

431 Class 25

`thomaseLove.github.io/431`

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Today's Agenda

- 1 What exactly is R doing if you ignore missing values when fitting models?
 - What does `type.convert()` do?
 - `na.omit` vs. `na.exclude` vs. `na.delete`
- 2 Use multiple imputation to deal with missing data in fitting a linear regression with `lm` using the `mice` package.

(MICE = Multiple Imputation through Chained Equations)

My Setup

My R project for Class 25 is located in the following folder...

```
"unique_stuff_to_me/2021-431/431-classes/class25"
```

```
library(here)
```

here() starts at C:/Users/Thomas/Dropbox/2021-431/431-classes/

```
library(magrittr); library(knitr)
```

```
library(janitor); library(naniar)
```

```
library(mice)
```

```
# mice = multiple imputation through chained equations
```

```
library(broom)
```

```
library(tidyverse)
```

```
theme_set(theme_bw())
```

What happens if you fit a regression model without doing anything at all about missing data?

What happens if you ignore NAs?

Let's open a small, simulated data set with 100 subjects and some missing values.

```
sim1 <- read_csv(here("data", "c12_sim1.csv")) %>%  
  type.convert(as.is = FALSE, na.strings = "NA")
```

```
head(sim1)
```

```
# A tibble: 6 x 6
```

	subject	out_q	out_b	pred1	pred2	pred3
	<fct>	<dbl>	<fct>	<dbl>	<dbl>	<fct>
1	S001	81.1	Yes	8.8	20.5	Middle
2	S002	105.	No	7.1	24.9	High
3	S003	NA	<NA>	9.9	17.4	Middle
4	S004	NA	No	8.9	31.8	<NA>
5	S005	75.9	<NA>	NA	22	High
6	S006	79.8	No	9.7	NA	<NA>

What does `type.convert()` actually do?

Tries to convert each column (individually) to either logical, integer, numeric, complex or (if a character vector) to factor.

- The first type (from that list) that can accept all non-missing values is chosen.
- If all of the values are missing, the column is converted to logical.
- Columns containing just F, T, FALSE, TRUE or NA values are made into logical.
- Use the `na.strings` parameter to add missing strings (default = "NA")
- `as.is = FALSE` converts characters to factors. `as.is = TRUE` will become the default soon, so you should specify.

Our sim1 data

Variable	Description
subject	Subject identifier
out_q	Quantitative outcome
out_b	Binary outcome with levels Yes, No
pred1	Predictor 1 (quantitative)
pred2	Predictor 2 (also quantitative)
pred3	Predictor 3 (categories are Low, Middle, High)

- Clean up the factors?

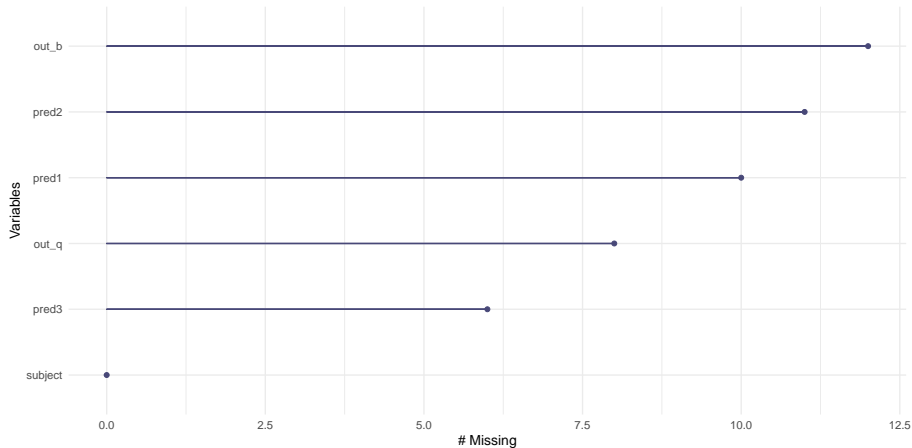
Cleaning up subject and pred3

```
sim1 <- sim1 %>%  
  mutate(subject = as.character(subject),  
         pred3 = fct_relevel(pred3, "Low", "Middle"))  
  
sim1 %>% tabyl(pred3, out_b)
```

pred3	No	Yes	NA_
Low	10	12	4
Middle	12	17	4
High	16	15	4
<NA>	4	2	0

How much missingness do we have?

```
gg_miss_var(sim1)
```



How much missingness do we have?

```
miss_var_summary(sim1)
```

```
# A tibble: 6 x 3
  variable n_miss pct_miss
  <chr>      <int>    <dbl>
1 out_b         12      12
2 pred2         11      11
3 pred1         10      10
4 out_q          8       8
5 pred3          6       6
6 subject        0       0
```

```
n_miss(sim1)
```

```
[1] 47
```

