#### 431 Class 11

thomaselove.github.io/431

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#### Today's Agenda

- Developing Four Models for sbp using dbp (and insurance)
- 2 Fitting a Bayesian Linear Model with default priors (m2)
- Including Insurance without (m3) and with (m4) interaction with dbp in linear models
- Visualizing Categorical Data
- Sessing Association in Cross-Tabulations

#### Today's Packages

```
library(broom)
library(equatiomatic) # new today
library(ggrepel) # sort of new today
library(glue) # sort of new today
library(janitor)
library(knitr)
library(magrittr)
library(patchwork)
library(rstanarm) # special today
library(tidyverse)
theme set(theme bw())
```

#### Today's Data

Again, we'll use an R data set (.Rds) to import the dm1000 data.

```
dm1000 <- read_rds("data/dm_1000.Rds")</pre>
```

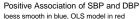
Then, we'll again partition the dm1000 cases with complete BP data into training and test samples.

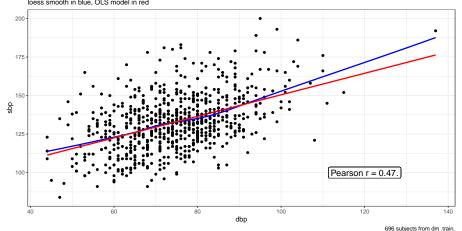
```
dm994 <- dm1000 %>% filter(complete.cases(sbp, dbp)) %>%
    select(subject, sbp, dbp, insurance)

set.seed(4312021) # for replicating the sampling later
dm_train <- dm994 %>% sample_frac(0.7)
dm_test <- dm994 %>% anti_join(dm_train, by = "subject")
```

Back to Regression: Can dbp predict sbp?

## Plotting sbp vs. dbp (training set)





## Model m1 for sbp using dbp (training set)

```
m1_train <- lm(sbp ~ dbp, data = dm_train)

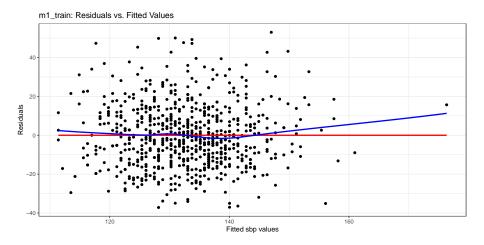
tidy(m1_train, conf.int = TRUE, conf.level = 0.90) %>%
    select(term, estimate, conf.low, conf.high) %>% kable()
```

| term        | estimate   | conf.low   | conf.high |
|-------------|------------|------------|-----------|
| (Intercept) | 80.6798905 | 74.4662421 | 86.893539 |
| dbp         | 0.6982168  | 0.6160396  | 0.780394  |

| nobs | r.squared | adj.r.squared | sigma    | AIC      | BIC      |
|------|-----------|---------------|----------|----------|----------|
| 696  | 0.2200811 | 0.2189573     | 16.11605 | 5848.663 | 5862.299 |

# m1\_train: Residuals vs. Predicted (Fitted) Values

m1\_train\_aug <- augment(m1\_train, data = dm\_train)</pre>



## Use model m1\_train to predict SBP in dm\_test

```
m1_test_aug <- augment(m1_train, newdata = dm_test)
mosaic::favstats(~ abs(.resid), data = m1_test_aug) %>%
    select(n, min, median, max, mean, sd) %>% kable(digits = 3)
```

| -  | n | min   | median | max    | mean   | sd     |
|----|---|-------|--------|--------|--------|--------|
| 29 | 8 | 0.028 | 9.822  | 71.066 | 12.139 | 10.177 |

```
sqrt(mean(m1_test_aug$.resid^2))
```

#### [1] 15.83017

| Summary                                    | Model m1 |
|--|----------|
| Mean Absolute Prediction Error             | 12.139   |
| Maximum Absolute Prediction Error          | 71.066   |
| Root Mean Squared Prediction Error (RMSPE) | 15.83    |

