

# Lecture 8: Point Level Models - Model Fitting

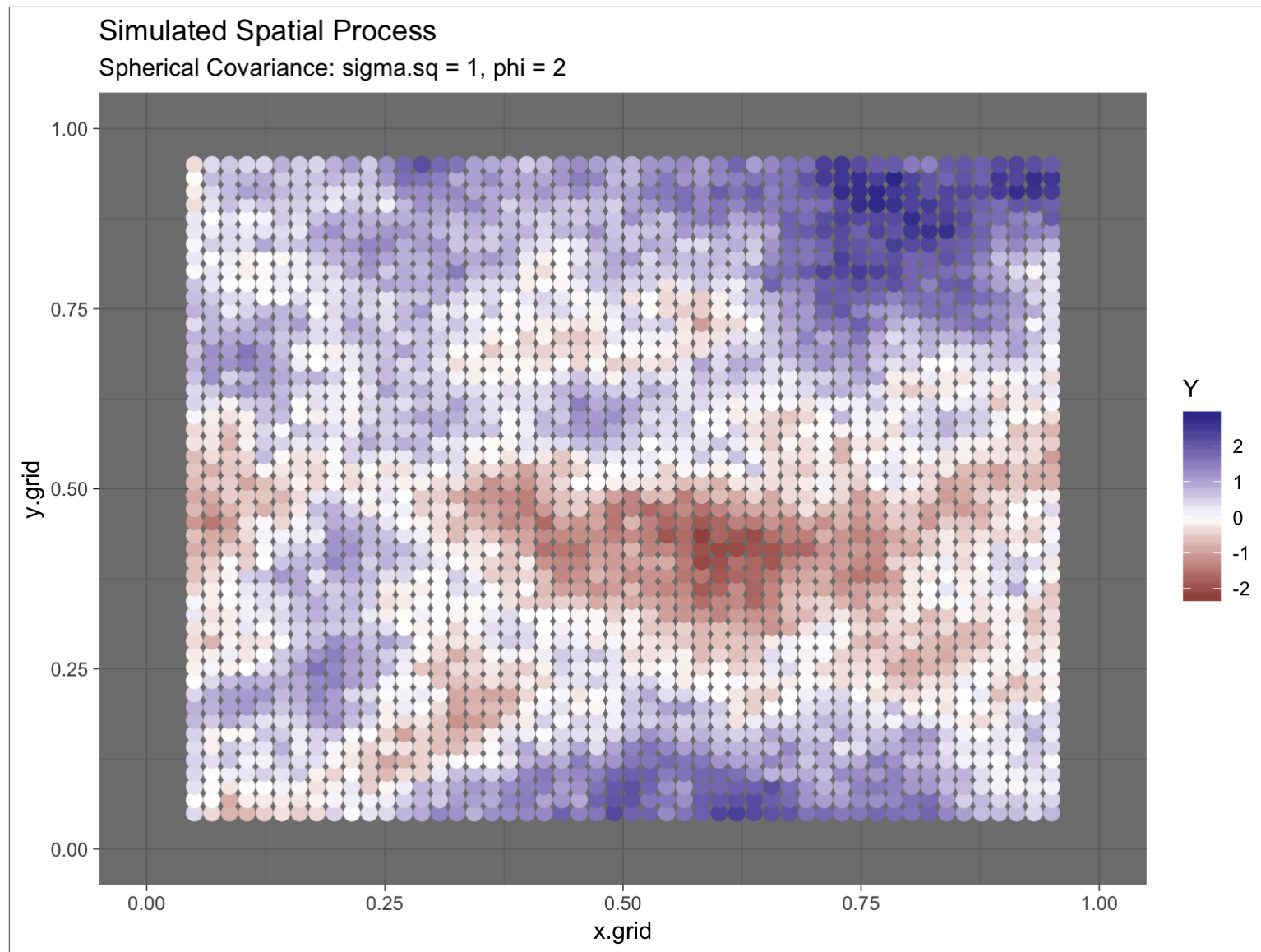
# Class Intro

# Intro Questions

- Why are we creating simulated spatial processes?
- For Today:
  - Model Fitting