How much missingness do we have?

```
prop_complete_case(sim1)
```

```
[1] 0.65
```

```
miss_case_table(sim1)
```

```
# A tibble: 4 x 3
```

	n_miss_in_case	n_cases	pct_cases
	<int>	<int>	<dbl>
1	0	65	65
2	1	25	25
3	2	8	8
4	3	2	2

Suppose we run a linear regression

without dealing with the missing data, so that we run:

```
mod1 <- lm(out_q ~ pred1 + pred2 + pred3, data = sim1)
```

How can we tell how many observations will be used?

What happens when we run a regression model?

```
mod1 <- lm(out_q ~ pred1 + pred2 + pred3, data = sim1)

anova(mod1)
```

Analysis of Variance Table

Response: out_q

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
pred1	1	209.8	209.81	0.5976	0.4423
pred2	1	132.1	132.14	0.3763	0.5417
pred3	2	86.5	43.24	0.1231	0.8843
Residuals	65	22821.9	351.11		

- How many observations were used to fit this model?

Summary of our linear model

```
> summary(mod1)

Call:
lm(formula = out_q ~ pred1 + pred2 + pred3, data = sim1)

Residuals:
    Min       1Q   Median       3Q      Max
-39.164 -13.900   2.419  15.541  34.156

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  105.2070    18.6185   5.651 3.82e-07 ***
pred1         -0.8361     1.3010  -0.643   0.523
pred2          0.2611     0.4614   0.566   0.573
pred3Middle   -1.3498     5.6802  -0.238   0.813
pred3High     -2.7443     5.5427  -0.495   0.622
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 18.74 on 65 degrees of freedom
(30 observations deleted due to missingness)
Multiple R-squared:  0.01843,    Adjusted R-squared:  -0.04198
F-statistic: 0.3051 on 4 and 65 DF,  p-value: 0.8736
```

Another way to see this

```
glance(mod1) %>% select(1:6)
```

```
# A tibble: 1 x 6
```

	r.squared	adj.r.squared	sigma	statistic	p.value	df
	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>
1	0.0184	-0.0420	18.7	0.305	0.874	4

```
glance(mod1) %>% select(7:12)
```

```
# A tibble: 1 x 6
```

	logLik	AIC	BIC	deviance	df.residual	nobs
	<dbl>	<dbl>	<dbl>	<dbl>	<int>	<int>
1	-302.	616.	629.	22822.	65	70

How could we have known this would be 70, in advance?

```
sim1 %>% select(out_q, pred1, pred2, pred3) %>%  
  miss_case_table()
```

```
# A tibble: 3 x 3  
  n_miss_in_case n_cases pct_cases  
      <int>      <int>      <dbl>  
1           0        70         70  
2           1        25         25  
3           2         5          5
```


Which observations were not used?

```
summary(mod1)$na.action
```

```
 3  4  5  6 13 16 19 26 27 29 30 34 39 48 51 56 62 66 67 68
 3  4  5  6 13 16 19 26 27 29 30 34 39 48 51 56 62 66 67 68
72 75 81 83 86 89 93 94 96 97
72 75 81 83 86 89 93 94 96 97
attr(,"class")
[1] "omit"
```

- A potentially more useful `na.action` setting in `lm` is `na.exclude` which pads out predicted values and residuals with NAs instead of omitting the 30 observations listed above.

```
lm(out_q ~ pred1 + pred2 + pred3,  
   data = sim1, na.action = na.exclude)
```

Predictions from mod1 with na.omit and na.exclude

```
mod1 <- lm(out_q ~ pred1 + pred2 + pred3, data = sim1)
      ## note: by default na.action = na.omit here
head(predict(mod1))
```

1	2	7	8	9	10
101.85279	103.02874	98.14391	96.57037	101.49208	101.01744

```
mod1_e <- lm(out_q ~ pred1 + pred2 + pred3, data = sim1,
             na.action = na.exclude)
head(predict(mod1_e))
```

1	2	3	4	5	6
101.8528	103.0287	NA	NA	NA	NA

Multiple Imputation: Potential and Pitfalls

Multiple imputation for missing data in epidemiological and clinical research: potential and pitfalls

In this article, we review the reasons why missing data may lead to bias and loss of information in epidemiological and clinical research. We discuss the circumstances in which multiple imputation may help by reducing bias or increasing precision, as well as describing potential pitfalls in its application. Finally, we describe the recent use and reporting of analyses using multiple imputation in general medical journals, and suggest guidelines for the conduct and reporting of such analyses.

