

Q1

```
# read data
WOOL <- read.csv("Warpbreaks.csv")

# model poisson regression
mod<-glm(breaks~., WOOL, family = poisson(link = "log"))
summary(mod)

##
## Call:
## glm(formula = breaks ~ ., family = poisson(link = "log"), data = WOOL)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -3.6871  -1.6503  -0.4269   1.1902   4.2616
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  3.17347    0.05567  57.002 < 2e-16 ***
## woolB       -0.20599    0.05157  -3.994 6.49e-05 ***
## tensionL     0.51849    0.06396   8.107 5.21e-16 ***
## tensionM     0.19717    0.06833   2.885 0.00391 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
##      Null deviance: 297.37  on 53  degrees of freedom
## Residual deviance: 210.39  on 50  degrees of freedom
## AIC: 493.06
##
## Number of Fisher Scoring iterations: 4

vcov(mod)

##              (Intercept)      woolB      tensionL      tensionM
## (Intercept)  0.003099518 -1.193312e-03 -2.564099e-03 -2.564099e-03
## woolB       -0.001193312  2.659585e-03  5.034078e-19  2.454025e-19
## tensionL    -0.002564099  5.034078e-19  4.090810e-03  2.564099e-03
## tensionM    -0.002564099  2.454025e-19  2.564099e-03  4.669354e-03

confint(mod,level=0.995)

## Waiting for profiling to be done...
##
##              0.3 %      99.8 %
## (Intercept)  3.013926429  3.32660462
## woolB       -0.351173215 -0.06152152
## tensionL     0.340192963  0.69951267
## tensionM     0.005811379  0.38973193
```