# Model Simulation

# Simulating Spatial Process





## Simulated Spatial Process: Exercise

How does the spatial process change with:

- another draw with same parameters?
- a different value of  $\phi$
- a different value of  $\sigma^2$
- adding a nugget term,  $\tau^2$

# Model Fitting



## Classical Model Fitting

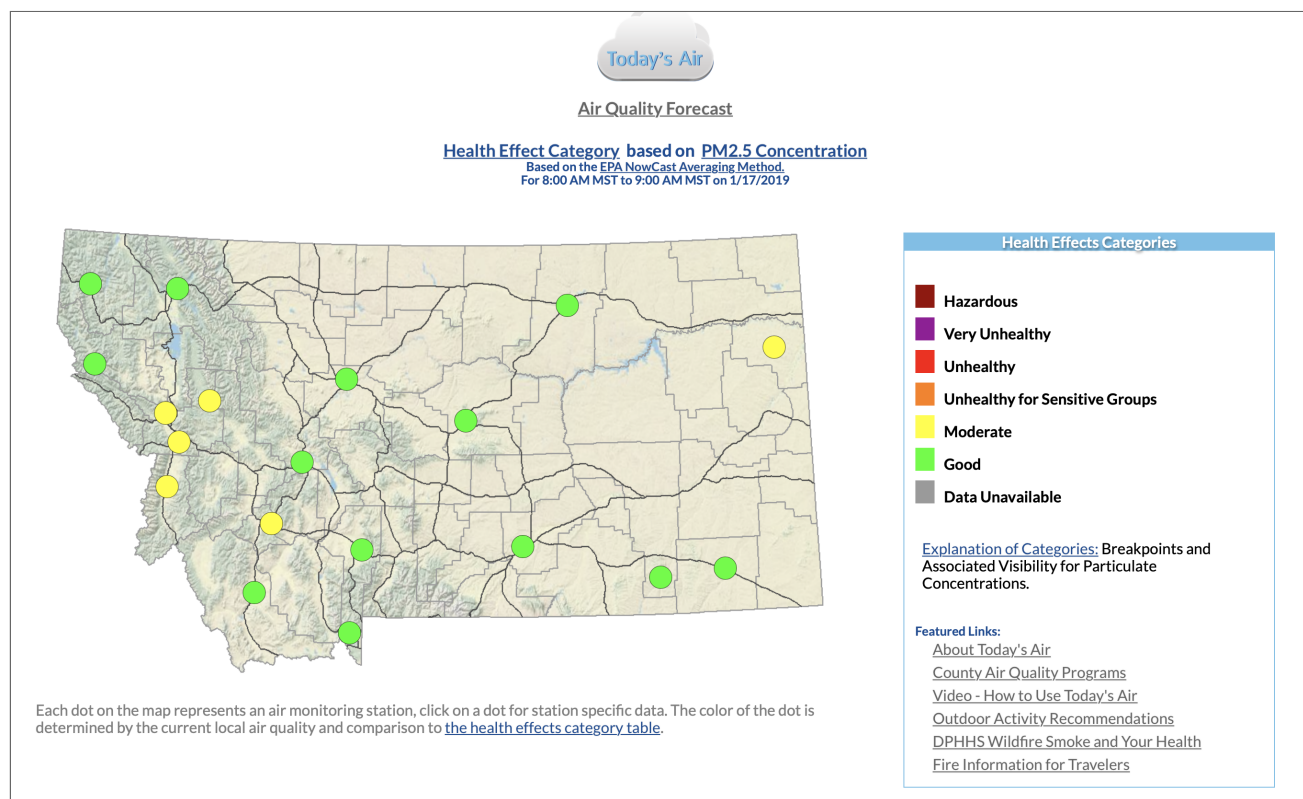
- The classical approach to spatial prediction is rooted in minimizing the mean-squared error.
- This approach is often referred to as *Kriging* in honor of D.G. Krige a South African mining engineer.
- As a result of Krige's work (along with others), point-level spatial analysis and geostatistical analysis are used interchangeably.