- <https://www.bmj.com/content/338/bmj.b2393>

Note: The next 7 slides are derived from Sterne et al.

An Example from Sterne et al.

Consider, for example, a study investigating the association of systolic blood pressure with the risk of subsequent coronary heart disease, in which data on systolic blood pressure are missing for some people.

The probability that systolic blood pressure is missing is likely to:

- decrease with age (doctors are more likely to measure it in older people),
- decrease with increasing body mass index, and
- decrease with history of smoking (doctors are more likely to measure it in people with heart disease risk factors or comorbidities).

If we assume that data are missing at random and that we have systolic blood pressure data on a representative sample of individuals within strata of age, smoking, body mass index, and coronary heart disease, then we can use multiple imputation to estimate the overall association between systolic blood pressure and coronary heart disease.

Missing Data Mechanisms

- **Missing completely at random** There are no systematic differences between the missing values and the observed values.
 - For example, blood pressure measurements may be missing because of breakdown of an automatic sphygmomanometer.
- **Missing at random** Any systematic difference between the missing and observed values can be explained by other observed data.
 - For example, missing BP measurements may be lower than measured BPs but only because younger people more often have a missing BP.
- **Missing not at random** Even after the observed data are taken into account, systematic differences remain between the missing values and the observed values.
 - For example, people with high BP may be more likely to have headaches that cause them to miss clinic appointments.

“Missing at random” is an **assumption** that justifies the analysis, and is not a property of the data.

Trouble: Data missing not at random

Sometimes, it is impossible to account for systematic differences between missing and observed values using the available data.

- In such (MNAR) cases, multiple imputation may give misleading results.
 - Those results can be either more or less misleading than a complete case analysis.
- For example, consider a study investigating predictors of depression.
 - If individuals are more likely to miss appointments because they are depressed on the day of the appointment, then it may be impossible to make the MAR assumption plausible, even if a large number of variables is included in the imputation model.

Where complete cases and multiple imputation analyses give different results, the analyst should attempt to understand why, and this should be reported in publications.

What if the data are MCAR?

If we assume data are MAR, then unbiased and statistically more powerful analyses (compared with analyses based on complete cases) can generally be done by including individuals with incomplete data.

There are circumstances in which analyses of **complete cases** will not lead to bias.

- Missing data in predictor variables do not cause bias in analyses of complete cases if the reasons for the missing data are unrelated to the outcome.
 - In such cases, imputing missing data may lessen the loss of precision and power resulting from exclusion of individuals with incomplete predictor variables but are not required in order to avoid bias.

Stages of Multiple Imputation (1 of 2)

Multiple imputation . . . aims to allow for the uncertainty about the missing data by creating several different plausible imputed data sets and appropriately combining results obtained from each of them.

The first stage is to create multiple copies of the dataset, with the missing values replaced by imputed values.

- The imputation procedure must fully account for all uncertainty in predicting the missing values by injecting appropriate variability into the multiple imputed values; we can never know the true values of the missing data.

Note that single Imputation of missing values usually causes standard errors to be too small, since it fails to account for the fact that we are uncertain about the missing values.

Stages of Multiple Imputation (2 of 2)

The second stage is to use standard statistical methods to fit the model of interest to each of the imputed datasets.

- Estimated associations in each of the imputed datasets will differ because of the variation introduced in the imputation of the missing values, and they are only useful when averaged together to give overall estimated associations.
- Standard errors are calculated using Rubin's rules, which take account of the variability in results between the imputed datasets, reflecting the uncertainty associated with the missing values.
- Valid inferences are obtained because we are averaging over the distribution of the missing data given the observed data.

Comparing Two Linear Models including Multiple Imputation

Framingham data

```
fram_raw <- read_csv(here("data/framingham.csv")) %>%  
  clean_names()
```

```
dim(fram_raw)
```

```
[1] 4238    17
```

```
n_miss(fram_raw)
```

```
[1] 645
```

- See <https://www.framinghamheartstudy.org/> for more details.

fram_sub Tibble for Today

```
fram_sub <- fram_raw %>%  
  mutate(educ = fct_recode(factor(education),  
                            "Some HS" = "1",  
                            "HS grad" = "2",  
                            "Some Coll" = "3",  
                            "Coll grad" = "4")) %>%  
  mutate(obese = as.numeric(bmi >= 30)) %>%  
  rename(smoker = "current_smoker",  
         sbp = "sys_bp") %>%  
  mutate(subj_id = as.character(subj_id)) %>%  
  select(sbp, educ, smoker, obese, glucose, subj_id)
```

Data Descriptions (variables we'll use today)

The variables describe $n = 4238$ adult subjects who were examined at baseline and then followed for ten years to see if they developed incident coronary heart disease during that time.