## Is this the only linear model R can fit to these data?

Nope.

```
library(rstanarm)
m2_train <- stan_glm(sbp ~ dbp, data = dm_train)</pre>
SAMPLING FOR MODEL 'continuous' NOW (CHAIN 1).
Chain 1:
Chain 1: Gradient evaluation took 0 seconds
Chain 1: 1000 transitions using 10 leapfrog steps per transit:
Chain 1: Adjust your expectations accordingly!
Chain 1:
Chain 1:
                                           (Warmup)
Chain 1: Iteration:
                         1 / 2000 [ 0%]
Chain 1: Iteration: 200 / 2000 [ 10%]
                                           (Warmup)
Chain 1: Iteration: 400 / 2000 [ 20%]
                                           (Warmup)
Chain 1: Iteration: 600 / 2000 [ 30%]
                                           (Warmup)
                                           (Warmup)
                      800 / 2000
                                    40%]
Chain 1: Iteration:
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```

#### **Default Prior Details**

#### prior\_summary(m2\_train)

```
prior_summary(m2_train)
Priors for model 'm2_train'
Intercept (after predictors centered)
  Specified prior:
    ~ normal(location = 133. scale = 2.5)
  Adjusted prior:
    ~ normal(location = 133. scale = 46)
Coefficients
  Specified prior:
    ~ normal(location = 0, scale = 2.5)
  Adjusted prior:
   ~ normal(location = 0, scale = 3.7)
Auxiliary (sigma)
  Specified prior:
    ~ exponential(rate = 1)
  Adjusted prior:
    ~ exponential(rate = 0.055)
See help('prior_summary.stanreg') for more details
```

## Bayesian fitted linear model for our sbp data

```
print(m2 train)
stan_glm
family: gaussian [identity]
formula: sbp ~ dbp
 observations: 696
predictors: 2
          Median MAD SD
(Intercept) 80.6 4.0
     0.7 0.1
dbp
Auxiliary parameter(s):
     Median MAD SD
sigma 16.2
            0.4
```

# Is the Bayesian model (with default prior) very different from our 1m in this situation?

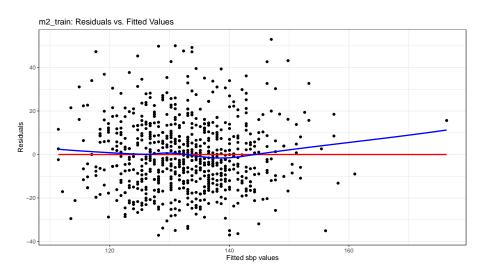
```
broom::tidy(m1_train) # fit with lm
# A tibble: 2 \times 5
 term estimate std.error statistic p.value
 <chr> <dbl> <dbl> <dbl> <dbl>
1 (Intercept) 80.7 3.77 21.4 2.47e-78
      0.698 0.0499 14.0 2.23e-39
2 dbp
broom.mixed::tidy(m2_train) # stan_glm with default priors
# A tibble: 2 x 3
 term estimate std.error
 <chr> <dbl> <dbl>
1 (Intercept) 80.6 3.97
      0.699 0.0522
2 dbp
```

#### Obtaining fits and residuals from Model m2

#### In the model training sample

#### In the model test sample

## Residuals vs. Fitted Values from Model m2 (training)



# Out-of-Sample (Test Set) Error Summaries (m2)

```
mosaic::favstats(~ abs(.resid), data = m2_test_aug) %>%
  select(n, min, median, max, mean, sd) %>% kable(digits = 3)
```

| _ | n   | min   | median | max    | mean  | sd     |
|---|-----|-------|--------|--------|-------|--------|
|   | 298 | 0.036 | 9.824  | 71.059 | 12.14 | 10.177 |

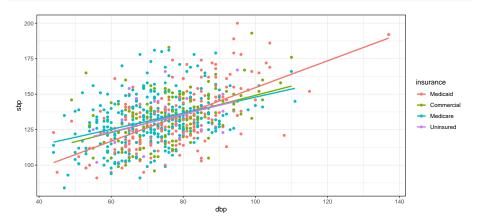
sqrt(mean(m2\_test\_aug\$.resid^2))

#### [1] 15.83019

| Test Set Error Summary             | OLS model m1 | Bayes model m2 |
|------------------------------------|--------------|----------------|
| Mean Absolute Prediction Error     | 12.139       | 12.14          |
| Maximum Absolute Prediction Error  | 71.066       | 71.059         |
| Root Mean Squared Prediction Error | 15.83        | 15.83          |

What if we add another predictor? (Insurance)

#### Plotting sbp vs. dbp and insurance



#### Two possible models

