions going beyond the class period), one chose a project outside of the sampling frame, and three opted out of the broader project prior to data collection.³ The remaining 11 students contributed code, data, registrations and materials to the final product. (In one case, a student did not complete the pre-registration procedure correctly but still submitted a pre-specified analysis plan to the instructors for review. We included this project in the final analysis.)

Each member of the instructor team coded the fidelity of replications, weighing whether the materials, sample, and task parameters differed from the original studies.

³Note that these students did not express doubts about the success of their projects prior to opting out, and 2/3 of these projects were judged by the instructor team to have been successful replications. Thus we have no evidence that these students opted out systematically due to their belief that their project would fail.

We coded projects independently on a scale of 1 – 7 (with 1 being a loose replication with substantive deviations from the original, and 7 being essentially identical), and then discussed our ratings and made adjustments (without coming to full consensus). The mean rating was 5.29 (Range = 4 to 6.67). Several studies were essentially identical to the originals, but there was a group of others that included differences in population or number of trials that might plausibly have had an effect on the results.

In two cases, in final review of the projects (after data analysis), the instructor team decided that the student’s choice of key statistical test was not in fact a strong test of the original authors’ primary theoretical claim. In both of these cases (Proudfoot et al., 2015; Sofer et al., 2015), the original authors had fit a multiple regression model to the data and there was not a single obvious test that corresponded directly to the authors’ hypothesis. After discussion, the instructor team converged on a statistical test that they thought better corresponded to the original authors’ intended hypothesis of interest. We report results from these corrected tests here. We return to this issue in the Discussion, as we believe that test selection is a critical theoretical issue in replication research (see e.g., Monin, 2016 for discussion).

Confirmatory analyses

Subjective judgments of replication. After data analysis, for each experiment, the student carrying out each experiment and the instructors all gave independent “replication ratings” to assess the subjective success of the replication. Our three-point rating system was based on the theoretical support for the original finding, “None” (no support; 0%), “Partial” (some support; 50%), and “Full” (replication consistent with original interpretation; 100%). We found that the ratings of the student carrying out the replication and the instructors were in agreement for 10 out of the 11

replications. The average rating given by instructors and students was 55% and 50% respectively.

The modal replication project in our sample was judged to be a “partial” replication, meaning that some aspects of the observed findings were different from those reported by the original authors. Since our sample only includes 11 studies, we describe general trends here rather than attempting to conduct statistical tests (which would be dramatically under-powered after correcting for multiple comparisons). Overall, projects had a slightly larger probability of being judged successful if they were judged to be closer to the original (70% vs. 42%). There were also small numerical effects of the original study being run on Mechanical Turk (60% vs. 50%) and having open materials (57% vs. 50%), though we do not believe these differences are interpretable given our sample size.

Significance of the key effect. We next turn to an assessment of the replication p -value to determine whether the key statistic found a significant effect at the traditional $p = .05$ level. Only 4 of the studies yielded a significant replication p -value. The well-documented “dance of the p -values” (Cumming, 2014) suggests that the inference drawn from a single replication p -value may not be informative, however, and other metrics such as effect sizes and Bayes factors may provide a more nuanced perspective.

Effect size and Bayes factor analysis. Effect sizes and Bayes factors improve upon inferences drawn from p -values in different ways. The effect size is an estimate of the *magnitude* of an effect, which is stable across sample sizes (assuming a constant experimental procedure, of course) and allows for a more fine-grained comparison across studies than a binary significance criterion. The Bayes Factor is an alternative hypothesis testing method that quantifies the *evidence* in favor of a hypothesis (Jeffreys,

1961). Among other advantages, it does not privilege the null hypothesis and can provide evidence either *for* or *against* (see also Scheibehenne, Jamil, & Wagenmakers, 2016; Wagenmakers, Morey, & Lee, 2016). While the Bayes Factor often agrees with the p -value as to which hypothesis is more likely (Wetzels et al., 2011), they often substantially disagree on the strength of the evidence.