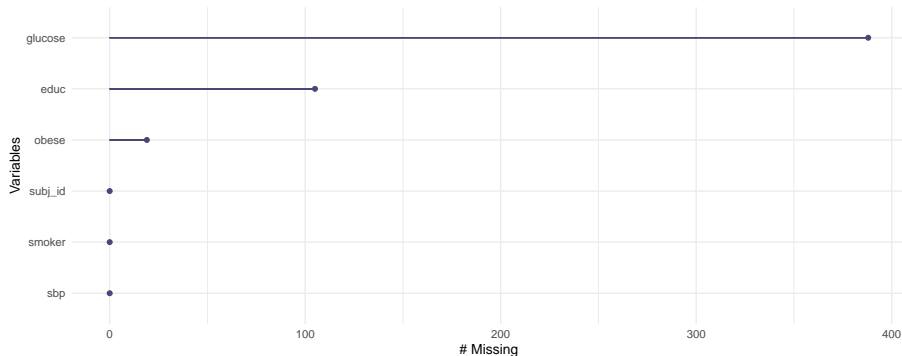
Variable	Description
educ	four-level factor: educational attainment
smoker	1 = current smoker at time of examination, else 0
sbp	systolic blood pressure (mm Hg)
obese	1 if subject's bmi is 30 or higher, else 0
glucose	blood glucose level in mg/dl

Today's Goal

Use linear regression to predict sbp using two different models, in each case accounting for missingness via multiple imputation, where the predictors of interest are glucose, obese, educ, and smoker.

Which variables are missing data?

```
gg_miss_var(fram_sub)
```



Track missingness with shadow

```
fram_sub_sh <- bind_shadow(fram_sub)
```

```
head(fram_sub_sh)
```

```
# A tibble: 6 x 12
```

	sbp	educ	smoker	obese	glucose	subj_id	sbp_NA	educ_NA
	<dbl>	<fct>	<dbl>	<dbl>	<dbl>	<chr>	<fct>	<fct>
1	106	Coll gr~	0	0	77	1	!NA	!NA
2	121	HS grad	0	0	76	2	!NA	!NA
3	128.	Some HS	1	0	70	3	!NA	!NA
4	150	Some Co~	1	0	103	4	!NA	!NA
5	130	Some Co~	1	0	85	5	!NA	!NA
6	180	HS grad	0	1	99	6	!NA	!NA

```
# ... with 4 more variables: smoker_NA <fct>,
```

```
# obese_NA <fct>, glucose_NA <fct>, subj_id_NA <fct>
```


Our Two Models

Model 2: predict sbp using glucose and obese.

Model 4: predict sbp using glucose, obese, educ, and smoker.

Model 2 (CC): Two-predictor model for sbp

Suppose we ignore the missingness and just run the model on the data with complete information on sbp, glucose and obese.

```
m2_cc <- fram_sub_sh %$% lm(sbp ~ glucose + obese)
```

```
tidy(m2_cc, conf.int = TRUE) %>% select(-statistic) %>%  
  kable(digits = 3)
```

term	estimate	std.error	p.value	conf.low	conf.high
(Intercept)	121.671	1.244	0	119.232	124.110
glucose	0.111	0.015	0	0.082	0.139
obese	13.532	1.045	0	11.484	15.580

Edited Summary of Model 2 (CC)

```
summary(m2_cc)    ## we'll just look at the bottom
```

```
Residual standard error: 21.42 on 3833 degrees of freedom  
(402 observations deleted due to missingness)  
Multiple R-squared:  0.05857,    Adjusted R-squared:  0.05808  
F-statistic: 119.2 on 2 and 3833 DF,  p-value: < 2.2e-16
```

```
glance(m2_cc) %>%  
  select(nobs, r.squared, adj.r.squared, AIC, BIC) %>%  
  kable(digits = c(0, 4, 4, 0, 0))
```

nobs	r.squared	adj.r.squared	AIC	BIC
3836	0.0586	0.0581	34401	34426

Model 4 (CC): Four-predictor model for sbp

```
m4_cc <- fram_sub_sh %$%  
  lm(sbp ~ glucose + obese + smoker + educ)  
  
tidy(m4_cc, conf.int = TRUE) %>% select(-statistic) %>%  
  kable(digits = 3)
```

term	estimate	std.error	p.value	conf.low	conf.high
(Intercept)	127.107	1.388	0	124.385	129.829
glucose	0.106	0.015	0	0.078	0.135
obese	12.304	1.066	0	10.213	14.395
smoker	-4.704	0.699	0	-6.075	-3.332
educHS grad	-3.698	0.833	0	-5.332	-2.065
educSome Coll	-4.724	1.010	0	-6.704	-2.744
educColl grad	-5.954	1.158	0	-8.225	-3.683