Part c

```
# Extracting the design matrix
X <- model.matrix(mod)
```

```

# betaest<-modest$coef
sigma <-vcov(mod)
y <- WOOL$breaks
p <-dim(X)[2]    #number of parameters

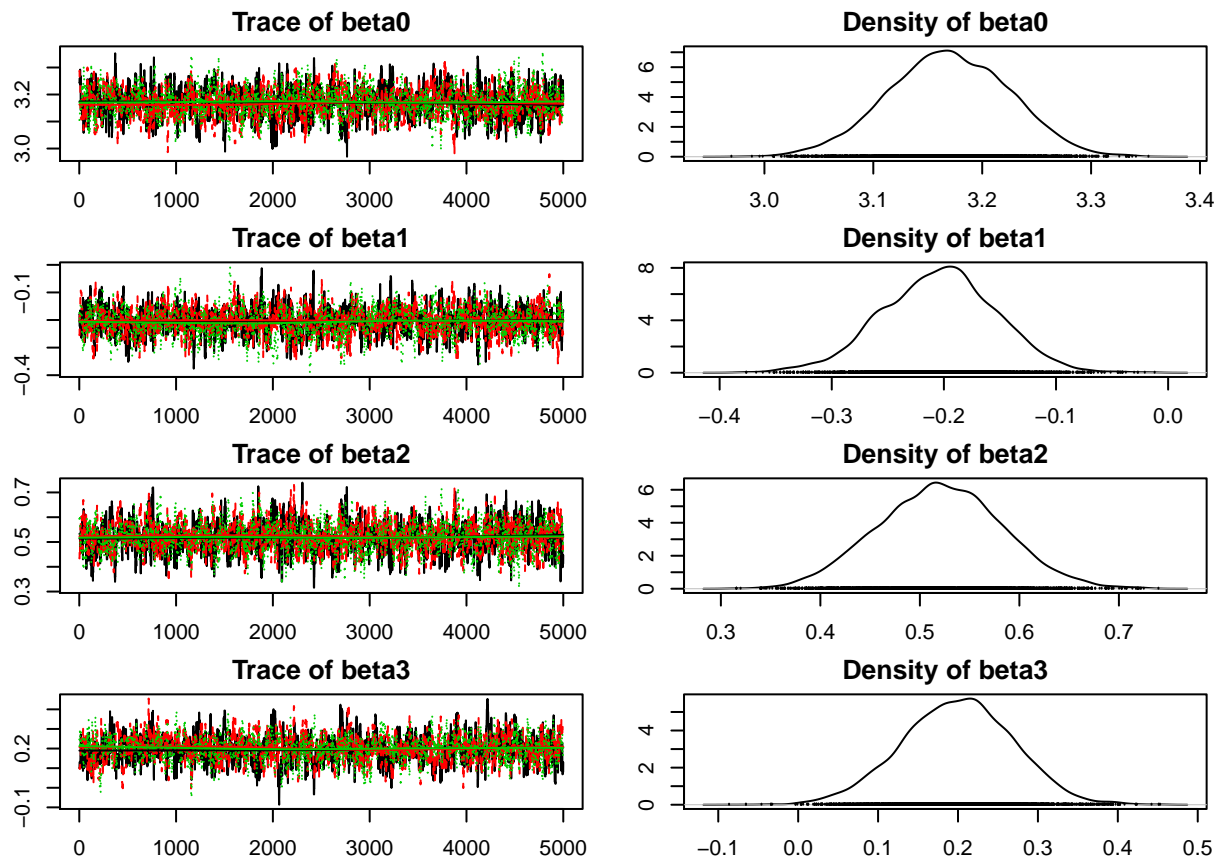
#Part one: function for performing Metropolis sampling for poisson regression normal random walk.
#Inputs:
#y: vector of responses
#n: vector (or scalar) of trial sizes.
#X: predictor matrix including intercept.
#c: rescaling for variance-covariance matrix, scalar  $J(\lambda/\lambda(t-1)) = N(\lambda(t-1), c^2 \Sigma)$ 
#Sigma is variance covariance matrix for parameters in J()
#iter: number of iterations
#burnin: number of initial iterations to throw out.
Metropolis.fn<-function(y,X,c,Sigma,iter,burnin){
p <-dim(X)[2]    #number of parameters
library(mvtnorm)
beta0<-rnorm(p) #initial values.
beta.sim<-matrix(0,iter,p) #matrix to store iterations
beta.sim[1,]<-beta0
for(i in 1:(iter-1)){
beta.cand <-rmvnorm(1,mean=beta.sim[i,],sigma=c^2*Sigma) #draw candidate (jointly)
beta.cand <-as.numeric(beta.cand)
xbc      <-X%%beta.cand
lambda.c <-exp(xbc)    #Calculating probability of success for candidates.
xb       <-X%%beta.sim[i,]
lambda.b <-exp(xb)     #Calculating probability of success for lambda(t-1).
#difference of log joint distributions.
r<-sum( dpois(y,lambda.c,log=TRUE)-dpois(y,lambda.b ,log=TRUE) )
#Draw an indicator whether to accept/reject candidate
ind<-rbinom(1,1,exp( min(c(r,0)) ) )
beta.sim[i+1,]<- ind*beta.cand + (1-ind)*beta.sim[i,]
}

#Removing the iterations in burnin phase
results<-data.frame(beta.sim[-c(1:burnin),])
names(results)<-c('beta0','beta1','beta2','beta3') #column names
return(results)
}

library(coda)

# c =1.6
par(mar=c(2,2,2,2))
results16_1 <- Metropolis.fn(y,X,c=1.6/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
results16_2 <- Metropolis.fn(y,X,c=1.6/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
results16_3 <- Metropolis.fn(y,X,c=1.6/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
c16_1 <- mcmc(results16_1[c('beta0','beta1','beta2','beta3')])
c16_2 <- mcmc(results16_2[c('beta0','beta1','beta2','beta3')])
c16_3 <- mcmc(results16_3[c('beta0','beta1','beta2','beta3')])
c16 <- mcmc.list(c16_1,c16_2,c16_3)
plot(c16)

```



```
# Acceptance Rate
```

```
1 - rejectionRate(c16)
```

```
##      beta0      beta1      beta2      beta3
## 0.4704941 0.4704941 0.4704941 0.4704941
```

```
# Gelman and Rubin diagnostic
```

```
gelman.diag(c16)
```

```
## Potential scale reduction factors:
```

```
##
```

```
##      Point est. Upper C.I.
```

```
## beta0          1          1.01
```

```
## beta1          1          1.00
```

```
## beta2          1          1.00
```

```
## beta3          1          1.01
```

```
##
```

```
## Multivariate psrf
```

```
##
```

```
## 1
```

```
# Effective sample size
```

```
effectiveSize(c16)
```

```
##      beta0      beta1      beta2      beta3
## 946.1410 980.9337 1016.6310 949.8586
```