These additional measures therefore give another sense in which replications can be successful: If we find roughly the same effect size or Bayes factor despite p -values on different sides of the arbitrary $p = 0.05$ threshold, we should increase our confidence in the effect’s replicability. To compute a standardized effect size, we followed Open Science Collaboration (2015) in converting test statistics to a measure of “correlation coefficient per df,” which is bounded between 0 and 1⁴. We first compared effect sizes from the original and replication studies (Figure 2A). Effect sizes were highly correlated between the two sets of studies ($r = 0.73$, $p < .0001$), but they were generally smaller in the replications (60% of the original; 95% CI [37 - 84]).

While we preregistered an analysis determining whether the original reported effect size is included in the 95% confidence interval found in the replication, we instead opted for a straightforward comparison of point estimates. This choice was motivated by recent concerns that using confidence intervals in this manner may lead to misleading interpretations (Morey, HoekstraLee, & Wagenmakers, 2016), as well as discrepancies in the authors’ reporting of effect sizes across the articles we included (which made computing confidence intervals difficult in some cases).

Next, following Etz & Vandekerckhove (2016), we compared Bayes Factors between the original and replication. We used the default test suggested by Rouder,

⁴For example, the conversion from a t statistic is given by $r = \sqrt{t^2/(t^2 + df)}$, where df is the degrees of freedom. Other formulae are given in Open Science Collaboration (2015), Supplemental Information.

Speckman, Sun, Morey, & Iverson (2009) and did not attempt to correct for publication bias.⁵ This analysis revealed a number of interesting findings. First, as with effect sizes, we saw generally smaller Bayes Factors for the replications than the originals. Second, it appeared that replication Bayes Factors generally tracked nicely with subjective replication judgments. Finally, the Bayes Factors for several of the original effects did not appear to show strong evidential value (e.g., $BF < 3$); it would have been somewhat surprising for our replications to show stronger evidence than the original.

Discussion

In this paper, we reported the results of a series of 11 student replications of previously-published experiments from the 2015 volume of *Psychological Science*. We had two interlocking goals: first, to create a proof of concept for pedagogical replications, and second, to provide evidence on the ability of students replicate important, recent findings in a top journal. Importantly, rather than attempting to assess the truth of particular effects, our project aimed to assess the challenges of replicating findings – selected on grounds of feasibility and interest – within the constraints of a course project. Despite this narrower scope, we believe that insights from our approach are applicable to the broader goal of assessing replicability at scale (cf. Open Science Collaboration, 2015)

One of the barriers to more reproducible work in psychology is pedagogical: Some of the tools (programmatic statistical analysis, version control) can feel inaccessible or intimidating to beginners. Indeed, our students, who came from a variety of disciplinary backgrounds, had relatively little experience with web experiments or literate, reproducible data analysis. Yet by the end of the course, all had gained sufficient

⁵A default BF for one paper could not be computed based on the statistical test that was used.

