

LO7: Particle Filter and Monte-Carlo Localization.

* Summary of LO6 localization.

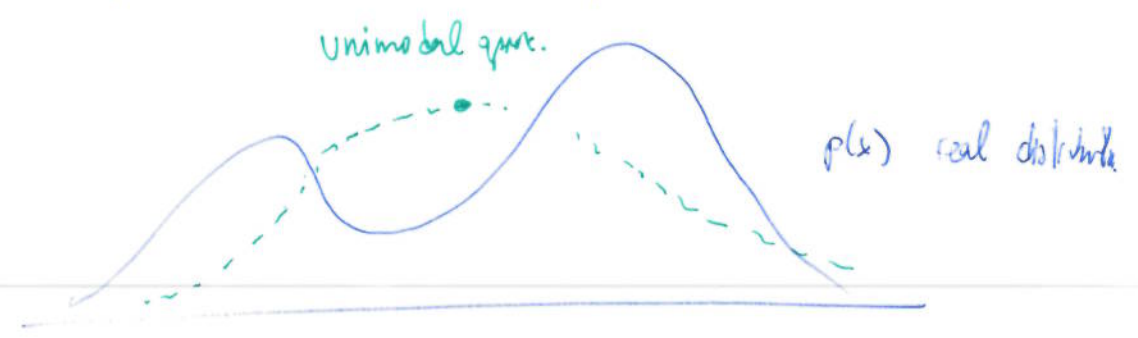
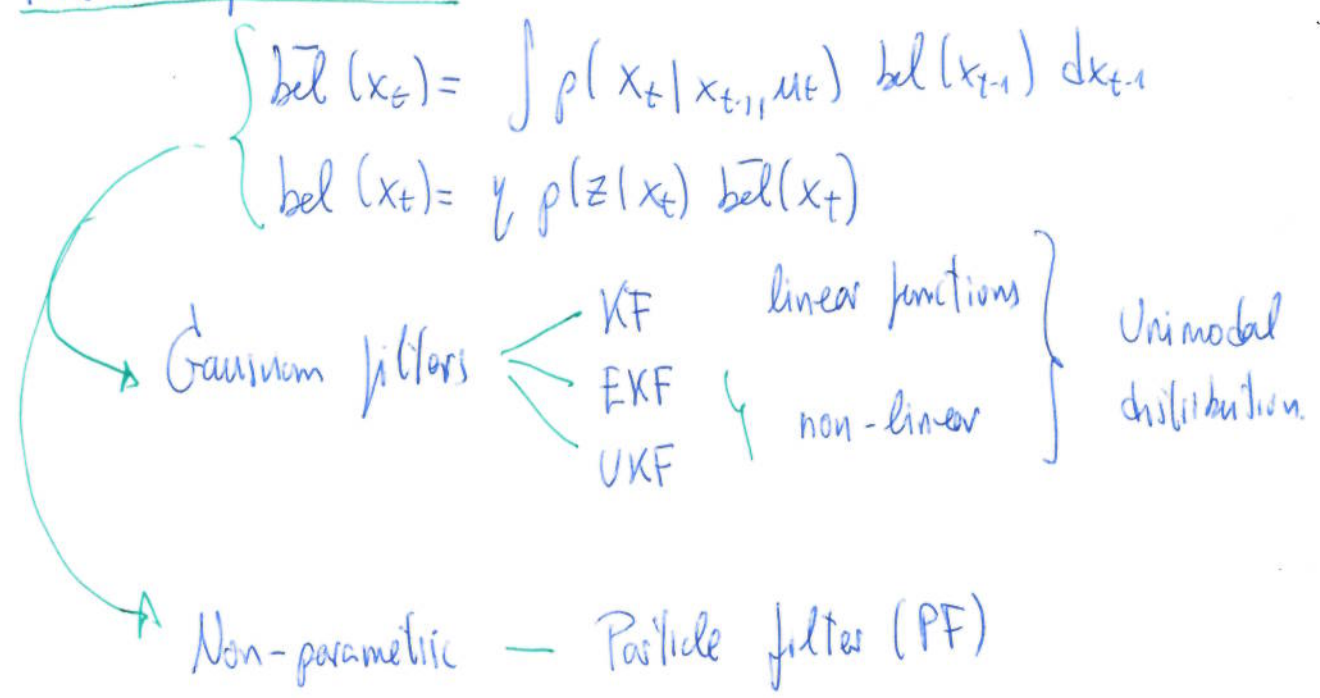
$$m = \left\{ \begin{bmatrix} m_{1,x} \\ m_{1,y} \end{bmatrix}, \begin{bmatrix} m_{2,x} \\ m_{2,y} \end{bmatrix}, \dots, \begin{bmatrix} m_{j,x} \\ m_{j,y} \end{bmatrix}, \dots \right\} \quad \text{map of known landmarks}$$