Edited Summary of Model 4 (CC)

```
summary(m4_cc) ## we'll just look at the bottom
```

```
Residual standard error: 21.2 on 3733 degrees of freedom  
(498 observations deleted due to missingness)  
Multiple R-squared: 0.08257, Adjusted R-squared: 0.0811  
F-statistic: 56 on 6 and 3733 DF, p-value: < 2.2e-16
```

```
glance(m4_cc) %>%  
  select(nobs, r.squared, adj.r.squared, AIC, BIC) %>%  
  kable(digits = c(0, 4, 4, 0, 0))
```

nobs	r.squared	adj.r.squared	AIC	BIC
3740	0.0826	0.0811	33466	33516

Variables used in our models 2 and 4

```
miss_var_summary(fram_sub)
```

```
# A tibble: 6 x 3  
  variable n_miss pct_miss  
  <chr>      <int>    <dbl>  
1 glucose    388     9.16  
2 educ      105     2.48  
3 obese      19     0.448  
4 sbp         0      0  
5 smoker      0      0  
6 subj_id     0      0
```

- Are we missing data on our outcome for these models?

Create multiple imputations for this subset

How many subjects have complete / missing data that affect this model?

```
pct_complete_case(fram_sub)
```

```
[1] 88.24917
```

```
pct_miss_case(fram_sub)
```

```
[1] 11.75083
```

Let's create 15 imputed data sets.

```
set.seed(431431)
```

```
fram_mice24 <- mice(fram_sub, m = 15, printFlag = FALSE)
```

Warning: Number of logged events: 1

- Using `printFlag = FALSE` eliminates a lot of unnecessary (and not particularly informative) output here.

Summary Information about Imputation Process

```
summary(fram_mice24)
```

```
Class: mids
```

```
Number of multiple imputations: 15
```

```
Imputation methods:
```

sbp	educ	smoker	obese	glucose	subj_id
""	"polyreg"	""	"pmm"	"pmm"	""

```
PredictorMatrix:
```

	sbp	educ	smoker	obese	glucose	subj_id
sbp	0	1	1	1	1	0
educ	1	0	1	1	1	0
smoker	1	1	0	1	1	0
obese	1	1	1	0	1	0
glucose	1	1	1	1	0	0
subj_id	1	1	1	1	1	0

```
Number of logged events: 1
```

it	im	dep	meth	out
----	----	-----	------	-----

Options within `mice` for imputation approaches

Default methods include:

- `pmm` predictive mean matching (default choice for quantitative variables)
- `logreg` logistic regression (default for binary categorical variables)
- `polyreg` polytomous logistic regression (for nominal multi-categorical variables)
- `polr` proportional odds logistic regression (for ordinal categories)

but there are `cart` methods and many others available, too.

What should we include in an imputation model?

- ❶ If things you are imputing are not Normally distributed, this can pose special challenges, and either a transformation or choosing an imputation method which is robust to these concerns is helpful.
- ❷ Include the outcome when imputing predictors. It causes you to conclude the relationship is weaker than it actually is, if you don't.
- ❸ The MAR assumption may only be reasonable when a certain variable is included in the model.
 - As a result, it's usually a good idea to include as wide a range of variables in imputation models as possible. The concerns we'd have about parsimony in outcome models don't apply here.

Store one (or more) of the imputed data sets

This will store the fifth imputed data set in `imp_5`.

```
imp_5 <- complete(fram_mice24, 5) %>% tibble()
```

```
dim(imp_5)
```

```
[1] 4238    6
```

```
n_miss(imp_5)
```

```
[1] 0
```

Run Model 2 on each imputed data frame

```
m2_mods <- with(fram_mice24, lm(sbp ~ glucose + obese))
```

```
> summary(m2_mods)
# A tibble: 45 x 6
  term          estimate std.error statistic  p.value  nobs
  <chr>         <dbl>      <dbl>    <dbl>    <dbl>   <int>
1 (Intercept)  122.        1.15     106.     0      4238
2 glucose       0.104    0.0134     7.74 1.27e-14 4238
3 obese        13.3     0.984     13.5 5.93e-41 4238
4 (Intercept)  121.        1.19     102.     0      4238
5 glucose       0.112    0.0139     8.03 1.23e-15 4238
6 obese        13.4     0.986     13.6 4.48e-41 4238
7 (Intercept)  121.        1.15     105.     0      4238
8 glucose       0.120    0.0135     8.93 6.49e-19 4238
9 obese        13.1     0.986     13.3 1.89e-39 4238
10 (Intercept) 121.        1.20     101.     0      4238
# ... with 35 more rows
```

- 3 coefficients in each model, times 15 imputations = 45 rows.