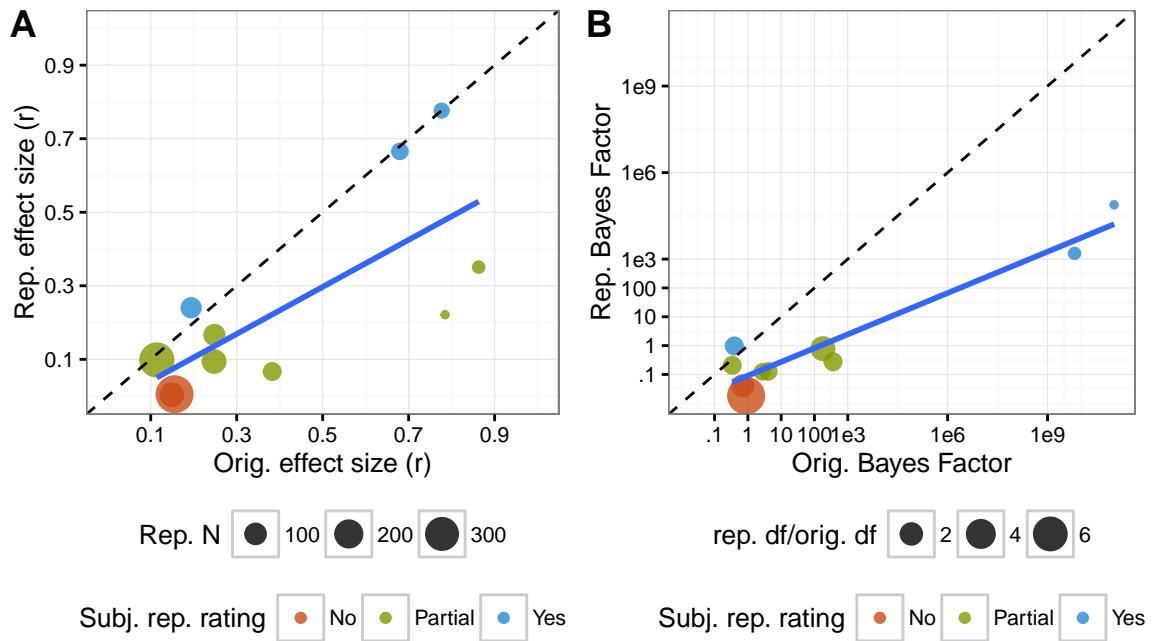


Figure 2. Replication effect size (A) and Bayes factor (B), plotted by the original effect size and Bayes factor, respectively. Point size shows replication N (A) and ratio of test degrees of freedom (B), color indicates subjective replication assessments by the authors. Note that the key statistic in the Sofer et al. (2015) replication was a multivariate F test, hence a comparable default Bayes factor could not be computed. Additionally, the original Bayes factor for Atir et al (2015) was many orders of magnitude greater than the others, so we displayed it at the edge of the range for visibility.

proficiency with the suite of technical and conceptual tools necessary to carry out a study independently. This experience suggests that it is possible to convey key concepts in sufficient depth that – within a supportive environment – students can quickly learn to use a reproducible toolset.

Our projects had a number of limitations, including at least: 1) limited time for iterative piloting and adjustment, 2) limits on funding for larger samples, and 3) limits on domain expertise with respect to the specific effects that were chosen. But within

these limitations, our group was able to produce a set of relatively close replications with generally good statistical power. Thus, our work here shows what is possible for motivated students using freely-available tools and resources. Such classes could become the backbone of future collaborative replication efforts (Everett & Earp, 2015; Frank & Saxe, 2012).

Statistically, however, our results were underwhelming. Despite our relatively close adherence to the original experimental protocols and relatively large sample sizes, the modal outcome of our projects was partial replication. These partial replications often were cases in which some hint of the original pattern was observed but the key statistical test was not statistically significant and showed both smaller effect size and lower evidential value than the same test in the original. We invite readers to browse the narrative descriptions of individual replication attempts in our SOM to see the ways that patterns of experimental findings can differ from one another beyond the significance of a single test.

These statistical findings mask a more optimistic message, however: In most of our projects, the next step for a motivated experimenter is clear. For example, in some projects we suspect that a followup could find strong evidence for the phenomenon of interest by titrating the particular planned analysis or the difficulty of the stimulus materials (Liverence & Scholl, 2015; Storm & Stone, 2015; Xu & Franconeri, 2015). In others, the difficulty appeared to be statistical power, since differences were in the predicted direction and sometimes reached significance in subsidiary analyses, so a followup would likely require a larger sample (N. A. Lewis & Oyserman, 2015; Scopelliti et al., 2015; Zaval et al., 2015). And in some, differences of population would mean that followups would require further stimulus or task adaptation (Sofer et al., 2015; Wang et al., 2015). More generally, we believe that our work here underscores the importance of

iterated replication (M. L. Lewis & Frank, in press).

Many statistical factors leading to lowered replicability have been discussed in past work, including analytic flexibility, context dependency, publication bias, and low statistical power among others (e.g., Button et al., 2013; Ioannidis, 2005; Van Bavel, Mende-Siedlecki, Brady, & Reiner, 2016). For both reasons of design and scale, our study here cannot disentangle these factors empirically. Nevertheless, the decreases in effect size and evidential value we report seem consistent with two explanations. First, we likely saw a “winner’s curse,” with initial publications tending to over-estimate effects. Second, we also see evidence that specific details of studies vary from context to context; creating successful adaptations can be challenging (Van Bavel et al., 2016). Both of these explain our observations, and we suspect both are present to some degree.