The localization problem becomes a state estimation problem.

$$\text{bel}(x_t) = p(x_t | \mathcal{U}, \mathcal{Z}, m) \quad \Leftrightarrow \text{EKF, UKF.}$$

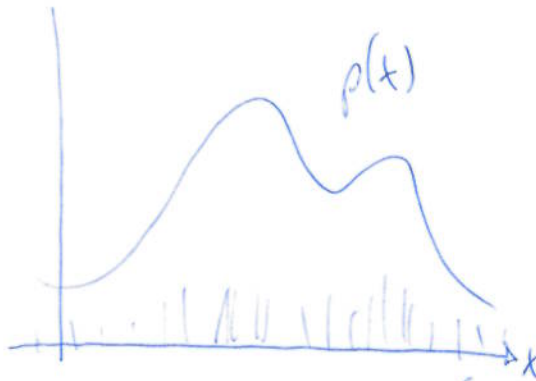
Assume (for now) $C_t^i = j \quad (z^i \rightarrow m_j)$ Known correspondence

Particle filter (PF)



Particle set $\chi_t = \{ \langle x_t^{[1]}, w_t^{[1]} \rangle, \dots, \langle x_t^{[M]}, w_t^{[M]} \rangle \}$

Ex:



$$x^{[m]} \sim p(x)$$

$$w^{[m]} = p_Z(x^{[m]})$$

(Importance factors)
Weighted samples. Particles become a good representation of distributions (if M is large enough)

Q: weights on PS1? Sample mean, sample covariance?

1. Particle filter (χ_{t-1}, u_t, z_t):

2: for $m = 1:M$

3: $x_t^{[m]} \sim p(x_t | u_t, x_{t-1}^{[m]})$

4: $w_t^{[m]} = p(z_t | x_t^{[m]})$

5: $\bar{\chi}_t = \bar{\chi}_t \cup \langle x_t^{[m]}, w_t^{[m]} \rangle$

6: $\chi_t = \text{resample}^*(\bar{\chi}_t)$

propagation $bel(x_t)$

correction

(better correction)

χ_t are "drawn" from $bel(x_t)$ and not bel

[Gordon] reading introduces resampling and required for the

PF to work properly.

* Bayes filter for full states

$$bel(x_{0:t}) = p(x_{0:t} | u_{1:t}, z_{1:t})$$

particles $x_{0:t}^{[m]} = x_0^{[m]}, x_1^{[m]}, \dots, x_t^{[m]}$ Sequence of samples of states.