More detailed regression results?

Consider working with the analysis done on the 4th imputed data set (of the 15 created)...

```
m2_a4 <- m2_mods$analyses[[4]]  
tidy(m2_a4) %>% kable(digits = 3)
```

term	estimate	std.error	statistic	p.value
(Intercept)	121.052	1.200	100.861	0
glucose	0.117	0.014	8.322	0
obese	13.243	0.985	13.441	0

Pool Results across the 15 imputations

```
m2_pool <- pool(m2_mods)
summary(m2_pool, conf.int = TRUE, conf.level = 0.95)
```

	term	estimate	std.error	statistic	df
1	(Intercept)	121.1962463	1.28910584	94.015745	408.3743
2	glucose	0.1153255	0.01529575	7.539707	332.3981
3	obese	13.2902350	0.99510223	13.355648	3704.3661
	p.value	2.5 %	97.5 %		
1	0.000000e+00	118.66213491	123.7303577		
2	4.527489e-13	0.08523679	0.1454141		
3	0.000000e+00	11.33923297	15.2412370		

Model 2 (Complete Cases vs. Multiple Imputation)

```
tidy(m2_cc, conf.int = TRUE) %>% kable(digits = 3)
```

term	estimate	std.error	statistic	p.value	conf.low	conf.high
(Intercept)	121.671	1.244	97.792	0	119.232	124.110
glucose	0.111	0.015	7.577	0	0.082	0.139
obese	13.532	1.045	12.954	0	11.484	15.580

```
summary(m2_pool, conf.int = TRUE, conf.level = 0.95) %>%  
  select(-df) %>% kable(digits = 2)
```

term	estimate	std.error	statistic	p.value	2.5 %	97.5 %
(Intercept)	121.20	1.29	94.02	0	118.66	123.73
glucose	0.12	0.02	7.54	0	0.09	0.15
obese	13.29	1.00	13.36	0	11.34	15.24

More Details on Multiple Imputation Modeling

m2_pool

```
> m2_pool
Class: mipo      m = 15
      term      m      estimate      ubar      b      t      dfcom      df
1 (Intercept) 15 121.1962463 1.3726312457 2.710900e-01 1.66179387 4235 408.3743
2      glucose 15   0.1153255 0.0001883469 4.276229e-05 0.00023396 4235 332.3981
3       obese 15  13.2902350 0.9703219834 1.866232e-02 0.99022845 4235 3704.3661
      riv      lambda      fmi
1 0.21066301 0.17400631 0.17802209
2 0.24217612 0.19496118 0.19976168
3 0.02051532 0.02010291 0.02063153
```

Definitions of all of these terms are available in the `mipo` help file.

- `riv` = relative increase in variance attributable to nonresponse
- `fmi` = fraction of missing information due to nonresponse

Model 4 run on each imputed data frame

```
m4_mods <- with(fram_mice24, lm(sbp ~ glucose +  
                                obese + smoker + educ))
```

```
> summary(m4_mods)
# A tibble: 105 x 6
```

	term <chr>	estimate <dbl>	std.error <dbl>	statistic <dbl>	p.value <dbl>	nobs <int>
1	(Intercept)	128.	1.27	100.	0	4238
2	glucose	0.0969	0.0133	7.28	3.93e-13	4238
3	obese	11.8	0.985	12.0	1.57e-32	4238
4	smoker	-4.46	0.656	-6.80	1.23e-11	4238
5	educHS grad	-3.68	0.780	-4.72	2.40e-6	4238
6	educSome Coll	-5.31	0.947	-5.61	2.13e-8	4238
7	educColl grad	-6.16	1.09	-5.65	1.71e-8	4238
8	(Intercept)	127.	1.31	97.2	0	4238
9	glucose	0.104	0.0138	7.57	4.70e-14	4238
10	obese	11.8	0.986	12.0	1.12e-32	4238

```
# ... with 95 more rows
```