Our experiences performing and compiling the studies here sheds light on one further concern that we believe has been under-reported in past studies (including Open Science Collaboration, 2015). Our standard model of replication is based on the notion of a single statistical test – and its associated effect size – being the key properties of a study that can be targeted for replication. Yet, as we mentioned above, selecting this key statistical test was difficult for most of the projects in our sample, and virtually impossible for some. We were often forced to consult other sections of the papers (e.g., abstracts, discussion sections) to gain clarity on what test was considered the critical one for the authors’ interpretation. It is likely that some of our decisions would not be ratified by the original authors.

Indeed, by the standards of clinical trials research, nearly every study we replicated in this report would be classified as exploratory research. No protocol was pre-registered, and the general pattern of evidence across studies was often more important to the authors’ conclusions than any particular test in any single study.

Almost every study conducted multiple statistical tests, often associated with several statistical models with differing specifications. And in some cases, there was no conventionally-reported and easily calculable effect size measures (e.g., for mixed- and random-effect models, cf. Bakeman, 2005). This set of features – which, in our experience, are endemic to published psychological work, including our own – makes replication research within the conventional statistical paradigm extremely challenging. In our view, two recommendations will improve this situation dramatically. First, pre-registration typically requires the selection of particular statistical analyses to be designated as critical for interpretation. Were this information available, it would have simplified the problem of selecting effects to replicate in several cases. Second, the widespread adoption of open data practices can dramatically facilitate the type of meta-research pursued here. In addition to clarifying the (often terse) reporting in a paper, the availability of data facilitates the computation of the relevant meta-analytic variables.⁶ No paper can perfectly predict the uses to which future investigators will put a piece of work; the availability of the source materials is thus an important route for more flexible reuse.

In conclusion, our results here demonstrate the practical possibility of performing replication research in the classroom. There are many challenges for ensuring high-quality replications in a pedagogical setting – from checking experimental methods to reviewing analytic code for errors – but we would argue that these are not just pedagogical challenges. They are challenges for psychological science. We believe that the openness and transparency we pursued here as part of our pedagogical goals should

⁶We note that, although we were grateful that several papers in our sample shared data and materials, the use of open (non-proprietary) formats for archiving data – and in particular analytic code – is often critical for investigators attempting to reproduce analyses.

be models not just for other classes but for future research more broadly.

SOM-R: Project-By-Project Methods Addendum

For each project, we present a summary of the finding, the key statistical test, and the impressionistic outcomes of the replication. Figure 3 shows a side-by-side comparison of the key visualization for each study.

Storm & Stone (2015)

This paper found that when participants had the opportunity to save a file containing a list of target words before studying another file, they retained more information from the second study session. Our target for replication was Experiment 3, in which the effect was found to depend on the amount of information that was studied in the first file, with a greater effect for an eight-word list than a two-word list. The key statistical test we selected was the interaction in a 2x2 mixed-design ANOVA (save vs. no-save condition and two- vs. eight-word condition). The replication study failed to find evidence of the interaction, but ironically this failure was because of a main effect of save condition: The authors' key manipulation succeeded for both load conditions. Perhaps even two words was sufficient to create interference for our online sample.

Lewis & Oyserman (2015)

This paper investigated motivations to save for retirement, and specifically whether changing participants' relationships to the future (making it feel more imminent) would lead them to report that they would begin saving for retirement sooner. Our target was Study 4, which found that seeing the time to retirement in days (10,950) rather than years (30) caused participants to say they would begin saving

sooner, even controlling for age, education, and income in a linear model (there was also a manipulation of whether saving was incremental that did not result in an effect). We failed to find a significant relationship between the time metric manipulation and saving time when including demographic controls (the key test we selected). However, when we did not include demographic controls, we did find a marginally-significant relationship. In exploratory analyses, we also observed an unpredicted effect of income such that participants who reported higher income selected to start saving sooner.