```
m3_train <- lm(sbp ~ dbp + insurance, data = dm_train)
m4_train <- lm(sbp ~ dbp * insurance, data = dm_train)</pre>
```

- What is the difference between m3 and m4?
  - Model m3 will allow the intercept term of the sbp-dbp relationship to vary depending on insurance.
  - Model m4 will allow both the slope and intercept of the sbp-dbp relationship to vary depending on insurance.

$$\begin{split} \widehat{\mathsf{sbp}} &= 77.58 + 0.72 (\mathsf{dbp}) \; + \\ &\quad 1.11 (\mathsf{insurance}_{\mathsf{Commercial}}) + 2.73 (\mathsf{insurance}_{\mathsf{Medicare}}) \; + \\ &\quad 1.16 (\mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

• Predicted sbp by m3 for a Commercial subject?

$$\begin{split} \widehat{\mathsf{sbp}} &= 77.58 + 0.72 (\mathsf{dbp}) \; + \\ &\quad 1.11 (\mathsf{insurance}_{\mathsf{Commercial}}) + 2.73 (\mathsf{insurance}_{\mathsf{Medicare}}) \; + \\ &\quad 1.16 (\mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- Predicted sbp by m3 for a Commercial subject?
- sbp = 77.58 + 0.72\*dbp + 1.11(1) + 2.73(0) + 1.16(0)

$$\begin{split} \widehat{\mathsf{sbp}} &= 77.58 + 0.72 (\mathsf{dbp}) \; + \\ &\quad 1.11 (\mathsf{insurance}_{\mathsf{Commercial}}) + 2.73 (\mathsf{insurance}_{\mathsf{Medicare}}) \; + \\ &\quad 1.16 (\mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- Predicted sbp by m3 for a Commercial subject?
- sbp = 77.58 + 0.72\*dbp + 1.11(1) + 2.73(0) + 1.16(0)
- sbp = 78.69 + 0.72\*dbp

$$\begin{split} \widehat{\mathsf{sbp}} &= 77.58 + 0.72 (\mathsf{dbp}) \; + \\ &\quad 1.11 (\mathsf{insurance}_{\mathsf{Commercial}}) + 2.73 (\mathsf{insurance}_{\mathsf{Medicare}}) \; + \\ &\quad 1.16 (\mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- Predicted sbp by m3 for a Commercial subject?
- sbp = 77.58 + 0.72\*dbp + 1.11(1) + 2.73(0) + 1.16(0)
- sbp = 78.69 + 0.72\*dbp
- $\bullet$  For a Medicaid subject, m3 predicts sbp = 77.58 + 0.72 dbp

$$\begin{split} \widehat{\mathsf{sbp}} &= 77.58 + 0.72 (\mathsf{dbp}) \; + \\ &\quad 1.11 (\mathsf{insurance}_{\mathsf{Commercial}}) + 2.73 (\mathsf{insurance}_{\mathsf{Medicare}}) \; + \\ &\quad 1.16 (\mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- Predicted sbp by m3 for a Commercial subject?
- sbp = 77.58 + 0.72\*dbp + 1.11(1) + 2.73(0) + 1.16(0)
- sbp = 78.69 + 0.72\*dbp
- ullet For a Medicaid subject, m3 predicts  ${ t sbp}=77.58+0.72~{ t dbp}$
- ullet For a Medicare subject, m3 predicts sbp = 80.31 + 0.72 dbp

$$\begin{split} \widehat{\mathsf{sbp}} &= 77.58 + 0.72 (\mathsf{dbp}) \; + \\ &\quad 1.11 (\mathsf{insurance}_{\mathsf{Commercial}}) + 2.73 (\mathsf{insurance}_{\mathsf{Medicare}}) \; + \\ &\quad 1.16 (\mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- Predicted sbp by m3 for a Commercial subject?
- sbp = 77.58 + 0.72\*dbp + 1.11(1) + 2.73(0) + 1.16(0)
- sbp = 78.69 + 0.72\*dbp
- ullet For a Medicaid subject, m3 predicts  ${ t sbp}=77.58+0.72~{ t dbp}$
- ullet For a Medicare subject, m3 predicts  ${\tt sbp} = 80.31 + 0.72 \; {\tt dbp}$
- For an uninsured subject, m3 predicts sbp = 78.74 + 0.72 dbp

$$\begin{split} \widehat{\mathsf{sbp}} &= 77.58 + 0.72 (\mathsf{dbp}) \; + \\ &\quad 1.11 (\mathsf{insurance}_{\mathsf{Commercial}}) + 2.73 (\mathsf{insurance}_{\mathsf{Medicare}}) \; + \\ &\quad 1.16 (\mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- Predicted sbp by m3 for a Commercial subject?
- sbp = 77.58 + 0.72\*dbp + 1.11(1) + 2.73(0) + 1.16(0)
- sbp = 78.69 + 0.72\*dbp
- ullet For a Medicaid subject, m3 predicts  ${ t sbp}=77.58+0.72~{ t dbp}$