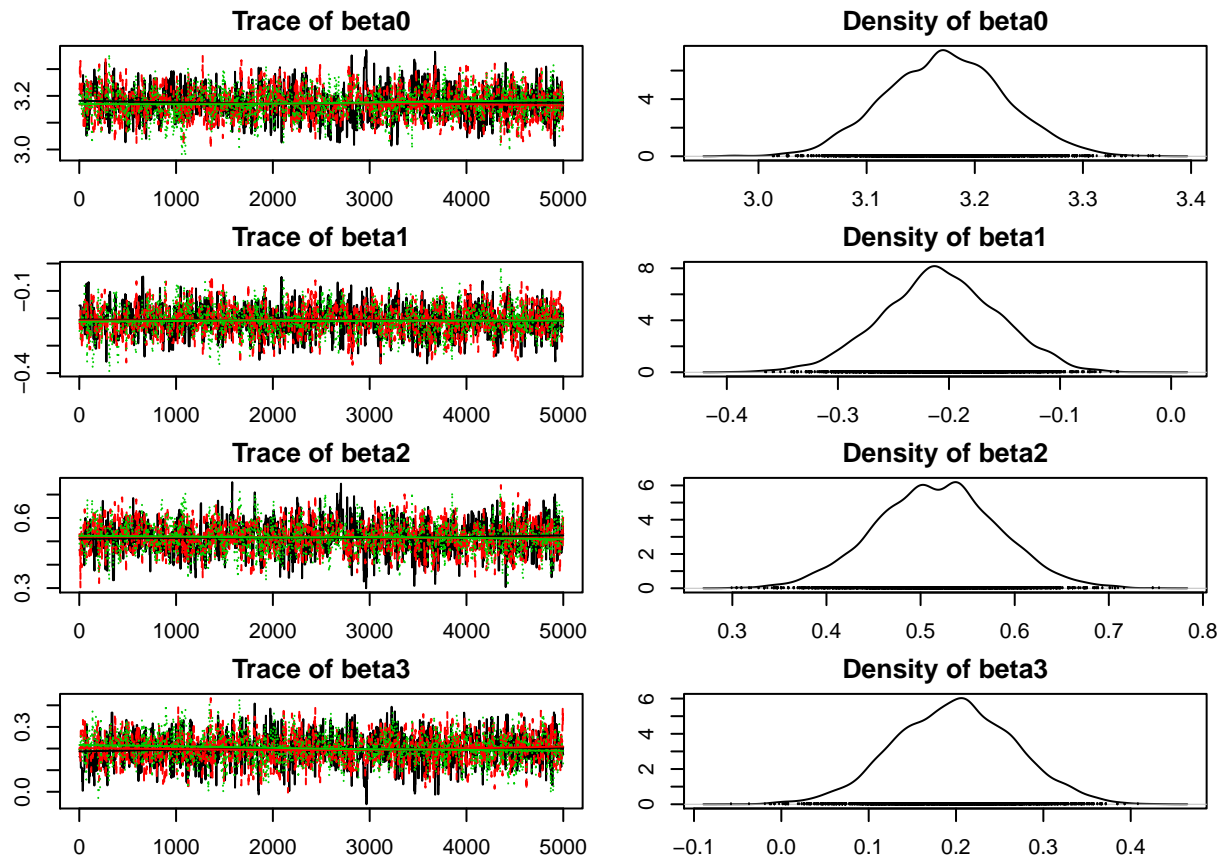
```
# c = 2.4
```

```
par(mar=c(2,2,2,2))
```

```

results24_1 <- Metropolis.fn(y,X,c=2.4/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
results24_2 <- Metropolis.fn(y,X,c=2.4/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
results24_3 <- Metropolis.fn(y,X,c=2.4/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
c24_1 <- mcmc(results24_1[c('beta0','beta1','beta2','beta3')])
c24_2 <- mcmc(results24_2[c('beta0','beta1','beta2','beta3')])
c24_3 <- mcmc(results24_3[c('beta0','beta1','beta2','beta3')])
c24 <- mcmc.list(c24_1,c24_2,c24_3)
plot(c24)