Scopelliti, Lowenstein & Vosgerau (2015)

The goal of this paper was to examine whether people are accurately calibrated in their estimates of how others will respond to their attempts at self-promotion. Our target was Experiment 3, which showed that participants, who were instructed to describe themselves in such a way that others would like them, expected to be liked more – but were actually liked less – than participants who were asked to describe themselves but not additionally instructed to maximize others’ interest in meeting them. The authors performed a series of regressions on four dependent variables capturing judgments about profile writers (interest, liking, bragging, and success), analyzing these as a function of condition (control vs. “maximize interest”) and who the evaluator was (the writer vs. an independent sample). Power analysis for this design was complex because the procedure had two stages: In a first stage participants generated descriptions and predicted judgments, while in a second stage a separate sample rated the descriptions.

We were faced with a choice: Either we could have re-rated the descriptions gathered in the original study (by contacting the author and treating these as “materials” for the study), effectively replicating only stage two, or we could redo the

entire two-step procedure. We elected to conduct a replication of the full procedure, but in our replication to decrease cost we focused on the liking judgments, which appeared to be central to the authors' conclusions. Our key test was the interaction of rater (writer vs. independent) and condition.⁷ However, to power even this smaller study adequately was outside of our budget, so instead of having each judge rate a subset of 10 out of 100 total profiles, we collected a smaller sample of profiles (18) and had every profile rated by each judge. We failed to find the predicted interaction in this new sample. Consistent with the original report, though, we found a main effect such that stage one participants predicted that raters would like them more than the stage two participants actually did. Despite our relatively high statistical power in the main test, however, we may have suffered from low power at the item level due to the relatively small number of profiles we collected in the replication study.

Liverence & Scholl (2015)

This paper tested the theory that persistent object representations could assist in spatial navigation, using a navigation paradigm in which participants used key presses to move through a grid of pictures that were visible only one at a time. In Experiment 1, our target, found that participants located targets faster when navigation involved persistence cues (via sliding animations) than when persistence was disrupted (via temporally matched fading animations). The original authors did not report training effects either within or across testing epochs (groups of 50 trials), so in our adaptation we decreased the length of the paradigm from four to two epochs.

The key test for our replication was a simple *t*-test comparing speed to find the

⁷We note that this project was submitted late and the student failed to preregister this analysis formally; however, the pre-written analysis protocol went through pre-data collection review.

target in the two conditions (slide vs. fade), not controlling for learning across the duration of the experiment. We failed to find this overall difference, but, in an exploratory followup analysis using a linear mixed effect model, we did find a highly-significant effect of condition when controlling for learning epoch. Our data may have been more variable for a number of reasons: unlike in the original paradigm, our participants were not required to click when they found targets; our stimuli differed slightly (images were different and featured thin black borders); and our participants differed in age and educational status from the undergraduates in the original study.

Wang et al. (2015)

This paper reported that the relationship between math anxiety and math performance follows an inverted-U pattern among college undergraduate students with higher intrinsic math motivation, whereas this same relationship is linearly negative among students with lower intrinsic math motivation. We targeted Study 2, the authors' own replication study, in which participants filled out three surveys about their math motivation and anxiety and completed a math task to measure performance. The key effect we targeted was the interaction of math anxiety squared and math motivation in predicting performance. We failed to find evidence for this effect. In contrast to the original study, we found a simple linear trend such that performance decreased with increasing math anxiety for both high and low math-motivation groups. Population differences provide one plausible explanation for the differences we found: Math performance was overall lower in our AMT sample than in the original undergraduate sample. It is possible that either the inverted U was undetectable given this lower level of performance, or that levels of motivation differed enough that the posited relationship did not hold.

Sofer et al. (2015)

This paper tested the hypothesis that face typicality affects trustworthiness judgments such that the most typical face will be judged to be the most trustworthy. In Study 1, our target, faces of varying typicality were rated on their trustworthiness and attractiveness; while trustworthiness peaked at the most typical face, less typical faces were judged more attractive. The original sample was Israeli women, and typical faces were constructed to be typical for that sample. Rather than attempting to construct a typical face for AMT workers, we instead used the stimuli provided by the original authors, recognizing that this decision limits the interpretation of our findings.