Pool Results across the five imputations

```
m4_pool <- pool(m4_mods)
```

```
summary(m4_pool, conf.int = TRUE, conf.level = 0.95) %>%  
  select(-df) %>% kable(digits = 2)
```

term	estimate	std.error	statistic	p.value	2.5 %	97.5 %
(Intercept)	126.91	1.41	90.25	0	124.14	129.67
glucose	0.11	0.02	7.16	0	0.08	0.14
obese	11.77	1.00	11.82	0	9.82	13.73
smoker	-4.43	0.66	-6.75	0	-5.72	-3.15
educHS grad	-3.66	0.79	-4.62	0	-5.21	-2.11
educSome Coll	-5.25	0.96	-5.47	0	-7.13	-3.37
educColl grad	-6.13	1.10	-5.58	0	-8.28	-3.97

Complete Cases Result (Model 4)

```
tidy(m4_cc, conf.int = TRUE) %>% select(-statistic) %>%  
  kable(digits = 3)
```

term	estimate	std.error	p.value	conf.low	conf.high
(Intercept)	127.107	1.388	0	124.385	129.829
glucose	0.106	0.015	0	0.078	0.135
obese	12.304	1.066	0	10.213	14.395
smoker	-4.704	0.699	0	-6.075	-3.332
educHS grad	-3.698	0.833	0	-5.332	-2.065
educSome Coll	-4.724	1.010	0	-6.704	-2.744
educColl grad	-5.954	1.158	0	-8.225	-3.683

Multiple Imputation Result (Model 4)

```
summary(m4_pool, conf.int = TRUE) %>%  
  select(-statistic, -df) %>% kable(digits = 3)
```

term	estimate	std.error	p.value	2.5 %	97.5 %
(Intercept)	126.905	1.406	0	124.143	129.668
glucose	0.107	0.015	0	0.078	0.137
obese	11.774	0.996	0	9.820	13.728
smoker	-4.435	0.657	0	-5.723	-3.147
educHS grad	-3.660	0.792	0	-5.214	-2.106
educSome Coll	-5.251	0.960	0	-7.133	-3.369
educColl grad	-6.126	1.097	0	-8.276	-3.975

More Details on Multiple Imputation Modeling

m4_pool

```
> m4_pool
Class: mipo      m = 15
      term m      estimate      ubar      b      t dfcom      df
1  (Intercept) 15 126.9054997 1.6777883057 2.806862e-01 1.9771868946 4231 521.7802
2    glucose 15  0.1074957 0.0001848288 3.778876e-05 0.0002251368 4231 387.9514
3    obese 15 11.7738693 0.9713228490 2.028081e-02 0.9929557100 4231 3628.0260
4    smoker 15 -4.4347370 0.4295295322 1.829583e-03 0.4314810871 4231 4184.1354
5  educHS grad 15 -3.6597613 0.6073110889 1.939744e-02 0.6280016922 4231 3105.0759
6 educSome Coll 15 -5.2509672 0.8952687033 2.432866e-02 0.9212192783 4231 3333.3562
7 educColl grad 15 -6.1257072 1.1823783257 1.974515e-02 1.2034398227 4231 3808.7663
      riv      lambda      fmi
1 0.178448370 0.151426549 0.154660564
2 0.218083002 0.179037883 0.183237700
3 0.022271546 0.021786330 0.022325139
4 0.004543471 0.004522921 0.004998414
5 0.034069200 0.032946732 0.033569016
6 0.028986353 0.028169813 0.028752383
7 0.017812824 0.017501080 0.018016589
```

Estimate R^2

```
pool.r.squared(m2_mods)
```

	est	lo 95	hi 95	fmi
R^2	0.05944041	0.04607775	0.07426549	NaN

```
pool.r.squared(m4_mods)
```

	est	lo 95	hi 95	fmi
R^2	0.08146963	0.06617182	0.09805258	NaN

- Can also calculate adjusted R^2 by using `pool.r.squared(m2_mods, adjusted = TRUE)`. See next slide.