```



```

# Acceptance Rate
1 - rejectionRate(c24)

```

```

##      beta0      beta1      beta2      beta3
## 0.3011936 0.3011936 0.3011936 0.3011936

```

```

# Gelman and Rubin diagnostic
gelman.diag(c16)

```

```

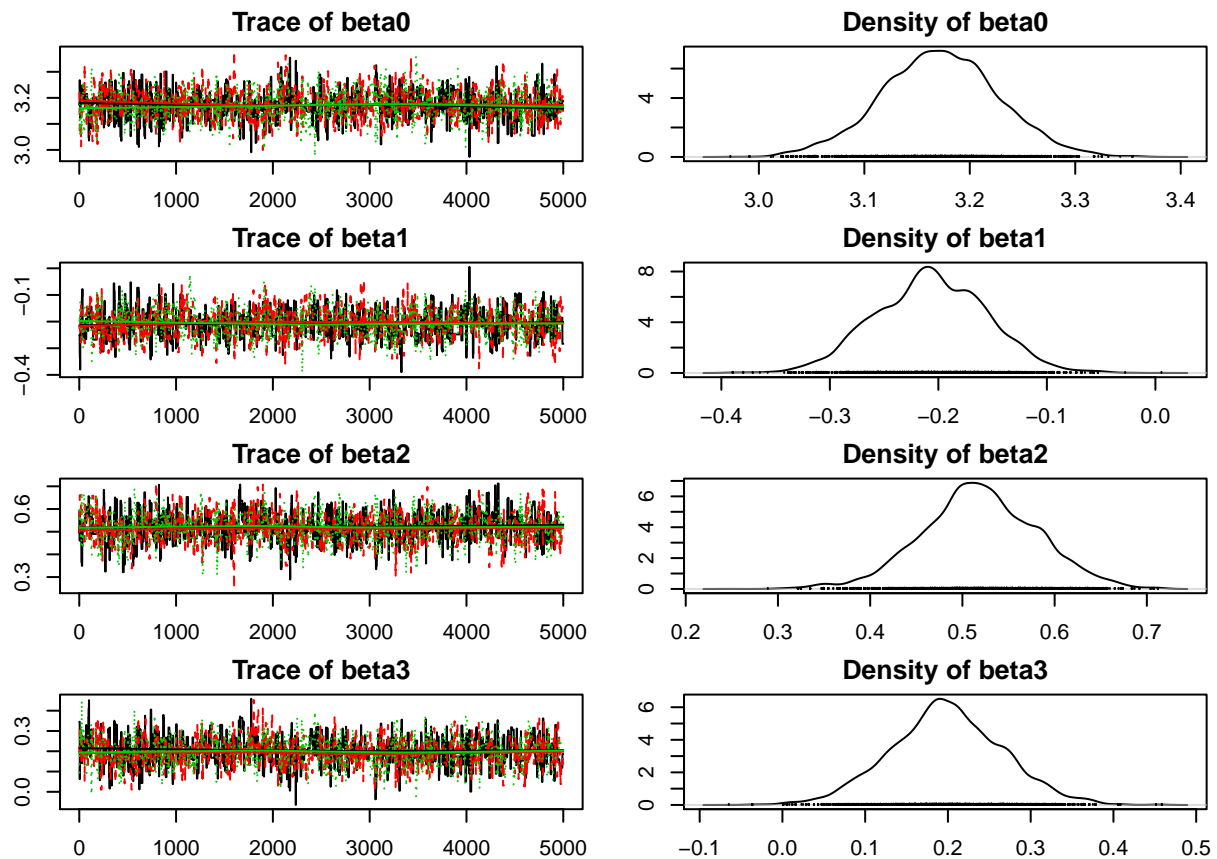
## Potential scale reduction factors:
##
##      Point est. Upper C.I.
## beta0          1          1.01
## beta1          1          1.00
## beta2          1          1.00
## beta3          1          1.01
##
## Multivariate psrf
##