## Mathematical Motivation

- Let  $Y = \{Y(s_1), \dots, Y(s_n)\}$  be observations of a spatial process and  $n$  sites.
- Then  $Y(s_0)$  is a site where the spatial process has not been observed.
- **The Goal:** What is the best predictor for  $Y(s_0)$  given that  $Y = \{Y(s_1), \dots, Y(s_n)\}$  was observed?

# Visual Motivation

- **The Goal:** What is the best predictor for  $Y(s_0)$  given that  $Y = \{Y(s_1), \dots, Y(s_n)\}$  was observed?



source: [airnow.gov](http://airnow.gov)

## Mathematical Notation

- A linear predictor for  $Y(\mathbf{s}_0)$ , given  $\mathbf{Y}$  takes the form  $\sum_i l_i Y(\mathbf{s}_i) + \delta_0$
- Using squared error loss, we'd seek to minimize

$$E \left[ Y(\mathbf{s}_0) - \left( \sum_i l_i Y(\mathbf{s}_i) + \delta_0 \right) \right]^2$$

as a function of  $l_i$  and  $\delta_0$ .

- Describe and interpret  $l_i$  and  $\delta_0$

## Connection to variogram

- Recall the intrinsic stationarity assumption

$$E[Y(\mathbf{s} + \mathbf{h}) - Y(\mathbf{s})] = 0,$$

thus  $\sum_i l_i = 1$  such that

$$E[Y(\mathbf{s}_0) - \sum_i l_i Y(\mathbf{s}_i)] = 0$$

- Following this logic, we would now minimize

$$E[Y(\mathbf{s}_0) - \sum_i l_i Y(\mathbf{s}_i)]^2 + \delta_0^2,$$

thus  $\delta_0 = 0$ .

## Connection to variogram: part 2

- Define  $a_0 = 1$  and  $a_i = -l_i$ , then we can rewrite

$$E[Y(s_0) - \sum_i l_i Y(s_i)]^2 \quad \text{as} \quad E\left[\sum_{i=0}^n a_i Y(s_i)\right]^2$$

- It turns out that

$$E\left[\sum_{i=0}^n a_i Y(s_i)\right]^2 = - \sum_i \sum_j a_i a_j \gamma(s_i - s_j)$$

- In other words, minimizing the squared error, under assumptions, justifies the variogram.
- This is a constrained optimization of a quadratic form that is typically handled with a Lagrange multiplier. **Khan Academy Refresher Video**

## Lagrange multipliers

To rewrite the constrained optimization in terms of  $l_i$  we get

$$-\sum_{i=0}^n \sum_{j=0}^n a_i a_j \gamma(s_i - s_j) = -\sum_{i=1}^n \sum_{j=1}^n l_i l_j \gamma_{ij} + 2 \sum_{i=1}^n l_i \gamma_{0i},$$

where  $\gamma_{ij} = \gamma(s_i - s_j)$  and hence  $\gamma_{0j} = \gamma(s_0 - s_j)$

- Solving equations with this constraint requires partial derivatives and the use of Lagrange multipliers, typically denoted as  $\lambda$ .

## BLUP

- It turns out that the solution for the vector  $\mathbf{l}$  is

$$\mathbf{l} = \Gamma^{-1} \left( \boldsymbol{\gamma}_0 + \frac{(1 - \mathbf{1}^T \Gamma^{-1} \boldsymbol{\gamma}_0)}{\mathbf{1}^T \Gamma^{-1} \mathbf{1}} \mathbf{1} \right),$$