Estimates of adjusted R^2

```
pool.r.squared(m2_mods, adjusted = TRUE)
```

	est	lo 95	hi 95	fmi
adj R ²	0.05899617	0.04567921	0.07377905	NaN

```
pool.r.squared(m4_mods, adjusted = TRUE)
```

	est	lo 95	hi 95	fmi
adj R ²	0.080167	0.06497656	0.09665148	NaN

Tests of Nested Fits after imputation

The models must be nested (same outcome, one set of predictors is a subset of the other) for this to be appropriate.

```
fit4 <- with(fram_mice24,  
             expr = lm(sbp ~ glucose + obese + smoker + educ))  
fit2 <- with(fram_mice24,  
             expr = lm(sbp ~ glucose + obese))
```


Comparing Model 4 to Model 2 fits

We'll use the Wald test after a linear regression fit.

```
D1(fit4, fit2)
```

	test	statistic	df1	df2	dfcom	p.value
1	~~ 2	24.85276	4	4093.675	4231	2.322448e-20
	riv					
		0.01982057				

Could also use a likelihood ratio test.

```
D3(fit4, fit2)
```

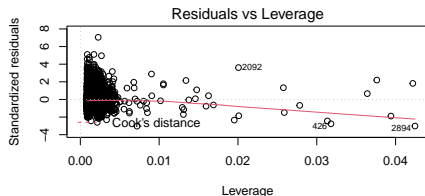
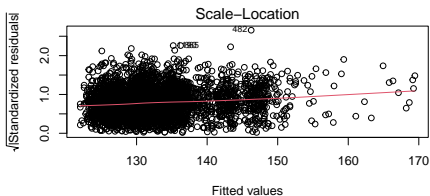
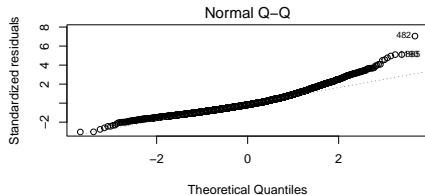
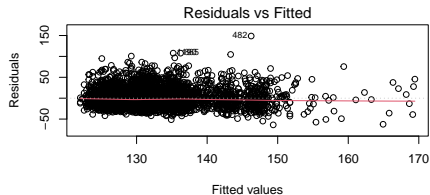
	test	statistic	df1	df2	dfcom	p.value	riv
1	~~ 2	24.76194	4	276830.7	4231	0	0.01339975

Residual Plots for mod4 with 6th imputed data set

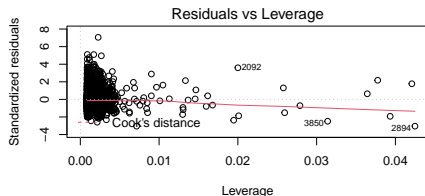
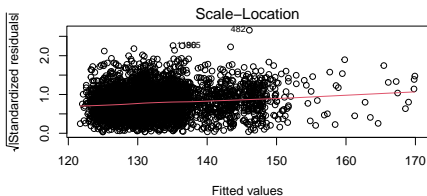
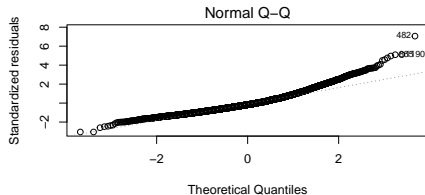
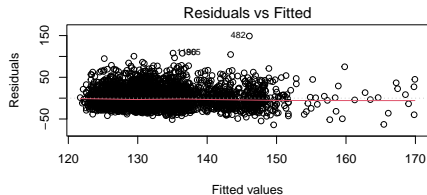
```
par(mfrow = c(2,2))  
plot(m4_mods$analyses[[6]])  
par(mfrow = c(1,1))
```

- See the next slide for the results.

Residual Plots for mod4 using imputation 6



Residual Plots for mod4 using imputation 7



Guidelines for Reporting

Guidelines for reporting, I (Sterne et al.)

How should we report on analyses potentially affected by missing data?

- Report the number of missing values for each variable of interest, or the number of cases with complete data for each important component of the analysis. Give reasons for missing values if possible, and indicate how many individuals were excluded because of missing data when reporting the flow of participants through the study. If possible, describe reasons for missing data in terms of other variables (rather than just reporting a universal reason such as treatment failure.)
- Clarify whether there are important differences between individuals with complete and incomplete data, for example, by providing a table comparing the distributions of key exposure and outcome variables in these different groups
- Describe the type of analysis used to account for missing data (eg, multiple imputation), and the assumptions that were made (eg, missing at random)

Guidelines for reporting, II (Sterne et al.)

How should we report on analyses that involve multiple imputation?

- Provide details of the imputation modeling (software used, key settings, number of imputed datasets, variables included in imputation procedure, etc.)
- If a large fraction of the data is imputed, compare observed and imputed values.
- Where possible, provide results from analyses restricted to complete cases, for comparison with results based on multiple imputation. If there are important differences between the results, suggest explanations.
- It is also desirable to investigate the robustness of key inferences to possible departures from the missing at random assumption, by assuming a range of missing not at random mechanisms in sensitivity analyses.