(new)

$$bel(x_{0:t}) = \gamma p(z_t | x_{0:t}, z_{1:t-1}, u_{1:t}) p(x_{0:t} | z_{1:t-1}, u_{1:t})$$

$$\begin{aligned} \text{Markov + Bayes} &= \gamma p(z_t | x_t) p(x_t | x_{0:t-1}, z_{1:t-1}, u_{1:t}) p(x_{0:t-1} | z_{1:t-1}, u_{1:t}) \\ &= \gamma p(z_t | x_t) p(x_t | x_{t-1}, u_t) \underbrace{p(x_{0:t-1} | z_{1:t-1}, u_{1:t-1})}_{bel(x_{0:t-1})} \end{aligned}$$

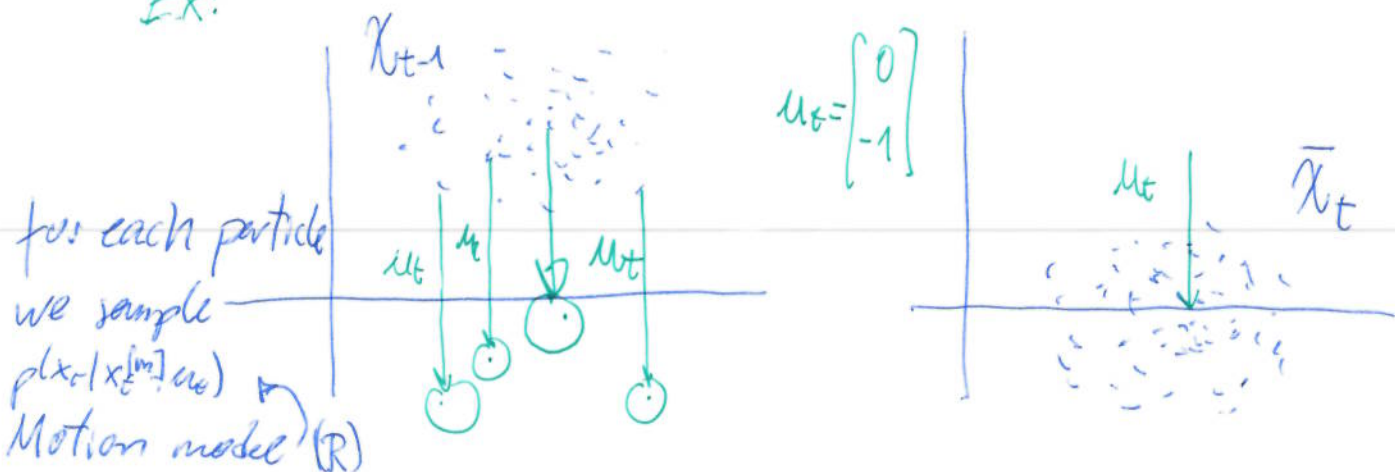
$$\begin{aligned} \bar{bel}(x_{0:t}) &= p(x_t | x_{t-1}, u_t) \bar{bel}(x_{0:t-1}) \\ \bar{bel}(x_{0:t}) &= \gamma p(z_t | x_t) \bar{bel}(x_{0:t}) \end{aligned}$$

From this full state Bayes (no marginalization)
given a particle $x_{t-1}^{[m]} \sim \bar{bel}(x_{0:t-1})$

$$\bar{bel}(x_{0:t}) \left\{ \begin{array}{l} \bar{x}_t^{[m]} \sim p(x_t | x_{t-1}^{[m]}, u_t) \cdot (w_{t-1}^{[m]}) \\ \bar{w}_t^{[m]} = 1 \cdot w_{t-1}^{[m]} \end{array} \right. \quad \begin{array}{l} \text{Sample.} \\ \text{importance factor from bel.} \end{array}$$

(Q: $w_{t-1}^{[m]}$ on single)

Ex:



* Importance sampling

$$E_p \{ I(x \in A) \} = \int I(x \in A) p(x) dx$$

$$= \int I(x \in A) \underbrace{\frac{p(x)}{q(x)}} \cdot q(x) dx$$

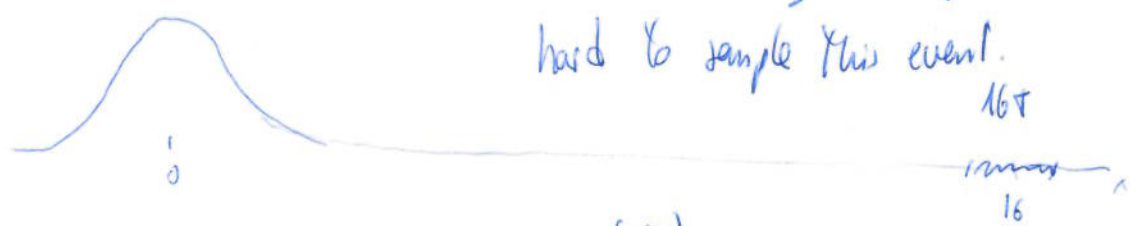
$$= E_q \{ I(x \in A) \cdot w(x) \}$$

$$w(x) = \frac{p(x)}{q(x)} \quad \text{Importance factor}$$

$p(x)$ target distribution (usually we can't use)
 $q(x)$ proposal distribution.

Ex: Probability of sample a 1d r.v X in $[15, 17]$ if $p(x) \propto N(0, 1)$? $A = \{x \mid 15 \leq x \leq 17\}$

$$1) P(x \in A) = \sum I(x^m \in A) \cdot p(x^m) \quad , \quad x^m \sim p(x^m)$$



hard to sample this event.

$$2) \text{I.S.: } P(x \in A) = \sum I(x^m \in A) \underbrace{\frac{p(x^m)}{q(x^m)}}_{w^m} \cdot q(x^m) \quad , \quad x^m \sim q(x)$$

for instance $q(x) = N(16, 1)$
 (proposal distribution.)

With this proposal distribution we don't need millions of samples but only hundreds.

$$w^m = \frac{N(x^m; 0, 1)}{N(x^m, 16, 1)} \quad \text{d d d}$$

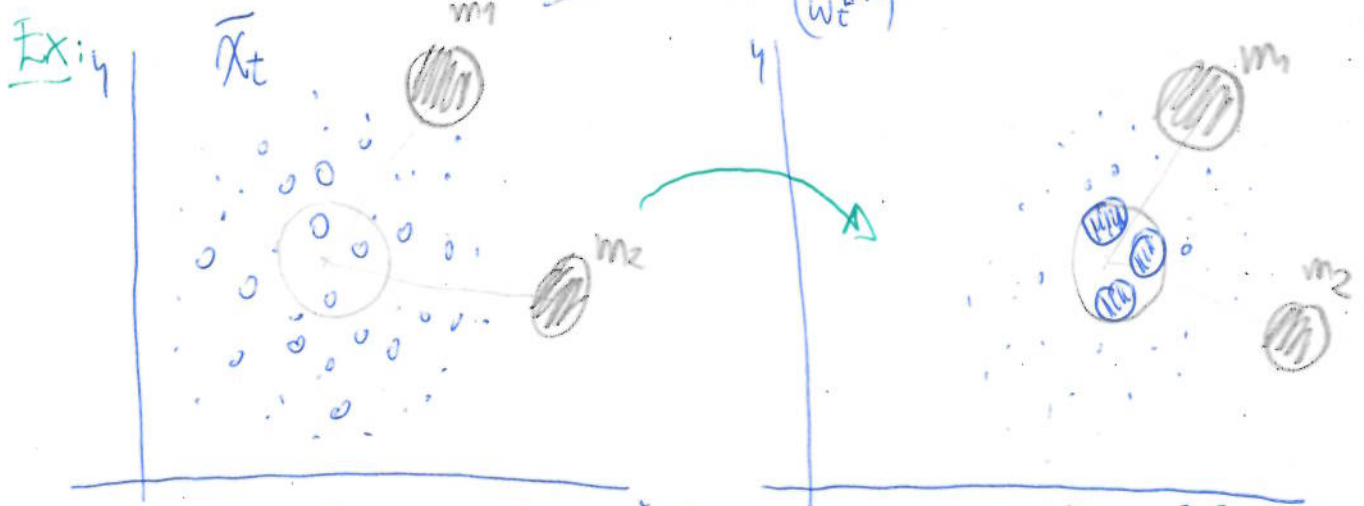
→ Correction step (line 4)