- ullet For a Medicare subject, m3 predicts sbp = 80.31 + 0.72 dbp
- For an uninsured subject, m3 predicts sbp = 78.74 + 0.72 dbp
- Note: only the intercept term varies by insurance in m3.

$$\begin{split} \widehat{\mathsf{sbp}} &= 60.26 + 0.94(\mathsf{dbp}) \, + \\ &\quad 23.54(\mathsf{insurance}_{\mathsf{Commercial}}) + 31.04(\mathsf{insurance}_{\mathsf{Medicare}}) \, + \\ &\quad 25.78(\mathsf{insurance}_{\mathsf{Uninsured}}) - 0.29(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Commercial}}) \, - \\ &\quad 0.38(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Medicare}}) - 0.32(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

• m4 predicts, for a Commercial subject...

$$\begin{split} \widehat{\mathsf{sbp}} &= 60.26 + 0.94(\mathsf{dbp}) + \\ &\quad 23.54(\mathsf{insurance}_{\mathsf{Commercial}}) + 31.04(\mathsf{insurance}_{\mathsf{Medicare}}) + \\ &\quad 25.78(\mathsf{insurance}_{\mathsf{Uninsured}}) - 0.29(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Commercial}}) - \\ &\quad 0.38(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Medicare}}) - 0.32(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- m4 predicts, for a Commercial subject...
- sbp = 60.26 + 0.94 \* dbp + 23.54 (1) + 31.04 (0) + 25.78 (0) 0.29 (dbp \* 1) 0.38 (dbp \* 0) 0.32 (dbp \* 0)

$$\begin{split} \widehat{\mathsf{sbp}} &= 60.26 + 0.94(\mathsf{dbp}) + \\ &\quad 23.54(\mathsf{insurance}_{\mathsf{Commercial}}) + 31.04(\mathsf{insurance}_{\mathsf{Medicare}}) + \\ &\quad 25.78(\mathsf{insurance}_{\mathsf{Uninsured}}) - 0.29(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Commercial}}) - \\ &\quad 0.38(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Medicare}}) - 0.32(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- m4 predicts, for a Commercial subject. . .
- sbp = 60.26 + 0.94 \* dbp + 23.54 (1) + 31.04 (0) + 25.78 (0) 0.29 (dbp \* 1) 0.38 (dbp \* 0) 0.32 (dbp \* 0)
- sbp = (60.26 + 23.54) + (0.94 0.29) \* dbp

$$\begin{split} \widehat{\mathsf{sbp}} &= 60.26 + 0.94(\mathsf{dbp}) + \\ &\quad 23.54(\mathsf{insurance}_{\mathsf{Commercial}}) + 31.04(\mathsf{insurance}_{\mathsf{Medicare}}) + \\ &\quad 25.78(\mathsf{insurance}_{\mathsf{Uninsured}}) - 0.29(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Commercial}}) - \\ &\quad 0.38(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Medicare}}) - 0.32(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- m4 predicts, for a Commercial subject. . .
- sbp = 60.26 + 0.94 \* dbp + 23.54 (1) + 31.04 (0) + 25.78 (0) 0.29 (dbp \* 1) 0.38 (dbp \* 0) 0.32 (dbp \* 0)
- sbp = (60.26 + 23.54) + (0.94 0.29) \* dbp
- sbp = 83.80 0.65 dbp for Commercial subjects

$$\begin{split} \widehat{\mathsf{sbp}} &= 60.26 + 0.94(\mathsf{dbp}) \, + \\ &\quad 23.54(\mathsf{insurance}_{\mathsf{Commercial}}) + 31.04(\mathsf{insurance}_{\mathsf{Medicare}}) \, + \\ &\quad 25.78(\mathsf{insurance}_{\mathsf{Uninsured}}) - 0.29(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Commercial}}) \, - \\ &\quad 0.38(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Medicare}}) - 0.32(\mathsf{dbp} \times \mathsf{insurance}_{\mathsf{Uninsured}}) \end{split}$$

- $\bullet$  For Medicaid subjects, sbp = 60.26 + 0.94 \* dbp
- ullet For Medicare subjects, sbp = 91.30 + 0.56 \* dbp
- $\bullet$  For the uninsured, sbp = 86.04 + 0.62 \* dbp
- So both the slope and the intercept are changing in m4

## **Training Sample Fit Quality**

#### Model m3 (no interaction)