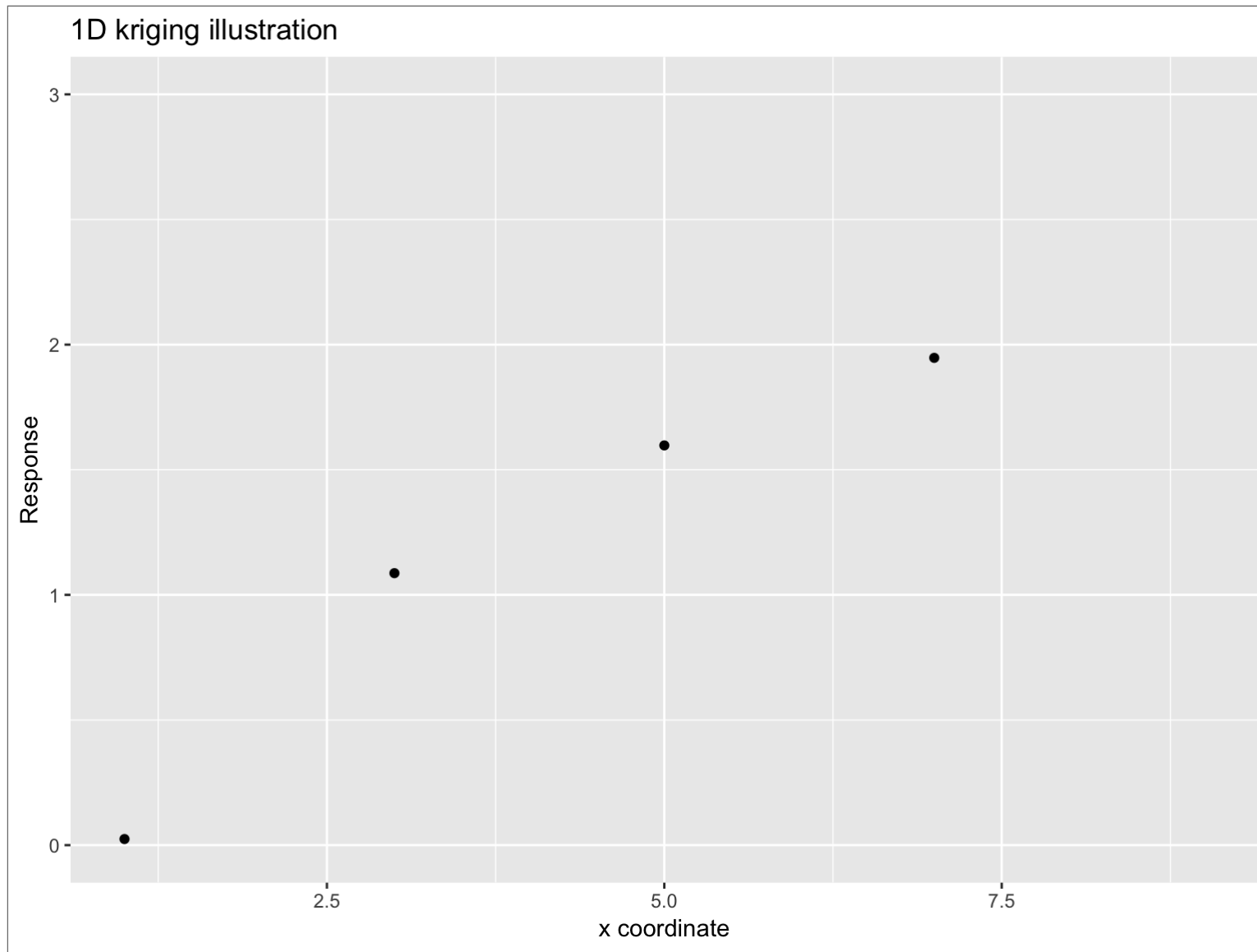
where  $\Gamma$  is an  $n \times n$  matrix with entries  $\Gamma_{ij} = \gamma_{ij}$  and  $\boldsymbol{\gamma}_0$  is the vector of  $\gamma_{0i}$  values.

- Then the Best Linear Unbiased Predictor is  $\mathbf{l}^T \mathbf{Y}$
- This BLUP also requires an estimate of  $\gamma(\mathbf{h})$