$$\underbrace{\text{bel}(x_{0:t})}_{\text{target distribution}} = \eta \underbrace{p(z_t | x_t)}_{\bar{X}_t \text{ particle } x_t \text{ representing this belief.}} \cdot \underbrace{\text{bel}(x_{0:t})}_{\text{Proposed distribution}}$$

In order to correctly characterize the posterior $\text{bel}(x_{0:t})$ we are going to weight the particles drawn (already) \bar{X}_t with proper weights or importance factors.

$$x_t^{[m]} = \bar{X}_t^{[m]} \quad (\text{previously drawn}) \text{ from proposal dist}$$

$$w_t^{[m]} = \eta \cdot \frac{\text{target distribution}}{\text{proposal distribution}} = \eta \frac{p(z_t | x_t^{[m]})}{\text{bel}(x_{0:t}^{[m]})} = \eta \cdot p(z_t | x_t^{[m]})$$



The size of the point corresponds to $w_t^{[m]}$

new weight $w_t^{[m]}$

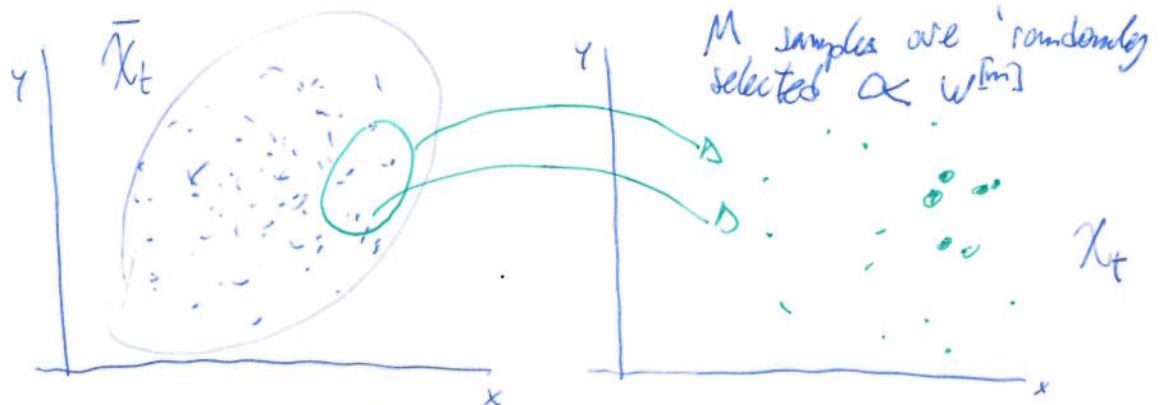
Problem: creates an almost empty set of particles with weights non-zero \Rightarrow and many particles with low weights.

Degeneracy over time.

* Resampling (The solution)

Idea: survival of the fittest. Only the most likely particles ($w^{[m]} \uparrow$) 'might' survive.

Ex:

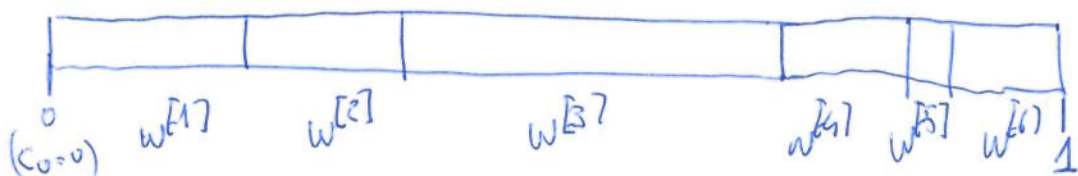


from M samples on \bar{X}_t we get M samples on X_t (closer to true belief)

