

# Modeling Point Patterns using NHPPs

Statistical modeling depends on a sampling model for the data and the associated likelihood function.

Conditional on the number of points,  $Num(D) = n$ , the location density can be specified as:

$$f(\mathbf{s}_1, \dots, \mathbf{s}_n | Num(D) = n) = \prod_i \frac{\lambda(\mathbf{s}_i)}{(\lambda(D))^n}$$

Then considering the location and number of points simultaneously, *we have*

$$f(\mathbf{s}_1, \dots, \mathbf{s}_n, Num(D) = n) = \prod_i \frac{\lambda(\mathbf{s}_i)}{(\lambda(D))^n} \times (\lambda(D))^n \frac{\exp(-\lambda(D))}{n!}$$

Alternatively, consider a fine partition for  $D$ , then using the Poisson assumption, the likelihood is the product of the counts across the partitions.

$$\prod_l \exp(-\lambda(A_l)) (\lambda(A_l))^{Num(A_l)}$$

As the grid becomes finer,  $Num(A_l) \in \{0, 1\}$ .

The likelihood is a function of the entire intensity surface  $\lambda(\mathbf{s})$ . Hence, a functional description of  $\lambda(\mathbf{s})$  is necessary.

There are two general approaches for modeling  $\lambda(\mathbf{s})$ : parametric and non-parametric.

A simple, but “silly” example of a parametric form would be  $\lambda(\mathbf{s}) = \sigma \lambda_0(\mathbf{s})$ . *Realistically,  $\lambda_0(\mathbf{s})$  would not be available.*

Suppose remote-sensed covariate information is available on a grid. How might this be used for constructing  $\lambda(\mathbf{s})$ ? *Markov random field properties (Areal data) could be used.*

In general a parametric function  $\lambda(\mathbf{s}; \theta)$  could be specified, but it would require a richly specified class that would need to be non-negative. A common approach for this type of problem is to use a set of basis functions such that  $\lambda(\mathbf{s}; \theta) = \sum_k a_k g_k(\mathbf{s})$ .

- Sketch a set of basis functions to create a multimodal curve in 1D.

*Sometimes a trend surface (as a function of latitude and longitude) can be used to specify  $\lambda(\mathbf{s})$*

The most common approach is to specify  $\log(\lambda(\mathbf{s})) = X^t(\mathbf{s})\gamma$ . *In otherwords, spatial covariates are used in a regression model for the intensity surface.*

Specifying covariates still requires integrating  $\int_D \exp(X^t(\mathbf{s})\gamma)d\mathbf{s}$ , which is challenging as  $X^t(\mathbf{s})$  is not often specified in a functional form.

Now, considering  $\lambda(\mathbf{s})$  through the non-parametric lens, *let*

$$\lambda(\mathbf{s}) = g(X^t(\mathbf{s})\gamma)\lambda_0(\mathbf{s})$$

*where  $g() \geq 0$  and  $\lambda_0(\mathbf{s})$  as the error process (with mean 1). This is referred to as a Cox process.*

Suppose  $\lambda(\mathbf{s}) = \exp(Z(\mathbf{s}))$  where  $Z(\mathbf{s})$  is a realization from a spatial Gaussian process with mean  $X^t(\mathbf{s})\gamma$  and covariance function  $\sigma^s\rho()$ . *Then this is referred to as a Log Gaussian Cox Process (LGCP).*

The LGCP provides a prior for  $\lambda(\mathbf{s})$ . This can be written as  $X^t(\mathbf{s})\gamma + w(\mathbf{s})$  where  $w(\mathbf{s}) = \log \lambda_0(\mathbf{s})$ .

In the parametric setting, *inference proceeds using  $\mathcal{L}(\theta|\mathbf{s}_\infty, \dots, \mathbf{s}_\setminus)$  in the normal manner.*

In the non-parametric setting,  $\lambda(\mathbf{s}) = \exp(Z(\mathbf{s}))$ , where  $Z(\mathbf{s}) = X^t(\mathbf{s})\gamma + w(\mathbf{s})$  is a realization of a GP (GP prior on  $\log \lambda_D$ )

*Thus priors are needed for the covariance function of  $Z(\mathbf{s})$ , likely  $\sigma^2$  and  $\phi$ .*

In general algorithms for this type of data are computationally intensive.

Due to the challenges in modeling LGCP, the data is sometimes considered on a grid. Then the intensity for grid  $i$  could be estimated as:

$$\log \lambda_i = X_i^t \beta + \phi_i.$$

However, the results can be sensitive to the grid specification.