While the original class projects selected a global, item-wise regression model as the key statistical test of interest, the instructor team (post-data collection) determined that the fit of this model did not in fact correspond most closely with the authors' stated hypothesis. We thus selected the interaction between face typicality and rating condition as our key test statistic and used the more standard by-participant regression model (which the original paper also reported subsequent to the by-items model) as the target model. Our replication study successfully showed this same interaction. We note that, despite the significant interaction, in our study trustworthiness judgments did not change substantially with typicality (though attractiveness did) and trustworthiness did not peak numerically at the most typical face. This result was plausibly caused by the difference in the typicality of the faces we used given our AMT participants.

Ko, Sadler, & Galinsky (2015)

This paper investigated vocal cues to social hierarchy. Our target for replication was Experiment 2, which assessed whether participants used particular acoustic cues to make inferences about speakers' level of social hierarchy, particularly that participants

from the high-rank condition in Experiment 1 were rated as more highly ranking. This paper employed a two-stage procedure in which participant-generated materials were rated by an independent sample. However, in contrast to the study of Scopelliti, Lowenstein & Vosgerau (2015), because materials were collected in Experiment 1 and were available openly, we elected to replicate only Experiment 2, the judgment study. The key test statistic was the main effect of hierarchy condition on “behavior score” (an index of whether speakers would engage in high-status behaviors). The authors were also interested in which particular vocal cues predicted participants judgments, but we judged these analyses to be more descriptive and exploratory. We successfully replicated the main effect of condition on hierarchy judgment, and additionally found a comparable main effect of speaker gender.

Atir, Rosenzweig, & Dunning (2015)

This paper investigated how perceived knowledge affects “overclaiming” (claiming knowledge that you do not have). We targeted Study 1b, which asked participants to rate their knowledge about finances, and then asked them to rate their knowledge about a set of financial terms, some of which were invented (and hence for which a positive knowledge statement would necessarily be an overclaim). The original study found that participants with higher self-rated financial knowledge also overclaimed at a higher rate. The key test statistic for this study was the coefficient in a regression model predicting overclaiming as a function of claimed personal financial knowledge, controlling for accuracy in the true financial knowledge questions. We successfully replicated this effect, despite our use of a much smaller sample (due to the extremely high post-hoc power of the original study).

Proudfoot, Kay, & Koval (2015)

This paper examined whether judgments of creativity are influenced by the gender of the creator. In Study 1, they tested the hypothesis that “out of the box” creativity was more associated with masculine traits than feminine traits by manipulating the definition of creativity that participants saw (convergent vs. divergent) and asking participants to judge the centrality of different traits to that definition. The original pre-registration selected a subsidiary analysis of a main effect as the key test statistic, but in post-data collection review the instructor team decided that a test of the interaction between trait gender and definition condition was closer to the central theoretical claim of the paper; we use this latter test as the key test statistic. Our replication study found an significant interaction of similar magnitude.

Zaval, Markowitz, & Weber (2015)

This paper investigated whether encouraging participants to think about their personal legacy could increase concern for the environment. The main study of the paper investigated whether a legacy priming writing exercise would increase a variety of environmental measures, including donations to an environmental charity, pro-environmental intentions, and climate change beliefs. The key test we selected was the effect of condition (priming vs. control) on behavioral intentions in an analysis of variance. We observed a marginally significant effect in the same direction in our replication, as well as support for the mediating relationship of legacy motives on behavioral intentions. Our study may have been under-powered due to budgetary constraints; we ran a sample of comparable size to the original sample of 312 participants. We also may have observed a smaller effect size due to the relatively smaller amount of time our participants spent on the legacy prime writing exercise. We

enforced a four minute writing time and a 20 word minimum; we found that most participants stayed relatively closer to these limits than participants in the original study.

Xu & Franconeri (2015)

This paper explored constraints on visual working memory in the context of mental rotation. In Experiment 1a, participants performed a verbal suppression task to control for use of verbal encoding while attempting to remember a cross with four colored bars, which sometimes switched colors; performing mental rotation significantly impaired ability to detect these bar swaps. To port the task to an online framework we decreased the number of trials and increased the number of participants we tested. The key test statistic was a comparison of K (a metric of the number of visual features remembered) between the rotation and no-rotation conditions. We observed a marginally-significant effect in our replication study. Our study showed a floor effect on capacity such that participants remembered on average very little; it is possible that small details in the displays (or the differing sample available online) led to this floor effect.