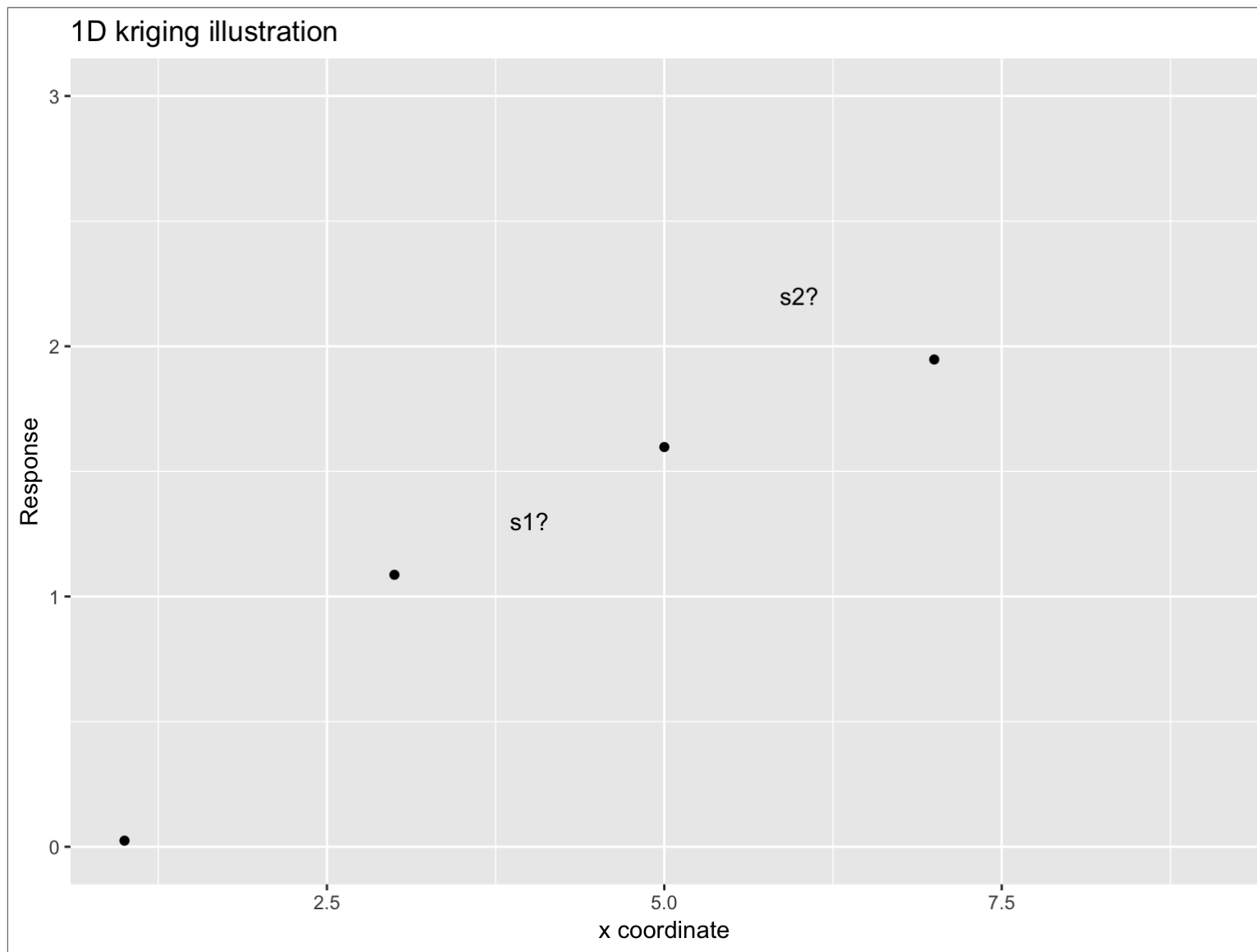
# So what does this all mean ...

Consider a small example on 1-dimension.



So what does this all mean ...