→ Independent Resampling. First solution

We create a cumulative distribution function:

$$C_m = C_{m-1} + w^{[m]} \quad (+ \text{normalization})$$



for $m = 1:M$
 $u \sim \mathcal{U}(0,1)$ (uniform distribution)
 $j = \text{find}(C_m, u)$
 $X_t = X_t \cup \langle x_t^{[j]}, w_t^{[j]} \rangle$

⚠ Problem: over time, independent resampling induces a loss of diversity in the particle population X_t

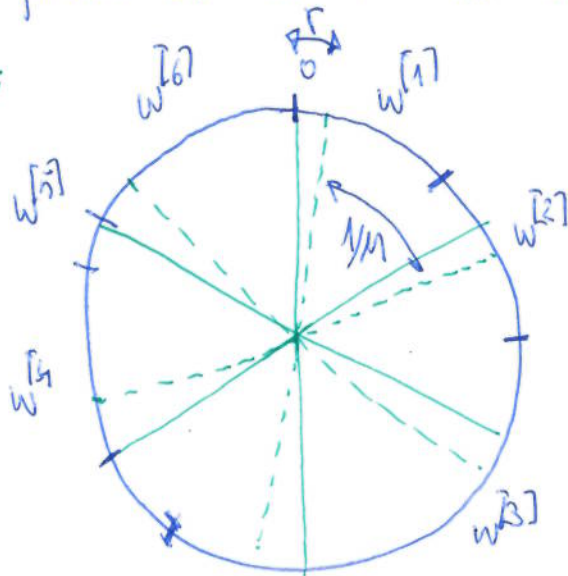
low-variance sampling

$$C_m = C_{m-1} + w^{[m]}$$



Idea: given an initial random configuration r , we add particles at intervals $1/M$ over the full set.

Ex:



$$r = 0$$

$$\chi_t = \{1, 2, 3, 3, 5\}$$

$$r \sim \mathcal{U}[0, 1/M]$$

$$\chi_t = \{$$

Algorithm: low-variance-sampler ($\bar{\chi}_t$): (ProbRob 110)

$$\chi_t = \emptyset, \quad C = w_t^{[1]}, \quad i = 1$$

$$r \sim \mathcal{U}[0, 1/M]$$

for $m = 1:M$

$$U = r + (m-1) \cdot \frac{1}{M}$$

while $U > C$

$$\begin{aligned} & i++ \\ & C = C + w^{[i]} \end{aligned}$$

$$\chi_t = \chi_t \cup \langle \bar{\chi}_t^{[i]}, \frac{1}{M} \rangle$$

return χ_t .

* Monte-Carlo localization (MCL)

@ readings : Dellaert'99

We want to solve the Markov localization (LOB) using PF.

$$\text{bel}(x_t) = p(x_t | \mathcal{U}, \mathcal{Z}, m) \longrightarrow x_t \quad \begin{array}{l} \text{Particle set} \\ \text{representing belief.} \end{array}$$

map of landmarks

Algorithm: MCL ($\mathcal{X}_{t-1}, \mu_t, \mathcal{Z}_t, m$):

- 1: $\bar{\mathcal{X}}_t = \mathcal{X}_t = \emptyset$
- 2: for $m = 1:M$
- 3: $x_t^{[m]} = \text{sample-motion-model}(\mu_t, x_{t-1}^{[m]})$ (LOS)
- 4: $w_t^{[m]} = \text{measurement-model}(\mathcal{Z}_t, x_t^{[m]}, m)$ (LOB, 5)
- 5: $\bar{\mathcal{X}}_t = \bar{\mathcal{X}}_t \cup \langle x_t^{[m]}, w_t^{[m]} \rangle$
- 6: $\mathcal{X}_t = \text{low-variance-sampling}(\bar{\mathcal{X}}_t)$

Next lecture SLAM: ProbRob Ch 10 + readings @ camera