```
glance(m3_train) %>%
  select(r.squared, adj.r.squared, sigma, AIC, BIC) %>%
  kable(digits = c(3, 3, 1, 1, 1))
```

| r.squared | adj.r.squared | sigma | AIC    | BIC    |
|-----------|---------------|-------|--------|--------|
| 0.224     | 0.22          | 16.1  | 5851.1 | 5878.4 |

#### Model m4 (with dbp-insurance interaction)

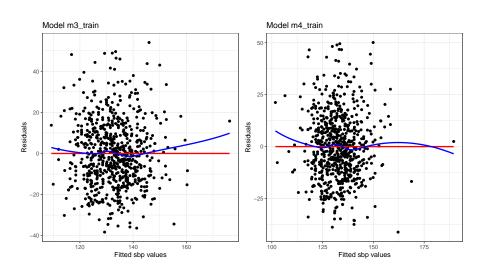
| r.squared | adj.r.squared | sigma | AIC  | BIC    |
|-----------|---------------|-------|------|--------|
| 0.236     | 0.229         | 16    | 5846 | 5886.9 |

#### Augmenting and Testing Models m3 and m4

```
m3_train_aug <- augment(m3_train, data = dm_train)
m3_test_aug <- augment(m3_train, newdata = dm_test)

m4_train_aug <- augment(m4_train, data = dm_train)
m4_test_aug <- augment(m4_train, newdata = dm_test)</pre>
```

#### Residuals vs. Fitted Values Plots



## Comparing performance on the training data

| modname | r2    | adj_r2 | sigma | AIC    | BIC    |
|---------|-------|--------|-------|--------|--------|
| m1      | 0.220 | 0.219  | 16.12 | 5848.7 | 5862.3 |
| m2      | NA    | NA     | 16.15 | NA     | NA     |
| m3      | 0.224 | 0.220  | 16.11 | 5851.1 | 5878.4 |
| m4      | 0.236 | 0.229  | 16.02 | 5846.0 | 5886.9 |

• The glance() function produces different results for a Bayesian stan\_glm() model like m2, so we'll ignore that for now.

## Comparing performance on the test data

Here are some fundamental summaries of absolute prediction error (APE) along with the root mean squared prediction error (RMSPE) for each of our models, in the **testing** sample.

| Summary                            | Mean APE | Max APE | RMSPE |
|------------------------------------|----------|---------|-------|
| m1_train: lm                       | 12.14    | 71.07   | 15.83 |
| m2_train: stan_glm                 | 12.14    | 71.06   | 15.83 |
| m3_train: dbp+insurance            | 12.04    | 72.37   | 15.78 |
| <pre>m4_train: dbp*insurance</pre> | 11.95    | 71.37   | 15.65 |

• Which of these models displays the strongest predictive performance in our test sample?

# Visualizing Categorical Data in dm1000

### 8 Categorical Variables from dm1000

#### Codebook

- subject = ID value (treat as character)
- sex = Female or Male (no missing data)
- insurance = Medicare, Commercial, Medicaid, Uninsured
- eye\_exam = 1 for eye examination in past year, else 0
- **statin** = 1 statin prescription in past year, else 0
- race\_ethnicity = 4 levels (Hispanic or Latinx, Non-Hispanic White, Non-Hispanic Black, Non-Hispanic Asian)
- residence = 2 levels (Suburbs, Cleveland), some NA
- tobacco = 3 levels (Current, Former, Never), some NA

## Using summary()

#### summary(dm\_cat)