What should the predictions be at  $s_1^* = 4$  and  $s_2^* = 6$



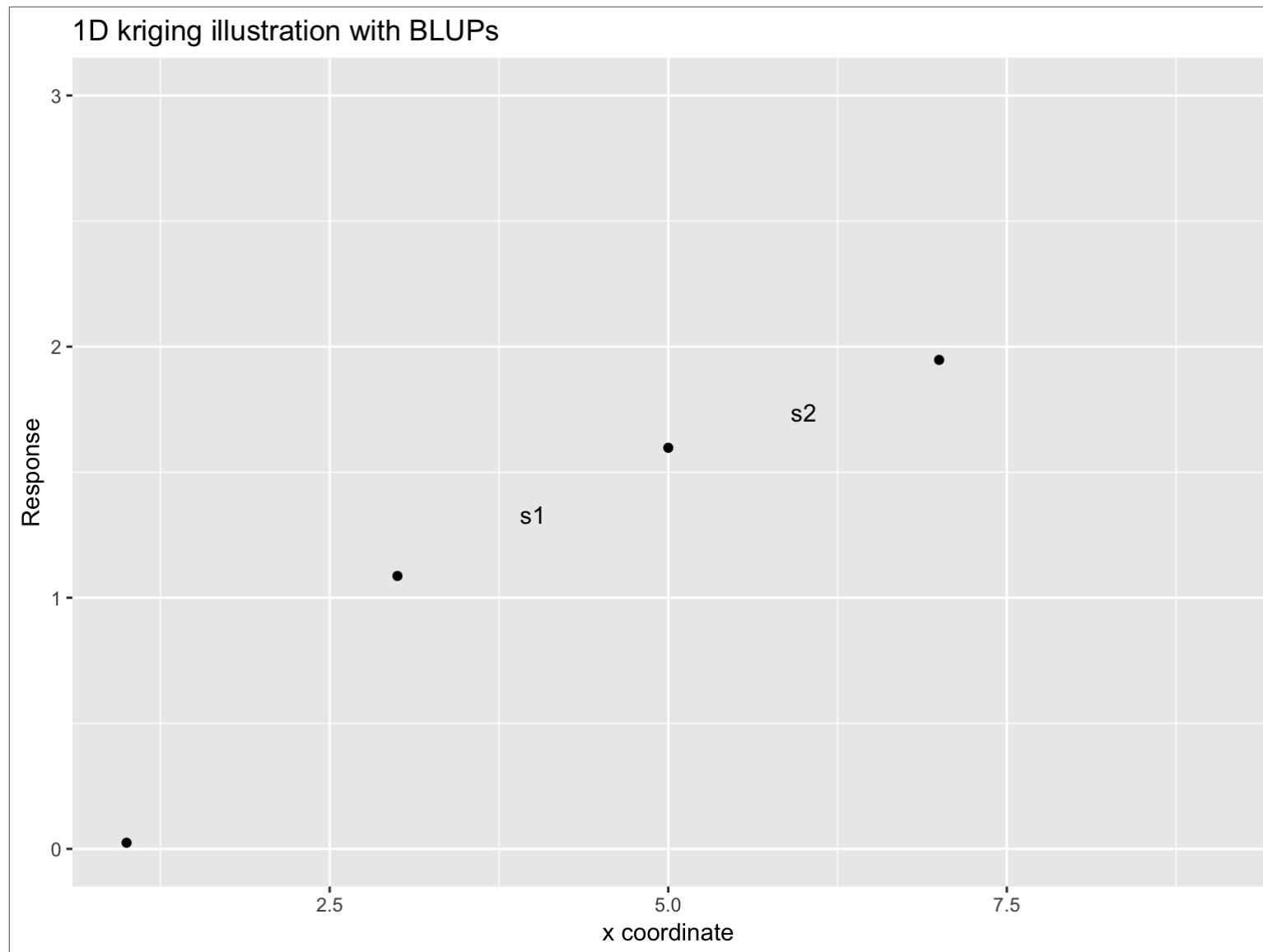
## Kriging Exercise:

- Recall

$$\boldsymbol{l} = \Gamma^{-1} \left( \boldsymbol{\gamma}_0 + \frac{(1 - \mathbf{1}^T \Gamma^{-1} \boldsymbol{\gamma}_0)}{\mathbf{1}^T \Gamma^{-1} \mathbf{1}} \mathbf{1} \right),$$

- Define  $\gamma(h) = 1 - \exp(-\frac{h}{3})$  and compute the BLUPs for  $s_1^*$  and  $s_2^*$
- Interpret and explain  $\boldsymbol{l}$  for each sample point.
- If you have time, fill in the line (rather than the surface) from (0.5, 7.5)