```

```
## 1
# Effective sample size
effectiveSize(c24)

##      beta0      beta1      beta2      beta3
## 1146.365 1151.456 1144.601 1183.105

# c = 3.2
par(mar=c(2,2,2,2))
results32_1 <- Metropolis.fn(y,X,c=3.2/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
results32_2 <- Metropolis.fn(y,X,c=3.2/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
results32_3 <- Metropolis.fn(y,X,c=3.2/sqrt(p),Sigma = sigma,iter=10000,burnin =5000)
c32_1 <- mcmc(results32_1[c('beta0','beta1','beta2','beta3')])
c32_2 <- mcmc(results32_2[c('beta0','beta1','beta2','beta3')])
c32_3 <- mcmc(results32_3[c('beta0','beta1','beta2','beta3')])
c32 <- mcmc.list(c32_1,c32_2,c32_3)
plot(c32)
```



```
# Acceptance Rate
1 - rejectionRate(c32)

##      beta0      beta1      beta2      beta3
## 0.1778356 0.1778356 0.1778356 0.1778356

# Gelman and Rubin diagnostic
gelman.diag(c16)

## Potential scale reduction factors:
##
```

```
##          Point est. Upper C.I.
## beta0      1      1.01
## beta1      1      1.00
## beta2      1      1.00
## beta3      1      1.01
##
## Multivariate psrf
##
## 1

# Effective sample size
effectiveSize(c32)

##      beta0      beta1      beta2      beta3
## 1043.194 1020.809 1056.539 1074.782

# comments
AcceptanceRate <- c(1 - rejectionRate(c16), 1 - rejectionRate(c24), 1 - rejectionRate(c32))
EffectiveSampleSize <- c(effectiveSize(c16), effectiveSize(c24), effectiveSize(c32))
rown <- c("c=1.6 beta0", "c=1.6 beta1", "c=1.6 beta2", "c=1.6 beta3",
          "c=2.4 beta0", "c=2.4 beta1", "c=2.4 beta2", "c=2.4 beta3",
          "c=3.2 beta0", "c=3.2 beta1", "c=3.2 beta2", "c=3.2 beta3")
summarychain <- data.frame(AcceptanceRate, EffectiveSampleSize, row.names = rown)
summarychain

##          AcceptanceRate EffectiveSampleSize
## c=1.6 beta0      0.4704941           946.1410
## c=1.6 beta1      0.4704941           980.9337
## c=1.6 beta2      0.4704941          1016.6310
## c=1.6 beta3      0.4704941           949.8586
## c=2.4 beta0      0.3011936          1146.3646
## c=2.4 beta1      0.3011936          1151.4565
## c=2.4 beta2      0.3011936          1144.6011
## c=2.4 beta3      0.3011936          1183.1047
## c=3.2 beta0      0.1778356          1043.1944
## c=3.2 beta1      0.1778356          1020.8089
## c=3.2 beta2      0.1778356          1056.5386
## c=3.2 beta3      0.1778356          1074.7824
```

As we increase c the Acceptance rate decreases. With a larger c we have that the jumping distribution can draw samples from a larger range as seen from the density plots above.

For the case $c=2.4/\sqrt{p}$ the is reported to be 0.2903247. This is in line with the results from the lectures as

```
# c=1.6/sqrt(p)
grc16 <- gelman.diag(c16)
grc16$psrf
```

```
##          Point est. Upper C.I.
## beta0      1.004507      1.010181
## beta1      1.001388      1.002239
## beta2      1.000946      1.001608
## beta3      1.003577      1.006501
```

```
# c=2.4/sqrt(p)
grc24 <- gelman.diag(c24)
grc24$psrf
```

```
##          Point est. Upper C.I.  
## beta0    1.007333    1.018318  
## beta1    1.001511    1.001875  
## beta2    1.001812    1.002133  
## beta3    1.002737    1.004956
```

```
#  $c=3.2/\sqrt{p}$   
grc32 <- gelman.diag(c32)  
grc32$psrf
```

```
##          Point est. Upper C.I.  
## beta0    1.003925    1.014018  
## beta1    1.003742    1.009635  
## beta2    1.007816    1.027640  
## beta3    1.005574    1.019073
```