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### Graded Homework 1 - Exercise 3

1.  $\gamma = E[R_1^3]$ . Based on the moment generating function, we could get:

$$\phi'''(0) = E[R_1^3]$$

Since  $R_1 \sim N(\mu, \sigma^2)$ , we could get the MGF:

$$\phi(t) = e^{t\mu + 0.5\sigma^2 t^2}$$

So,

$$\phi'''(t) = (2\sigma^4 t + 2\mu\sigma^2)e^{t\mu + 0.5\sigma^2 t^2} + (\mu^2 + \sigma^4 t^2 + 2\mu\sigma^2 t + \sigma^2)(\mu + \sigma^2 t)e^{t\mu + 0.5\sigma^2 t^2}$$

Plug in  $t = 0$  we could get that,

$$\phi'''(0) = \mu^3 + 3\mu\sigma^2$$

So,  $\gamma = E[R_1^3] = \mu^3 + 3\mu\sigma^2$ .

2. (a) By CLT, we have

$$\sqrt{n}(\bar{R} - \mu) \xrightarrow[n \rightarrow \infty]{D} N(0, \sigma^2)$$

By  $\delta$ -method, we have

$$\sqrt{n}((\bar{R})^3 - \mu^3) \xrightarrow[n \rightarrow \infty]{D} N(0, \sigma^2 \cdot (3\mu^2)^2) = N(0, 9\mu^4\sigma^2)$$

Therefore,  $E[\hat{\gamma}] = E[(\bar{R})^3] = \mu^3$ ,

$$\text{bias}(\hat{\gamma}) = E[\hat{\gamma}] - \gamma = \mu^3 - \mu^3 - 3\mu\sigma^2 = -3\mu\sigma^2.$$

- (b)  $\hat{\gamma}$  is not consistent.

Since  $\hat{\gamma} = (\bar{R})^3$ , as  $n \rightarrow \infty$ , due to WLLN and Continuous Mapping Theorem,  $\hat{\gamma} \rightarrow \mu^3$ .

While  $\gamma = \mu^3 + 3\mu\sigma^2$ . So, this estimator is not consistent.

3. We claim that  $\hat{\gamma}' = (\bar{R})^3 + 3\bar{R}S^2$  is an unbiased estimator, where  $S = \frac{1}{n-1} \sum (R_i - \bar{R})^2$  and  $E[S^2] = \sigma^2$ .

Since  $R_i$  is *i.i.d.* normal,  $\bar{R}$  and  $S^2$  are independent. That is,  $E[\bar{R}S^2] = E[\bar{R}] \cdot E[S^2] = \mu\sigma^2$ .

Therefore,  $E[(\bar{R})^3 + 3\bar{R}S^2] = E[(\bar{R})^3] + 3E[\bar{R}] \cdot E[S^2] = \mu^3 + 3\mu\sigma^2 = \gamma$ .

Thus we have proven that  $\hat{\gamma}'$  is unbiased.

If we want to estimate  $\mu^3$ , we can use  $(\frac{1}{n} \sum_{i=1}^n R_i)^3$ . Actually, this estimator is just  $\hat{\gamma}$  in point 2, and we have already shown that  $E[\hat{\gamma}] = \mu^3$ , so it is an unbiased estimator of  $\mu^3$ .

4. (a) We want to calculate  $E[\tilde{\gamma}] - \gamma$ .

$$E[\tilde{\gamma}] = E\left[\frac{1}{n} \sum_{i=1}^n R_i^3\right] = \frac{1}{n} \sum_{i=1}^n E[R_i^3]$$

Since  $E[R_i^3] = E[R_1^3]$ , we have  $E[\tilde{\gamma}] = \gamma$ . So, the bias of this estimator is

$$E[\tilde{\gamma}] - \gamma = 0$$

- (b) By WLLN, we have

$$\tilde{\gamma} = \frac{1}{n} \sum_{i=1}^n R_i^3 \xrightarrow[n \rightarrow \infty]{P} E[R_1^3] = \gamma$$

Therefore,  $\tilde{\gamma}$  is consistent.

5. Using the Rao-Blackwell Theorem, now we have an unbiased estimator

$$\hat{\gamma}' = (\bar{R})^3 + 3\bar{R}S^2$$

According to the slides, the minimal sufficient statistic for normal distribution is that

$$T(X) = \left(\sum_{i=1}^n X_i, \sum_{i=1}^n (X_i - \bar{X})^2\right) \propto (\bar{R}, S^2)$$

Then we could get a new unbiased estimator  $\hat{\gamma} = E[\tilde{\gamma}|T(X)]$ .

$$\begin{aligned} \hat{\gamma} &= E[\hat{\gamma}'|T(X)] \\ &= E[(\bar{R})^3 + 3\bar{R}S^2|\bar{R}, S^2] \\ &= (\bar{R})^3 + 3\bar{R}S^2 \\ &= \hat{\gamma}' \end{aligned}$$

That is, the unbiased estimator  $\hat{\gamma}'$  has already had the minimum variance.