# Kriging Solution





# Kriging with Gaussian Processes



## A Gaussian Process

- The BLUP does not contain a distributional assumptions, but rather comes from an optimization framework.
- Now assume that
$$Y = \mu \mathbf{1} + \epsilon, \quad \text{where } \epsilon \sim N(\mathbf{0}, \Sigma)$$
- With no nugget, let  $\Sigma = \sigma^2 H(\phi)$ , where  $(H(\phi))_{ij} = \rho(\phi; d_{ij})$ , where  $d_{ij}$  is the distance between  $s_i$  and  $s_j$ .
- A nugget can be included by modifying  $\Sigma$  to be  $\Sigma = \sigma^2 H(\phi) + \tau^2 I$

# Minimizing Mean-Square Prediction Error

- **Goal:** find  $h(\mathbf{y})$  that minimizes  $E[(Y(\mathbf{s}_0) - h(\mathbf{y}))^2 | \mathbf{y}]$
- $E[(Y(\mathbf{s}_0) - h(\mathbf{y}))^2 | \mathbf{y}]$
- $= E[(Y(\mathbf{s}_0) - h(\mathbf{y}) \pm E[(Y(\mathbf{s}_0) | \mathbf{y})])^2 | \mathbf{y}]$
- $= E\{(Y(\mathbf{s}_0) - E[(Y(\mathbf{s}_0) | \mathbf{y})])^2 | \mathbf{y}\} + \{E[(Y(\mathbf{s}_0) | \mathbf{y})] - h(\mathbf{y})\}^2$

## Minimizing Mean-Square Prediction Error: Part 2

- As  $\{E[(Y(s_0)|y) - h(y)]^2 \geq 0$
- we have
$$E[(Y(s_0) - h(y))^2 | y] \geq E\{(Y(s_0) - E[(Y(s_0)|y)])^2 | y\}$$
- Hence to minimize  $E[(Y(s_0) - h(y))^2 | y]$ , we set ...
- $h(y) = E[(Y(s_0)|y)]$
- Hence,  $h(y)$  that minimizes the error is the conditional expectation of  $Y(s_0)$
- Note this is also the *posterior mean* of  $Y(s_0)$

## Multivariate Normal Theory

- For consider partitioning a multivariate normal distribution into two parts

$$\begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} = N \left( \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \Omega_{11} & \Omega_{12} \\ \Omega_{21} & \Omega_{22} \end{pmatrix} \right),$$

where  $\Omega_{12} = \Omega_{21}^T$

## Conditional Multivariate Normal Theory

- The conditional distribution,  $p(Y_1|Y_2)$  is normal with:
- $E[Y_1|Y_2] = \mu_1 + \Omega_{12}\Omega_{22}^{-1}(Y_2 - \mu_2)$
- $Var[Y_1|Y_2] = \Omega_{11} - \Omega_{12}\Omega_{22}^{-1}\Omega_{21}$
- Thus with  $Y_1 = Y(s_0)$  and  $Y_2 = y \setminus$   
 $\Omega_{11} = \sigma^2 + \tau^2$ ,  $\Omega_{12} = (\sigma^2 p(\phi; d_{01})), \dots, p(\phi; d_{0n}))$ ,  $\Sigma_{22} = \sigma^2 H(\phi) + \tau^2$

## Universal Kriging

- When covariate information is available for inclusion in the analysis, this is often referred to as *universal kriging*
- Now we have
$$Y = X\boldsymbol{\beta} + \boldsymbol{\epsilon}, \quad \text{where } \boldsymbol{\epsilon} \sim N(\mathbf{0}, \Sigma)$$
- The conditional distributions are very similar to what we have derived above, watch for HW question.
- In each case, kriging or universal kriging, it is still necessary to estimate the following parameters:  $\sigma^2$ ,  $\tau^2$ ,  $\phi$ , and  $\mu$  or  $\beta$ .
- This can be done with least-squares methods or in a Bayesian framework.

# Gaussian Process Exercise:

# Overview

- Similar to the previous exercise, we will simulate data from a 1D process and make predictions at unobserved locations.
- In this situation, please plot the mean of the distribution as well as some uncertainty metric.
- You do not need to estimate  $\sigma^2$ ,  $\tau^2$ , and  $\phi$  but can use the known values in the R code.



# Data Sampling

