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Graded Homework 2- Exercise 4

1.

$$P\{N_A = n_A, N_C = n_C, N_G = n_G, N_T = n_T\}$$

$$= \binom{n}{n_A} \binom{n - n_A}{n_C} \binom{n - n_A - n_C}{n_G} (1 - \theta)^{n_A} (\theta - \theta^2)^{n_C} (\theta^2 - \theta^3)^{n_G} (\theta^3)^{n_T}$$

2.

$$\begin{split} \log L(\theta) = & n_A \log(1-\theta) + n_C \log(\theta-\theta^2) + n_G \log(\theta^2-\theta^3) + n_T \log \theta^3 + C \\ = & n_A \log(1-\theta) + n_C \log(1-\theta) + n_G \log(1-\theta) \\ & + n_C \log \theta + 2n_G \log \theta + 3n_T \log \theta + C \\ \text{where } C \text{ is a constant} \end{split}$$

MLE equation:

$$0 = \frac{d}{d\theta} \log L(\theta)$$
$$= -\frac{n_A + n_C + n_G}{1 - \theta} + \frac{n_C + 2n_G + 3n_T}{\theta}$$

Solve for θ , and we have

$$\hat{\theta} = \frac{N_C + 2N_G + 3N_T}{N_A + 2N_C + 3N_G + 3N_T}$$

3. $\hat{\theta}$ is consistent and asymptotically normal distributed, so we only need to calculate $I(\theta)$.

$$\begin{split} I(\theta) &= -E \left[\frac{d^2}{(d\theta)^2} \log L(\theta) \right] \\ &= -E \left[\frac{d}{d\theta} - \frac{n_A + n_C + n_G}{1 - \theta} + \frac{n_C + 2n_G + 3n_T}{\theta} \right] \\ &= E \left[\frac{n_A + n_C + n_G}{(1 - \theta)^2} + \frac{n_C + 2n_G + 3n_T}{\theta^2} \right] \\ &= \frac{E[n_A] + E[n_C] + E[n_G]}{(1 - \theta)^2} + \frac{E[n_C] + 2E[n_G] + 3E[n_T]}{\theta^2} \\ &= n \cdot \frac{1 + \theta + \theta^2}{\theta - \theta^2} \end{split}$$

Therefore,

$$\begin{split} \hat{\theta} \xrightarrow[n \to \infty]{D} N\left(\theta, \frac{1}{n \cdot I(\theta)}\right) \\ = N\left(\theta, \frac{\theta(1 - \theta)}{n^2(1 + \theta + \theta^2)}\right) \end{split}$$

4. We claim $a_A = 0, a_C = a_G = a_T = \frac{1}{n}$ can satisfy that $E[T] = \theta$. In this case, $T = \frac{N_C + N_G + N_T}{n}$.

$$\begin{split} E[T] &= \frac{E[N_C + N_G + N_T]}{n} \\ &= \frac{E[n - N_A]}{n} \\ &= 1 - \frac{E[N_A]}{n} \\ &= 1 - \frac{n(1 - \theta)}{n} \\ &= \theta \end{split}$$

Therefore, T is an unbiased estimator.

5.

$$var(T) = var\left(\frac{N_C + N_G + N_T}{n}\right)$$
$$= var\left(\frac{n - N_A}{n}\right)$$
$$= \frac{1}{n^2} \cdot var\left(N_A\right)$$
$$= \frac{1}{n^2} \cdot n(1 - \theta)\theta$$
$$= \frac{\theta(1 - \theta)}{n}$$

$$var(\hat{\theta}) \xrightarrow[n \to \infty]{} \frac{1}{n \cdot I(\theta)} = \frac{\theta(1-\theta)}{n^2(1+\theta+\theta^2)}$$

Asymptotic_Eff(
$$T, \hat{\theta}$$
) = $\frac{\text{MSE}(T)}{\text{MSE}(\hat{\theta})} = \frac{var(T)}{var(\hat{\theta})} = n(1 + \theta + \theta^2) \xrightarrow[n \to \infty]{} \infty$

Thus T is much worse than $\hat{\theta}$.

6. MLE without common parameter θ is still consistent and asymptotically normal distributed. This MLE includes four estimators, $\hat{p_i} = \frac{n_i}{n}$, where $i \in \{A, C, G, T\}$ and $\sum \hat{p_i} = 1$.

Since $\hat{\theta}$ knows more information than $\hat{p_i}$, we think $\hat{\theta}$ has a lower asymptotic variance.

Previous point shows that $\hat{\theta}$ is "infinitely" better than T, so we think \hat{p}_i should also be better than T.

In conclusion, the new MLE estimator is consistent. As for asymptotic efficiency, it is worse than $\hat{\theta}$ but better than T.

7. We will use likelihood ratio test.

 $H_0: \mathbf{p} = \mathbf{p}(\theta)$ v. $H_1: \mathbf{p}$ does not depend on a common parameter θ .

 H_0 means $\hat{\theta}$ is correct, and H_1 means $\hat{p_i}$ in previous point is correct.

Test statistic $\Lambda = 2 \log L(H_1) - 2 \log L(H_0)$. $\log L(H_0)$ refers to the log likelihood under H_0 , which has already been shown earlier. $\log L(H_1)$ is very similar.

$$\Lambda \sim \chi_2^2$$
 and $p - value = P\{\chi_2^2 \ge \Lambda\}$.