```
summary(dm_cat)
 subject
                      insurance
                                     tobacco
                                                   statin
Length: 1000
                  Medicaid :330
                                  Current:274
                                               Min.
                                                      :0.000
Class :character
                  Commercial:196
                                  Never :343
                                               1st Qu.:1.000
                  Medicare :432 Former :367
                                               Median :1.000
Mode :character
                  Uninsured: 42
                                  NA'S
                                         : 16
                                               Mean
                                                      :0.758
                                               3rd Ou.:1.000
                                                      :1.000
                                               Max.
                                      race_ethnicity
                  sex
                                                        residence
  eye_exam
Min.
      :0.000 Female:550
                           Non-Hispanic Black:533
                                                    Suburbs :371
1st Qu.:0.000
               Male :450
                           Hispanic or Latinx: 91
                                                    Cleveland: 601
Median :1.000
                           Non-Hispanic White:356
                                                    NA's
                                                             : 28
                           Non-Hispanic Asian: 20
Mean
      :0.562
3rd Ou.:1.000
Max.
      :1.000
```

## Using taby1 to tabulate a categorical variable

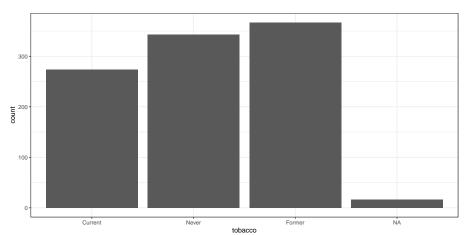
```
dm_cat %>% tabyl(tobacco) %>%
  adorn_pct_formatting() %>%
  adorn_totals()
```

```
tobacco n percent valid_percent Current 274 27.4% 27.8% Never 343 34.3% 34.9% Former 367 36.7% 37.3% <NA> 16 1.6% - Total 1000 - -
```

#### Using count to create a tibble of counts

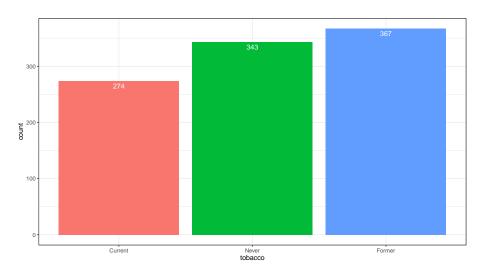
# Using geom\_bar to show a distribution

```
ggplot(dm_cat, aes(x = tobacco)) +
  geom_bar()
```



# Augmenting the geom\_bar result (code)

# Augmenting the geom\_bar result



### Using taby1 to cross-tabulate two variables

```
dm_cat %>% tabyl(insurance, residence) %>%
 adorn totals(where = c("row", "col"))
  insurance Suburbs Cleveland NA Total
                             15
  Medicaid
               114
                         201
                                   330
                75
                         119 2 196
```

601

259 10 432

22 1 42

28

1000

Commercial

Medicare 163

371

Uninsured 19

Total

### Using count to create a tibble of counts

```
      statin residence
      n

      <dbl> <fct>< (int>

      1
      0 Suburbs
      77

      2
      0 Cleveland
      157

      3
      0 <NA>
      8

      4
      1 Suburbs
      294

      5
      1 Cleveland
      444

      6
      1 <NA>
      20
```

# Were suburban residents more likely to have a statin prescription?

```
dm_cat %>%
  filter(complete.cases(statin, residence)) %>%
  tabyl(residence, statin)

residence 0 1
```

Suburbs 77 294 Cleveland 157 444

# Revise the order of the statin levels, add percentages

```
dm_cat %>% filter(complete.cases(statin, residence)) %>%
 mutate(statin = fct_relevel(factor(statin), "1", "0")) %>%
 tabyl(residence, statin)
residence 1 0
   Suburbs 294 77
Cleveland 444 157
dm cat %>% filter(complete.cases(statin, residence)) %>%
 mutate(statin = fct_relevel(factor(statin), "1", "0")) %>%
 tabyl(residence, statin) %>%
 adorn_percentages(denom = "row") %>%
 adorn_pct_formatting()
```

### Create using table instead

```
tab1 <- dm_cat %>%
  filter(complete.cases(statin, residence)) %>%
  mutate(statin = fct_relevel(factor(statin), "1", "0")) %$%
  table(residence, statin)
```

#### Assess 2x2 table (results on next slide)