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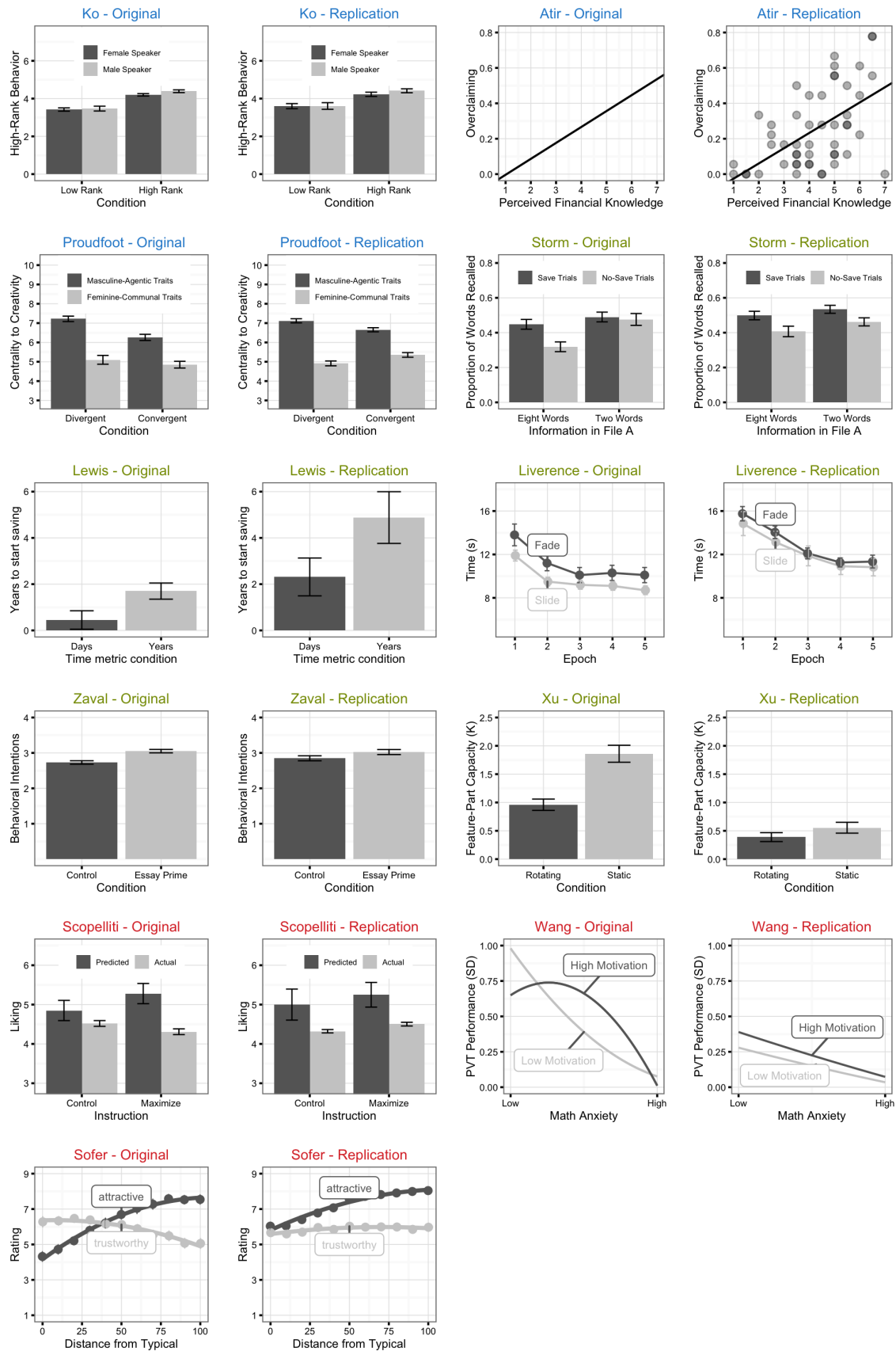


Figure 3. Side-by-side plots for each attempted replication. Error bars show standard error of the mean. Original data estimated from figures when not otherwise available. Replications are sorted by subjective replication success.