More Bayesian Inference and STAN

Key Concepts

- Visual Overview of Bayesian Inference
- Writing Stan Code

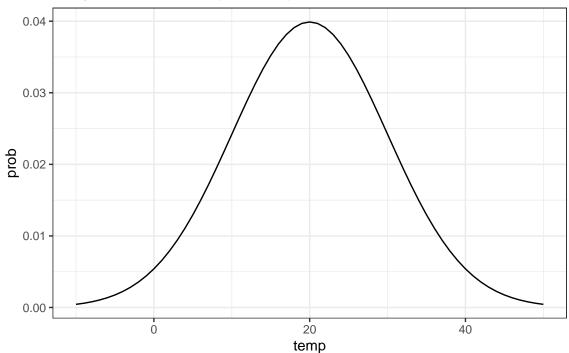
Visual Overview of Bayesian Inference

Using some Bridger Bowl weather data we will provide a visual overview of Bayesian Inference. The goal will be to model the average winter high temperature at the base of Bridger Bowl.

- 1. Prior Specification
- First sketch a prior distribution that encapsulates your belief about what you believe the average high temperature would be. Note this should obey law of total probability
- Next we generally need to parameterize (perhaps approximately) this belief with some sort of probability distribution.

```
temp_seq <- -10:50
prob_seq <- dnorm(temp_seq, mean = 20, sd = 10)
tibble(temp = temp_seq, prob = prob_seq) %>% ggplot(aes(temp, prob)) + geom_line() + theme_bw() +
    ggtitle("Andy's Prior Belief: N(20, 10^2)")
```

Andy's Prior Belief: N(20, 10^2)



Formally, my prior is on the mean high temp, which we will denote μ .

$$\mu \sim N(20, 10^2)$$

- 2. Specify the sampling distribution for the data or perhaps in more familiar language, state the likelihood for the statistical model
- We will assume that the temperature readings are continuous (or "nearly continuous")
- It seems reasonable to start with a normal distribution, so:

$$X|\mu,\sigma^2 \sim N(\mu,\sigma^2)$$

- Note that we also need to estimate σ in this model and need a prior for that parameter too.
- Grab some weather data from Bridger Bowl (roughly the first half of January 2021)

• Any concerns about using this data to inform our research question?

3. Posterior Inference

• Using classical inference, how would you estimate μ .

• Using maximum likelihood, $\hat{\mu}_{MLE} = \bar{X} = 33$.

• With Bayesian inference, our posterior belief is based on the data and our prior belief. Note this can be a blessing or a curse.

• Formally, we have a distribution for the maximum temperature (a posterior distribution):

$$p(\mu|x) = \int \frac{p(x|\mu,\sigma) \times p(\mu)p(\sigma)}{p(\mu)} d\sigma$$

, note solving this is not trivial and isn't something we will handle in this class.

• Luckily, there is an elegant computational procedure that will allow us to approximate $p(\mu|x)$ by taking samples from the distribution. This is, of course, MCMC.

STAN code for this situation can be written as below. Note that the prior values are hard coded, these could also be passed in as arguments to the model.

```
data {
  int<lower=0> N;
  vector[N] y;
}

parameters {
  real mu;
  real<lower=0> sigma;
}

model {
  y ~ normal(mu, sigma);
  mu ~ normal(20, 10);
}
```

```
print(temp_data)
```

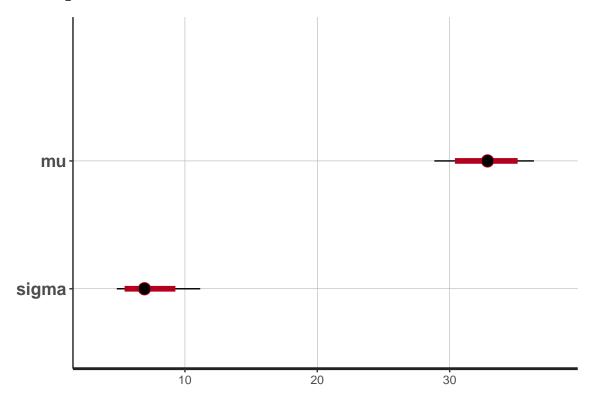
```
## Inference for Stan model: normal.
## 4 chains, each with iter=2000; warmup=1000; thin=1;
## post-warmup draws per chain=1000, total post-warmup draws=4000.
##
##
                               2.5%
                                       25%
                                                      75%
           mean se_mean
                          sd
                                              50%
                                                          97.5% n_eff Rhat
## mu
          32.81
                   0.04 1.90
                              28.85
                                     31.63
                                            32.86
                                                   34.06
                                                          36.37
                                                                  2567
          7.22
                               4.85
                                     6.08
                                             6.94
                                                                          1
## sigma
                   0.04 1.64
                                                    8.07
                                                          11.15
                                                                  2083
## lp__
        -32.82
                   0.03 1.04 -35.61 -33.21 -32.50 -32.08 -31.81
                                                                          1
##
## Samples were drawn using NUTS(diag_e) at Fri Jan 22 11:39:24 2021.
## For each parameter, n_eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor on split chains (at
## convergence, Rhat=1).
```

temp_data <- stan("normal.stan", data=list(N = length(temp), y=temp))</pre>

plot(temp_data)

ci_level: 0.8 (80% intervals)

outer_level: 0.95 (95% intervals)



We can also view the posterior and prior beliefs together on a single figure.

Prior and posterior belief for winter temperature in Bozeman