```
Epi::twoby2(tab1)
```

#### twoby2 results

```
> Epi::twoby2(tab1)
2 by 2 table analysis:
Outcome : 1
Comparing: Suburbs vs. Cleveland
           1 0 P(1) 95% conf. interval
Suburbs 294 77 0.7925 0.7482 0.8307
Cleveland 444 157 0.7388 0.7022 0.7723
                                95% conf. interval
            Relative Risk: 1.0727 0.9996 1.1510
        Sample Odds Ratio: 1.3501 0.9903 1.8407
Conditional MLE Odds Ratio: 1.3497 0.9805 1.8679
   Probability difference: 0.0537 -0.0018 0.1065
            Exact P-value: 0.0638
       Asymptotic P-value: 0.0577
```

### A three-by-four two-way table

```
dm_cat %>% filter(complete.cases(tobacco, insurance)) %>%
  tabyl(tobacco, insurance) %>%
  adorn_totals(where = c("row", "col"))
```

```
tobacco Medicaid Commercial Medicare Uninsured Total
             118
                         44
                                   99
                                             1.3
                                                  274
Current
  Never
             105
                         80
                                  140
                                             18
                                                  343
                                             11
                                                  367
 Former
             103
                         70
                                  183
  Total
             326
                         194
                                  422
                                             42
                                                  984
```

- 3 rows, 4 columns: hence, this is a 3 x 4 table
- It's a two-way table, because we are studying the association of two variables (tobacco and insurance)
- Can we compare the insurance percentages by tobacco group?

## Compare insurance rates by tobacco group

```
dm_cat %>% filter(complete.cases(tobacco, insurance)) %>%
  tabyl(tobacco, insurance) %>%
  adorn_percentages(denominator = "row") %>%
  adorn_totals(where = "col") %>% kable(digits = 3)
```

| tobacco | Medicaid | Commercial | Medicare | Uninsured | Total |
|---------|----------|------------|----------|-----------|-------|
| Current | 0.431    | 0.161      | 0.361    | 0.047     | 1     |
| Never   | 0.306    | 0.233      | 0.408    | 0.052     | 1     |
| Former  | 0.281    | 0.191      | 0.499    | 0.030     | 1     |

- Note that these are actually **proportions** and not percentages.
- Proportions fall between 0 and 1: multiply by 100 for percentages.

# Insurance rates by tobacco group?

```
tab2 <- dm_cat %>%
  filter(complete.cases(tobacco, insurance)) %$%
  table(tobacco, insurance)
```

tab2

```
insurance
tobacco Medicaid Commercial Medicare Uninsured
Current 118 44 99 13
Never 105 80 140 18
Former 103 70 183 11
chisq.test(tab2)
```

Pearson's Chi-squared test

data: tab2

X-squared = 25.592, df = 6, p-value = 0.0002651

#### Using count for three variables

```
dm_cat %>% count(sex, statin, eye_exam)
# A tibble: 8 x 4
        statin eye_exam
  sex
  <fct> <dbl> <dbl> <int>
1 Female
                            68
                       0
2 Female
                            65
3 Female
                           176
4 Female
                           241
5 Male
                       0
                            61
6 Male
                            48
7 Male
                           133
8 Male
                           208
```

#### A three-way table

```
dm_cat %>% tabyl(statin, residence, sex) %>%
adorn_title()
```

```
$Female
```

```
residence
```

```
statin Suburbs Cleveland NA_

0 42 87 4

1 160 245 12
```

#### \$Male

#### residence

```
statin Suburbs Cleveland NA_

0 35 70 4

1 134 199 8
```

### Flattening a three-way table

```
dm_cat %$%
  ftable(sex, residence, statin)
```

|        |                   | statin | 0  | 1   |
|--------|-------------------|--------|----|-----|
| sex    | ${\tt residence}$ |        |    |     |
| Female | Suburbs           |        | 42 | 160 |
|        | ${\tt Cleveland}$ |        | 87 | 245 |
| Male   | Suburbs           |        | 35 | 134 |
|        | ${\tt Cleveland}$ |        | 70 | 199 |

 Note that ftable() excludes the missing residence values by default.

## Reminder of Today's Agenda

- Developing Four Models for sbp using dbp (and insurance)
- 2 Fitting a Bayesian Linear Model with default priors (m2)
- Including Insurance without (m3) and with (m4) interaction with dbp in linear models
- Visualizing Categorical Data
- Sessing Association in Cross-Tabulations