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Graded Homework 2- Exercise 5

1. We want to calculate the conditional distribution, which is $P(y^{(k)}|(n_A^{(k)}, n_B^{(k)}, n_d^{(k)})$. So the total number of relapsed $n_d^{(k)}$ is given.

We have $n_A^{(k)} + n_B^{(k)}$ patients, among which $n_A^{(k)}$ are from group A, $n_B^{(k)}$ are from group B. We know that there are $n_d^{(k)}$ patients of relapses from $n_A^{(k)} + n_B^{(k)}$.

Under H_0 , the survival function of the two groups are the same $S_A = S_B$. So the probability of individual patient being relapse is equal among the two groups.

So the occurance of relapses is random in the larger group of $n_A^{(k)} + n_B^{(k)} = m_A + m_B$, which is randomly selecting m out of $m_A + m_B$.

In the $n_d^{(k)} = m_d$ relapse cases, the probability of $y^{(k)} = m$ coming from group A is:

$$P(m|(m_A, m_B, m_d)) = \frac{\binom{m_A}{m}\binom{m_B}{m_d - m}}{\binom{m_A + m_B}{m_d}}$$

which is:

$$P(y^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)}) = \frac{\binom{n_A^{(k)}}{y^{(k)}})\binom{n_B^{(k)}}{n_d^{(k)}-y^{(k)}})}{\binom{n_A^{(k)}+n_B^{(k)}}{n_d^{(k)}}}$$

So it's a HyperGeometric distribution with parameter ($n_A^{(k)} + n_B^{(k)}, n_A^{(k)}, n_d^{(k)})$

2. From point 1, we know that $P(y^{(k)}|(n_A^{(k)}, n_B^{(k)}, n_d^{(k)})$ has a hypergeometric distribution.

$$P(y^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)}) = \frac{\binom{n_A^{(k)}}{y^{(k)}})\binom{n_B^{(k)}}{n_d^{(k)}-y^{(k)}})}{\binom{n_A^{(k)}+n_B^{(k)}}{n_d^{(k)}}}$$

So the conditional expectation is given by:

$$E^{(k)} = \frac{m \times m_A}{m_A + m_B} = \frac{y^{(k)} \times n_A^{(k)}}{n_A^{(k)} + n_B^{(k)}} = \frac{y^{(k)} n_A^{(k)}}{n^{(k)}}$$

as
$$n^{(k)} = n_A^{(k)} + n_B^{(k)}$$
.

The conditional variance is given by:

$$\begin{split} V^{(k)} &= n_d^{(k)} \frac{n_A^{(k)}}{n_A^{(k)} + n_B^{(k)}} \frac{n_A^{(k)} + n_B^{(k)} - n_A^{(k)}}{n_A^{(k)} + n_B^{(k)}} \frac{n_A^{(k)} + n_B^{(k)} - n_d^{(k)}}{n_A^{(k)} + n_B^{(k)}} \\ V^{(k)} &= \frac{n_A^{(k)} n_B^{(k)} n_A^{(k)} n_S^{(k)}}{(n_A^{(k)} + n_B^{(k)})(n_A^{(k)} + n_B^{(k)})(n_A^{(k)} + n_B^{(k)} - 1)} = \frac{n_A^{(k)} n_B^{(k)} n_A^{(k)} n_S^{(k)}}{(n^{(k)})^2 (n^{(k)} - 1)} \\ \text{as } n_S^{(k)} &= n_A^{(k)} + n_B^{(k)} - n_A^{(k)}. \end{split}$$

3. According to conditional variance formula, we can fraction $Var[y^{(k)} - E^{(k)}]$ under H_0 :

$$\begin{split} &Var[y^{(k)}-E^{(k)}] = Var[E[y^{(k)}-E^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)})]] + E[Var[y^{(k)}-E^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)})]] \\ &= Var[E[y^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)})] - E[E^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)})]] + E[Var[y^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)})]] \\ &= Var(0) + E[Var[y^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)})]] \\ &= E[V^{(k)}] \\ &\text{as } E[y^{(k)}|(n_A^{(k)},n_B^{(k)},n_d^{(k)})] = E^{(k)}. \end{split}$$

4. Under H_0 we have:

$$\begin{split} &Var[\sum_{k=1}^{K}(y^{(k)}-E^{(k)})]\\ &=Var[E[\sum_{k=1}^{K}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]+E[Var[\sum_{k=1}^{K}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[E[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})+(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]+\\ &E[Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})+(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[E[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]+E[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]+\\ &E[Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]+Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &(According\ to\ iid\ assumption\ of\ k)\\ &=Var[E[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]+0]+E[Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]+\\ &Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})]+E[Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})]+E[Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})]+E[Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})]+E[Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})]+E[Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]]\\ &=Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})]+E[Var[(y^{(K)}-E^{(K)})|(n_A^{(K)},n_B^{(K)},n_d^{(K)})]\\ &=Var[\sum_{k=1}^{K-1}(y^{(k)}-E^{(k)})]+E[Var[(y^{(K)}-E^{(K$$

Similarly, we can apply this process to every other value of k, and get that:

$$Var[\sum_{k=1}^{K} (y^{(k)} - E^{(k)})] = \sum_{k=1}^{K} E[V^{(K)}]$$

5. H_0 : the survival functions of group A and group B are identical. We implement the test to the sample in 1.2 using r.

```
df <- df[order(df$time);]
df$k <- floor(df$time)
for (i in c(1:nrow(df))){
    df$n_a[i] <- nrow(df %% filter((types == 1 & k == df$k[i])))
    df$n_b[i] <- nrow(df %% filter((types == 2 & k == df$k[i])))
    df$n_d[i] <- nrow(df %% filter(survive == 1 & k == df$k[i]))
    df$y_k[i] <- nrow(df %% filter(types == 1 & k == df$k[i]))
}
df$y_k[i] <- nrow(df %% filter(types == 1 & k == df$k[i] & survive == 1)))
}
df_5 <- distinct(df[c("k","n_a","n_b","n_d","y_k")])
for (i in c(1:nrow(df_5))){
    df_5$E_k <- (df_5$n_a*df_5$n_d)/(df_5$n_a + df_5$n_b)
    df_5$V_k <- (df_5$n_a*df_5$n_b*df_5$n_d*(df_5$n_a+df_5$n_b-df_5$n_d))/(((df_5$V_k)))
}
df_5$V_k[is.na(df_5$V_k)] <- 0

Z = (sum(df_5$y_k-df_5$E_k))/sqrt(sum(df_5$V_k))
Z</pre>
```

Here we compute the Z value and get -1.32141.

By CLT we know that Z has N(0,1) distribution. As -1.32141 > -1.96, we failed to reject H_0 . So we cannot reject the hypothesis that the two survival functions are identical.